

## Chapter 6

### Appendix

#### 6.1 Appendix-1

Derivation of Bayes Filter using Bayes theorem and Markov's Assumptions:

$$P(B|A_1) = \frac{P(A_1 \cap B)}{P(A_1)} \Rightarrow P(A_1 \cap B) = P(B|A_1) * P(A_1)$$

similarly

$$P(B|A_2) = \frac{P(A_2 \cap B)}{P(A_2)} \Rightarrow P(A_2 \cap B) = P(B|A_2) * P(A_2)$$

Thus finally we see

$$\Rightarrow P(B) = \sum P(A_i) * P(A_i \cap B)$$

$$\Rightarrow P(A_1|B) = \frac{P(A_1 \cap B)}{P(B)} = \frac{P(B|A_1) * P(A_1)}{\sum P(A_i) * P(A_i \cap B)}$$

Probability of area shown in figure 6.1 can be represented as

$$P(x) = \int P(x, y) dy ; \text{ or } = \sum P(x_i, y)$$

but we know  $P(x, y) = P(x|y).P(y)$

$$\Rightarrow P(x) = \int P(x|y).P(y) \quad -1$$

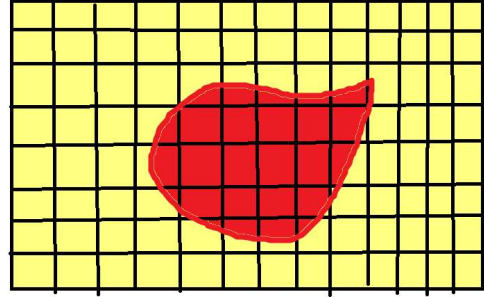


Figure-6.1 y spread in x domain

On extending the result -

$$P(A|B) = \frac{P(A, B)}{P(B)}$$

$$P(A|B, C) = \frac{P(B|A, C) * P(A|C)}{P(B|C)}, P(B|C) \text{ is usually } \eta-2$$

$$P(A, B) = \sum P(A, B, C_i) = \sum P(A|B, C_i) * P(B, C_i)$$

$$\Rightarrow \sum P(A|B, C_i) * P(C_i|B) * P(B) -3$$

Thus finally

$$P(A|B) = \frac{P(A, B)}{P(B)} = \frac{\sum P(A|B, C_i) * P(C_i|B) * P(B)}{P(B)}$$

$$\Rightarrow \sum P(A|B, C_i) * P(C_i|B) -4$$

Now on combining the above result and Markov Assumption we will get-

Bel- denotes belief of being at

$$Bel(X_t) = P(X_t|U_1, Z_1, \dots, U_t, Z_t)$$

we may assume  $X_t$  as A and  $U_1, Z_1, \dots$  as B and  $U_t, Z_t$  as C  
then apply eqn. 2

$$Bel(X_t) = \eta P(Z_t|X_t, U_1, Z_1, \dots, U_t) * P(X_t|U_1, Z_1, \dots, U_t)$$

then apply eqn. 5

$$Bel(X_t) = \eta P(Z_t|X_t) * P(X_t|U_1, Z_1, \dots, U_t)$$

then apply eqn. 4

$$Bel(X_t) = \eta P(Z_t|X_t) * \int P(X_t|U_1, Z_1, \dots, U_t, X_{t-1}) * P(X_{t-1}|U_1, Z_1, \dots, U_t) dX_{t-1}$$

then apply eqn. 6

$$Bel(X_t) = \eta P(Z_t|X_t) * \int P(X_t|X_{t-1}, U_t) * P(X_{t-1}|U_1, Z_1, \dots, U_t) dX_{t-1}$$

$$Bel(X_t) = \eta P(Z_t|X_t) * \int P(X_t|X_{t-1}, U_t) * P(X_{t-1}|U_1, Z_1, \dots, U_{t-1}, Z_{t-1}) dX_{t-1}$$

thus we see  $Be(X_{t-1})$  as second term in integration hence reaching t-1 state shows recursion

Thus we will obtain the final equation of as

$$Bel(X_t) = \eta P(Z_t|X_t) * \int P(X_t|X_{t-1}, U_t) * Bel(X_{t-1}) dX_{t-1}$$

## 6.2 Appendix-2

The linear state may be represented as-

$$x_{k+1} = A_k x_k + B_k u_k + G_k v_k$$

$$z_k = H_k x_k + w_k$$

→ Relation between two states  $x_{k-1}, x_k$  is linear

→ Gaussian distribution

→ for one dimensional 1-D →  $\mu$  (mean) ;  $\sigma^2$  (is variance)

→ for N dimensional N-D →  $\vec{\mu}$  [matrix] (mean) ;  $\sum$  (is covariance matrix)

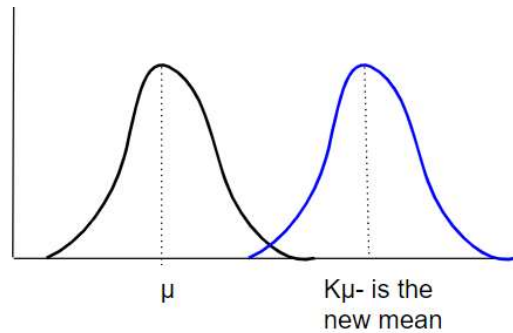


Figure-6.2- state multiplied by a scalar K

$$\mu' = K\mu \quad ; \quad \sigma'^2 = K^2 \sigma^2$$

while for N-Dimensional gaussian multiplied by F matrix

$$\vec{\mu}' = F\vec{\mu} \quad ; \quad \sum' = F \sum F^T$$

Now on modeling the belief as a Gaussian distribution-

→ belief is modeled as Gaussian

$$N(x, \mu, \sigma) = \frac{1}{\sigma\sqrt{2\Pi}} \exp\left(-\frac{(x - \mu)^2}{2\sigma^2}\right)$$

→ multiplication of two beliefs for a state :

$$N(x, \mu_0, \sigma_0) * N(x, \mu_1, \sigma_1) = N(x, \mu', \sigma')$$

$$\mu' = \mu_0 + \frac{\sigma_0^2(\mu_1 - \mu_0)}{\sigma_0^2 + \sigma_1^2}$$

$$\sigma'^2 = \sigma_0^2 - \frac{\sigma_0^4}{\sigma_0^2 + \sigma_1^2}$$

$$\text{if } K = \frac{\sigma_0^2}{\sigma_0^2 + \sigma_1^2} \text{ then } \mu' = \mu_0 + K(\mu_1 - \mu_0) ;$$

$$\text{and } \sigma'^2 = \sigma_0^2 - K\sigma_0^2$$

→ K is called as gain

⇒ if  $\sigma_0 \rightarrow \infty$  ;  $\mu' \rightarrow \mu_1$  this means higher uncertainty in  $\mu_0$

⇒ if  $\sigma_1 \rightarrow \infty$  ;  $\mu' \rightarrow \mu_0$  this means higher uncertainty in  $\mu_1$

→ for N-Dimensional Gaussian

$$K = \sum_0 (\sum_0 + \sum_1)^{-1} \text{ -A}$$

$$\vec{\mu}' = \vec{\mu}_0 + K(\vec{\mu}_1 - \vec{\mu}_0) \text{ -B}$$

$$\sum' = \sum_0 - K \sum_0 \text{ -C}$$

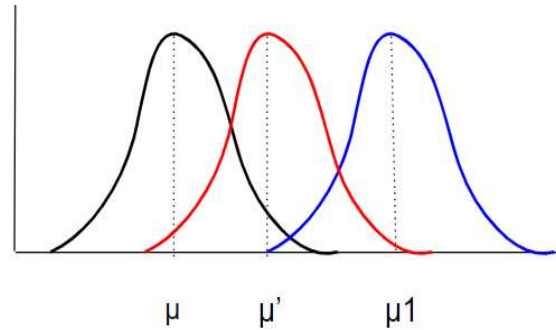


Figure-6.3-represents gaussian multiplication

### 6.2.1 Prediction Step

if we are to consider acceleration then the equation becomes

$$X_k = X_{k-1} + \Delta t + \frac{1}{2}a\Delta t^2$$

$$\widehat{x}_k = F_k \widehat{x}_{k-1} + \begin{bmatrix} \frac{\Delta t^2}{2} \\ \Delta t \end{bmatrix} a$$

$$\Rightarrow \widehat{x}_k = F_k \widehat{x}_{k-1} + B_k \vec{U}_k$$

while the new covariance matrix becomes :  $P_k = F_k P_{k-1} F_k^T + Q_k$

$B_k$  - is the control matrix

$\vec{U}_k$  - is the control vector or command input

$Q_k$  - is the Process noise matrix

### 6.2.2 Updation Step

$$\vec{\mu}_0 = H_k X_k \quad ; \quad \sum_0 = H_k P_k H_k^T$$

$$\text{Sensor data : } Z_k \quad ; \quad R_k \quad ; \quad \sum_k = H_k P_k H_k^T + R_k$$

$$(\mu_0, \sum_0) = (H_k \widehat{x}_k, H_k P_k H_k^T)$$

$$(\mu_1, \sum_1) = (Z_k, R_k)$$

$$H_k \widehat{x}_k' = H_k \widehat{x}_k + K(\vec{Z}_k - H_k \widehat{x}_k)$$

$$H_k P_k' H_k^T = H_k P_k H_k^T - K(H_k P_k H_k^T)$$

$$K = H_k P_k H_k^T (H_k P_k H_k^T + R_k)^{-1}$$

‘Thus the final optimum equation obtained is-

$$\widehat{x}_k' = \widehat{x}_k + K'(Z_k - H_k \widehat{x}_k)$$

$$P_k' = P_k - K' H_k P_k$$

$$K' = P_k H^T (H_k P_k H_k^T + R_k)^{-1}$$

### 6.3 Appendix-3

Taking a one-dimensional random variable  $x$  having a gaussian distribution.

- Mean  $\bar{x}$  ( $\mu$ ) and variance  $\sigma^2$  characterize the equation as

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

- Three sample points called as Sigma Points are selected

$$\tilde{x}^0 = \bar{x}$$

$$\tilde{x}^1 = \bar{x} + \sqrt{1+\kappa} \cdot \sigma$$

$$\tilde{x}^2 = \bar{x} - \sqrt{1+\kappa} \cdot \sigma$$

$$W_0 = \frac{\kappa}{1+\kappa}$$

$$W_1 = W_2 = \frac{1}{2(1+\kappa)}$$

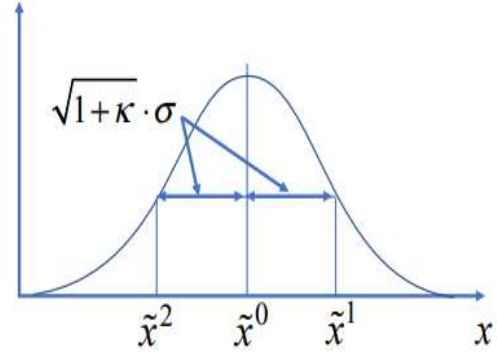


Figure 6.4 - One dimensional gaussian distribution showing sigma points. [5]

- Here  $\kappa$  is a parameter of sigma points to be tuned, and  $W_i$  is the weight of the  $i$ th sigma point used for computing mean and variance
- The weighted mean of the three Sigma points agrees with the true mean of the

$$\begin{aligned} \sum_{i=0}^2 W_i \tilde{x}^i &= \frac{\kappa}{1+\kappa} \bar{x} + \frac{1}{2(1+\kappa)} \left\{ (\bar{x} + \sqrt{1+\kappa} \cdot \sigma) + (\bar{x} - \sqrt{1+\kappa} \cdot \sigma) \right\} \\ &= \frac{\kappa}{1+\kappa} \bar{x} + \frac{2}{2(1+\kappa)} \bar{x} = \bar{x} \end{aligned}$$

- The weighted variance of the three Sigma points agrees with the true variance

$$\begin{aligned} \sum_{i=0}^2 W_i (\tilde{x}^i - \bar{x})^2 &= \frac{\kappa}{1+\kappa} (\bar{x} - \bar{x}) + \frac{1}{2(1+\kappa)} \left\{ (\bar{x} + \sqrt{1+\kappa} \cdot \sigma - \bar{x})^2 + (\bar{x} - \sqrt{1+\kappa} \cdot \sigma - \bar{x})^2 \right\} \\ &= \frac{2}{2(1+\kappa)} (\sqrt{1+\kappa} \cdot \sigma)^2 = \sigma^2 \end{aligned}$$

### 6.3.1 Unscented Transform

- Consider a non-linear transformation of  $x$  to  $y$  by  $g(x)$ .
- Here the function is

$$y = g(x) = \sum_{k=0}^{\infty} \frac{g^{(k)}(\bar{x})}{k!} (x - \bar{x})^k$$

- The distribution of  $y$  is no longer Gaussian, but its mean  $E[y]$  and variance  $E[(y-E[y])^2]$  can characterize the distribution

$$\bar{y}_{sample} \triangleq \sum_{i=0}^2 W_i \tilde{y}^i = E[y]$$

$$\sigma_{sample}^2 \triangleq \sum_{i=0}^2 W_i (\tilde{y}^i - \bar{y}_{sample})^2 = E[(y - E[y])^2]$$

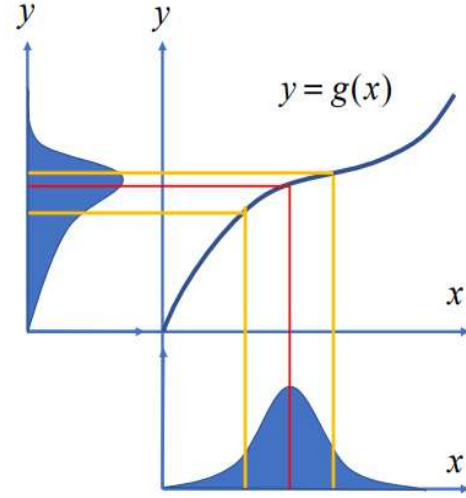


Figure 6.5 -Graph showing sigma points . [5]

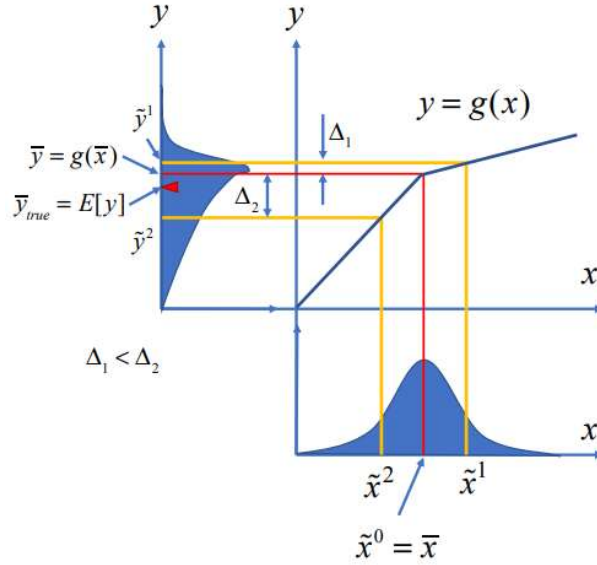


Figure 6.6 -Graph showing sigma points and their propagation . [5]

Thus we see that weighted mean of sigma points gives better mean value.

$$\begin{aligned} \bar{y}_{sample} &\triangleq \sum_{i=0}^2 W_i \tilde{y}^i = \frac{\kappa}{1+\kappa} \tilde{y}^0 + \frac{1}{2(1+\kappa)} (\tilde{y}^1 + \tilde{y}^2) \\ &= \frac{\kappa}{1+\kappa} \bar{y} + \frac{1}{2(1+\kappa)} (\bar{y} + \Delta_1 + \bar{y} - \Delta_2) = \bar{y} + \frac{1}{2(1+\kappa)} (\Delta_1 - \Delta_2) < \bar{y} \end{aligned}$$

### 6.3.2 Unscented Transform in a multidimensional Gaussian distribution

- For an n-dimensional Gaussian distribution, we use  $(2n + 1)$  Sigma points
- The covariance matrix  $P_x$  is real, symmetric, and positive-definite. Therefore, it can be diagonalized

$$P_X = V D V^T$$

$$V = (\mathbf{v}_1 \cdots \mathbf{v}_n), \mathbf{D} = \text{diag}(\sigma_1^2 \cdots \sigma_n^2)$$

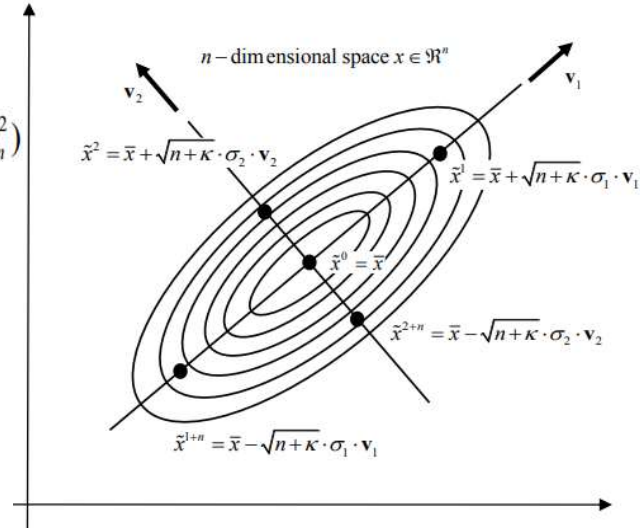


Figure 6.7- N dimensional distribution [5]

- Thus all points propagate through the function

$$\bar{y}_{sample} = \sum_{i=0}^{2n} W_i \tilde{y}^i$$

- The weighted mean is
- While weighted covariance is

$$P_{y,sample} = \sum_{i=0}^{2n} W_i (\tilde{y}^i - \bar{y}_{sample})(\tilde{y}^i - \bar{y}_{sample})^T$$

- weighted mean can approximate the true mean to the third order, and the weighted covariance to the second order.

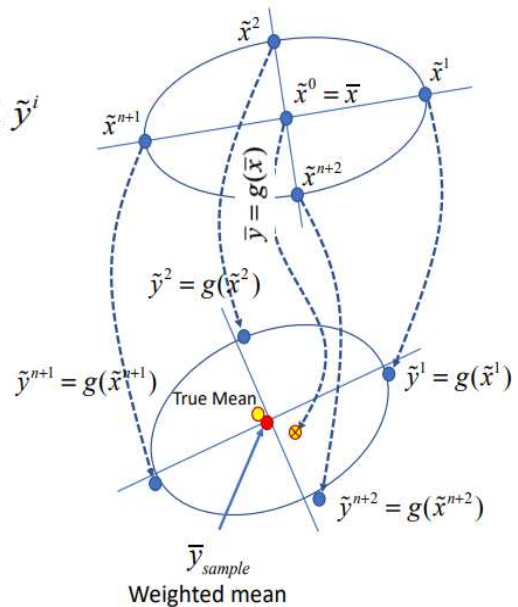


Figure 6.8 -Propagation via function. [5]



### 6.3.3 State Space and Prediction

- Let us consider a dynamic system represented as

$$\begin{aligned} u_t &= 0 \\ x_{t+1} &= f(x_t, u_t, t) + w_t \\ y_t &= h(x_t, t) + v_t \end{aligned}$$

- Process noise and measurement noise are zero-mean, uncorrelated (white) noise with covariance  $Q_t$  and  $R_t$   $w_t \sim N(0, Q_t)$ ,  $v_t \sim N(0, R_t)$
- At time  $t-1$  the state is  $x_{t-1}$  and the covariance is given by

$$P_{t-1} = E[(\hat{x}_{t-1} - x_{t-1})(\hat{x}_{t-1} - x_{t-1})^T]$$

- Then Using the eigen values and eigen vector of state  $P_{t-1}$  is found along with the sigma points  $\tilde{x}_{t-1}^0 = \hat{x}_{t-1} : \bar{x}$

$$\begin{aligned} \tilde{x}_{t-1}^i &= \hat{x}_{t-1} + \sqrt{n+K} \cdot \sigma_i \cdot v_i, \quad \tilde{x}_{t-1}^{i+n} = \hat{x}_{t-1} - \sqrt{n+K} \cdot \sigma_i \cdot v_i \\ i &= 1, \dots, n \end{aligned}$$

- Sigma points are propagated through the state equation

$$\tilde{x}_{t|t-1}^{i*} = f(\tilde{x}_{t-1}^i, t-1) + w_{t-1}, \quad i = 0, \dots, 2n$$

- The weighted mean is given as

$$\hat{x}_{t|t-1, sample} = \sum_{i=0}^{2n} W_i \hat{x}_{t|t-1}^{i*}$$

- The covariance is propagated as

$$P_{t|t-1, sample} = \sum_{i=0}^{2n} W_i (\tilde{x}_{t|t-1}^i - \hat{x}_{t|t-1, sample})(\tilde{x}_{t|t-1}^i - \hat{x}_{t|t-1, sample})^T + Q_{t-1}$$

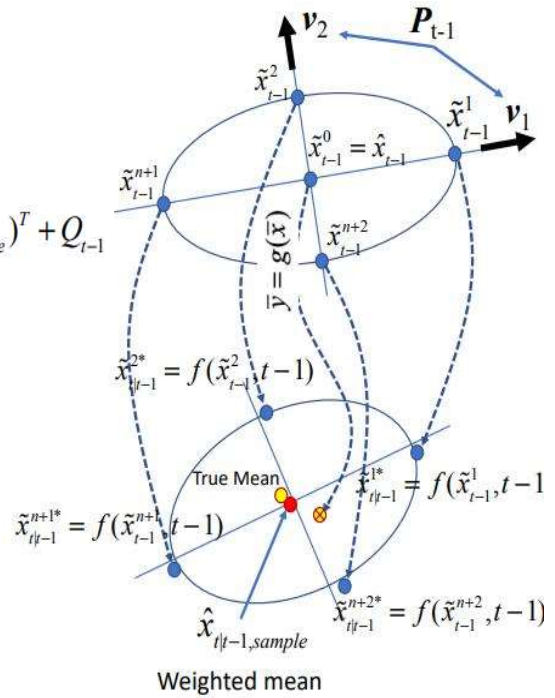


Figure 6.9 -Propagation of State. [5]

### 6.3.4 State Update

- State and covariance can be updated by the Unscented Transform on  $P_{t|t-1}$
- We first obtain the innovation covariance by examining the distribution of the output created through the measurement equation
- We compute the eigenvalues and eigen-vectors of the a priori covariance,
- Then, we generate  $(2n+1)$  Sigma points, and estimate the mean and covariance of the distribution of output  $y$  with the Sigma points
- Using the deterministic part of the measurement function, the Sigma points are mapped

$$\tilde{y}_t^i = h(\tilde{x}_{t|t-1}^i, t), \quad i = 0, \dots, 2n$$

- The weighted mean of the Sigma points is given by

$$\hat{y}_{t,sample} = \sum_{i=0}^{2n} W_i \tilde{y}_t^i$$

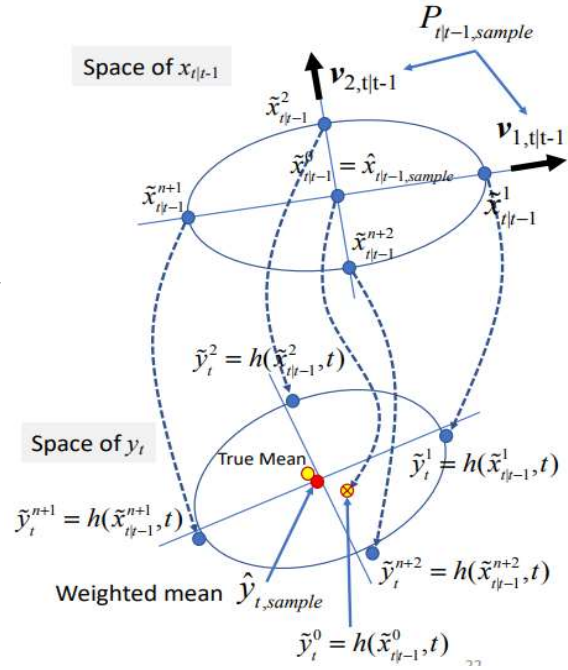


Figure 6.10 -Update of State. [5]

- The next step is computing the first term and cross variance

$$P_{xy} = \sum_{i=0}^{2n} W_i (\tilde{x}_{t|t-1}^i - \hat{x}_{t|t-1,sample}) (\tilde{y}_t^i - \hat{y}_{t,sample})^T \quad P_y = \sum_{i=0}^{2n} W_i (\tilde{y}_t^i - \hat{y}_{t,sample}) (\tilde{y}_t^i - \hat{y}_{t,sample})^T + R_t$$

- Kalman gain  $K$  can be computed by  $K_t = P_{xy} P_y^{-1}$
- State update given as  $\hat{x}_t = \hat{x}_{t|t-1,sample} + K_t [y_t - \hat{y}_{t,sample}]$
- Finally new covariance is given by  $P_t \cong P_{t|t-1,sample} - K_t P_y K_t^T$