

COMPSCI 589

Lecture 19: Principal Components Analysis

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Outline

1 Review

2 Linear Algebra

3 PCA

4 Connection to SVD

The Dimensionality Reduction Task

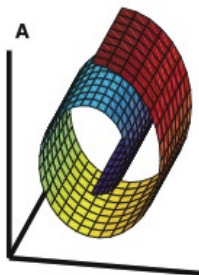
Definition: The Dimensionality Reduction Task

Given a collection of feature vectors $\mathbf{x}_i \in \mathbb{R}^D$, map the feature vectors into a lower dimensional space $\mathbf{z}_i \in \mathbb{R}^K$ where $K < D$ while preserving certain properties of the data.

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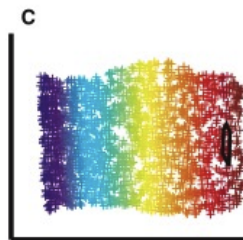
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high-dim distribution



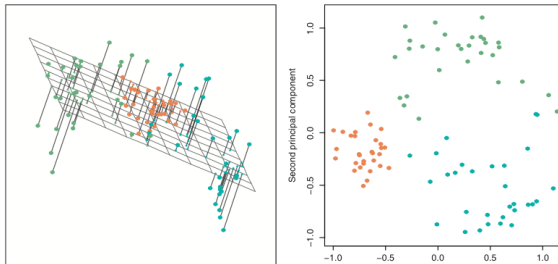
high-dim samples



estimated manifold

Linear Dimensionality Reduction

- The simplest dimensionality reduction methods assume that the observed high dimensional data vectors $\mathbf{x}_i \in \mathbb{R}^D$ lie on a K -dimensional linear manifold within \mathbb{R}^D .
- Mathematically, the linear sub-space assumption can be written as $\mathbf{X} = \mathbf{Z} \times \mathbf{B}$



Learning

- The learning problem for linear dimensionality reduction is to estimate values for both \mathbf{Z} and \mathbf{B} given only the noisy observations \mathbf{X} .
- One possible learning criteria is to minimize the sum of squared errors when reconstructing \mathbf{X} from \mathbf{Z} and \mathbf{B} . This leads to:

$$\arg \min_{\mathbf{Z}, \mathbf{B}} \|\mathbf{X} - \mathbf{Z}\mathbf{B}\|_F$$

where $\|\mathbf{A}\|_F$ is the Frobenius norm of matrix \mathbf{A} (the sum of the squares of all matrix entries).

Singular Value Decomposition

- We can pick a unique representation for the subspace by specifying additional criteria. Classical Rank- K Singular Value Decomposition (K-SVD) corresponds to the following restriction:

$$\arg \min_{\mathbf{U}, \mathbf{S}, \mathbf{V}} \|\mathbf{X} - \mathbf{USV}^T\|_F$$

where \mathbf{S} is a $K \times K$ diagonal matrix with positive elements, \mathbf{U} is an $N \times K$ matrix such that $\mathbf{U}^T \mathbf{U} = \mathbf{I}$, and \mathbf{V} is a $D \times K$ matrix such that $\mathbf{V}^T \mathbf{V} = \mathbf{I}$.

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- The matrix product $\mathbf{Z} = \mathbf{US}$ gives the optimal rank-K representation of \mathbf{X} with respect to Frobenius norm minimization, with \mathbf{V}^T acting as the basis for the space.

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- A full-rank (invertible) matrix $\mathbf{A} \in \mathbb{R}^{D \times D}$ will have D linearly independent eigenvectors.

Eigendecomposition

- Let $\mathbf{V} \in \mathbb{R}^{D \times D}$ be a matrix whose columns \mathbf{v}_d are D linearly independent eigenvectors of \mathbf{A} with Λ the corresponding diagonal matrix of eigenvalues such that $\Lambda_{dd} = \lambda_d$. Then:

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- Without loss of generality, we can assume that

$$\lambda_1 > \lambda_2 > \dots > \lambda_n$$

Eigendecomposition of a Symmetric Matrix

- If \mathbf{A} is symmetric, we can choose D orthonormal eigenvectors so that $\|\mathbf{v}_d\|_2 = 1$, $\mathbf{v}_d^T \mathbf{v}_{d'} = 0$ and D real eigenvalues $\lambda_d \in \mathbb{R}$. This representation of \mathbf{A} is unique. As a result, we have:

$$\mathbf{A} = \mathbf{V} \mathbf{\Lambda} \mathbf{V}^T = \sum_{d=1}^D \lambda_d \mathbf{v}_d \mathbf{v}_d^T$$

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Representation of a Vector in the Eigen Basis

- Similarly, if \mathbf{a} is an arbitrary vector, then we can also represent \mathbf{a} using the basis provided by the eigenvectors \mathbf{V} of a real symmetric matrix \mathbf{A} . We obtain:

$$\mathbf{a} = \sum_{d=1}^D \alpha_d \mathbf{v}_d \quad (1)$$

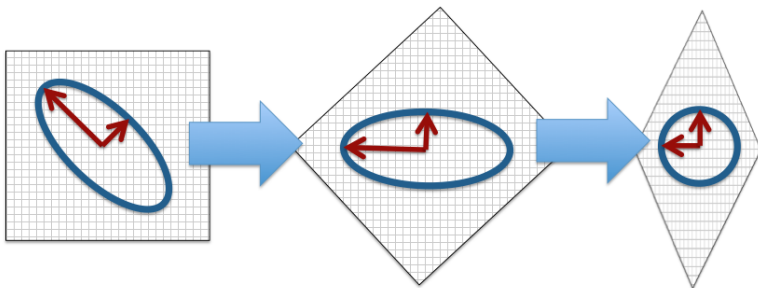
$$\alpha_d = \mathbf{a}^T \mathbf{v}_d \quad (2)$$

Geometry

- If \mathbf{A} is a real symmetric matrix with positive eigenvalues, then the quadratic equation $\mathbf{x}^T \mathbf{A} \mathbf{x} = 0$ defines an ellipsoid in a D -dimensional space, which provides a different way of thinking about these operations:

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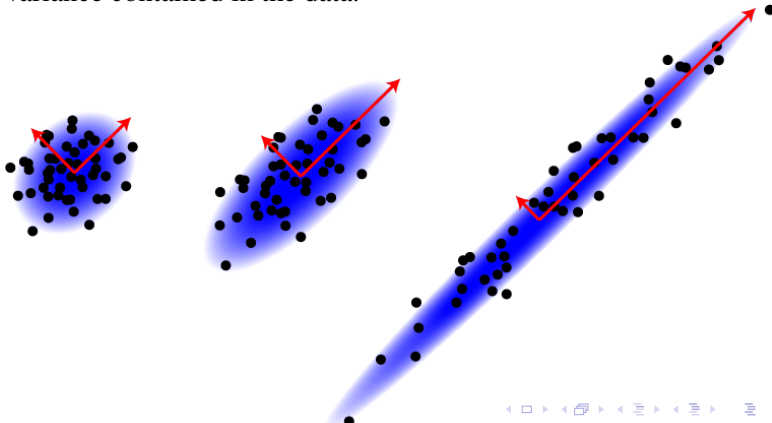
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Principal Component Analysis

- Given a data matrix $\mathbf{X} \in \mathbb{R}^{N \times D}$, the goal of Principal Component Analysis (PCA) is to identify the directions of maximum variance contained in the data.

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- Let $\mathbf{w} \in \mathbb{R}^D$ such that $\|\mathbf{w}\|_2 = \sqrt{\mathbf{w}^T \mathbf{w}} = 1$.
- The sample estimate of the variance in the direction \mathbf{w} given the data set \mathbf{X} is given by the expression:

$$\frac{1}{N} \sum_{i=1}^N (\mathbf{X}_i \mathbf{w} - \mu)^2 \quad \text{where} \quad \mu = \frac{1}{N} \sum_{i=1}^N \mathbf{X}_i \mathbf{w}$$

Pre-Centering

- Under the assumption that the data are pre-centered so that $\frac{1}{N} \sum_{i=1}^N \mathbf{X}_i = 0$, this expression simplifies to:

$$\frac{1}{N} \sum_{i=1}^N (\mathbf{X}_i \mathbf{w})^2 = (\mathbf{X} \mathbf{w})^T (\mathbf{X} \mathbf{w}) = \mathbf{w}^T \mathbf{X}^T \mathbf{X} \mathbf{w}$$

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- How can we solve this problem?

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- $\mathbf{V}_d \in \mathbb{R}^D$ are the eigenvectors of Σ . They satisfy:

$$\|\mathbf{V}_d\|_2 = \sqrt{\mathbf{V}_d^T \mathbf{V}_d} = 1 \dots \text{for all } d$$

$$\mathbf{V}_d^T \mathbf{V}_{d'} = 0 \dots \text{for all } d \neq d'$$

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- Using this result, we can write the optimization problem as:

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- The constraint that $\|\mathbf{w}\|_2 = 1$ becomes $\sqrt{\sum_{d=1}^D \omega_d^2} = 1$.
- This means $\sum_{d=1}^D \omega_d^2 = 1$ and $\omega_d^2 > 0$, so the ω_d^2 values act like a discrete probability distribution.

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- Plugging this back into the objective function, we have:

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- Working our way back to \mathbf{w}_1 , we put all our weight on the maximum eigenvalue, so $\mathbf{w} = \sum_{d=1}^D \omega_d \mathbf{V}_d = \mathbf{V}_1$.
- **This shows that the maximum variance direction given a data matrix \mathbf{X} is the eigenvector of $\mathbf{X}^T \mathbf{X}$ with the largest eigenvalue.**

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- Suppose instead of just the direction of maximum variance, we want the K largest directions of variance that are all mutually orthogonal.
- Finding the second-largest direction of variance corresponds to solving the problem:

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- **In general, the top K directions of variance $\mathbf{w}_1, \dots, \mathbf{w}_K$ are given by the K eigenvectors corresponding to the K largest eigenvalues of $\mathbf{X}^T \mathbf{X}$.**

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- 3 Stack the eigenvectors together into a $D \times K$ matrix \mathbf{W} where each column k of \mathbf{W} corresponds to \mathbf{w}_k .
- 4 Project the matrix \mathbf{X} into the rank- K sub-space of maximum variance by computing the matrix product $\mathbf{Z} = \mathbf{X}\mathbf{W}$.
- 5 To reconstruct \mathbf{X} given \mathbf{Z} and \mathbf{W} , we use $\hat{\mathbf{X}} = \mathbf{Z}\mathbf{W}^T$.

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Connection to SVD

- Last class we saw that the minimum Frobenius norm linear dimensionality reduction problem could be solved using the the rank-K SVD of \mathbf{X} :

$$\arg \min_{\mathbf{U}, \mathbf{S}, \mathbf{V}} \|\mathbf{X} - \mathbf{USV}^T\|_F$$

where the matrix product $\mathbf{Z} = \mathbf{US}$ gives the optimal rank-K representation of \mathbf{X} with respect to Frobenius norm minimization.

Connection to SVD

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- This means that the K largest singular values and K largest eigenvalues correspond to the same K basis vectors.

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- Finally, note that if the decompositions are based only on the K leading basis vectors, which are identical under both PCA and SVD, the projections $\mathbf{Z} = \mathbf{XW}$ and $\mathbf{Z} = \mathbf{US}$ will still be identical.

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- As a result, generic linear dimensionality reduction simultaneously minimizes the Frobenius norm of the reconstruction error of \mathbf{X} and maximizes the retained variance in the learned sub-space.
- Both SVD and PCA provide the same refinement of generic linear dimensionality reduction: an orthogonal basis for exactly the same optimal linear subspace.

Issues

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- If $K \ll D$, then PCA can also be computed iteratively, as can SVD.
- The basic SVD and PCA algorithms are not suitable for large-scale data. Instead, randomized algorithms are often used.
- The value of K can sometimes be chosen based on looking for eigenvalue gaps in the eigenspectrum of the covariance matrix. Otherwise, a supervised end/side-task is needed or a criteria like AIC/BIC must be applied.