Linear Discriminant Analysis

• Example: binary (two class) classification

Logit:
$$\log \frac{\Pr(G=1|X=x)}{1-\Pr(G=1|X=x)} = \log \frac{\Pr(G=1|X=x)}{\Pr(G=2|X=x)} = \beta_0 + x^T \beta$$

The posterior probability

$$Pr(G = 1|X = x) = \frac{\exp(\beta_0 + x^T \beta)}{1 + \exp(\beta_0 + x^T \beta)}, \exp(x) = e^x$$

$$Pr(G = 2|X = x) = \frac{1}{1 + \exp(\beta_0 + x^T \beta)}, X_2 \uparrow_{>0.5}^{\Pr(G = 1|X = x)} \downarrow_{>0.5}^{\Pr(G = 1|X = x)}$$
ion boundary

Decision boundary

$$\{x|\beta_0 + x^T\beta = 0\}$$

 $\begin{array}{c}
\operatorname{Core}(\beta_0 + x^T \beta) \\
\operatorname{Decision boundary} \\
\operatorname{Pr}(G = 1 | X = x) \\
= \operatorname{Pr}(G = 2 | X = x) = 0.5
\end{array}$ Green: class 1
Blue: class 2 $\begin{array}{c}
\operatorname{Pr}(G = 2 | X = x) \\
> 0.5
\end{array}$

Fisher's Formulation of Discriminant Analysis

LDA: Approach 1

- 1. Estimating $\hat{\Sigma}$, $\hat{\mu}_k$ and $\hat{\pi}_k$
- 2. Discriminant function

$$\delta_k(x) = x^T \widehat{\mathbf{\Sigma}}^{-1} \widehat{\mu}_k - \frac{1}{2} \widehat{\mu}_k^T \widehat{\mathbf{\Sigma}}^{-1} \widehat{\mu}_k + \log \widehat{\pi}_k$$

3. Classify to class *k* that maximizes the discriminant function

$$\widehat{G}(x) = \operatorname*{argmax}_{k \in \mathcal{G}} \delta_k(x)$$

Q: why data sphering makes $\widehat{\Sigma}^* = I$?

Hint:
$$\widehat{\boldsymbol{\Sigma}} = \frac{\sum_{k=1}^{K} \sum_{g_i=k} (x_i - \widehat{\mu}_k) (x_i - \widehat{\mu}_k)^T}{N - K}$$

LDA: Approach 2

- 1. Estimating $\hat{\Sigma}$, $\hat{\mu}_k$ and $\hat{\pi}_k$
- 2. Eigen-decomposition:

$$\widehat{\Sigma} = \mathbf{U}\mathbf{D}\mathbf{U}^T$$

3. Data sphering $(\hat{\Sigma}^* = I)$

$$\mathbf{D}^{-\frac{1}{2}}\mathbf{U}^{T}x = \widehat{\mathbf{\Sigma}}^{-\frac{1}{2}}x$$

$$\hat{\mu}_{k}^{*} = \mathbf{D}^{-\frac{1}{2}}\mathbf{U}^{T}\hat{\mu}_{k} = \widehat{\mathbf{\Sigma}}^{-\frac{1}{2}}\hat{\mu}_{k}$$

4. Classify to its closest class centroid in the transformed space

$$G(x) = \underset{k \in \mathcal{G}}{\operatorname{argmin}} \frac{1}{2} \|x^* - \hat{\mu}_k^*\|^2 - \ln \hat{\pi}_k$$

Eg. 2 Dice roll problem (6 outcomes instead of 2)



Likelihood is \sim Multinomial($\theta = \{\theta_1, \theta_2, ..., \theta_k\}$)

$$P(\mathcal{D} \mid \theta) = \theta_1^{\alpha_1} \theta_2^{\alpha_2} \dots \theta_k^{\alpha_k}$$

If prior is Dirichlet distribution,

$$P(\theta) = \frac{\theta_1^{\beta_1 - 1} \ \theta_2^{\beta_2 - 1} \dots \theta_k^{\beta_k - 1}}{B(\beta_1, \dots, \beta_k)} \sim \text{Dirichlet}(\beta_1, \dots, \beta_k)$$

Then posterior is Dirichlet distribution

$$P(\theta|D) \sim \mathsf{Dirichlet}(\beta_1 + \alpha_1, \dots, \beta_k + \alpha_k)$$

and MAP estimate is therefore

$$\hat{\theta_i}^{MAP} = \frac{\alpha_i + \beta_i - 1}{\sum_{j=1}^k (\alpha_j + \beta_j - 1)}$$

Estimating Parameters: Y, X, discrete-valued

Maximum likelihood estimates:

$$\frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c} 1 \\ 1 \end{array} \right\} = \frac{1}{2} \left\{ \begin{array}{c$$

$$\widehat{\underline{\chi}}_{k} = \widehat{P}(Y = y_{k}) = \frac{\#D\{Y = y_{k}\}}{|D|}$$

$$\hat{\theta}_{ijk} = \hat{P}(X_i = x_j | Y = y_k) = \frac{\#D\{X_i = x_j \land Y = y_k\}}{\#D\{Y = y_k\}}$$

MAP estimates (Beta, Dirichlet priors):

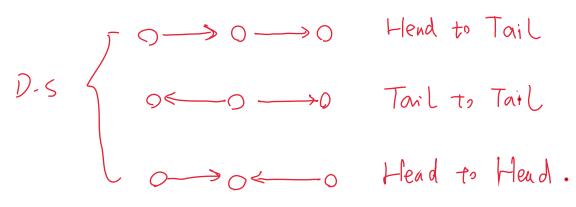
$$\hat{\pi}_k = \hat{P}(Y = y_k) = \frac{\#D\{Y = y_k\} + (\beta_k - 1)}{|D| + \sum_m (\beta_m - 1)}$$
 "imaginary" examples
$$\hat{\theta}_{ijk} = \hat{P}(X_i = x_j | Y = y_k) = \frac{\#D\{X_i = x_j \land Y = y_k\} + (\beta_k - 1)}{\#D\{Y = y_k\} + \sum_m (\beta_m - 1)}$$

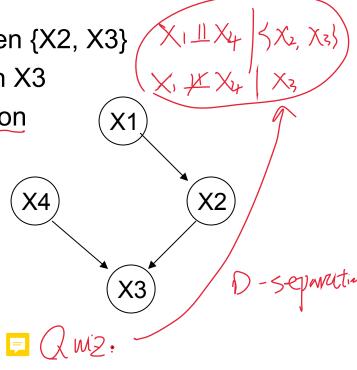
What is the Bayes Network for Naïve Bayes?



Conditional Independence, Revisited

- We said:
 - Each node is conditionally independent of its non-descendents, given its immediate parents.
- Does this rule give us all of the conditional independence relations implied by the Bayes network?
 - No!
 - E.g., X1 and X4 are conditionally indep given {X2, X3}
 - But X1 and X4 not conditionally indep given X3
 - For this, we need to understand D-separation





EM Algorithm - Precisely

EM is a general procedure for learning from partly observed data

Given observed variables X, unobserved Z (X={F,A,H,N}, Z={S})/

Define
$$Q(\theta'|\theta) = E_{P(Z|X,\theta)}[\log P(X,Z|\theta')]$$
 next step current current current

Iterate until convergence:

- E Step: Use X and current θ to calculate $P(Z|X,\theta) = \frac{Y(X,E|\theta)}{\sum Y(X,E|\theta)}$
- M Step: Replace current θ by

$$\frac{\theta \leftarrow \arg\max_{\theta'} Q(\theta'|\theta)}{Q(\theta)|\theta|} = \frac{Q(\theta'|\theta)}{Q(\theta)|\theta|} = \frac{Q(\theta'|\theta)}{Q(\theta)} = \frac{Q(\theta'|\theta$$

Guaranteed to find local maximum.

Each iteration increases $E_{P(Z|X,\theta)}[\log P(X,Z|\theta')]$





$$\frac{\partial Q}{\partial Q_f} = 0 \Rightarrow Q_f = 0$$

$$\frac{\partial Q}{\partial Q_g} = 0 \Rightarrow Q_g = 0$$

Sample Complexity for Supervised Learning

Consistent Learner

- Input: S: $(x_1,c^*(x_1)),...,(x_m,c^*(x_m))$
- · Output: Find h in H consistent with the sample (if one exits).

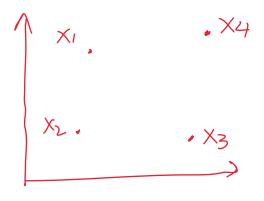
Theorem

$$m \geq \frac{1}{\varepsilon} \left[\ln(|H|) + \ln\left(\frac{1}{\delta}\right) \right] \qquad \text{or} \qquad \blacksquare$$

labeled examples are sufficient so that with prob. $1 - \delta$, all $h \in H$ with $err_D(h) \ge \varepsilon$ have $err_S(h) > 0$.

Contrapositive: if the target is in H, and we have an algo that can find consistent fns, then we only need this many examples to get generalization error $\leq \epsilon$ with prob. $\geq 1 - \delta$

Today's Quiz



$$(1)$$
 $H(s) = \frac{7}{2}$

Analyzing Training Error: Proof Math

Step 1: unwrapping recurrence: $D_{T+1}(i) = \frac{1}{m} \left(\frac{\exp(-y_i f(x_i))}{\prod_t Z_t} \right)$ where $f(x_i) = \sum_t \alpha_t h_t(x_i)$.

Step 2: $\operatorname{err}_{S}(H_{final}) \leq \prod_{t} Z_{t}$.

Step 3:
$$_t Z_t = \prod_t 2\sqrt{\epsilon_t(1-\epsilon_t)} = _t \sqrt{1-4\gamma_t^2} \le e^{-2\sum_t \gamma_t^2}$$

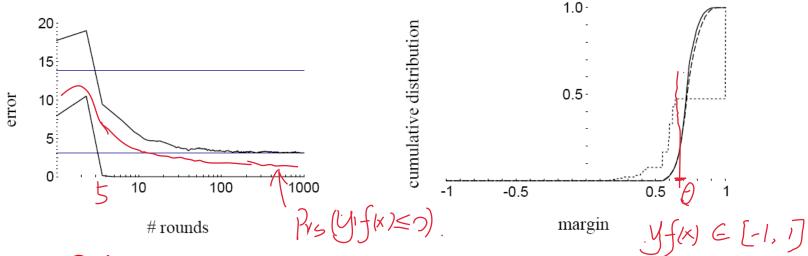
Note: recall $Z_t = (1 - \epsilon_t)e^{-\alpha_t} + \epsilon_t e^{\alpha_t} = 2\sqrt{\epsilon_t(1 - \epsilon_t)}$ $\alpha_t \text{ minimizer of } \alpha \to (1 - \epsilon_t)e^{-\alpha} + \epsilon_t e^{\alpha} \quad \blacksquare \quad \mathbf{Quiz}$

$$\begin{array}{l} e\gamma\gamma_{o}(g) \leq e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ e\gamma\gamma_{o}(g) = P_{ro}(y(x) \neq y) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)} \neq y) \end{array} \begin{array}{l} e\gamma\gamma_{s}(g) + \beta\left(\sqrt{dT}\right) \\ = P_{ro}(H_{f}^{(k)}$$

Theorem: VCdim(H) = d, then with prob. $\geq 1 - \delta$, $\forall f \in co(H)$, $\forall \theta > 0$,

$$\Pr_{D}[yf(x) \neq \emptyset] \leq \Pr_{S}[yf(x) \leq \theta] + O\left(\frac{1}{\sqrt{m}} \sqrt{\frac{d \ln^{2} \frac{m}{d}}{\theta^{2}} + \ln \frac{1}{\delta}}\right) = \tilde{O}\left(\sqrt{\frac{d}{m}} \sqrt{\frac{d \ln^{2} \frac{m}{d}}{\theta^{2}}}\right)$$
Threshold

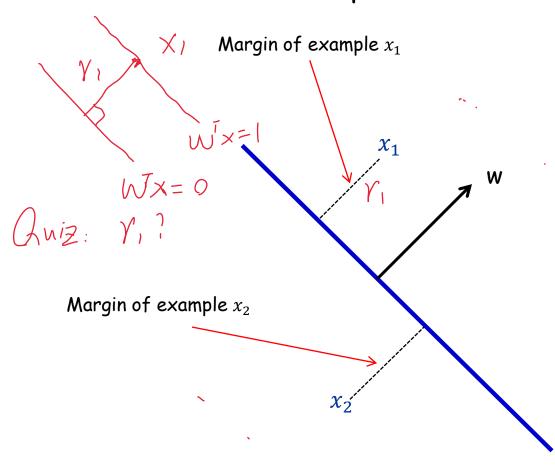
Note: bound does not depend on T (the # of rounds of boosting), depends only on the complex. of the weak hyp space and the margin!



Quiz: according to this slide, explain why adaboost keeps decreasing testing error, even if training error equals to zero.

Geometric Margin

Definition: The margin of example x w.r.t. a linear sep. w is the distance from x to the plane $w \cdot x = 0$.



If ||w|| = 1, margin of x w.r.t. w is $|x \cdot w|$.

Support Vector Machines (SVMs)

Question: what if data isn't perfectly linearly separable? Replace "# mistakes" with upper bound called "hinge loss"

```
Input: S=\{(x_1, y_1), ..., (x_m, y_m)\};
  Find \operatorname{argmin}_{W,\xi_1,...,\xi_m} ||w||^2 + C \sum_i \xi_i \text{ s.t.}:
               For all i, y_i w \cdot x_i \ge 1 - \xi_i
                          \xi_i \geq 0
                                             D y; w'x; ≥1, ξ;=0
              \xi_i are "slack variables" y_i y_i y_i^T x_i < 1
C controls the relative weighting between the
twin goals of making the ||w||^2 small (margin is
large) and ensuring that most examples have
functional margin \geq 1.
             hinge loss: [= max (0, 1- ywx)
                                                                 l(w, x, y) = \max(0, 1 - y w \cdot x)
```

Modern ML: New Learning Approaches

Modern applications: massive amounts of raw data.

Techniques that best utilize data, minimizing need for expert/human intervention.

Paradigms where there has been great progress.

· Semi-supervised Learning, (Inter)active Learning.









An Easy Case for k-means: k=1

Input: A set of n datapoints $x^1, x^2, ..., x^n$ in R^d

Output: $c \in \mathbb{R}^d$ to minimize $\sum_{i=1}^n ||\mathbf{x}^i - \mathbf{c}||^2$

Solution: The optimal choice is $\mu = \frac{1}{n} \sum_{i=1}^{n} x^{i}$

Idea: bias/variance like decomposition

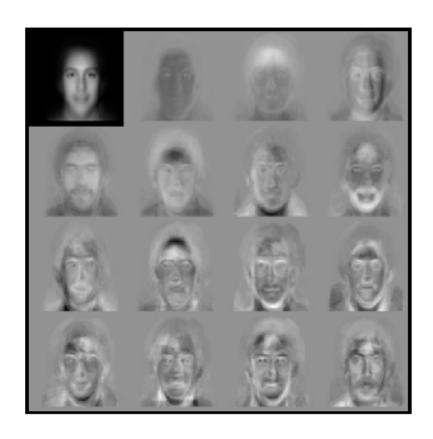
$$\frac{1}{n}\sum_{i=1}^{n} ||\mathbf{x}^{i} - \mathbf{c}||^{2} = ||\mu - \mathbf{c}||^{2} + \frac{1}{n}\sum_{i=1}^{n} ||\mathbf{x}^{i} - \mu||^{2}$$
 Quiz

Avg k-means cost wrt c

Avg k-means cost wrt μ

So, the optimal choice for c is μ .

Example: faces



Figenfaces from 7562 images:

top left image is linear combination of rest.

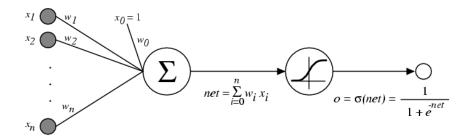
Sirovich & Kirby (1987) Turk & Pentland (1991)

Can represent a face image using just 15 numbers!





Error Gradient for a Sigmoid Function



 $x_d = input$

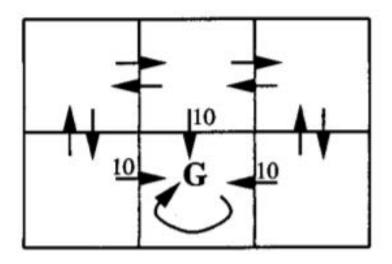
t_d = target output

o_d = observed unit output

 w_i = weight i



Quiz



\gamma=0.9

- (1) Give an optimal policy for the above problem;
- (2) Calculate the V*(s) values;
- (3) Calculate the Q(s,a) values.

Positive semidefinite cone

notation:

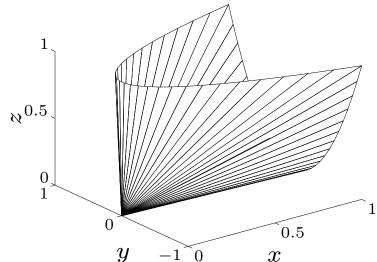
- S^n is set of symmetric $n \times n$ matrices
- $\mathbf{S}_{+}^{n} = \{X \in \mathbf{S}^{n} \mid X \succeq 0\}$: positive semidefinite $n \times n$ matrices

$$X \in \mathbf{S}^n_+ \iff z^T X z \ge 0 \text{ for all } z$$

 \mathbf{S}^n_+ is a convex cone

• $\mathbf{S}_{++}^n = \{X \in \mathbf{S}^n \mid X \succ 0\}$: positive definite $n \times n$ matrices

example: $\begin{bmatrix} x & y \\ y & z \end{bmatrix} \in \mathbf{S}_{+}^{2}$



Quiz:

Is S_{++}^n a convex cone? Show the reason.

Affine function

suppose $f: \mathbf{R}^n \to \mathbf{R}^m$ is affine $(f(x) = Ax + b \text{ with } A \in \mathbf{R}^{m \times n}, b \in \mathbf{R}^m)$

ullet the image of a convex set under f is convex

$$S \subseteq \mathbf{R}^n \text{ convex} \implies f(S) = \{f(x) \mid x \in S\} \text{ convex}$$

ullet the inverse image $f^{-1}(C)$ of a convex set under f is convex

Quiz $C \subseteq \mathbf{R}^m$ convex $\implies f^{-1}(C) = \{x \in \mathbf{R}^n \mid f(x) \in C\}$ convex

Pointwise supremum

if f(x,y) is convex in x for each $y \in \mathcal{A}$, then

$$g(x) = \sup_{y \in \mathcal{A}} f(x, y)$$

is convex Quiz

examples

- support function of a set C: $S_C(x) = \sup_{y \in C} y^T x$ is convex
- distance to farthest point in a set C:

$$f(x) = \sup_{y \in C} ||x - y||$$

ullet maximum eigenvalue of symmetric matrix: for $X \in \mathbf{S}^n$,

$$\lambda_{\max}(X) = \sup_{\|y\|_2 = 1} y^T X y$$

Log-concave and log-convex functions

a positive function f is log-concave if $\log f$ is concave:

$$f(\theta x + (1 - \theta)y) \ge f(x)^{\theta} f(y)^{1-\theta}$$
 for $0 \le \theta \le 1$

f is log-convex if $\log f$ is convex

- powers: x^a on \mathbf{R}_{++} is log-convex for $a \leq 0$, log-concave for $a \geq 0$
- \bullet many common probability densities are log-concave, e.g., normal:

Quiz
$$f(x) = \frac{1}{\sqrt{(2\pi)^n \det \Sigma}} e^{-\frac{1}{2}(x-\bar{x})^T \Sigma^{-1}(x-\bar{x})}$$

ullet cumulative Gaussian distribution function Φ is log-concave

$$\Phi(x) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{x} e^{-u^2/2} \, du$$

LP and SOCP as SDP

LP and equivalent SDP

LP: minimize c^Tx SDP: minimize c^Tx subject to $Ax \leq b$ subject to $\operatorname{diag}(Ax - b) \leq 0$

(note different interpretation of generalized inequality \leq)

SOCP and equivalent SDP

SOCP: minimize f^Tx subject to $\|A_ix + b_i\|_2 \le c_i^Tx + d_i, \quad i = 1, \dots, m$

SDP: minimize f^Tx subject to $\begin{bmatrix} (c_i^Tx+d_i)I & A_ix+b_i \\ (A_ix+b_i)^T & c_i^Tx+d_i \end{bmatrix} \succeq 0, \quad i=1,\ldots,m$

Quiz: how to represent QP as SDP?

Lagrange dual and conjugate function

minimize
$$f_0(x)$$

subject to $Ax \leq b$, $Cx = d$

dual function

$$g(\lambda, \nu) = \inf_{x \in \text{dom } f_0} \left(f_0(x) + (A^T \lambda + C^T \nu)^T x - b^T \lambda - d^T \nu \right)$$
$$= -f_0^* (-A^T \lambda - C^T \nu) - b^T \lambda - d^T \nu$$

- recall definition of conjugate $f^*(y) = \sup_{x \in \mathbf{dom}\ f} (y^T x f(x))$
- ullet simplifies derivation of dual if conjugate of f_0 is known

example: entropy maximization

$$f_0(x) = \sum_{i=1}^n x_i \log x_i, \qquad f_0^*(y) = \sum_{i=1}^n e^{y_i - 1}$$

Quiz: derive the dual function of entropy maximization problem.