## Serie 12

### Aufgabe 12.1

We investigate further the data set Boston from the last exercise sheet.

In order to fit a multiple linear regression model using least squares, we again use the ols() function. The syntax ols(" $y \sim x1 + x2 + x3$ ", data=...) is used to fit a model with three predictors, x1, x2, and x3. The summary() function now outputs the regression coefficients for all the predictors.

a) Fit a multiple linear regression model with response variable **medv** and predictors **lstat** and **age**.

Define the model and interpret all values in the **summary** () output which we discussed in class (coefficients, its P values,  $R^2$  value, P value of the F-statistics).

b) The Boston data set contains 13 variables, and so it would be cumbersome to have to type all of these in order to perform a regression using all of the predictors. Instead, we can use the following short-hand

```
all_columns = "+".join(df.columns.drop("medv"))
all_columns
fit = ols("medv~" + all_columns , data=df).fit()
```

In the **summary** () output interpret the coefficient of **age** and the corresponding *P* value compare this with the output in a) and explain the difference.

- c) The  $R^2$  value is bigger than the one calculated in a). Explain.
- d) It is easy to include interaction terms in a linear model using the lm() function. The syntax lstat:black tells R to include an interaction term between lstat and black.

The syntax lstat \* age simultaneously includes lstat, age, and the interaction term lstat \* age as predictors; it is a shorthand for lstat + age + lstat:age.

Again, discuss all the values in the **summary ()** of **lstat\*age** as in a).

### Aufgabe 12.2

Wir führen noch eine multiple lineare Regression für Auto aus der letzten Übung durch.

Bevor wir damit beginnen, entfernen wir alle Variablen, "Unnamed: 0", "X1", "name", die qualtitativ oder nicht relevant sind. Dies machen wir mit der .drop()-Methode.

```
import pandas as pd
import statsmodels.api as sm
from statsmodels.graphics.regressionplots import abline_plot
from statsmodels.formula.api import ols
import matplotlib.pyplot as plt
import numpy as np
df = pd.read_csv("../../Themen/Einfache_Lineare_Regression/Jupyter_Notebooks_de/Auto.csv")
df.columns
## Index(['Unnamed: 0', 'X1', 'mpg', 'cylinders', 'displacement', 'horsepower',
## 'weight', 'acceleration', 'year', 'origin', 'name'],
##
        dtype='object')
df = df.drop(["Unnamed: 0", "X1", "name"], axis=1)
df.head()
       mpg cylinders displacement ... acceleration year origin
##
## 0 7.650 8 307.0 ... 12.0 70 1
## 1 6.375 8 350.0 ... 11.5 70 1
## 2 7.650 8 318.0 ... 11.0 70 1
## 3 6.800 8 304.0 ... 12.0 70 1
## 4 7.225 8 302.0 ... 10.5 70 1
##
## [5 rows x 8 columns]
```

a) Produzieren Sie mit .pairplot () Streudiagramme, die alle Variablen des Datensatzes enthält.

```
import seaborn as sns
sns.pairplot(df)
```

Welche Abhängigkeiten stellen Sie fest?

b) Berechnen Sie die Korrelationsmatrix zwischen den Variablen mit df.corr().

Interpretieren Sie die Werte für horsepower und displacement mit den Streudiagrammen oben.

- c) Wir verwenden **lm()** um eine multiple Regression mit der Zielgrösse **mpg** und allen anderen Variablen (ausser **name**) als Prädiktoren durchzuführen. Verwenden Sie wieder Output des **summary()**-Befehls zu interpretieren.
  - i) Gibt es einen Zusammenhang zwischen den Prädiktoren und der Zielvariable? Begründen Sie dies mit dem *p*-Wert zum *F*-Wert.
  - ii) Welche Prädiktoren scheinen statistisch signifikant einen Einfluss auf die Zielvariable zu haben?
  - iii) Was deutet der Koeffizient für year an?
- d) Untersuchen das Modell aus c) noch auf Interaktionseffekte.

# Kurzlösungen vereinzelter Aufgaben

## Musterlösungen zu Serie 12

## Lösung 12.1

a) Model:

$$medv = \beta_0 + \beta_1 \cdot lstat + \beta_2 \cdot age$$

```
import pandas as pd
import statsmodels.api as sm
from statsmodels.formula.api import ols
\textbf{from} \ \texttt{statsmodels.graphics.regressionplots} \ \textbf{import} \ \texttt{abline\_plot}
from scipy.stats import uniform, norm
import matplotlib.pyplot as plt
import numpy as np
df = pd.read_csv("../../Themen/Einfache_Lineare_Regression/Jupyter_Notebooks_de/Boston.csv").drop("Topton de la company de la co
0", axis=1)
fit = ols("medv~lstat+age", data=df).fit()
## <class 'statsmodels.iolib.summary.Summary'>
## """
                                                                                       OLS Regression Results
## Dep. Variable: medv R-squared:
## Model: OLS Adj. R-squared:
## Method: Least Squares F-statistic:
## Date: Mon, 11 May 2020 Prob (F-statistic): 2
## Time: 10:18:15 Log-Likelihood:
## No. Observations: 506 AIC:
## Df Residuals: 503 BIC:
## Df Model: 2
                                                                                                                                                                                                                                 0.551
0.549
309.0
                                                                                                                                                                                                                          2.98e-88
                                                                                                                                                                                                                                -1637.5
                                                                                                                                                                                                                                          3294.
                                                                                                   2
 ## Df Model:
## Covariance Type: nonrobust
                coef std err t P>|t| [0.025 0.975]
##
## -----
## Intercept 33.2228 0.731 45.458 0.000 31.787 34.659
## 1stat -1.0321 0.048 -21.416 0.000 -1.127 -0.937
## age 0.0345 0.012 2.826 0.005 0.011 0.059
## ----
## Omnibus: 124.288 Durbin-Watson: 0.945
## Prob(Omnibus): 0.000 Jarque-Bera (JB): 244.026
## Skew: 1.362 Prob(JB): 1.02e-53
                                                                                                           5.038 Cond. No.
## Kurtosis:
##
## Warnings:
 ## [1] Standard Errors assume that the covariance matrix of the errors is correctly specified.
```

#### The estimates are

$$\hat{\beta}_0 = 33.22;$$
  $\hat{\beta}_1 = -1.03;$   $\hat{\beta}_2 = 0 - 03$ 

We get for the model

$$medv = 33.22 - 1.03 \cdot lstat + 0.03 \cdot age$$

Interpretation of the estimates:

•  $\hat{\beta}_0 = 33.22$ 

In neighborhoods where there is no population of lower status and no units build before 1940, the medium value of houses is \$ 33 220.

•  $\hat{\beta}_1 = -1.03$ 

For each additional percent of population of lower status, the medium value decreases by \$ 1030.

•  $\hat{\beta}_2 = 0.03$ 

For each additional percent of units build before 1949, the medium value increases by \$ 30.

- All *p*-values are significant (below the significance level of 5 %), so all estimates individually contribute significantly to the model.
- The  $R^2$  value is 0.5513, therefore about 55 % of the variation is explained by the model.
- The p-value of the F value is below the significance level and therefore significant. The null hypothesis  $H_0$

$$\beta_1 = \beta_2 = 0$$

is rejected. One of  $\beta$ 's is significantly different from 0. At least one variables contributes significantly to the model.

The p-value is almost 1, so not significant at all. But in a), the p-value is 0.005, which is significant. That means that the variable **age** must correlate strongly with other variables (see d)).

- c) The more variables you have the bigger the  $R^2$  value. That means that the  $R^2$  is not a good indicator to compare different models.
- d) Model:

$$medv = \beta_0 + \beta_1 \cdot lstat + \beta_2 \cdot age + \beta_{12} \cdot lstat*age$$

Remark: \* in **lstat\*age** does *not* signify multiplication, it just means interaction.

The estimates are

$$\widehat{\beta}_0 = 36.10;$$
  $\widehat{\beta}_1 = -1.39;$   $\widehat{\beta}_2 = -0.0007;$   $\widehat{\beta}_{12} = 0.004$ 

We get for the model

$$medv = 36.10 - 1.39 \cdot lstat - 0.00072 \cdot age + 0.0041 \cdot lstat*age$$

Interpretation of the estimates:

• 
$$\hat{\beta}_0 = 36.10$$

In neighborhoods where there is no population of lower status and no units build before 1940, the medium value of houses is \$ 36 100.

• 
$$\hat{\beta}_1 = -1.39$$

For each additional percent of population of lower status, the medium value decreases by \$ 1930.

• 
$$\hat{\beta}_2 = -0.00072$$

For each additional percent of units build before 1949, the medium value decreases by \$ 0.27.

As you can imagine, this value is not significant, as you can see from the output.

• 
$$\hat{\beta}_{12} = 0.004$$

This coefficient is somewhat difficult to interpret and we did't do it in class.

• Not all *p*-values are significant (below the significance level of 5%) anymore.

The p value for **age** ist 0.97, so this not significance anymore, whereas without interaction it was. What is the reason for this?

The p-value of the interaction term is 0.0252 which is below the significance level of 5%. The null hypothesis  $H_0$ , that there is no interaction, is rejected. There is statistically significant interaction.

Now, let's take a look at the correlation coefficient of the two explanatory variables **lstat** and **age**.

```
df.corr()

## crim zn indus ... black lstat medv

## crim 1.000000 -0.200469 0.406583 ... -0.385064 0.455621 -0.388305

## zn -0.200469 1.000000 -0.533828 ... 0.175520 -0.412995 0.360445

## indus 0.406583 -0.533828 1.000000 ... -0.356977 0.603800 -0.483725

## chas -0.055892 -0.042697 0.062938 ... 0.048788 -0.053929 0.175260

## nox 0.420972 -0.516604 0.763651 ... -0.380051 0.590879 -0.427321

## rm -0.219247 0.311991 -0.391676 ... 0.128069 -0.613808 0.695360

## age 0.352734 -0.569537 0.644779 ... -0.273534 0.602339 -0.376955

## dis -0.379670 0.664408 -0.708027 ... 0.291512 -0.496996 0.249929

## rad 0.625505 -0.311948 0.595129 ... -0.444413 0.488676 -0.381626

## tax 0.582764 -0.314563 0.720760 ... -0.44418 0.543993 -0.468536

## ptratio 0.289946 -0.391679 0.383248 ... -0.177383 0.374044 -0.507787

## black -0.385064 0.175520 -0.356977 ... 1.000000 -0.366087 0.333461

## lstat 0.455621 -0.412995 0.603800 ... -0.366087 1.000000 -0.737663

## medv -0.388305 0.360445 -0.483725 ... 0.333461 -0.737663 1.000000
```

This value is quite high. An explanation *could* be that in the poorer neighborhoods, people didn't have the money to build new houses, so there are more houses built before 1940.

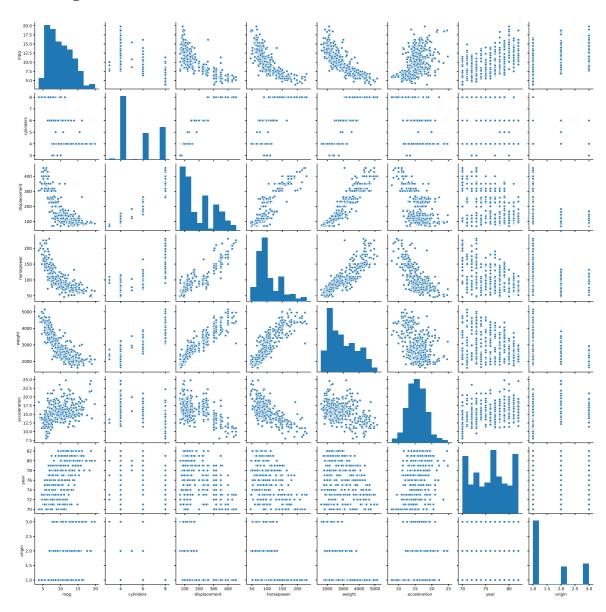
- The  $R^2$  value is 0.56, therefore about 56 % of the variation is explained by the model.
- The p-value of the F value is below the significance level and therefore significant. The null hypothesis  $H_0$

$$\beta_1 = \beta_2 = \beta_{12} = 0$$

is rejected. One of  $\beta$ 's is significantly different from 0. At least one variables contributes significantly to the model.

## Lösung 12.2

## a) Streudiagramme:



#### b) Korrelationsmatrix

```
df.corr()
##
                    mpg cylinders
                                           year
                                                  origin
## mpg
               1.000000 -0.777618
                                   ... 0.580541 0.565209
                                   ... -0.345647 -0.568932
              -0.777618 1.000000
## cylinders
## displacement -0.805127
                          0.950823
                                   ... -0.369855 -0.614535
              -0.778427
                                   ... -0.416361 -0.455171
## horsepower
                         0.842983
## weight
              -0.832244
                        0.897527
                                   ... -0.309120 -0.585005
  acceleration 0.423329 -0.504683 ... 0.290316 0.212746
               0.580541 -0.345647 ... 1.000000 0.181528
## year
```

```
## origin    0.565209 -0.568932 ... 0.181528 1.000000
##
## [8 rows x 8 columns]

df.corr().loc["horsepower", "displacement"]
## 0.897257001843467
```

#### c) Output

```
all_columns = "+".join(df.columns.drop("mpg"))
all columns
## 'cylinders+displacement+horsepower+weight+acceleration+year+origin'
fit = ols("mpg~" + all_columns , data=df).fit()
fit.summary()
## <class 'statsmodels.iolib.summary.Summary'>
## """
                        OLS Regression Results
##
## -----
                                                                                0.818
                                                                           252.4
2.04e-139
## Time: 10:18:15
## No. Observations: 392
## Df Residuals: 384
                                   392 AIC:
                                                                                 1392.
## Df Residuals:
                                       384 BIC:
                                                                                 1424.
## Df Model:
                                        7
## Covariance Type: nonrobust
coef std err t P>|t| [0.025 0.975]
## -----
## Intercept -7.3178 1.974 -3.707 0.000 -11.199 -3.437
## cylinders -0.2097 0.137 -1.526 0.128 -0.480 0.060
## displacement 0.0085 0.003 2.647 0.008 0.002 0.015
## horsepower -0.0072 0.006 -1.230 0.220 -0.019 0.004
## weight -0.0028 0.000 -9.929 0.000 -0.003 -0.002
## acceleration 0.0342 0.042 0.815 0.415 -0.048 0.117
## year 0.3191 0.022 14.729 0.000 0.276 0.362
## origin 0.6061 0.118 5.127 0.000 0.374 0.839
## -----
## Omnibus: 31.906 Durbin-Watson: 1.309
## Prob(Omnibus): 0.000 Jarque-Bera (JB): 53.100
## Skew: 0.529 Prob(JB): 2.95e-12
## Kurtosis: 4.460 Cond. No. 8.59e+04
                                                                             8.59e+04
                                    4.460 Cond. No.
##
##
## [1] Standard Errors assume that the covariance matrix of the errors is correctly specified.
\#\# [2] The condition number is large, 8.59e+04. This might indicate that there are
## strong multicollinearity or other numerical problems.
```

- i) Der *p*-Wert zum zugehörigen *F*-Wert ist praktisch 0 und somit besteht ein statistisch signifikanter Zusammenhang zwischen Zielvariable und den Prädiktoren.
- ii) Dies sind die Koeffizienten (displacement, weight, year und origin).

iii) Der Koeffizient für **year** ist positiv. Das heisst, man mit jüngeren Autos weiter pro Gallone Benzin kommt. Die neueren Autos sind als im Allgemeinen sparsamer.

#### d) Output:

```
fit = ols("mpg~weight*year" , data=df).fit()
fit.summary()
## <class 'statsmodels.iolib.summary.Summary'>
##
                                    OLS Regression Results
## Dep. Variable: mpg R-squared: 0.834
## Model: OLS Adj. R-squared: 0.833
## Method: Least Squares F-statistic: 649.3
## Date: Mon, 11 May 2020 Prob (F-statistic): 8.06e-151
## Time: 10:18:15 Log-Likelihood: -673.91
## No. Observations: 392 AIC: 1356.
## Df Residuals: 388 BIC: 1372.
## Df Model: 3
## Covariance Type: nonrobust
## -----
             coef std err t P>|t| [0.025 0.975]
##
## -----
## Intercept -46.9421 5.502 -8.531 0.000 -57.760 -36.124
## weight 0.0117 0.002 6.242 0.000 0.008 0.015
## year 0.8672 0.073 11.876 0.000 0.724 1.011
## weight:year -0.0002 2.51e-05 -7.752 0.000 -0.000 -0.000
## =
## Omnibus: 48.151 Durbin-Watson: 1.345
## Prob(Omnibus): 0.000 Jarque-Bera (JB): 86.841
## Skew: 0.722 Prob(JB): 1.39e-19
## Kurtosis: 4.798 Cond. No. 1.87e+07
                                         4.798 Cond. No.
## -----
## Warnings:
## [1] Standard Errors assume that the covariance matrix of the errors is correctly specified.
\#\# [2] The condition number is large, 1.87e+07. This might indicate that there are
## strong multicollinearity or other numerical problems.
fit.pvalues
## Intercept 3.295213e-16
## weight 1.136624e-09
## year 5.881890e-28
## year 5.881890e-28
## weight:year 8.015486e-14
## dtype: float64
```

Der p-Wert des Interaktionsterm ist von der Grössenordnung  $10^{-14}$ , also sehr nahe bei 0. Die Nullhypothese, dass keine Interaktion vorliegt, wird also verworfen.

Dies lässt sich damit erklären, dass das Gewicht mit den jüngeren Autos immer kleiner geworden ist.