Modelos Lineares e Aplicações Formulário – 2021/2022

Genéricos

Cov[aX + b, Y] = aCov[X, Y]

$$S_{XX} = \sum_{i=1}^{n} (x_i - \overline{x})^2 = \sum_{i=1}^{n} x_i^2 - n\overline{x}^2$$
 $S_{YY} = \sum_{i=1}^{n} (y_i - \overline{y})^2 = \sum_{i=1}^{n} y_i^2 - n\overline{y}^2$

$$S_{XY} = \sum_{i=1}^{n} (x_i - \overline{x})(y_i - \overline{y}) = \sum_{i=1}^{n} x_i y_i - n\overline{x} \overline{y}$$

Coeficientes de correlação de Pearson: $r_{xy} = \frac{S_{XY}}{\sqrt{S_{XX}S_{YY}}}$

$$H_0: \rho = 0 \text{ vs } H_1: \rho \neq 0 \text{ sendo } ET = \frac{\sqrt{n-2}}{\sqrt{1-r_{xy}^2}} \times r_{xy} \sim t_{n-2} \text{ e } RC = \{t: |t| > t_{\frac{\alpha}{2};n-2}\}$$

Coeficientes de correlação de Spearman: $r_S = 1 - \frac{6\sum_{i=1}^n d_i^2}{n(n^2-1)}$ em que d_i são as diferenças entre as ordens de x_i e y_i

Regressão linear simples

$$Y_i = \beta_0 + \beta_1 x_i + \epsilon_i$$

$$\hat{\beta}_1 = \frac{S_{XY}}{S_{XX}} \qquad \qquad \hat{\beta}_0 = \overline{Y} - \hat{\beta}_1 \overline{x} \qquad \qquad SSE = \sum_{i=1}^n (Y_i - \hat{Y}_i)^2$$

$$H_0: \beta_1 = 0 \text{ vs } H_1: \beta_1 \neq 0 \text{ sendo } ET = \frac{\hat{\beta}_1}{\sqrt{\hat{\sigma}^2/S_{XX}}} \sim t_{n-2} \text{ e } RC = \{t: |t| > t_{\frac{\alpha}{2};n-2}\}$$

$$H_0: \beta_0 = 0 \text{ vs } H_1: \beta_0 \neq 0 \text{ sendo } ET = \frac{\hat{\beta}_0}{\sqrt{\hat{\sigma}^2 \left(\frac{1}{n} + \frac{\overline{x}^2}{S_{XX}}\right)}} \sim t_{n-2} \text{ e } RC = \{t: |t| > t_{\frac{\alpha}{2}; n-2}\}$$

onde
$$\hat{\sigma}^2 = \frac{1}{n-2} \left(S_{YY} - \frac{S_{XY}^2}{S_{XX}} \right) = \frac{SSE}{n-2}$$

Coeficientes de determinação: $R^2 = \frac{S_{XY}^2}{S_{XX}S_{YY}} = r_{xy}^2$

Intervalo de Confiança a 100(1-
$$\alpha$$
)% para $E[Y_0]$: $\hat{E}[Y_0] \pm t_{\frac{\alpha}{2};n-2} \times \sqrt{\hat{\sigma}^2 \left(\frac{1}{n} + \frac{(\overline{x} - x_0)^2}{S_{XX}}\right)}$ onde $\hat{E}[Y_0] = E[Y|x = x_0] = \hat{\beta}_0 + \hat{\beta}_1 x_0$

Intervalo de Predição a 100(1-
$$\alpha$$
)% para Y_0 : $\hat{Y}_0 \pm t_{\frac{\alpha}{2};n-2} \times \sqrt{\hat{\sigma}^2 \left(1 + \frac{1}{n} + \frac{(\overline{x} - x_0)^2}{S_{XX}}\right)}$ onde $\hat{Y}_0 = \hat{\beta}_0 + \hat{\beta}_1 x_0$

ANOVA: SST=SSR+SSE
$$\iff \sum_{i=1}^{n} (y_i - \overline{y})^2 = \sum_{i=1}^{n} (\hat{y}_i - \overline{y})^2 + \sum_{i=1}^{n} (y_i - \hat{y}_i)^2$$

$$H_0: \beta_1 = \dots = \beta_p = 0 \text{ vs } H_1: \exists \beta_j \neq 0, \quad j = 1 \dots p \quad \text{sendo} \quad ET = \frac{MSR}{MSE} \sim F_{(p,n-p-1)} \quad e$$

$$RC = \{F : F > F_{\alpha;p,n-p-1}\}$$
 onde $MSR = \frac{SSR}{p}$ e $MSE = \frac{SSE}{n-p-1}$

Análise de diagnóstico. Observações discordantes.

- outlier, se $|r_i| \geq 3$, onde $r_i \sim T$ _student
- ponto com elevado "leverage", se $|h_{ii}| > \frac{2(p+1)}{n}$
- ponto influente, se distância de Cook $|c_i| > F(0.5, p+1, n-p-1)$ ou $|\text{dffits}_i| > 2\sqrt{\frac{p+1}{n-p-1}}$

Regressão linear múltipla

$$Y_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_p x_{ip} + \epsilon_i \iff Y = X\beta + \epsilon \text{ (notação matricial)}$$

$$\hat{\beta} = (X^T X)^{-1} X^T Y \qquad \qquad \Sigma_{\beta} = \hat{\sigma}^2 (X^T X)^{-1} \quad \text{onde} \quad \hat{\sigma}^2 = \frac{\left(Y - \hat{Y}\right)^T \left(Y - \hat{Y}\right)}{n - p - 1} = \frac{SSE}{n - p - 1} \quad \text{e} \quad \hat{Y} = X \hat{\beta}$$

$$H_0: \beta_j = b_j \text{ vs } H_1: \beta_j \neq b_j \text{ sendo } ET = \frac{\hat{\beta}_j - b_j}{\sqrt{\hat{\sigma}^2 C_{ij}}} \sim t_{n-p-1} \text{ e } RC = \{t: |t| > t_{\frac{\alpha}{2}; n-p-1}\}$$

onde C_{jj} é o j-ésimo elemento da diagonal principal da matriz $C=(X^TX)^{-1}$

Coeficientes de determinação:
$$R^2 = 1 - \frac{SSE}{SST} = \frac{SSR}{SST}$$

Coeficientes de determinação ajustado:
$$R_a^2 = 1 - \frac{n-1}{n-p-1}(1-R^2)$$

Intervalo de Confiança a 100(1-
$$\alpha$$
)% para $E[Y_0]$: $\hat{E}[Y_0] \pm t_{\frac{\alpha}{2};n-p-1} \times \sqrt{\hat{\sigma}^2 \left(\vec{x_0}^T C \vec{x_0}\right)}$ onde $\hat{E}[Y_0] = E[Y|x = \vec{x_0}] = \hat{\beta}_0 + \hat{\beta}_1 x_{01} + \dots + \hat{\beta}_p x_{0p} = \vec{x_0}^T \hat{\beta}$

Intervalo de Predição a 100(1-
$$\alpha$$
)% para Y_0 : $\hat{Y}_0 \pm t_{\frac{\alpha}{2};n-p-1} \times \sqrt{\hat{\sigma}^2 \left(1 + \vec{x_0}^T C \vec{x_0}\right)}$ onde $\hat{Y}_0 = \hat{\beta}_0 + \hat{\beta}_1 x_{01} + \dots + \hat{\beta}_p x_{0p} = \vec{x_0}^T \hat{\beta}$

Teste de Fisher-parcial

$$H_0: r$$
 parâmetros são nulos $(r < p)$ vs $H_1: Y_i = \beta_0 + \beta_1 x_{i1} + \dots + \beta_p x_{ip} + \epsilon_i$

sendo
$$ET = F_0 = \frac{n-p-1}{r} \times \frac{SSE(H_0) - SSE(H_1)}{SSE(H_1)} \sim F_{(r,n-p-1)}$$
 e $RC = \{F: F > F_{\alpha;r,n-p-1}\}$

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