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- A rather different approach to sampling.
- Part of the Markov Chain Monte Carlo (MCMC) family of algorithms.
- Don't generate each sample from scratch.
- Generate samples by making a random change to the previous sample.

- Gibbs sampling for Bayesian networks starts with an arbitrary state.
- So pick state with evidence variables fixed at observed values. (If we know Cloudy = true, we pick that.)
- Generate next state by randomly sampling from a non-evidence variable.
- Do this sampling conditional on the current values of the Markov blanket.
- "The algorithm therefore wanders randomly around the state space ... flipping one variable at a time, but keeping the evidence variables fixed".

• Consider the query:

$$\textbf{P}(Cloudy|Sprinker=true,WetGrass=true)$$

- The evidence variables are fixed to their observed values.
- The non-evidence variables are initialised randomly.

$$Cloudy = true$$

$$Rain = false$$

State is thus:

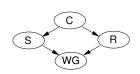
$$[Cloudy = true, Sprinkler = true, Rain = false, WetGrass = true].$$

- First we sample Cloudy given the current state of its Markov blanket.
- Markov blanket is *Sprinkler* and *Rain*.
- So, sample from:

$$\mathbf{P}(Cloudy|Sprinkler = true, Rain = false)$$

■ Suppose we get Cloudy = false, then new state is:

$$[Cloudy = false, Sprinkler = true, \\ Rain = false, WetGrass = true].$$



- Next we sample Rain given the current state of its Markov blanket.
- \blacksquare Markov blanket is Cloudy, Sprinkler and WetGrass.
- So, sample from:

$$\mathbf{P}(Rain|Cloudy = false, Sprinkler = true, WetGrass = false)$$

■ Suppose we get Rain = true, then new state is:

$$[Cloudy = false, Sprinkler = true, \\ Rain = true, WetGrass = true].$$

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Each state visited during this process contributes to our estimate for:

$$\mathbf{P}(Cloudy|Sprinkler=true,WetGrass=true)$$

- Say the process visits 80 states.
- In 20, Cloudy = true
- In 60, Cloudy = false
- Then

$$\mathbf{P}(Cloudy|Sprinkler = true, WetGrass = true) = \alpha \begin{pmatrix} 20\\60 \end{pmatrix} = \begin{pmatrix} 0.25\\0.75 \end{pmatrix}$$

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\mathbf{Z}, the nonevidence variables in bn \mathbf{x}, the current state of the network, initially copied from \mathbf{e} initialize \mathbf{x} with random values for the variables in \mathbf{Z} for j=1 to N do for each Z_i in \mathbf{Z} do sample the value of Z_i in \mathbf{x} from \mathbf{P}(Z_i|mb(Z_i)) given the values of MB(Z_i) in \mathbf{x} \mathbf{N}[x] \leftarrow \mathbf{N}[x] + 1 where x is the value of X in \mathbf{x} return NORMALIZE(\mathbf{N}[X])
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function GIBBS-ASK(X, \mathbf{e} , bn, N) **returns** an estimate of $P(X|\mathbf{e})$ **local variables**: $\mathbf{N}[X]$, a vector of counts over X, initially zero

- All of this begs the question:
 - "How do we sample a variable given the state of its Markov blanket?"
- For a value x of a variable X:

$$\mathbf{P}(X|mb(X)) = \alpha \mathbf{P}(X|parents(X)) \prod_{Y \in Children(X)} \mathbf{P}(Y|parents(Y))$$

where mb(X) is the Markov blanket of X.

• Given P(X|mb(X)), we can sample from it just as we have before.

To summarise

- Bayesian networks exploit conditional independence to create a more compact set of information.
- Reasonably efficient computation for some problems.
- Five approaches to inference in Bayesian networks.
 - Exact: Inference by enumeration.
 - · Approximate: Prior sampling
 - Approximate: Rejection sampling
 - Approximate: Importance sampling/likelihood weighting
 - Approximate: Gibbs sampling
- Can answer a simple query for any BN.