

Maximum Likelihood & Bayes Rule

Topics

- Review Maximum Likelihood Concept
- Consider Prior Beliefs
- Continuous valued Random Variables
- Posterior probability and Maximum Likelihood

Click-Through Rates in Online Advertisizing

Example from Xueri Wang et al

Results

Model:

X = Number of click-throughs in 200 views

$X \sim \text{Binomial}(200, p)$

In 200 views, 25 click-throughs occurred

Let's estimate p using the likelihood approach

Maximum Likelihood

Consider the chance of 25 successes if $p=0.01$

$$P(X = 25 | p=0.01) = C(200, 25) 0.01^{25} 0.99^{200-25} \\ = 7.7e-20$$

Let's consider other possible values for p ,

Maximum Likelihood

If $p=0.05$, then $P(X = 25 | p=0.05) = 1.7e-05$

If $p=0.10$, then $P(X = 25 | p=0.10) = 0.04$

if $p=0.15$, then $P(X = 25 | p=0.10) = 0.05$

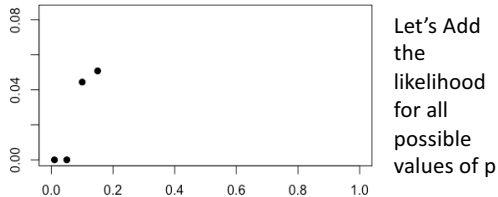
Let's place these values on a plot

Maximum Likelihood

If $p=0.05$, then $P(X = 25 | p=0.05) = 1.7e-05$

If $p=0.10$, then $P(X = 25 | p=0.10) = 0.04$

if $p=0.15$, then $P(X = 25 | p=0.10) = 0.05$

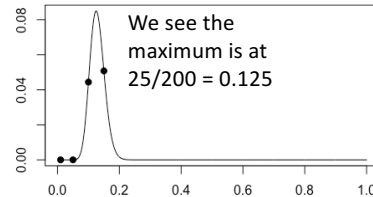


Maximum Likelihood

If $p=0.05$, then $P(X = 25 | p=0.05) = 1.7e-05$

If $p=0.10$, then $P(X = 25 | p=0.10) = 0.04$

if $p=0.15$, then $P(X = 25 | p=0.10) = 0.05$



Likelihood

- These likelihoods can be viewed as a function of p given the data

$$L(p) = C(200, 25) p^{25} (1-p)^{200-25}$$

Find the p that maximizes the likelihood for our data and use it to estimate p .

Likelihood

$$L(p) = C(200, 25) p^{25} (1-p)^{200-25}$$

here $C(200, 25)$ is the factorial

It is often easier to maximize the log of the likelihood function:

$$\log(L(p)) = C(200, 25) + 25\log(p) + (200-25)\log(1-p)$$

We can differentiate the log-likelihood and set to 0 to solve for p to get 0.125 as our estimate

Practice Problem

Geometric(p)

- X = number of failures until first success
- Trials are independent with the same probability of success
- $P(k) = P(k \text{ failures followed by a success})$
 $= p(1-p)^k$ for $k = 0, 1, 2, \dots$

Suppose you observe the geometric n times, and record k_1, \dots, k_n . Which value of p maximizes the likelihood of our data?

Find the Likelihood

$$\begin{aligned}
 P(X_1=k_1, X_2=k_2, \dots, X_n=k_n) \\
 &= P(X_1=k_1)P(X_2=k_2) \dots P(X_n=k_n) \quad \text{independence} \\
 &= p(1-p)^{k_1} \times p(1-p)^{k_2} \times \dots \times p(1-p)^{k_n} \quad \text{geometric}(p) \\
 &= p^n(1-p)^{k_1+k_2+\dots+k_n} \quad \text{The likelihood function}
 \end{aligned}$$

Let's maximize the log likelihood

$$\log(L(p)) = n\log(p) + (k_1+\dots+k_n)\log(1-p)$$

Differentiate wrt p

$$n/p - (k_1+\dots+k_n)/(1-p)$$

Set to 0 and solve for p

$$0 = n/p - (k_1+\dots+k_n)/(1-p)$$

$$0 = (1-p)n - p(k_1+\dots+k_n)$$

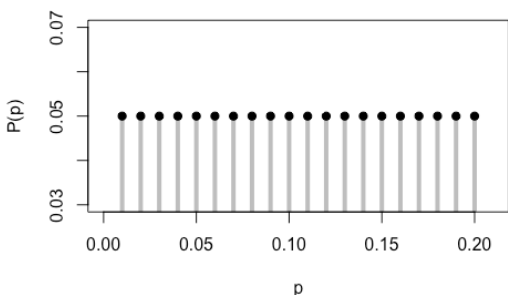
$$\hat{p} = n/(n + k_1+\dots+k_n) = 1/(1 + \text{avg})$$

number of successes / number of trials

Return to our click-through example

- Suppose that you were certain that p was either 0.01, 0.02, ..., 0.20
- You figure it might be any one of these values
- And you think of them as equally likely values for p
- This is similar to the vitamin problem where you reach into a barrel of coins that have different chances of success and pick one to flip

Here's a pmf for the possible values of p



Given the data: 25 click-throughs on 200 visits, how would you update your pmf for p ?

For example, what is

$$P(p = 0.01 \mid \text{data} = 25) ?$$

$$P(p = 0.01 \mid \text{data} = 25)$$

Let's apply the conditional probability rule

$$= P(p = 0.01 \text{ and } \text{data} = 25) / P(\text{data})$$

Let's apply the rule again to the numerator

$$= P(p = 0.01) P(\text{data} = 25 \mid p = 0.01) / P(\text{data})$$

$$= 0.05 C(200, 25) 0.01^{25} 0.99^{200-25} / P(\text{data})$$

Bayes Rule

$$P(A \mid B) = P(A \text{ and } B) / P(B)$$

$$P(B \mid A) = P(A \text{ and } B) / P(A)$$

$$\text{So } P(A \text{ and } B) = P(A)P(B \mid A)$$

Plug this into the numerator above

$$\text{Bayes RULE: } P(A \mid B) = P(A)P(B \mid A) / P(B)$$

Bayes Rule

- We just used Bayes rule

$$P(A \mid B) = P(A)P(B \mid A) / P(B)$$

- $A = \{p = 0.1\}$

- $B = \{\text{data} = 25\}$

$$P(p = 0.01 \mid \text{data} = 25) =$$

$$P(p = 0.01) P(\text{data} = 25 \mid p = 0.01) / P(\text{data} = 25)$$

The denominator is
the average over all
possible values for p .

$P(p \mid \text{data} = 25)$

We can compute $P(p = x \mid \text{data} = 25)$ for all possible values of p .

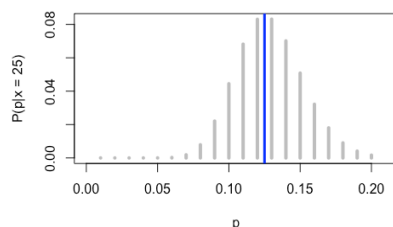
In general, for $x = 0.01, 0.02, \dots, 0.20$,

$$P(p = x \mid \text{data} = 25) =$$

$$= P(p = x) P(\text{data} = 25 \mid p = x) / P(\text{data})$$

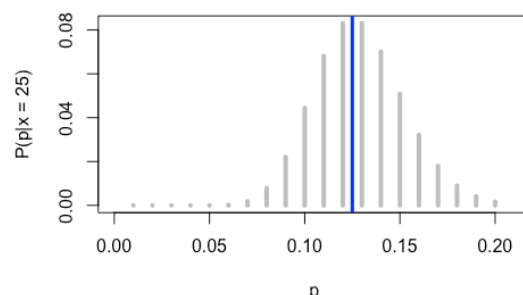
$$= 0.05 C(200, 25) x^{25} (1-x)^{200-25} / P(\text{data})$$

$$P(p = x \mid \text{data} = 25) = C x^{25} (1-x)^{200-25}$$



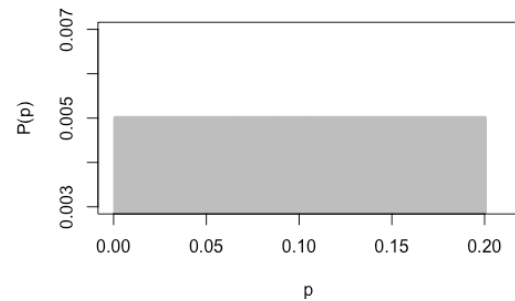
We have updated our probabilities for p given the data.

Posterior probability of p given our data

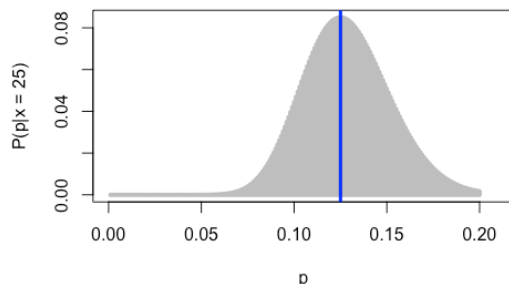


What happens if we change our initial (prior) probabilities on p to be discrete uniform on 0.001, 0.002, ..., 0.200?

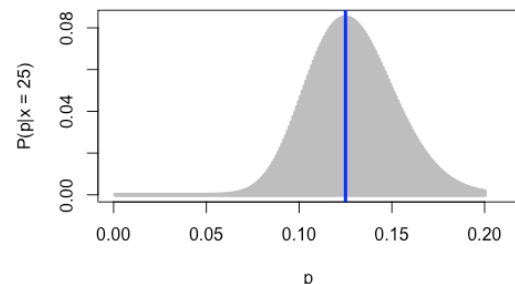
What probability mass function is so dense that we can't distinguish the individual values for p .



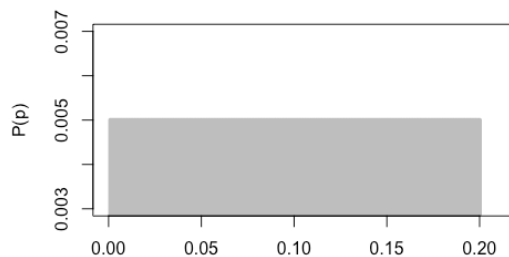
This posterior distribution for p places the highest probability on 0.125, but the small values between 0 and 0.05 still have a chance (a small chance) of being the true p .



We can use this posterior distribution to estimate p (choose the one with the highest chance). We can provide credible regions for p , e.g., interval with a 95% posterior probability.



This pmf is so dense that we wonder, can we just say that p can be equally likely to be any value in $(0, 1)$?



One small problem:
There are uncountably many p
possible values in $(0,1)$

Continuous Random Variables

Continuous Uniform Distribution

- Uniform(0, 1) distribution
- We say the random variable X has a Uniform(0, 1) distribution, if

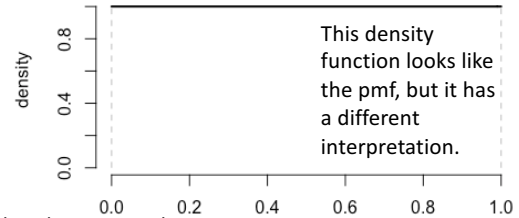
$$P(a < X < b) = (b-a) \text{ for any } a \text{ \& } b \text{ in } (0,1)$$

For example,

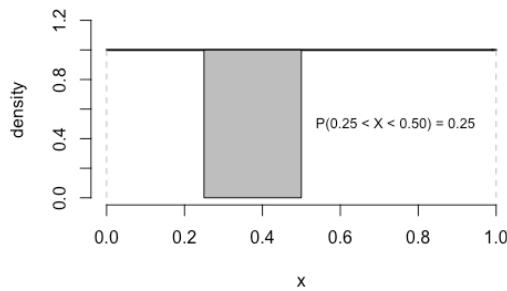
$$\begin{aligned} P(1/4 < X < 1/2) &= P(1/8 < X < 3/8) \\ &= P(3/4 < X < 1) \\ &= 1/4 \end{aligned}$$

We compute probabilities on intervals, rather than exact values

Probability Density Function



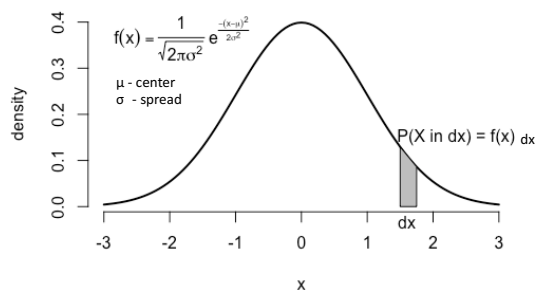
Like a histogram, the y-axis is a density scale. The area under the curve corresponds to chance.



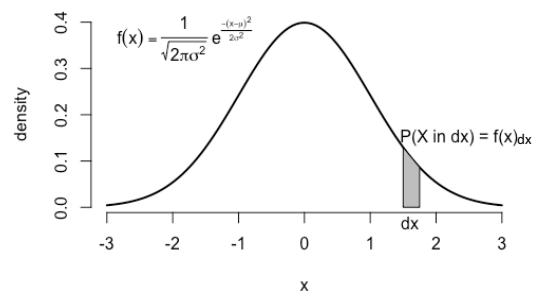
Another Continuous Distribution

The Normal(μ , σ^2)

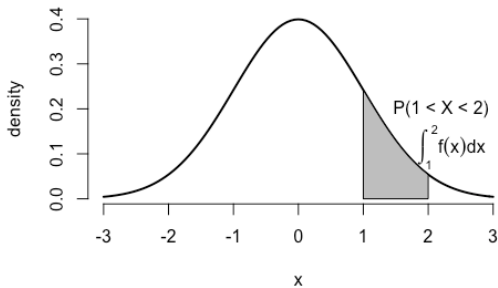
Normal Distribution



For a small region of width dx ,
 $P(X \text{ in } dx) = f(x)dx$



We find probabilities by integration

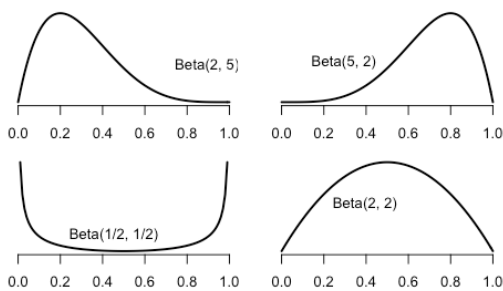


This integral doesn't have a closed form solution so we need the computer to approximate the area.

Another Continuous Distribution

The Beta(α, β)

The Family of Beta Distributions



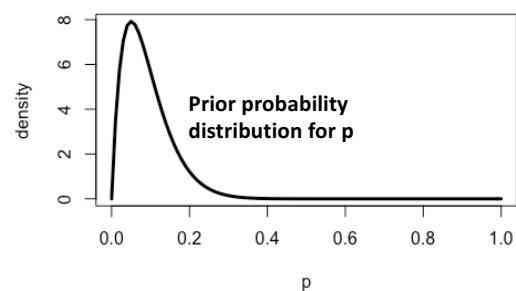
Beta(α, β)

- The Beta distribution is for random variables in (0, 1)
- The Beta(1,1) is the Uniform(0,1) distribution
- The Beta includes symmetric, skewed, U-shaped distributions
- The probability density function is $f(p) = B(\alpha, \beta) p^{\alpha-1} (1-p)^{\beta-1}$ for p in (0, 1)

Beta(α, β)

- Let's return to our click-through example
- Let's consider another prior distribution for p , such as the Beta(2, 10)
- The Beta(2, 10) probability density function is $f(p) = B(\alpha, \beta) p^{2-1} (1-p)^{10-1}$ for p in (0, 1)

Beta(2, 10)



Posterior for p given the data

Recall that we used Bayes rule to compute the

$$P(p = 0.01 \mid \text{data} = 25) =$$

$$P(p = 0.01) P(\text{data} = 25 \mid p = 0.01) / P(\text{data} = 25)$$

We use it again with the Beta density giving us the probability $P(p \text{ in } dp) = f(p) dp$

Posterior for p given the data

If we have a continuous distribution for the possible values of p, then

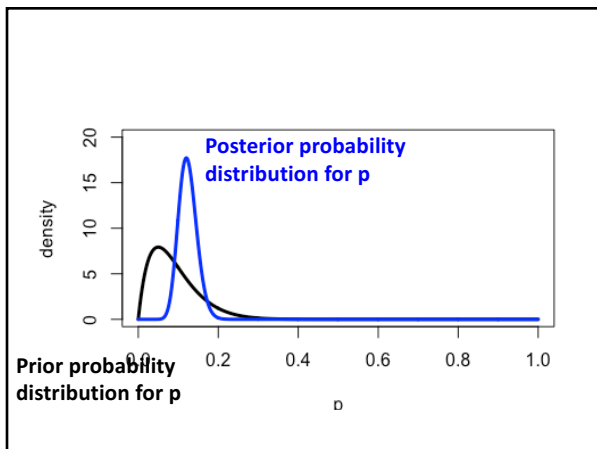
$$P(p \text{ in } dp \mid \text{data} = 25)$$

$$= P(p \text{ in } dp) P(\text{data} = 25 \mid p \text{ in } dp) / P(\text{data} = 25)$$

$$= B(\alpha, \beta) p(1-p)^9 C(200, 25) p^{25}(1-p)^{175}$$

$$= B(\alpha, \beta) C(200, 25) p^{25+1}(1-p)^{175+9}$$

This is the Beta(26, 184) distribution for p



Using the prior to provide a single point estimate for p

- Our posterior distribution for p is B(26,184)
- If we want to provide an estimate for p, then we could provide the p with the highest posterior probability, i.e., we would maximize

$$B(\alpha, \beta) C(200, 25) p^{25+1}(1-p)^{175+9}$$

- We know this is $p_{\text{hat}} = (25+1)/(200+10)$

Laplace Smoothing

- $p_{\text{hat}} = (25+1)/(200+10)$
- Notice that we have taken the MLE of 25/200 and added a small bit to the numerator and a small bit to the denominator
- This is especially useful for rare events, i.e., when p is small
- We can re-express p_{hat} as a weighted average of the MLE and the Ev of the prior:

$$\frac{200}{210} \frac{25}{200} + \frac{10}{210} \frac{1}{10}$$

Step Back

- Real world phenomena can often be modeled with a probability distribution
- Many probability distributions can be expressed as probability mass (or density) functions that depend on parameter(s)
- We can use Maximum Likelihood to estimate these parameters, i.e., to find the parameters that are most likely to have produced our data

- We may have prior information about the parameters
- This prior information can be expressed as a probability distribution on the parameter values
- We can use Bayes rule to find the posterior distribution of our parameters given our data
- Sometimes the prior fits nicely with the data distribution, e.g., the Beta and Binomial
- Other times we use computational methods to compute the posterior

Spam Detection

Spam

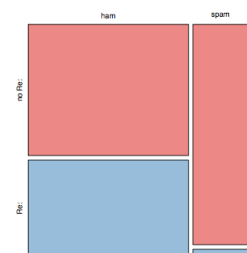
- Spam appears in our email, comments on blogs, reviews on Yelp, etc.
- We can develop detectors to help us programmatically identify spam
- In the case of email, Spam Assassin provided 9000 email messages that are hand-classified as spam or ham

Email Corpus

- Later in the semester we will discuss how to build classifiers
- We look at a simple example today
- From the 9000 email messages we determined
 - whether or not the subject line starts Re:
 - the percentage of capital letters in the email

Re: in the subject line of the email

	ham	spam	
Re:	2400	300	2700
No Re:	3600	2700	6300
	6000	3000	9000



A new email arrives, it has an Re: in its subject line.

What is the chance it is spam?

Is the presence of Re: a useful indicator of ham?

Re: in the subject line of the email

What are we assuming to answer this question?

	ham	spam	
Re:	2400	300	2700
No Re:	3600	2700	6300
	6000	3000	9000

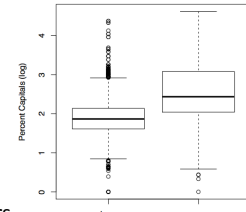
New email has a similar distribution of Re: within spam and ham as the corpus

We have enough data to accurately estimate this probability

$$\begin{aligned}\text{Prop}(\text{spam} | \text{Re}:) &= \\ \text{Prop}(\text{spam and Re:}) / \text{prop}(\text{Re:}) &= \\ = 300 / 2700 = 0.11\end{aligned}$$

Capitalization in the email

	ham	spam	
<20	4000	750	4750
20-30	1500	1500	3000
>30	500	750	1250
	6000	3000	9000



A new email arrives, it has more than 30% capital letters.

What is the chance it is spam?

Is the percent capitals a useful indicator of ham?

> 30% capital letters in the email

	ham	spam	
<20	4000	750	4750
20-30	1500	1500	3000
>30	500	750	1250
	6000	3000	9000

$$\begin{aligned}\text{Prop}(\text{spam} | > 30\%) &= \\ \text{Prop}(\text{spam and } >30\%) / \text{prop}(>30\%) &= \\ = 750 / 1250 = 0.6\end{aligned}$$

- A new email arrives.
- It has an Re: in the subject line and fewer than 20% of the letters are capitalized
- What is the chance it is spam?
- Can we answer this question?

$P(\text{spam} | \text{Re: and } <20\% \text{ caps})$

		ham	spam	
Re:	<20			
	20-30			
	>30			
No Re:	<20			
	20-30			
	>30			
		6000	3000	9000

$$\begin{aligned}P(\text{spam} | \text{Re: and } <20\%) &= \\ P(\text{spam \& Re: \& } <20\%) / & \\ P(\text{Re: \& } <20\%) &= \\ = 50 / 2050\end{aligned}$$

$P(\text{spam} | \text{Re: and } <20\% \text{ caps})$

		ham	spam	
Re:	<20	2000	50	2050
	20-30	300	100	400
	>30	100	150	250
No Re:	<20	2000	700	2700
	20-30	1200	1400	2600
	>30	400	600	1000
		6000	3000	9000

In practice

- We have many features x_1, x_2, \dots, x_m
- We observe x_1, x_2, \dots, x_m
- We want $P(\text{spam} | x_1, x_2, \dots, x_m)$
- Building a probability model is quite complex
- We don't have enough data to estimate the joint distribution of m random variables

$$P(\text{spam} | x_1, x_2, \dots, x_m)$$

$$= P(\text{spam and } x_1, x_2, \dots, x_m) / P(x_1, x_2, \dots, x_m)$$

Why? Definition of conditional probability

$$= P(\text{spam})P(x_1, x_2, \dots, x_m | \text{spam}) / P(x_1, x_2, \dots, x_m)$$

Why?

$$\text{Bayes Rule } P(A | B) = P(A)P(B | A) / P(B)$$

Naively assume independence

$$P(\text{spam} | x_1, x_2, \dots, x_m)$$

$$= P(\text{spam})P(x_1, x_2, \dots, x_m | \text{spam}) / P(x_1, x_2, \dots, x_m)$$

$$= P(\text{spam})P(x_1 | \text{spam}) * \dots * P(x_m | \text{spam}) / P(x_1, x_2, \dots, x_m)$$

Naïve Bayes Estimation of $P(\text{spam} | x_1, x_2, \dots, x_m)$

Computational Considerations

Take log to turn product of small probabilities into sums

$$\begin{aligned} \text{Log}(P(\text{spam})) &= \log(P(\text{spam})) + \sum \log(P(x_i | \text{spam})) \\ &\quad - \log(P(x_1, x_2, \dots, x_m)) \end{aligned}$$

$$\text{approx } \log(3/9) + \sum \log(\#x_i \text{ in spam}) / \# \text{spam} - C$$

$$\text{approx } \log(3/9) + \sum \log(\#x_i \text{ in spam} + 1) / (\# \text{spam} + 1) - C$$

Examine the likelihood ratio,

$$\text{Log}(P(\text{spam})/P(\text{ham}))$$

We don't need to compute $P(x_1, x_2, \dots, x_m)$

Values above 0 indicate $P(\text{spam}) > P(\text{ham})$

Take Aways

- In practice,
 - We might not have a named probability distribution so we resort to estimating probabilities with proportions
 - We might not have enough data so we smooth our proportions and make naïve assumptions
 - Computational considerations can be important for accuracy and efficiency