

# Residuals and residual variation

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#### Residuals

- · Model  $Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$  where  $\epsilon_i \sim N(0, \sigma^2)$ .
- Observed outcome i is  $Y_i$  at predictor value  $X_i$
- Predicted outcome i is  $\hat{Y}_i$  at predictor valuve  $X_i$  is

$$\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_i$$

· Residual, the between the observed and predicted outcome

$$e_i = Y_i - \hat{Y}_i$$

- The vertical distance between the observed data point and the regression line
- Least squares minimizes  $\sum_{i=1}^n\,e_i^2$
- · The  $e_i$  can be thought of as estimates of the  $\epsilon_i$ .

### Properties of the residuals

- $\cdot E[e_i] = 0.$
- · If an intercept is included,  $\sum_{i=1}^{n} e_i = 0$
- · If a regressor variable,  $X_i$ , is included in the model  $\sum_{i=1}^{n} e_i X_i = 0$ .
- · Residuals are useful for investigating poor model fit.
- · Positive residuals are above the line, negative residuals are below.
- · Residuals can be thought of as the outcome (Y) with the linear association of the predictor (X) removed.
- · One differentiates residual variation (variation after removing the predictor) from systematic variation (variation explained by the regression model).
- · Residual plots highlight poor model fit.

#### Code

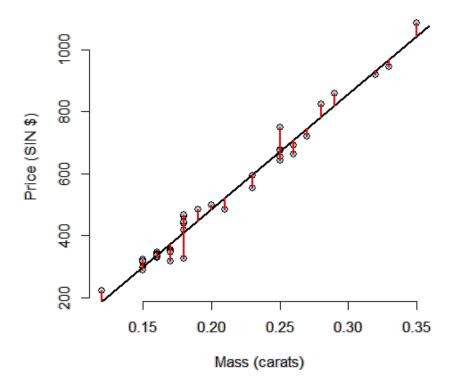
```
data(diamond)
y <- diamond$price; x <- diamond$carat; n <- length(y)
fit <- lm(y ~ x)
e <- resid(fit)
yhat <- predict(fit)
max(abs(e -(y - yhat)))</pre>
```

```
[1] 9.486e-13
```

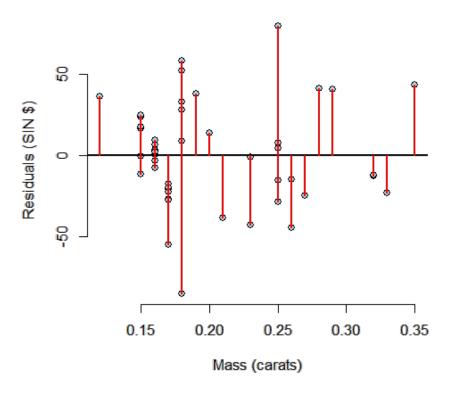
```
\max(abs(e - (y - coef(fit)[1] - coef(fit)[2] * x)))
```

```
[1] 9.486e-13
```

## Residuals are the signed length of the red lines

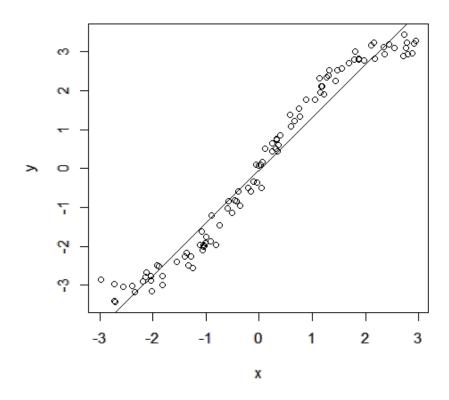


### Residuals versus X



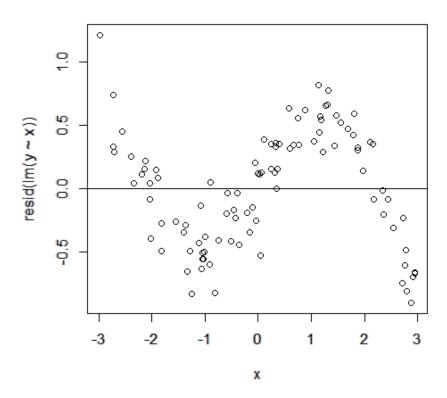
#### Non-linear data

```
x <- runif(100, -3, 3); y <- x + sin(x) + rnorm(100, sd = .2); plot(x, y); abline(lm(y ~ x))
```



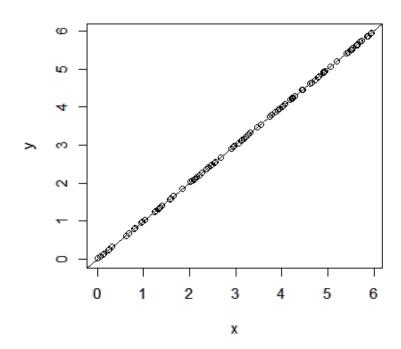
```
plot(x, resid(lm(y \sim x)));

abline(h = 0)
```



## Heteroskedasticity

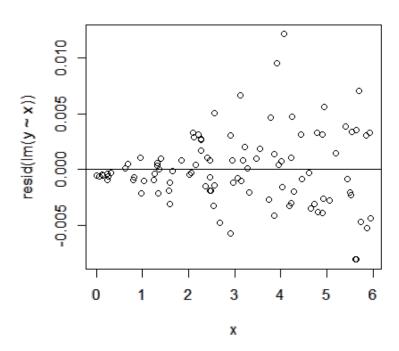
```
x \leftarrow runif(100, 0, 6); y \leftarrow x + rnorm(100, mean = 0, sd = .001 * x); plot(x, y); abline(lm(y ~ x))
```



## Getting rid of the blank space can be helpful

```
plot(x, resid(lm(y \sim x)));

abline(h = 0)
```



## **Estimating residual variation**

- · Model  $Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$  where  $\epsilon_i \sim N(0, \sigma^2)$ .
- · The ML estimate of  $\sigma^2$  is  $\frac{1}{n}\sum_{i=1}^n\,e_i^2$ , the average squared residual.
- · Most people use

$$\hat{\sigma}^2 = \frac{1}{n-2} \sum_{i=1}^n e_i^2$$
.

• The n – 2 instead of n is so that  $E[\hat{\sigma}^2] = \sigma^2$ 

## Diamond example

```
[1] 31.84
```

```
sqrt(sum(resid(fit)^2) / (n - 2))
```

```
[1] 31.84
```

#### **Summarizing variation**

$$\begin{split} \sum_{i=1}^{n} (Y_i - \bar{Y})^2 &= \sum_{i=1}^{n} (Y_i - \hat{Y}_i + \hat{Y}_i - \bar{Y})^2 \\ &= \sum_{i=1}^{n} (Y_i - \hat{Y}_i)^2 + 2 \sum_{i=1}^{n} (Y_i - \hat{Y}_i)(\hat{Y}_i - \bar{Y}) + \sum_{i=1}^{n} (\hat{Y}_i - \bar{Y})^2 \end{split}$$

#### Scratch work

$$\begin{split} &(Y_{i} - \hat{Y}_{i}) = \{Y_{i} - (\bar{Y} - \hat{\beta}_{1}\bar{X}) - \hat{\beta}_{1}X_{i}\} = (Y_{i} - \bar{Y}) - \hat{\beta}_{1}(X_{i} - \bar{X}) \\ &(\hat{Y}_{i} - \bar{Y}) = (\bar{Y} - \hat{\beta}_{1}\bar{X} - \hat{\beta}_{1}X_{i} - \bar{Y}) = \hat{\beta}_{1}(X_{i} - \bar{X}) \\ &\sum_{i=1}^{n} (Y_{i} - \hat{Y}_{i})(\hat{Y}_{i} - \bar{Y}) = \sum_{i=1}^{n} \{(Y_{i} - \bar{Y}) - \hat{\beta}_{1}(X_{i} - \bar{X})\} \{\hat{\beta}_{1}(X_{i} - \bar{X})\} \\ &= \hat{\beta}_{1} \sum_{i=1}^{n} (Y_{i} - \bar{Y})(X_{i} - \bar{X}) - \hat{\beta}_{1}^{2} \sum_{i=1}^{n} (X_{i} - \bar{X})^{2} \\ &= \hat{\beta}_{1}^{2} \sum_{i=1}^{n} (X_{i} - \bar{X})^{2} - \hat{\beta}_{1}^{2} \sum_{i=1}^{n} (X_{i} - \bar{X})^{2} = 0 \end{split}$$

#### **Summarizing variation**

$$\sum_{i=1}^{n} (Y_i - \bar{Y})^2 = \sum_{i=1}^{n} (Y_i - \hat{Y}_i)^2 + \sum_{i=1}^{n} (\hat{Y}_i - \bar{Y})^2$$

Or

Total Variation = Residual Variation + Regression Variation

Define the percent of total varation described by the model as

$$R^2 = \frac{\sum_{i=1}^{n} (\hat{Y}_i - \bar{Y})^2}{\sum_{i=1}^{n} (Y_i - \bar{Y})^2} = 1 - \frac{\sum_{i=1}^{n} (Y_i - \hat{Y}_i)^2}{\sum_{i=1}^{n} (Y_i - \bar{Y})^2}$$

## Relation between $\mathbb{R}^2$ and $\mathbb{R}$ (the corrrelation)

Recall that  $(\hat{Y}_i - \bar{Y}) = \hat{\beta}_1(X_i - \bar{X})$  so that

$$R^{2} = \frac{\sum_{i=1}^{n} (\hat{Y}_{i} - \bar{Y})^{2}}{\sum_{i=1}^{n} (Y_{i} - \bar{Y})^{2}} = \hat{\beta}_{1}^{2} \frac{\sum_{i=1}^{n} (X_{i} - \bar{X})}{\sum_{i=1}^{n} (Y_{i} - \bar{Y})^{2}} = Cor(Y, X)^{2}$$

Since, recall,

$$\hat{\beta}_1 = Cor(Y, X) \frac{Sd(Y)}{Sd(X)}$$

So,  $R^2$  is literally r squared.

#### Some facts about $\mathbb{R}^2$

- $\cdot$  R<sup>2</sup> is the percentage of variation explained by the regression model.
- $0 \le R^2 \le 1$
- $\cdot$  R<sup>2</sup> is the sample correlation squared.
- $\cdot$  R<sup>2</sup> can be a misleading summary of model fit.
  - Deleting data can inflate R<sup>2</sup>.
  - (For later.) Adding terms to a regression model always increases  $R^2$ .
- · Do example (anscombe) to see the following data.
  - Basically same mean and variance of X and Y.
  - Identical correlations (hence same  $R^2$  ).
  - Same linear regression relationship.

### data(anscombe);example(anscombe)

#### Anscombe's 4 Regression data sets

