# Linear Regression Normal Equation – Additional Results

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#### Abstract

We derive the normal equation for linear regression and show that the mean-squared-error is a convex function.

### 1 Notation

Let  $(\boldsymbol{x}^{(1)}, y^{(1)}), \dots, (\boldsymbol{x}^{(m)}, y^{(m)}) \in \mathbb{R}^n \times \mathbb{R}$  denote the collection of training examples. The *i*th training example consists of the feature vector  $\boldsymbol{x}^{(i)} = (x_1^{(i)}, \dots, x_n^{(i)})^T \in \mathbb{R}^n$  and the label  $y^{(i)} \in \mathbb{R}$ .

Set

$$\boldsymbol{X} = \begin{pmatrix} x_1^{(1)} & \dots & x_n^{(1)} \\ x^{(2)} & \dots & x_n^{(2)} \\ \vdots & \ddots & \vdots \\ x^{(m)} & \dots & x_n^{(m)} \end{pmatrix} \in \mathbb{R}^{m \times n}$$

and

$$oldsymbol{y} = \left(egin{array}{c} y^{(1)} \ y^{(2)} \ dots \ y^{(m)} \end{array}
ight) \in \mathbb{R}^m.$$

 $\boldsymbol{X}$  is called the design matrix. Its rows correspond to the feature vectors of the training examples.

# 2 Normal equation

To simplify the discussion, consider first the case that the bias of the linear regression model is set to 0, that is, only the weight vector  $\mathbf{w} = (w_1, \dots, w_n)^T \in \mathbb{R}^n$  needs to be determined.

**Theorem 1** (Normal equation). The optimal weight vector  $\mathbf{w} = (w_1, \dots, w_n)^T \in \mathbb{R}^n$ , that is, the one that minimizes the mean squared error is given by the formula

$$\boldsymbol{w} = (\boldsymbol{X}^T \boldsymbol{X})^{-1} \boldsymbol{X}^T \boldsymbol{y}.$$

This is proved in 5.1.4 Example: Linear Regression in [1]. I have included this proof with additional results so you can understand every step of the proof.

### 3 Additional results

We introduce some abbreviations. Let  $[n] = \{1, \ldots, n\}$ . Let  $\partial w_r$  denote the partial derivative operator

 $\frac{\partial}{\partial w_r}$ .

**Lemma 1** (Gradient of quadratic form). Let  $\mathbf{A} = (a_{rs}) \in \mathbb{R}^{n \times n}$  be an arbitrary symmetric matrix and  $\mathbf{w} = (w_1, \dots, w_n)^T \in \mathbb{R}^n$  an arbitrary column vector. Define the quadratic form  $f(\mathbf{w}) = \mathbf{w}^T \mathbf{A} \mathbf{w}$ . Its gradient is given by

$$\nabla_{\boldsymbol{w}} f(\boldsymbol{w}) = 2\boldsymbol{A}\boldsymbol{w}$$
.

*Proof.* The right hand side is the column vector whose entries are given by

$$2\sum_{s=1}^{n}a_{rs}w_{s}.$$

for  $r \in [n]$ . This follows simply by carrying out the matrix-vector-multiplication.

The left hand side of the above equation is the column vector whose entries are the partial derivatives

$$\partial w_r f(\boldsymbol{w})$$

for  $r \in [n]$ . This follows from the definition of the nabla operator

$$\nabla_{\boldsymbol{w}} = \left(\begin{array}{c} \partial w_1 \\ \vdots \\ \partial w_n \end{array}\right) .$$

We have

$$\partial w_r f(\boldsymbol{w}) = \partial w_r \left( \sum_{t,s=1} w_t a_{ts} w_s \right)$$

$$= \partial w_r \left( w_r^2 a_{rr} + 2 \sum_{s \neq r} w_r a_{rs} w_s \right)$$

$$= 2 w_r a_{rr} + 2 \sum_{s \neq r} a_{rs} w_s$$

$$= 2 \sum_{s=1}^n a_{rs} w_s.$$

We use that either t = r and s = r or t = r and  $s \neq r$ . Otherwise the partial derivative  $\partial w_r(w_t a_{ts} w_s)$  is equal to 0. We also use that  $\mathbf{A}$  is symmetric, that is,  $a_{rs} = a_{sr}$ .

**Lemma 2.** Let  $\mathbf{w} = (w_1, \dots, w_n)^T \in \mathbb{R}^n$  and  $\mathbf{v} = (v_1, \dots, v_n)^T \in \mathbb{R}^n$  be an arbitrary column vectors. Define the function  $g(\mathbf{w}) = \mathbf{w}^T \mathbf{v}$ . Its gradient is given by

$$\nabla_{\boldsymbol{w}}g(\boldsymbol{w}) = \boldsymbol{v}$$
.

*Proof.* This is easy. Prove it yourself.

# 4 Proof of normal equation

Let  $\hat{y}^{(i)} = \boldsymbol{w}^T \boldsymbol{x}^{(i)}$  denote the prediction of the linear regression model with weight vector  $\boldsymbol{w}$  when given the *i*th feature vector  $\boldsymbol{x}^{(i)}$ . Define

$$\hat{m{y}} = \left( egin{array}{c} \hat{y}^{(1)} \ dots \ \hat{y}^{(n)} \end{array} 
ight)$$

The mean squared error (MSE) is given by

MSE = 
$$\frac{1}{m} \sum_{i=1}^{m} (\hat{y}^{(i)} - y^{(i)})^2$$
  
=  $\frac{1}{m} ||\hat{\boldsymbol{y}} - \boldsymbol{y}||_2^2$ ,

so the error increases whenever the Euclidean distance between the predictions and the targets (labels) increases.

To minimize MSE, we can simply solve for where its gradient is 0:

$$\nabla_{\boldsymbol{w}} MSE = \boldsymbol{0} \tag{1}$$

$$\Rightarrow \nabla_{\boldsymbol{w}} \frac{1}{m} \|\hat{\boldsymbol{y}} - \boldsymbol{y}\|_{2}^{2} = \mathbf{0}$$
 (2)

$$\Rightarrow \frac{1}{m} \nabla_{\boldsymbol{w}} \| \boldsymbol{X} \boldsymbol{w} - \hat{\boldsymbol{y}} \|_{2}^{2} = \boldsymbol{0}$$
 (3)

$$\Rightarrow \frac{1}{m} \nabla_{\boldsymbol{w}} (\boldsymbol{X} \boldsymbol{w} - \hat{\boldsymbol{y}})^T (\boldsymbol{X} \boldsymbol{w} - \hat{\boldsymbol{y}}) = \boldsymbol{0}$$
 (4)

$$\Rightarrow \frac{1}{m} \nabla_{\boldsymbol{w}} (\boldsymbol{w}^T \boldsymbol{X}^T \boldsymbol{X} \boldsymbol{w} - 2 \boldsymbol{w}^T \boldsymbol{X}^T \boldsymbol{y} + \boldsymbol{y}^T \boldsymbol{y}) = \boldsymbol{0}$$
 (5)

$$\Rightarrow 2\boldsymbol{X}^{T}\boldsymbol{X}\boldsymbol{w} - 2\boldsymbol{X}^{T}\boldsymbol{y} = \boldsymbol{0} \tag{6}$$

$$\Rightarrow \boldsymbol{w} = (\boldsymbol{X}^T \boldsymbol{X} \boldsymbol{w})^{-1} \boldsymbol{X}^T \boldsymbol{y} \tag{7}$$

Note that  $\hat{y}^{(i)}$  is the inner product between the weight vector  $\boldsymbol{w}$  and the feature vector  $\boldsymbol{x}^{(i)}$ . We also have  $\hat{y}^{(i)} = (\boldsymbol{x}^{(i)})^T \boldsymbol{w}$  because the inner product is symmetric.

In eq. (5), we can use Lemma 1 with  $\mathbf{A} = \mathbf{X}^T \mathbf{X}$  and Lemma 2 with  $\mathbf{v} = \mathbf{X}^T \mathbf{y}$  to compute the gradient. The system of equations whose solutions is given by eq. 7 is known as normal equations.

The general case of linear regression with non-zero bias b can also be solved with the help of the normal equation. Define the augmented weight vector  $\mathbf{w}_b = (b, w_1, \dots, w_n)^T \in \mathbb{R}^{n+1}$  and the augmented feature vectors  $\mathbf{x}_b^{(i)} = (1, x_1^{(i)}, \dots, x_n^{(i)})^T \in \mathbb{R}^{n+1}$ . We have

$$\hat{y}^{(i)} = \boldsymbol{w}_b^T \boldsymbol{y}_b^{(i)} = \boldsymbol{w}^T \boldsymbol{y}^{(i)} + b.$$

# 5 Proof of convexity of mean-squared-error

We now show that MSE is convex in  $\boldsymbol{w}$ . We can write

$$MSE(\boldsymbol{w}) = \phi(\|\boldsymbol{X}\boldsymbol{w} - \boldsymbol{y}\|),$$

where  $\phi:[0,\infty)\to\mathbb{R}, \phi(x)=\frac{1}{m}x^2$  is non-decreasing and convex. Therefore, it suffices to show that  $f(\boldsymbol{w})=\|\boldsymbol{X}\boldsymbol{w}-\boldsymbol{y}\|$  is convex. This is seen by invoking the following simple lemma.

**Lemma 3.** Suppose  $\phi:[0,\infty)\to\mathbb{R}$  is non-decreasing and convex and  $f:\mathbb{R}^n\to[0,\infty)$  is convex. For  $p\in[0,1]$  and  $r,s\in\mathbb{R}^n$ , we have

$$\phi(f(p\mathbf{r} + (1-p)\mathbf{s})) \leq \phi(f(p\mathbf{r}) + (1-p)f(\mathbf{s}))$$
  
$$\leq p\phi(f(\mathbf{r})) + (1-p)\phi(f(\mathbf{s})).$$

Hence,  $\phi \circ f$  is convex.

Let  $\boldsymbol{w}$  and  $\tilde{\boldsymbol{w}}$  be two weight vectors. We have

$$f(p\boldsymbol{w} + (1-p)\tilde{\boldsymbol{w}})) = \|\boldsymbol{X}(p\boldsymbol{w} + (1-p)\tilde{\boldsymbol{w}}) - \boldsymbol{y}\|$$

$$\leq \|p(X\boldsymbol{w} - y) + (1-p)(\boldsymbol{X}\tilde{\boldsymbol{w}} - \boldsymbol{y})\|$$

$$\leq p\|X\boldsymbol{w} - y\| + (1-p)\|\boldsymbol{X}\tilde{\boldsymbol{w}} - \boldsymbol{y}\|$$

$$= pf(\boldsymbol{w}) + (1-p)f(\tilde{\boldsymbol{w}}).$$

In the above derivation, we have used the triangle inequality  $\|\mathbf{r} + \mathbf{s}\| \leq \|\mathbf{r}\| + \|\mathbf{s}\|$  and  $\|\lambda \mathbf{r}\| = |\lambda| \|\mathbf{r}\|$ , which hold for arbitrary  $\lambda \in \mathbb{R}$  and  $\mathbf{r}, \mathbf{s} \in \mathbb{R}^n$ .

We see that the mean-squared-error is a convex function by combining all results.

## References

[1] I. Goodfellow, Y. Bengio, and A. Courville, *Deep learning*, MIT Press, 2006, http://www.deeplearningbook.org