

## Functions of a Random Variable

1. Let  $X$  be a geometric random variable with  $E[X] = 1/p$ . Find the PMF of the following functions of  $X$ :

(a)  $Y = X^2$ ;

*Ans:* Since  $S_X = \{1, 2, 3, \dots\}$ ,  $S_Y = \{1, 4, 9, 16, \dots\} = \{k^2 : k = 1, 2, \dots\}$ . The PMF of  $Y$  is

$$p_Y(k^2) = p_X(k) = p(1-p)^{k-1}, \quad k = 1, 2, 3, \dots$$

Alternatively, we can say that

$$p_Y(y) = p(1-p)^{\sqrt{y}-1}, \quad y \in S_Y.$$

(b)  $Z = \ln(X)$ ;

*Ans:* Now we have  $Z \in \{\ln k : k = 1, 2, \dots\}$ , and the PMF of  $Z$  is

$$p_Z(\ln k) = p_X(k) = p(1-p)^{k-1}, \quad k = 1, 2, 3, \dots$$

Alternatively, we can say that

$$p_Z(z) = p(1-p)^{e^z-1}, \quad z \in \{\ln k : k = 1, 2, \dots\}.$$

(c)  $V = e^X$ .

*Ans:* The range of  $V$  is  $S_V = \{e^k : k = 1, 2, \dots\}$ , and its PMF is

$$p_V(e^k) = p_X(k) = p(1-p)^{k-1}, \quad k = 1, 2, \dots$$

We can also write

$$p_V(v) = p(1-p)^{\ln v-1}, \quad v \in S_V.$$

2. The number  $X$  is drawn at random from the unit interval  $[0, 1]$ , and  $Y$  is defined as  $X$  rounded to the nearest tenth, i.e. if  $X = 0.12$  then  $Y = 0.1$ , if  $X = 0.87$  then  $Y = 0.9$ , etc. Find the PMF of  $Y$ .

*Ans:* The range of  $Y$  is  $S_Y = \{0, 0.1, 0.2, \dots, 0.9, 1.0\}$ . The event  $\{Y = 0\}$  is equivalent to  $\{0 \leq X < 0.05\}$  and therefore  $p_Y(0) = P[0 \leq X < 0.05] = 0.05$ . Similarly, we have  $\{Y = 0.1\}$  being equivalent to  $\{0.05 \leq X < 0.15\}$  and hence  $p_Y(0.1) = 0.1$ . Proceeding this way we can see that the PMF of  $Y$  is

$$p_Y(y) = \begin{cases} 0.05, & y = 0, 1.0 \\ 0.1, & y = 0.1, 0.2, 0.3, 0.4, 0.5, 0.6, 0.7, 0.8, 0.9 \end{cases} \quad (1)$$

3. A wire has length  $X$ , an exponential random variable with mean  $5\pi$  cm. The wire is cut to make rings of diameter 1 cm. Let  $Y$  be the number of complete rings made from the wire. Find the PMF of  $Y$ .

*Ans:* Each ring requires  $\pi$  cm of wire. Therefore, with  $X$  cm of wire, we can make  $Y = \lfloor X/\pi \rfloor$  complete rings, where  $\lfloor x \rfloor =$  the largest whole number smaller than  $x$ , e.g.  $\lfloor 4.5 \rfloor = 4$ ,  $\lfloor 2.9 \rfloor = 2$ , etc.

The event  $\{Y = 0\}$  is equivalent to  $\{X < \pi\}$ , while  $\{Y = 1\}$  is equivalent to  $\{\pi \leq X < 2\pi\}$ . Generalizing we have

$$\{Y = k\} \equiv \{k\pi \leq X < (k+1)\pi\},$$

and therefore the PMF of  $Y$  is

$$p_Y(k) = P[k\pi \leq X < (k+1)\pi].$$

Using the CDF of the exponential random variable with parameter  $\lambda = 1/5\pi$ , we have finally

$$p_Y(k) = e^{-k/5} - e^{-(k+1)/5}, \quad k = 0, 1, 2, \dots$$

4. Let  $X$  be uniformly distributed in  $(0, 1]$ . Sketch the function  $g(x)$  in the following cases, and then find the PDF of  $Y = g(X)$ .

- (a)  $g(x) = x^2$ ;

*Ans:*  $Y = g(X)$  has the range  $(0, 1]$ . Using the direct PDF method, and noting that  $g'(x) = 2x$ ,

$$f_Y(y) = \frac{f_X(\sqrt{y})}{2\sqrt{y}} \quad (2)$$

$$= \frac{1}{2\sqrt{y}}, \quad 0 < y \leq 1. \quad (3)$$

- (b)  $g(x) = e^{-x}$ ;

*Ans:* Now, let's try the CDF route.

$$F_Y(y) = P[e^{-X} \leq y] \quad (4)$$

$$= P[X \geq -\ln y] \quad (5)$$

$$= \begin{cases} 0 & y < e^{-1} \\ 1 + \ln y, & e^{-1} \leq y < 1 \\ 1 & y \geq 1 \end{cases} \quad (6)$$

Therefore the PDF of  $Y$  is

$$f_Y(y) = \frac{1}{y}, \quad e^{-1} \leq y < 1.$$

(c)  $g(x) = \cos 2\pi x$ .

*Ans:* For any  $y \in [-1, 1]$ ,  $g(x) = y$ ,  $x \in (0, 1]$ , has two solutions:  $x_0 = \frac{1}{2\pi} \cos^{-1} y$  and  $x_1 = 1 - \frac{1}{2\pi} \cos^{-1} y$ . We find that the derivative of  $g(x)$  is  $g'(x) = -2\pi \sin 2\pi x$  and thus

$$g'(x_0) = -2\pi \sqrt{1 - y^2}, \quad g'(x_1) = 2\pi \sqrt{1 - y^2}.$$

Using the formula  $f_Y(y) = \sum_{i=1}^N f_X(x_i)/|g'(x_i)|$ , we have

$$f_Y(y) = \frac{1}{2\pi \sqrt{1 - y^2}} + \frac{1}{2\pi \sqrt{1 - y^2}} \quad (7)$$

$$= \frac{1}{\pi \sqrt{1 - y^2}}, \quad -1 \leq y \leq 1. \quad (8)$$

5. Let  $X \sim \mathcal{N}(2, 4)$ , and

$$Y = (X)^+ = \begin{cases} X & X > 0 \\ 0 & X \leq 0 \end{cases}$$

Find the CDF and hence the PDF of  $Y$ .

*Ans:* We have

$$P[Y = 0] = P[X < 0] = 1 - P[X \geq 0]. \quad (9)$$

Recall that for any Gaussian r.v.  $X$ ,  $P[X > x] = Q((x - \mu)/\sigma)$ , where  $Q(x) = P[\mathcal{N}(0, 1) > x]$ . Therefore

$$P[Y = 0] = 1 - Q\left(\frac{-2}{2}\right) = 1 - Q(-1) = Q(1). \quad (10)$$

Since  $P[Y = 0] = Q(1) \neq 0$ , there must be a step increase in  $F_Y(y)$  at  $y = 0$ , with  $F_Y(0) = Q(1)$  and  $F_Y(y) = 0$  for  $y < 0$ .

The CDF of  $Y$ , for  $y > 0$ , would be (from a sketch of  $y = (x)^+$ )

$$F_Y(y) = P[Y \leq y] = P[X \leq y] \quad (11)$$

$$= 1 - P[X > y] = 1 - Q\left(\frac{y - 2}{2}\right) \quad (12)$$

Thus, we have the entire CDF of  $Y$ .

The PDF of  $Y$  is obtained by differentiating  $F_Y(y)$ . Note that

$$Q(x) = 1 - \Phi(x)$$

where  $\Phi(x)$  is the CDF of the  $\mathcal{N}(0, 1)$  random variable, and therefore

$$\frac{dQ(x)}{dx} = -\frac{d\Phi(x)}{dx} = -\frac{1}{\sqrt{2\pi}} e^{-x^2/2}.$$

For  $y > 0$ , we have

$$\begin{aligned} f_Y(y) &= \frac{dF_Y(y)}{dy} \\ &= -\frac{1}{2} \left[ -\frac{1}{\sqrt{2\pi}} e^{-(y-2)^2/8} \right] \\ &= \frac{1}{2\sqrt{2\pi}} \exp\left(-\frac{(y-2)^2}{8}\right). \end{aligned}$$

There is also a delta function at  $y = 0$  of strength  $Q(1)$  to account for  $P[Y = 0] = Q(1)$ . In total,

$$f_Y(y) = Q(1)\delta(y) + \frac{1}{2\sqrt{2\pi}} \exp\left(-\frac{(y-2)^2}{8}\right) u(y).$$

**Note:** The formula for finding the PDF of  $g(X)$  directly cannot be used here, because  $g(x)$  is flat, i.e.  $g'(x) = 0$ , over a range of values.

6. Let  $X$  be a Rayleigh random variable with PDF

$$f_X(x) = xe^{-x^2/2}, \quad x > 0.$$

Find the PDF of  $Z = X^2$  by (i) first finding the CDF, and (ii) finding the PDF directly.

*Ans:* (i) In the first method, we find the CDF and then the PDF through differentiation. From first principles,

$$F_Z(z) = P[Z \leq z] = P[X^2 \leq z] \tag{13}$$

$$= P[X \leq \sqrt{z}] \tag{14}$$

$$= \int_0^{\sqrt{z}} xe^{-x^2/2} dx, \quad z > 0. \tag{15}$$

Using the substitution  $u = \frac{x^2}{2}$ , we have

$$F_Z(z) = \int_0^{z/2} e^{-u} du = 1 - e^{-z/2}, \quad z > 0. \tag{16}$$

Therefore,

$$f_Z(z) = \frac{dF_Z(z)}{dz} = \frac{1}{2} e^{-z/2}, \quad z > 0. \tag{17}$$

(ii) We can find the PDF directly from

$$f_Z(z) = \sum_{i=1}^N \frac{f_X(x_i)}{|g'(x_i)|}$$

by first identifying that  $g(x) = z$  has only one solution, i.e.  $x = \sqrt{z}$  within the range of  $X$  (i.e.  $(0, \infty)$ ). Therefore  $N = 1$  in the above formula. Also  $g'(x) = 2x$ , and hence  $g'(\sqrt{z}) = 2\sqrt{z}$ . Finally, we obtain

$$f_Z(z) = \frac{\sqrt{z}e^{-z/2}}{2\sqrt{z}} = \frac{1}{2}e^{-z/2}, \quad z > 0$$

as in method (i).

7. A random variable  $Y$  is said to be log-normally distributed if  $X = \ln Y$  is an  $\mathcal{N}(\mu, \sigma^2)$  random variable. Find the PDF of  $Y$ , and hence show that  $E[Y] = \exp\left[\mu + \frac{\sigma^2}{2}\right]$ .

*Ans:* Essentially we want to find the PDF of  $Y = e^X$ , when  $X \sim \mathcal{N}(\mu, \sigma^2)$ . Using the direct method, we solve for  $x$  in  $e^x = y$ , i.e.  $x = \ln y$ ; then find the derivative of  $e^x$  at  $x = \ln y$ , i.e.  $g'(\ln y) = y$ . Substituting these into the formula gives us

$$f_Y(y) = \frac{1}{\sqrt{2\pi\sigma^2}} \frac{e^{-(\ln y - \mu)^2/2\sigma^2}}{y}, \quad y > 0. \quad (18)$$

To find the mean of  $Y$ , we can use the definition of  $E[Y]$ :

$$E[Y] = \int_0^\infty \frac{1}{\sqrt{2\pi\sigma^2}} e^{-(\ln y - \mu)^2/2\sigma^2} dy \quad (19)$$

$$= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^\infty e^{-(v - \mu)^2/2\sigma^2} e^v dv \quad (20)$$

where we substituted  $v = \ln y$  to obtain the second line. This integral is dealt with by “completing the squares” in the exponent, i.e.

$$\frac{(v - \mu)^2}{2\sigma^2} - v = \frac{1}{2\sigma^2} [(v - \mu)^2 - 2\sigma^2 v] \quad (21)$$

$$= \frac{1}{2\sigma^2} [v^2 - 2(\mu + \sigma^2)v + \mu^2] \quad (22)$$

$$= \frac{1}{2\sigma^2} [(v - (\mu + \sigma^2))^2 - (\mu + \sigma^2)^2 + \mu^2] \quad (23)$$

$$= \frac{1}{2\sigma^2} [(v - (\mu + \sigma^2))^2] - \mu - \frac{\sigma^2}{2}. \quad (24)$$

Substitution of this final expression into (20) yields

$$E[Y] = e^{\mu + \sigma^2/2} \cdot \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^\infty e^{-(v - (\mu + \sigma^2))^2/2\sigma^2} dv \quad (25)$$

$$= e^{\mu + \sigma^2/2} \quad (26)$$

because the integral on the RHS above is the area under the  $\mathcal{N}(\mu + \sigma^2, \sigma^2)$  PDF, which is 1 by definition. (Completing the squares in the exponent is an often-used tool that helps to solve certain problems involving Gaussian densities.)

8. The input to a full-wave rectifier is  $X$  and its output is  $Y = |X|$ . Find the PDF of  $Y$  if  $X \sim \mathcal{N}(0, \sigma^2)$ .

*Ans:* Let  $Y = g(X) = |X|$ , then  $|g'(x)| = 1$  for all  $x \neq 0$ . Furthermore,  $g(x) = y$ ,  $y > 0$ , has two solutions  $x_1 = -y$  and  $x_2 = +y$ . Therefore, using the direct method for finding the PDF of  $Y$ , we get

$$f_Y(y) = \frac{2}{\sqrt{2\pi\sigma^2}} e^{-y^2/2\sigma^2}, \quad y > 0. \quad (27)$$