Definition

Let $\{X_i\}_{i\in I}$ be a renewal process with inter-arrival times $\{t_\ell\}_{\ell=1}^\infty$ and rewards $\{r_\ell\}_{\ell=1}^\infty$. If the random variables $\{(t_\ell,r_\ell)\}_{\ell=1}^\infty$ are i.i.d. then we have a renewal-reward process.

Theorem

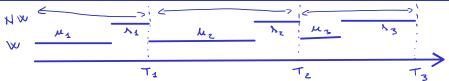
Let $\{X_i\}_{i\in I}$ be a renewal process with $E[t_1]>0$, and $P(t_1<\infty)=1$. Then, with probability 1,

$$\lim_{i\to\infty}\frac{R_i}{i}=\frac{\underline{E}[r_1]}{E[t_1]},$$

where

$$R_i = \sum_{\ell=1}^{X_i} r_{\ell}$$

is the cumulative reward up to time i.



Assume a machine create a revenue of 10 euros per day when it works, repairing it costs 50 euros of transportation plus a repairing cost of approximately 40 euros per day. The machine does not break down for times with mean 30 days and repairing it takes on average 5 days. Assume that when the machine is repaired, it can be considered as new.

After 10 years of usage, how much revenue per day has the machine produced?

M: = tempo la voco i-ESIHO INT. S: = tempo tipaxa Zione

$$\frac{R_{40.365}}{40.365} \approx \lim_{i \to +\infty} \frac{R_{i}}{i} = \frac{E[t_{4}]}{E[t_{4}]} = \frac{E[10.\mu_{4} - 50 - 40.\lambda_{4}]}{E[\mu_{4} + \lambda_{4}]}$$

$$= \frac{10.30 - 50 - 40.5}{30 + 5}$$

$$= \frac{300 - 50 - 200}{35} = \frac{50}{35} E/aims$$

In the previous question, how does the answer change if repair times are geometric random variables, and times until break-downs are Poisson?

Suppose the manager of a restaurant should pay $\underline{\$c_1}$ per day for health and safety maintenance, enough to guarantee that there will be no health violations. However, the manager is evil and wants to budget less, say $\underline{c_2} < c_1$. We want to figure out when this will make him money. Here is what we know:

- The restaurant is inspected, on average, every 45 days, and the number of days between two consecutive inspections are independent and identically distributed random variables (maybe not realistic).
- There is a probability $p = p(c_2)$ that a violation will be found on a given visit. It is a monotonically decreasing function of c_2 , with $p(c_1) = 0$.
- The fines have an expected value of $c_3 > 0$ and the fine sizes are all independent and identically distributed, and independent on the inspection times.

$$M_{4}=0$$
 $M_{2}>0$ $M_{3}>0$ $|M_{4}=0$ $|M_{5}=0$
 $+_{3}$ $|M_{5}=0$
 $+_{4}$ $|M_{5}=0$
 $+_{5}$ $|M_{5}=0$
 $|M_{5}=0$

$$\lim_{\hat{i}\to b+\infty} \frac{R_i}{\hat{i}} = \frac{E[\lambda_i]}{E[+,1]}$$

RENEWAL-REWARD PROCESS

$$E[Y_{4}] = -C_{2} \cdot E[Y_{4}] - E[Y_{4}] = -C_{2} \cdot 45 - p(C_{2}) \cdot C_{3}$$

$$\lim_{z \to z} \frac{R_{1}}{z} = \frac{-C_{2} \cdot 45 - p(C_{2}) \cdot C_{3}}{45} = -C_{2} - \frac{p(C_{2}) \cdot C_{3}}{45}$$

$$Z_i = \# GioRni$$

 $\widetilde{C}_i = -C_z - \widetilde{\widetilde{M}}_i$

ENTRO : = i

R.R. PROCESS?

Assume

$$p(c_2) = \frac{1+c_1}{c_1} \cdot \frac{1}{1+c_2} - \frac{1}{c_1}$$

We want to minimize

$$g(c_2) = c_2 + c_3 \frac{p(c_2)}{45} = c_2 + \frac{c_3}{45} \left(\frac{1+c_1}{c_1} \cdot \frac{1}{1+c_2} - \frac{1}{c_1} \right).$$

We consider

$$g'(c_2) = 1 - \frac{c_3}{45} \cdot \frac{1+c_1}{c_1} \cdot \frac{1}{(1+c_2)^2}$$

and note that

$$g''(c_2) = \frac{c_3}{45} \cdot \frac{1+c_1}{c_1} \cdot \frac{2}{(1+c_2)^3}$$

is positive for any c_2 in $[0, c_1]$.

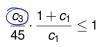


Solving $g'(c_2^*) = 0$ yields

$$c_2^* = \sqrt{\frac{c_3}{45} \cdot \frac{1+c_1}{c_1}} - 1.$$

We have three cases:

- if c_2^* is between 0 and c_1 , then it corresponds to the optimal policy the evil manager is looking for;
- if



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then the fee to pay for the violation is so low that the optimal strategy is to pay nothing for maintenance;

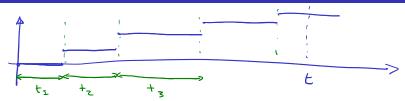
if

$$\frac{c_3}{45} \cdot \frac{1+c_1}{c_1} \ge (1+c_1)^2$$

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then the fee is adequate, and the optimal policy is to pay c_1 dollars for maintenance.

Poisson processes



Definition

A Poisson process with rate λ is a counting process $(N(s))_{s\in[0,\infty)}$ with N(0)=0, whose inter-arrival times are i.i.d. exponential random variables with rate λ .

$$t_s, t_z, t_3, \dots$$
 i.i.d
 $\sim \exp(\lambda)$

A review of the Exponential distribution

2 RATE

We say that $\tau \sim \text{exp}(\lambda)$ if any of the following holds:



$$f_{\tau}(t) = \begin{cases} \lambda e^{-\lambda t} & \text{if } t \ge 0 \\ 0 & \text{if } t < 0 \end{cases} \qquad \boxed{\mathsf{T}(t)} = P(\tau \le t) = \begin{cases} 1 - e^{-\lambda t} & \text{if } t \ge 0 \\ 0 & \text{if } t < 0 \end{cases}$$

$$P(\tau > t) = \begin{cases} e^{-\lambda t} & \text{if } t \ge 0 \\ 1 & \text{if } t < 0 \end{cases} = 4 - \boxed{\mathsf{T}(t)}$$

$$E[\tau] = 1/\lambda$$
, $Var(\tau) = 1/\lambda^2$

Memoryless property

2 BURATA

For any
$$t \not > s \ge 0$$
 we have $P(\tau > t + s | \tau > s) = P(\tau > t)$

$$P(\tau > t + s | \tau > s) = \frac{P(\tau > t + s)}{P(\tau > s)} = \frac{P(\tau > t + s)}{P(\tau > s)} = \frac{e^{-\lambda(t + s)}}{e^{-\lambda s}} = e^{-\lambda t}$$

Exponential random variables are the **only** continuous random variables with the memoryless property (geometric random variables are the only discrete random variables with the memoryless property).