

# Inequality Extensions

## SI252 Reinforcement Learning

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# Cauchy-Schwarz

## Theorem

*For any r.v.s  $X$  and  $Y$  with finite variances,*

$$|\mathbb{E}(XY)| \leq \sqrt{\mathbb{E}(X^2)\mathbb{E}(Y^2)}.$$

# Example: Second Moment Method

Let  $X$  be a nonnegative random variable, then

$X \geq 0$

r.v.

$$\Pr(X = 0) \leq \frac{\text{Var}(X)}{\mathbb{E}(X^2)}.$$

Examples:

①  $X$ : number of questions that Fred gets wrong on an exam

$P(X=0) = P\{\text{Fred gets a perfect score}\}$

②  $X$ : number of pairs of people at a party with the same birthday

$P(X=0) = P\{\text{No birthday matches}\}$

Proof: Since  $X \geq 0$

$$\underline{X = X I\{X > 0\}} \quad \vee \quad I\{X > 0\} : \text{indicator of event } X > 0$$

$$\text{If } X > 0 : \quad X I\{X > 0\} = X \cdot 1 = X$$

$$\text{If } X = 0 : \quad X I\{X > 0\} = X \cdot 0 = 0 \quad I\{\cdot\}^2 = I\{\cdot\}$$

By Cauchy-Schwarz inequality,

$$E[I\{A\}] = P(A)$$

$$\begin{aligned} E(X) &= E(X I\{X > 0\}) \leq \sqrt{E(X^2) \cdot E[I\{X > 0\}^2]} \\ &= \sqrt{E(X^2) \cdot E[I\{X > 0\}]} \\ &= \sqrt{E(X^2) \cdot P(X > 0)} \end{aligned}$$

$$E^2(X) \leq E(X^2) \cdot P(X > 0) \Rightarrow P(X > 0) \geq \frac{E^2(X)}{E(X^2)}$$

$$\underline{P(X=0) = 1 - P(X > 0)} \leq 1 - \frac{E^2(X)}{E(X^2)} = \frac{E(X^2) - E^2(X)}{E(X^2)} = \frac{\text{Var}(X)}{E(X^2)}$$

# Example: Application of Second Moment Method

Covariance : rvs  $X$  and  $Y$

$$\text{Cov}(X, Y) = E[(X - EX)(Y - EY)] = E(XY) - E(X)E(Y)$$

Correlation : rvs  $X$  and  $Y$

$$\text{Corr}(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)\text{Var}(Y)}}$$

$$E(I_j) = \Pr(I_j=1)E(0,1)$$

Assume  $X = I_1 + \dots + I_n$ , where the  $I_j$  are uncorrelated indicator r.v.s. Let  $p_j = \mathbb{E}(I_j)$ . Upper bound of  $\Pr(X = 0)$ ?

Uncorrelated :  $X, Y$  uncorrelated  $\Leftrightarrow \text{Cov}(X, Y) = 0$  or  $\text{Corr}(X, Y) = 0$ .

Property : If rvs.  $X$  and  $Y$  uncorrelated, then  $\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y)$

Proof :  $\text{Cov}(X+Y, Z) = E[(X+Y)Z] - E(X+Y)E(Z)$

$$\begin{aligned} &= E(XZ) + E(YZ) - (E(X)E(Z) + E(Y)E(Z)) \\ &= \text{Cov}(X, Z) + \text{Cov}(Y, Z) \end{aligned}$$

$$\begin{aligned}\text{Var}(X+Y) &= E[(X+Y)^2] - E^2(X+Y) \\ &= \text{Cov}(X+Y, X+Y)\end{aligned}$$

$$\underline{\text{Cov}(X, X) = \text{Var}(X)}$$

$$\begin{aligned}&= \text{Cov}(X, X) + \text{Cov}(X, Y) + \text{Cov}(Y, X) + \text{Cov}(Y, Y) \\ &= \text{Var}(X) + 0 + 0 + \text{Var}(Y) = \text{Var}(X) + \text{Var}(Y)\end{aligned}$$

$$P(X=0) \leq \frac{\text{Var}(X)}{E(X^2)}$$

$$\begin{aligned}P_j \in (0, 1) \quad \mu &= \sum_{j=1}^n P_j > \sum_{j=1}^n P_j^2 = c \\ \mu - c > 0 \quad \frac{\mu^2}{\mu - c} &> \frac{\mu(\mu - c)}{\mu - c} = \mu\end{aligned}$$

Since  $\{I_j\}$  are uncorrelated, we have

$$\begin{aligned}\text{Var}(X) &= \text{Var}\left(\sum_{j=1}^n I_j\right) = \sum_{j=1}^n \text{Var}(I_j) = \sum_{j=1}^n [E(I_j^2) - E^2(I_j)] \\ &= \sum_{j=1}^n [E(I_j) - E^2(I_j)] = \sum_{j=1}^n (P_j - P_j^2) = \mu - c\end{aligned}$$

$$\mu \triangleq E(X) = \sum_{j=1}^n E(I_j) = \sum_{j=1}^n P_j \quad c \triangleq \sum_{j=1}^n P_j^2$$

$$E(X^2) = \text{Var}(X) + (EX)^2 = \mu - c + \mu^2 \quad P(X=0) \leq \frac{\mu - c}{\mu - c + \mu^2} = \frac{1}{1 + \frac{\mu^2}{\mu - c}}$$

# Example: Application of Second Moment Method

$I_{ij} \triangleq I \{ \text{person } i \text{ and } j \text{ have near birthdays} \}$

$X \triangleq$  number of "near birthday" pairs

$$X = \sum_{i \neq j} I_{ij}$$

$d_i$ :  $i$ 's birthday

$$P(X=0) \leq \frac{1}{1+E(X)}$$

$$E(X) = E\left(\sum_{i \neq j} I_{ij}\right) = \sum_{i \neq j} E(I_{ij}) = \sum_{i \neq j} \frac{3}{365} = \binom{14}{2} \frac{3}{365}$$

Suppose there are 14 people in a room. How likely is it that there are two people with the same birthday or birthdays one day apart?

$$E(I_{ij}) = P(I_{ij}=1) = P(i \text{ and } j \text{ have near birthdays})$$

$$= \sum_{t=1}^{365} P(d_i \in \{(t-1), t, (t+1)\} \% 365 \mid d_j = t) \underbrace{P(d_j = t)}$$

$$= \sum_{t=1}^{365} \frac{3}{365} \times \frac{1}{365} = \frac{3}{365}$$

$$P(X=0) \leq \frac{1}{1+E(X)} < 0.573.$$

# Markov's Inequality & Chebyshev's Inequality

## Theorem (Markov's Inequality)

$\in [0,1]$

For any r.v.  $X$  and constant  $a > 0$ ,

$$\bar{P}(|X| \geq a) \leq \frac{E|X|}{a} \leq 1$$

## Theorem (Chebyshev's Inequality)

Let  $X$  have mean  $\mu$  and variance  $\sigma^2$ . Then for any  $a > 0$ ,

$$P(|X - \mu| \geq a) \leq \frac{\sigma^2}{a^2}.$$



# Example: Coin Flipping

$$n \geq 4$$

① Using Chebyshev's inequality:

$X_i \triangleq I\{\text{the } i\text{-th coin flip is head}\}$   $X \triangleq \text{number of heads}$

$$X = X_1 + \dots + X_n \quad X_i \sim \text{Bern}\left(\frac{1}{2}\right) \quad E(X_i) = \frac{1}{2} \quad \text{Var}(X_i) = \frac{1}{4}$$

$$E(X) = nE(X_i) = \frac{n}{2} \quad \text{Var}(X) = n\text{Var}(X_i) = \frac{n}{4}$$

Find bounds on the probability of having no more than  $n/4$  heads or fewer than  $3n/4$  heads in a sequence of  $n$  fair coin flips.

$$P\left(X \leq \frac{n}{4} \text{ or } X \geq \frac{3n}{4}\right) = P\left(|X - \frac{n}{2}| \geq \frac{n}{4}\right) = P\left(|X - E(X)| \geq \frac{n}{4}\right) \leq \frac{\text{Var}(X)}{\left(\frac{n}{4}\right)^2} = \frac{\frac{n}{4}}{n} = \frac{1}{4}$$

② Using Markov's inequality:

$$P\left(X \geq \frac{3n}{4}\right) \leq \frac{E(X)}{\frac{3n}{4}} = \frac{n/2}{3n/4} = \frac{2}{3}$$

$$\begin{array}{l} X \sim \text{Bin}\left(n, \frac{1}{2}\right) \\ n - X \sim \text{Bin}\left(n, \frac{1}{2}\right) \end{array}$$

$$P\left(X \leq \frac{n}{4}\right) = P\left(n - X \geq \frac{3n}{4}\right) = P\left(X \geq \frac{3n}{4}\right) \leq \frac{2}{3}$$

$$P\left(X \geq \frac{3n}{4} \text{ or } X \leq \frac{n}{4}\right) \leq \frac{2}{3} + \frac{2}{3} = \frac{4}{3} > 1$$





# Chernoff's Technique

## Theorem

For any r.v.  $X$  and constant  $a$ ,

$$P(X \geq a) \leq \inf_{t>0} \frac{E(e^{tX})}{e^{ta}},$$

$$P(X \leq a) \leq \inf_{t<0} \frac{E(e^{tX})}{e^{ta}}.$$

# Example: Sum of Independent Bernoulli R.V.s

Let  $X_1, \dots, X_n$  be independent Bernoulli random variables such that  $\Pr(X_i = 1) = p_i$ ,  $\Pr(X_i = 0) = 1 - p_i$ . Let  ~~$X = n - 1 - X_i$~~  and  $\mu = \mathbb{E}(X)$ . Then for  $0 < \delta < 1$ ,

$$X = \sum_{i=1}^n X_i$$

$$\Pr(|X - \mu| \geq \delta \mu) \leq 2e^{-\mu \delta^2 / 3}.$$

Moment Generating Function (MGF):

r.v.  $X$   $M(t) = E(e^{tx})$   $t$  open interval  $(-a, a)$

MGF of Bernoulli: r.v.  $X \sim \text{Bern}(p)$ :  $M(t) = E(e^{tx}) = pe^t + 1-p$

MGF of a sum of independent r.v.s:

$X, Y$  independent,  $M_{X+Y}(t) = M_X(t) \cdot M_Y(t)$

$$M_{X+Y}(t) = E[e^{t(X+Y)}] = E[e^{tx} \cdot e^{ty}] = \underbrace{E(e^{tx})}_{M_X(t)} \cdot \underbrace{E(e^{ty})}_{M_Y(t)}$$

$$\begin{aligned} P(|X-\mu| \geq \delta\mu) &= P(X-\mu \geq \delta\mu) + P(X-\mu \leq -\delta\mu) \\ &= \underbrace{P(X \geq (1+\delta)\mu)}_{\textcircled{1}} + \underbrace{P(X \leq (1-\delta)\mu)}_{\textcircled{2}} \end{aligned}$$

①  $P(X \geq (1+\delta)\mu)$  by Chernoff's technique,

$$P(X \geq (1+\delta)\mu) \leq \frac{E(e^{tX})}{e^{t(1+\delta)\mu}} = \frac{M_X(t)}{e^{t(1+\delta)\mu}} \quad \forall t > 0$$

Since  $X = \sum_{i=1}^n X_i$ ,  $\{X_i\}$  independent  $\Rightarrow M_X(t) = \prod_{i=1}^n M_{X_i}(t)$

$$\forall t \in \mathbb{R}, \quad \underline{M_{X_i}(t) = E(e^{tX_i}) = p_i e^t + (1-p_i) = 1 + p_i(e^t - 1) \leq e^{p_i(e^t - 1)}} \quad \underline{1+x \leq e^x \quad \forall x \in \mathbb{R}}$$

$$\begin{aligned} M_X(t) &\leq \prod_{i=1}^n e^{p_i(e^t - 1)} \\ &= e^{(e^t - 1) \cdot \sum_{i=1}^n p_i} \end{aligned}$$

$$\begin{aligned} \text{Since } \mu &\triangleq E(X) = E\left(\sum_{i=1}^n X_i\right) \\ &= \sum_{i=1}^n E(X_i) = \sum_{i=1}^n p_i \end{aligned}$$

Proof:

(if  $x > 0$ ) (if  $x < 0$ )

① Taylor's theorem:  $\exists \theta \in (0, x)$  or  $(x, 0)$  s.t.  
 $e^x = 1 + x + \frac{1}{2} f''(\theta) x^2 = 1 + x + \frac{1}{2} e^\theta x^2 \geq 1 + x.$

② Convexity:  $f(x) = e^x - x - 1 \quad \forall x \in \mathbb{R}$   
 $f(x) \geq f(0) = 0 \quad f(x) \geq 0 \quad \forall x \in \mathbb{R}$

$$\Rightarrow \underline{M_X(t) \leq e^{(e^t - 1) \cdot \mu}}$$

$$\text{Thus, } P(X \geq (1+\delta)\mu) \leq \frac{N_X(t)}{e^{t(1+\delta)\mu}} = \frac{e^{(e^t-1)\mu}}{e^{t(1+\delta)\mu}} \quad \forall t > 0$$

$$\text{Let } t = \ln(1+\delta) \quad \forall \delta \in (0, 1) \text{ we have } t > 0$$

$$\begin{aligned} e^t &\rightarrow 1+\delta \\ \Rightarrow \frac{e^\delta}{(1+\delta)^{(1+\delta)}} &\leq e^{-\delta^2/3} \quad \forall \delta \in (0, 1) \end{aligned}$$

$$\Rightarrow \delta - (1+\delta) \ln(1+\delta) \leq -\delta^2/3 \quad \forall \delta \in (0, 1)$$

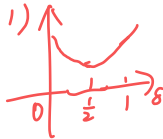
$$\Rightarrow f(\delta) = \delta - (1+\delta) \ln(1+\delta) + \delta^2/3 \leq 0 \quad \forall \delta \in (0, 1)$$

$$f'(\delta) = -\ln(1+\delta) + \frac{2}{3}\delta \quad f''(\delta) = -\frac{1}{1+\delta} + \frac{2}{3}$$

$$f''(\delta) < 0 \quad \forall \delta \in (0, \frac{1}{2}) \quad f''(\delta) > 0 \quad \forall \delta \in (\frac{1}{2}, 1)$$

$$f'(0) = 0 \quad f'(1) < 0 \quad f'(\delta) \leq 0 \quad \forall \delta \in (0, 1)$$

$$f(\delta) \leq f(0) = 0 \quad \forall \delta \in (0, 1)$$







## Example: Revisit Example of Coin Flipping

$$X_i \triangleq \mathbb{I} \{ i\text{-th coin flip is head} \} \quad X_i \sim \text{Bern}(\frac{1}{2}) \quad E(X_i) = \frac{1}{2}$$

$$X = X_1 + \dots + X_n \quad \mu = E(X) = nE(X_i) = \frac{n}{2}$$

$$P(X \leq \frac{n}{4} \text{ or } X \geq \frac{3n}{4}) = P(|X - \frac{n}{2}| \geq \frac{n}{4}) = P(|X - \mu| \geq \delta \mu)$$

$\delta = \frac{1}{2}$

Find bounds on the probability of having no more than  $n/4$  heads or fewer than  $3n/4$  heads in a sequence of  $n$  fair coin flips.

Chebyshev :  $\left(\frac{4}{n}\right)$

Markov :  $\frac{4}{3} > 1$

$$\leq 2e^{-\mu \delta^2 / 3} = 2e^{-\mu / 12}$$

$$= 2e^{-\frac{n}{24}}$$

$$n \geq 4$$

$$n = 100$$



