

# SI231 Matrix Analysis and Computations

## Positive Semidefinite Matrices

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# Positive Semidefinite Matrices

- positive semidefinite (PSD) matrices
- properties of PSD matrices
- PSD matrix inequalities
- Schur complement
- application: factor models
- application: graph matrices
- application: log-determinant function
- application: subspace method for super-resolution spectral analysis
- application: Euclidean distance matrices

## Highlights

- a matrix  $\mathbf{A} \in \mathbb{S}^n$  is said to be **positive semidefinite (PSD)** if

$$\mathbf{x}^T \mathbf{A} \mathbf{x} \geq 0, \quad \text{for all } \mathbf{x} \in \mathbb{R}^n;$$

and **positive definite (PD)** if

$$\mathbf{x}^T \mathbf{A} \mathbf{x} > 0, \quad \text{for all } \mathbf{x} \in \mathbb{R}^n \text{ with } \mathbf{x} \neq \mathbf{0}$$

- in this case, we are interested in symmetric positive semidefinite matrices
- a matrix  $\mathbf{A} \in \mathbb{S}^n$  is PSD (resp. PD)
  - if and only if its eigenvalues are all non-negative (resp. positive);
  - if and only if it can be factored as  $\mathbf{A} = \mathbf{B}^T \mathbf{B}$  for some  $\mathbf{B} \in \mathbb{R}^{m \times n}$
- there are more general definitions of definiteness, including real non-symmetric matrices, or non-Hermitian ones
- in this lecture, we will deal with the real symmetric matrices—the Hermitian case follows along the same lines

# Quadratic Form

Let  $\mathbf{A} \in \mathbb{S}^n$ . For  $\mathbf{x} \in \mathbb{R}^n$ , the matrix product

$$\mathbf{x}^T \mathbf{A} \mathbf{x}$$

is called a **quadratic form**.

- some basic facts (try to verify):

- $\mathbf{x}^T \mathbf{A} \mathbf{x} = \sum_{i=1}^n \sum_{j=1}^n a_{ij} x_i x_j = \sum_{i=1}^n a_{ii} x_i^2 + \sum_{i=1}^{n-1} \sum_{j=i+1}^n 2a_{ij} x_i x_j$
- $\mathbf{x}^T \mathbf{A} \mathbf{x} = \sum_{i=1}^n a_{ii} x_i^2 + \sum_{i=1}^{n-1} \sum_{j=i+1}^n (a_{ij} + a_{ji}) x_i x_j$  for general  $\mathbf{A} \in \mathbb{R}^{n \times n}$ , there may exist  $\mathbf{A}_1$  and  $\mathbf{A}_2$  s.t.  $\mathbf{x}^T \mathbf{A}_1 \mathbf{x} = \mathbf{x}^T \mathbf{A}_2 \mathbf{x}$

- \* it suffices to consider unique symmetric  $\mathbf{A}$  for general  $\mathbf{A} \in \mathbb{R}^{n \times n}$  since

$$\mathbf{x}^T \mathbf{A} \mathbf{x} = \mathbf{x}^T \left[ \frac{1}{2} (\mathbf{A} + \mathbf{A}^T) \right] \mathbf{x}$$

- complex case:

- \* the (complex) quadratic form is defined as  $\mathbf{x}^H \mathbf{A} \mathbf{x}$ , where  $\mathbf{x} \in \mathbb{C}^n$
- \* for  $\mathbf{A} \in \mathbb{H}^n$ ,  $\mathbf{x}^H \mathbf{A} \mathbf{x}$  is real for any  $\mathbf{x} \in \mathbb{C}^n$

# Positive Semidefinite Matrices

A matrix  $\mathbf{A} \in \mathbb{S}^n$  is said to be

- **positive semidefinite (PSD)** if  $\mathbf{x}^T \mathbf{A} \mathbf{x} \geq 0$  for all  $\mathbf{x} \in \mathbb{R}^n$  (not very often called nonnegative definite)
- **positive definite (PD)** if  $\mathbf{x}^T \mathbf{A} \mathbf{x} > 0$  for all  $\mathbf{x} \in \mathbb{R}^n$  with  $\mathbf{x} \neq \mathbf{0}$
- **indefinite** if both  $\mathbf{A}$  and  $-\mathbf{A}$  are not PSD

Notation:

- $\mathbf{A} \succeq \mathbf{0}$  means that  $\mathbf{A}$  is PSD
- $\mathbf{A} \succ \mathbf{0}$  means that  $\mathbf{A}$  is PD
- $\mathbf{A} \not\succeq \mathbf{0}$  means that  $\mathbf{A}$  is indefinite
- if  $\mathbf{A}$  is PD, then it is also PSD
- a quadratic form is called PSD (resp. PD) if  $\mathbf{A}$  is PSD (resp. PD)
- Concepts negative semidefinite (NSD) and negative definite (ND) may be defined by reversing the inequalities or, equivalently, by saying  $-\mathbf{A}$  is PSD or PD, resp.
- Positive (semi)definite and negative (semi)definite matrices together are called **definite matrices**.

## Example: Covariance Matrices

- let  $\mathbf{y}_0, \mathbf{y}_2, \dots, \mathbf{y}_{T-1} \in \mathbb{R}^n$  be a sequence of multi-dimensional data samples
  - examples: multivariate features in machine learning, patches in image processing, multi-channel signals in signal processing, history of returns of assets in finance [\[Brodie-Daubechies-et al.'09\]](#), ...
- sample mean:  $\hat{\boldsymbol{\mu}} = \frac{1}{T} \sum_{t=0}^{T-1} \mathbf{y}_t$
- sample covariance:  $\hat{\boldsymbol{\Sigma}} = \frac{1}{T} \sum_{t=0}^{T-1} (\mathbf{y}_t - \hat{\boldsymbol{\mu}})(\mathbf{y}_t - \hat{\boldsymbol{\mu}})^T$
- a sample covariance is PSD:  $\mathbf{x}^T \hat{\boldsymbol{\Sigma}} \mathbf{x} = \frac{1}{T} \sum_{t=0}^{T-1} |(\mathbf{y}_t - \hat{\boldsymbol{\mu}})^T \mathbf{x}|^2 \geq 0$
- the (statistical) covariance matrix (or variance-covariance mat.) of  $\mathbf{y}_t$  is also PSD
  - to put into context, assume that  $\mathbf{y}_t$  is a wide-sense (or weakly) stationary random process
  - the covariance, defined as  $\boldsymbol{\Sigma} = \mathbb{E}[(\mathbf{y}_t - \boldsymbol{\mu})(\mathbf{y}_t - \boldsymbol{\mu})^T]$  where the mean  $\boldsymbol{\mu} = \mathbb{E}[\mathbf{y}_t]$ , is PSD:  $\mathbf{x}^T \boldsymbol{\Sigma} \mathbf{x} = \mathbf{x}^T \mathbb{E}[(\mathbf{y}_t - \boldsymbol{\mu})(\mathbf{y}_t - \boldsymbol{\mu})^T] \mathbf{x} = \mathbb{E}[|(\mathbf{y}_t - \boldsymbol{\mu})^T \mathbf{x}|^2] \geq 0$

## Example: Covariance Matrices

- define  $\mathbf{Y} = [\mathbf{y}_0 \ \mathbf{y}_1 \ \dots \ \mathbf{y}_{T-1}] \in \mathbb{R}^{n \times T}$
- the sample mean:  $\hat{\boldsymbol{\mu}} = \frac{1}{T} \mathbf{Y} \mathbf{1}$
- subtracting  $\hat{\boldsymbol{\mu}}$  from each column gives the row-centered data matrix

$$\mathbf{Y}_c = \mathbf{Y} - \hat{\boldsymbol{\mu}} \mathbf{1}^T = \mathbf{Y} - \frac{1}{T} \mathbf{Y} \mathbf{1} \mathbf{1}^T = \mathbf{Y} (\mathbf{I} - \frac{1}{T} \mathbf{1} \mathbf{1}^T) = \mathbf{Y} \mathbf{C}$$

where  $\mathbf{C} = \mathbf{I} - \frac{1}{T} \mathbf{1} \mathbf{1}^T$  is called the **centering matrix**

- a symmetric, PSD, and idempotent matrix, and  $\text{rank}(\mathbf{C}) = T - 1$
  - singular and has the eigenvalue 1 of multiplicity  $n - 1$  and eig. 0 of multip. 1
  - has a nullspace of dimension 1, along the vector  $\mathbf{1}$
  - an orthogonal projection matrix (projection to a subspace of all  $n$ -dim. vectors whose components sum to zero)
  - left-multiplying it with a vector has the same effect as subtracting the mean of the components of the vector from every component of that vector
- the sample covariance matrix based on  $\mathbf{Y}_c$  (a Gram matrix) is

$$\hat{\boldsymbol{\Sigma}} = \frac{1}{T} \mathbf{Y}_c \mathbf{Y}_c^T = \frac{1}{T} \mathbf{Y} \mathbf{C} \mathbf{C}^T \mathbf{Y}^T = \frac{1}{T} \mathbf{Y} \mathbf{C}^2 \mathbf{Y}^T = \frac{1}{T} \mathbf{Y} \mathbf{C} \mathbf{Y}^T = \frac{1}{T} \mathbf{Y} \mathbf{Y}^T - \hat{\boldsymbol{\mu}} \hat{\boldsymbol{\mu}}^T$$

## Properties of PSD Matrices

results that immediately follow from the definition: let  $\mathbf{A}, \mathbf{B}, \mathbf{C} \in \mathbb{S}^n$ .

- $\mathbf{A} \succeq \mathbf{0}, \alpha \geq 0$  (resp.  $\mathbf{A} \succ \mathbf{0}, \alpha > 0$ )  $\implies \alpha \mathbf{A} \succeq \mathbf{0}$  (resp.  $\alpha \mathbf{A} \succ \mathbf{0}$ )
- $\mathbf{A}, \mathbf{B} \succeq \mathbf{0}$  (resp.  $\mathbf{A} \succeq \mathbf{0}, \mathbf{B} \succ \mathbf{0}$ )  $\implies \mathbf{A} + \mathbf{B} \succeq \mathbf{0}$  (resp.  $\mathbf{A} + \mathbf{B} \succ \mathbf{0}$ )
- $\mathbf{A} \succ \mathbf{0}, \mathbf{B} \succeq \mathbf{0} \implies \mathbf{A} + \mathbf{B} \succeq \mathbf{0}$
- $\mathbf{A} \succ \mathbf{0} \iff \mathbf{A}^{-1} \succ \mathbf{0}$
- $\mathbf{A} \succeq \mathbf{0}$  (resp.  $\mathbf{A} \succ \mathbf{0}$ )  $\implies \text{tr}(\mathbf{A}) \geq 0$  and  $\det(\mathbf{A}) \geq 0$  (resp.  $\text{tr}(\mathbf{A}) > 0$  and  $\det(\mathbf{A}) > 0$ )
- The set of PSD matrices is convex. (the basis of semidefinite programming (SDP), a subfield of convex optimization in matrix variables)
- ...



## Properties of PSD Matrices: Eigenvalues

**Theorem 1.** Let  $\mathbf{A} \in \mathbb{S}^n$ , and let  $\lambda_1, \dots, \lambda_n$  be the eigenvalues of  $\mathbf{A}$ . We have

1.  $\mathbf{A} \succeq \mathbf{0} \iff \lambda_i \geq 0$  for  $i = 1, \dots, n$
  2.  $\mathbf{A} \succ \mathbf{0} \iff \lambda_i > 0$  for  $i = 1, \dots, n$
  3.  $\mathbf{A} \not\succeq \mathbf{0} \iff \lambda_i > 0$  for some  $i$  and  $\lambda_i < 0$  for some  $i$
- proof: let  $\mathbf{A} = \mathbf{V}\mathbf{\Lambda}\mathbf{V}^T$  be the eigendecomposition of  $\mathbf{A}$ .

$$\begin{aligned}\mathbf{A} \succeq \mathbf{0} &\iff \mathbf{x}^T \mathbf{V} \mathbf{\Lambda} \mathbf{V}^T \mathbf{x} \geq 0, \quad \text{for all } \mathbf{x} \in \mathbb{R}^n \\ &\iff \mathbf{z}^T \mathbf{\Lambda} \mathbf{z} \geq 0, \quad \text{for all } \mathbf{z} \in \mathcal{R}(\mathbf{V}^T) = \mathbb{R}^n \\ &\iff \sum_{i=1}^n \lambda_i |z_i|^2 \geq 0, \quad \text{for all } \mathbf{z} \in \mathbb{R}^n \\ &\iff \lambda_i \geq 0 \text{ for all } i\end{aligned}$$

The PD and indefinite cases are proven by the same manner.

## Example: Hessian

- let  $f : \mathbb{R}^n \rightarrow \mathbb{R}$  be a twice differentiable function
- the **Hessian** of  $f$ , denoted by  $\nabla^2 f(\mathbf{x}) \in \mathbb{S}^n$ , is a matrix whose  $(i, j)$ th entry is given by

$$[\nabla^2 f(\mathbf{x})]_{i,j} = \frac{\partial^2 f}{\partial x_i \partial x_j}$$

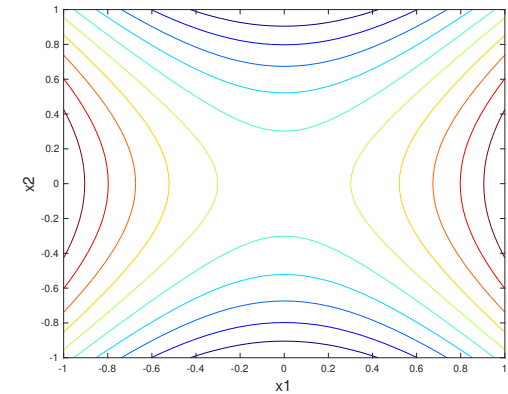
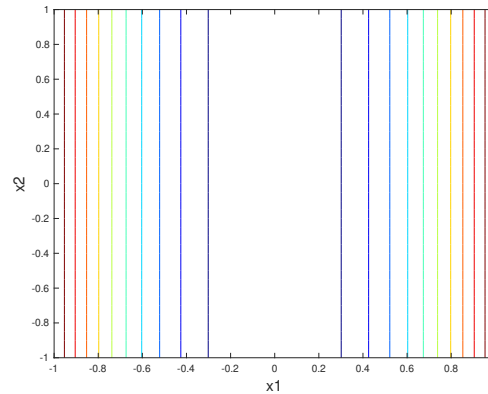
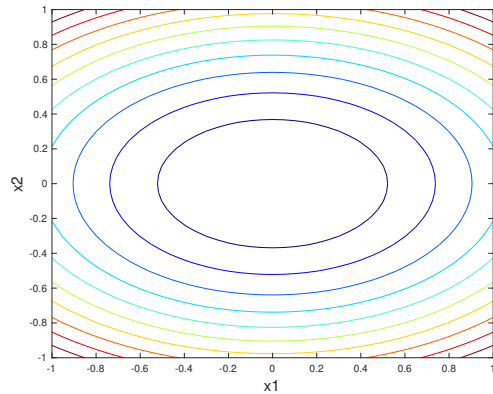
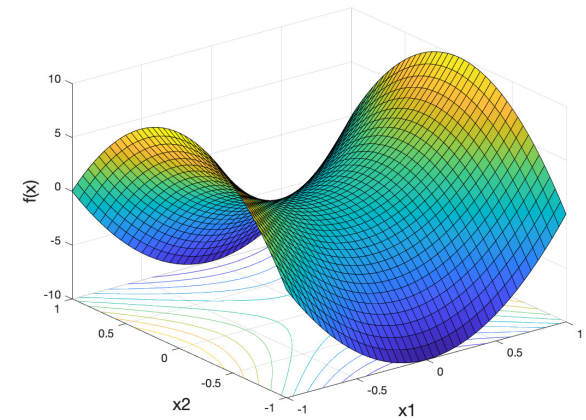
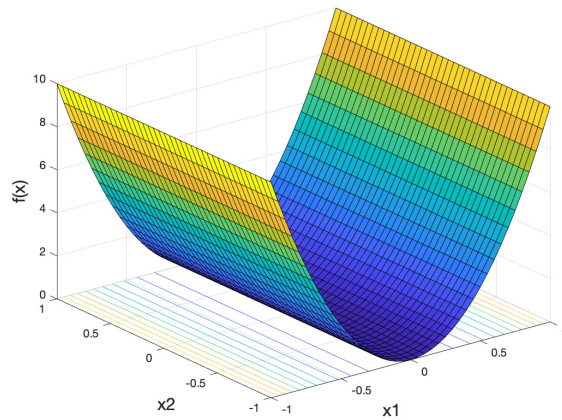
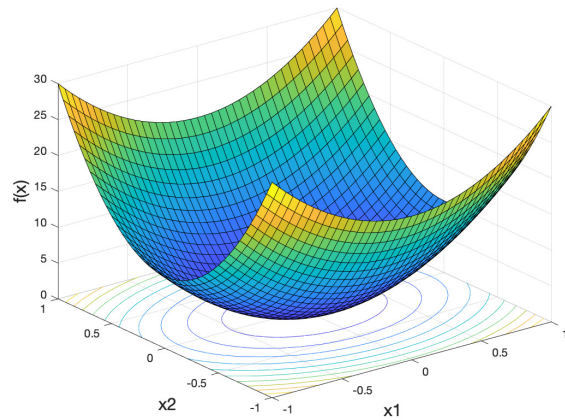
- **Fact:**  $f$  is convex if and only if  $\nabla^2 f(\mathbf{x}) \succeq \mathbf{0}$  for all  $\mathbf{x}$  in the problem domain
- example: consider the quadratic function with  $\mathbf{R} \in \mathbb{S}^n$

$$f(\mathbf{x}) = \frac{1}{2} \mathbf{x}^T \mathbf{R} \mathbf{x} + \mathbf{q}^T \mathbf{x} + c$$

It can be verified that  $\nabla^2 f(\mathbf{x}) = \mathbf{R}$ . Thus,  $f$  is convex (resp. strictly convex) if and only if  $\mathbf{R} \succeq \mathbf{0}$  (resp.  $\succ \mathbf{0}$ )

# Illustration of Quadratic Functions

$$f(\mathbf{x}) = \mathbf{x}^T \mathbf{A} \mathbf{x}$$



(a) PSD  $\mathbf{A}$  with all  $\lambda_i > 0$ . (b) PSD  $\mathbf{A}$  with some  $\lambda_i = 0$ .

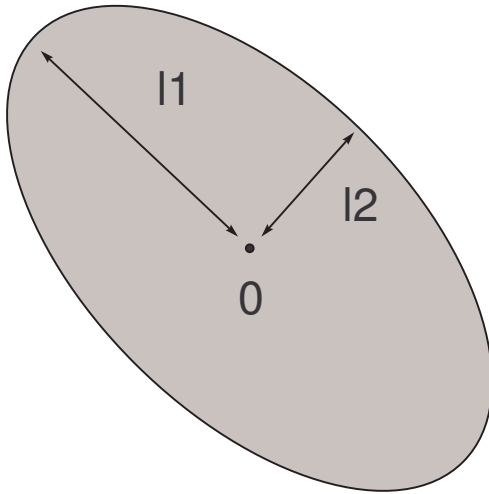
(c) indefinite  $\mathbf{A}$ .

## Example: Ellipsoid

- an ellipsoid of  $\mathbb{R}^n$  centered at  $\mathbf{0}$  is defined as

$$\mathcal{E}(\mathbf{P}) = \{ \mathbf{x} \in \mathbb{R}^n \mid \mathbf{x}^T \mathbf{P}^{-1} \mathbf{x} \leq 1 \},$$

for some PD  $\mathbf{P} \in \mathbb{S}^n$



let  $\mathbf{P} = \mathbf{V} \mathbf{\Lambda} \mathbf{V}^T$  be the eigendecomposition

- $\mathbf{V}$  determines the directions of the semi-axes
- $\lambda_1, \dots, \lambda_n$  determine the lengths of the semi-axes
- $\ell_i = \lambda_i^{\frac{1}{2}} \mathbf{v}_i$

- note:
  - in direction  $\mathbf{v}_1$ ,  $\mathbf{x}^T \mathbf{P}^{-1} \mathbf{x}$  is large, hence ellipsoid is fat in direction  $\mathbf{v}_1$
  - in direction  $\mathbf{v}_n$ ,  $\mathbf{x}^T \mathbf{P}^{-1} \mathbf{x}$  is small, hence ellipsoid is thin in direction  $\mathbf{v}_n$
  - $\sqrt{\lambda_1 / \lambda_n}$  gives maximum eccentricity

## Example: Multivariate Gaussian Distribution

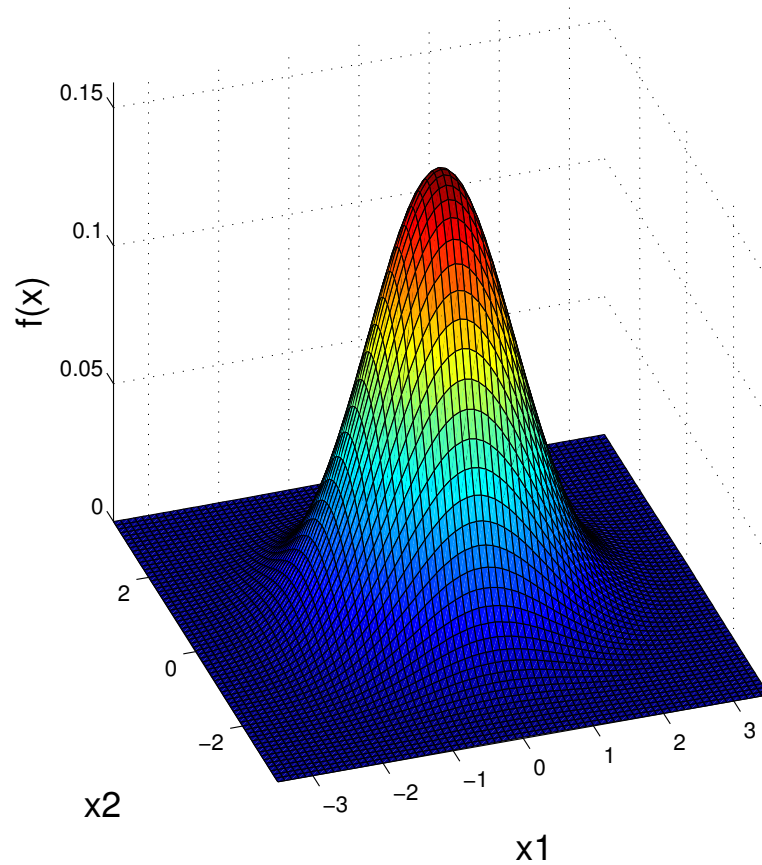
- probability density function for a Gaussian-distributed vector  $\mathbf{x} \in \mathbb{R}^n$ :

$$p(\mathbf{x}; \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{\frac{n}{2}} (\det(\boldsymbol{\Sigma}))^{\frac{1}{2}}} \exp \left( -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right)$$

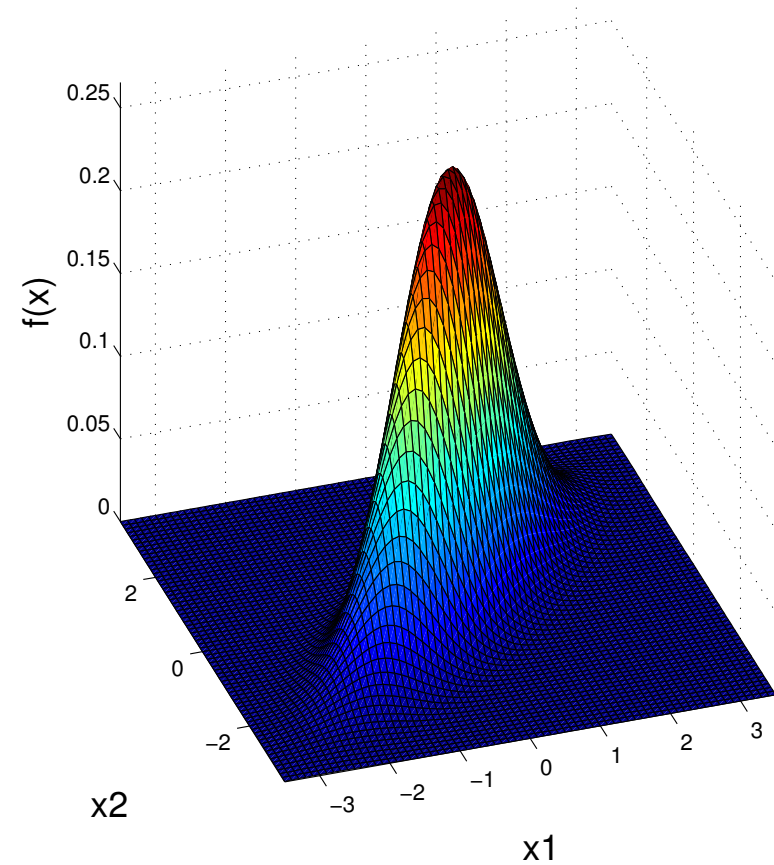
where  $\boldsymbol{\mu}$  and  $\boldsymbol{\Sigma}$  are the mean and covariance matrix of  $\mathbf{x}$ , resp.

- $\boldsymbol{\Sigma}$  is PD
- $\boldsymbol{\Sigma}$  determines how  $\mathbf{x}$  is spread, by the same way as in ellipsoid

## Example: Multivariate Gaussian Distribution



(a)  $\mu = 0, \Sigma = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}.$



(b)  $\mu = 0, \Sigma = \begin{bmatrix} 1 & 0.8 \\ 0.8 & 1 \end{bmatrix}.$

## Example: Multivariate Generalized Gaussian Distribution

- probability density function for a generalized Gaussian-distributed vector  $\mathbf{x} \in \mathbb{R}^n$ :

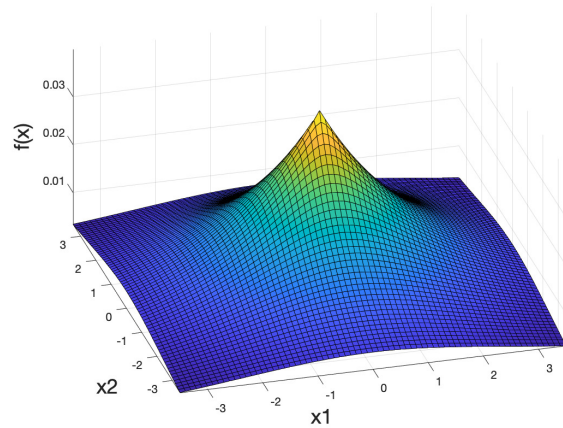
$$p(\mathbf{x}; \mathbf{m}, \mathbf{C}, \nu) = \frac{\frac{\nu}{2}\Gamma(\frac{n}{2})}{(2^{\frac{2}{\nu}}\pi)^{\frac{n}{2}}\Gamma(\frac{n}{\nu})(\det(\mathbf{C}))^{\frac{1}{2}}}\exp\left(-\frac{1}{2}\left((\mathbf{x} - \mathbf{m})^T\mathbf{C}^{-1}(\mathbf{x} - \mathbf{m})\right)^{\frac{\nu}{2}}\right)$$

where  $\mathbf{m}$  and  $\mathbf{C}$  are the location and scatter matrix (a.k.a. scale matrix) of  $\mathbf{x}$ , resp., and  $\nu$  is the shape parameter (a.k.a. form parameter)

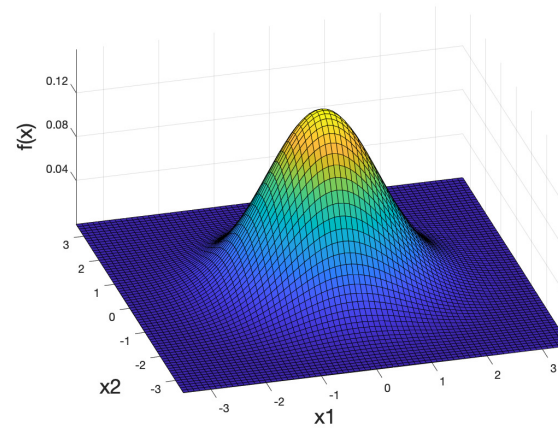
- $\mathbf{C}$  is PD and determines how  $\mathbf{x}$  is spread
- $\nu$  affect the shape of the distribution
  - \* it becomes the multivariate Gaussian distribution when  $\nu = 2$
  - \* more peaky with heavy tails if  $\nu < 2$  (it becomes a multivariate Laplacian distribution when  $\nu = 1$ )
  - \* less peaky with light tails if  $\nu > 2$  (it tends to converge to a multivariate uniform distribution when  $\nu \rightarrow \infty$ )
- For Gaussian distribution, the location and scatter parameters correspond to the mean and covariance, resp. But this is not necessarily true for other distributions.

# Example: Multivariate Generalized Gaussian Distribution

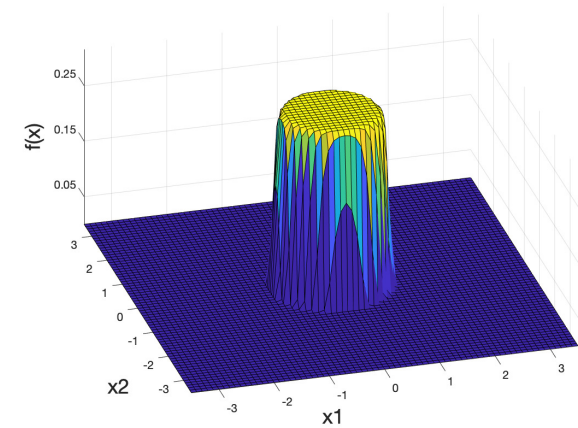
$$\mathbf{m} = \mathbf{0}, \quad \mathbf{C} = \begin{bmatrix} 1 & 0 \\ 0 & 1 \end{bmatrix}$$



(a)  $\nu = 1$ .



(b)  $\nu = 2$ .



(c)  $\nu = 30$ .



## Properties of PSD Matrices

- it can be directly seen from the definition that
  - $\mathbf{A} \succeq \mathbf{0} \implies a_{ii} \geq 0$  for all  $i$
  - $\mathbf{A} \succ \mathbf{0} \implies a_{ii} > 0$  for all  $i$
- extension (also direct): partition  $\mathbf{A}$  as

$$\mathbf{A} = \begin{bmatrix} \mathbf{A}_{11} & \mathbf{A}_{12} \\ \mathbf{A}_{21} & \mathbf{A}_{22} \end{bmatrix}.$$

Then,  $\mathbf{A} \succeq \mathbf{0} \implies \mathbf{A}_{11} \succeq \mathbf{0}, \mathbf{A}_{22} \succeq \mathbf{0}$ . Also,  $\mathbf{A} \succ \mathbf{0} \implies \mathbf{A}_{11} \succ \mathbf{0}, \mathbf{A}_{22} \succ \mathbf{0}$

- further extension:
  - a **principal submatrix** of  $\mathbf{A}$ , denoted by  $\mathbf{A}_{\mathcal{I}}$ , where  $\mathcal{I} = \{i_1, \dots, i_m\} \subseteq \{1, \dots, n\}$ ,  $m < n$ , is a submatrix obtained by keeping only the rows and columns indicated by  $\mathcal{I}$ ; i.e.,  $[\mathbf{A}_{\mathcal{I}}]_{jk} = a_{i_j, i_k}$  for all  $j, k \in \{1, \dots, m\}$
  - if  $\mathbf{A}$  is PSD (resp. PD), then any principal submatrix of  $\mathbf{A}$  is PSD (resp. PD), and then any principal minor of  $\mathbf{A}$  is nonnegative (resp. positive)

## Properties of PSD Matrix

- (Sylvester's criterion). Let  $\mathbf{A} \in \mathbb{S}^n$ .
  - $\mathbf{A}$  is PD  $\iff$  all its leading (resp., trailing) principal minors are positive (for  $\mathbf{A} \in \mathbb{S}^n$ , the positivity of the leading principal minors implies the positivity of all its principal minors)
  - $\mathbf{A}$  is PSD  $\iff$  all its principal minors are nonnegative
  - If the first  $n - 1$  leading principal minors (resp., the last  $n - 1$  trailing principal minors) of  $\mathbf{A}$  are positive and  $\det(\mathbf{A}) \geq 0$ , then  $\mathbf{A}$  is PSD.
- $\mathbf{A}$  is ND  $\iff$  its odd leading principal minors are negative and even are positive
- $\mathbf{A}$  is NSD  $\iff$  its odd principal minors are nonpositive and even are nonnegative
- $\mathbf{A}$  is indefinite  $\iff$  there are two of its odd leading principal minors that have different signs or there is one of its even leading principal minors that is negative

## Properties of PSD Matrix

- To obtain conditions for a matrix to be PD or ND, we need to examine the leading principal minors.
- To obtain conditions for a matrix to be PSD or NSD, we need to examine all the principal minors.
- Procedures for checking the definiteness of a matrix
  - find the leading principal minors and check if the conditions for positive or negative definiteness are satisfied; if they are, the the matrix is PD or ND
  - if the conditions are not satisfied, check if they are strictly violated; if they are, then the matrix is indefinite
  - if the conditions are not strictly violated, find all its principal minors and check if the conditions for positive or negative semidefiniteness are satisfied

## Properties of PSD Matrix

- $\mathbf{A}$  is PSD,  $\mathbf{x}^T \mathbf{A} \mathbf{x} = 0 \iff \mathbf{A} \mathbf{x} = \mathbf{0}$  for an  $\mathbf{x}$ . (how to prove it?)
  - proved by eigenvalue properties of PSD matrices
  - alternative proof: the “if” part is easy; the “only if” part: constructing

$$p(\lambda) = (\mathbf{x} + \lambda \mathbf{y})^T \mathbf{A} (\mathbf{x} + \lambda \mathbf{y}) = \lambda^2 \mathbf{y}^T \mathbf{A} \mathbf{y} + 2\lambda \mathbf{y}^T \mathbf{A} \mathbf{x} + \mathbf{x}^T \mathbf{A} \mathbf{x}.$$

Since for all  $\mathbf{x}$ ,  $\lambda$ , and  $\mathbf{y}$ ,  $p(\lambda) \geq 0$ , we have the discriminant for  $p(\lambda)$  should be nonpositive, i.e.,

$$4(\mathbf{y}^T \mathbf{A} \mathbf{x})^2 - 4(\mathbf{y}^T \mathbf{A} \mathbf{y})(\mathbf{x}^T \mathbf{A} \mathbf{x}) \leq 0$$

If  $\mathbf{x}^T \mathbf{A} \mathbf{x} = 0$ , the discriminant is nonpositive only if  $\mathbf{y}^T \mathbf{A} \mathbf{x} = 0$  for all  $\mathbf{y}$  or, equivalently, if  $\mathbf{A} \mathbf{x} = \mathbf{0}$ .

- $\mathbf{A}$  is PSD and nonsingular  $\iff \mathbf{A}$  is PD
- for a PSD  $\mathbf{A}$ , it is PD  $\iff \mathbf{A}$  is nonsingular
- $\mathbf{A} \succ \mathbf{0} \iff \mathbf{A}^{-1} \succ \mathbf{0}$