Recap: A,h, A is a Kxr integer matrix with 1 in its rowspan $h \in \mathbb{R}^r$.

Want to test Ho: p∈MA, h vs. Ho: p & MA, h.

To do: Compute p-value. asymptotic vs. exact approach.

$$\frac{\text{Prop 1:}}{\text{P(U=u)}} \quad \text{If } \rho \in \mathcal{M}_{A}, \text{ and } \rho(x) = \theta_{i}^{a_{ix}} \cdot \theta_{k}^{a_{kx}}, \text{ $x \in [r]$, then} \\ P(U=u) = \frac{n!}{\prod_{i \in [r]}} \quad \text{and the probability } P(U=u \mid AU=Au) \\ \text{does not depend on } \rho.$$

· Based on Prop 1, we generalize Fisher's exact test by computing the p-value

$$P(X^2(U) > X^2(u) \mid AU = Au)$$

Here
$$X^{2}(U) = \sum_{i \in \mathcal{U}_{i}} \frac{(U_{i} - \hat{u}_{i})^{2}}{\hat{u}_{i}}$$

 $\hat{\mathcal{U}}_i = n \hat{\rho}_i$ where

Hence the p-value is:

$$\frac{\sum_{v \in \mathcal{F}(u)} 1_{\chi^{2}(v)}, \chi^{2}(u)}{\left(\prod_{i \in r} v_{i}! \right)}$$

$$\frac{\sum_{v \in \mathcal{F}(u)} 1 / \left(\prod_{i \in r} v_{i}! \right)}{\left(\prod_{i \in r} v_{i}! \right)}$$

→ Exact computation of this quantity is prohibitive. Thus we sample from elements in the fiber. Def: Let MA be the log-linear model associated with a matrix A, 1 Erowspan(A) A finite subset BcKer, A is a Markov basis for MA if for all $u \in T(n)$ and all pairs $v, v' \in F(u)$ there exists a sequence u,..., uzeB such that

 $v' = v + \sum_{\kappa=1}^{L} u_{\kappa}$ and $v + \sum_{\kappa=1}^{l} u_{\kappa} > 0$ for all l = 1, ..., L

This is saying entries in the transition tables always need to be positive

Algorithm: (Metropolis- Hastings)

INPUT: A table $u \in T(n)$ and a Markov basis for MA

OUTPUT: A sequence $(\times^2(v_i), \times^2(v_2), \dots, \infty)$ for tables Vt in the fiber 子(u).

STEP 1: VI = U.

STEP 2: For t= 1,2,...

(i) Select uniformly at random a move ute8 (ii) If any coordinate of v+u+ is negative, set Vt+1 = Vt, else set

$$V_{t+1} = \begin{cases} v_t + U_t \\ v_t \end{cases} \quad \text{with probability} \quad \begin{cases} q \\ 1 - q \end{cases}$$

where
$$q = \min \left\{ 1, \frac{P(U = v_t + u_t \mid AU = Au)}{P(U = v_t \mid AU = Au)} \right\}$$

4 The fact that this (iii) Compute X2(Vt) is a ratio means we

don't have to comput

Thm: The output $(X^2(v_t))_{t=1}^{\infty}$ of ALG. is an aperiodic, reversible and irreducible Markov chain that has stationary distribution equal to the conditional distribution of X2(U) given AU= Au.

Coroll 1: With probability one, the output sequence $(X^2(v_t))_{t=1}^{\infty}$ of ALG satisfies

$$\lim_{M\to\infty} \frac{1}{M} \sum_{t=1}^{M} 1_{\left\{ \chi^{2}(v_{t}) \right\} \chi^{2}(u) \right\}} = \rho(\chi^{2}(U) \chi^{2}(u))$$

$$AU = Au$$

Thm: (Fundamental Thm of Markov bases). A finite subset BS Kerz A is a Markov basis for A

 $\{\rho^{\nu^{+}}-\rho^{\nu^{-}}: \nu \in \mathcal{B}\}$ is a generating set for the

$$\int$$
 is a generating set for the toric ideal I_{A} .

Conditional independence models Ch.4. Sullivant

Let $X=(X_1,...,X_m)$ be a random vector with state space $X=\prod_{i=1}^m x_i$.

 $f(x) = f(x_1, ..., x_m)$ is the joint pdf of X.

• For $A \subseteq [m]$, $X_A = (X_a)_{a \in A}$, $X_A = \prod_{a \in A} X_a$.

Def: · Let A = [m]. The marginal density fA(XA) of XA is obtained by integrating out X[m]\A:

$$f_A(x_A) := \int f(x_A, x_{[m]\setminus A}) d\nu_{[m]\setminus A}(x_{[m]\setminus A}), x_A \in X_A$$

• Let A,B \subseteq [m] be disjoint and $x_B \in X_B$. The conditional density of X_A given $X_B = x_B$ is

$$f_{A|B}(x_{A}|x_{B}) = \begin{cases} \frac{f_{AUB}(x_{A}, x_{B})}{f_{B}(x_{B})}, & \text{if } f_{B}(x_{B}) > 0 \\ 0 & \text{otherwise} \end{cases}$$

Let A, B, C ⊆ [m] be pairwise disjoint. The random vector XA
is conditionally independent of XB given Xc if and only if

for all XA, XB, Xc. Notation: XA ! XB | Xc

$$f_{A|BUC}(x_A|x_B,x_c) = \frac{f_{AUB|C}(x_A,x_B|x_c)}{f_{B|C}(x_B|x_c)} = f_{A|C}(x_A|x_c)$$

"Given Xc, knowing XB does not give any information about XA"

Suppose $X=(X_1,...,X_m)$ is discrete. X_j has outcome space $[r_j]$. $X=\prod_{j\in [m]} r_{j}$

<u>Proposition:</u> If X is a discrete random vector, then the condition. Independence statement XAILXBIXc holds if and only if

 $\begin{aligned}
\rho_{i,A,i_B,i_C,+} \cdot \rho_{j,A,j_B,i_C,+} - \rho_{i,A,j_B,i_C,+} \cdot \rho_{j,A,i_B,i_C,+} &= 0 \\
\text{for all } i_A, i_B \in X_A \quad i_A, i_B \in X_B \quad i_A \in X_C
\end{aligned}$

for all $i_{A,j_A} \in X_A$, $i_{B,j_B} \in X_B$, $i_C \in X_C$.

 $P_{i,a,i_B,i_C,+} = P(X_A = i_A, X_B = i_B, X_C = i_C)$ $P_{i,a,i_B,i_C,+} = P(X_A = i_A, X_B = i_B, X_C = i_C)$ $P_{i,a,i_B,i_C,+} = P(X_A = i_A, X_B = i_B, X_C = i_C)$ These can be realized as minors of $i_{i_B} = P(X_B = x_B)$ $p_{i_B} = P(X_B = x_B)$ $p_{i_B} = P(X_B = x_B)$ $p_{i_B} = p_{i_B} = p_{i_B}$ $p_{i_B} = p_{i_B} = p_{i_B} = p_{i_B}$ $p_{i_B} = p_{i_B} = p_{i_$

 $\rho_{i_A, i_B, i_C, +} = \sum_{\substack{\text{over all subindices } \overline{1} \\ \text{in } TTX_i \\ i \in [m] \setminus (AUBUC)}}$

 $\rho f. \quad P(X_A = i_A, X_B = i_B, X_C = i_C) \cdot P(X_A = j_A, X_B = j_B, X_C = i_C)$ $= P(X_A = i_A \mid X_B = i_B, X_C = i_C) P(X_B = i_B, X_C = i_C)$ $\cdot P(X_A = j_A \mid X_B = j_B, X_C = i_C) P(X_B = j_B, X_C = i_C)$

Def: The conditional independence ideal IAMBIC = C[P] is generated by all the quadratic polynomials in the Prop. The discrete

conditional independence model, $r=1\times 1$ MALLBIC = $V_{\Delta}(I_{\Delta} \sqcup B_{1}) \subseteq \Delta_{r-1}$ is the set of all probability distributions in Δ_{r-1} satisfying the polyn in Prop.

If $C = \{A_1 \coprod B_1 \mid C_1, A_2 \coprod B_2 \mid C_2, \dots \}$ is a set of conditional

independence statements, we construct the conditional indepideal

which is the sum of ideals generated by quadrics The model $\mathcal{M}_{\mathcal{C}} := V_{\Delta}(\mathbf{I}_{\mathcal{C}}) \subseteq \Delta r$ consists of all probability

(1e) = 21r consists of all probabile distributions satisfying the constraints in C.

Example: Binary contraction axiom.

$$C = \{ 1 \perp 1 \mid 2 \mid 3, 2 \perp 1 \mid 3 \}$$
 $\Rightarrow \{ 2 \perp 1 \mid 1,3 \}$