

Gaussian Discriminant Analysis

We assume that $p(x|y) \sim \mathcal{N}(\mu, \Sigma)$ Multivariate Normal

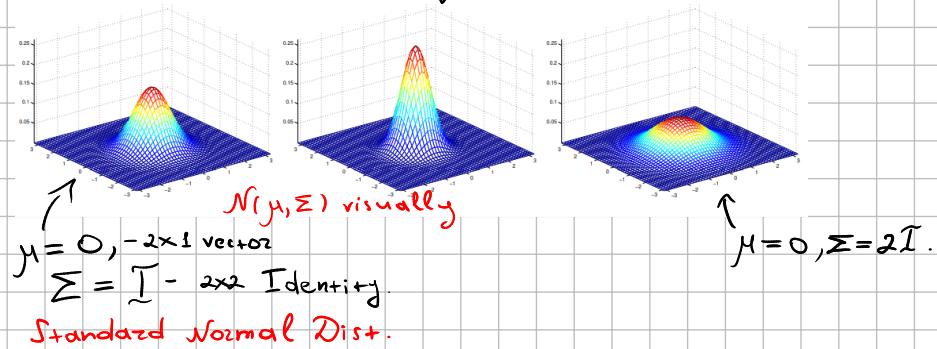
1. Multivariate Normal Distribution

μ - mean vector $\in \mathbb{R}^d$, Σ - cov. matrix $\in \mathbb{R}^{d \times d}$, it is symmetric, positive semi-definite > 0 .

$$\text{Its p.d.f. } p(x; \mu, \Sigma) = \frac{1}{2\pi^{d/2} |\Sigma|^{1/2}} \exp\left(-\frac{1}{2} (x-\mu)^T \Sigma^{-1} (x-\mu)\right)$$

determinant of Σ .

$$x = \mathbb{E}(X) = \int_x x p(x; \mu, \Sigma) dx$$



The GDA Model.

We have a classification problem, x - continuous-valued r.v.'s, we have a dataset $D = \{(x_i, y_i)\}_{i=1}^n$

$$y \sim \text{Bernoulli}(\varphi)$$

$$x|y=0 \sim \mathcal{N}(\mu_0, \Sigma)$$

$$x|y=1 \sim \mathcal{N}(\mu_1, \Sigma)$$

Both classes are normally distributed.

$$\text{Thus, } p(y) = \varphi^y (1-\varphi)^{1-y} ; p(x|y=0) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp\left(-\frac{1}{2} (x-\mu_0)^T \Sigma^{-1} (x-\mu_0)\right) ;$$

$$p(x|y=1) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp\left(-\frac{1}{2} (x-\mu_1)^T \Sigma^{-1} (x-\mu_1)\right)$$

usually applied using one covariance matrix

Parameters of our model are $\varphi, \Sigma, \mu_0, \mu_1$

The likelihood of our model: $\mathcal{L} = \prod_{i=1}^n p(x_i, y_i; \varphi, \mu_0, \mu_1, \Sigma)$

The log-likelihood: $\ell = \sum_{i=1}^n \log(p(x_i, y_i; \varphi, \mu_0, \mu_1, \Sigma)) = \sum_{i=1}^n \log(p(x_i|y_i; \mu_0, \mu_1, \Sigma)p(y_i; \varphi))$

We need to maximize it w.r.t. params.

Our objective: $\underset{\varphi, \mu_0, \mu_1, \Sigma}{\operatorname{argmax}} \log \left\{ \prod_{i=1}^n (p(x_i|y_i; \mu_0, \mu_1, \Sigma)p(y_i; \varphi)) \right\}$

$$\left. \frac{\partial \ell}{\partial \varphi} \right|_{\varphi=0} = \frac{\partial}{\partial \varphi} \left(\log \left(\prod_{i=1}^n p(y_i; \varphi) \right) \right) = \frac{\partial}{\partial \varphi} \left(\log \left(\prod_{i=1}^n (\varphi^y_i (1-\varphi)^{1-y_i}) \right) \right) = \frac{\partial}{\partial \varphi} \left(\sum_{i=1}^n \log (\varphi^{y_i} (1-\varphi)^{1-y_i}) \right) = \frac{\partial}{\partial \varphi} \left(\sum_{i=1}^n (\log(\varphi^{y_i}) + (1-y_i) \log(1-\varphi)) \right) =$$

spec log = ln

$$= \sum_{i=1}^n \left(\frac{\partial \log(\varphi)}{\partial \varphi} \cdot y_i + (1-y_i) \frac{\partial \log(1-\varphi)}{\partial \varphi} \right) = \sum_{i=1}^n \left(y_i \cdot \frac{1}{\varphi} - (1-y_i) \cdot \frac{1}{1-\varphi} \right) = \sum_{i=1}^n \left(\frac{y_i}{\varphi} - \frac{1-y_i}{1-\varphi} \right) = 0 \Rightarrow n_1 = \varphi(n_1 + n_0) \Rightarrow \varphi = \frac{n_1}{n_1 + n_0}$$

My understanding to $\frac{\partial \ln(x)}{\partial x} = \frac{1}{x}$

$$\frac{\partial \ln(1-\varphi)}{\partial \varphi} = \frac{1}{1-\varphi} \cdot \frac{1}{1-\varphi} = (-1) \frac{1}{1-\varphi}$$

$$\sum_{i=1}^n y_i (1-\varphi) - (1-y_i) \varphi = 0 \Rightarrow n_1 (1-\varphi) - n_0 \varphi = 0$$

n_1 observed $y_i = 1$

n_0 observed $y_i = 0$

$$n_1 - \varphi(n_1 + n_0) = 0$$

$$2) \frac{\partial \ell}{\partial \mu_0} = \frac{\partial}{\partial \mu_0} \left\{ \log \left(\prod_{i=1}^{n_0} p(x_i | y_i=0; \mu_0, \Sigma) \right) \right\} = \frac{\partial}{\partial \mu_0} \left\{ \log \left(\prod_{i=1}^{n_0} \left(\frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp \left[-\frac{1}{2} (x_i - \mu_0)^\top \Sigma^{-1} (x_i - \mu_0) \right] \right) \right) \right\} =$$

$$= \frac{\partial}{\partial \mu_0} \left\{ \sum_{i=1}^{n_0} \log \left(\frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp \left[-\frac{1}{2} (x_i - \mu_0)^\top \Sigma^{-1} (x_i - \mu_0) \right] \right) \right\}$$

$$\sum_{i=1}^{n_0} \left(\log \left(\frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \right) - \underbrace{\frac{1}{2} (x_i - \mu_0)^\top \Sigma^{-1} (x_i - \mu_0)}_{\text{line 1}} \log \left(\frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \right) \right) = \sum_{i=1}^{n_0} \sum^{-1} (x_i - \mu_0) = 0 \quad / \cdot \Sigma \Rightarrow \sum_{i=1}^{n_0} (x_i - \mu_0) = 0$$

$$\frac{\partial}{\partial \mu_0} = 0 \quad \frac{\partial}{\partial \mu_0} = \frac{1}{2} \sum^{-1} \cdot 2 (x_i - \mu_0) (-1) = -\sum^{-1} (x_i - \mu_0)$$

3) Same procedure for $\mu_1 \Rightarrow \mu_1^* = \frac{n_0 x_i}{n_0}$

$$4) \text{ For } \Sigma, \quad \frac{\partial \ell}{\partial \Sigma} = \frac{\partial}{\partial \Sigma} \left(\log \left(\prod_{i=1}^n \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp \left[-\frac{1}{2} (x_i - \mu_i)^\top \Sigma^{-1} (x_i - \mu_i) \right] \right) \right) = \frac{\partial}{\partial \Sigma} \left(\sum_{i=1}^n \left(\log \left(\frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \right) + \right. \right.$$

$$\left. \left. + \left(-\frac{1}{2} (x_i - \mu_i)^\top \Sigma^{-1} (x_i - \mu_i) \right) \right) \right) = \frac{n}{2} \log |\Sigma| - \frac{1}{2} S_i^\top = \frac{n}{2} \log |\Sigma| - \frac{1}{2} (x_i - \mu_i) (x_i - \mu_i)^\top = 0$$

$$\log((2\pi)^{-d/2} |\Sigma|^{-1/2}) = \log(2\pi) \cdot \left(-\frac{d}{2} \right) - \frac{1}{2} \log(|\Sigma|) = \left| \sum_{i=1}^n \right| = -\frac{n}{2} \log |\Sigma| = \frac{n}{2} \log |\Sigma|$$

$$\sum_{i=1}^n \frac{\partial}{\partial \Sigma} \left(-\frac{1}{2} (x_i - \mu_i)^\top \Sigma^{-1} (x_i - \mu_i) \right) = \sum_{i=1}^n \frac{\partial}{\partial \Sigma} \left(-\frac{1}{2} \text{tr}((x_i - \mu_i) (x_i - \mu_i)^\top \Sigma^{-1}) \right) = \text{let's differentiate w.r.t. } \Sigma^{-1} = W =$$

$$= \frac{\partial}{\partial W} \left(\sum_{i=1}^n -\frac{1}{2} \text{tr}((x_i - \mu_i) (x_i - \mu_i)^\top W) \right) = -\frac{1}{2} \sum_{i=1}^n \frac{\partial \text{tr}((x_i - \mu_i) (x_i - \mu_i)^\top W)}{\partial W} = -\frac{1}{2} \sum_{i=1}^n \frac{\partial \text{tr}(S_i W)}{\partial W} = -\frac{1}{2} S_i^\top$$

let $(x_i - \mu_i) (x_i - \mu_i)^\top = S_i$

$$\frac{n}{2} \log |\Sigma| = \frac{1}{2} \sum_{i=1}^n (x_i - \mu_i) (x_i - \mu_i)^\top \Rightarrow \Sigma^* = \frac{1}{n} \sum_{i=1}^n (x_i - \mu_i) (x_i - \mu_i)^\top$$

The optimal values are:

$$\varphi = \frac{1}{n} \sum_{i=1}^n \mathbb{1}\{y_i=1\}; \quad \mu_0 = \frac{\sum_{i=1}^n \mathbb{1}\{y_i=0\} x_i}{\sum_{i=1}^n \mathbb{1}\{y_i=0\}}; \quad \mu_1 = \frac{\sum_{i=1}^n \mathbb{1}\{y_i=1\} x_i}{\sum_{i=1}^n \mathbb{1}\{y_i=1\}}; \quad \Sigma = \frac{1}{n} \sum_{i=1}^n (x_i - \mu_i) (x_i - \mu_i)^\top$$

Two Gaussians have contours that are the same shape and orientation, since they share a cov. matrix Σ .

