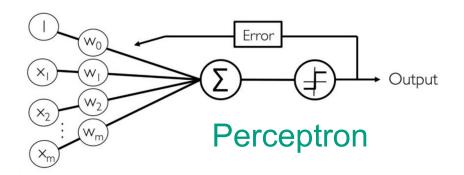
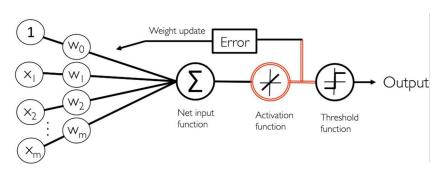
Quadratic loss/cost function



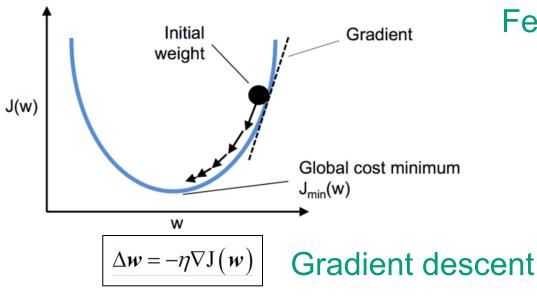
Recap

$$J(\mathbf{w}) = \frac{1}{2} \sum_{i} \left(y^{(i)} - \phi(z^{(i)}) \right)^{2}$$

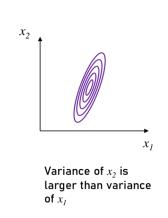


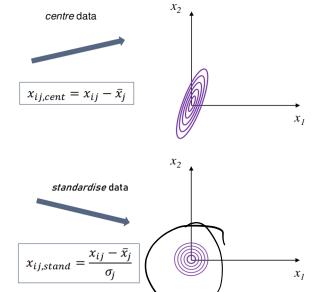


Adaptive linear neuron (Adaline)



Feature scaling







Linear regression & logistic regression



Basic notation (repetition)



Vectors and Matrices

We often represent raw (numeric) data as vectors and matrices

Example: Iris data can be represented as a 150 by 4 matrix: $X \in \mathbb{R}^{150x4}$

- Superscript means i-th training sample
- Subscript means j-th feature (dimension)
- Lowercase boldface \rightarrow vectors ($x \in \mathbb{R}^{nx_1}$)
- Uppercase boldface \rightarrow matrices ($X \in \mathbb{R}^{mxn}$)
- Single element in a vector $x^{(i)}$
- Single element in a matrix $x_j^{(i)}$

$$\boldsymbol{X} = \begin{bmatrix} x_1^{(1)} & x_2^{(1)} & x_3^{(1)} & x_4^{(1)} \\ x_1^{(2)} & x_2^{(2)} & x_3^{(2)} & x_4^{(2)} \\ \vdots & \vdots & \vdots & \vdots \\ x_1^{(150)} & x_2^{(150)} & x_3^{(150)} & x_4^{(150)} \end{bmatrix}$$



Vectors and Matrices

Row vectors (e.g. one row in *X*, "flower" sample in Iris data set)

$$\mathbf{x}^{(i)} = \begin{bmatrix} x_1^{(i)} & x_2^{(i)} & x_3^{(i)} & x_4^{(i)} \end{bmatrix} \qquad \mathbf{x}^i \in \mathbb{R}^{1\times 4}$$

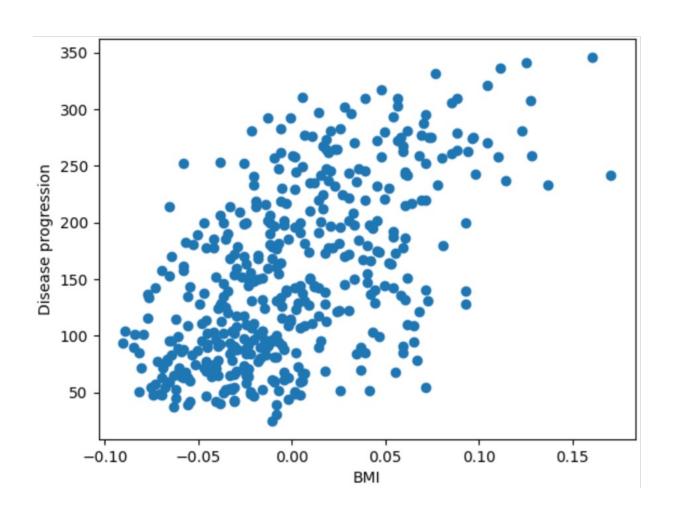
Column vectors (e.g. one column in *X*, one feature)

$$\boldsymbol{x}_{j} = \begin{bmatrix} x_{j}^{(1)} \\ x_{j}^{(2)} \\ \vdots \\ x_{j}^{(150)} \end{bmatrix} \qquad \boldsymbol{x}_{j} \in \mathbb{R}^{150x}$$

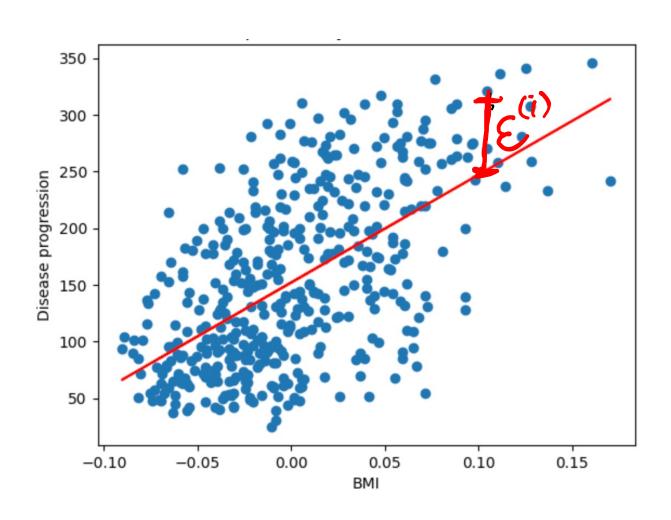


Simple linear regression (and least squares)









Model (parametric)

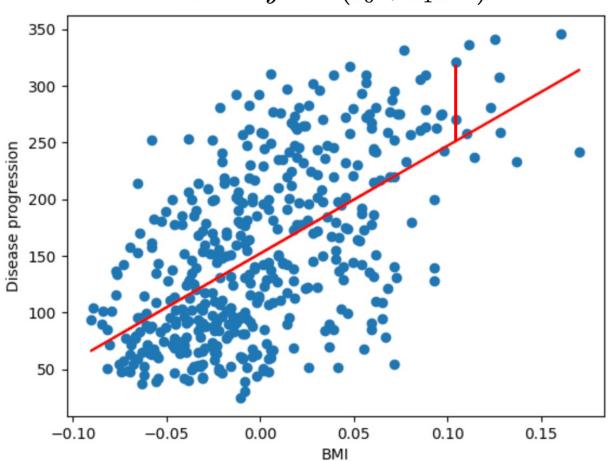
$$f(x) = \theta_0 + \theta_1 x \qquad \mathbf{y} = \mathbf{X} \boldsymbol{\theta}$$

Error

$$\varepsilon^{(i)} = y^{(i)} - (\theta_0 + \theta_1 x^{(i)})$$



$$\varepsilon^{(i)} = y^{(i)} - (\theta_0 + \theta_1 x^{(i)})$$



Generalization to m features

$$\mathbf{x}^{(i)} = y^{(i)} - \mathbf{x}^{(i)}\boldsymbol{\theta} \qquad \text{vector}$$

$$\mathbf{x}^{(i)} = \begin{bmatrix} 1, x_1 \end{bmatrix} \quad \boldsymbol{\theta} = \begin{bmatrix} \theta_0 \\ \theta_1 \end{bmatrix}$$

$$\mathbf{x}^{(i)}\boldsymbol{\theta} : \neq \sum_{j=0}^{i} \boldsymbol{\theta}_{i}^{(j)}\boldsymbol{\theta}_{j}$$

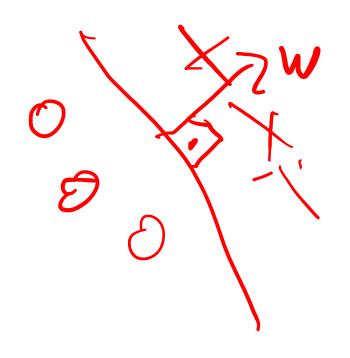


Alternative equivalent parameter representation

$$\mathbf{x}^{(i)}\boldsymbol{\theta} = b + \mathbf{x}^{(i)}\mathbf{w}$$

$$oldsymbol{ heta} = egin{bmatrix} heta_0 \ heta_1 \ dots \ heta_m \end{bmatrix} := egin{bmatrix} b \ \mathbf{w} \end{bmatrix}$$

$$oldsymbol{ heta} oldsymbol{ heta} = egin{bmatrix} heta_0 \ heta_1 \ dots \ heta_m \end{bmatrix} := egin{bmatrix} b \ extbf{w} \end{bmatrix} & oldsymbol{ extbf{w}} & oldsymbol{ extbf{w}} = egin{bmatrix} heta_1 \ heta_2 \ dots \ heta_m \end{bmatrix}, \quad b = heta_0$$



"weights"

"bias"

"normal"

"offset"

(hyperplanes)

Error in matrix notation

$$\xi = \begin{cases} \xi(1) \\ \xi(3) \end{cases}$$

$$0 = \begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}$$

$$\begin{bmatrix} 1 \\ 1 \\ 1 \end{bmatrix}$$

$$\begin{bmatrix} 1 \\ 1 \\ 1 \end{bmatrix}$$

$$\begin{bmatrix} 1 \\ 1 \\ 1 \end{bmatrix}$$



Goal: Minimize the error

We have one error for each sample, how to measure a global error?

Quadratic / Squared-error loss function

$$L = \epsilon^T \epsilon = ||\epsilon||_2^2 = \sum_{i=1}^n \varepsilon^{(i)^2}$$

$$||\xi||_2 = \sqrt{\xi_1^2 + \dots \xi_N^2}$$

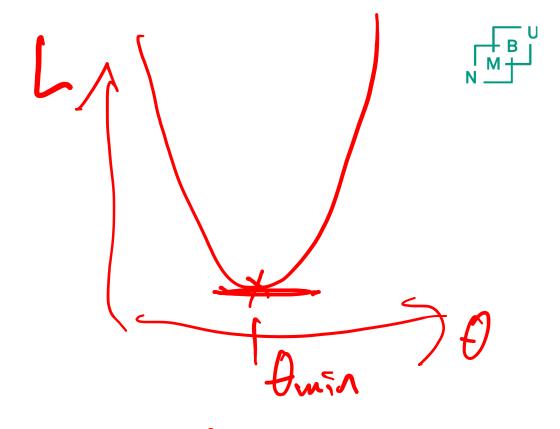
Goal: Minimize the sum of squared errors

$$L = \boldsymbol{\epsilon}^T \boldsymbol{\epsilon} = ||\boldsymbol{\epsilon}||_2^2 = \sum_{i=1}^n \varepsilon^{(i)^2}$$

Insert error

$$L(\boldsymbol{\theta}) = (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^T (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})$$

$$= \mathbf{y}^T \mathbf{y} - \mathbf{y}^T \mathbf{X}\boldsymbol{\theta} - \boldsymbol{\theta}^T \mathbf{X}^T \mathbf{y} + \boldsymbol{\theta}^T \mathbf{X}^T \mathbf{X}\boldsymbol{\theta}$$





Goal: Minimize the sum of squared errors

$$L(\boldsymbol{\theta}) = (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^T (\mathbf{y} - \mathbf{X}\boldsymbol{\theta}) = \mathbf{y}^T \mathbf{y} - \mathbf{y}^T \mathbf{X}\boldsymbol{\theta} - \boldsymbol{\theta}^T \mathbf{X}^T \mathbf{y} + \boldsymbol{\theta}^T \mathbf{X}^T \mathbf{X}\boldsymbol{\theta}$$

$$\frac{\partial L}{\partial \boldsymbol{\theta}} = \frac{\partial \left(\mathbf{y}^T \mathbf{y} - \mathbf{y}^T \mathbf{X} \boldsymbol{\theta} - \boldsymbol{\theta}^T \mathbf{X}^T \mathbf{y} + \boldsymbol{\theta}^T \mathbf{X}^T \mathbf{X} \boldsymbol{\theta} \right)}{\partial \boldsymbol{\theta}} = \mathbf{0}$$

$$= -2\mathbf{X}^T\mathbf{y} + 2\mathbf{X}^T\mathbf{X}\boldsymbol{\theta} = \mathbf{0}$$



$$\frac{\partial L}{\partial \boldsymbol{\theta}} = -2\mathbf{X}^T \mathbf{y} + 2\mathbf{X}^T \mathbf{X} \boldsymbol{\theta} = \mathbf{0}$$

$$\mathbf{X}^T \mathbf{X} \boldsymbol{\theta} = \mathbf{X}^T \mathbf{y}$$

(Normal equations)

$$\boldsymbol{\theta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

Least squares solution



Learning / Training / Fitting

$$\boldsymbol{\theta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

Prediction

$$\mathbf{y} = \mathbf{X}\boldsymbol{\theta}$$

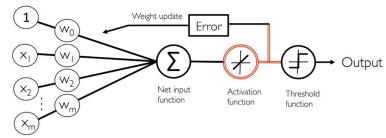
Remark: remember that we added a "1" feature to X

lin_regression.ipynb



Linear regression and Adaline?

→ Adaline is linear regression with a threshold



Adaptive Linear Neuron (Adaline)

$$L = \sum_{i=1}^{n} \varepsilon^{(i)^2} = \sum_{i=1}^{n} \left(y^{(i)} - \sum_{k=0}^{m} x_k^{(i)} \theta_k \right)^2 = \sum_{i=1}^{n} \left(y^{(i)} - \mathbf{x}^{(i)} \boldsymbol{\theta} \right)^2$$

Why is Adaline not always so good in practice?

sensitive to outliers

practice?

1: dead

0: a live

3: M



Linear regression again Probabilistic interpretation



Linear model

$$y^{(i)} = \mathbf{x}^{(i)}\boldsymbol{\theta} + \varepsilon^{(i)}$$



Error model (assumption)

$$\varepsilon^{(i)} \sim \mathcal{N}(0, \sigma^2)$$

$$p(\varepsilon^{(i)}) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{\varepsilon^{(i)^2}}{2\sigma^2}\right)$$

$$\varepsilon^{(i)}$$
 are i.i.d.

independent and identically distributed



Probability of a single outcome, given the sample, and parametrized by θ

$$P(y^{(i)} \mid x^{(i)}; \boldsymbol{\theta}) = \frac{1}{\sigma \sqrt{2\pi}} \exp\left(-\frac{\left(y^{(i)} - \mathbf{x}^{(i)} \boldsymbol{\theta}\right)^2}{2\sigma^2}\right)$$

Likelihood of θ (defined in terms of probability of the data)

$$\mathcal{L}(\boldsymbol{\theta}) = P(\mathbf{y} \mid \mathbf{X}; \boldsymbol{\theta})$$



Goal: We want to maximize the likelihood of our parameters

$$\mathcal{L}(\boldsymbol{\theta}) = P(\mathbf{y} \mid \mathbf{X}; \boldsymbol{\theta})$$

$$= \prod_{i=1}^{n} P(y^{(i)} \mid x^{(i)}; \boldsymbol{\theta}), \quad \text{(since } \varepsilon^{(i)} \text{ are i.i.d.)}$$

$$= \prod_{i=1}^{n} \frac{1}{\sigma \sqrt{2\pi}} \exp\left(-\frac{\left(y^{(i)} - \mathbf{x}^{(i)} \boldsymbol{\theta}\right)^{2}}{2\sigma^{2}}\right)$$

$$TT = 414243.4$$

$$i=11$$



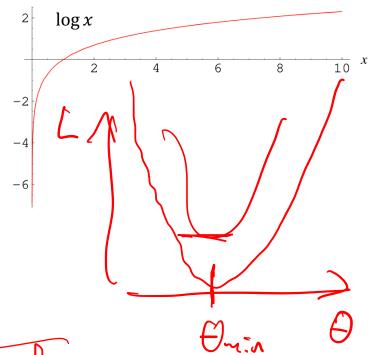
$$\mathcal{L}(oldsymbol{ heta}) = \prod_{i=1}^n P(y^{(i)} \mid x^{(i)}; oldsymbol{ heta})$$

In practice it's often beneficial to look at the log-likelihood

Since the natural logarithm (log) is a strictly monotone function, likelihood and log-likelihood attain maximum at the same θ

$$egin{aligned} \ell(oldsymbol{ heta}) &:= \log \mathcal{L}(oldsymbol{ heta}) = \log \prod_{i=1}^n P(y^{(i)} \mid x^{(i)}; oldsymbol{ heta}) \\ &= \log \prod_{i=1}^n rac{1}{\sigma \sqrt{2\pi}} \exp\left(-rac{(y^{(i)} - \mathbf{x}^{(i)} oldsymbol{ heta})^2}{2\sigma^2}
ight) \end{aligned}$$

$$= n \log \frac{1}{\sigma \sqrt{2\pi}} - \frac{1}{2\sigma^2} \sum_{i=1}^{n} \left(y^{(i)} - \mathbf{x}^{(i)} \boldsymbol{\theta} \right)^2$$





Maximizing a function is the same as minimizing the negative function

$$\ell(\boldsymbol{\theta}) := \log \mathcal{L}(\boldsymbol{\theta}) = n \log \frac{1}{\sigma \sqrt{2\pi}} - \frac{1}{2\sigma^2} \sum_{i=1}^n \left(y^{(i)} - \mathbf{x}^{(i)} \boldsymbol{\theta} \right)^2$$

New goal: Minimize negative log-likelihood

(leave away scaling factors and constant)

$$L(\boldsymbol{\theta}) = \sum_{i=1}^{n} \left(y^{(i)} - \mathbf{x}^{(i)} \boldsymbol{\theta} \right)^{2} = \sum_{i=1}^{n} \varepsilon^{(i)^{2}}$$

→ For example, solve with least squares



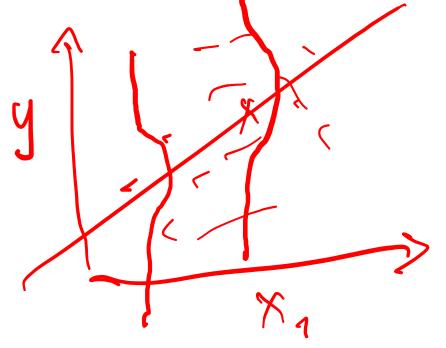
Under the assumption that the errors are Gaussian and i.i.d., the **maximum likelihood estimator** for θ is given by the least squares solution

$$\boldsymbol{\theta}_{\mathrm{MLE}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

What is the variance σ^2 ?

$$\frac{\partial \ell(\boldsymbol{\theta}, \sigma)}{\partial \sigma} = -\frac{n}{\sigma} + \frac{1}{\sigma^3} \sum_{i=1}^{n} \left(y^{(i)} - \mathbf{x}^{(i)}^T \boldsymbol{\theta} \right)^2 = 0.$$

$$\rightarrow \quad \sigma^2 = \frac{1}{n} \sum_{i=1}^n \left(y^{(i)} - \mathbf{x}^{(i)T} \boldsymbol{\theta} \right)^2 = \frac{1}{n} \sum_{i=1}^n \varepsilon^{(i)^2}$$



variance is the mean squared error





Logistic regression, a binary classifier

Labels

$$y^{(i)} \in \{0, 1\}$$

Probability of the data

$$p := P(y^{(i)} = 1 \mid \mathbf{x}^{(i)}) \rightarrow P(y^{(i)} = 0 \mid \mathbf{x}^{(i)}) = 1 - p.$$

What are the odds?

Useful in practice: log-odds

$$\frac{p}{p-1} \qquad \qquad \log \operatorname{it}(p) := \log \frac{p}{1-p}, \quad \operatorname{logit}: (0,1) \to \mathbb{R}$$



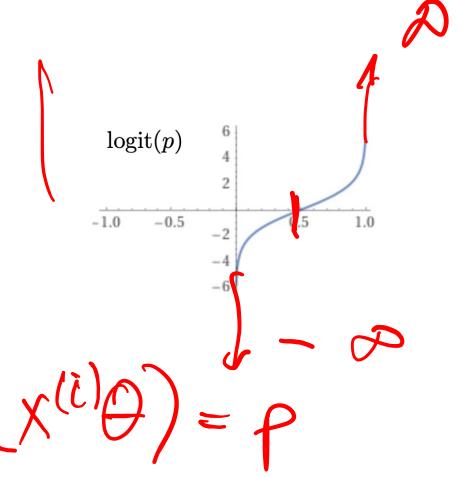
Logistic regression, the model

The log-odds are modelled by a linear function

$$logit(p) = \mathbf{x}^{(i)}\boldsymbol{\theta} = b + \mathbf{x}^{(i)}\mathbf{w}$$

$$\operatorname{logit}(p) := \log \frac{p}{1 - p}, \quad \operatorname{logit}: (0, 1) \to \mathbb{R}$$

$$2(\log(i)) = P d(x(i))$$





Logistic regression, the model

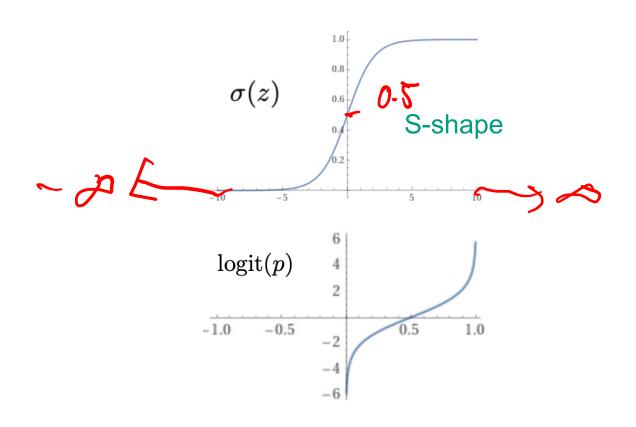
But we are interested in the probability

$$p = \sigma(z) = \frac{1}{1 + \exp(-z)}$$

Logistic function / sigmoid function

$$\sigma: \mathbb{R} \to (0,1), \quad \sigma(z) = \frac{1}{1 + \exp(-z)}$$

Sigmoid is the inverse of logit





Logistic regression, the plan

Use linear model for the log-odds

Convert to probability using logistic function

Use thresholding to predict class label

$$\hat{y}(z) = \begin{cases} 1 & \text{if } \sigma(z) \le 0.5 \\ 0 & \text{otherwise} \end{cases} = \begin{cases} 1 & \text{if } z \le 0 \\ 0 & \text{otherwise} \end{cases}$$



Probabilities of the data

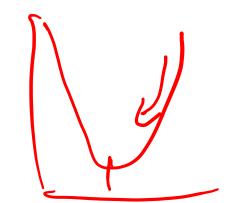
$$P(y^{(i)}=1|\mathbf{x}^{(i)}; \boldsymbol{\theta}) = \sigma(z)$$
 $P(y^{(i)}=0 \mid \mathbf{x}^{(i)}; \boldsymbol{\theta}) = 1-\sigma(z)$
Combined

$$P(y^{(i)} \mid \mathbf{x}^{(i)}; \boldsymbol{\theta}) = \sigma(z)^{y^{(i)}} (1 - \sigma(z))^{(1 - y^{(i)})}$$

 $y^{(i)} \in \{0, 1\}$

Like for linear regression: maximize like hood of the parameters.

$$\mathcal{L}(\boldsymbol{\theta}) = P(\mathbf{y} \mid \mathbf{X}; \boldsymbol{\theta})$$





Goal: maximize likelihood of the parameters

$$\mathcal{L}(\boldsymbol{\theta}) = P(\mathbf{y} \mid \mathbf{X}; \boldsymbol{\theta})$$

$$= \prod_{i=1}^{n} P(y^{(i)} \mid x^{(i)}; \boldsymbol{\theta}), \quad \text{(samples are i.i.d.)}$$

$$= \prod_{i=1}^{n} \left[\sigma(z)^{y^{(i)}} (1 - \sigma(z))^{(1-y^{(i)})} \right].$$



Goal: maximize likelihood of the parameters

Use log-likelihood instead

$$\ell(\boldsymbol{\theta}) = \log \mathcal{L}(\boldsymbol{\theta}) = \log P(\mathbf{y} \mid \mathbf{X}; \boldsymbol{\theta})$$

$$= \log \prod_{i=1}^{n} \left[\sigma(z)^{y^{(i)}} (1 - \sigma(z))^{(1 - y^{(i)})} \right]$$

$$= \sum_{i=1}^{n} \left[y^{(i)} \log(\sigma(z)) + (1 - y^{(i)}) \log(1 - \sigma(z)) \right]$$



Goal: maximize likelihood of the parameters

Use log-likelihood instead

$$\ell(\boldsymbol{\theta}) = \log \mathcal{L}(\boldsymbol{\theta}) = \log P(\mathbf{y} \mid \mathbf{X}; \boldsymbol{\theta})$$

$$= \log \prod_{i=1}^{n} \left[\sigma(z)^{y^{(i)}} (1 - \sigma(z))^{(1 - y^{(i)})} \right]$$

$$= \sum_{i=1}^{n} \left[y^{(i)} \log(\sigma(z)) + (1 - y^{(i)}) \log(1 - \sigma(z)) \right]$$



Goal: maximize likelihood of the parameters \rightarrow minimize negative log-likelihood

Use log-likelihood instead

$$L(oldsymbol{ heta}) = -\ell(oldsymbol{ heta})$$

$$\ell(oldsymbol{ heta}) = \sum_{i=1}^n \left[y^{(i)} \log(\sigma(z)) + (1-y^{(i)}) \log(1-\sigma(z)) \right]$$

$$\frac{\partial L(\boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = \mathbf{0}$$
 condition for minimum



$$\sum_{i=1}^{n} \left[y^{(i)} \log(\sigma(z)) + (1 - y^{(i)}) \log(1 - \sigma(z)) \right]$$

$$\frac{\partial \sigma(z)}{\partial z} =$$

$$rac{\partial (\log \sigma(z))}{\partial oldsymbol{ heta}} =$$

$$\frac{\partial (\log(1-\sigma(z)))}{\partial \boldsymbol{\theta}} =$$



Goal: maximize likelihood of the parameters \rightarrow minimize negative log-likelihood

Use log-likelihood instead

$$L(oldsymbol{ heta}) = -\ell(oldsymbol{ heta})$$
 $\ell(oldsymbol{ heta}) = \sum_{i=1}^n \left[y^{(i)} \log(\sigma(z)) + (1-y^{(i)}) \log(1-\sigma(z))
ight]$

$$\frac{\partial L(\boldsymbol{\theta})}{\partial \boldsymbol{\theta}} = -\sum_{i=1}^{n} \left[y^{(i)} (1 - \sigma(z)) \mathbf{x}^{(i)} - (1 - y^{(i)}) \sigma(z) \mathbf{x}^{(i)} \right]$$

$$= -\sum_{i=1}^{n} \left[(y^{(i)} - \sigma(z)) \mathbf{x}^{(i)} \right] = \mathbf{0},$$
The point is the second state of the point in the point in

Component-wise

$$\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} = -\sum_{i=1}^n \left(x_j^{(i)} \left[(y^{(i)} - \sigma(\mathbf{x}^{(i)}\boldsymbol{\theta})) \right] \right) = 0$$

Same gradient as for linear regression / Adaline, except for $\sigma(z)$!

But nonlinear, no explicit solution!

Logistic regression summary

Learn / train / fit

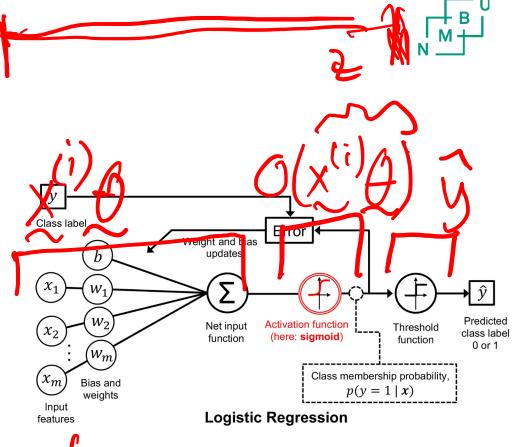
$$\frac{\partial L(\boldsymbol{\theta})}{\partial \theta_j} = -\sum_{i=1}^n \left(x_j^{(i)} \left[(y^{(i)} - \sigma(\mathbf{x}^{(i)}\boldsymbol{\theta})) \right] \right) = 0.$$

$$\boldsymbol{\theta}^{n+1} = \boldsymbol{\theta} - \eta \frac{\partial L(\boldsymbol{\theta})}{\partial \boldsymbol{\theta}}$$
 (Batch) gradient descent

Predict

$$z = \mathbf{x}\boldsymbol{\theta}$$

Predicted Activation function Threshold class label (here: sigmoid) function 0 or 1 Class membership probability. $p(y = 1 \mid x)$ **Logistic Regression** log_regression.ipynb

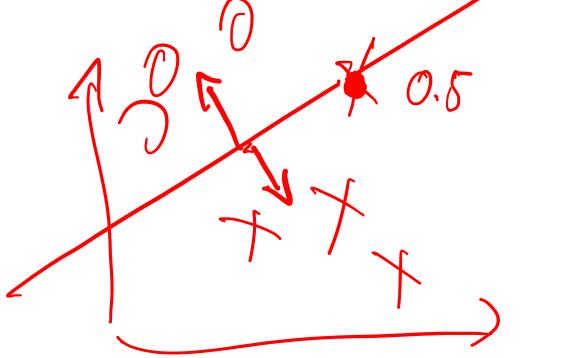




Logistic regression summary

Logistic regression gives us label and probability

Very popular e.g. in health section, but really everywhere



log_regression.ipynb



