

Machine Learning-Based Estimation of Monthly GDP*

Yonggeun Jung[†]

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Abstract

This paper proposes a scalable framework to estimate monthly GDP using machine learning methods. We apply Multi-Layer Perceptron (MLP), Long Short-Term Memory networks (LSTM), Extreme Gradient Boosting (XGBoost), and Elastic Net regression to map monthly indicators to quarterly GDP growth, and reconcile the outputs with actual aggregates. Using data from China, Germany, the UK, and the US, our method delivers robust performance across varied data environments. Benchmark comparisons with prior US studies and UK official statistics validate its accuracy. We also explore nighttime light as a proxy, finding its usefulness varies by economic structure. The approach offers a flexible and data-driven tool for high-frequency macroeconomic monitoring and policy analysis.

JEL Codes: C53, C55, E01

Keywords: Monthly GDP, Machine Learning, Temporal Disaggregation.

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[†]Ph.D. Student in Economics. Department of Economics, University of Florida, PO Box 117140, Gainesville, Florida, USA, 32611. Email: yonggeun.jung@ufl.edu.

1 Introduction

Real gross domestic product (GDP) is a core economic indicator, widely valued for its intuitive and comprehensive representation of economic activity within a given region. While GDP is traditionally published at a quarterly frequency, the increasing demand for timely and granular economic insights has led to growing interest in both nowcasting and high-frequency estimation (Chu & Qureshi, 2023; Ghosh & Ranjan, 2023; Giannone et al., 2008; Jansen et al., 2016; Mitchell et al., 2005; Richardson et al., 2021).

In particular, the expanding availability of high-frequency data across various domains has intensified efforts to estimate monthly GDP (Brave et al., 2019; Koop et al., 2021, 2023; Mariano & Murasawa, 2003, 2010; Mitchell et al., 2005). Monthly estimates may offer distinct advantages. For policymakers, they provide a more comprehensive snapshot of the economy, enabling more agile and precise interventions, particularly during periods of rapid change.¹ For financial market participants, they can serve as a valuable benchmark that reduces uncertainty by synthesizing the noise from various individual indicators. Furthermore, they allow researchers to capture short-run dynamics with greater precision, such as identifying economic turning points more accurately and analyzing the immediate effects of policy shocks.²

To address this gap, Koop et al. (2023) developed a Bayesian Mixed-Frequency Vector Auto-Regressive (MF-VAR) model that integrates quarterly GDP with a range of monthly indicators to estimate monthly GDP in the US. Their approach offers high accuracy in identification and inference, particularly under a well-specified prior. However, its Bayesian structure can pose implementation challenges, especially when applied to countries with limited prior information.

In response to these limitations, this study proposes an alternative framework based on machine learning techniques.³ Our objective is to develop a scalable method for monthly GDP estimation that is applicable across a wide range of country contexts without relying on strong prior assumptions. Specifically, we employ four dis-

¹Koop et al. (2023) analyze detailed dynamics—such as the timing of economic troughs and the speed of recovery—that cannot be observed using quarterly GDP data. The UK’s Office for National Statistics (ONS) evaluates the effects of policy changes such as lockdowns and fiscal support, as well as the recovery patterns across industries, using monthly GDP data (Scruton et al., 2018).

²Since monthly GDP is often difficult to observe directly, practitioners may rely on proxy indicators. However, such proxies can be biased depending on the structural characteristics of the economy.

³For a discussion on the usefulness of the underlying features driving the gains from machine learning over standard macroeconometric methods, see Goulet Coulombe et al. (2022).

tinct classes of models—a Multi-Layer Perceptron (MLP), a Long Short-Term Memory (LSTM) network, XGBoost, and Elastic Net—to learn the relationship between quarterly GDP growth and monthly economic indicators. The MLP is well-suited for capturing nonlinear patterns, while the LSTM leverages its ability to retain information from sequential observations. In contrast to these neural network-based approaches, XGBoost and Elastic Net provide greater interpretability with respect to the roles of explanatory variables. Furthermore, we explore the potential of incorporating nighttime light (NTL) data as an additional input for estimating monthly GDP.

After a quarterly model is trained, it is applied to the monthly indicators to generate a preliminary high-frequency signal of GDP growth. This signal is then adjusted using the Denton (1971) proportional benchmarking method to ensure that the resulting monthly estimates are consistent with the observed quarterly GDP.⁴ In doing so, the approach preserves intra-quarter dynamics implied by the machine learning model while maintaining consistency with national accounts.

This paper makes several key contributions. First, it offers a scalable and generalizable framework. The proposed method is successfully applied to four countries with diverse data environments—China, Germany, the UK, and the US—demonstrating its adaptability beyond a single-country context. Second, it adopts a data-driven modeling approach. By leveraging deep learning models capable of capturing complex nonlinear relationships, the framework allows the data to uncover hidden patterns without relying on rigid structural assumptions.⁵ Third, the framework addresses interpretability. Recognizing the “black box” limitations of neural networks, the study also incorporates XGBoost and Elastic Net—models that provide transparency through variable importance scores and coefficient estimates. This provides interpretive insights into which indicators are most associated with GDP growth, and allows cross-validation of results across different model classes. Finally, the paper provides empirical validation. The proposed methodology is rigorously benchmarked against monthly estimates from Koop et al. (2023) for the US and the official monthly GDP

⁴However, the application of such methods inherently involves adjusting model outputs using already published quarterly GDP figures, which limits their usefulness for another important task: nowcasting.

⁵See Chow and Lin (1971) and Litterman (1983) for traditional temporal disaggregation methods: Chow and Lin (1971) propose a regression-based approach, while Litterman (1983) uses a state-space variant. These methods typically rely on auxiliary regressors and impose dynamic structures to interpolate or extrapolate monthly values. While powerful, they require strong assumptions about variable relationships and are not directly applicable for reconciling model-generated high-frequency signals with official aggregates.

series from the UK’s Office for National Statistics (ONS), reinforcing the credibility of the results.

The remainder of the paper is structured as follows. Section 2 outlines the methodological framework, including model architecture and data preprocessing procedures. Section 3 presents the estimation results, focusing on out-of-sample predictive performance and benchmarking against existing monthly GDP series for the US and UK. It also discusses model interpretability using XGBoost and Elastic Net, highlighting cross-country insights and robustness checks. Section 4 explores additional estimation exercises incorporating NTL data. Section 5 concludes and outlines directions for future research. Supplementary information, including data descriptions, algorithmic details, reconciliation methods, and additional figures, is provided in the appendices.

2 Methodology

2.1 Model

This section describes the model⁶ employed to estimate monthly GDP. Let Y_q denote the observed quarterly real GDP growth for quarter q , and let $X_m \in \mathbb{R}^k$ represent the vector of k explanatory variables observed in month m . Typical explanatory variables include industrial production, retail sales, the unemployment rate, consumer sentiment, and financial market indicators.⁷

To relate these high-frequency monthly indicators to the observed quarterly GDP growth, we first aggregate the monthly variables to construct quarterly-level variables $X_q \in \mathbb{R}^k$. For each variable j , the quarterly value $X_q^{(j)}$ is constructed as follows:

(i) For variables measured in levels (e.g., the unemployment rate), we compute the quarterly average:

$$X_q^{(j)} = \frac{1}{3} \sum_{m \in q} X_m^{(j)}. \quad (1)$$

(ii) For variables expressed as growth rates (e.g., first differences of logarithms),

⁶The algorithms used to implement this model are detailed in Appendix C.

⁷For example, Koop et al. (2023) use Average Weekly Hours: Manufacturing, CPI, Industrial Production, Real Personal Consumption Expenditures, the Federal Funds Rate, the 10-Year Treasury Rate, the S&P Stock Price Index, and the Civilian Unemployment Rate.

we compute the quarterly sum, consistent with the cumulative growth identity:

$$X_q^{(j)} = \sum_{m \in q} X_m^{(j)}. \quad (2)$$

In practice, we first determine which variables require log-differencing based on the Augmented Dickey-Fuller (ADF) test (Said & Dickey, 1984). After constructing the differenced series and aggregating all variables to a quarterly frequency, we split the dataset into training and test sets. To prevent data leakage, we standardize the non-log-differenced level variables only after this split.⁸ Specifically, we apply Z-score standardization.⁹

We specify a general regression framework to model the relationship between quarterly GDP growth and the aggregated explanatory variables:

$$Y_q = f(X_q; \theta) + e_q, \quad (3)$$

where $f(\cdot)$ denotes the potentially nonlinear function to be estimated, θ represents the set of model parameters, and e_q is the error term. To approximate the unknown function f , we consider four distinct classes of machine learning models that vary in flexibility, structure, and interpretability.

First, we employ a **Multi-Layer Perceptron (MLP)**, a class of feedforward artificial neural networks capable of capturing complex nonlinear relationships (Rumelhart et al., 1986). Second, to account for the time-series nature of the data, we utilize a **Long Short-Term Memory (LSTM)** network, a type of recurrent neural network (RNN) designed to capture long-range dependencies through gated memory cells (Hochreiter & Schmidhuber, 1997). Third, we include **XGBoost** (Extreme Gradient Boosting), a high-performance implementation of the gradient boosting framework that combines an ensemble of decision trees (T. Chen & Guestrin, 2016). In addition to its strong predictive accuracy, XGBoost provides variable importance scores, which help alleviate the interpretability limitations of neural networks and allow for cross-model robustness checks. Finally, as a linear benchmark, we consider the **Elastic**

⁸Standardizing the entire dataset before splitting could allow information from the test set—such as its mean and standard deviation—to leak into the training process.

⁹Each variable X_j is transformed into z_j as follows:

$$z_j = \frac{X_j - \mu_j}{\sigma_j},$$

where μ_j and σ_j denote the sample mean and standard deviation of variable X_j .

Net, a regularized linear regression model that combines LASSO (λ_1) and Ridge (λ_2) penalties (Zou & Hastie, 2005). Its linear structure enables direct interpretation of each explanatory variable’s contribution.^{10 11}

Once the quarterly model $f(\cdot; \hat{\theta})$ is estimated, we generate preliminary monthly estimates by applying it to the original monthly covariate vectors X_m :

$$\tilde{Y}_m = f(X_m; \hat{\theta}). \quad (4)$$

Importantly, since the model is trained on quarterly data, each \tilde{Y}_m reflects a prediction of the *total quarterly* growth rate, conditional on the economic environment observed in month m . The resulting series $\{\tilde{Y}_m\}$ thus serves as a high-frequency, model-based signal that captures intra-quarter variation as inferred from monthly movements in the explanatory variables.

Because this series does not, by construction, sum to the observed quarterly total, we apply a proportional benchmarking adjustment, following the approach of Denton (1971).¹² For each quarter q , we compute an adjustment factor k_q , defined as the ratio of the actual quarterly growth rate to the sum of the preliminary monthly estimates within that quarter:

$$k_q = \frac{Y_q}{\sum_{j \in q} \tilde{Y}_j}. \quad (5)$$

We then obtain the adjusted monthly growth rates, \hat{y}_m , by scaling each \tilde{Y}_m using the corresponding adjustment factor:¹³

$$\hat{y}_m = k_q \cdot \tilde{Y}_m. \quad (6)$$

This procedure ensures that the adjusted monthly estimates are consistent with the observed quarterly aggregates—i.e., $\sum_{m \in q} \hat{y}_m = Y_q$ —while preserving the intra-quarter variation implied by the model. However, since the model output reflects a

¹⁰LASSO (Least Absolute Shrinkage and Selection Operator) regularization (λ_1 penalty) performs variable selection by shrinking some coefficients to zero, while Ridge regularization (λ_2 penalty) shrinks coefficients continuously without eliminating any (James et al., 2023).

¹¹Detailed model specifications are provided in Appendix B.

¹²See technical details in Appendix D.

¹³In our implementation, we first derive a naive monthly estimate (\hat{p}_m) by dividing the preliminary quarterly-scale estimate (\tilde{Y}_m) by 3. This scaling factor, however, mathematically cancels out in the final adjustment formula, yielding the same result as the more concise expression presented in Equations (5) and (6). For the detailed algorithm, please see Appendix C.

projection of total quarterly growth conditional on monthly information, rather than a direct estimate of each month’s contribution, the resulting high-frequency series should be interpreted as a model-based, internally consistent approximation.

2.2 Data

For the empirical analysis, we collect data for four countries—China, Germany, the UK, and the US. While our target sample spans from 1991 to 2024, the actual coverage varies by country depending on the availability of data. In all cases, we ensure that the quarterly GDP series and all explanatory variables are temporally aligned within a consistent time frame for each country.

Quarterly real GDP growth is calculated as the log difference of seasonally adjusted real GDP. The monthly explanatory variables cover various dimensions of the economy—including consumption, production, investment, capital markets, labor, and trade—subject to each country’s data availability. These variables are collected from central banks, capital market authorities, national statistical offices, and other third-party providers.

While the specific variables—and even the presence of certain categories—vary across countries, we select variables flexibly within and across categories to maximize relevance and comparability. In some cases, certain economic dimensions (e.g., investment or trade) may be missing or sparsely represented, while others may be relatively rich in data. Table 1 summarizes the datasets used for each country, with detailed descriptions provided in Appendix A.

For the US, monthly GDP growth estimates produced by Koop et al. (2023) are available,¹⁴ while the UK’s ONS publishes official monthly GDP figures.¹⁵ These two cases allow us to benchmark the performance of our framework. For the remaining countries, we demonstrate that the model performs robustly even when the set of available indicators differs.

¹⁴Koop et al. (2023) provide both the replication package and data, which are available on the journal’s website: <https://www.tandfonline.com/doi/full/10.1080/07350015.2022.2044336>.

¹⁵Available at <https://www.ons.gov.uk/economy/grossdomesticproductgdp>.

Table 1. Summary of Datasets (Sources)

Variable	China	Germany	UK	US
CPI	FRED	FRED	FRED	FRED
Retail Trade		FRED		
M1 Money Stock			BOE	FRED
M2 Money Stock				FRED
Total Reserves	FRED	FRED	FRED	
Exchange Rate	FRED		FRED	
Unemployment Rate		FRED	ONS	FRED
Employment Rate			ONS	
Labor Force				FRED
Avg. Weekly Hours				FRED
Industrial Production				FRED
Production Volume		FRED	FRED	
Manufacturing		FRED		
PPI	FRED			
Real Personal Consum.				FRED
Interest Rate (10Y Gov)			FRED	
Interest Rate (Gov Bond)			FRED	
Interest Rate (10Y Treas.)				FRED
Effect. Fund Rate				FRED
Effect. Ex. Rate	FRED	FRED		
S&P/ FTSE / DAX / SSEC	IV	YF	YF	YF
Share Prices			FRED	
Moody's AAA Bond				FRED
(Net) Exports	FRED	FRED	FRED	FRED
Imports	FRED	FRED	FRED	
US Imports from China	FRED			
Price Competitiv.		DB		
Policy Uncertainty	FRED	FRED		
Personal Savings Rate				FRED
Nonfarm Payroll Emp.				FRED
Real GDP (Quarterly)	FRED	FRED	FRED	FRED
Total Variables Used	11	13	14	16
Data Span	01.94-09.23	01.91-12.23	01.91-03.24	01.92-12.24

Note: The entries in the table indicate the primary data source used to fetch each variable. FRED: Federal Reserve Economic Data (<https://fred.stlouisfed.org/>); YF: Yahoo Finance (<https://finance.yahoo.com/>); ONS: Office for National Statistics (<https://www.ons.gov.uk/>); BOE: Bank of England (<https://www.bankofengland.co.uk/boeapps/database/>); DB: Deutsche Bundesbank (<https://www.bundesbank.de/en>); IV: Investing.com (<https://www.investing.com/>). They may serve as a data platform; the original source of each series may differ by variable.

Table 2. Test Performance

Country	Model	MSE	RMSE	MAE	R^2	Sign Acc.	sMAPE	Theil U1
China	MLP	0.006	0.077	0.060	0.628	0.875	78.218	0.335
	LSTM	0.001	0.033	0.026	0.936	0.950	31.807	0.132
	XGB	0.011	0.107	0.084	0.289	0.750	99.990	0.542
	ENET	0.011	0.106	0.091	0.299	0.750	134.608	0.636
Germany	MLP	0.000	0.018	0.011	0.507	0.667	120.109	0.425
	LSTM	0.000	0.021	0.014	0.432	0.435	162.949	0.536
	XGB	0.001	0.024	0.012	0.121	0.741	104.188	0.740
	ENET	0.001	0.025	0.013	0.041	0.667	123.126	0.724
UK	MLP	0.001	0.030	0.017	0.712	0.815	115.059	0.314
	LSTM	0.003	0.057	0.025	0.086	0.696	134.470	0.852
	XGB	0.003	0.053	0.021	0.072	0.741	114.271	0.858
	ENET	0.003	0.054	0.022	0.030	0.556	137.120	0.900
US	MLP	0.000	0.009	0.008	0.826	0.889	79.447	0.203
	LSTM	0.000	0.017	0.011	0.529	0.739	95.018	0.427
	XGB	0.000	0.021	0.009	0.165	0.926	67.349	0.668
	ENET	0.000	0.014	0.008	0.600	0.889	102.448	0.391

Note: Mean Squared Error (MSE), Root Mean Squared Error (RMSE), Mean Absolute Error (MAE), and symmetric Mean Absolute Percentage Error (sMAPE, %) are all scale-dependent error metrics where lower values indicate better predictive accuracy. The coefficient of determination (R^2) measures the proportion of variance explained by the model and is better when closer to one. Sign Accuracy indicates how often the model correctly predicts the direction of change and is also better when higher. Theil’s U1 statistic compares the forecast accuracy of a model to that of a naive (no-change) forecast; values closer to zero indicate better performance.

3 Estimation Results

3.1 Estimated Monthly GDP by Country

As discussed in Section 2.1, a key starting point of our framework is its ability to accurately predict quarterly GDP growth, \tilde{Y}_q , from the transformed quarterly covariates X_q . Table 2 summarizes the test performance of each model across the four countries, using a hold-out test set. Notably, the test period includes the COVID-19 era (2020–2022), enabling an evaluation of how well each model captures large, unexpected economic shocks.¹⁶

Model performance is assessed using multiple evaluation metrics: Mean Squared Error (MSE), Root Mean Squared Error (RMSE), Mean Absolute Error (MAE), R^2 ,

¹⁶See Appendix H for time series plots by country.

Sign Accuracy, Symmetric Mean Absolute Percentage Error (sMAPE), and Theil’s U1. Lower values indicate better performance for MSE, RMSE, MAE, sMAPE, and Theil’s U1, while higher values are preferred for R^2 and Sign Accuracy.

Among the models, the MLP delivers the most accurate point forecasts overall, generally yielding the lowest MSE and highest R^2 , though the LSTM model shows superior performance for China. XGBoost, while exhibiting higher errors, performs relatively well in predicting the direction of change (high sign accuracy).

LSTM, despite its theoretical advantage in modeling sequential dependencies, underperforms in this setting, except for China. One strength of LSTM is its ability to learn from long sequences of past observations, potentially capturing persistent dynamics. However, in this application, the model must be trained on quarterly GDP targets using sequences of monthly covariates. As the input sequence length t increases, information loss becomes more likely or significant, which introduces a trade-off between capturing long-run dependencies and preserving data quality. In contrast, MLPs are not affected by this sequence alignment issue and thus appear more suitable for the structure and size of the current dataset.

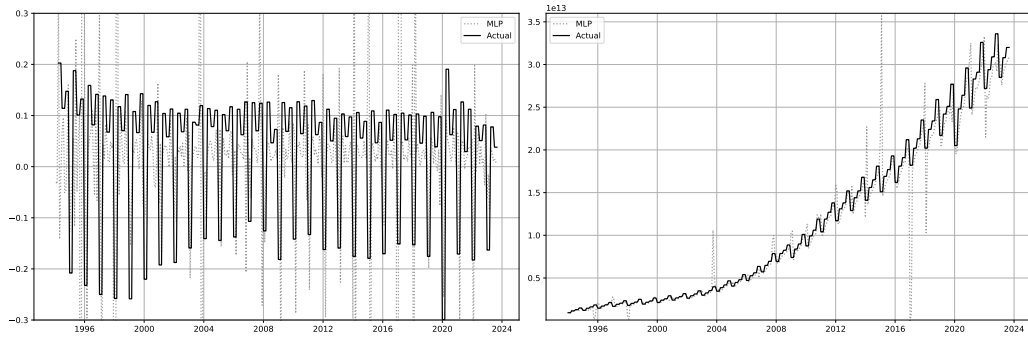
To investigate this trade-off further, we conducted a simulation study with synthetic data. The simulation revealed a nuanced dynamic: while LSTM excels in idealized, trend-free environments, the MLP demonstrated superior robustness in a more realistic scenario involving co-trending variables with complex lagged relationships. This suggests that the MLP’s strong empirical performance stems from its ability to capture the robust, low-frequency relationships that persist after temporal aggregation damages more granular sequential information.¹⁷

Across all countries, the MLP model demonstrates relatively stable and robust performance, successfully capturing general trends in monthly GDP—even during periods of economic shocks. In contrast, alternative models such as XGBoost and Elastic Net tend to produce larger outliers during shock periods, particularly in countries where fewer explanatory variables are available.¹⁸ Therefore, while the MLP is preferred for accurate prediction, models like XGBoost or Elastic Net may offer additional insights when interpretability or variable importance is of interest. While they may not outperform MLPs in terms of prediction accuracy, their transparency can help uncover which indicators drive model predictions.

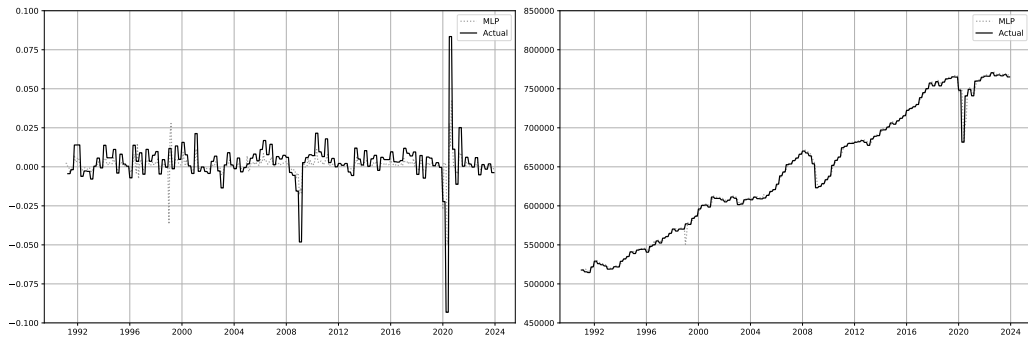
Finally, the Elastic Net model performs the worst in terms of point forecast-

¹⁷A detailed discussion of the simulation design and results is provided in Appendix E.

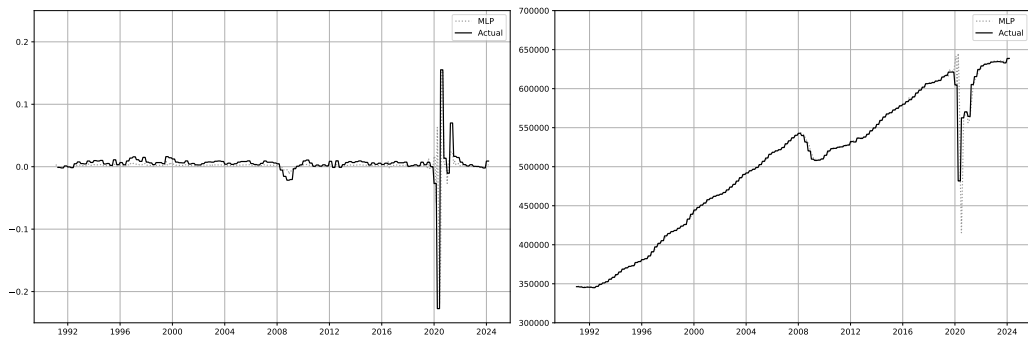
¹⁸See Appendix H.



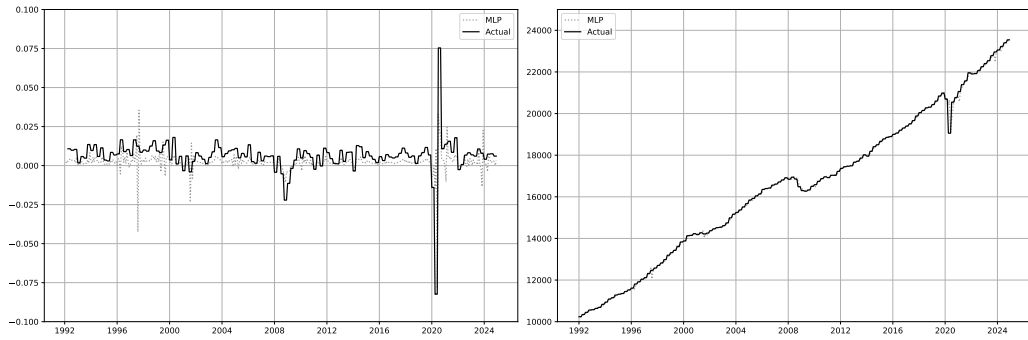
(a) China



(b) Germany



(c) UK



(d) US

Figure 1. Best-Performance Models (MLP): Growth (L) and Level (R)

ing accuracy, likely due to its inherent linearity and limited flexibility in capturing nonlinear relationships. These limitations become more pronounced in some countries—especially China—where the lack of rich data further constrains model performance.

For China, the LSTM model yields the best out-of-sample performance in the quarterly prediction task, as shown in Table 2, an exceptional result given the MLP’s dominance in other countries. This finding can be interpreted through our simulation study in Appendix E, which demonstrates that LSTM’s capacity to learn sequential patterns gives it a distinct advantage in idealized data environments with clear and persistent trends. The LSTM’s success may therefore suggest that China’s monthly indicators provide a cleaner, more stable signal compared to the more complex, cointrending data environments of the other economies we study. However, a separate issue arises when we inspect the disaggregated monthly estimates. The monthly series produced by the LSTM, while consistent with its accurate quarterly predictions, tends to exhibit higher volatility compared to the series from the MLP. This higher monthly volatility is unlikely a result of simple overfitting, given the LSTM’s strong quarterly test performance. Instead, it may point to other factors, such as limitations in the quality of the underlying monthly data or, more fundamentally, concerns about the reliability of the official GDP statistics used for the final reconciliation (Holz, 2014; Rawski, 2001). To investigate this latter possibility, we re-estimate the MLP model using the OECD’s Composite Leading Indicator (CLI) as an alternative target. This analysis, detailed in Appendix F, reveals that the model’s test performance improves noticeably when predicting the CLI instead of the official GDP, lending further support to the hypothesis that the quality of the official GDP series may be a contributing factor.

3.1.1 US Case: Koop et al. (2023)

Koop et al. (2023) construct monthly GDP estimates by embedding a Bayesian MF-VAR model that incorporates quarterly expenditure-side (GDP_E) and income-side (GDP_I) GDP, a latent true GDP process, and multiple monthly indicators of economic activity. The model treats GDP_E and GDP_I as noisy observations of true GDP and enforces temporal aggregation constraints to reconcile monthly and quarterly data. Identification is achieved using either instrumental variables (e.g., unemployment) or set-identification via variance bounds.

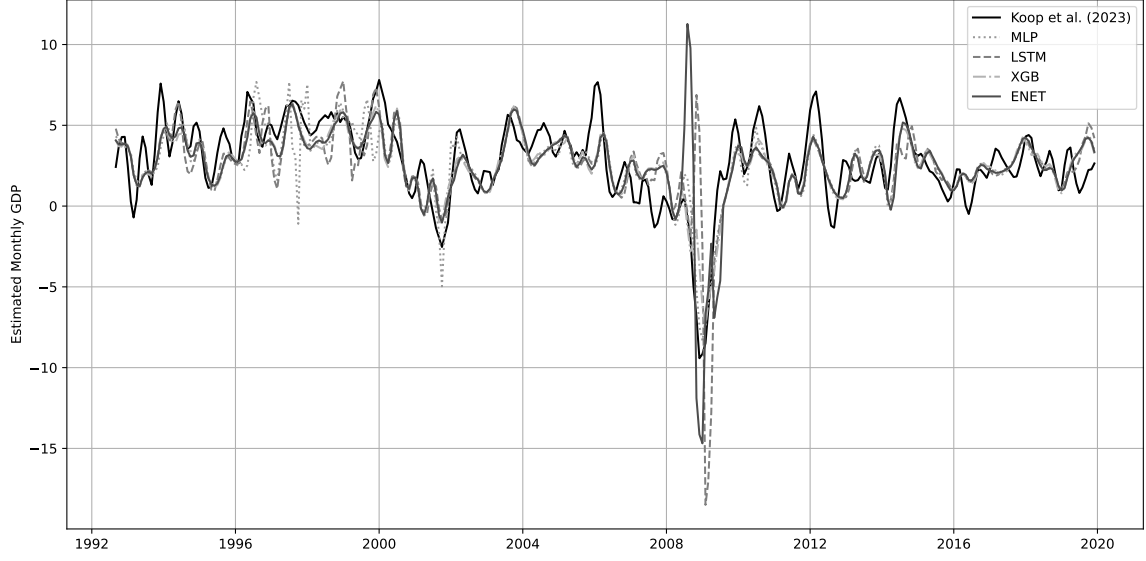


Figure 2. Benchmark to Koop et al. (2023): US Estimated Monthly GDP

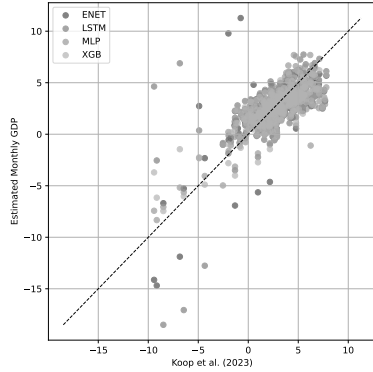
Following Koop et al. (2023)—and building on the methodology of Mariano and Murasawa (2003, 2010) and Mitchell et al. (2005)—we transform our model’s final monthly growth rate \hat{y}_m into an annualized rate for comparison using the same five-month weighted average formula:

$$\hat{y}_{m,t}^A = \left(\frac{1}{3}\hat{y}_{m,t} + \frac{2}{3}\hat{y}_{m,t-1} + \hat{y}_{m,t-2} + \frac{2}{3}\hat{y}_{m,t-3} + \frac{1}{3}\hat{y}_{m,t-4} \right) \times 4 \times 100 \quad (7)$$

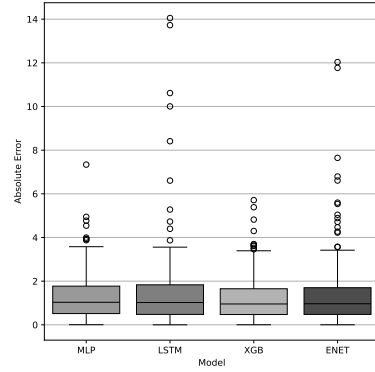
While Koop et al. (2023) provide estimates from 1960 to 2019, our estimation spans from 1992 to 2024. For comparability, we focus on the overlapping period from 1992 to 2019. As shown in Figure 2, our estimates broadly track the benchmark series over time. Despite some differences in magnitude and timing, especially around turning points, the overall trajectory—including during the 2008 Great Recession—is notably similar.

To quantify the differences, Figure 3 shows a scatter plot (left) and error distribution (right), comparing our estimates to the benchmark. Across all models, approximately 75% of the errors lie within a 2 percentage point range. However, larger deviations are observed for the LSTM and Elastic Net models, which align with their relatively weaker test set performance.

As reported in Table 3, the XGBoost model attains the lowest mean absolute error (MAE) and highest correlation with the benchmark, with the MLP model following closely. All models achieve a sign accuracy of approximately 93%, indicating a high



(a) Scatter plot with benchmark values



(b) Distribution of forecast errors

Figure 3. Benchmark to Koop et al. (2023): Scatter (L) and Error Distribution (R)

level of directional consistency with the benchmark series.

Table 3. US Benchmark Comparison

Model	MAE	RMSE	Corr.	Sign Acc.
MLP	1.269	1.624	0.773	0.933
LSTM	1.401	2.127	0.642	0.924
XGB	1.168	1.504	0.810	0.933
ENET	1.311	1.918	0.708	0.927

Note: MAE (Mean Absolute Error) and RMSE (Root Mean Squared Error) are scale-dependent measures of forecast accuracy where lower values indicate better performance. Correlation measures linear association between predicted and actual values. Sign Accuracy refers to the proportion of times the model correctly predicts the direction of change.

3.1.2 UK Case: Monthly GDP from ONS and Nighttime Light

The UK’s ONS publishes official monthly GDP estimates from January 1997 onward, with the most recent data available as of April 2025. The series is indexed to 100 at October 2022. For comparison, we rescale our predicted monthly GDP series to match this base and focus on the overlapping period. Figure 4 shows the comparison between the official ONS series and our predicted values using the MLP model.¹⁹

Overall, the MLP model closely tracks the general pattern of the official monthly GDP series. Slight deviations are observed during periods of heightened volatility, such as the COVID-19 shock, where the predicted series tends to display greater

¹⁹We present only the MLP model in the main text, as including all models in a single figure obscures the dynamics due to extreme outliers. Full model comparisons are available in Appendix H.

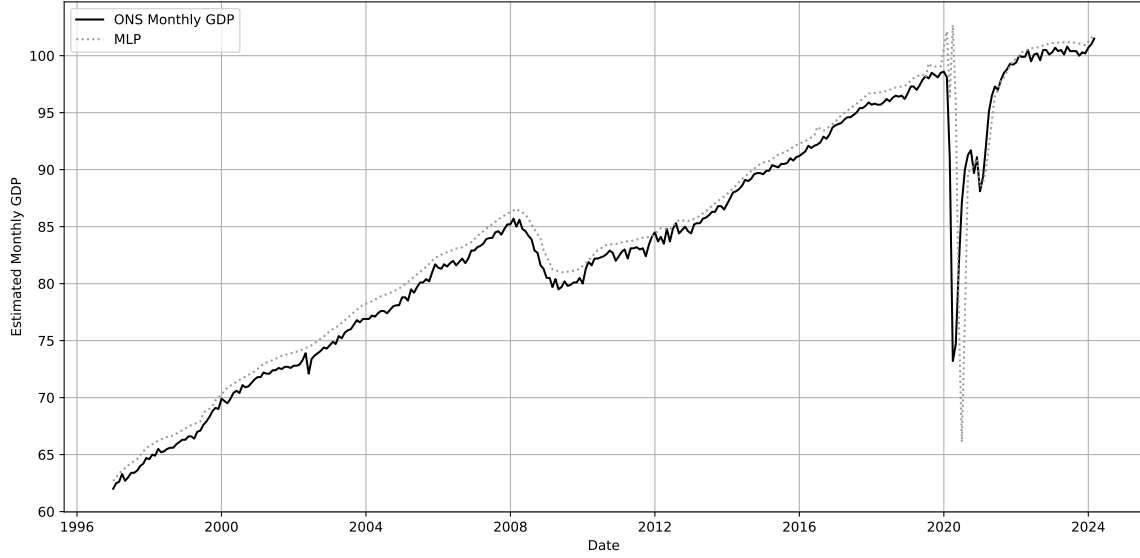
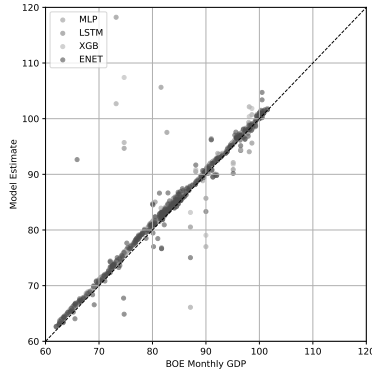
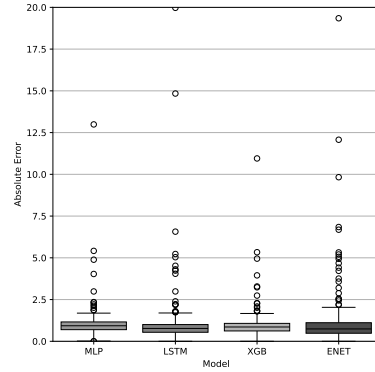


Figure 4. Benchmark to ONS: UK Estimated Monthly GDP (MLP)

amplitude and noise. As reported in Table 4, benchmark performance for the UK is somewhat lower than that for the US but still acceptable. This performance gap may be partly attributed to data availability: the US model utilizes 15 explanatory variables, while the UK model includes only 13.



(a) Scatter plot with benchmark values



(b) Distribution of forecast errors

Figure 5. Benchmark to ONS: Scatter (L) and Error Distribution (R)

Figure 5 presents the corresponding scatter and error distribution plots. The Elastic Net model performs poorly in this case, due in part to extreme outliers observed in certain periods. Since the UK's ONS reports monthly GDP as an index series, we transform the data to enable a meaningful comparison of metrics like sign accuracy and to maintain consistency with the US benchmark. We first compute the monthly

Table 4. UK Benchmark Comparison

Model	MAE	RMSE	Corr.	Sign Acc.
MLP	1.213	2.664	0.972	0.929
LSTM	1.176	3.335	0.954	0.907
XGB	1.323	5.733	0.873	0.910
ENET	49.365	779.262	0.056	0.807

Note: MAE (Mean Absolute Error) and RMSE (Root Mean Squared Error) are scale-dependent measures of forecast accuracy where lower values indicate better performance. Correlation measures the linear association between predicted and actual values. Sign Accuracy refers to the proportion of times the model correctly predicts the direction of change. Since the UK benchmark is based on level data, sign accuracy is not directly comparable. To enable meaningful comparison, we transformed the series according to Equation (7), as done in the US benchmark.

growth rate by log-differencing the index, and then annualize it using Equation (7). Under this transformation, the MLP model achieves a sign accuracy of approximately 93%, on par with the US case. Other models perform less well, with the Elastic Net model falling as low as 80%.

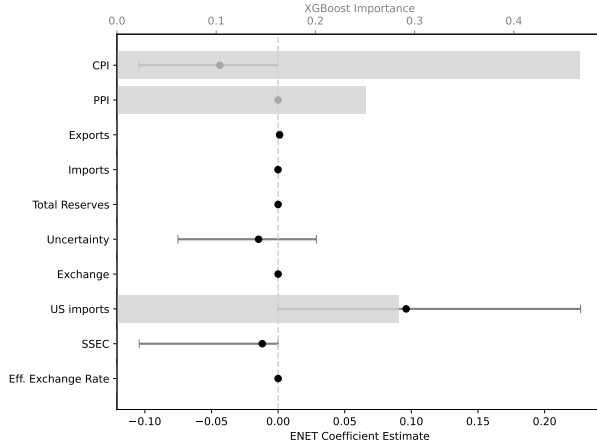
3.2 Interpretation with XGBoost and Elastic Net

As demonstrated in our results, neural network models generally exhibit strong out-of-sample predictive performance. However, their “black box” nature presents a key limitation in academic research, where interpretability is often as important as accuracy.²⁰

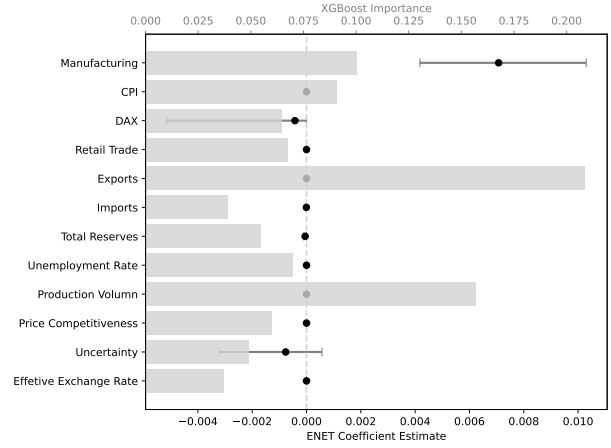
In contrast, models based on decision trees or linear regressions offer more transparent interpretations by providing variable importance scores or direct coefficient estimates. These models can thus serve a complementary role, especially when the differences in predictive performance across models are not substantial. For instance, XGBoost and Elastic Net allow for further examination of the role each explanatory variable plays in the prediction task.

XGBoost, being tree-based, provides intuitive variable importance rankings that are accessible even to non-specialists. However, it does not yield information on coefficient signs or confidence intervals, limiting its utility. On the other hand, while Elastic Net often underperforms in terms of prediction accuracy, it offers notable advantages in interpretability. Through its combination of LASSO (λ_1) and Ridge

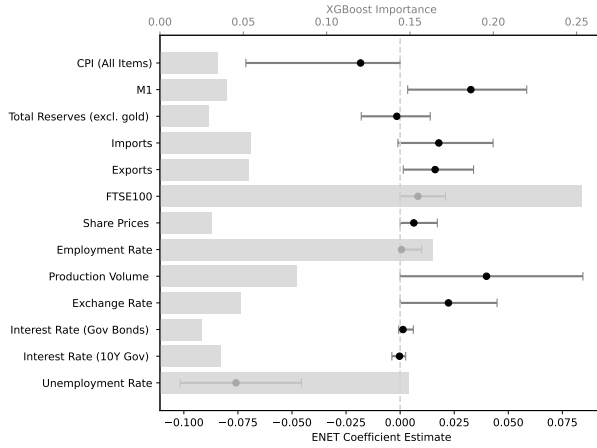
²⁰Although recent advances—such as SHAP (SHapley Additive exPlanations)—have made strides in opening these models to greater scrutiny and interpretability. See Lundberg and Lee (2017).



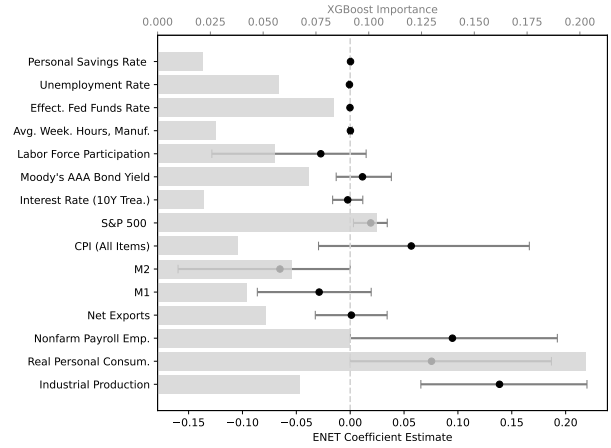
(a) China



(b) Germany



(c) UK



(d) US

Figure 6. Elastic Net Coefficients and XGBoost Importance

(λ_2) penalties,²¹ it performs effective variable selection and regularization, allowing researchers to estimate coefficient signs and confidence intervals while mitigating the influence of irrelevant predictors.

Figure 6 presents a comparative forest plot of variable importance and estimated coefficients across countries.²² Discrepancies between the two approaches are occasionally observed. For example, in the UK case, the employment rate receives low weight in the Elastic Net model, likely due to high correlation with the unemployment rate; yet, XGBoost assigns it notable importance. These differences highlight how each model processes information: Elastic Net, based on a regular-

²¹See Appendix B for further details.

²²See Appendix H for detailed values.

ized linear structure, tends to suppress one of the correlated variables to mitigate multicollinearity, effectively selecting among them. In contrast, XGBoost can accommodate highly correlated variables simultaneously, as it evaluates splits sequentially and non-parametrically in a tree structure.

This distinction can be particularly informative for applied researchers. When one aims to investigate the joint role of closely related economic indicators—such as employment and unemployment rates—Elastic Net may not capture their combined effect due to collinearity penalties, whereas XGBoost may still reflect their joint contribution to prediction. Therefore, contrasting the outputs of these two models provides additional interpretative value, especially in settings where the relationships among predictors are of substantive interest rather than merely nuisance correlations.

These considerations highlight the importance of using multiple models not only for prediction but also for robust interpretability. Cross-model validation of variable relevance can help ensure that empirical findings are not artifacts of a single modeling choice but are instead supported by diverse analytical approaches.

4 Nighttime Light

Table 5. Benchmark Comparison using Nighttime Light

Country	Model	MAE	RMSE	Corr.	Sign Acc.
China	MLP (main)	1.969	2.128	0.651	0.804
	NTL	46.937	57.011	0.065	0.509
	MLP w/ NTL	0.833	1.516	0.732	0.811
	MLP w/ only NTL	0.266	1.216	0.645	0.919
UK	MLP (main)	1.536	4.114	0.705	0.907
	NTL	13.960	19.622	-0.271	0.492
	MLP w/ NTL	1.937	9.598	0.791	0.932
	MLP w/ only NTL	1.801	7.832	0.805	0.880

Note: MAE (Mean Absolute Error) and RMSE (Root Mean Squared Error) are scale-dependent measures of forecast accuracy where lower values indicate better performance. Correlation measures the linear association between predicted and actual values. Sign Accuracy refers to the proportion of times the model correctly predicts the direction of change. Since the UK benchmark is based on level data, sign accuracy is not directly comparable. To enable meaningful comparison, we transformed the series according to Equation (7), as done in the US benchmark.

Nighttime light (NTL) has garnered increasing attention as a proxy for economic indicators such as GDP (K. Chen et al., 2025; X. Chen & Nordhaus, 2011; Donaldson

& Storeygard, 2016; Elvidge et al., 1997; Henderson et al., 2012; Pinkovskiy & Sala-i-Martin, 2016). Given its availability at a daily frequency, NTL may serve as a valuable alternative or complement to monthly GDP. The most widely used NTL datasets are from the US Air Force Defense Meteorological Satellite Program (DMSP) and the Visible Infrared Imaging Radiometer Suite (VIIRS).²³ While DMSP data cover the period from 1992 to 2013 and VIIRS data are available from 2014 to the present, the two are not directly comparable due to differences in imaging time and spatial resolution. Therefore, to maximize the temporal coverage and ensure sufficient training data for our main models, we exclude NTL from the primary analysis.

Nonetheless, since VIIRS NTL data are available at a daily frequency,²⁴ we conduct supplementary analyses to evaluate their utility. Specifically, we assess the predictive performance of NTL by (i) comparing raw NTL data directly to our model’s MLP-based predictions,²⁵ (ii) incorporating NTL data as an explanatory variable within the MLP model for the overlapping period, and (iii) training MLP models with NTL as the sole input. To evaluate these approaches, we compare the resulting predictions against the UK’s ONS monthly GDP and China’s CLI index. This comparison further enables us to assess the heterogeneous informativeness of NTL data across differing economic structures.

As shown in Figure 7 and Table 5, directly using raw NTL data—without any preprocessing other than rescaling to benchmark series—results in large prediction errors.²⁶ In contrast, incorporating NTL as an input variable into the MLP model substantially enhances predictive accuracy, despite the reduction in data coverage by up to 20 years compared to traditional datasets. However, the magnitude of improvement varies by country. In the UK, the baseline MLP model (without NTL) performs better, whereas in China, models that include NTL perform more accurately, with the best results obtained when NTL is the only explanatory variable.

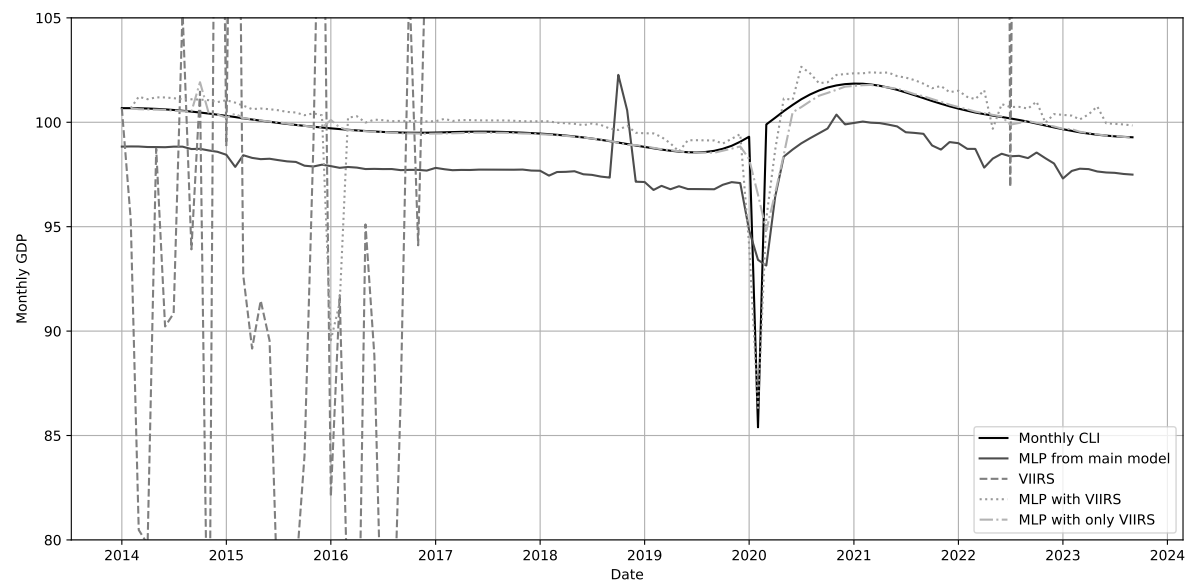
Interestingly, in the UK, models incorporating NTL show stable performance during normal periods but significantly overpredict GDP during the COVID-19 pandemic. In contrast, the Chinese model performs well across the entire period, including the pandemic. This divergence may reflect underlying structural differences in the two economies. In China’s manufacturing-centered economy, reductions in NTL (e.g.,

²³See <https://eogdata.mines.edu/products/vnl/>.

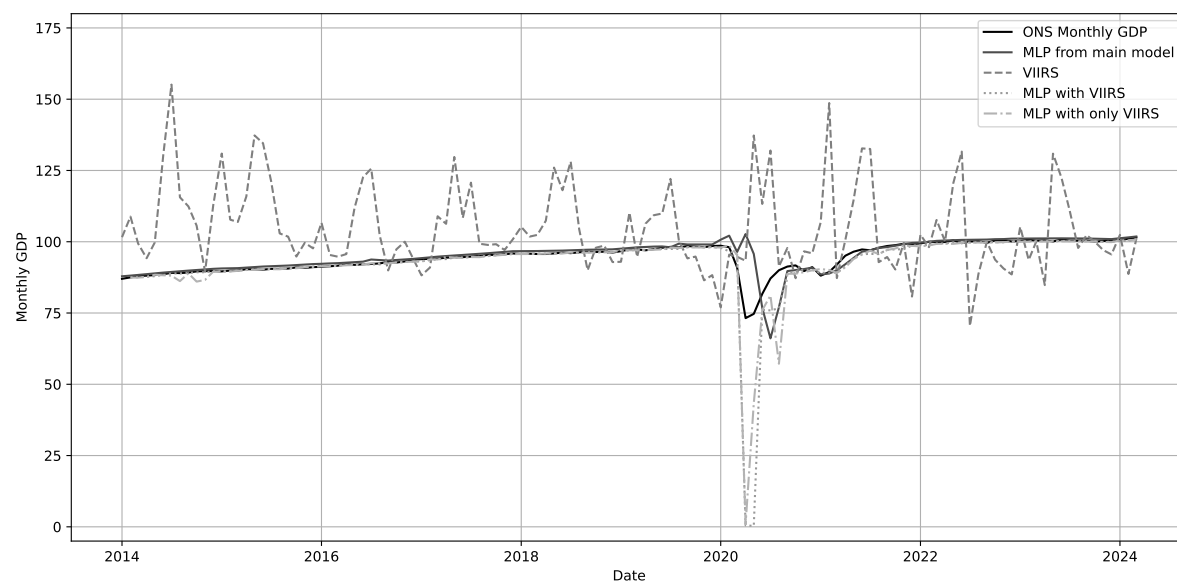
²⁴Appendix G provides further details on the data.

²⁵We use the results from Section 3.

²⁶In practice, many studies that utilize NTL data apply log-differencing or control for various factors such as fixed effects and weather conditions (K. Chen et al., 2025; Pinkovskiy & Sala-i-Martin, 2016).



(a) China



(b) UK

Figure 7. Benchmark using Nighttime Light

factory lights turning off) closely correspond to production halts and GDP declines, making NTL a highly informative proxy.

In contrast, the UK’s economy, dominated by high value-added services, experienced a shift in core economic activity during the lockdowns—from high-illumination commercial spaces to low-illumination residential areas. This spatial shift likely led to an overrepresentation of economic shutdown in NTL data, resulting in significant overprediction of GDP decline. These findings imply that the informativeness of NTL data is contingent on a country’s economic structure. Given the trade-off between temporal coverage and predictive power, the applicability of NTL data in monthly GDP estimation should be carefully considered in the context of each country’s structural characteristics.

5 Conclusion

This paper proposes a scalable, data-driven framework for estimating monthly GDP using a suite of machine learning models. By training four models—MLP, LSTM, XGBoost, and Elastic Net—on quarterly GDP data and monthly economic indicators, we generate high-frequency GDP estimates for China, Germany, the UK, and the US. The methodology is validated through benchmarking against official monthly GDP figures for the UK and academic estimates for the US, demonstrating both credibility and robustness across heterogeneous data environments. Among the models, MLP stands out for its balance of predictive accuracy and computational efficiency.

The primary contribution of this study is to offer a flexible alternative to traditional econometric approaches, which often rely on strong structural assumptions. While the machine learning framework is relatively more computationally intensive, it adapts well to varying levels of data availability, making it broadly applicable across different countries. Moreover, by integrating interpretable models such as XGBoost and Elastic Net with “black box” neural networks, the framework enables model-based validation of variable importance and economic mechanisms. This hybrid design enhances both transparency and trust in the estimated outputs, providing policymakers with timely and interpretable tools for high-frequency economic monitoring. Additionally, we show that the framework’s applicability can be extended to data-scarce regions by incorporating novel inputs such as nighttime light (NTL) data.

Despite its promising performance, the framework has several limitations that point to directions for future research. A key constraint is its reliance on official

quarterly GDP figures for the final reconciliation step, which limits the model’s utility to retrospective estimation. To support real-time nowcasting, a natural extension would be to generate out-of-sample forecasts for the current quarter’s GDP, which could then serve as the target for monthly reconciliation.²⁷ Such an extension would preserve the model’s structure while expanding its utility to real-time monitoring.

Another potential improvement lies in refining the variable selection process. In this study, we use the full set of available indicators for each country; however, a two-stage strategy could improve both model performance and computational efficiency. For instance, Elastic Net could be employed in the first stage for variable selection, leveraging its regularization properties to eliminate noisy or redundant predictors. A nonlinear model such as MLP could then be trained on this reduced feature set, enhancing both the signal-to-noise ratio and overall interpretability.

²⁷Except in exceptional circumstances such as the COVID-19 pandemic, NTL data alone demonstrates strong predictive power for GDP. This suggests that one viable approach could be to first estimate GDP using NTL as a primary input and then apply corrections or adjustments using additional data sources.

References

- Baker, S. R., Bloom, N., & Davis, S. J. (n.d.-a). Economic Policy Uncertainty Index for Europe [EUEPUINDXM] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025]. Based on: Baker, Scott, Nicholas Bloom and Steven Davis (2012), “Measuring Economic Policy Uncertainty” .].
- Baker, S. R., Bloom, N., & Davis, S. J. (n.d.-b). Economic Policy Uncertainty Index: Mainland Papers for China [CHNMAINLANDEPU] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025]. Based on: “Economic Policy Uncertainty in China Since 1949: The View from Mainland Newspapers” by Davis, Liu, and Sheng.].
- BEA. (n.d.-a). Personal Saving Rate [PSAVERT] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- BEA. (n.d.-b). Real Gross Domestic Product [GDPC1] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- BEA. (n.d.-c). Real Personal Consumption Expenditures [DPCERA3M086SBEA] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- BIS. (n.d.-a). Real Broad Effective Exchange Rate Index for China [RBCNBIS] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- BIS. (n.d.-b). Real Broad Effective Exchange Rate Index for Germany [RNDEBIS] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- BLS. (n.d.-a). All Employees: Total Nonfarm Payrolls [PAYEMS] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- BLS. (n.d.-b). Average Weekly Hours of Production and Nonsupervisory Employees, Manufacturing [AWHMAN] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- BLS. (n.d.-c). Consumer Price Index for All Urban Consumers: All Items in U.S. City Average [CPIAUCSL] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- BLS. (n.d.-d). Labor Force Participation Rate [CIVPART] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- BLS. (n.d.-e). Producer Price Index by Commodity: Furniture and Household Durables: Vitreous China, Fine Earthenware, and Other Pottery Products [WPU1261] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 22, 2025].
- BLS. (n.d.-f). Unemployment Rate [UNRATE] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].

- Board of Governors of the Federal Reserve System. (n.d.-a). Federal Funds Effective Rate [FEDFUNDS] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- Board of Governors of the Federal Reserve System. (n.d.-b). Industrial Production: Total Index [INDPRO] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- Board of Governors of the Federal Reserve System. (n.d.-c). M1 [M1SL] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- Board of Governors of the Federal Reserve System. (n.d.-d). M2 [M2SL] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- Board of Governors of the Federal Reserve System. (n.d.-e). Market Yield on U.S. Treasury Securities at 10-Year Constant Maturity, Quoted on an Investment Basis [GS10] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- Board of Governors of the Federal Reserve System. (n.d.-f). U.S. Dollars to U.K. Pound Sterling Spot Exchange Rate [EXUSUK] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 27, 2025].
- Brave, S. A., Butters, R. A., Kelley, D., et al. (2019). A New “Big Data” Index of US Economic Activity. *Economic Perspectives*, 43(1), 1–30.
- Chen, K., Wang, Q., Xu, J., & Yao, J. (2025). The International Spillover of Monetary Policy Shock: New Evidence from Nighttime Light.
- Chen, T., & Guestrin, C. (2016). XGBoost: A Scalable Tree Boosting System. *Proceedings of the 22nd ACM SIGKDD International Conference on Knowledge Discovery and Data Mining*, 785–794.
- Chen, X., & Nordhaus, W. D. (2011). Using Luminosity Data as a Proxy for Economic Statistics. *Proceedings of the National Academy of Sciences*, 108(21), 8589–8594.
- Chow, G. C., & Lin, A. (1971). Best Linear Unbiased Interpolation, Distribution, and Extrapolation of Time Series by Related Series. *The Review of Economics and Statistics*, 372–375.
- Chu, B., & Qureshi, S. (2023). Comparing Out-of-Sample Performance of Machine Learning Methods to Forecast US GDP Growth. *Computational Economics*, 62(4), 1567–1609.

- Denton, F. T. (1971). Adjustment of Monthly or Quarterly Series to Annual Totals: An Approach Based on Quadratic Minimization. *Journal of the American Statistical Association*, 66(333), 99–102.
- Donaldson, D., & Storeygard, A. (2016). The View from Above: Applications of Satellite Data in Economics. *Journal of Economic Perspectives*, 30(4), 171–198.
- Elvidge, C. D., Baugh, K. E., Kihn, E. A., Kroehl, H. W., Davis, E. R., & Davis, C. W. (1997). Relation between Satellite Observed Visible-near Infrared Emissions, Population, Economic Activity and Electric Power Consumption. *International Journal of Remote Sensing*, 18(6), 1373–1379.
- Eurostat. (n.d.). Real GDP for Germany [CLVMNACSCAB1GQDE] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 27, 2025].
- Ghosh, S., & Ranjan, A. (2023). A Machine Learning Approach to GDP Nowcasting: An Emerging Market Experience. *Bulletin of Monetary Economics and Banking*, 26, 33–54.
- Giannone, D., Reichlin, L., & Small, D. (2008). Nowcasting: The Real-Time Informational Content of Macroeconomic Data. *Journal of Monetary Economics*, 55(4), 665–676.
- Goulet Coulombe, P., Leroux, M., Stevanovic, D., & Surprenant, S. (2022). How is Machine Learning Useful for Macroeconomic Forecasting? *Journal of Applied Econometrics*, 37(5), 920–964.
- Henderson, J. V., Storeygard, A., & Weil, D. N. (2012). Measuring Economic Growth from Outer Space. *American Economic Review*, 102(2), 994–1028.
- Hochreiter, S., & Schmidhuber, J. (1997). Long Short-term Memory. *Neural Computation*, 9(8), 1735–1780.
- Holz, C. A. (2014). The Quality of China’s GDP Statistics. *China Economic Review*, 30, 309–338.
- IMF. (n.d.-a). Interest Rates, Government Securities, Government Bonds for United Kingdom [INTGSBGBM193N] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- IMF. (n.d.-b). Real GDP for Great Britain [NGDPRSAXDCGBQ] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- IMF. (n.d.-c). Total Reserves excluding Gold for China [TRESEGCNM052N] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 22, 2025].
- IMF. (n.d.-d). Total Reserves excluding Gold for United Kingdom [TRESEGGBM052N] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].

- James, G., Witten, D., Hastie, T., Tibshirani, R., & Taylor, J. (2023). *An Introduction to Statistical Learning with Applications in Python*. Springer.
- Jansen, W. J., Jin, X., & de Winter, J. M. (2016). Forecasting and Nowcasting Real GDP: Comparing Statistical Models and Subjective Forecasts. *International Journal of Forecasting*, 32(2), 411–436.
- Koop, G., McIntyre, S., Mitchell, J., & Poon, A. (2021). Nowcasting ‘True’ Monthly US GDP During the Pandemic. *National Institute Economic Review*, 256, 44–70.
- Koop, G., McIntyre, S., Mitchell, J., & Poon, A. (2023). Reconciled Estimates of Monthly GDP in the United States. *Journal of Business & Economic Statistics*, 41(2), 563–577.
- Litterman, R. B. (1983). A Random Walk, Markov Model for the Distribution of Time Series. *Journal of Business & Economic Statistics*, 1(2), 169–173.
- Lundberg, S. M., & Lee, S.-I. (2017). A Unified Approach to Interpreting Model Predictions. *Advances in Neural Information Processing Systems*, 30.
- Mariano, R. S., & Murasawa, Y. (2003). A New Coincident Index of Business Cycles Based on Monthly and Quarterly Series. *Journal of Applied Econometrics*, 18(4), 427–443.
- Mariano, R. S., & Murasawa, Y. (2010). A Coincident Index, Common Factors, and Monthly Real GDP. *Oxford Bulletin of Economics and Statistics*, 72(1), 27–46.
- Mitchell, J., Smith, R. J., Weale, M. R., Wright, S., & Salazar, E. L. (2005). An Indicator of Monthly GDP and an Early Estimate of Quarterly GDP Growth. *The Economic Journal*, 115(501), F108–F129.
- Moody’s. (n.d.). Moody’s Seasoned Aaa Corporate Bond Yield [AAA] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-a). 10-Year Government Bond Interest Rate for UK [IRLTLT01GBM156N] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-b). Consumer Price Index for Germany [DEUCPIALLMINMEI] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-c). Consumer Price Index for the United Kingdom [GBRCPIALLMINMEI] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-d). Consumer Price Index: All Items: Total for China [CPALTT01CNM657N] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].

- OECD. (n.d.-e). Exports of Goods and Services for China [XTEXVA01CNM664S] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-f). Exports of Goods and Services for Germany [XTEXVA01DEM664S] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-g). Exports of Goods and Services for UK [XTEXVA01GBM664S] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-h). Gross Domestic Product for China [CHNGDPNQDSMEI] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-i). Imports of Goods and Services for China [XTIMVA01CNM664S] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-j). Imports of Goods and Services for Germany [XTIMVA01DEM664S] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-k). Imports of Goods and Services for UK [XTIMVA01GBM664S] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-l). Industrial Production Index for Germany [DEUPROINDMISMEI] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-m). Industrial Production Index for the UK [GBRPROINDMISMEI] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-n). Manufacturing Production for Germany [DEUPRMNT001GPSAM] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-o). Retail Trade for Germany [DEUSARTMISMEI] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-p). Share Prices in the UK [SPASTT01GBM661N] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-q). Total Reserves Excluding Gold for Germany [TRESEGDEM052N] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- OECD. (n.d.-r). Unemployment Rate in Germany [LRHUTTTTTDEM156S] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- Pinkovskiy, M., & Sala-i-Martin, X. (2016). Lights, Camera... Income! Illuminating the National Accounts-household Surveys Debate. *The Quarterly Journal of Economics*, 131(2), 579–631.
- Rawski, T. G. (2001). What is Happening to China’s GDP Statistics? *China Economic Review*, 12(4), 347–354.

- Richardson, A., van Florenstein Mulder, T., & Vehbi, T. (2021). Nowcasting GDP using Machine-learning Algorithms: A Real-time Assessment. *International Journal of Forecasting*, 37(2), 941–948.
- Rumelhart, D. E., Hinton, G. E., & Williams, R. J. (1986). Learning Representations by Back-propagating Errors. *Nature*, 323(6088), 533–536.
- Said, S. E., & Dickey, D. A. (1984). Testing for Unit Roots in Autoregressive-Moving Average Models of Unknown Order. *Biometrika*, 71(3), 599–607.
- Scruton, J., O'Donnell, M., & Dey-Chowdhury, S. (2018). Introducing a New Publication Model for GDP. *Office for National Statistics*, 1–14.
- U.S. Census Bureau. (n.d.). U.S. Imports from China [IMPCH] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- U.S. Census Bureau and BEA. (n.d.). Exports of Goods and Services, Balance of Payments Basis [BOPTEXP] [Retrieved from FRED, Federal Reserve Bank of St. Louis, May 2, 2025].
- Zou, H., & Hastie, T. (2005). Regularization and Variable Selection via the Elastic Net. *Journal of the Royal Statistical Society Series B: Statistical Methodology*, 67(2), 301–320.

Appendices

A Data Descriptions

All data used in this study are publicly available from open-source platforms. Detailed instructions for data acquisition—including Python scripts for API-based access—are provided in the replication package.²⁸ Most datasets are retrieved via the FRED API, which requires user registration (free of charge), and the Yahoo Finance API, which does not require registration. As of June 2025, all sources remain accessible. However, it is important to note that API providers may not be the original data publishers, and access could be discontinued without prior notice. The processed datasets used in this study are stored as `master_<country>.csv` files in model folder within the replication package. All summary statistics are based on raw data prior to any log-differencing or standardization.

A.1 China Data

Table A1. China Dataset Description

Variable	Code	Unit Root	Transform.
CPI (All Items, Growth)	CPALTT01CNM657N	No	$(X_m - \mu_j)/\sigma_j$
PPI (Furniture, Household)	WPU1261	Yes	$\Delta \log(X_m)$
Total Reserves (excl. gold)	TRESEGCNM052N	Yes	$\Delta \log(X_m)$
Exports	XTEXVA01CNM664S	Yes	$\Delta \log(X_m)$
Imports	XTIMVA01CNM664S	No	$\Delta \log(X_m)$
US Imports	IMPCH	Yes	$\Delta \log(X_m)$
Effective Exchange Rate	RBCNBIS	Yes	$\Delta \log(X_m)$
Exchange Rate (RMB/USD)	EXCHUS	Yes	$\Delta \log(X_m)$
Policy Uncertainty	CHNMAINLANDEPU	Yes	$\Delta \log(X_m)$
SSEC Index	-	Yes	$\Delta \log(X_m)$
Real GDP	CHNGDPNQDSMEI	-	$\Delta \log(X_m)$

Note: Variable codes are retrieved from FRED (<https://fred.stlouisfed.org/>). SSEC Index is obtained from Investing.com (<https://www.investing.com/>).

²⁸https://github.com/Yonggeun-Jung/monthly_gdp

Table A2. China Data Summary Statistics

Variable	Obs.	Mean	Std. Dev.	Min	Max
CPI (All Items, Growth)	357	0.248	0.796	-1.403	4.048
PPI (Furniture, Household)	357	220.260	43.093	156.200	330.018
Total Reserves (excl. gold)	357	1.81e+06	1.44e+06	26355	4.01e+06
Exports	357	8.04e+11	5.87e+11	5.75e+10	2.15e+12
Imports	357	6.64e+11	4.63e+11	7.21e+10	1.61e+12
US Imports	357	24764.98	14770.62	2183	52081.07
Exchange Rate (RMB/USD)	357	7.381	0.859	6.051	8.725
Effective Exchange Rate	357	82.613	13.719	52.800	106.380
Policy Uncertainty	357	145.334	112.203	10.111	661.828
SSEC Index	357	2256.469	1017.002	333.920	5954.770
Real GDP	119	1.14e+13	9.59e+12	9.38e+11	3.36e+13

Note: Summary statistics are calculated using raw values before any transformation (e.g., differencing or standardization).

A.2 Germany Data

Table A3. Germany Dataset Description

Variable	Code	Unit Root	Transform.
CPI (All Items)	DEUCPIALLMINMEI	Yes	$\Delta \log(X_m)$
Retail Trade	DEUSARTMISMEI	Yes	$\Delta \log(X_m)$
Unemployment Rate	LRHUTTTTDEM156S	Yes	$\Delta \log(X_m)$
Production Volume	DEUPROINDMISMEI	Yes	$\Delta \log(X_m)$
Manufacturing	DEUPRMNTO01GPSAM	No	$(X_m - \mu_j)/\sigma_j$
Exports	XTEXVA01DEM664S	Yes	$\Delta \log(X_m)$
Imports	XTIMVA01DEM664S	Yes	$\Delta \log(X_m)$
Price Competitiveness	-	Yes	$\Delta \log(X_m)$
Effective Exchange Rate	RNDEBIS	Yes	$\Delta \log(X_m)$
Total Reserves (excl. gold)	TRESEGDEM052N	Yes	$\Delta \log(X_m)$
DAX Index	GDAXI	Yes	$\Delta \log(X_m)$
Policy Uncertainty	EUEPUINDXM	Yes	$\Delta \log(X_m)$
Real GDP	CLVMNACSCAB1GQDE	-	$\Delta \log(X_m)$

Note: Variable codes are retrieved from FRED (<https://fred.stlouisfed.org/>) and Yahoo Finance (<https://finance.yahoo.com/>). Price competitiveness index is obtained from the Deutsche Bundesbank (<https://www.bundesbank.de/en>).

Table A4. Germany Data Summary Statistics

Variable	Obs.	Mean	Std. Dev.	Min	Max
CPI (All Items)	396	90.15	13.72	63.998	124.19
Retail Trade	396	100.12	7.16	90.71	122.61
Unemployment Rate	396	6.66	2.48	2.90	11.20
Production Volume	396	87.85	11.68	65.25	107.38
Manufacturing	396	0.07	2.15	-22.02	11.94
Exports	396	7.21e+10	3.13e+10	2.58e+10	1.39e+11
Imports	396	6.08e+10	2.66e+10	2.33e+10	1.34e+11
Price Competitiveness	396	98.72	3.90	90.90	111.90
Effective Exchange Rate	396	102.63	4.33	93.89	116.97
Total Reserves (excl. gold)	396	64913.49	15365.64	39972.28	120255.30
DAX Index	396	7025.15	4126.87	1373.61	16681.73
Policy Uncertainty	396	153.02	77.56	47.69	433.28
Real GDP	132	644258.50	80194.45	514515.30	770584.30

Note: Summary statistics are calculated using raw values before any transformation (e.g., differencing or standardization).

A.3 UK Data

Table A5. UK Dataset Description

Variable	Code	Unit Root	Transform.
CPI (All Items)	GBRCPIALLMINMEI	Yes	$\Delta \log(X_m)$
Production Volume	GBRPROINDMISMEI	Yes	$\Delta \log(X_m)$
Unemployment Rate	MG SX	Yes	$\Delta \log(X_m)$
Employment Rate	LF24	Yes	$\Delta \log(X_m)$
Exports	XTEXVA01GBM664S	Yes	$\Delta \log(X_m)$
Imports	XTIMVA01GBM664S	Yes	$\Delta \log(X_m)$
Exchange Rate	EXUSUK	Yes	$\Delta \log(X_m)$
Interest Rate (Gov Bonds)	INTGSBGBM193N	Yes	$\Delta \log(X_m)$
Interest Rate (10Y Gov)	IRLTLT01GBM156N	Yes	$\Delta \log(X_m)$
Total Reserves (excl. gold)	TRESEGGBM052N	Yes	$\Delta \log(X_m)$
M1 Money Stock	LPMVWYE	Yes	$\Delta \log(X_m)$
Share Prices	SPASTT01GBM661N	Yes	$\Delta \log(X_m)$
FTSE 100 Index	FTSE	Yes	$\Delta \log(X_m)$
Real GDP	NGDPRSAXDCGBQ	—	$\Delta \log(GDP_q)$

Note: Codes, except for FTSE 100 Index, M1 Money Stock, Unemployment rate, and Employment Rate, are from FRED (<https://fred.stlouisfed.org/>), FTSE 100 Index is retrieved from Yahoo Finance (<https://finance.yahoo.com/>). Total Reserves and M1 series are from BOE (<https://www.bankofengland.co.uk/boeapps/database/>). Unemployment and employment rates are from ONS (<https://www.ons.gov.uk/>). Since BOE and ONS do not offer data in API format, we downloaded full-period datasets manually from their websites and processed them by hand.

Table A6. UK Data Summary Statistics

Variable	Obs.	Mean	Std. Dev.	Min	Max
CPI (All Items)	399	87.01	18.22	57.00	131.60
Production Volume	399	96.73	9.62	75.20	113.32
Unemployment Rate	399	6.24	1.88	3.60	10.70
Employment Rate	399	72.36	2.05	68.30	76.50
Exports	399	1.96e+10	6.42e+09	8.32e+09	3.32e+10
Imports	399	2.67e+10	1.14e+10	9.35e+09	5.87e+10
Exchange Rate	399	1.56	0.20	1.13	2.07
Share Prices	399	83.34	22.75	31.94	120.02
Interest Rate (Gov Bonds)	399	4.34	2.55	0.21	10.59
Interest Rate (10Y Gov)	399	4.34	2.55	0.21	10.59
Total Reserves (excl. gold)	399	82244.13	46307.86	28912.71	176024.00
M1 Money Stock	399	992469.80	680818.10	174324.00	2575027.00
Share Prices	399	83.34	22.75	31.94	120.02
FTSE 100 Index	399	5495.65	1500.32	2106.64	7915.01
Real GDP	133	500639.00	88311.65	344987.00	638746.00

Note: Summary statistics are calculated using raw values before any transformation (e.g., differencing or standardization).

A.4 US Data

Table A7. US Dataset Description

Variable	Code	Unit Root	Transform.
Avg. Week. Hours, Manuf.	AWHMAN	No	$(X_m - \mu_j)/\sigma_j$
CPI (All Items)	CPIAUCSL	Yes	$\Delta \log(X_m)$
Industrial Production	INDPRO	Yes	$\Delta \log(X_m)$
Real Personal Consum.	DPCERA3M086SBEA	Yes	$\Delta \log(X_m)$
Effect. Fed Funds Rate	FEDFUNDS	No	$(X_m - \mu_j)/\sigma_j$
Interest Rate (10Y Trea.)	GS10	Yes	$\Delta \log(X_m)$
Unemployment Rate	UNRATE	No	$(X_m - \mu_j)/\sigma_j$
M2 Money Stock	M2SL	Yes	$\Delta \log(X_m)$
M1 Money Stock	M1SL	Yes	$\Delta \log(X_m)$
Labor Force Participation	CIVPART	Yes	$\Delta \log(X_m)$
Personal Savings Rate	PSAVERT	No	$(X_m - \mu_j)/\sigma_j$
Nonfarm Payroll Emp.	PAYEMS	Yes	$\Delta \log(X_m)$
Moody's AAA Bond Yield	AAA	Yes	$\Delta \log(X_m)$
Net Exports	BOPTEXP	Yes	$\Delta \log(X_m)$
S&P 500 Index	GSPC	Yes	$\Delta \log(X_m)$
Real GDP	GDPC1	—	$\Delta \log(GDP_q)$

Note: All data, except for the S&P 500 Index, are from FRED (<https://fred.stlouisfed.org/>). The S&P 500 Index is obtained from Yahoo Finance (<https://finance.yahoo.com/>).

Table A8. US Data Summary Statistics

Variable	Obs.	Mean	Std. Dev.	Min	Max
Avg. Week. Hours, Manuf.	396	41.19	0.61	38.40	42.40
CPI (All Items)	396	211.53	46.71	138.30	317.60
Industrial Production	396	91.84	11.41	61.48	104.10
Real Personal Consum.	396	83.12	19.72	49.07	123.08
Effect. Fed Funds Rate	396	2.60	2.18	0.05	6.54
Interest Rate (10Y Trea.)	396	4.01	1.75	0.62	7.96
Unemployment Rate	396	5.68	1.79	3.40	14.80
M2 Money Stock	396	9526.84	5663.39	3381.20	21749.60
M1 Money Stock	396	4237.49	5996.20	910.40	20727.10
Labor Force Participation	396	64.85	1.89	60.10	67.30
Personal Savings Rate	396	5.94	2.92	1.40	32.00
Nonfarm Payroll Emp.	396	134619.49	12153.07	108313.00	158942.00
Moody's AAA Bond Yield	396	5.40	1.60	2.14	8.68
Net Exports	396	142927.29	64637.35	50044.00	273516.00
S&P 500 Index	396	1776.10	1244.48	407.36	6017.37
Real GDP	132	16525.10	3582.27	10236.43	23542.35

Note: Summary statistics are calculated using raw values before any transformation (e.g., differencing or standardization).

B Model Descriptions

B.1 MLP: Multi-Layer Perceptron

In the Multi-Layer Perceptron (MLP) networks we estimate the regression model

$$Y_q = f(X_q; \theta) + e_q, \quad (\text{B1})$$

where $f(X_q; \theta)$ is the function learned by an MLP which is a class of feedforward artificial neural networks. An MLP models the dependent (target) variable as a nested series of nonlinear transformations of the explanatory variables (features). For instance, a simple MLP with one hidden layer can be represented as:

$$f(X_q; \theta) = W_2 \sigma(W_1 X_q + b_1) + b_2, \quad (\text{B2})$$

where W_1, W_2 are weight matrices, b_1, b_2 are bias vectors, σ is a nonlinear activation function, and $\theta = \{W_1, b_1, W_2, b_2\}$ represents the set of all model parameters. The model parameters θ are estimated by minimizing the mean squared error (MSE) between the actual quarterly growth Y_q and the model's predictions $f(X_q; \theta)$. This minimization is performed using the Adam optimizer.²⁹ The architecture and training process of the MLP are governed by several hyperparameters, which we optimize through a Bayesian optimization search to find the best-performing model specification.³⁰

B.2 LSTM: Long Short-Term Memory

The Long Short-Term Memory (LSTM) network is a type of Recurrent Neural Network (RNN) specifically designed to model time-series and sequential data (Hochreiter & Schmidhuber, 1997). Unlike feedforward networks such as the MLP, LSTM contains feedback connections, allowing information to persist. Its key innovation is the gating mechanism (comprising forget, input, and output gates), which regulates the flow of information through the network's cell state. This structure enables the model to effectively capture long-range dependencies and mitigate the vanishing

²⁹Adam (Adaptive Moment Estimation) is a stochastic optimizer that adapts learning rates using estimates of first and second moments of gradients (Kingma, 2014).

³⁰In practice, we utilized the `Keras-Tuner` package for this purpose. Bayesian optimization is a sequential strategy that finds the optimal hyperparameter combination by building a probabilistic model of the function mapping hyperparameters to model performance, making the search more efficient than random or grid search. See https://keras.io/keras_tuner/.

gradient problem common in simple RNNs.

In our implementation, we structure the quarterly data into sequences of a fixed length t .³¹ The LSTM model is then trained to predict the GDP growth of the subsequent quarter, Y_q , based on the sequence of explanatory variables from the preceding four quarters, $\{X_{q-1}, X_{q-2}, \dots, X_{q-t}\}$. Similar to the MLP, the model parameters are estimated by minimizing the MSE using the Adam optimizer, and the optimal network architecture is determined through Bayesian hyperparameter optimization.

B.3 XGBoost: Extreme Gradient Boosting

We also employ XGBoost (Extreme Gradient Boosting), a high-performance, tree-based ensemble learning method based on the gradient boosting framework (Chen & Guestrin, 2016). The core idea of gradient boosting is to build models sequentially, with each new model correcting the errors (residuals) of its predecessor. XGBoost builds on this by using an optimized and regularized implementation of decision trees as base learners.

The model is trained to predict quarterly GDP growth, Y_q , from the aggregated quarterly features, X_q . We treat it as a standard regression problem and optimize its key hyperparameters—such as the number of trees (`n_estimators`), tree depth (`max_depth`), and `learning_rate`—using a cross-validated search to find the model with the best predictive performance. In addition to its predictive power, XGBoost also provides feature importance scores, offering a nonlinear perspective on the key drivers of GDP growth.

B.4 Elastic Net

This section describes the Elastic Net model that is employed mainly to explain the role of each explanatory variable in linear projection setting. We estimate the regression model

$$Y_q = f(X_q; \beta) + e_q, \quad (\text{B3})$$

where $f(X_q; \beta)$ is a linear function of the aggregated regressors and e_q is an error term. The parameter vector $\beta \in \mathbb{R}^k$ denotes the marginal effect of each explanatory variable on quarterly GDP growth. Specifically, we minimize the following penalized

³¹We set 4 as the default

loss function:

$$\mathcal{L}(\beta) = \sum_q (Y_q - X_q^\top \beta)^2 + \lambda_1 \sum_{j=1}^k |\beta_j| + \lambda_2 \sum_{j=1}^k \beta_j^2, \quad (\text{B4})$$

where the first term penalizes the discrepancy between the observed quarterly GDP growth (Y_q) and the predictions derived from the quarter-level regressors ($X_q^\top \beta$). The second and third terms introduce regularization penalties to control model complexity: specifically, (i) $\sum_{j=1}^k |\beta_j|$ is the sum of the absolute values of the elements, which encourages sparsity in the coefficients. This enables automatic selection of a subset of relevant variables, setting others to zero; (ii) $\sum_{j=1}^k \beta_j^2$ is the sum of squared coefficients, which shrinks large coefficients and mitigates overfitting when explanatory variables are highly correlated. This combination of penalties defines the Elastic Net regularization method, which balances the strengths of both LASSO (Least Absolute Shrinkage and Selection Operator) and Ridge regularization (James et al., 2023).³² The regularization parameters λ_1 and λ_2 are treated as hyperparameters, which must be tuned by the researcher using a validation set. Their values determine the model’s bias-variance tradeoff and affect the selection and shrinkage behavior of explanatory variables.

Furthermore, a nonparametric bootstrap approach is also employed to improve the model’s predictive accuracy (Hansen, 2022; James et al., 2023). Specifically, this involves generating a number of bootstrap samples ($B = 50,000$) by resampling with replacement from the original dataset. The model is then re-estimated on each of these samples. This process not only helps in enhancing the stability of the predictions, often by averaging outcomes from the B models, but also allows for the empirical estimation of confidence intervals for the predicted monthly GDP figures.

C Algorithms

This section illustrates algorithms described in Section 2.1. We implement the algorithms by Python. The Python scripts are available at the replication package.³³

³²LASSO regularization (λ_1 penalty) performs variable selection by shrinking some coefficients to exactly zero, while Ridge regularization (λ_2 penalty) shrinks coefficients continuously without eliminating any (James et al., 2023).

³³https://github.com/Yonggeun-Jung/monthly_gdp

C.1 Data Preprocessing

Algorithm 1: Monthly Transformation of Explanatory Variables

Data: Master file `master_country.csv` with columns: `DATE`, X_1, \dots, X_k , `Y`;
pre-defined `log_diff_cols`

Result: Transformed monthly explanatory variables `X_prime_m_df`

```

1 Load master_country.csv into DataFrame df.;
2 Convert DATE column to datetime format.;
3 Create quarter column by extracting quarter from DATE.;
4 Initialize X_prime_m_df with columns DATE and quarter.;
5 for each explanatory variable  $X_j \in \{X_1, \dots, X_k\}$  do
6   | Conduct ADF test on  $X_j$  to check stationarity and define log_diff_cols.;
7   | if  $X_j \in \text{log\_diff\_cols}$  then
8   |   |  $X'_{m,j} \leftarrow \log(X_j) - \log(X_{j,\text{prev}})$ ;
9   | else
10  |   |  $X'_{m,j} \leftarrow X_j$ ;
11  | end
12  | Add  $X'_{m,j}$  to X_prime_m_df.;
13 end

```

In Algorithm 1, we preprocess the dataset. The data collected for each country are stored in files named `master_<country_name>.csv`, each containing columns for `DATE`, X_1, \dots, X_k , and `Y`. The primary task is to determine how to transform the X variables. For each X_j , we conduct an ADF test using `tsa.stattools.adfuller` function from `statsmodels` package to detect a unit root. If X_j exhibits a unit root and is not included in the exception list of variables that should be treated as growth rates, we apply a log-difference transformation. Otherwise, we retain X_j in its original form.

Algorithm 2: Quarterly Aggregation and Final Processing

Data: Transformed monthly `X_prime_m_df`; `log_diff_cols`

Result: Quarterly dataset `X_q_processed`, `Y_q_processed`

```
1 quarters ← unique sorted quarters from X_prime_m_df;
2 for each quarter q in quarters do
3   X_current_q ← rows in quarter q from X_prime_m_df;
4   for each variable j in explanatory variables do
5     if j ∈ log_diff_cols then
6       |  $X_{q,j} \leftarrow \sum_{m \in q} X'_{m,j}$ ;
7     else
8       |  $X_{q,j} \leftarrow \frac{1}{3} \sum_{m \in q} X'_{m,j}$ ;
9     end
10    X_q_row[j] ←  $X_{q,j}$ ;
11  end
12 end
13 Extract  $Y_q$  as the first month of each quarter in df.;
14 Y_q_processed ←  $Y_q \leftarrow \log(Y_q) - \log(Y_{q-1})$ , dropping first value.;
15 X_q_processed ← concatenate log-diff and non-log columns;
```

In Algorithm 2, we aggregate the monthly transformed variables into quarterly observations to match the frequency of the dependent variable Y . For each quarter q , we sum the monthly values of log-differenced variables and average the monthly values of non-log-differenced variables.³⁴ The Y_q is computed as the log difference.

C.2 Main Body of Models

C.2.1 MLP: Multi-Layer Perceptron

In Algorithm 3, we identify the optimal set of hyperparameters by iteratively training multiple MLP network specifications. Among the candidate models, we retain the one that achieves the lowest validation MSE as the final model. This procedure is implemented using the `keras-tuner` library provided by `TensorFlow`.

³⁴In case of the non-log-differenced variables, we standardize it only after splinting data into training/test sets to avoid data leakage.

Algorithm 3: Neural Network Training and Hyperparameter Optimization

Data: `X_q_processed`; `Y_q_processed`; `non_log_cols`

Result: The best-performing trained model, $\hat{f}(\cdot; \hat{\theta})$

```
/* Split Dataset and Standardization */
1 Split X_q_processed and Y_q_processed into 80% training sets;;
2 ( $X_q^{\text{train}}$ ,  $Y_q^{\text{train}}$ ) and 20% test sets ( $X_q^{\text{test}}$ ,  $Y_q^{\text{test}}$ );
3 if non_log_cols is not empty then
4     scaler  $\leftarrow$  StandardScaler();
5     Fit scaler on the non_log_cols of  $X_q^{\text{train}}$  and transform them;
6     Transform the non_log_cols of  $X_q^{\text{test}}$  using the fitted scaler;
7 end

/* Hyperparameter Tuning */
8 Define a hyperparameter search space  $H$  for MLP architecture;
9 Initialize a Bayesian Optimization tuner to search space  $H$ , with the
   objective of minimizing the validation mean squared error (val_mse);

/* Model Search and Selection */
10 Execute the hyperparameter search on the training set, using the test set for
    validation;
    // The tuner iteratively builds, compiles, and trains MLP
       models. Each model is trained using the Adam optimizer and
       employs Early Stopping to prevent overfitting.
11  $\hat{f}(\cdot; \hat{\theta}) \leftarrow$  Retrieve the best model with the lowest val_mse;
```

While the search space can be expanded for better performance, We impose certain constraints to account for computational limitations. These constraints primarily involve setting upper bounds on the number of hidden layers, neurons, and training epochs. The default configuration uses up to 5 hidden layers and 512 neurons per layer, with the number of epochs³⁵ set to 2,000. While the exact values may vary by country, these defaults provide a reasonable balance between model flexibility and computational efficiency. The candidate activation functions include Rectified Linear Unit (ReLU), hyperbolic tangent (Tanh), Exponential Linear Unit (ELU), Scaled Exponential Linear Unit (SELU), and Swish, with the final layer using a linear function to reflect the continuous nature of the GDP growth.

³⁵An epoch refers to one complete pass through the entire training dataset during model training (James et al., 2023).

C.2.2 LSTM: Long Short-Term Memory

Algorithm 4: LSTM Training and Hyperparameter Optimization

Data: `X_q_processed`; `Y_q_processed`; `non_log_cols`; `TIME_STEPS`

Result: The best-performing trained LSTM model, $\hat{f}_{\text{LSTM}}(\cdot; \hat{\theta})$

```

/* Split Dataset and Standardization */
1 Split X_q_processed and Y_q_processed into training and test sets;
2 Standardize non_log_cols using a scaler fitted only on the training data;
/* Create Time Series Dataset */
3 for each dataset (training and test) do
4   Transform the 2D data  $(N, k)$  into a 3D sequential dataset  $(N - T, T, k)$ 
   using a sliding window of size TIME_STEPS;
5   Store as  $X_{\text{LSTM}}^{\text{train}}, Y_{\text{LSTM}}^{\text{train}}, X_{\text{LSTM}}^{\text{test}}, Y_{\text{LSTM}}^{\text{test}}$ ;
6 end
/* Hyperparameter Tuning */
7 Define a search space  $H$  for LSTM architecture (e.g., layers, units), dropout
  rates, and learning rate;
8 Initialize a Bayesian Optimization tuner to search space  $H$ , with the
  objective of minimizing val_mse;
/* Model Search and Selection */
9 Execute the hyperparameter search on  $X_{\text{LSTM}}^{\text{train}}$ , using  $X_{\text{LSTM}}^{\text{test}}$  for validation;
  // The tuner iteratively builds and trains LSTM models using the
  Adam optimizer and Early Stopping.
10  $\hat{f}_{\text{LSTM}}(\cdot; \hat{\theta}) \leftarrow$  Retrieve the best model with the lowest val_mse;

```

The hyperparameter tuning for the LSTM is also conducted using TensorFlow’s `keras-tuner` library, making the underlying code structure very similar to that of the MLP. However, since the LSTM takes into account historical data, we must define the length of the input sequence, denoted as t , and reshape the dataset into a 3D structure accordingly. A larger value of t allows the model to learn from a longer history, but it also results in the loss of more observations—especially important in our case, as training is performed on quarterly data. Taking this trade-off into account, we set the default value of t to 4.

C.2.3 XGBoost: Extreme Gradient Boosting

Algorithm 5: XGBoost Training and Hyperparameter Optimization

Data: `X_q_processed`; `Y_q_processed`; `non_log_cols`
Result: The best-performing trained XGBoost model, $\hat{f}_{\text{XGB}}(\cdot; \hat{\theta})$

```

/* Split Dataset and Standardization */
1 Split X_q_processed and Y_q_processed into training and test sets;
2 Standardize non_log_cols using a scaler fitted only on the training data;
/* Hyperparameter Tuning */
3 Define a hyperparameter grid  $H$  for XGBoost (e.g., n_estimators,
   max_depth, learning_rate, subsample);
4 Initialize a GridSearchCV or RandomizedSearchCV instance with an
   XGBRegressor estimator, the search space  $H$ , and  $K$ -fold cross-validation;
/* Model Search and Selection */
5 Execute the hyperparameter search by fitting the search instance on the
   training set  $(X_q^{\text{train}}, Y_q^{\text{train}})$ ;
6  $\hat{f}_{\text{XGB}}(\cdot; \hat{\theta}) \leftarrow$  Retrieve the best estimator (model) from the search results that
   yielded the best cross-validation score;
```

XGBoost is implemented using the `xgboost` library in Python. As with other models, its key hyperparameters—such as tree depth (which controls the ability to capture complex relationships), step size (to prevent overfitting), and the minimum loss reduction required for further partitioning—must be carefully tuned by the researcher.

To reduce computation time, one can use `RandomizedSearchCV`, which tests a random subset of hyperparameter combinations and selects the best among them. Alternatively, `GridSearchCV` allows the user to explicitly define a wide search space, offering more precise tuning at the cost of significantly greater computation time.

A hybrid approach—conducting a randomized search first, then performing a grid search around the best-performing region—can strike a balance between efficiency and accuracy.

In this study, we prioritize predictive accuracy and therefore apply `GridSearchCV` from the outset. However, `RandomizedSearchCV` or a combined approach may be more appropriate in cases where computational efficiency is critical.

C.2.4 Elastic Net

Algorithm 6: Hyperparameter Tuning and Initial Estimation

Data: Quarterly dataset `X_q_processed`, `Y_q_processed`

Result: Optimal hyperparameters $(\lambda_1^*, \lambda_2^*)$ and coefficient estimates $\hat{\beta}_{\text{orig}}$

```

/* Split into Train and Test Sets */
1 Define training and test set:  $(X_{\text{train}}, Y_{\text{train}}), (X_{\text{test}}, Y_{\text{test}})$ ;
2 Standardize non_log_cols using StandardScaler in each set;
/* Define Penalized Loss Function */
3 Elastic Net loss:  $\mathcal{L}(\beta) = \sum_q (Y_q - X_q^\top \beta)^2 + \lambda_1 \sum_j |\beta_j| + \lambda_2 \sum_j \beta_j^2$ ;
/* Cross Validation on Training Set */
4 Define candidate values for  $(\alpha, \rho)$ .;
5 Map to  $(\lambda_1 = \alpha\rho, \lambda_2 = \alpha(1 - \rho))$ .;
6 Split  $(X_{\text{train}}, Y_{\text{train}})$  into  $K$  sequential folds.;
7 for each pair  $(\alpha, \rho)$  in the grid do
8     Initialize total_cv_error  $\leftarrow 0$ ;
9     for each fold  $k = 1$  to  $K$  do
10         Use folds 1 to  $k$  as training set; fold  $k + 1$  as validation set.;
11         Fit ElasticNet $(\alpha, \rho)$  and obtain  $\hat{\beta}^{(k)}$ ;
12         Predict on validation set and compute MSE  $e_k$ ;
13         total_cv_error  $\leftarrow$  total_cv_error +  $e_k$ ;
14     end
15     if total_cv_error < min_cv_error then
16         Update best  $(\alpha^*, \rho^*)$ ;
17     end
18 end

/* Initial Model Estimation on the Training Set */
19 Fit ElasticNet $(\alpha^*, \rho^*)$  on  $(X_{\text{train}}, Y_{\text{train}})$  to obtain  $\hat{\beta}_{\text{initial}}$ ;
20 Return  $(\lambda_1^* = \alpha^* \rho^*, \lambda_2^* = \alpha^* (1 - \rho^*))$  and  $\hat{\beta}_{\text{initial}}$ ;

```

Algorithm 7: Bootstrap Estimation

Data: Original training data ($\mathbf{X_q_train}, \mathbf{Y_q_train}$), optimized hyperparameters (λ_1^*, λ_2^*), transformed monthly explanatory variables $\mathbf{X_prime_m_df}$

Result: Monthly GDP estimates \hat{y}_m and optional confidence intervals

```
/* 3.2 Bootstrap Procedure and Final Coefficient Estimation */
1  $B \leftarrow 50000$  ; /* Number of bootstrap */
2 Initialize list beta_hat_bootstrap_list;
3 for  $b = 1$  to  $B$  do
4    $(X_{q,b}^*, Y_{q,b}^*) \leftarrow$  Resample with replacement from  $\mathbf{X\_q\_train}, \mathbf{Y\_q\_train}$ ;
5    $\hat{\beta}_b^* \leftarrow \text{ElasticNet}(X_{q,b}^*, Y_{q,b}^*, \lambda_1^*, \lambda_2^*)$ ;
6   Add  $\hat{\beta}_b^*$  to beta_hat_bootstrap_list;
7 end
8  $\hat{\beta} \leftarrow \frac{1}{B} \sum_{b=1}^B \hat{\beta}_b^*$  ; /* Use average */
/* 4. Monthly GDP Growth Estimation */
9  $\mathbf{X\_prime\_m\_for\_prediction} \leftarrow \mathbf{X\_prime\_m\_df}$ ;
10  $\hat{y}_m \leftarrow (\mathbf{X}_m')^\top \hat{\beta}$ ;
11 for each month  $m$  do
12   Initialize list y_m_bootstrap_predictions;
13   for each  $\hat{\beta}_b^*$  in beta_hat_bootstrap_list do
14     Add  $(\mathbf{X}_m')^\top \hat{\beta}_b^*$  to y_m_bootstrap_predictions;
15   end
16   Compute 2.5th and 97.5th percentiles  $\rightarrow$  CI for  $\hat{y}_m$ ;
17 end
```

C.3 Monthly GDP Estimation

Algorithm 8 produces a preliminary $\tilde{Y}_m = \hat{f}(X_m; \hat{\theta})$. Each \tilde{Y}_m is an estimate on a quarterly scale and their aggregation is not arithmetically consistent with the observed quarterly totals. For each quarter, an adjustment factor (k_q) is calculated by taking the ratio of the actual quarterly growth to the sum of the preliminary monthly estimates within that quarter. This factor is then applied to scale each preliminary estimate, yielding the final adjusted monthly growth rate, \hat{y}_m .

Algorithm 8: Monthly GDP Growth Estimation and Reconciliation

Data: $\hat{f}(\cdot; \hat{\theta})$; `X_prime_m_df`; `Y_q_processed`; `non_log_cols`
Result: The final series of adjusted monthly GDP growth estimates \hat{y}_m

/ Prepare Monthly Input Data */*

1 `X_prime_m_scaled` \leftarrow Create a copy of the relevant columns from `X_prime_m_df`;
2 **if** `non_log_cols` is not empty **then**
3 Transform the `non_log_cols` of `X_prime_m_scaled`;
4 **end**

/ Generate Preliminary Monthly Estimates */*

5 $\tilde{Y}_m \leftarrow \hat{f}(\text{X_prime_m_scaled}; \hat{\theta})$ for all months m ;
6 $\hat{p}_m \leftarrow \tilde{Y}_m/3$ for all months m ; *// Calculate naive monthly estimates*
/ Reconcile Monthly Estimates with Quarterly Actuals */*

7 Initialize an empty series for final monthly estimates, \hat{y}_m ;
8 **for** each quarter q **do**
9 $\text{sum_naive} \leftarrow \sum_{j \in q} \hat{p}_j$; *// Sum of naive estimates in the quarter*
10 **if** quarter q has a corresponding value in `Y_q_processed` **then**
11 $Y_q \leftarrow$ The actual observed growth for quarter q ;
12 **if** $\text{sum_naive} \neq 0$ **then**
13 $k_q \leftarrow Y_q/\text{sum_naive}$; *// Calculate adjustment factor*
14 **for** each month m in quarter q **do**
15 $\hat{y}_m \leftarrow k_q \cdot \hat{p}_m$; *// Apply adjustment factor*
16 **end**
17 **else**
18 *// Distribute actual growth evenly if sum of naive estimates is zero*
19 $\hat{y}_m \leftarrow Y_q/(\text{number of months in } q)$ for each month $m \in q$;
20 **end**
21 **end**
22 *// Use naive estimates if no actual quarterly data*
23 $\hat{y}_m \leftarrow \hat{p}_m$ for each month $m \in q$;
24 **end**

D Reconciliation with Quarterly Aggregates

This section demonstrate the reconciliation method described in Section 2.1. Denton (1971) provides the main idea of it. The preliminary monthly estimates, contained in the vector $\tilde{Y} \in \mathbb{R}^{T \times 1}$ (where T is the total number of months), are reconciled to be consistent with the observed quarterly growth totals, given by the vector $Y \in \mathbb{R}^{Q \times 1}$ (where Q is the total number of quarters, and $T = 3Q$). This consistency is enforced by the temporal aggregation constraint:

$$J\hat{y} = Y, \quad (\text{D1})$$

where $\hat{y} \in \mathbb{R}^{T \times 1}$ is the vector of final adjusted monthly estimates we wish to find. The matrix $J \in \mathbb{R}^{Q \times T}$ is the temporal aggregation matrix that sums the monthly data into quarterly totals. For monthly-to-quarterly aggregation, J takes the form:

$$J = \begin{pmatrix} 1 & 1 & 1 & 0 & 0 & 0 & \cdots & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 1 & 1 & \cdots & 0 & 0 & 0 \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots & \ddots & \vdots & \vdots & \vdots \\ 0 & 0 & 0 & 0 & 0 & 0 & \cdots & 1 & 1 & 1 \end{pmatrix} = I_Q \otimes \mathbf{1}_3^\top, \quad (\text{D2})$$

where I_Q is the $Q \times Q$ identity matrix, $\mathbf{1}_3^\top$ is a 1×3 row vector of ones, and \otimes denotes the Kronecker product.

Our proportional benchmarking method assumes that the final adjusted estimate for each month, \hat{y}_m , is proportional to its preliminary estimate, \tilde{Y}_m . The proportionality factor is constant for all months within the same quarter. This relationship can be expressed as:

$$\hat{y}_m = k_q \cdot \tilde{Y}_m, \quad \forall m \in q, \quad (\text{D3})$$

where k_q is the adjustment factor for quarter q . In matrix notation, this is equivalent to:

$$\hat{y} = \text{diag}(\tilde{Y}) J^\top k, \quad (\text{D4})$$

where $\text{diag}(\tilde{Y})$ is a $T \times T$ diagonal matrix with the preliminary estimates \tilde{Y}_m on its diagonal, and $k \in \mathbb{R}^{Q \times 1}$ is the vector of quarterly adjustment factors.

To find the adjustment factors, we substitute this expression for \hat{y} back into the

aggregation constraint:

$$J \left(\text{diag}(\tilde{Y}) J^\top k \right) = Y. \quad (\text{D5})$$

By solving for k , we obtain the vector of quarterly adjustment factors:

$$k = \left(J \text{diag}(\tilde{Y}) J^\top \right)^{-1} Y. \quad (\text{D6})$$

The matrix $(J \text{diag}(\tilde{Y}) J^\top)$ is a $Q \times Q$ diagonal matrix where each diagonal element is the sum of the preliminary monthly estimates for the corresponding quarter, i.e., $\sum_{j \in q} \tilde{Y}_j$. This makes its inversion straightforward.

Finally, substituting the solution for k back gives the closed-form solution for the final adjusted monthly estimates:

$$\hat{y} = \text{diag}(\tilde{Y}) J^\top \left(J \text{diag}(\tilde{Y}) J^\top \right)^{-1} Y. \quad (\text{D7})$$

It ensures that the final monthly estimates \hat{y} satisfy the quarterly constraints while preserving the relative intra-quarter dynamics provided by the preliminary indicator series \tilde{Y} .

E Simulation on MLP and LSTM with Synthetic Data

This section presents a simulation study designed to compare the performance of MLP and LSTM models under distinct data generating processes. We construct synthetic datasets to examine how temporal aggregation from monthly to quarterly frequency may influence the relative performance of sequence-based models such as LSTM and MLP. We consider two versions of the explanatory variables: one with stationary white noise and the other with nonstationary random walk processes.

E.1 Explanatory Variables

Let $\mathbf{X}_m \in \mathbb{R}^{n \times k}$ denote a panel of monthly explanatory variables, where $k = 5$. We consider two versions of \mathbf{X}_m :

- (a) **Stationary specification:** each explanatory variable follows an i.i.d. standard normal distribution:

$$X_{k,m} \sim \mathcal{N}(0, 1), \tag{E1}$$

independently across k and m .

- (b) **Nonstationary specification:** each variable follows a random walk process:

$$X_{k,m} = X_{k,m-1} + \epsilon_{k,m}, \quad \epsilon_{k,m} \sim \mathcal{N}(0, 1), \tag{E2}$$

with $X_{k,0} = 0$ for all k .

Although five regressors are generated in both cases, only X_0 , X_1 , and X_2 are used to construct the target variable. The remaining two regressors (X_3 , X_4) are included as irrelevant covariates to mimic the presence of uninformative controls in empirical datasets.

E.2 Data Generating Processes

We generate two distinct versions of the latent monthly GDP series y_m^* , each reflecting a different structural relationship between the explanatory variables and the outcome.

E.2.1 Scenario A: Contemporaneous Relationship (MLP-Advantaged)

In this case, monthly GDP depends non-linearly on explanatory variables³⁶ from the same period:

$$y_m^* = 1.5 \sin(x_{1,m}) + 0.8x_{2,m}^2 - 0.5x_{3,m} + 100 + \nu_m, \quad \nu_m \sim \mathcal{N}(0, 1^2). \quad (\text{E3})$$

This specification captures a static but non-linear dependence structure that should be favorable to feedforward models such as MLP.

E.2.2 Scenario B: Long-Term Dependency (LSTM-Advantaged)

In contrast, this scenario introduces a long-range dependency by making y_m^* depend on a 4-month moving average of past values of X_0 and X_1 :

$$y_m^* = 2 \cos \left(\frac{1}{L} \sum_{l=1}^L x_{1,m-l} \right) + \left(\frac{1}{L} \sum_{l=1}^L x_{2,m-l} \right) + 100 + \xi_m, \quad \xi_m \sim \mathcal{N}(0, 1^2), \quad (\text{E4})$$

where $L = 4$. Note that the outcome is undefined for the first L months and is excluded from the evaluation sample in this case.

E.3 Quarterly Aggregation

To generate quarterly GDP, we aggregate the monthly series using the conventional identity:

$$Y_q^* = \sum_{m \in q} y_m^*. \quad (\text{E5})$$

Only the quarterly aggregates Y_q^* are observable to the model during training, whereas the underlying monthly series y_m^* remain latent.

E.4 Simulation Results

The simulation results, summarized in Table E1, present a nuanced picture that challenges our initial hypotheses but ultimately provides a deeper insight into the relative strengths of MLP and LSTM in a mixed-frequency context. Contrary to the expectation that MLP would be favored in contemporaneous settings, the LSTM

³⁶The additive constant of 100 ensures that y_m^* remains in a positive GDP-level range.

model demonstrates superior performance in three out of the four scenarios, judged by lower RMSE and MAE, and higher correlation. The most striking result is LSTM’s strong outperformance in the Scenario A Stationary case, which was designed as a simple non-linear function approximation task.

Table E1. Simulation Performance by Scenario and Model

Scenario	Type	Model	MAE	RMSE	Corr.	Sign Acc.
A	Stationary	MLP	5.720	35.202	0.008	0.510
		LSTM	2.947	4.859	0.053	0.524
	Nonstationary	MLP	41.036	135.279	0.730	0.668
		LSTM	22.731	55.980	0.933	0.628
B	Stationary	MLP	5.975	48.775	0.050	0.518
		LSTM	1.515	2.390	0.148	0.541
	Nonstationary	MLP	1.592	2.201	0.953	0.538
		LSTM	2.357	4.819	0.825	0.555

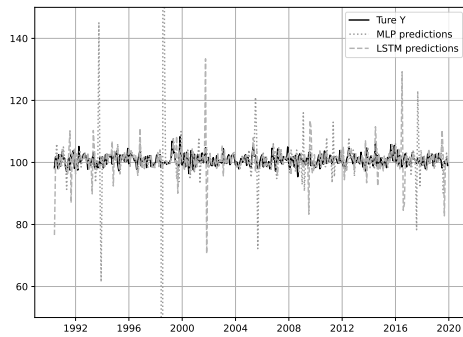
Note: Scenario A and B represent contemporaneous relationship and long-term dependency, respectively. MAE (Mean Absolute Error) and RMSE (Root Mean Squared Error) are scale-dependent measures of forecast accuracy where lower values indicate better performance. Correlation measures the linear association between predicted and actual values. Sign Accuracy refers to the proportion of times the model correctly predicts the direction of change.

The sole exception, and the key to understanding these results, is the Scenario B Nonstationary case, where the MLP model decisively outperforms the LSTM. This scenario is arguably the most realistic, as it models an economy where multiple trending (nonstationary) variables interact with complex time lags. The MLP’s success here suggests that while temporal aggregation from monthly to quarterly data may obscure the precise, high-frequency lag structure that LSTMs are designed to capture, it preserves the stronger, low-frequency signal of co-movement among the variables. The MLP, acting as a powerful static function approximator, effectively learns this robust relationship between the quarterly-aggregated levels, leading to superior performance. This finding may support our main paper’s hypothesis: the information loss inherent in temporal aggregation can neutralize the traditional advantages of sequence models.

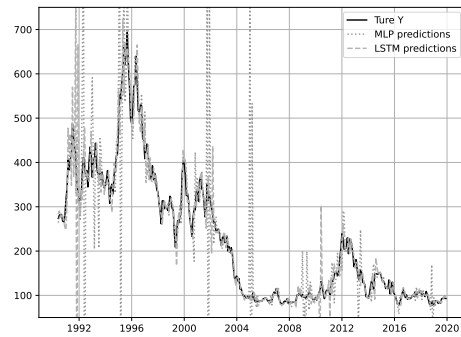
The LSTM’s dominance in the other three scenarios can be understood in this light. In the stationary cases, the absence of strong underlying trends creates an idealized environment. In the Scenario B Stationary case, LSTM easily captures

the clean, textbook long-term dependency. In the Scenario A Stationary case, its victory is likely attributable to its greater inherent complexity, allowing it to better approximate the specific non-linear function than the tested MLP architecture. In the Scenario A Nonstationary case, although the relationship between y_m^* and \mathbf{X}_m is contemporaneous, the inputs \mathbf{X}_m themselves follow a strong random-walk process. The LSTM exploits its sequential nature to learn the dynamics of the input series itself, giving it a decisive edge over the non-sequential MLP.

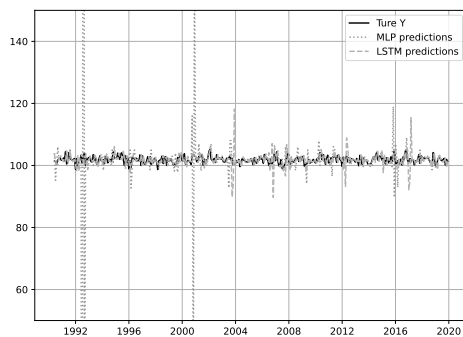
In conclusion, this simulation study provides compelling evidence that the empirical success of the MLP framework in our paper is not accidental. It highlights that the MLP's strength lies in its ability to robustly model the strong, static relationships that persist between co-trending macroeconomic variables, even after the nuances of high-frequency dynamics are lost through aggregation. The LSTM, while powerful, may be handicapped in this specific mixed-frequency setting, where the very act of preparing its target variable damages the granular information it is best suited to process.



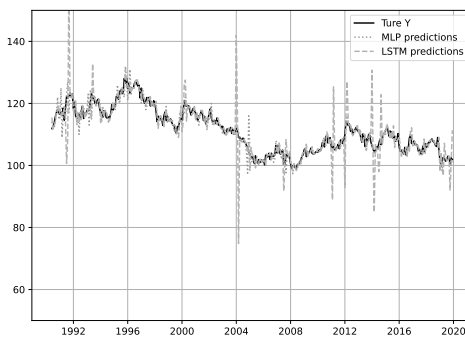
(a) Scenario A – Stationary



(b) Scenario A – Nonstationary



(c) Scenario B – Stationary



(d) Scenario B – Nonstationary

Figure E1. Simulation Results by Scenario and Model

F Estimation Results of MLP Model with Chinese CLI data

Concerns have been raised regarding the reliability of China’s official statistics, including GDP (Holz, 2014; Rawski, 2001). In response, alternative measures of economic activity have been proposed, such as the so-called “Li Keqiang Index,” which was inspired by remarks from Premier Li Keqiang and aims to proxy China’s actual economic conditions in place of official GDP figures (Song & He, 2015).³⁷

In Section 3, China shows the weakest test performance, which we primarily attribute to the limited number of explanatory variables and shorter time coverage. However, if the official GDP data fail to reflect true economic conditions—that is, if there are concerns not only about the quantity but also the quality of data—then the issue lies on a different level.

To investigate this, we re-estimate the MLP model using the same set of explanatory variables but replacing the official GDP series with the Composite Leading Indicator (CLI) for China, published by the OECD.³⁸ Although the CLI is a monthly index, we take quarterly averages to construct a comparable quarterly target series. As with the original GDP series, we apply log-differencing to the index level, and we keep all other experimental settings—such as timeframes and hyperparameter search settings—unchanged. We then compare the results with the original model using the official GDP.

Table F1. MLP Test Performance (China)

Dependent Var.	MSE	RMSE	MAE	R^2	Sign Acc.	sMAPE	Theil U1
GDP	0.006	0.077	0.060	0.628	0.875	78.218	0.335
CLI	0.000	0.007	0.006	0.794	0.792	118.143	0.224

Note: Mean Squared Error (MSE), Root Mean Squared Error (RMSE), Mean Absolute Error (MAE), and symmetric Mean Absolute Percentage Error (sMAPE, %) are all scale-dependent error metrics where lower values indicate better predictive accuracy. The coefficient of determination (R^2) measures the proportion of variance explained by the model and is better when closer to one. Sign Accuracy indicates how often the model correctly predicts the direction of change and is also better when higher. Theil’s U1 statistic compares the forecast accuracy of a model to that of a naive (no-change) forecast; values closer to zero indicate better performance.

³⁷Although the Economist magazine is often credited with constructing the Li Keqiang Index, it appears to use the index primarily in reporting rather than publishing the underlying data. Therefore, we rely on alternative sources for this study.

³⁸The CLI series is available through the FRED API under the code CHNLORSGPNOSTSAM. <https://fred.stlouisfed.org/series/CHNLORSGPNOSTSAM>.

Table F2. China Benchmark with CLI

Model	MAE	RMSE	Corr.	Sign Acc.
MLP	0.004	0.018	0.040	0.818

Note: MAE (Mean Absolute Error) and RMSE (Root Mean Squared Error) are scale-dependent measures of forecast accuracy where lower values indicate better performance. Correlation measures the linear association between predicted and actual values. Sign Accuracy refers to the proportion of times the model correctly predicts the direction of change. Since the CLI is based on level data, sign accuracy is not directly comparable. To enable meaningful comparison, we transformed the series according to equation (7), as done in the US benchmark.

As shown in Table F1, when the dependent variable is changed from official GDP to the CLI while keeping all other conditions constant, the test performance of the MLP model improves noticeably. While a direct comparison between the two dependent variables is inherently difficult due to their structural differences, this result suggests that part of the weaker performance observed in the China model in Section 3 may be attributable to the quality of the official GDP data, rather than solely to the quantity or coverage of explanatory variables.³⁹

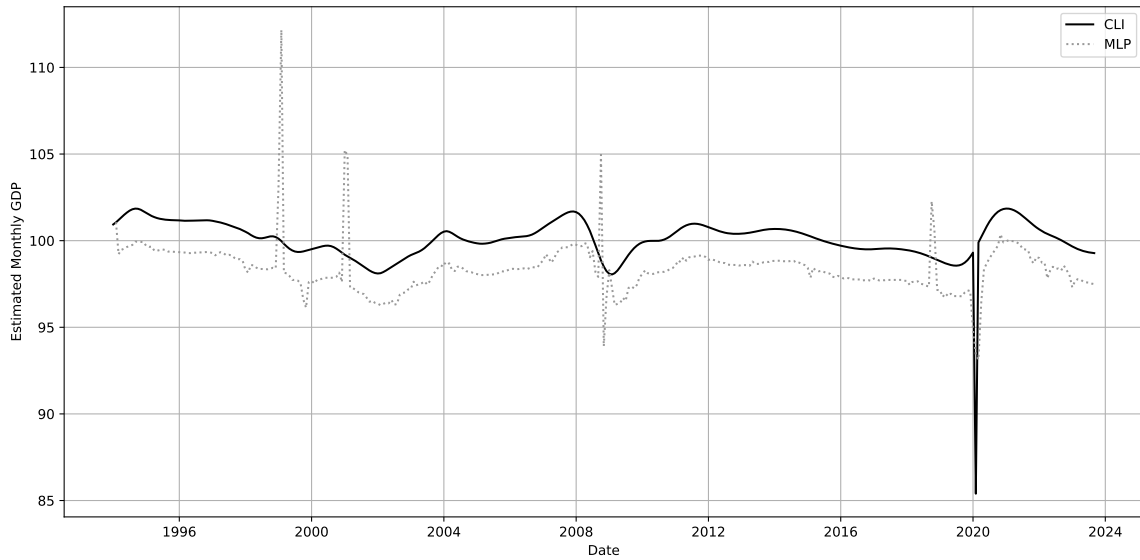


Figure F1. Actual and Predicted CLI (MLP)

In addition, the MLP model's predictions of monthly CLI growth follow the overall trends of the actual CLI series reasonably well. Notably, the model captures the

³⁹This finding may also be interpreted as additional circumstantial support for concerns regarding the reliability of China's official GDP statistics. However, a detailed discussion of this issue lies beyond the scope of this paper.

sharp contraction and rebound during the COVID-19 period—despite this period not being included in the training data—suggesting a degree of generalizability to large economic shocks. However, outside major episodes, the model tends to exhibit somewhat exaggerated responses to minor fluctuations compared to the actual CLI, indicating potential oversensitivity in less volatile periods.

G VIIRS Nighttime Light Data

This section describes the nighttime light (NTL) data used in the paper. While there are several sources of nighttime light data, we obtain the data from the Earth Observation Group at the Payne Institute for Public Policy, Colorado School of Mines,⁴⁰ via the Google Earth Engine API.⁴¹

There are two major sources of NTL data: the Visible and Infrared Imaging Suite (VIIRS) and the US Air Force Defense Meteorological Satellite Program (DMSP). DMSP data are available from 1992 to 2013, while VIIRS data are available from 2014 to the present. VIIRS offers higher spatial resolution. However, due to differences in imaging time and other technical specifications, the two datasets are not directly comparable and thus cannot be combined into a single continuous series without proper preprocessing. Therefore, we use the VIIRS dataset in this study.⁴²

We extract the country’s land boundary using the FAO Global Administrative Unit Layers (GAUL), thereby excluding brightness over oceans. Then, for each month, we aggregate the brightness values of all pixels within the boundary. Since the dataset includes basic preprocessing procedures such as cloud masking, it can be used directly without major adjustments. However, because the aggregation involves summing pixel values, countries with larger land areas naturally show higher total brightness. Therefore, proper normalization is required for cross-country comparisons.

The data are available starting from January 2014. To ensure compatibility with other datasets, we collect VIIRS data up to March 2024 for the US and to September 2023 for China, which corresponds to the time span of the main data used in our study. A statistical description of the data is provided in Table G1.

Table G1. NTL Data Summary Statistics

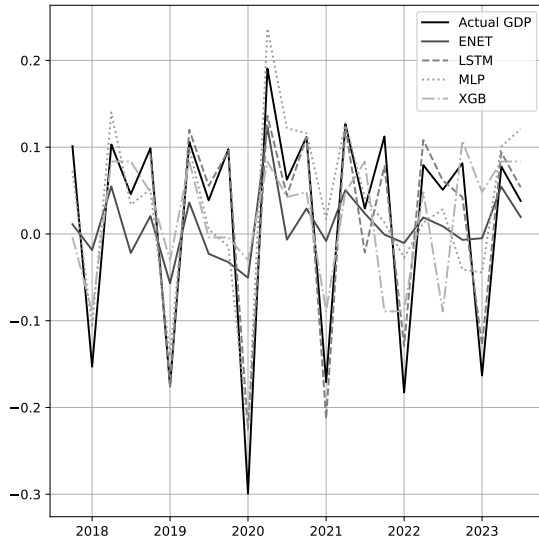
Country	Variable	Obs.	Mean	Std. Dev.	Min	Max
China	NTL	117	3.00e+07	8747270	1.48e+07	4.92e+07
UK	NTL	123	3355302	477328.9	2261410	4977409

⁴⁰<https://eogdata.mines.edu/products/vnl/#monthly>.

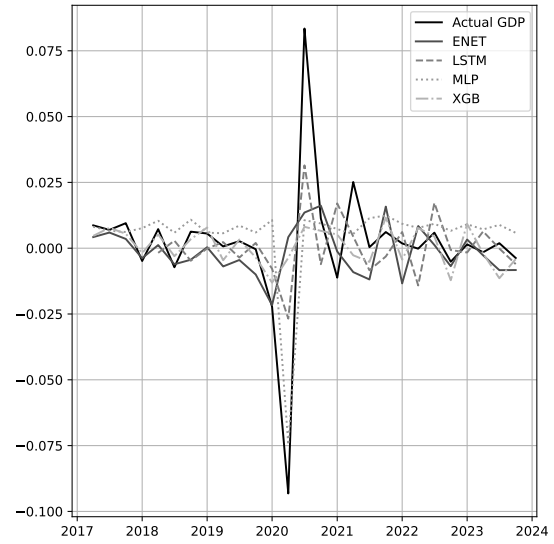
⁴¹https://developers.google.com/earth-engine/datasets/catalog/NOAA_VIIRS_DNB_MONTHLY_V1_VCMSLCFG.

⁴²The code for downloading the data via the API is provided in the replication package which is available at https://github.com/Yonggeun-Jung/monthly_gdp. Access to the Google Earth Engine API requires separate user registration, which is free for research purposes.

H Supplementary Figures and Tables



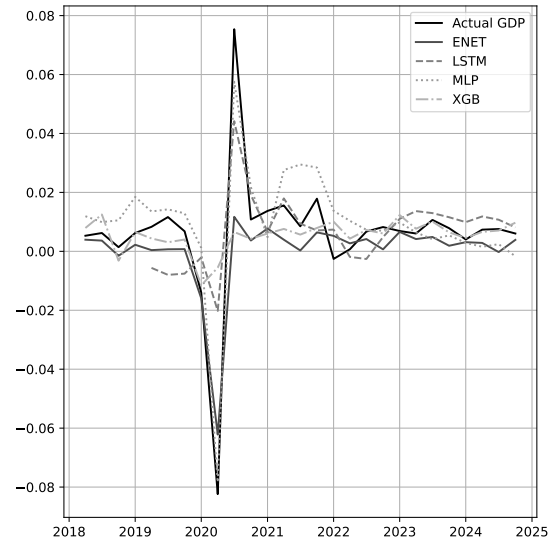
(a) China



(b) Germany

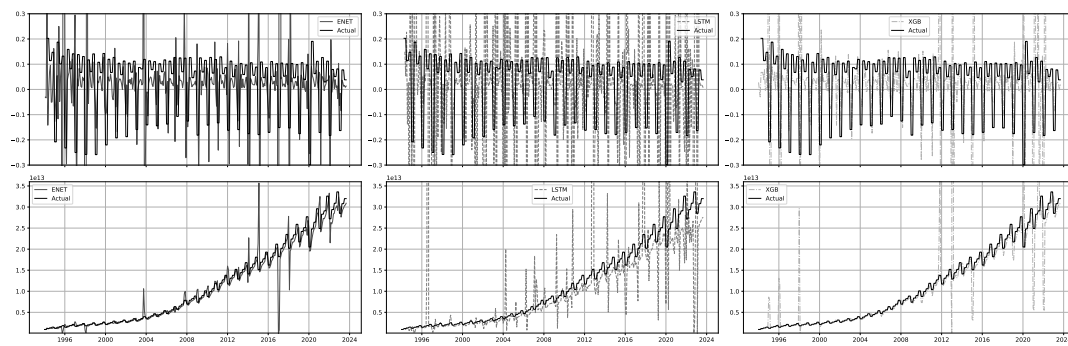


(c) UK

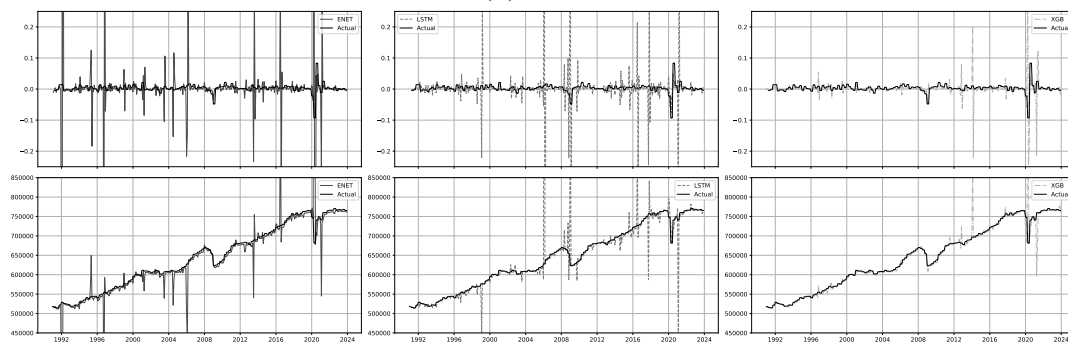


(d) US

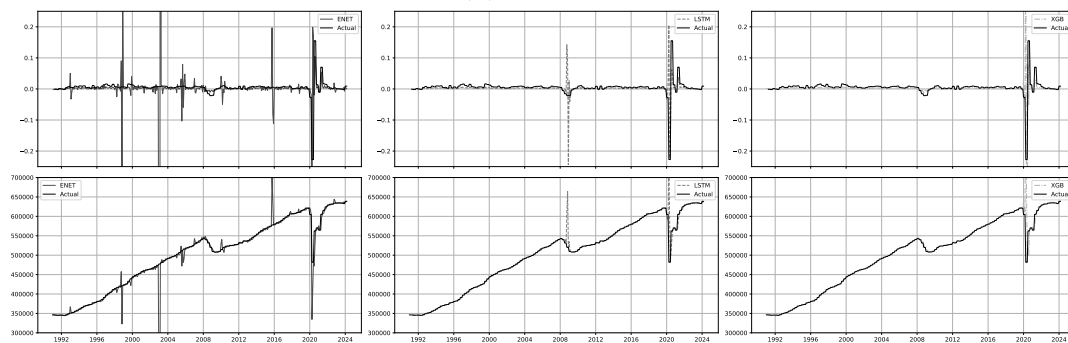
Figure H1. Test Performance of Each Model



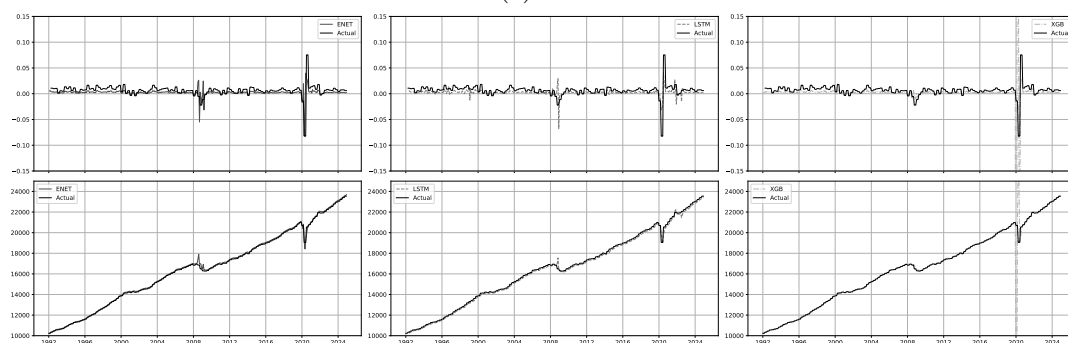
(a) China



(b) Germany



(c) UK



(d) US

Figure H2. Monthly GDP Estimation Results (Ex. MLP): Growth (U) and Level (D)

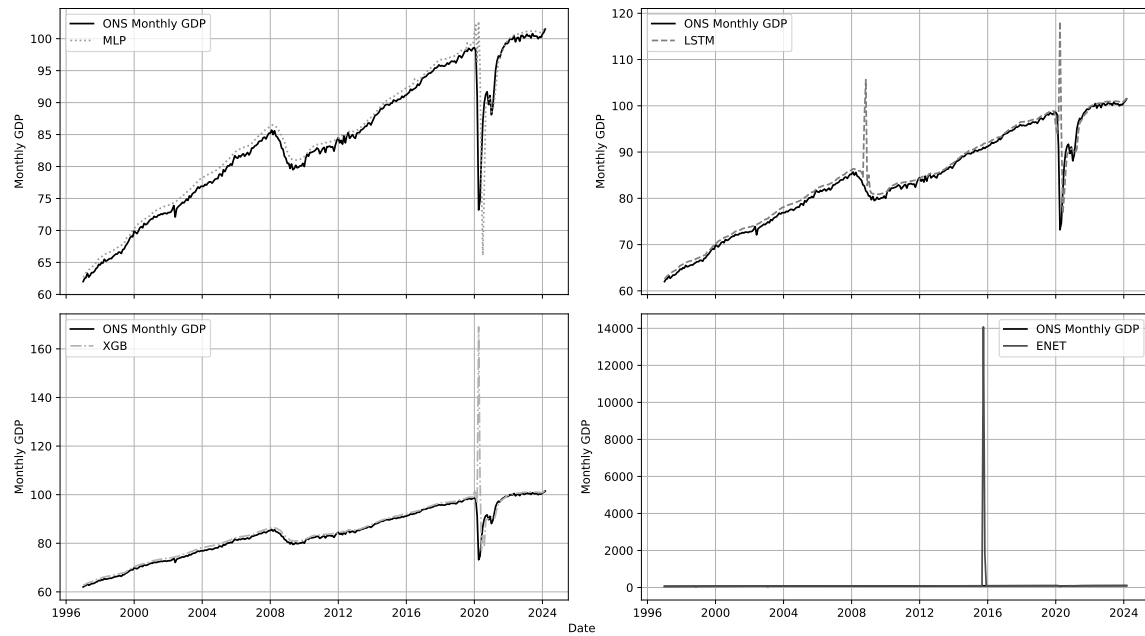


Figure H3. Benchmark to ONS: UK Estimated Monthly GDP (All Models)

Table H1. Elastic Net and XGBoost Results

Variable	China		Germany		UK		US	
	ENET	XGB	ENET	XGB	ENET	XGB	ENET	XGB
CPI	-0.044 (-0.10,0.00)	0.47	0.000 (0.00, 0.00)	0.09	-0.018 (-0.07, 0.00)	0.03	0.057 (-0.03, 0.17)	0.04
Retail Trade			0.000 (0.00, 0.00)	0.07				
M1 Money Stock					0.033** (0.00, 0.06)	0.04	-0.029 (-0.09, 0.02)	0.04
M2 Money Stock							-0.065 (-0.16, 0.00)	0.06
Total Reserves	0.000 (0.00, 0.00)	0.00	-0.000 (0.000,0.000)	0.05	-0.002 (-0.002,0.018)	0.03		
Exchange Rate	0.000 (0.00,0.00)	0.00			0.022* (0.00, 0.05)	0.05		
Unemployment Rate			0.000 (0.00, 0.00)	0.07	-0.076*** (-0.10, -0.05)	0.15	-0.001* (-0.00, -0.00)	0.06
Employment Rate					0.001 (0.00, 0.01)	0.16		
Labor Force							-0.027 (-0.13, 0.02)	0.06
Avg. Weekly Hours							0.000 (-0.00, 0.00)	0.03
Industrial Production							0.139*** (0.07, 0.22)	0.07
Production Volume			0.000 (0.00, 0.00)	0.16	0.040* (0.00, 0.09)	0.08		
Manufacturing			0.007*** (0.00, 0.01)	0.10				
PPI	0.000 (0.00,0.00)	0.25						
Real Personal Consum.							0.075 (0.00, 0.19)	0.20
Interest Rate (10Y Gov)					-0.000 (-0.00, 0.00)	0.04	-0.002 (-0.02, 0.01)	0.02
Interest Rate (Gov Bond)					0.001 (-0.00, 0.01)	0.02		
Effect. Fund Rate							-0.000 (-0.00, 0.00)	0.08
Effect. Ex. Rate	0.000 (0.00, 0.00)	0.00	0.000 (0.00, 0.00)	0.04				
S&P/FTSE/DAX/SSEC	-0.012 (-0.10,0.00)	0.00	-0.000 (-0.01, 0.00)	0.06	0.008 (0.00, 0.02)	0.25	0.019** (0.00, 0.03)	0.10
Share Prices					0.006 (0.00, 0.02)	0.03		
Moody's AAA Bond							0.011 (-0.01, 0.04)	0.07
(Net) Exports	0.001 (0.00,0.00)	0.00	0.000 (0.00, 0.00)	0.21	0.016** (0.00, 0.03)	0.05	0.001 (-0.03, 0.03)	0.05
Imports	-0.000 (0.00,0.00)	0.00	-0.000 (0.00, 0.00)	0.04	0.018 (-0.00, 0.04)	0.05		
US Imports from China	0.096 (0.00, 0.23)	0.28						
Price Competitiv.			0.000 (0.00, 0.00)	0.06				
Policy Uncertainty	-0.015 (-0.08,0.03)	0.00	-0.001 (-0.00, 0.00)					
Personal Savings Rate							0.000 (-0.00, 0.00)	0.02
Nonfarm Payroll Emp.							0.095* (0.00, 0.19)	0.09

Note: * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$. 95% bootstrap confidence intervals in parentheses. Coefficients rounded to 3 decimal places. Bootstrap based on 50,000 replications.

References for Appendices

- Chen, T., & Guestrin, C. (2016). XGBoost: A Scalable Tree Boosting System. *Proceedings of the 22nd ACM SIGKDD International Conference on Knowledge Discovery and Data Mining*, 785–794.
- Denton, F. T. (1971). Adjustment of Monthly or Quarterly Series to Annual Totals: An Approach Based on Quadratic Minimization. *Journal of the American Statistical Association*, 66(333), 99–102.
- Hansen, B. (2022). *Econometrics*. Princeton University Press.
- Hochreiter, S., & Schmidhuber, J. (1997). Long Short-term Memory. *Neural Computation*, 9(8), 1735–1780.
- Holz, C. A. (2014). The Quality of China’s GDP Statistics. *China Economic Review*, 30, 309–338.
- James, G., Witten, D., Hastie, T., Tibshirani, R., & Taylor, J. (2023). *An Introduction to Statistical Learning with Applications in Python*. Springer.
- Kingma, D. P. (2014). Adam: A Method for Stochastic Optimization. *arXiv preprint arXiv:1412.6980*.
- Rawski, T. G. (2001). What is Happening to China’s GDP Statistics? *China Economic Review*, 12(4), 347–354.
- Song, X., & He, Z. (2015). The Keqiang Index: a New Benchmark for China’s Development. *Social Indicators Research*, 123, 661–676.