# **PROJECT**

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## 23/11/22

### Question

Question. State the response variable you would like to model.

Ans. The response variable that I will model is the number of innings played by an player.

Question. State your proposed explanatory variables and explain why you think they would be accurate predictors.

Ans. The explanatory variables that I will model are "number of times a player scores 50" and "number of times a player scores 0". 0 means that the player gets out at 0 score.

Number of times a player scores 50 should help predict number of innings played by a player as if a player scores 50 'n' number of times then the player must had played at least n innings and it is rarely possible that in every innings a player score 50. So, a player must had played more than n matches if that players scores 50 in n matches.

Number of times a player scores 0 should help predict number of innings played by a player as if a player scores 0 'n' number of times then the player must had played at least n innings and it is rarely possible that in every innings a player gets out scoring 0. So, a player must had played more than n matches if that players scores 0 in n matches.

### Data Set

#### Dataset in tabular form

```
cric <- read.csv("stats.csv")
knitr::kable(cric, "latex", align=c("l", "c", "c"))</pre>
```

SR Tendulkar (INDIA)       452       96       2         V Kohli (INDIA)       245       62       1         RT Ponting (AUS/ICC)       365       82       2         RG Sharma (INDIA)       220       43       1         ST Jayasuriya (Asia/SL)       433       68       3         HM Amla (SA)       178       39       4         AB de Villiers (Afr/SA)       218       53       7         CH Gayle (ICC/WI)       294       54       2         KC Sangakkara (Asia/ICC/SL)       380       93       1	
V Kohli (INDIA)       245       62       1         RT Ponting (AUS/ICC)       365       82       2         RG Sharma (INDIA)       220       43       1         ST Jayasuriya (Asia/SL)       433       68       3         HM Amla (SA)       178       39       4         AB de Villiers (Afr/SA)       218       53       7         CH Gayle (ICC/WI)       294       54       2         KC Sangakkara (Asia/ICC/SL)       380       93       1	13 20 13 34
RT Ponting (AUS/ICC)       365       82       2         RG Sharma (INDIA)       220       43       1         ST Jayasuriya (Asia/SL)       433       68       3         HM Amla (SA)       178       39       4         AB de Villiers (Afr/SA)       218       53       7         CH Gayle (ICC/WI)       294       54       2         KC Sangakkara (Asia/ICC/SL)       380       93       1	20 13 34 4
RG Sharma (INDIA)       220       43       1         ST Jayasuriya (Asia/SL)       433       68       3         HM Amla (SA)       178       39       4         AB de Villiers (Afr/SA)       218       53       7         CH Gayle (ICC/WI)       294       54       2         KC Sangakkara (Asia/ICC/SL)       380       93       1	13 34 1
ST Jayasuriya (Asia/SL)       433       68       3         HM Amla (SA)       178       39       4         AB de Villiers (Afr/SA)       218       53       7         CH Gayle (ICC/WI)       294       54       2         KC Sangakkara (Asia/ICC/SL)       380       93       1	34 1
HM Amla (SA)       178       39       4         AB de Villiers (Afr/SA)       218       53       7         CH Gayle (ICC/WI)       294       54       2         KC Sangakkara (Asia/ICC/SL)       380       93       1	1
AB de Villiers (Afr/SA)         218         53         7           CH Gayle (ICC/WI)         294         54         2           KC Sangakkara (Asia/ICC/SL)         380         93         1	
CH Gayle (ICC/WI)         294         54         2           KC Sangakkara (Asia/ICC/SL)         380         93         1	
KC Sangakkara (Asia/ICC/SL) 380 93 1	25
	15
SC Ganguly (Asia/INDIA) 300 72 1	16
	l1
	)
	22
` '	15
	16
. , , ,	28
DA Warner (AUS) 126 23 2	2
ME Waugh (AUS) 236 50 1	16
AJ Finch (AUS) 128 29 1	11
S Dhawan (INDIA) 142 33 5	5
	13
JH Kallis (Afr/ICC/SA) 314 86 1	17
Q de Kock (SA) 124 26 4	1
JE Root (ENG) 142 35 5	5
MJ Guptill (NZ) 183 37 1	15
NJ Astle (NZ) 217 41 1	19
AC Gilchrist (AUS/ICC) 279 55 1	19
	17
	14
Mohammad Yousuf (Asia/PAK) 273 64 1	l5
Babar Azam (PAK) 81 17 3	3
Tamim Iqbal (BAN) 217 51 1	19
EJG Morgan (ENG/IRE) 228 47 1	16
Yuvraj Singh (Asia/INDIA) 278 52 1	18
KS Williamson (NZ) 144 39 5	<u> </u>
G Kirsten (SA) 185 45 1	l1
ME Trescothick (ENG) 122 21 1	13
PR Stirling (IRE) 131 26 1	10
F du Plessis (SA) 136 35 3	}
` ' ' '	13
, ,	3
SPD Smith (AUS) 113 25 5	5
	3
WTS Porterfield (IRE) 142 20 1	10
	l1
. ,	7
	l5
	19
S Chanderpaul (WI) 251 59 6	3

The data represents the record of 49 cricketers till the year 2021 in all the three formats. (ODI, T20 and test)

#### Reference

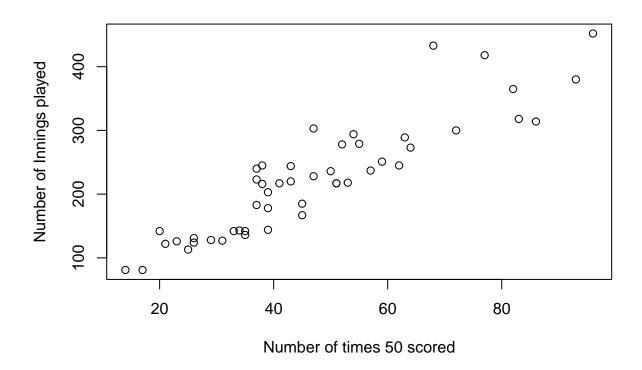
KRISHNA. (2021). Cricket Dataset. Retrieved from Cricket Dataset: https://www.kaggle.com/datasets/notkrishna/cricket-statistics-for-all-formats?select=odb.csv

#### Explanation of variable and the units of measurement.

The variable Inns is the number of innings a cricketer played. The variable X50 is the number of times a player scores 50. The variable X0 is the number of times a player scores 0.

### Scatter plot between Inns and X50.

plot(cric\$X50,cric\$Inns, xlab = "Number of times 50 scored",ylab = "Number of Innings played")



Calculating degree of variability between Inns and  ${\rm x50}(r^2)$ 

```
cor(cric$Inns,cric$X50)^(2)
```

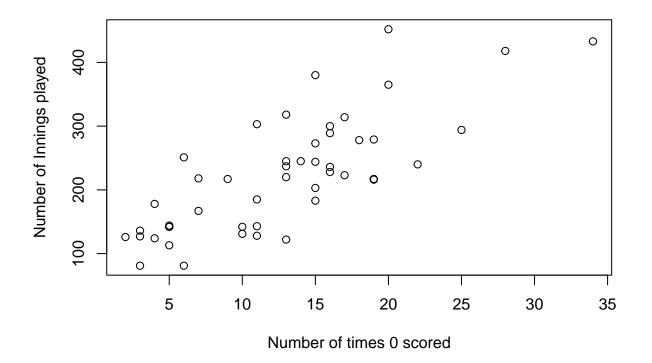
## [1] 0.8207977

 $r^2 = 0.8207977$ 

So, there exists a positive relationship between number of times 50 scored and number of innings played.

## Scatter plot between Inns and X0.

```
plot(cric$X0,cric$Inns, xlab = "Number of times 0 scored",ylab = "Number of Innings played")
```



Calculating degree of variability between Inns and  $x0(r^2)$ 

```
cor(cric$Inns,cric$X0)^(2)
```

## [1] 0.5943858

 $r^2 = 0.5943858$ 

So, there exists a postive relationship between number of times 0 scored and number of innings played.

## Preliminary Model

```
cric.lm <- lm(Inns ~ X0, data = cric)
summary(cric.lm)</pre>
```

```
##
## Call:
## lm(formula = Inns ~ XO, data = cric)
```

```
##
## Residuals:
##
       Min
                  1Q
                      Median
                      -1.044
## -102.244 -43.340
                                32.718 157.823
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
                                      5.359 2.47e-06 ***
## (Intercept)
                 94.368
                            17.608
                             1.204
## XO
                  9.990
                                      8.299 9.15e-11 ***
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
## Residual standard error: 57.8 on 47 degrees of freedom
## Multiple R-squared: 0.5944, Adjusted R-squared: 0.5858
## F-statistic: 68.87 on 1 and 47 DF, p-value: 9.148e-11
Regression Line:
\hat{y} = 94.368 + 9.99X1
cric2.lm <- lm(Inns ~ X50, data = cric)</pre>
summary(cric2.lm)
##
## Call:
## lm(formula = Inns ~ X50, data = cric)
##
## Residuals:
##
     Min
              1Q Median
                            3Q
                                  Max
## -67.96 -26.52 -6.90 23.65 124.55
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 30.7219
                           14.2349
                                      2.158
                                              0.0361 *
## X50
                 4.0842
                            0.2784 14.672
                                              <2e-16 ***
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
## Residual standard error: 38.42 on 47 degrees of freedom
## Multiple R-squared: 0.8208, Adjusted R-squared: 0.817
## F-statistic: 215.3 on 1 and 47 DF, p-value: < 2.2e-16
Regression Line:
\hat{y} = 30.7219 + 4.0842X2
cric3.lm <- lm(Inns ~ X0 + X50, data = cric)</pre>
summary(cric3.lm)
##
## Call:
## lm(formula = Inns ~ X0 + X50, data = cric)
##
```

```
## Residuals:
##
       Min
                 1Q Median
                                  30
                                         Max
## -50.861 -16.225 -1.501 10.758 89.794
##
## Coefficients:
##
                Estimate Std. Error t value Pr(>|t|)
## (Intercept) 11.4820
                              9.5984
                                       1.196
                                                 0.238
## XO
                  5.0304
                              0.6275
                                       8.017 2.79e-10 ***
## X50
                  3.1147
                              0.2183 14.268 < 2e-16 ***
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
## Residual standard error: 25.08 on 46 degrees of freedom
## Multiple R-squared: 0.9252, Adjusted R-squared: 0.922
## F-statistic: 284.7 on 2 and 46 DF, p-value: < 2.2e-16
RegressionLine: \hat{y} = 5.0304X1 + 3.1147X2 + 11.4820
For Inns ~ X0, R_{adi}^2 = 0.5858
For Inns ~ X50, R_{adi}^2 = 0.817
For Inns ~ X0 + X50, R_{adj}^2 = 0.922
So, R_{adj}^2 has increased from both Inns ~ X0 and Inns ~ X50.
```

### Second Order Model

```
cric.full <- lm(Inns ~ X0 + X50 + I(X0^2) + I(X50^2) + X0 * X50 , data = cric)
summary(cric.full)
##
## Call:</pre>
```

```
## lm(formula = Inns \sim X0 + X50 + I(X0^2) + I(X50^2) + X0 * X50,
##
       data = cric)
##
## Residuals:
      Min
                1Q Median
                                30
                                       Max
## -48.139 -11.953 -4.397
                                    93.778
                             7.728
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 32.725784 22.070708
                                      1.483 0.14543
## XO
               1.133768
                          1.916151
                                      0.592 0.55716
## X50
               3.122231
                           0.885361
                                      3.527 0.00102 **
## I(X0^2)
               0.099216
                           0.092283
                                      1.075 0.28831
## I(X50^2)
               -0.001708
                           0.011398
                                     -0.150
                                             0.88161
## X0:X50
               0.017499
                           0.063021
                                      0.278 0.78260
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
## Residual standard error: 24.41 on 43 degrees of freedom
```

```
## Multiple R-squared: 0.9338, Adjusted R-squared: 0.9261  
## F-statistic: 121.4 on 5 and 43 DF, p-value: < 2.2e-16  
Regression Line: \hat{y} = 32.725784 + 1.133768X1 + 3.122231X2 + 0.099216X1^2 - 0.001708X2^2 + 0.017499X1X2
```

#### Conducting ANOVA Test

```
LEVEL OF SIGNIFICANCE: 0.05
```

```
HYPOTHESES: H_o: \beta_1 = \beta_2 = 0 H_a: At least one \beta_i != 0 where i = 1,2
```

DECISION RULE: if p-value  $< \alpha$  then reject  $H_0$ . if P-value  $> \alpha$  then fail to reject  $H_0$ .

TEST STATISTIC: 121.4

P-VALUE:  $\approx 0$ 

CONCLUSION: p-value  $< \alpha$ . So, we reject  $H_0$ . So, we have sufficient evidence to conclude that at least one of the model terms is significant.

#### Model Refinement

```
summary(cric.full)
```

```
##
## Call:
## lm(formula = Inns \sim X0 + X50 + I(X0^2) + I(X50^2) + X0 * X50,
##
       data = cric)
##
## Residuals:
               1Q Median
                                3Q
##
      Min
                                      Max
## -48.139 -11.953 -4.397
                            7.728 93.778
##
## Coefficients:
               Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 32.725784 22.070708
                                    1.483 0.14543
               1.133768
                          1.916151
                                     0.592 0.55716
## XO
## X50
               3.122231
                          0.885361
                                     3.527 0.00102 **
                          0.092283
                                     1.075 0.28831
## I(X0^2)
               0.099216
## I(X50^2)
              -0.001708
                          0.011398 -0.150 0.88161
## X0:X50
               0.017499
                           0.063021
                                    0.278 0.78260
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
## Residual standard error: 24.41 on 43 degrees of freedom
## Multiple R-squared: 0.9338, Adjusted R-squared: 0.9261
## F-statistic: 121.4 on 5 and 43 DF, p-value: < 2.2e-16
```

Only the coefficient of X50 has significant p-value as it is less than 0.05. X0 has p-value greater than 0.05 but we will keep it in model as it is first order term.

p-value of X50: 0.00102 p-value of X0: 0.55716

#### Proposing Reduced Model

```
cric.reduced <- lm(Inns ~ X0 + X50, data= cric)</pre>
```

#### Performing nested F-Test

```
anova(cric.reduced,cric.full)
## Analysis of Variance Table
## Model 1: Inns ~ X0 + X50
## Model 2: Inns \sim X0 + X50 + I(X0^2) + I(X50^2) + X0 * X50
               RSS Df Sum of Sq
     Res.Df
                                     F Pr(>F)
## 1
         46 28944
         43 25622 3
## 2
                          3321.6 1.8582 0.151
LEVEL OF SIGNIFICANCE: 0.05
HYPOTHESES: H_o: \beta_3 = \beta_4 = \beta_5 = 0 H_a: At least one \beta_i != 0 where i = 3,4,5
DECISION RULE: if p-value < \alpha then reject H_0. if P-value > \alpha then fail to reject H_0.
TEST STATISTIC: 1.8582
P-VALUE: 0.151
```

CONCLUSION: p-value  $> \alpha$ . So, we fail to reject  $H_0$ . So, we have insufficient evidence to conclude that the coefficients of the terms  $X50^2$ ,  $X0^2$  and X50 \* X0 are significant.

#### Final Model and Assessment

#### Conducting ANOVA test on reduced model

```
summary(cric.reduced)
```

```
##
## Call:
## lm(formula = Inns ~ X0 + X50, data = cric)
##
## Residuals:
##
      Min
               1Q Median
                               3Q
                                      Max
## -50.861 -16.225 -1.501 10.758 89.794
##
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 11.4820
                           9.5984
                                    1.196
                                             0.238
                                    8.017 2.79e-10 ***
## XO
                5.0304
                           0.6275
## X50
                3.1147
                           0.2183 14.268 < 2e-16 ***
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
```

```
##
## Residual standard error: 25.08 on 46 degrees of freedom
## Multiple R-squared: 0.9252, Adjusted R-squared: 0.922
## F-statistic: 284.7 on 2 and 46 DF, p-value: < 2.2e-16</pre>
```

LEVEL OF SIGNIFICANCE: 0.05

HYPOTHESES:  $H_o$ :  $\beta_1 = \beta_2 = 0$   $H_a$ : At least one  $\beta_i != 0$  where i = 1,2

DECISION RULE: if p-value  $< \alpha$  then reject  $H_0$ . if P-value  $> \alpha$  then fail to reject  $H_0$ .

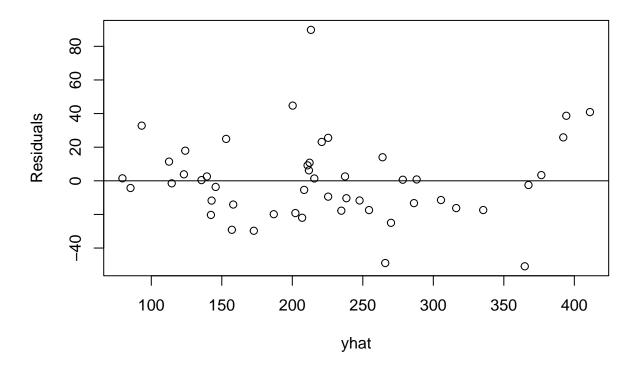
TEST STATISTIC: 284.7

P-VALUE:  $\approx 0$ 

CONCLUSION: p-value  $< \alpha$ . So, we reject  $H_0$ . So, we have sufficient evidence to conclude that at least one of the model terms is significant.

#### **Residual Plot:**

```
cric.res <- resid(cric.reduced)
cric.fitted <- fitted.values(cric.reduced)
plot(cric.fitted,cric.res,ylab = "Residuals", xlab="yhat")
abline(h = 0)</pre>
```



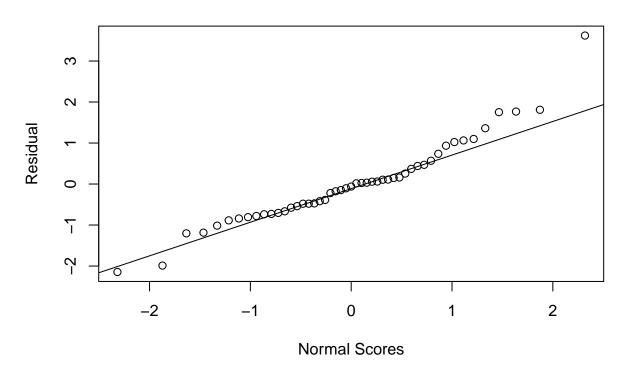
Linearity: There is no discernible pattern in the points so the data is linear.

Not Constant Variance: The width of the error terms is not similar throughout the plot and be the same above and below the axis.

### Normal Quantile Plot:

```
cric.stdres = rstandard(cric.reduced)
qqnorm(cric.stdres, ylab = "Residual", xlab = "Normal Scores")
qqline(cric.stdres)
```

## Normal Q-Q Plot



Normality of Error terms: error terms follow a normal distribution as all observations are near the quantilequantile line.

### Conclusion

According to my opinion, the explanatory variables were able to predict the response variable as the final reduced model has those variables. Moreover, the model passed the nested f test. Another thing that I want to add is that the data follows normal distribution according to the normal quantile quantile plot and residual plot.

### Final Regression Line

```
\hat{y} = 5.0304X1 + 3.1147X2 + 11.4820
```