

Homework 1

1 Exercise 1

The so defined time inversion process

$$B(t; \alpha) = \begin{cases} t^\alpha W(1/t) & t > 0 \\ 0 & t = 0 \end{cases} \quad (1)$$

in order to be a Brownian Motion has to satisfy the following properties:

1. with probability 1, the mapping $t \mapsto W(t)$ is continuous and $W(0) = 0$;
2. if $0 = t_0 < t_1 < \dots < t_N = T$, then the increments

$$W(t_N) - W(t_{N-1}), \dots, W(t_1) - W(t_0)$$

are *independent*; and

3. for all $t > s$ the increment $W(t) - W(s)$ has the *normal* distribution, with $E[W(t) - W(s)] = 0$ and $E[(W(t) - W(s))^2] = t - s$

We start by looking at the property (1). It's clear from the property of the Brownian Motion W that $B(t, \alpha)$ is continuous in $(0, \infty)$. The only point that we need to check the continuity is in fact at $t = 0$.

As the hint suggests, we start by computing the following for an arbitrary $h > 0, t > 0$.

$$\text{Cov}[B(t+h, \alpha), B(t, \alpha)] = \mathbb{E} \left[(t+h)^\alpha t^\alpha W \left(\frac{1}{t} \right) W \left(\frac{1}{t+h} \right) \right] - \mathbb{E}[B(t+h, \alpha)] \mathbb{E}[B(t, \alpha)].$$

We note that $\mathbb{E}[B(t, \alpha)] = t^\alpha \mathbb{E}[W(\frac{1}{t})] = 0$. Thus, the right hand side can be simplified as follows.

$$\begin{aligned} \text{Cov}[B(t+h, \alpha), B(t, \alpha)] &= \mathbb{E} \left[(t+h)^\alpha t^\alpha W \left(\frac{1}{t} \right) W \left(\frac{1}{t+h} \right) \right] \\ &= (t+h)^\alpha t^\alpha \frac{1}{t+h} \\ &= (t+h)^{\alpha-1} t^\alpha \end{aligned}$$

Where in the one to last equality we used the fact that $\mathbb{E}[W_s W_t] = \min(s, t)$. This gives a necessary condition for $B(t, \alpha)$ to be a Brownian Motion, that is $\alpha = 1$. We will proceed in showing the other properties of $B(t, \alpha)$ assuming $\alpha = 1$.

For a mesh $\Pi = t_1 < t_2 < \dots < t_n$ we write the process B_t as follows.

$$\begin{bmatrix} B(t_1) \\ \vdots \\ B(t_n) \end{bmatrix} = A \begin{bmatrix} W\left(\frac{1}{t_1}\right) \\ \vdots \\ W\left(\frac{1}{t_n}\right) \end{bmatrix} \quad (2)$$

Where just by the definition of the process B_t , we have $A = \text{diag}(t_1, \dots, t_n)$. Moreover, we can write the matrix on the right hand side as follows

$$\begin{bmatrix} W\left(\frac{1}{t_1}\right) \\ \vdots \\ W\left(\frac{1}{t_n}\right) \end{bmatrix} = O^n + \begin{bmatrix} \frac{1}{\sqrt{t_1}} \\ \vdots \\ \frac{1}{\sqrt{t_n}} \end{bmatrix} \begin{bmatrix} \mathcal{N}(0, 1) \\ \vdots \\ \mathcal{N}(0, 1) \end{bmatrix}. \quad (3)$$

Upon defining the matrix D^n as follows

$$D_n = \begin{bmatrix} 0 & \dots & 0 \\ & I_{n-1} & \end{bmatrix}. \quad (4)$$

We can then write the matrix of the increments with the ingredients we have prepared so far.

$$B' = \begin{bmatrix} B_{t_1} \\ \vdots \\ B_{t_n} \end{bmatrix} - \begin{bmatrix} 0 \\ B_1 \\ \vdots \\ B_{t_{n-1}} \end{bmatrix} \quad (5)$$

It follows that $B' = AW - D^n AW = (I - D^n)AW$, where $W = \begin{bmatrix} W\left(\frac{1}{t_1}\right) \\ \vdots \\ W\left(\frac{1}{t_n}\right) \end{bmatrix}$. From

this representation we get that B' is a multivariate Gaussian. Since, implying that each marginal is also Gaussian. Moreover, we have that $\mathbb{E}[B_{t_n} - B_{t_{n-1}}] = 0$. Moreover, we have that

$$\begin{aligned} \text{Cov}[B'_i, B'_j] &= \mathbb{E}[B'_i B'_j] \\ &= \mathbb{E}\left[\left(W\left(\frac{1}{t_i}\right) - W\left(\frac{1}{t_{i-1}}\right)\right)\left(W\left(\frac{1}{t_j}\right) - W\left(\frac{1}{t_{j-1}}\right)\right)\right] \\ &= \mathbb{E}\left[W\left(\frac{1}{t_i}\right)W\left(\frac{1}{t_j}\right)\right] - \mathbb{E}\left[W\left(\frac{1}{t_i}\right)W\left(\frac{1}{t_{j-1}}\right)\right] - \mathbb{E}\left[W\left(\frac{1}{t_{i-1}}\right)W\left(\frac{1}{t_j}\right)\right] + \mathbb{E}\left[W\left(\frac{1}{t_{i-1}}\right)W\left(\frac{1}{t_{j-1}}\right)\right] \\ &= 0 \end{aligned}$$

From the previous point, by taking the limit of the mesh \mathcal{P} , we have that for every a_t

and $t \in [0, \infty) \cap \mathbb{Q}$, the following holds true: $P(\{B(t) < a_t, a_t \in \mathbb{R}, t \in [0, \infty) \cap \mathbb{Q}\}) = P(\{B(t) < a_t, a_t \in \mathbb{R}, t \in [0, \infty) \cap \mathbb{Q}\})$. To conclude the proof that B is continuous at $t = 0$ we use the following proposition.

Proposition 1 *Let $(x_n)_{n \in \mathbb{Q}}$ with $x_n \rightarrow 0$ and let $X(x_n) \xrightarrow{d} Y(x_n)$, $Y(x_n) \rightarrow c$ a.s.. Then, $X(x_n) \rightarrow c$ a.s..*

2 Exercise 2

2.1

In this exercise we will prove the following equality:

$$\int_0^T t dW(t) = TW(T) - \int_0^T W(t) dt \quad (6)$$

Look at the left hand side, by taking it's forward Euler we obtain

$$I_1 = \int_0^T t dW(t) = \sum_{n=0}^{N-1} t_n (W(t_{n+1}) - W(t_n)). \quad (7)$$

Applying the hint, i.e. using the Abel's summation by parts we get

$$\begin{aligned} \sum_{n=0}^{N-1} t_n (W(t_{n+1}) - W(t_n)) &= t_N W(t_N) - t_0 W(t_0) - \sum_{k=1}^{N-1} W(t_k) (t_k - t_{k-1}) \\ &= TW(T) - \sum_{k=1}^{N-1} W(t_k) (t_k - t_{k-1}). \end{aligned}$$

What is left to prove is the convergence in L_2 of the right hand side, i.e. the following equation.

$$\lim_{n \rightarrow \infty} \mathbb{E} \left[\left(\int_0^T W(t) dt - \sum_{k=1}^{N-1} W(t_k) (t_k - t_{k-1}) \right)^2 \right] = 0. \quad (8)$$

Let's fix $t_n - t_{n-1} = \Delta t$ for every $n \geq 0$. Using the linearity of the integral, we can rewrite it as follows

$$\begin{aligned} \mathbb{E} \left[\left(\sum_{n=0}^{N-1} \int_{t_{n-1}}^{t_n} W(t) dt - \sum_{k=1}^{N-1} W(t_k) (t_k - t_{k-1}) \right)^2 \right] &= \mathbb{E} \left[\left(\sum_{n=0}^{N-1} \int_{t_{n-1}}^{t_n} W(t) - W(t_n) dt \right)^2 \right] \\ &= \sum_{n=0}^{N-1} \mathbb{E} \left[\int_{t_{n-1}}^{t_n} W(t) - W(t_n) dt \right]^2 \\ &= \sum_{n=0}^{N-1} \mathbb{E} [e_n^2]. \end{aligned}$$

Where e_n is n-th error term. Moreover, in the last equation we dropped the cross terms since $\mathbb{E}[e_i e_j] = 0$ by the properties of the Brownian motion.

Next, we compute $\mathbb{E}[e_n^2]$ for every $n \geq 0$.

$$\begin{aligned} \mathbb{E}[e_n^2] &= \mathbb{E} \left[\left(\int_{t_n}^{t_{n+1}} W(t) - W(t_n) dt \right)^2 \right] \\ &= \mathbb{E} \left[\int_{t_n}^{t_{n+1}} \int_{t_n}^{t_{n+1}} (W(t) - W(t_n)) (W(s) - W(t_n)) dt ds \right] \\ &= \mathbb{E} \left[\int_{t_n}^{t_{n+1}} \int_{t_n}^{t_{n+1}} (W(t) - W(t_n)) (W(s) - W(t_n)) dt ds \right]. \end{aligned}$$

Where we used twice Fubini's Theorem. We also use Fubini for the following equality.

$$\begin{aligned} \mathbb{E} \left[\int_{t_n}^{t_{n+1}} \int_{t_n}^{t_{n+1}} (W(t) - W(t_n)) (W(s) - W(t_n)) dt ds \right] &= \int_{t_n}^{t_{n+1}} \int_{t_n}^{t_{n+1}} \mathbb{E} [(W(t) - W(t_n)) (W(s) - W(t_n))] dt ds \\ &= \int_{t_n}^{t_{n+1}} \int_{t_n}^{t_{n+1}} \min(t, s) - t_n dt ds \\ &= \frac{1}{3} (\Delta t)^3. \end{aligned}$$

Finally, by summing all the values, we get

$$\sum_{n=0}^T \mathbb{E}[e_n] = \sum_{n=0}^t \frac{1}{3} (\Delta t)^3 = \frac{T^3}{n^2}. \quad (9)$$

It follows that by letting n go to infinity, the Forward Euler converges in L_2 .

2.2

As in the hint, we now prove the following

$$\sum_{n=0}^{N-1} W(t_n)(W(t_{n+1}) - W(t_n)) = \sum_{n=0}^{N-1} \frac{W(t_{n+1})^2 - W(t_n)^2}{2} - \frac{(W(t_{n+1}) - W(t_n))^2}{2} \quad (10)$$

We start by looking at the right hand side.

$$\begin{aligned} \sum_{n=0}^{N-1} \frac{W(t_{n+1})^2 - W(t_n)^2}{2} - \frac{(W(t_{n+1}) - W(t_n))^2}{2} &= \\ &= \sum_{n=0}^{N-1} \frac{W(t_{n+1})^2 - W(t_n)^2}{2} - \frac{W(t_{n+1})^2 - 2W_{t_{n+1}}W(t_n) + W_{t_n}^2}{2} \\ &= \sum_{n=0}^{N-1} \frac{2W_{t_n}W_{t_{n+1}} - 2W(t_n)}{2} = \sum_{n=0}^{N-1} W(t_n)(W(t_{n+1}) - W(t_n)) \end{aligned}$$

Note that the

$$\sum_{n=0}^{N-1} \frac{W(t_{n+1})^2 - W(t_n)^2}{2} = \frac{W(T)}{2} \quad (11)$$

as it is a telescopic sum. To finalize the proof, we need to show that the second part converges to $\frac{T}{2}$ or in other words that

$$\mathbb{E} \left[\left(\sum_{n=0}^{N-1} [W(t_{n+1}) - W(t_n)]^2 - T \right)^2 \right] \rightarrow 0. \quad (12)$$

We define the summation $S_N := \sum_{n=0}^{N-1} [W(t_{n+1}) - W(t_n)]^2$. It follows that

$$\mathbb{E} [(S_N - T)^2] = \mathbb{E} [S_N^2] - T^2 - 2T\mathbb{E} [S_N]. \quad (13)$$

We focus now on $\mathbb{E} [S_N]$ as follows.

$$\begin{aligned} \mathbb{E} [S_N] &= \sum_{n=0}^{N-1} \mathbb{E} [W_{t_{n+1}}^2 + W_{t_n}^2 - 2W_{t_{n+1}}W_{t_n}] \\ &= \sum_{n=0}^{N-1} t_{n+1} + t_n - 2t_n = \sum_{n=0}^{N-1} t_{n+1} - t_n = T. \end{aligned}$$

Where we used the fact that $\mathbb{E} [W_s W_t] = \min(s, t)$. It follows that $\mathbb{E} [(S_N - T)^2] = \mathbb{E} [S_N^2] - T^2 = \mathbb{E} [S_N^2] - \mathbb{E} [S_N]^2 = \text{Var} [S_N]$.

The problem reduces then to compute $\mathbb{E}[S_N^2]$.

$$\mathbb{E}[S_N^2] = \sum_n^{N-1} \sum_m^{N_1} \mathbb{E}[\Delta W_n^2 \Delta W_m^2]$$

For $m = n$ we get $\mathbb{E}[\Delta W_m^4] = 3\Delta t^2$ by using the forth moment of the standard normal. On the other hand, for $n \neq m$ we get the following

$$\mathbb{E}[\Delta W_m^w \Delta W_n^2] = \mathbb{E}[\Delta W_m^2] \mathbb{E}[\Delta W_n^w] = (\Delta t)^2. \quad (14)$$

Putting everything together we obtain

$$\begin{aligned} \mathbb{E}[S_N^2] &= 3N(\Delta t)^2 + N(N-1)(\Delta t)^2 \\ &= \frac{2T^2}{N} - T^2N \\ \implies \mathbb{E}[(S_N - T)^2] &= \text{Var}[S_N] = \frac{2T^2}{N} \end{aligned}$$

Which clearly converges to 0 as $n \rightarrow \infty$.

3 Exercise 3

3.1 Mean and Variance

Given the stochastic process

$$X(t) = x_\infty + e^{-at}(x_0 - x_\infty) + b \int_0^t e^{-a(t-s)} dW(s). \quad (15)$$

We compute its expectation and variance as follows.

$$\mathbb{E}[X_t] = x_\infty + e^{-at}(x_0 - x_\infty) + b\mathbb{E}[(f(t) \cdot W_s)_t]$$

Using the fact that the ito integral of an adapted (since it is deterministic) process with respect to the Brownian Motion has 0 expectation, it further simplifies to.

$$\mathbb{E}[X_t] = x_\infty + e^{-at}(x_0 - x_\infty). \quad (16)$$

We now take its limit for $t \rightarrow \infty$ we obtain

$$\begin{aligned}\lim_{t \rightarrow \infty} \mathbb{E}[X_t] &= \lim_{t \rightarrow \infty} x_\infty + e^{-at}(x_0 - x_\infty) \\ &= x_\infty.\end{aligned}$$

We now proceed onto computing the variance of the process and its limit as $t \rightarrow \infty$.

$$\begin{aligned}\text{Var}[X_t] &= \mathbb{E}[(X_t - \mathbb{E}[X_t])^2] = \mathbb{E}[X_t^2] - \mathbb{E}[X_t]^2 \\ &= b^2 \mathbb{E}\left[\left(\int_0^t e^{-a(t-s)} dW_s\right)^2\right].\end{aligned}$$

In the last equality we simply used the fact that the Ito integral has again zero expectation, therefore it annihilates all the terms that are multiplied to as they are bounded. We shall now use the Ito isometry.

$$\begin{aligned}b^2 \mathbb{E}\left[\left(\int_0^t e^{-a(t-s)} dW_s\right)^2\right] &= b^2 \mathbb{E}\left[\int_0^t e^{-2a(t-s)} ds\right] \\ &= b^2 e^{-2at} \mathbb{E}\left[\int_0^t e^{-2as} ds\right] \\ &= \frac{b^2 e^{-2at}}{2a} (e^{2at} + 1) \\ &= \frac{b^2}{2a} (1 - e^{-2at}).\end{aligned}$$

Thus

$$\lim_{t \rightarrow \infty} \text{Var}[X_t] = \lim_{t \rightarrow \infty} \frac{b^2}{2a} (1 - e^{-2at}) = \frac{b^2}{2a}. \quad (17)$$

3.2 Covariance

For a more general settings, we compute the variance between X_t and X_s where w.log. we assume $t > s$.

$$\begin{aligned}\text{Cov}[X_s, X_t] &= b^2 \text{Cov}\left[\int_0^s e^{a(t-u)} dW_u, \int_0^t e^{-a(t-u)} dW_u\right] \\ &= \mathbb{E}\left[\int_0^s e^{-a(t-u)} dW_u \int_0^t e^{-a(t-u)} dW_u\right] \\ &= b^2 e^{-a(s+t)} \mathbb{E}\left[\int_0^s e^{au} dW_u \int_0^t e^{au} dW_u\right] \\ &= b^2 e^{-a(s+t)} \mathbb{E}\left[\int_0^s e^{au} dW_u \left(\int_0^s e^{au} dW_u + \int_s^t e^{au} dW_u\right)\right]\end{aligned}$$

By using the independence of the increments of Ito integral, we have the expectation of the product as the product of the expectation as follows.

$$\begin{aligned}
b^2 e^{-a(s+t)} \mathbb{E} \left[\int_0^s e^{au} dW_u \left(\int_0^s e^{au} dW_u + \int_s^t e^{au} dW_u \right) \right] &= \\
&= b^2 e^{-a(s+t)} \mathbb{E} \left[\left(\int_0^s e^{au} dW_u \right)^2 + \int_s^t e^{au} dW_u \int_0^s e^{au} dW_u \right] \\
&= b^2 e^{-a(s+t)} \left(\mathbb{E} \left[\left(\int_0^s e^{au} dW_u \right)^2 \right] + \mathbb{E} \left[\int_s^t e^{au} dW_u \int_0^s e^{au} dW_u \right] \right) \\
&= b^2 e^{-a(s+t)} \mathbb{E} \left[\left(\int_0^s e^{au} dW_u \right)^2 \right]
\end{aligned}$$

By Ito isometry again

$$\begin{aligned}
b^2 e^{-a(s+t)} \mathbb{E} \left[\left(\int_0^s e^{au} dW_u \right)^2 \right] &= b^2 e^{-a(s+t)} \mathbb{E} \left[\int_0^s e^{2au} du \right] \\
&= b^2 e^{-a(s+t)} \frac{1}{2a} (e^{2at} - 1) \\
&= \frac{b^2}{2a} (e^{-a(t-s)} - e^{-a(s+t)})
\end{aligned}$$

3.3 Interpretation and numerics

The first thing to notice is the limit of the expectation. This is x_∞ , therefore, the process will converge in mean to such value.

To corroborate this, by running numerical experiments, we can see that indeed it converges in mean to x_∞ . In the plot 1 we plotted only 10 sample paths, however, for the mean we use a sample size of 1000 to make it converge.

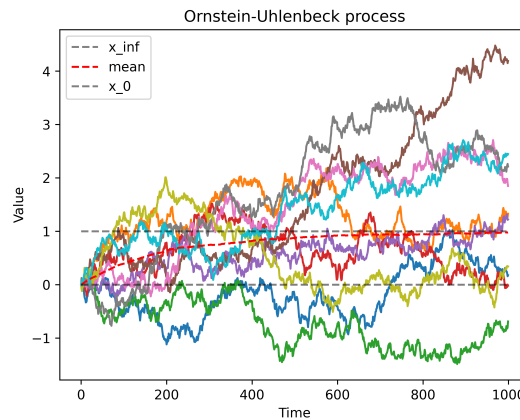


Figure 1: Plot showing some sample paths from the OU process, their mean and the x_∞ .