

ANOVA approach to hypothesis testing

The F Test

- In this section, we will talk about a different strategy to testing $H_0: \beta_1 = 0$

In simple linear regression, this results in the *exact same-value* as the test that uses the t-statistic.

However, this strategy is more applicable to general linear hypotheses that we'll discuss in multiple linear regression.

- Testing $H_0: \beta_1 = 0$

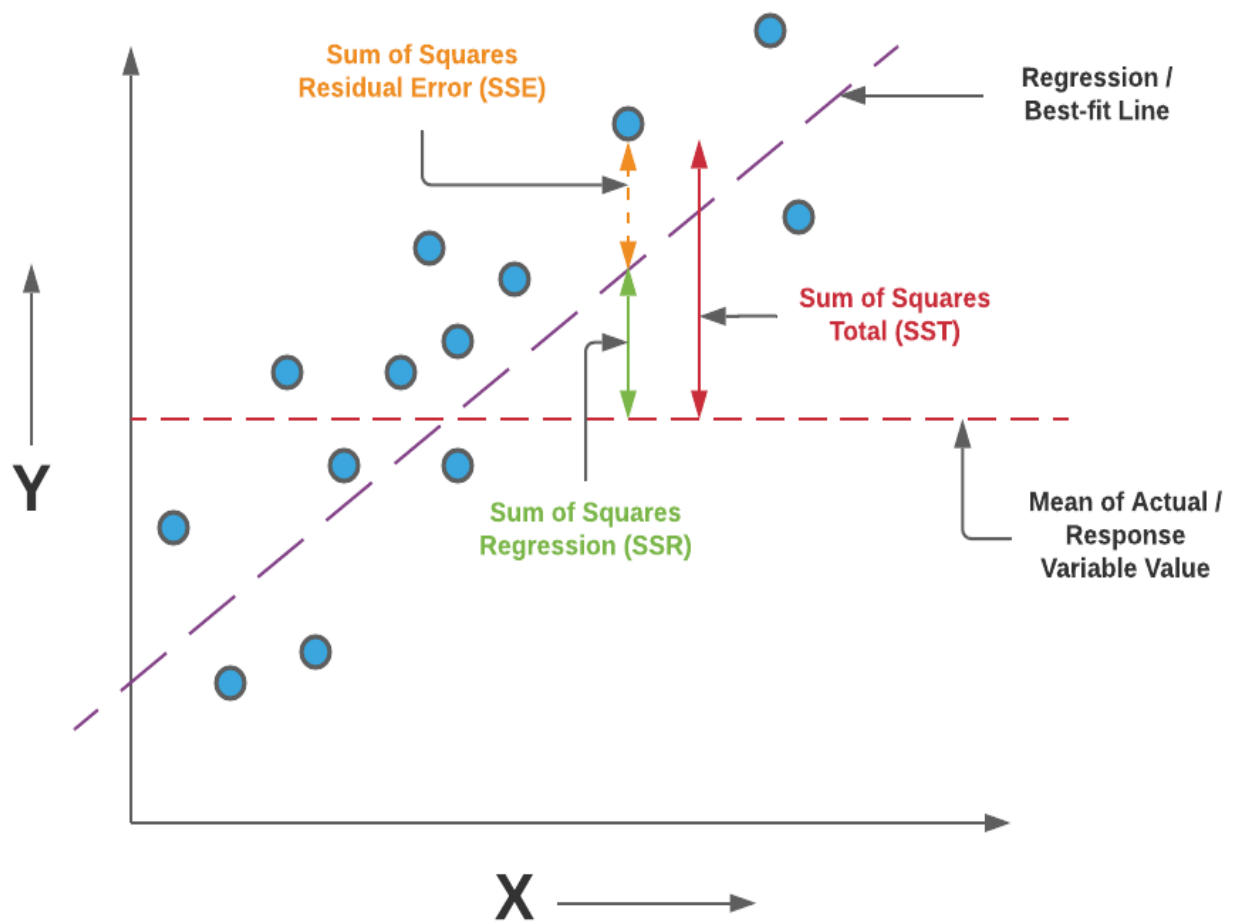
is really a comparison between the two models:

$$\begin{aligned}H_0: Y_i &= \beta_0 + \epsilon_i \\ H_A: Y_i &= \beta_0 + \beta_1 X_i + \epsilon_i\end{aligned}$$

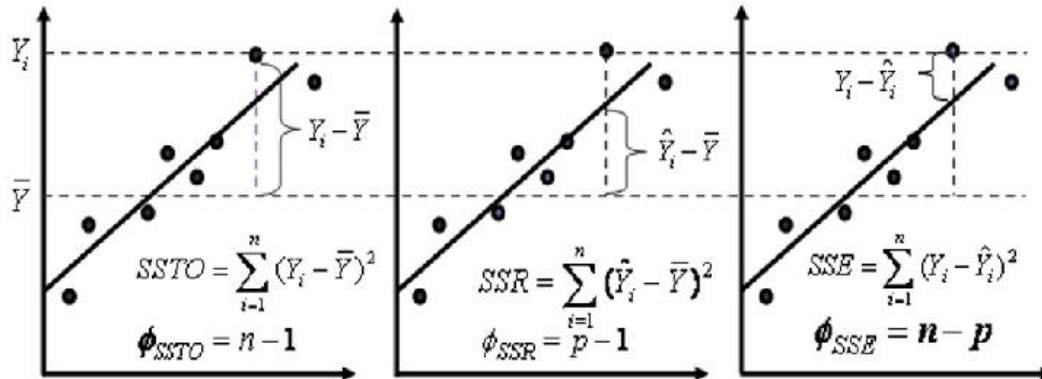
- We will call the first model the **reduced** model and the second model the **full** model. This is because the reduced model is a subset of the full model (you get the reduced from the full by setting $\beta_1 = 0$).
- Our strategy will be to compare the residuals under H_0 and H_A . If H_A were true, we would expect the those residuals to be much smaller than the residuals under H_0 (because the line fits a lot better).

If H_0 were true, then we would expect the residuals under H_A to only be a little bit smaller than those under H_0
- We fit H_A by the method of least squares, obtaining the OLS estimates and the corresponding residuals.
- We fit H_0 also by least squares. It turns out that under H_0 , the OLS estimate is just \bar{Y} . We measure how small the residuals are by the sum of squared residuals.

PICTORIAL and PLOT representation SSTO, SSE, and SSR



Example:



X	Y	\hat{Y}_i (hat) Predicted Value	$Y_i - \hat{Y}_i$ (hat) Residual	$(Y_i - \bar{Y})^2$	$(Y_i - \hat{Y}_i)^2$	$(\hat{Y}_i - \bar{Y})^2$
34	5	4.1505	.8496	25	.72165	34.21665
108	17	14.9693	2.0307	49	4.12374	24.69394
64	11	8.5365	2.4635	1	6.06883	2.14183
88	8	12.0453	-4.0453	4	16.3644	4.18325
99	14	13.6535	.3465	16	.12006	13.34906
51	5	6.6359	-1.6359	25	2.6762	11.31717
TOTALS				120	30.07488	89.9019

$\bar{Y} = 10$

The following Regression Model for the data is found by using R coding: `lm(y ~ x)`

$\hat{Y} = -0.8203 + 0.1462X$

The following ANOVA table is generated by using R coding: `anova(lm(y ~ x))`

Analysis of Variance Table

Response: y

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
x	1	89.925	89.925	11.96	0.02586 *

Residuals 4 30.075 7.519

Signif. codes:

0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Now let us conduct the F test to determine if $H_0: Y_i = \beta_0 + \epsilon_i$ should be rejected.

If the F_{value} is greater than the F_{critical} , we reject H_0 and conclude H_a

If the F_{value} is less than or equal to F_{critical} , we fail to reject H_0

The F_{value} is found in the ANOVA table.

The F_{critical} value is found in the F distribution table, using the construction $F(1 - \alpha; 1, n - 2)$ or by using

For our case, F_{value} is 11.96 (found in the ANOVA output table).

The F_{critical} value can be found as follows;

Let $\alpha = .05$, and our designated degrees of freedom are 1 and 6.

We therefore require $F(.95, 1, 4)$, now going to the F table in the back of the we get 7.71

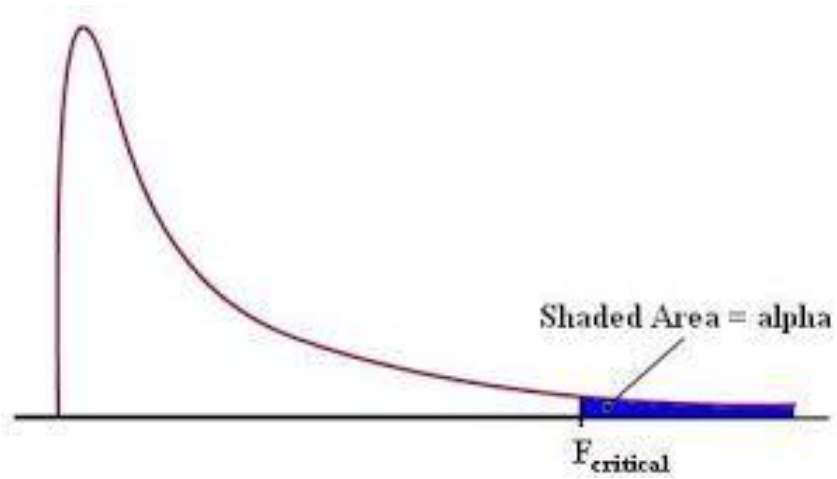
We can also use R code as follows to get the F_{critical} value.

```
qf(p=.05, df1=1, df2=4, lower.tail=FALSE)
```

```
7.708647
```

Since F_{value} (11.96) is greater than F_{critical} (7.71), we will reject the null Hypothesis that $B_1 = 0$ and conclude that B_1 does not = 0. Moreover, a linear relationship does exist between Y and X.

Graphical Interpretation: F Distribution Curve



If F_{value} is greater than (to the right of F_{critical} you are to reject H_0

If F_{value} is less than (to the left of F_{critical} or equal to F_{critical} , you fail to reject H_0

F Distribution Table (Percentiles)

		1	2	3	4	5	6	7	8	9	10
0.95	1	161.5	199.5	215.7	224.6	230.2	234.0	236.8	238.9	240.5	241.9
0.975		647.8	799.5	864.2	899.6	921.8	937.1	948.2	956.7	963.3	968.7
0.99		4052.2	4999.5	5403.4	5624.6	5763.7	5859.0	5928.4	5981.07	6022.4	6055.9
0.95	2	18.51	19.00	19.16	19.25	19.30	19.33	19.35	19.37	19.38	19.40
0.975		38.51	39.00	39.17	39.25	39.30	39.33	39.36	39.37	39.39	39.40
0.99		98.50	99.00	99.17	99.25	99.30	99.33	99.36	99.37	99.39	99.40
0.95	3	10.13	9.55	9.28	9.12	9.01	8.94	8.89	8.85	8.81	8.79
0.975		17.44	16.04	15.44	15.10	14.88	14.73	14.62	14.54	14.47	14.42
0.99		34.12	30.82	29.46	28.71	28.24	27.91	27.67	27.49	27.35	27.23
0.95	4	7.71	6.94	6.59	6.39	6.26	6.16	6.09	6.04	6.00	5.96
0.975		12.22	10.65	9.98	9.60	9.36	9.20	9.07	8.98	8.90	8.84
0.99		21.20	18.00	16.69	15.98	15.52	15.21	14.98	14.80	14.66	14.55
0.95	5	6.61	5.79	5.41	5.19	5.05	4.95	4.88	4.82	4.77	4.74
0.975		10.01	8.43	7.76	7.39	7.15	6.98	6.85	6.76	6.68	6.62
0.99		16.26	13.27	12.06	11.39	10.97	10.67	10.46	10.29	10.16	10.05
0.95	6	5.99	5.14	4.76	4.53	4.39	4.28	4.21	4.15	4.10	4.06
0.975		8.81	7.26	6.60	6.23	5.99	5.82	5.70	5.60	5.52	5.46
0.99		13.75	10.92	9.78	9.15	8.75	8.47	8.26	8.10	7.98	7.87
0.95	7	5.59	4.74	4.35	4.12	3.97	3.87	3.79	3.73	3.68	3.64
0.975		8.07	6.54	5.89	5.52	5.29	5.12	4.99	4.90	4.82	4.76
0.99		12.25	9.55	8.45	7.85	7.46	7.19	6.99	6.84	6.72	6.62
0.95	8	5.32	4.46	4.07	3.84	3.69	3.58	3.50	3.44	3.39	3.35
0.975		7.57	6.06	5.42	5.05	4.82	4.65	4.53	4.43	4.36	4.30
0.99		11.26	8.65	7.59	7.01	6.63	6.37	6.18	6.03	5.91	5.81
0.95	9	5.12	4.26	3.86	3.63	3.48	3.37	3.29	3.23	3.18	3.14
0.975		7.21	5.71	5.08	4.72	4.48	4.32	4.20	4.10	4.03	3.96
0.99		10.56	8.02	6.99	6.42	6.06	5.80	5.61	5.47	5.35	5.26
0.95	10	4.96	4.10	3.71	3.48	3.33	3.22	3.14	3.07	3.02	2.98
0.975		6.94	5.46	4.83	4.47	4.24	4.07	3.95	3.85	3.78	3.72
0.99		10.04	7.56	6.55	5.99	5.64	5.39	5.20	5.06	4.94	4.85
0.95	11	4.84	3.98	3.59	3.36	3.20	3.09	3.01	2.95	2.90	2.85
0.975		6.72	5.26	4.63	4.28	4.04	3.88	3.76	3.66	3.59	3.53
0.99		9.65	7.21	6.22	5.67	5.32	5.07	4.89	4.74	4.63	4.54
0.95	12	4.75	3.89	3.49	3.26	3.11	3.00	2.91	2.85	2.80	2.75
0.975		6.55	5.10	4.47	4.12	3.89	3.73	3.61	3.51	3.44	3.37
0.99		9.33	6.93	5.95	5.41	5.06	4.82	4.64	4.50	4.39	4.30
0.95	13	4.67	3.81	3.41	3.18	3.03	2.92	2.83	2.77	2.71	2.67
0.975		6.41	4.97	4.35	4.00	3.77	3.60	3.48	3.39	3.31	3.25
0.99		9.07	6.70	5.74	5.21	4.86	4.62	4.44	4.30	4.19	4.10
0.95	14	4.60	3.74	3.34	3.11	2.96	2.85	2.76	2.70	2.65	2.60
0.975		6.30	4.86	4.24	3.89	3.66	3.50	3.38	3.29	3.21	3.15
0.99		8.86	6.51	5.56	5.04	4.69	4.46	4.28	4.14	4.03	3.94
0.95	15	4.54	3.68	3.29	3.06	2.90	2.79	2.71	2.64	2.59	2.54
0.975		6.20	4.77	4.15	3.80	3.58	3.41	3.29	3.20	3.12	3.06
0.99		8.68	6.36	5.42	4.89	4.56	4.32	4.14	4.00	3.89	3.80