# **Martingales**

### Conditional expectation

### Definition (Conditional expectation)

Given a random variable X that is  $\mathcal{F}$ —measurable and a  $\sigma$ -algebra  $\mathcal{G}\subseteq\mathcal{F}$ , the conditional expectation  $\mathbb{E}[X|\mathcal{G}]$  is a *random variable* such that

- $\mathbb{E}[X|\mathcal{G}]$  is  $\mathcal{G}$ —measurable;
- ②  $\mathbb{E}\left[\mathbb{1}_A\mathbb{E}[X|\mathcal{G}]\right] = \mathbb{E}[\mathbb{1}_AX]$  for all  $A \in \mathcal{G}$ . Equivalently,

$$\int_{A} \mathbb{E}[X|\mathcal{G}](\omega)dP(\omega) = \int_{A} X(\omega)dP(\omega).$$

Existence and uniqueness (almost everywhere) of  $\mathbb{E}[X|\mathcal{G}]$  follow from the Radon-Nikodym Theorem.

### Conditional expectation

Intuitively,  $\mathbb{E}[X|\mathcal{G}]$  is the best approximation of X given the information contained in  $\mathcal{G}$ , or the projection of X onto the sub- $\sigma$ -algebra  $\mathcal{G}$ . It can be proven that if  $\mathbb{E}[X^2] < \infty$ , then  $\mathbb{E}[X|\mathcal{G}]$  minimizes

$$\mathbb{E}[(X-\xi)^2]$$

over all random variables  $\xi$  such that

- lacktriangle  $\xi$  is  $\mathcal{G}$ -measurable;

If  $\mathcal{G}$  is the trivial  $\sigma$ -algebra  $\{\varnothing,\Omega\}$  then  $\mathbb{E}[X|\mathcal{G}]=\mathbb{E}[X]$ .

# Properties of conditional expectation

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Linearity \mathbb{E}[aX+bY|\mathcal{G}]=a\mathbb{E}[X|\mathcal{G}]+b\mathbb{E}[Y|\mathcal{G}] a.s. . Monotonicity if X\leq Y a.s. then \mathbb{E}[X|\mathcal{G}]\leq \mathbb{E}[Y|\mathcal{G}] a.s. . Identity if X is \mathcal{G}—measurable, then \mathbb{E}[X|\mathcal{G}]=X a.s. . Jensen's inequality if \phi is convex and \mathbb{E}[|\phi(X)|]<\infty, then \mathbb{E}[\phi(X)|\mathcal{G}]\geq \phi\Big(\mathbb{E}[X|\mathcal{G}]\Big) a.s. . Pulling out what's known if Y is measurable and \mathbb{E}[|Y|]<\infty then \mathbb{E}[XY|\mathcal{G}]=Y\mathbb{E}[X|\mathcal{G}] a.s. .
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# Properties of conditional expectation

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Tower property if \mathcal{G}'\subseteq\mathcal{G} then \mathbb{E}[\mathbb{E}[X|\mathcal{G}]|\mathcal{G}']=\mathbb{E}[X|\mathcal{G}'] a.s. .
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- Irrelevance of independent information if  $\mathcal H$  is independent of  $\sigma(X,\mathcal G)$  then  $\mathbb E[X|\sigma(\mathcal G,\mathcal H)]=\mathbb E[X|\mathcal G]$  a.s. . In particular, if  $\mathcal H$  is independent of X then  $\mathbb E[X|\mathcal H]=\mathbb E[X]$  a.s. .
- Monotone convergence if  $X_n \uparrow X$  a.s. and  $\mathbb{E}[|X_n|], \mathbb{E}[X] < \infty$  then  $\mathbb{E}[X_n|\mathcal{G}] \uparrow \mathbb{E}[X|\mathcal{G}]$  a.s. .
- Fatou's lemma if  $X_n \geq 0$  and  $\mathbb{E}[|\liminf_{n \to \infty} X_n|] < \infty$  then  $\mathbb{E}[\liminf_{n \to \infty} X_n | \mathcal{G}] \leq \liminf_{n \to \infty} \mathbb{E}[X_n | \mathcal{G}]$  a.s. .
- Dominated convergence if  $X_n \leq Y$  with  $\mathbb{E}[|Y|] < \infty$  and  $X_n \to X$  a.s. then  $\mathbb{E}[X_n | \mathcal{G}] \to \mathbb{E}[X | \mathcal{G}]$  a.s. .

### Definition (Martingale)

A stochastic process  $(X_i)_{i \in I}$  is a martingale with respect to a filtration  $(\mathcal{F}_i)_{i \in I}$  if

- $X_i$  is  $\mathcal{F}_i$ -measurable for each  $i \in I$ ;
- $\mathbb{E}[|X_i|] < \infty$  for each  $i \in I$ ;
- $\mathbb{E}[X_j | \mathcal{F}_i] = X_i$  a.s. for each  $i < j \in I$ .

We simply say that a stochastic process is a martingale if it is so with respect to its natural filtration.

Martingales are *fair games*: if we are playing in a fair game, with  $\mathcal{F}_i$  being the information on what happened in the game up to time i, and  $X_i$  is our wealth at time i, then the expected value of our wealth in the future, given what has happened so far, is our current wealth.

By the tower property,  $\mathbb{E}[X_j] = \mathbb{E}[\mathbb{E}[X_j|\mathcal{F}_i]] = \mathbb{E}[X_i]$  for all  $i < j \in I$ .

### Discrete time case

#### Lemma

A discrete time stochastic process  $(X_n)_{n=1}^{\infty}$  adapted to  $(\mathcal{F}_n)_{n=1}^{\infty}$  satisfies  $\mathbb{E}[X_{n+m}|\mathcal{F}_n]=X_n$  a.s. for all n,m>0 if and only if  $\mathbb{E}[X_{n+1}|\mathcal{F}_n]=X_n$  a.s. for all n>0.

#### Proof.

The forward direction is trivial.

We prove the reverse direction by induction on m > 0. The case m = 1 is given by the assumption. If we have  $\mathbb{E}[X_{n+m}|\mathcal{F}_n] = X_n$ , then by the tower property

$$\mathbb{E}[X_{n+m+1}|\mathcal{F}_n] = \mathbb{E}[\underbrace{\mathbb{E}[X_{n+m+1}|\mathcal{F}_{n+m}]}_{X_{n+m}}|\mathcal{F}_n] = \mathbb{E}[X_{n+m}|\mathcal{F}_n] = X_n.$$



### Example 1: random walk

$$\cdots \underbrace{1/2}_{1/2} -1 \underbrace{1/2}_{1/2} 0 \underbrace{1/2}_{1/2} 1 \underbrace{1/2}_{1/2} 2 \underbrace{1/2}_{1/2} \cdots$$

The random walk with p=1/2 is a martingale: let  $(B_i)_{i=1}^{\infty}$  be a sequence of i.i.d. random variables with  $P(B_i=-1)=P(B_i=1)=1/2$ . Then the random walk can be written as

$$X_n = X_0 + \sum_{i=1}^n B_i.$$

Hence,

$$\mathbb{E}[|X_n|] \le n + \mathbb{E}[|X_0|] < \infty$$
 if  $\mathbb{E}[|X_0|] < \infty$ 

and

$$\mathbb{E}[X_{n+m}|\mathcal{F}_n] = \mathbb{E}[X_n|\mathcal{F}_n] + \sum_{i=n+1}^{n+m} \mathbb{E}[B_i|\mathcal{F}_n] = X_n + \sum_{i=n+1}^{n+m} \mathbb{E}[B_i] = X_n.$$

### Example 2: a multiplicative example

Let 
$$(B_i)_{i=1}^{\infty}$$
 be a sequence of i.i.d. random variables with  $P(B_i = -1) = P(B_i = 1) = 1/2$ . Consider 
$$X_n = \prod_{i=1}^n (1 + B_i) = \begin{cases} 2^n & \text{if } B_1 = B_2 = \dots = B_n = 1\\ 0 & \text{otherwise} \end{cases}$$

We have

$$\mathbb{E}[|X_n|] \leq 2^n < \infty$$

$$\mathbb{E}[|X_n|] \leq \sum_{i=n+1}^{n+m} (1 + \mathbb{E}[B_i|T_n])$$

$$\mathbb{E}[X_{n+m}|\mathcal{F}_n] = X_n \prod_{i=n+1}^{n+m} (1 + \mathbb{E}[B_i]) = X_n.$$

$$\times_{n+m} = \times_n \cdot \prod_{i=n+1}^{n+m} (1 + B_i)$$

### Example 3: compensated Poisson process

Let  $(M_t)_{t\in[0,\infty)}$  be a non-homogeneous Poisson process with rate  $\lambda(t)$ , such that

$$\int_0^t \lambda(s) ds < \infty \quad \text{for all } t \ge 0.$$

Consider

$$W_t = M_t - \int_0^t \lambda(s) ds.$$

The natural filtrations of  $(W_t)_{t\in[0,\infty)}$  and  $(M_t)_{t\in[0,\infty)}$  coincide (their difference is deterministic at any time point). By the triangular property,

$$\mathbb{E}[|W_t|] \leq \mathbb{E}[M_t] + \int_0^t \lambda(s) ds = 2 \int_0^t \lambda(s) ds < \infty.$$

### Example 3: compensated Poisson process

Let  $(M_t)_{t\in[0,\infty)}$  be a non-homogeneous Poisson process with rate  $\lambda(t)$ , such that

$$\int_0^t \lambda(s) ds < \infty \quad \text{for all } t \ge 0.$$

Consider

$$W_t = M_t - \int_0^t \lambda(s) ds.$$

By independence of the increments of  $(M_t)_{t \in [0,\infty)}$ ,

$$\mathbb{E}[W_{t+h}|\mathcal{F}_t] = \mathbb{E}\left[M_{t+h} - \int_0^{t+h} \lambda(s)ds \middle| \mathcal{F}_t\right] - M_{\epsilon} + M_{\epsilon}$$

$$= \mathbb{E}[M_{t+h} - M_t|\mathcal{F}_t] - \int_0^{t+h} \lambda(s)ds + M_t$$

$$= \int_t^{t+h} \lambda(s)ds - \int_0^{t+h} \lambda(s)ds + M_t = M_t - \int_0^t \lambda(s)ds = W_t.$$

# Example 4: a non-Markovian example

Let  $(B_i)_{i=1}^{\infty}$  be a sequence of i.i.d. random variables with  $P(B_i=-1)=P(B_i=1)=1/2$ . Consider  $X_0\sim \text{Pois}(10)$  and

$$X_n = X_{n-1} + B_n \sum_{i=1}^{n-1} X_i \quad n \ge 1.$$

We have

•  $\mathbb{E}[|X_n|] < \infty$ .

We prove it by induction:  $\mathbb{E}[|X_0|] = 10 < \infty$  and if  $\mathbb{E}[|X_i|] < \infty$  for all  $i \le n-1$  then

$$\mathbb{E}[|X_n|] \leq \mathbb{E}[|X_{n-1}|] + \sum_{i=1}^{n-1} \mathbb{E}[|X_i|] < \infty.$$

# Example 4: a non-Markovian example

Let  $(B_i)_{i=1}^{\infty}$  be a sequence of i.i.d. random variables with  $P(B_i=-1)=P(B_i=1)=1/2$ . Consider  $X_0\sim \text{Pois}(10)$  and

$$X_n = X_{n-1} + B_n \sum_{i=1}^{n-1} X_i \quad n \ge 1.$$

We have

• 
$$\mathbb{E}[X_{n+m}|\mathcal{F}_n] = X_n$$
.

Here it is easier to prove

$$\mathbb{E}[X_{n+1}|\mathcal{F}_n] = \mathbb{E}\left[X_n + B_{n+1} \sum_{i=1}^n X_i \middle| \mathcal{F}_n\right]$$
$$= X_n + \mathbb{E}[B_{n+1}] \sum_{i=1}^n X_i = X_n.$$

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# Example 4: a non-Markovian example

Let  $(B_i)_{i=1}^{\infty}$  be a sequence of i.i.d. random variables with  $P(B_i=-1)=P(B_i=1)=1/2$ . Consider  $X_0\sim \text{Pois}(10)$  and

$$X_n = X_{n-1} + B_n \sum_{i=1}^{n-1} X_i \quad n \ge 1.$$

We have

•  $(X_n)_{n=1}^{\infty}$  is not Markovian.

$$P(X_0 = 5, X_2 = 0, X_2 = 6) = 0$$
, hence

$$\times_{z} = X_{1} + \mathcal{B}_{z} \cdot (\times_{0} + \times_{3})$$

$$= 0 + \mathcal{B}_{z} \cdot 5 < \frac{5}{-5} \neq 6$$

$$= 0.$$

 $P(X_2 = 6 | X_1 = 0, X_0 = 5) = 0.$ 

However, it is possible that  $X_1 = 0$  and  $X_2 = 6$  (it can happen if  $X_0 = 6$ ), hence

$$P(X_2 = 6 | X_1 = 0) > 0.$$

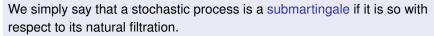
$$\times_{0} = 6$$
  $B_{1} = -1$   
 $\times_{1} = 6 - 6 = 0$   
 $B_{2} = 1$   
 $\times_{1} = 0 + 6 = 6$ 

# Submartingales

#### Definition (Submartingale)

A stochastic process  $(X_i)_{i \in I}$  is a submartingale with respect to a filtration  $(\mathcal{F}_i)_{i \in I}$  if

- $X_i$  is  $\mathcal{F}_i$ -measurable for each  $i \in I$ ;
- $\mathbb{E}[|X_i|] < \infty$  for each  $i \in I$ ;
- $\mathbb{E}[X_j | \mathcal{F}_i] \ge X_i$  a.s. for each  $i < j \in I$ .



Submartingales are favourable games.

By the tower property,  $\mathbb{E}[X_j] = \mathbb{E}[\mathbb{E}[X_j|\mathcal{F}_i]] \ge \mathbb{E}[X_i]$  for all  $i < j \in I$ .



### Supermartingales

### Definition (Supermartingale)

A stochastic process  $(X_i)_{i \in I}$  is a supermartingale with respect to a filtration  $(\mathcal{F}_i)_{i\in I}$  if

- $X_i$  is  $\mathcal{F}_i$ -measurable for each  $i \in I$ ;
- $\mathbb{E}[|X_i|] < \infty$  for each  $i \in I$ ;
- $\mathbb{E}[X_j | \mathcal{F}_i] \leq X_i$  a.s. for each  $i < j \in I$ .



Supermartingales are unfavourable games.

By the tower property, 
$$\mathbb{E}[X_j] = \mathbb{E}[\mathbb{E}[X_j|\mathcal{F}_i]] \leq \mathbb{E}[X_i]$$
 for all  $i < j \in I$ .

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### Example: random walks

$$\cdots \qquad q \qquad -1 \qquad \stackrel{p}{\rightleftharpoons} 0 \qquad \stackrel{p}{\rightleftharpoons} 1 \qquad \stackrel{p}{\rightleftharpoons} 2 \qquad \stackrel{p}{\rightleftharpoons} \cdots$$

The random walk with p < 1/2 is a supermartingale: let  $(B_i)_{i=1}^{\infty}$  be a sequence of i.i.d. random variables with  $P(B_i = -1) = 1 - p$  and  $P(B_i = 1) = p$ . Then the random walk can be written as

$$X_n = X_0 + \sum_{i=1}^n B_i.$$

Hence,

$$\mathbb{E}[|X_n|] \leq n + \mathbb{E}[|X_o|] < \infty \quad \text{if } \mathbb{E}[|X_o|] < \infty$$

and

$$\mathbb{E}[X_{n+m}|\mathcal{F}_n] = X_n + \sum_{i=n+1}^{n+m} \mathbb{E}[B_i] < X_n.$$
  $= \{p-1 < 0\}$ 

Similarly, a random walk with p > 1/2 is a submartingale.

### Example: Poisson process

A Poisson process with rate  $\boldsymbol{\lambda}$  is a submartingale: by the independence of the increments

$$\mathbb{E}[X_{t+h}|\mathcal{F}_t] = \mathbb{E}[X_{t+h} - X_t|\mathcal{F}_t] + X_t = \lambda h + X_t > X_t.$$

# Jensen's inequality

#### Theorem

A function  $f: \mathbb{R} \to \mathbb{R}$  is convex if for any  $x_1, x_2 \in \mathbb{R}$  and  $\alpha \in [0, 1]$ 

$$f(\alpha x_1 + (1-\alpha)x_2) \leq \alpha f(x_1) + (1-\alpha)f(x_2).$$



We say that  $f: \mathbb{R} \to \mathbb{R}$  is concave if for any  $x_1, x_2 \in \mathbb{R}$  and  $\alpha \in [0, 1]$ 

$$f(\alpha x_1 + (1-\alpha)x_2) \ge \alpha f(x_1) + (1-\alpha)f(x_2).$$



### Theorem (Jensen's inequality)

If f is convex and X is a real, Fineasurable random variable with 9 = F  $\mathbb{E}[|X|], \mathbb{E}[|f(X)|] < \infty$ , then almost surely

$$f(\mathbb{E}[X|\mathcal{G}]) \leq \mathbb{E}[f(X)|\mathcal{G}].$$

#### Theorem

Let  $(X_i)_{i\in I}$  be a real-valued process, and let  $f: \mathbb{R} \to \mathbb{R}$  be such that  $\mathbb{E}[f(X_i)] < \infty$  for all  $i \in I$ .

- If  $(X_i)_{i \in I}$  is a submartingale w.r.t.  $(\mathcal{F}_i)_{i \in I}$  and f is convex and non-decreasing, then  $(f(X_i))_{i\in I}$  is a submartingale w.r.t.  $(\mathcal{F}_i)_{i\in I}$ ;
- If  $(X_i)_{i \in I}$  is a supermartingale w.r.t.  $(\mathcal{F}_i)_{i \in I}$  and f is concave and non-decreasing, then  $(f(X_i))_{i\in I}$  is a supermartingale w.r.t.  $(\mathcal{F}_i)_{i\in I}$ ;
- If  $(X_i)_{i \in I}$  is a martingale w.r.t.  $(\mathcal{F}_i)_{i \in I}$  and f is convex, then  $(f(X_i))_{i \in I}$  is a submartingale w.r.t.  $(\mathcal{F}_i)_{i \in I}$ ;
- If  $(X_i)_{i \in I}$  is a martingale w.r.t.  $(\mathcal{F}_i)_{i \in I}$  and f is concave, then  $(f(X_i))_{i \in I}$  is a supermartingale w.r.t.  $(\mathcal{F}_i)_{i \in I}$ .

### Jensen's inequality

#### Proof.

If  $(X_i)_{i \in I}$  is a submartingale w.r.t.  $(\mathcal{F}_i)_{i \in I}$  and f is convex and non-decreasing, then by Jensen's inequality

equality 
$$\mathbb{E}[f(X_{i+j})|\mathcal{F}_i] \geq f(\mathbb{E}[X_{i+j}|\mathcal{F}_i]) \geq f(X_i). \quad -\triangleright \quad (8(x_i)).$$

If  $(X_i)_{i\in I}$  is a supermartingale w.r.t.  $(\mathcal{F}_i)_{i\in I}$  and f is concave and non-decreasing, then -f is convex and non-increasing, so by Jensen's inequality

$$-\mathbb{E}[f(X_{i+j})|\mathcal{F}_i] \geq -f(\mathbb{E}[X_{i+j}|\mathcal{F}_i]) \geq -f(X_i),$$

which implies  $\mathbb{E}[f(X_{i+j})|\mathcal{F}_i] \leq f(X_i)$ .

### Jensen's inequality

#### Proof.

If  $(X_i)_{i\in I}$  is a martingale w.r.t.  $(\mathcal{F}_i)_{i\in I}$  and f is convex, then by Jensen's inequality

$$\mathbb{E}[f(X_{i+j})|\mathcal{F}_i] \stackrel{\text{def}}{\stackrel{\text{def}}{=}} f(\mathbb{E}[X_{i+j}|\mathcal{F}_i]) = f(X_i). \qquad \frac{\left( \ \S(\times_i) \right)_{i \in \mathcal{I}}}{\text{Subhart in Galce}}$$

If  $(X_i)_{i\in I}$  is a supermartingale w.r.t.  $(\mathcal{F}_i)_{i\in I}$  and f is concave, then -f is convex so by Jensen's inequality

$$-\mathbb{E}[f(X_{i+j})|\mathcal{F}_i] \geq -f(\mathbb{E}[X_{i+j}|\mathcal{F}_i]) = -f(X_i),$$

which implies  $\mathbb{E}[f(X_{i+j})|\mathcal{F}_i] \leq f(X_i)$ .

### Examples

Let  $(X_n)_{n=1}^{\infty}$  be the random walk with p=1/2. Then,  $(X_n)_{n=1}^{\infty}$  is a martingale. We have that

- $(X_n^2)_{n=1}^{\infty}$  and  $(|X_n|)_{n=1}^{\infty}$  are submartingales;
- $(-X_n^4 + X_n + 7)_{n=1}^{\infty}$  and  $(-|X_n|)_{n=1}^{\infty}$  are supermartingales.

### Examples

Let  $(B_i)_{i=1}^{\infty}$  be a sequence of i.i.d. random variables with  $P(B_i = -1) = P(B_i = 1) = 1/2$ . Consider

$$X_n = \prod_{i=1}^n (1 + B_i) = \begin{cases} 2^n & \text{if } B_1 = B_2 = \dots = B_n = 1\\ 0 & \text{otherwise} \end{cases}$$

Then

- $((X_n+1)^{-1})_{n=1}^{\infty}$  is a submartingale;
- $(\sqrt{X_n})_{n=1}^{\infty}$  and  $(\log(X_n+1))_{n=1}^{\infty}$  are supermartingales.