

# CS182 - Introduction to Machine Learning, 2021-22 Fall

## Final Exam

9:00AM - 11:00AM, Saturday, Jan. 8th, 2022

12 pages, 6 problems, and 120 points (including 20 points for bonus) in total

(NOTE: your exam grade will be counted as  $\min\{100, \text{"the grade on test paper"}\}$ )

### Problem 1. (20 points) (Bayesian Decision Theory and Linear Discrimination)

- 1) Consider a binary classification problem, with the class conditional density

$$p(\mathbf{x} | C_i) = \frac{1}{(2\pi)^{\frac{d}{2}} \det(\Sigma_i)^{\frac{1}{2}}} \exp \left[ -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu}_i)^T \Sigma_i^{-1} (\mathbf{x} - \boldsymbol{\mu}_i) \right], \quad i = 1, 2.$$

Assume that  $\Sigma_i = \sigma^2 \mathbf{I}$  for all  $i$ , and that the priors for the two classes are not equal, i.e.,  $P(C_1) \neq P(C_2)$ .

If our target is to minimize the expected loss, a.k.a. conditional risk, (as in the lecture notes, we denote by  $\lambda_{ij}$  the loss incurred for assigning an input  $\mathbf{x}$  to class  $C_i$  when the actual state is  $C_j$ ), derive the decision boundary, and explain how geometrically it differs from that when one minimizes the misclassification error, a.k.a. probability of error. (10 points)

- 2) Assume that  $\mathbf{x}$  is 2-dimensional, and that the two dimensions are independent following the Laplace distribution:

$$p(x_j | C_i) = \frac{1}{2\sigma} \exp\left(-\frac{|x_j - \mu_{ij}|}{\sigma}\right), \quad j = 1, 2.$$

By minimizing the misclassification error, obtain and draw the decision boundary when  $\mu_{11} = 1$ ,  $\mu_{12} = 1$ ,  $\mu_{21} = 3$ ,  $\mu_{22} = 5$ ,  $\sigma = 1$ , and  $P(C_1) = P(C_2)$ . (10 points)

### Solution:

- 1) When using the loss functions in the binary classification case, we will decide  $\omega_1$  if  $(\lambda_{21} - \lambda_{11})p(\mathbf{x}|C_1)P(C_1) > (\lambda_{22} - \lambda_{12})p(\mathbf{x}|C_2)P(C_2)$ . (1') Taking the logarithm and the discriminant function can be obtained by

$$\begin{aligned} g_i(\mathbf{x}) &= \ln c_i + \ln p(\mathbf{x} | C_i) + \ln P(C_i) \\ &= -\frac{\|\mathbf{x} - \boldsymbol{\mu}_i\|^2}{2\sigma^2} + [\ln c_i + \ln P(C_i) + \text{const.}], \quad (1') \end{aligned}$$

where  $c_1 = (\lambda_{21} - \lambda_{11})$ ,  $c_2 = (\lambda_{22} - \lambda_{12})$ ,  $\text{const.}$  is a constant.

By applying the result in slides, the decision boundary is a line:  $\mathbf{w}^T(\mathbf{x} - \mathbf{x}_0) = 0$ , (1') where

$$\begin{aligned} \mathbf{w} &= \boldsymbol{\mu}_1 - \boldsymbol{\mu}_2, \\ \mathbf{x}_0 &= \frac{1}{2}(\boldsymbol{\mu}_1 + \boldsymbol{\mu}_2) - \frac{\sigma^2}{\|\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2\|^2} \ln \frac{c_1 P(C_1)}{c_2 P(C_2)} (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2). \quad (2') \end{aligned}$$

We know the decision boundary obtained by minimizing the error is the line  $\mathbf{w}^T(\mathbf{x} - \mathbf{x}_{e0}) = 0$ , where

$$\mathbf{x}_{e0} = \frac{1}{2}(\boldsymbol{\mu}_1 + \boldsymbol{\mu}_2) - \frac{\sigma^2}{\|\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2\|^2} \ln \frac{P(C_1)}{P(C_2)} (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2). \quad (1')$$

Since

$$\begin{aligned}
 \mathbf{x}_0 &= \frac{1}{2} (\boldsymbol{\mu}_1 + \boldsymbol{\mu}_2) - \frac{\sigma^2}{\|\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2\|^2} \ln \frac{c_1 P(C_1)}{c_2 P(C_2)} (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2) \\
 &= \frac{1}{2} (\boldsymbol{\mu}_1 + \boldsymbol{\mu}_2) - \frac{\sigma^2}{\|\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2\|^2} \ln \frac{P(C_1)}{P(C_2)} (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2) - \frac{\sigma^2}{\|\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2\|^2} \ln \frac{c_1}{c_2} (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2) \quad (2') \\
 &= \mathbf{x}_{e0} - \frac{\sigma^2}{\|\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2\|^2} \ln \frac{c_1}{c_2} (\boldsymbol{\mu}_1 - \boldsymbol{\mu}_2),
 \end{aligned}$$

The decision boundary obtained by minimizing the loss is a translation of its error minimization counterpart. (2')

2) When minimize error, the discriminant functions are

$$g_i(\mathbf{x}) = -\ln(2\sigma) - \frac{\|\mathbf{x} - \boldsymbol{\mu}_i\|_1}{\sigma} + \ln P(C_i), \quad i = 1, 2. \quad (2')$$

The decision boundary satisfies the equation

$$g_1(\mathbf{x}) - g_2(\mathbf{x}) = 0.$$

That is,

$$\|\mathbf{x} - \boldsymbol{\mu}_1\|_1 - \|\mathbf{x} - \boldsymbol{\mu}_2\|_1 - \sigma \ln \frac{P(C_1)}{P(C_2)} = 0. \quad (2')$$

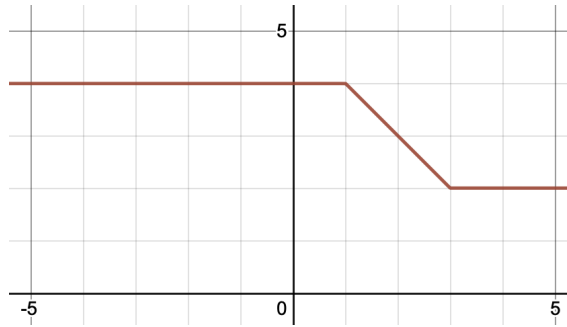
Using those values, the boundary satisfies

$$(|x_1 - 1| - |x_1 - 3|) + (|x_2 - 1| - |x_2 - 5|) = 0. \quad (2')$$

The decision boundary is

$$\begin{cases} x_1 + x_2 = 5, & 1 \leq x_1 \leq 3, \\ x_2 = 2, & x_1 > 3, \\ x_2 = 4, & x_1 < 1. \end{cases} \quad (2')$$

The following figure shows the decision boundary (2')



**Problem 2. (20 points) (SVM and Decision Trees)**

A novel multi-class classification method is called Decision Trees SVM (DT-SVM), which is a binary decision tree where every binary split is trained by an SVM. Given the input feature  $\mathbf{x}$ , a classification result from the DT-SVM is given in the following figure where  $C_i$  denotes the region of class  $i$ .

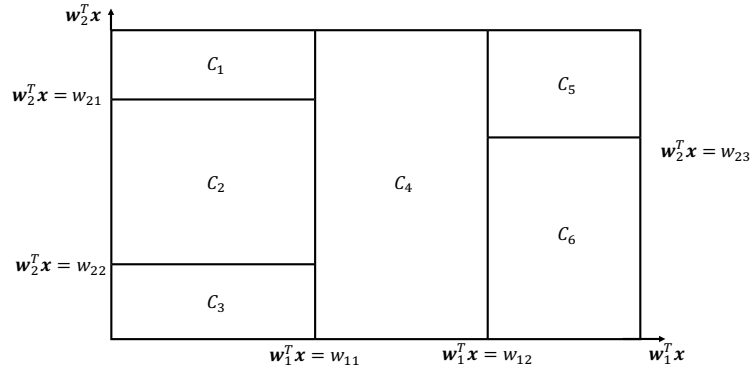
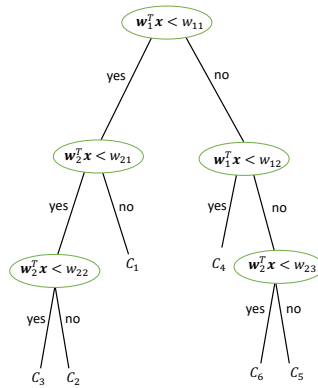


Figure: A classification result from the DT-SVM.

- 1) Draw the corresponding binary decision tree according to the above figure. (6 points)
- 2) You have learned the hard-margin SVMs and the soft-margin SVMs. Discuss the differences between them. (7 points)
- 3) Suppose the decision boundary  $w_1^T x = w_{11}$  is determined based on a soft-margin SVM, write down the corresponding primal optimization problem. (7 points)

**Solution:**

- 1) The decision tree is



(6')

- 2) Hard margin: We use the hard-margin SVM, when the data is linearly separable. It does not allow misclassifications. (3.5')

Soft margin: We use the soft-margin SVM in these two cases. (a) The separating hyperplane does not exist, because the data is not linearly separable or it has a high noise level which causes a large overlap of the

classes. (b) The separating hyperplane exists, but it is not the best solution when there exist outliers in the data. (3.5')

- 3) Define  $r^{(\ell)} = -1$  for  $y^{(\ell)} \in \{C_1, C_2, C_3\}$ , and  $r^{(\ell)} = 1$  for  $y^{(\ell)} \in \{C_4, C_5, C_6\}$ . (3') The primal optimization problem is

$$\begin{aligned} & \underset{\mathbf{w}_1, w_{11}, \xi_\ell}{\text{minimize}} && \frac{1}{2} \|\mathbf{w}_1\|^2 + \eta \sum_{\ell} \xi_\ell \\ & \text{subject to} && r^{(\ell)} \left( \mathbf{w}_1^T \mathbf{x}^{(\ell)} + w_{11} \right) \geq 1 - \xi_\ell, \quad \forall \ell \end{aligned}$$

(A correct structure of soft-margin SVM. (1') Specification of  $\mathbf{w}_1$  and  $w_{11}$  in the optimization problem. (3'))

**Problem 3. (20 points)** (Multilayer Perceptron (MLP) and CNN)

- 1) A single output MLP is considered with its output given by

$$y^{(\ell)} = \sum_{h=1}^H v_h z_h^{(\ell)} + v_0,$$

where

$$z_h^{(\ell)} = \text{sigmoid}(\mathbf{w}_h^T \mathbf{x}^{(\ell)}) = \text{sigmoid}\left(\sum_j w_{hj} x_j^{(\ell)}\right),$$

with  $\text{sigmoid}(a) = \frac{1}{1+\exp(-a)}$ . Given sample  $\{\mathbf{x}^{(\ell)}, r^{(\ell)}\}_\ell$ , the loss function in learning is defined as

$$\mathcal{L} = \frac{1}{2} \sum_{\ell} (r^{(\ell)} - y^{(\ell)})^2.$$

Compute  $\frac{\partial \mathcal{L}}{\partial w_{hj}}$  and explain why the learning algorithm is called “back-propagation”. (10 points)

- 2) Given a 2-dimensional convolution layer in a CNN, the input matrix  $\mathbf{X}$  and the convolution kernel  $\mathbf{W}$  are defined by

$$\mathbf{X} = \begin{bmatrix} x_{11} & x_{12} & x_{13} \\ x_{21} & x_{22} & x_{23} \\ x_{31} & x_{32} & x_{33} \end{bmatrix},$$

and

$$\mathbf{W} = \begin{bmatrix} w_{11} & w_{12} \\ w_{21} & w_{22} \end{bmatrix}.$$

- a) Write down the output of the convolution

$$\mathbf{Y} = \mathbf{X} * \mathbf{W}$$

with no kernel flipping (you do not need to flip the kernel in computation) and no zero padding. (5 points)

- b) Discuss the properties of the convolution operation in comparison with the “fully-connected” weights in the conventional neural networks like the MLP discussed in 1). (5 points)

**Solution:**

- 1) Observed that

$$\frac{\partial \mathcal{L}}{\partial w_{hj}} = \sum_{\ell} \frac{\partial \mathcal{L}^{(\ell)}}{\partial y^{(\ell)}} \frac{\partial y^{(\ell)}}{\partial z_h^{(\ell)}} \frac{\partial z_h^{(\ell)}}{\partial w_{hj}} \quad (1')$$

$$= - \sum_{\ell} \left( r^{(\ell)} - y^{(\ell)} \right) \times v_h \times z_h^{(\ell)} \left( 1 - z_h^{(\ell)} \right) x_j^{(\ell)} = - \sum_{\ell} \left( r^{(\ell)} - y^{(\ell)} \right) v_h z_h^{(\ell)} \left( 1 - z_h^{(\ell)} \right) x_j^{(\ell)}. \quad (4')$$

The  $(r^{(\ell)} - y^{(\ell)})v_h$  acts like the error term for hidden unit  $h$ .  $(r^{(\ell)} - y^{(\ell)})$  is the error in the output which is backpropagated from the output to the hidden unit weighted by the “responsibility” of the hidden unit as given by its weight  $v_h$ . (5')

- 2) (a)  $(\mathbf{y})_{mn} = (\mathbf{X} * \mathbf{W})_{mn} = \sum_{i=-1}^1 \sum_{j=-1}^1 x_{i+m, j+n} w_{ij}$ . (5') (b) Your answer should cover the following two points: sparse connectivity (2.5') and parameter sharing (2.5').

**Problem 4. (15 + 5 points)** (*Parameter Estimation, Clustering, and Nonparametric Methods*)

Given sample  $\mathcal{X} = (\mathbf{x}^{(1)}, \dots, \mathbf{x}^{(N)})$  generated from some unknown distribution  $p(\mathbf{x})$ .

- 1) Suppose  $p(\mathbf{x})$  is a multivariate Gaussian distribution model  $p(\mathbf{x}) = \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}, \boldsymbol{\Sigma})$ , derive the estimates of parameters  $\boldsymbol{\mu}$  and  $\boldsymbol{\Sigma}$  based on MLE. (5 points)
- 2) Suppose  $p(\mathbf{x})$  is a multivariate Gaussian mixture model

$$p(\mathbf{x}) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k), \quad \sum_{k=1}^K \pi_k = 1, \quad \pi_k \geq 0, \quad \forall k = 1, \dots, K,$$

with parameters  $\pi_k$ ,  $\boldsymbol{\mu}_k$ , and  $\boldsymbol{\Sigma}_k$ ,

- a) discuss the advantages of this model over the multivariate Gaussian distribution model in 1) and why it can be used for clustering; (5 points)
- b) (**Bonus Question**) describe the idea of expectation-maximization (EM) algorithm for MLE of the parameters  $\pi_k$ ,  $\boldsymbol{\mu}_k$ , and  $\boldsymbol{\Sigma}_k$ , and how the MLE results derived in 1) can be used in the EM algorithm. (5 points)

(Hint: You are not required to write down the steps of the EM algorithm.)

- 3) Suppose dimensions of  $\mathbf{x}$  are independent from each other, provide an approach to estimating  $p(\mathbf{x})$  based on univariate nonparametric density estimation method. (5 points)

**Solution:**

- 1) Suppose  $\mathbf{x}$  is a  $M$ -dimensional vector, then the log-likelihood is

$$\begin{aligned} \mathcal{L} &= \log \prod_{n=1}^N \mathcal{N}(\mathbf{x}^{(n)} | \boldsymbol{\mu}, \boldsymbol{\Sigma}) \\ &= \sum_{n=1}^N \log \frac{1}{\sqrt{(2\pi)^M |\boldsymbol{\Sigma}|}} e^{-\frac{1}{2}(\mathbf{x}^{(n)} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x}^{(n)} - \boldsymbol{\mu})} \\ &= -\frac{1}{2} \sum_{n=1}^N \left( \log(2\pi)^M + \log |\boldsymbol{\Sigma}| \right) - \frac{1}{2} \sum_{n=1}^N \left( \mathbf{x}^{(n)} - \boldsymbol{\mu} \right)^T \boldsymbol{\Sigma}^{-1} \left( \mathbf{x}^{(n)} - \boldsymbol{\mu} \right) \\ &= \frac{1}{2} \sum_{n=1}^N \log |\boldsymbol{\Sigma}^{-1}| - \frac{1}{2} \sum_{n=1}^N \log (2\pi)^M - \frac{1}{2} \sum_{n=1}^N \text{tr} \left( \left( \mathbf{x}^{(n)} - \boldsymbol{\mu} \right) \left( \mathbf{x}^{(n)} - \boldsymbol{\mu} \right)^T \boldsymbol{\Sigma}^{-1} \right). \quad (1') \end{aligned}$$

By setting  $\frac{\partial \mathcal{L}}{\partial \boldsymbol{\mu}} = \mathbf{0}$ , we have

$$\frac{\partial \mathcal{L}}{\partial \boldsymbol{\mu}} = \sum_{n=1}^N \boldsymbol{\Sigma}^{-1} \left( \mathbf{x}^{(n)} - \boldsymbol{\mu} \right) = \mathbf{0}, \quad (1')$$

leading to the MLE estimator

$$\hat{\boldsymbol{\mu}} = \frac{1}{N} \sum_{n=1}^N \mathbf{x}^{(n)}. \quad (1')$$

By setting  $\frac{\partial \mathcal{L}}{\partial \boldsymbol{\Sigma}^{-1}} = \mathbf{0}$ , we have

$$\frac{\partial \mathcal{L}}{\partial \boldsymbol{\Sigma}^{-1}} = \frac{1}{2} \sum_{n=1}^N \boldsymbol{\Sigma} - \frac{1}{2} \sum_{n=1}^N \left( \mathbf{x}^{(n)} - \boldsymbol{\mu} \right) \left( \mathbf{x}^{(n)} - \boldsymbol{\mu} \right)^T = \mathbf{0}, \quad (1')$$

leading to the MLE estimator

$$\hat{\Sigma} = \frac{1}{N} \sum_{n=1}^N \left( \mathbf{x}^{(n)} - \boldsymbol{\mu} \right) \left( \mathbf{x}^{(n)} - \boldsymbol{\mu} \right)^T. \quad (1')$$

- 2) (a) The multivariate Gaussian distribution model in 1) can only approximate Gaussian distributed data, while the multivariate Gaussian mixture model has the ability to approximate many naturally occurring real-world data thanks to the law of large numbers. (3') The clustering ability is naturally obtained for mixture model as it classified the data into  $K$  components, i.e., clusters. (2')
- (b) The EM algorithm is an iterative method to find (local) maximum likelihood or maximum a posteriori (MAP) estimates of parameters in statistical models, where the model depends on unobserved latent variables. The EM iteration alternates between performing an expectation (E) step, which creates a function for the expectation of the log-likelihood evaluated using the current estimate for the parameters, and a maximization (M) step, which computes parameters maximizing the expected log-likelihood found on the E step. These parameter-estimates are then used to determine the distribution of the latent variables in the next E step. The results in 1) can be used in the M step. (5')
- 3) Histogram estimator: First divide input space into equal-sized bins:  $[x_0 + mh, x_0 + (m + 1)h]$  with  $x_0$  being the origin,  $h$  being the bin width, and  $m$  being an integer. Then the histogram estimator gives

$$\hat{p}(x) = \prod_{m=1}^M \frac{\# \left\{ x_m^{(n)} \text{ in the same bin as } x \right\}}{Nh}. \quad (5')$$

Other estimators like naive estimator, kernel estimator, and  $k$ -nearest neighbor estimator are also fine.

**Problem 5. (15 + 5 points)** (*Dimension Reduction and Matrix Factorization*)

Let  $\mathbf{X}$  be a centered (i.e., zero-mean) sample matrix where each row is one observation and  $\Sigma$  be the sample covariance matrix. Assuming unit vector  $\mathbf{w}_1$  is the eigenvector of  $\Sigma$  corresponding to the largest eigenvalue.

- 1) Prove that  $\mathbf{w}_1$  is the principal component in principal component analysis (PCA) in that the projection onto direction  $\mathbf{w}_1$  leads to the maximum variance for  $\mathbf{X}$ , i.e.,  $\mathbf{w}_1$  maximizes  $\mathbf{w}^T \Sigma \mathbf{w}$  for any  $\mathbf{w}$  satisfying  $\|\mathbf{w}\|_2 = 1$ . (10 points)
- 2) Suppose the sample  $\mathbf{X}$  is labeled into two classes, briefly describe the idea of linear discriminant analysis (LDA) and discuss the similarities and differences between PCA and LDA. (5 points)  
(Hint: You are not required to write down the detailed derivations of LDA.)
- 3) (**Bonus Question**) Show that  $\mathbf{w}_1$  minimizes the reconstruction error  $\|\mathbf{X} - \mathbf{X}\mathbf{w}\mathbf{w}^T\|_F^2$  for any  $\mathbf{w}$  satisfying  $\|\mathbf{w}\|_2 = 1$ . (5 points)  
(Hint: You can directly use the result in 1).)

**Solution:**

- 1) The variance maximization problem is

$$\begin{aligned} & \underset{\mathbf{w}}{\text{maximize}} && \mathbf{w}^T \Sigma \mathbf{w} \\ & \text{subject to} && \|\mathbf{w}\| = 1. \end{aligned}$$

By setting the Lagrangian's derivative to zero as follows:

$$\frac{\partial}{\partial \mathbf{w}} - \mathbf{w}^T \Sigma \mathbf{w} + \lambda (\mathbf{w}^T \mathbf{w} - 1) = 0,$$

where  $\lambda$  is the Lagrange multiplier, leading to

$$\Sigma \mathbf{w} = \lambda \mathbf{w}. \quad (5')$$

Observe that

$$\mathbf{w}^T \Sigma \mathbf{w} = \mathbf{w}^T \lambda \mathbf{w} = \lambda,$$

as  $\mathbf{w}_1$  corresponds to the largest  $\lambda$ , we can conclude that  $\mathbf{w}_1$  maximizes the variance  $\mathbf{w}^T \Sigma \mathbf{w}$ . (5')

- 2) LDA finds the vector  $\mathbf{w}$  on which the data are projected such that the examples from the two classes are as well separated as possible. (2')

Similarity of LDA and PCA: both are linear transformation techniques for dimension reduction. (1')

Differences of LDA and PCA: LDA is a supervised technique that attempts to find a feature subspace that maximizes class separability, while PCA is unsupervised that focuses on capturing the direction of maximum variation in the data set. (2')

- 3) The reconstruction error minimization problem is

$$\begin{aligned} & \underset{\mathbf{w}}{\text{maximize}} && f(\mathbf{w}) = \|\mathbf{X} - \mathbf{X}\mathbf{w}\mathbf{w}^T\|_F^2 \\ & \text{subject to} && \|\mathbf{w}\| = 1. \end{aligned}$$



Observe that

$$\begin{aligned}
 f(\mathbf{w}) &= \|\mathbf{X} - \mathbf{X}\mathbf{w}\mathbf{w}^T\|_F^2 \\
 &= \text{tr} \left( (\mathbf{X} - \mathbf{X}\mathbf{w}\mathbf{w}^T)^T (\mathbf{X} - \mathbf{X}\mathbf{w}\mathbf{w}^T) \right) \\
 &= \text{tr} \left( \mathbf{X}^T \mathbf{X} - \mathbf{X}^T \mathbf{X} \mathbf{w} \mathbf{w}^T - \mathbf{w}_1 \mathbf{w}^T \mathbf{X}^T \mathbf{X} + \mathbf{w} \mathbf{w}^T \mathbf{X}^T \mathbf{X} \mathbf{w} \mathbf{w}^T \right) \\
 &= \text{tr} \left( \mathbf{X}^T \mathbf{X} + \mathbf{w} \mathbf{w}^T \mathbf{w} \mathbf{w}^T \mathbf{X}^T \mathbf{X} \right) - 2 \text{tr} \left( \mathbf{w} \mathbf{w}^T \mathbf{X}^T \mathbf{X} \right) \\
 &= \text{tr} \left( \mathbf{X}^T \mathbf{X} \right) - \text{tr} \left( \mathbf{w}^T \boldsymbol{\Sigma} \mathbf{w} \right).
 \end{aligned}$$

Therefore, it is obvious that the reconstruction error minimization problem is equivalent to the variance maximization problem, through which the proof is completed. (5')

**Problem 6. (10 + 10 points)** (*Ensemble Learning and Model Selection*)

- 1) In ensemble learning, suppose there are  $L$  independent two-class classifiers used for simple voting and the output of classifier  $j$  ( $j = 1, \dots, L$ ) is denoted as  $d_j$ . From the point of view that the mean squared error of an estimator can be decomposed into the bias part and the variance part, explain why increasing  $L$  can lead to the increase of the classification accuracy. (10 points)
- 2) (**Bonus Question**) Suppose we carry out a  $K$ -fold cross-validation on a dataset and obtain the classification error rates  $\{p_i\}_{i=1}^K$ , describe the steps of a one-sided  $t$  test on testing the null hypothesis  $H_0$  that the classifier has error percentage  $p_0$  or less at a significance level  $\alpha$ . (10 points)

**Solution:**

- 1) Let  $\mathbb{E}[d_j]$  and  $\text{Var}(d_j)$  are the expected value and variance of  $d_j$  for classifier  $j$ , where  $d_j$  is the output of the  $j$ th classifier. Expected value and variance of output for independent classifiers:

$$\mathbb{E}[y] = \mathbb{E}\left[\sum_j \frac{1}{L} d_j\right] \geq \frac{1}{L} L \min_j \{\mathbb{E}[d_j]\} = \min_j \{\mathbb{E}[d_j]\} \quad (2')$$

$$\text{Var}(y) = \text{Var}\left(\sum_j \frac{1}{L} d_j\right) = \frac{1}{L^2} \text{Var}\left(\sum_j d_j\right) \leq \frac{1}{L^2} L \max_j \{\text{Var}(d_j)\} = \frac{1}{L} \max_j \{\text{Var}(d_j)\} \quad (2')$$

As  $L$  increases, the expected value (and hence the bias) does not change (3') but the variance decreases (3'), and hence the mean squared error of the estimator  $y$  decreases, leading to an increase in accuracy.

- 2) Let

$$m = \frac{\sum_{i=1}^K p_i}{K} \quad (2'), \quad S^2 = \frac{\sum_{i=1}^K (p_i - m)^2}{K - 1} \quad (2'),$$

we have

$$\sqrt{K} \frac{(m - p_0)}{S} \sim \tau_{K-1} \quad (2')$$

and hence we will fail to reject at significance level  $\alpha$  if

$$\sqrt{K} \frac{m - p_0}{S} \in (-\infty, t_{\alpha, K-1}). \quad (4')$$