## State Legislatures, Term Limits, and Polarization\*

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#### Abstract

How do term limits affect the ideological composition of state legislatures? While existing work documents increased polarization in term-limited incumbents' voting records, little is known about how term limits affect the candidate pipeline, electoral selection, and incumbents' ideology over time. Pairing a first-of-its-kind dataset of state legislative election returns for 1992-2020 with novel roll-call-based candidate ideology scalings introduced in Handan-Nader, Myers, and Hall (2021) and a difference-indifferences design, I implement the first comprehensive study of the ideological effects of term limits in state legislatures. I find that term limits generate increased polarization among candidates at all stages of the candidate pipeline, from the pool of primary and general election candidates to eventual race winners. Contrary to pundits' expectations, I show that this effect is not mediated by asymmetric polarization. Term limits also appear to systematically shift the electorate's preferences, resulting in a decline in the electoral return to moderation in general election races. Finally, I present evidence that term limits do not significantly induce incumbents to shift their ideological positions. In sum, term-limited legislatures simultaneously attract more extreme candidates and reward extremity at a higher rate at the ballot box. These findings have important implications for models of electoral accountability and incentives.

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#### 1 Introduction

Over the past three decades, twenty-one states have enacted legislative term limits, often with the intent of increasing legislator responsiveness and reducing the influence of special interest groups.<sup>1</sup> Contemporary gridlock and ideological extremity, however, have generated new interest in legislative term limits as a mechanism to reduce partisan ideological polarization. A diverse group of politicians, from presidential hopefuls to former members of congress, have endorsed legislative term limits as a means to reduce ideological extremity.<sup>2</sup> Even among the American public, interest in term limits is high. Facing increasing partisan gridlock and polarization, a 2013 Gallup poll reported that 75% of American adults supported term limits for Congress (Gallup, 2013).

Recent research, however, casts doubt on the ability of term limits to moderate ideological extremity. In an analysis of states for 1993-2016, Olson and Rogowski (2020) illustrate that state legislative term limits have instead increased partisan polarization among sitting legislators. While informative as a diagnostic tool, this finding leaves important unanswered questions related to the effect of term limits on the aggregate candidate pipeline, electoral selection, and incumbents' ideology over time.<sup>3</sup>

In this paper, I investigate how term limits affect the ideological composition of state legislatures. Specifically, I aim to trace the effect of term limits on the supply of state legislative candidates, the role voters play in shaping the candidate pool and selecting winners, and the evolution of term-limited incumbents' ideology over time.

My research follows in a rich literature on the effects of legislative incentives on incumbent behavior and candidate ideology. Previous work underscores how factors such as legislator pay (Hall, 2019), electoral competitiveness (Fiorina, 1993; Ansolabehere, Brady, and Fiorina, 1992; Griffin, 2006), primary election format (Norrander, 1989; Gerber and Morton, 1998; Kaufmann, Gimpel, and Hoffman, 2003; McGhee et al., 2014), and primary challengers (Ansolabehere, Snyder, and Stewart, 2001; Brady, Han, and Pope, 2007) shape candidate behavior.

As an important determinant of legislative incentives, previous research evaluates the extent to which legislative term limits affect a variety of behavioral and policy outcomes. Outcomes of interest include legislative productivity (Fouirnaies and Hall, 2021), fiscal policy (Johnson and Crain, 2004; Erler, 2007), women and minority groups' representation (Casellas, 2010; Carroll, 2005; Robert, 1996), bills' policy complexity (Kousser, 2006), voter turnout (Nalder, 2007), and the electoral advantage of incumbents (Rogers, 2014).

Existing scholarship also probes the effect of term limits on incumbents' ideology and measures of states' polarization. Olson and Rogowski (2020) provide the most-comprehensive evidence on the effect of legislative term limits on partisan polarization. Analyzing NP-Scores for incumbents, Olson and Rogowski (2020) find that term limits increase the ideological

<sup>&</sup>lt;sup>1</sup>Legislative term limits imposed by voters in six states were nullified by court or legislative action. Fifteen states currently have legislative term limits in effect.

<sup>&</sup>lt;sup>2</sup>See, for example, comments by former Utah governor Jon Huntsman and former U.S. Senator Joseph Lieberman reported in Olson and Rogowski (2020).

<sup>&</sup>lt;sup>3</sup>Henceforth, I will use the term "candidate pipeline" to refer to the aggregate population of candidates who run for legislative office. This term includes primary and general election candidates in addition to eventual election winners.

gap between Democrats' and Republicans' voting records. Olson and Rogowski argue that term limits increase the role of parties in the legislative process and alter legislators' career incentives. Although informative for incumbents, this work does not address the effect of term limits on the broader pool of candidates who run for office.

Other related studies report null results. Wright (2007), comparing nation-wide legislative roll-call voting for the 1999-2000 session, finds no evidence that term limits increase levels of partisan polarization among state legislatures. Similarly, in the context of the Arkansas state senate (Titiunik and Feher, 2018) and California legislature (Cain and Kousser, 2004), scholars find no significant effect of the introduction of term limits on candidates' ideological positions. Nevertheless, this work is hampered by its cross-sectional design and focus on individual states, precluding extrapolation to other settings.

Finally, my research also complements scholarship on the electoral returns to moderation and electoral accountability (Handan-Nader, Myers, and Hall, 2021; Caughey and Warshaw, 2019; Hall, 2019, 2015). Given the growing scholarly and public concern with polarization in the American political system, it is important to fully understand how term limits contribute to this effect.

I begin by showing how term limits increase partisan polarization within the pipeline of legislative candidates. As Hall (2019) demonstrates in the context of U.S. House elections, the ideological composition of office-seekers shapes overall legislative polarization. If the legislative candidate pipeline becomes more polarized, legislative polarization will increase in tandem.

Subsequently, I demonstrate that, contrary to scholarly expectations, state legislative polarization is not driven by asymmetric polarization among Republican candidates. My analysis shows that Democratic and Republican candidates contribute equally to the polarization of state legislatures. While surprising, this finding parallels recent work in Handan-Nader, Myers, and Hall (2021).

Candidates are not the only source of increased polarization among term-limited states. I illustrate that voters in term-limited states punish extreme candidates at half the rate of their non-term-limited counterparts in general elections. Thus, extremist candidates face less-significant electoral penalties at the ballot box and prospective extremists receive a strong indicator of their potential success in term-limited states.

Finally, I find that legislative term limits do not significantly induce incumbents to alter their ideological positions, complementing similar work by Fournaies and Hall (2021). Overall, I conclude that the polarization of term-limited states can be traced to a changing candidate pool as well as the preferences of voters.

The remainder of the paper is organized as follows. Section 2 underscores the importance of studying the relationship between term limits and candidate ideology. In Section 3, I outline my solution to the methodological challenges of studying candidate pool ideology and introduce a new dataset on state legislative election results. Section 4 documents the polarizing effects of term limits across all stages of legislative elections. Section 5 illustrates how the electorates' ideological preferences change under term limits. In Section 6, I investigate whether term limits cause incumbents to change their ideological positions. Section 7 concludes.

#### 2 Term Limits and Polarization

Before presenting my data and analysis, I outline why the relationship between term limits and legislative polarization is an important object of study.

State legislative term limits and their effects on polarization are important to study for at least three reasons. First, state legislatures are increasingly consequential policy-making bodies. Many of today's most controversial political issues—including abortion rights, voting access, education policy, and election certification—originate and are decided in statehouses. If term limits alter the ideological composition of state legislatures, they will also impact a host of essential policy outcomes.

Second, state legislatures are a key source of members of Congress. By one count, nearly half of the members of the 116th Congress were former state legislators.<sup>4</sup> Thus, policies that affect the composition of state legislatures are certain to shape policy-making and polarization at the federal level (Hall, 2019; Thomsen, 2014).

Finally, the study of state legislative term limits offers insights into age-old questions about electoral accountability and legislative incentives. The direct link between a legislator's actions and her hopes for re-election form the foundation of models of electoral accountability (e.g. Barro 1973; Fearon 1999). To the extent that they alter legislators' incentives, term limits comprehensively shape the democratic system. Hence, careful study of legislative term limits has implications for key state-level policy outcomes, the composition of Congress, and models of democratic representation.

#### 3 Data and Methods

To implement my study, I combine data on state legislative primary and general election returns with novel roll-call-based candidate ideology scores. Overall, these data cover 49 states for the years 1992-2020, ensuring comprehensive coverage of term-limited as well as non-term-limited state legislative candidates. In accordance with existing work, I exclude non-partisan Nebraska from the analysis and focus on Democratic and Republican candidates. Table 1 summarizes the relevant characteristics of term-limited states included in my analysis.

General election data were extracted from the State Legislative Election Returns dataset (SLERs) (Klarner, 2021) and includes full coverage of this study's window of analysis. The majority of the primary election returns data was aggregated in Handan-Nader, Myers, and Hall (2021) with supplementary data collected by the author for this study.<sup>5</sup> After merging primary and general election returns, the combined dataset features 75, 479 distinct general election candidates and 42,068 distinct primary election candidates across 146,855 races with a total of 208,589 candidate-year observations. See Appendix Table A.1 for a state-by-year matrix of my data's coverage.

The ideal measure of ideology for this analysis captures how legislators would cast roll-call

<sup>&</sup>lt;sup>4</sup>https://www.ncsl.org/blog/2018/11/02/how-many-former-state-legislators-serve-in-congress.aspx.

<sup>&</sup>lt;sup>5</sup>Although every effort was made to construct a complete primary returns dataset, returns for a small number of primary races were not available online. Overall, my primary dataset covers approximately 86% of all state-year-chambers.

Table 1: Summary of Term-Limited States in Analysis

State	Year Enacted	Type	Term Limit House	Term Limit Senate							
AR	1992	$\begin{cases} \text{Lifetime} & t < 2020\\ \text{Consecutive} & t \ge 2020 \end{cases}$	$\begin{cases} 6 & t < 2014 \\ 16 & t \in [2014, 2020) \end{cases}$	$\begin{cases} 8 & t < 2014 \\ 16 & t \in [2014, 2020) \\ 12 & t \ge 2020 \end{cases}$							
AZ	1992	Consecutive	$\begin{cases} 12 & t \ge 2020 \\ 8 & \end{cases}$	$\begin{cases} 12 & t \ge 2020 \\ 8 & \end{cases}$							
CA	1990	Lifetime	$\begin{cases} 6 & t < 2012 \\ 12 & t \ge 2012 \end{cases}$	$\begin{cases} 8 & t < 2012 \\ 12 & t \ge 2012 \end{cases}$							
CO	1990	Consecutive	8	8							
$\operatorname{FL}$	1992	Consecutive	8	8							
LA	1995	Consecutive	12	12							
ME	1993	Consecutive	8	8							
MI	1992	Lifetime	6	8							
MO	1992	Lifetime	8	8							
MT	1992	Consecutive	8	8							
NV	1996	Lifetime	12	12							
ОН	1992	Consecutive	8	8							
OK	1990	Lifetime	12	12							
SD	1992	Consecutive	8	8							

votes in office. Unfortunately, a roll-call-based ideology scaling, such as Shor and McCarthy's (2011) NP-Scores, is only available for the subset of state legislative candidates who become sitting legislators. In response, Handan-Nader, Myers, and Hall (2021) apply a supervised machine learning scaling procedure to predict roll-call-behavior, as measured by NP-Scores, using candidate campaign contribution records. The resulting scalings—referred to as candidates' "estimated ideology"—correlate highly with NP-Scores (r = 0.97), but are available for election losers in addition to winners and are dynamic over time. This scaling is the main measure of candidate ideology employed throughout my analysis.

Finally, after merging ideology scores to the election returns dataset, I construct indicators for candidate-level and chamber-level term limits using data from the National Conference of State Legislatures (NCSL).

## 4 Term Limits Generate A More-Polarized Candidate Pool

In this section, I explore the effect of legislative term limits on candidate-pool partisan polarization. I aim to evaluate whether term limits increase the overall level of ideological extremity among legislative candidates. In closely related work, Olson and Rogowski (2020) analyze the impact of legislative term limits on the state-level partisan polarization of incumbents. They find that legislative term limits are associated with increased partisan po-

larization among incumbents. I complement this analysis by studying partian polarization at all stages of the election process.

The analysis in this section proceeds in two stages. First, I consider the effect of term limits on state-level partisan polarization. Studying state-level trends provides an understanding of the overall ideological effects of term limits. In the second subsection, I decompose the effects of term limits by party.

#### 4.1 State-Level Analysis

Consistent with Olson and Rogowski (2020), I employ a state-level difference-in-differences design for the years 1992-2020. Specifically, I model

$$Y_{st} = \beta_0 + \beta_1 Term \ Limited_{st} + \Omega X_{st} + \alpha_s + \delta_t + \epsilon_{st}$$
 (1)

where  $Y_{st}$  is the level of partisan polarization in state s in year t,  $Term\ Limited_{st}$  indicates whether state s in time t had term limits in effect,  $X_{st}$  is a vector of controls, and  $\alpha_s$  and  $\delta_t$  are state and year fixed effects, respectively. The error term,  $\epsilon_{st}$ , is clustered at the state level. This specification allows me to make comparisons within state-year units.

I define partisan polarization,  $Y_{st}$ , as the difference between the median Republican and Democratic candidates' ideology scores in state s in year t. The term limits variable,  $Term\ Limited_{st}$ , indicates state-years for which tenured incumbents are no longer eligible to run for re-election.<sup>6</sup> When years of impact differ between a state's house and senate, I code treatment as beginning on the first year of impact.<sup>7</sup> Both  $Y_{st}$  and  $Term\ Limited_{st}$  mirror definitions in Olson and Rogowski (2020).

My difference-in-differences research design requires a parallel trends assumption. Functionally, this assumption dictates that control states (i.e. states that never enacted legislative term limits) would have responded to the implementation of term limits in the same manner as treatment states (i.e. states that eventually implemented term limits). Since legislative term limits were most often implemented by means of citizen-led referendums, and I control for dynamic state-specific political attributes, the following results may be meaningfully interpreted as the ideological effect of term limits.

Table 2 shows the results for my state-level analysis. For all specifications I present a univariate model and, to guard against the possibility of attributing non-static state features to the effect of term limits, a model with state governance controls. The battery of controls was first introduced in Olson and Rogowski (2020). Legislative Professionalism (Squire, 2017)

Term Limited<sub>st</sub> = 
$$\begin{cases} 0, & t < 2000 \\ 1, & t \ge 2000 \end{cases}.$$

 $<sup>^6</sup>$ An alternate definition might operationalize  $Term\ Limited_{st}$  based on term limits' enactment date. Unfortunately, campaign finance data limitations preclude this possibility. Existing research by Keele, Malhotra, and McCubbins (2013), however, find similar results using enactment and implementation dates. Further, as Olson and Rogowski (2020) note, defining  $Term\ Limited_{st}$  based on implementation date will likely serve to attenuate coefficient estimates, making the ensuing analysis particularly rigorous.

<sup>&</sup>lt;sup>7</sup>For example,  $Term\ Limited_{st}$  for Michigan, which implemented 6 year house and 8 year senate term limits beginning in 1992 and has biannual house elections, is defined

<u>Table 2: Term Limits and Partisan Polarization</u>

	Н	andan-Na I	NP-Score					
	Candida	ate Pool		Election lidates	Ge	eneral El	ection Wi	nners
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Term Limited	0.105*	0.107*	0.103*	0.105*	0.079	0.082	0.197**	0.202***
	(0.06)	(0.05)	(0.06)	(0.06)	(0.06)	(0.06)	(0.08)	(0.07)
Log(Leg Prof)		0.025		0.014		0.048		0.075
		(0.05)		(0.05)		(0.05)		(0.07)
Divided Government		-0.005		-0.010		-0.007		-0.006
		(0.01)		(0.01)		(0.01)		(0.02)
Party Competetiveness		0.000		0.000		0.000		-0.000
		(0.00)		(0.00)		(0.00)		(0.00)
N	591	591	591	591	588	588	485	485
Standard Deviation	.39	.39	.39	.39	.4	.4	.47	.47
Year FEs	Y	Y	Y	Y	Y	Y	Y	Y
State FEs	Y	Y	Y	Y	Y	Y	Y	Y

<sup>\*</sup> p<.1, \*\* p<.05, \*\*\* p<.01

Note: Standard errors in parentheses are clustered by state. The outcome—either "esimated ideology" or NP-Score—is listed in the table header.

combines information on legislator salary, session length, and staffing resources to quantify legislator engagement in policy making. *Divided Government* indicates whether one party simultaneously controls the governorship, house, and senate. Finally, *Party Competitiveness* measures the absolute two-party difference in control of legislative seats.

Columns 1 and 2 in Table 2 estimate the effect of legislative term limits on partisan polarization across the candidate pipeline, including primary winners and losers and general election candidates. The coefficients on *Term Limited* in columns 1 and 2 are positive and statistically significant, indicating that partisan polarization is greater, on average, among candidate pools with legislative term limits. These results are robust to the inclusion of control variables in column 2, a pattern that holds across all specifications in Table 2.

Next, I restrict the sample to general election candidates in columns 3 and 4. I find continued evidence that term limits increase partisan polarization, but in this case among candidates who reach the general election. Finally, columns 5 and 6 analyze only general election winners or, equivalently, candidates who become legislators. While the coefficients on *Term Limited* among general election winners are not significant at traditional levels, the point estimates are strikingly similar to those of columns 1-4. Note that the effects outlined in Table 2—approximately one quarter of one standard deviation—are substantively meaningful in addition to statistically significant.

As an external validity check, I replicate the incumbent analysis using NP-Scores in columns 7 and 8. The resulting coefficients mirror those of Olson and Rogowski (2020), lending external validity to the estimated ideology score findings.<sup>8</sup>

<sup>&</sup>lt;sup>8</sup>That the NP-Score coefficients are larger in magnitude than the estimated ideology scores of Handan-Nader, Myers, and Hall (2021) suggests that the results in columns 1-6 constitute a lower bound estimate of

The results in Table 2 establish an important finding: state legislative term limits produce meaningfully higher levels of partisan polarization among all office seekers, rather than only among office holders. Hence, legislative term limits not only cause incumbents to take more-extreme positions in the legislature, but also increase the overall level of extremity of legislative candidates.

#### 4.2 Party-Level Analysis

In this section, I evaluate whether the ideological effects of legislative term limits vary by party.

A growing body of scholarship explores the prevalence of asymmetric polarization in American elections. In the standard account, scholars argue that ideological polarization is disproportionately driven by rising Republican extremity (e.g Grossman and Hopkins 2016; McCarty, Poole, and Rosenthal 2007). Recent evidence at the state level, reaches different conclusions. Olson and Rogowski (2020) find no evidence of asymmetric polarization among incumbent legislators in term-limited states. Handan-Nader, Myers, and Hall (2021) also find little evidence of asymmetric polarization among the pool of general election candidates, but illustrate that Democratic primaries favor extremists at a higher rate than Republican primaries.

Table 3: Asymmetric Polarization in the Candidate Pipeline

		Democrats	}	F	Republicar	ns
	(1)	(2)	(3)	$\overline{(4)}$	(5)	(6)
Term Limited	-0.066**	-0.061**	-0.067**	0.034	0.036	0.031
	(0.03)	(0.03)	(0.03)	(0.03)	(0.03)	(0.03)
Log(Leg Prof)	0.001	-0.022	0.028	0.019	0.027	0.012
	(0.03)	(0.04)	(0.03)	(0.02)	(0.02)	(0.02)
Divided Government	-0.004	-0.001	-0.006	-0.005	-0.011*	0.001
	(0.01)	(0.01)	(0.01)	(0.01)	(0.01)	(0.01)
Party Competetiveness	-0.000	-0.000	-0.000	-0.000	-0.000	0.000
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)
N	1,128	579	546	1,136	583	548
Standard Deviation	.85	.85	.85	.85	.85	.85
Specification	Pooled	House	Senate	Pooled	House	Senate
Year FEs	Y	Y	Y	Y	Y	Y
State FEs	Y	Y	Y	Y	Y	Y

<sup>\*</sup> p<.1, \*\* p<.05, \*\*\* p<.01

Note: Standard errors in parentheses are clustered by state. The outcome variable is Handan-Nader, Myers, and Hall's (2021) estimated ideology score for the complete candidate pool.

In Table 3, I re-estimate Equation 1 on the candidate pipeline after defining  $Y_{st}$  separately for Democrats and Republicans. Pooling across chambers, columns 1 and 4 demonstrate that

the true effect.

term limits are associated with a shift to the left among Democratic office-seeks and a shift to the right among Republican office-seekers compared to non-term-limited candidates. Only the effect for Democratic candidates is statistically distinguishable from zero, a surprising finding in light of national-level asymmetric polarization research. These results hold when I restrict my sample to state house candidates (columns 2 and 5) and state senate candidates (columns 3 and 6).

The coefficient estimates in Table 3 suggest that term limits' effects are approximately 40% larger among Democrats than Republicans. In Appendix Table A.3, I illustrate that this asymmetric polarization does not approach statistical significance.

The findings presented in this section suggest a more-nuanced picture of the ideological effects of legislative term limits, including an important role of electoral selection to which I now turn.

### 5 The Electorate's Changing Ideological Preferences

The origins of the effects observed in Section 3 can be divided into two constituent parts. First, term limits may directly influence the ideological positions of legislative candidates and incumbents. From this perspective, termed-out legislators may systematically shift their ideology once electoral incentives are removed, or term limits may foster ideological selection into the candidate pool. Second, term limits may alter the preferences and behavior of the electorate. In the first setting, candidates and legislators would drive the increased polarization observed in Section 3, while voters would play the same role in the second. In this section I consider the latter explanation, leaving analysis of the former to Section 5.

To assess the role term-limited states' electorates play in producing increased partisan polarization, I employ two modeling strategies. Because general election races feature direct two-party competition, it is possible to compare the ideology of competing Democratic and Republican general election candidates and predict their electoral returns to changes in ideological platform. To do so, I adopt the midpoint method of Ansolabehere, Snyder, and Stewart (2001). Consequently, I estimate an equation of the form

$$Y_{dct} = \beta_0 + \beta_1 Midpoint_{dct} + \beta_2 Distance_{dct} + \Omega X_{dct} + \alpha_d + \delta_t + \epsilon_{dct}$$
 (2)

where  $Y_{dct}$  is either the Democratic candidate's general election vote share or a victory indicator in district d in chamber c in year t.  $^9$  Midpoint and Distance are the midpoint and distance between Democratic and Republican candidates, respectively. Finally,  $X_{dct}$  is an optional vector of controls,  $\alpha_d$  and  $\delta_t$  are district and year fixed effects, respectively, and the error term,  $\epsilon_{dct}$ , is clustered by district d.  $^{10}$ 

The coefficient of interest is  $\beta_1$ , the estimated electoral return for the Democratic candidate arising from a rightward (i.e. positive) shift in *Midpoint*. Previous research on Congress

<sup>&</sup>lt;sup>9</sup>Since this design requires competition between one Democratic and one Republican candidate, I restrict my sample to single-member districts when using the midpoint model.

<sup>&</sup>lt;sup>10</sup>The midpoint model requires the ideology of districts' mean voter to be held constant. Ansolabehere, Snyder, and Stewart (2001) use presidential vote share for this purpose. Because presidential vote share is not available at the level of state legislative districts, I employ district fixed effects to hold the median voter constant.

(Hall, 2019; Ansolabehere, Snyder, and Stewart, 2001) as well as state legislatives (Handan-Nader, Myers, and Hall, 2021; Caughey and Warshaw, 2019) suggest that  $\beta_1$  is positive and significant. After replicating existing findings, I test whether  $\beta_1$  differs significantly between term-limited and non-term-limited states.

Since the midpoint model cannot be applied to races with multiple candidates from the same party, I apply a candidate ideological extremism model to study primary election returns. Specifically, I estimate the model

$$Y_{jpdt} = \beta_0 + \beta_1 Extremism_{jpdt} + \alpha_{pd} + \delta_{pt} + \eta_{pdt} + \epsilon_{jpdt}$$
(3)

where  $Y_{jpdt}$  is the vote share or a victory indicator for candidate j in party p in district d in year t. Extremism is the absolute value of a candidate's estimated ideology score. The variables  $\alpha_{pd}$ ,  $\delta_{pt}$ , and  $\eta_{pdt}$  are party-by-district, party-by-year, and number of primary candidates fixed effects, respectively.

The coefficient  $\beta_1$  captures the electoral return to becoming more extreme. Previous work finds a positive coefficient (Handan-Nader, Myers, and Hall, 2021; Ansolabehere, Snyder, and Stewart, 2001; Brady, Han, and Pope, 2007), indicating that primary candidates receive an electoral boost from ideological extremism.

Table 4: Term Limits and Electoral Selection

		te Share		y Votes	Dem	Win	Prima	ry Win
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Midpoint	0.047***	0.058***			0.300***	0.327***		
	(0.01)	(0.01)			(0.06)	(0.07)		
Distance	0.017*	0.017*			0.135***	0.135***		
	(0.01)	(0.01)			(0.05)	(0.05)		
Dem Contributions	0.441***	0.440***			1.074***	1.070***		
	(0.01)	(0.01)			(0.05)	(0.05)		
Rep Contributions	-0.408***	-0.407***			-1.103***	-1.102***		
	(0.01)	(0.01)			(0.06)	(0.05)		
Term Limits		0.008		0.013		0.007		-0.033
		(0.01)		(0.02)		(0.06)		(0.04)
Term Limits $\cdot$ Midpoint		-0.033*				-0.082		
		(0.02)				(0.10)		
Extremism			0.020***	0.022***			0.052***	0.043***
			(0.01)	(0.01)			(0.01)	(0.02)
Contributions			0.083***	0.083***			0.213***	0.213***
			(0.00)	(0.00)			(0.00)	(0.00)
Term Limits $\cdot$ Extremism				-0.008				0.035
				(0.01)				(0.03)
N	21,702	21,702	42,595	$42,\!595$	21,702	21,702	49,134	49,134
District FEs	Y	Y	N	N	Y	Y	N	N
Year FEs	Y	Y	N	N	Y	Y	N	N
District-by-Party FEs	N	N	Y	Y	N	N	Y	Y
Party-by-Year FEs	N	N	Y	Y	N	N	Y	Y

<sup>\*</sup> p<.1, \*\* p<.05, \*\*\* p<.01

Note: Robust standard errors in parentheses are clustered by district. The outcome is Handan-Nader, Myers, and Hall's (2021) estimated ideology score.

Results are reported in Table 4. The baseline general election models, listed in columns 1 and 5, provide compelling evidence that general election candidates are punished by voters

for ideological extremity. These estimates closely approximate estimates in Handan-Nader, Myers, and Hall (2021), providing credibility to my subsequent extension.

I am interested in the difference in *Midpoint* coefficients between term-limited and non-term-limited state-years. To estimate this difference, columns 2 and 6 interact *Midpoint* with *Term Limits*, an indicator for the presence of legislative term limits defined in Section 3. If voters in term-limited states reward ideological extremity at a higher rate than their peers in non-term-limited states, the interaction term would be negative. Conversely, if, relative to non-term-limited-states, voters in term-limited states punish candidates for ideological extremism more, the interaction term would be positive.

It turns out that the interaction terms in both models are negative, indicating that extreme ideological positions are penalized at a lower rate in term-limited states. These effects are decidedly large. In fact, the vote-share penalty for ideological extremism is more than 50% lower in term-limited states. Similarly, the decline in term-limited candidates' win probability as a function of ideological extremism is 30% lower in term-limited states.

Next, I consider primary elections. Columns 3 and 7 explore the overall relationship between extremism and primary election outcomes. Contrary to general elections, I find that, on average, primary election candidates receive an electoral advantage from taking more extreme ideological positions. The coefficient estimates are on par with existing research (Handan-Nader, Myers, and Hall, 2021). I find no evidence that legislative term limits affect voters' primary election behavior. The interaction terms in columns 4 and 7 are small and not statistically significant.

In Appendix Table A.2, I re-estimate equations 2 and 3 using only open-seat races. The coefficient on *Midpoint* vote share is almost three times as large for only open-seat races compared to all races. Although not a dispositive test, this finding suggests that polarization in term-limited legislatures can be partially attributed to cyclic replacement of more-extreme legislators.

Taken as a whole, I find compelling evidence that voters contribute to the elevated ideological polarization in term-limited states. This relationship, however, is limited to general elections. It now remains to investigate the potential for term-limited legislators' actions in office to shape partisan polarization.

### 6 Termed-Out Legislators' Ideology is Constant

In the final portion of this paper, I evaluate whether incumbents contribute to the heightened partisan polarization observed in states with legislative term limits. To do so, I consider whether termed-out legislators systematically shift their ideological positions.

Legislative term limits comprehensively alter candidates' electoral incentives. Candidates who are termed-out need not consider the electoral implications of their policy actions, at least as long as they don't plan to run for higher office. As a result, some observers suggest that term-limited candidates may pander to their partisan bases, leading to increased ideological extremity (Canes-Wrone, Herron, and Shotts, 2001). Conversely, if candidates become more responsive to their constituents in their final term—a perspective supported by the "ideological shirking" and "marginality hypothesis" literatures—we might expect term

limits to reduce term-limited candidates' ideological extremism. <sup>11</sup>

Ultimately, this is an empirical question, but one that has received little direct scholarly attention. Related research on the effect of district electoral incentives on legislator responsiveness reaches mixed conclusions (e.g. Fiorina 1993; Ansolabehere, Brady, and Fiorina 1992; Burden 2004; Griffin 2006). Further, most research focuses on Congressional elections rather than state legislative elections.

The notable exception is Fouirnaies and Hall (2021) which analyzes how legislative term limits influence legislators' behavior in office, including ideological positioning. Using W-NOMINATE scores derived from roll-call records as well as interest group ratings, Fouirnaies and Hall (2021) find no evidence that termed-out legislators systematically shift their ideological platforms. Instead, they argue that termed-out legislators allocate less time to legislative activities. Due to data and modeling limitations, this analysis applies to 16 of the 29 term-limited state-chambers and uses only 2 non-term-limited control states.

Since Handan-Nader, Myers, and Hall's (2021) estimated ideology scores are dynamic over time and feature near-complete coverage, I extend the analysis framework of Fouirnaies and Hall to all term-limited state-chambers (with the exception of non-partisan Nebraska) and use all possible control states. I implement a difference-in-differences design within individuals and states, allowing me to control for legislator time effects.<sup>12</sup>

Specifically, I estimate the equation

$$Y_{jct} = \beta_0 + \beta_1 Term \ Limited_{jct} + \alpha_j + \Omega + \epsilon_{jct}$$
 (4)

where  $Y_{jct}$  is the absolute value of legislator j's estimated ideology score in time t in chamber c,  $Term\ Limited_{jct}$  indicates whether a legislator is serving in their final term before term limits apply,  $\alpha_i$  are legislator fixed effects, and  $\Omega$  stands in for either chamber-by-year or chamber-by-party-by-year fixed effects. The error term,  $\epsilon_{jct}$ , is clustered at the legislator level.

Table 5 reports the results. In the first column, which includes legislator and chamber-by-year fixed effects, the effect of being term-limited is calculated within groups of legislators serving in the same chamber and year. Thus, the institutional factors that define term limits, as well as other unobserved static features, are held constant. The estimate on *Term Limits* in column 1 is negative and significant, suggesting that candidates moderate their ideological platforms in their final term. However, the coefficient—representing roughly 1.5% of one standard deviation—is quite small, especially in comparison to the results in Tables 2 and 4.

To ensure the results are not confounded by inter-party differences, the second column substitutes in chamber-by-party-by-year fixed effects. In this model, the counterfactual for a term-limited candidate is a candidate in the same chamber and party and year. Again, while the coefficient estimate in column 2 is statistically significant, it is not meaningful.

Overall, I find no evidence that termed-out legislators meaningfully shift their ideological

<sup>&</sup>lt;sup>11</sup>For a review of work related to ideological shirking, see Bender and Lott (1996). See Griffin (2006) for an overview of electoral competition and marginality-hypothesis-related research.

<sup>&</sup>lt;sup>12</sup>Since estimated ideology scores are constructed with reference to candidates' campaign contributions, and final-term candidates likely accrue fewer donations, these results should be considered in conjunction with those of Fouirnaies and Hall (2021).

Table 5: Incumbent Ideological Positioning

	Incumben	t Ideology
	(1)	(2)
Term Limited	-0.006**	-0.006**
	(0.00)	(0.00)
N	64,159	64,115
Standard Deviation	.31	.31
Legislator FEs	Y	Y
Chamber-by-Year FEs	Y	N
Chamber-by-Party-by-Year FEs	N	Y

\* p<.1, \*\* p<.05, \*\*\* p<.01

Note: Robust standard errors in parentheses are clustered by legislator. The outcome is the absolute value of Handan-Nader, Myers, and Hall's (2021) estimated ideology score.

platforms. This finding is consistent with empirical results reported in Fourinaies and Hall (2021) and the theoretical expectations of the marginality hypothesis literature.

#### 7 Discussion and Conclusion

Political leaders, scholars, and the general public alike are increasingly concerned about partisan polarization and the accompanying legislative gridlock. Proponents argue that term limits will reduce legislative polarization by tempering the incumbent advantage, reducing the influence of lobbyists and special interests, and redirecting legislators' priorities from reelection to policy. Recent research suggests that term limits fail to achieve this objective, at least among incumbents.

Using ideological scalings and election returns introduced in Handan-Nader, Myers, and Hall (2021), I conduct the first comprehensive analysis of the effects of legislative term limits on legislative ideology. I establish three empirical findings.

First, I demonstrate that legislative term limits produce increased polarization at all stages of the political process. The average state legislative politician is significantly more extreme in legislatures with term limits compared to non-term-limited legislatures. This pattern holds for the aggregate pool of candidates, general election candidates, and eventual office holders. Contrary to theories of asymmetric polarization, I find no evidence that either political party disproportionately contributed to term-limit-driven polarization at the state level.

Second, I illustrate that term limits systematically shift voters' preferences. Voters in term-limited states punish ideological extremity at less than half the rate of non-term-limited states in general elections. I find no evidence that term limits shape voters' primary election preferences.

Finally, in accord with recent work by Fouirnaies and Hall (2021), I show that legislative term limits do not meaningfully impact incumbents' ideology. Taken together, my analysis suggests that term-limit-included polarization is driven by voters and candidates, but not meaningfully by incumbents.

From a broader perspective, my paper contributes to an expansive literature on electoral incentives. Over the last thirty years, stronger electoral incentives—in the form of legislative term limits—have led to increased polarization. Future research should investigate the causal relationships that mediate the effects outlined in this paper. Research of this kind might explore how term limits affect voter preferences and candidate selection. As noted in Section 2, state legislatures shape important domestic policies. Future research that explores how term limits affect these key policy-making bodies will be well-rewarded.

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A.1 Data Descriptives

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Table A.1: Data Coverage Matrix

## A.2 Open Seat Electoral Selection

Table A.2: Term Limits and Electoral Selection - Only Open Seat Races

	Dem Vo	te Share	Primar	y Votes	Dem	Win	Prima	ry Win
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Midpoint	0.028	0.087**			0.659***	0.701**		
	(0.03)	(0.04)			(0.22)	(0.29)		
Distance	0.000	0.004			0.077	0.093		
	(0.03)	(0.03)			(0.17)	(0.17)		
Dem Contributions	0.448***	0.439***			1.358***	1.336***		
	(0.03)	(0.03)			(0.18)	(0.18)		
Rep Contributions	-0.401***	-0.395***			-1.533***	-1.518***		
	(0.03)	(0.03)			(0.22)	(0.22)		
Term Limits		0.035		-0.016		-0.039		-0.112
		(0.03)		(0.02)		(0.20)		(0.07)
Term Limits $\cdot$ Midpoint		-0.098**				-0.069		
		(0.05)				(0.30)		
Extremism			0.047***	0.049***			0.133***	0.107***
			(0.01)	(0.01)			(0.03)	(0.03)
Contributions			0.089***	0.089***			0.240***	0.240***
			(0.00)	(0.00)			(0.00)	(0.00)
Term Limits $\cdot$ Extremism				-0.005				0.067
				(0.02)				(0.05)
N	2,102	2,102	18,893	18,893	2,102	2,102	19,337	19,337
District FEs	Y	Y	N	N	Y	Y	N	N
Year FEs	Y	Y	N	N	Y	Y	N	N
District-by-Party FEs	N	N	Y	Y	N	N	Y	Y
Party-by-Year FEs	N	N	Y	Y	N	N	Y	Y

<sup>\*</sup> p<.1, \*\* p<.05, \*\*\* p<.01

Note: Robust standard errors in parentheses are clustered by district. The outcome is Handan-Nader, Myers, and Hall's (2021) estimated ideology score. Sample is restricted to open seat races.

#### Testing Asymmetric Polarization A.3

Table A.3: Asymmetric Polarization in the Candidate Pipeline

	(1)	(2)	(3)
Term Limited	0.116	0.117	0.114
	(0.07)	(0.07)	(0.07)
Dem	0.094	0.094	0.095
	(0.07)	(0.07)	(0.07)
Term Limited $\cdot$ Dem	-0.135	-0.138	-0.132
	(0.14)	(0.14)	(0.14)
Log(Leg Prof)	0.003	0.022	-0.018
	(0.02)	(0.02)	(0.02)
Divided Government	0.000	-0.004	0.005
	(0.00)	(0.00)	(0.01)
Party Competetiveness	0.000	-0.000	0.000
	(0.00)	(0.00)	(0.00)
N	2,265	1,164	1,100
Standard Deviation	.3	.3	.3
Specification	Pooled	House	Senate
Year FEs	Y	Y	Y
State FEs	Y	Y	Y

value of Handan-Nader, Myers, and Hall's (2021) estimated ideology score for the complete candidate pool.