Homework 3

Algorithms for Big Data CS498ABD Spring 2019

Due: 10am, Wednesday, March 27th

Instructions:

- Each home work can be done in a group of size at most two. Only one home work needs to be submitted per group. However, we recommend that each of you think about the problems on your own first.
- Homework needs to be submitted in pdf format on Gradescope. See https://courses.engr.illinois.edu/cs374/fa2018/hw-policies.html for more detailed instructions on Gradescope submissions.
- Follow academic integrity policies as laid out in student code. You can consult sources but cite all of them including discussions with other class mates. Write in your own words. See the site mentioned in the preceding item for more detailed policies.

Exercise 1: Frequent items and Misra-Greis Algorithm We saw the deterministic Misra-Greis algorithm that uses k counters and outputs an estimate \hat{f}_i for each f_i such that $f_i - \frac{m}{k+1} \le \hat{f}_i \le f_i$. Here m is the total number of elements in the stream.

• Let m' be the sum of the counters at the end of the algorithm. Show that the actual estimate provided by the algorithm is slightly stonger, namely, for each i,

$$f_i - \frac{m - m'}{k + 1} \le \hat{f}_i \le f_i.$$

- Suppose we have run the (one-pass) Misra-Gries algorithm on two streams σ_1 and σ_2 thereby obtaining a summary for each stream consisting of k counters. Consider the following algorithm for merging these two summaries to produce a single k-counter summary.
 - 1. Combine the two sets of counters, adding up counts for any common items.
 - 2. If more than k counters remain:
 - (a) $c \leftarrow \text{value of } (k+1)\text{th counter, based on decreasing order of value.}$
 - (b) Reduce each counter by c and delete all keys with non-positive values.

Prove that the resulting summary is good for the combined stream $\sigma_1 \cdot \sigma_2$ (concatenation of the two streams) in the sense that frequency estimates obtained from it satisfy the bounds given in the previous part.

Solution (sketch). For part 1, let ℓ be the number of times k+1 counters are decremented. In lecture, it was argued that $\ell(k+1)$ cannot exceed m since we cannot decrement more than there are items. We can argue a stronger bound: $\ell(k+1) + m' \leq m$ because after decrementing, there are still m' items left and this total cannot exceed the length of the stream. Rearranging, we obtain $\ell \leq \frac{m-m'}{k+1}$. The same analysis from class implies that $f_i - \hat{f}_i \le \ell$.

For part 2, if $f_{1,i}$, $f_{2,i}$, f_i are the frequencies of i and $\hat{f}_{1,i}$, $\hat{f}_{2,i}$, \hat{f}_i are the estimated frequencies of i in $\sigma_1, \sigma_2, \sigma_1 \cdot \sigma_2$ respectively, then $f_i = f_{1,i} + f_{2,i}$, $\hat{f}_i \geq \hat{f}_{1,i} + \hat{f}_{2,i} - c$. Similarly, if m_1, m_2, m are the lengths and m'_1, m'_2, m' are the sum of the resulting counters for $\sigma_1, \sigma_2, \sigma_1 \cdot \sigma_2$, respectively, then $m = m_1 + m_2$ and $m' \leq m'_1 + m'_2 - c(k+1)$. Putting these together with the previous part gives

$$f_i - \hat{f_i} \le f_{1,i} + f_{2,i} - \hat{f}_{1,i} - \hat{f}_{2,i} + c \le \frac{m_1 - m_1'}{k+1} + \frac{m_2 - m_2'}{k+1} + \frac{c(k+1)}{k+1} \le \frac{m - m_1'}{k+1}.$$

Exercise 2: Count Sketch In the Count-Sketch analysis we showed that if we choose $w = 3/\epsilon^2$ and $d = \Omega(\log(n))$ that for each i we obtain an estimate \tilde{x}_i such that with high probability $|\tilde{x}_i - x_i| \leq \epsilon ||x||_2$. This can be pessimistic in situations where the data is highly skewed with most of the $||x||_2$ is concentrated in a few coordinates. To make this precise, for some fixed parameter $\ell \in \mathbb{N}$, let $y_i \in \mathbb{R}^n$ be the vector defined by the ℓ largest coordinates (by absolute value) of x, as well as the *i*th coordinate of x, to 0. (All other coordinates are the same as x) Prove that for $\ell = 1/\epsilon^2$, if w is chosen to be $6/\epsilon^2$ and $d = O(\log n)$, then for all $i \in [n]$, with high probability, we have

$$|\tilde{x}_i - x_i| \le \epsilon ||y_i||_2.$$

Solution (sketch). Instead of ℓ biggest, we will use k biggest so we can use notation consistent with that seen in class.

Fix i. Let $T_{\text{big}} = \{i' \mid i' \text{ is one of the } k \text{ biggest coordinates in } x \}$ and $T_{\text{small}} = [n] \setminus$ $(T \cup \{i\})$. Notice that T_{small} differs slightly from lecture, and is set up this way so that $||y_i||_2^2 = \sum_{i' \in T_{\text{small}}} x_{i'}^2$.

As in lecture, set

$$Z_{\ell} = g_{\ell}(i)C[\ell, h_{\ell}(i)] = x_i + \sum_{\substack{i' \in T_{\text{big}} \\ i' \neq i}} g_{\ell}(i)g_{\ell}(i')x_{i'}Y_{i'} + \sum_{\substack{i' \in T_{\text{small}} \\ i' \neq i}} g_{\ell}(i)g_{\ell}(i')x_{i'}Y_{i'}.$$

We will show that $|Z_{\ell}-x_i| \leq \epsilon ||y_i||_2$ with probability at least $\frac{1}{2}$, so that taking the median of $d = O(\log n)$ copies and applying a Chernoff bound gives the desired result.

As before, $|Z_{\ell} - x_i| > \epsilon ||y_i||_2$ means that a large coordinate collides with i or $|Z'_{\ell}| \ge \epsilon ||y_i||_2$ where $Z'_{\ell} = \sum_{i' \in T} g_{\ell}(i)g_{\ell}(i')x_{i'}Y_{i'}$.

Let A be the event that $h_{\ell}(i') = h_{\ell}(i)$ for some $i' \in T_{\text{big}} \setminus \{i\}$. Then just as in class, $P[A] \le \frac{k}{w}$. By the assumption $k = 1/\epsilon^2$, $w = 6/\epsilon^2$, we get $P[A] \le \frac{1}{6}$.

Straightforward calculation similar to the one done in lecture implies $\operatorname{Var}[Z'_{\ell}] \leq \frac{\|y_i\|_2^2}{w}$.

Thus
$$P[|Z'_{\ell}| \ge \epsilon ||y_i||_2] \le \frac{||y_i||_2^2}{w\epsilon^2 ||y_i||_2^2} = \frac{1}{6}.$$

Putting the previous two bounds together gives $P[|Z_{\ell} - x_i| \le \epsilon ||y_i||_2] \le \frac{1}{3}$ which is less than $\frac{1}{2}$, as desired.

Exercise 3. JL preserves angles Recall that the distributional JL lemma implies that a projection matrix Π chosen from an appropriate distribution preserves length of any fixed vector x to within a $(1 \pm \epsilon)$ -factor with probability $1 - \delta$ if the number of dimensions in the projection is $O(\log(1/\delta)/\epsilon^2)$.

- 1. Suppose we have two unit vectors u, v. Prove that Π preserves the dot product between u and v to within a ϵ -additive factor with probability 1δ with a slight increase in dimensions.
- 2. Show that with a slight increase in dimension, Π preserves the angle between any two vectors u, v up to a $\pm \epsilon$ -additive factor.

Hint: Taylor expansion...

Solution (sketch). Recall that

$$||u + v||^2 = ||u||^2 + ||v||^2 + 2\langle u, v \rangle.$$

Rewriting the above w/r/t $\langle u,v\rangle$ ϵ -additive approximations to $\|u\|^2$, $\|v\|^2$, and $\|u+v\|^2$ lead to $c\epsilon$ -additive approximations for $\langle u,v\rangle$ for some constant c>0. JL preserves $\|u\|^2$, $\|v\|^2$, and $\|u+v\|^2$ up to $(1\pm\epsilon)$ -multiplicative factors, which leads to 2ϵ -additive approximations for these quantities because each has norm at most 2.

Part 2 was not very well-designed question. For example, The Taylor expansion is good for some, but not all, values of θ . One can get slightly weaker bounds by other means but it is not easy (although some students still did it)! So we will make this part extra credit.

Exercise 4. Distinct elements in strict turnstile streams. In class, we say how to estimate F_0 in the cash register setting where each item increases the frequency of a coordinate. Here we develop a different algorithm that allows us to estimate F_0 in the strict turnstile model where each item in the stream may increase or decrease a coordinate value, but every coordinate is always nonnegative.

More explicitly, the stream might look like

$$(i_1, \Delta_1), (i_2, \Delta_2), (i_3, \Delta_3), \dots, (i_m, \Delta_m)$$

where for each $j, i_j \in [n]$ and $\Delta_j \in \mathbb{R}$ (positive or negative). The resulting frequency vector is

$$f = \sum_{j=1}^{m} \Delta_{i_j} e_{i_j}.$$

(Here $e_i \in \{0,1\}^n$ is the *i*th indicator vector.) In the strict turnstile model, we are promised that $f_i \geq 0$ for all *i*. The number of distinct elements, *d*, becomes the number of coordinates *i* with $f_i > 0$. Note that this is at most, and not necessarily equal to, the number of distinct elements to every appear in the stream.

Our goal is to develop an algorithm for this harder streaming model. You may assume that $d \geq c/\epsilon^2$ for some conveniently large constant c (or alternatively, the algorithm is allowed to output " $d \leq c/\epsilon^2$ " when d is too small.)

- 1. Consider an ideal hash function $h:[n] \to [k]$, where we assume that k is sufficiently large relative to $\frac{1}{\epsilon}$ (say $k \ge c_0/\epsilon^2$ for any convenient constant c_0 and sufficiently small $\epsilon > 0$). Let p be the probability that at none of the d distinct elements have hash value 1. We first observe that p is increasing in k and decreasing in d.
 - (a) Show that for $k \leq d/(1+\epsilon)$, we have

$$p \le (1 - c\epsilon)e^{-1}$$

for some constant c > 0.

(b) Show that for $k \geq (1 + \epsilon)d$, we have

$$p \ge (1 + c\epsilon)e^{-1}$$

for some constant c > 0.

2. As a warmup, consider the unit incremental case $(\Delta_j = 1 \text{ for all } j)$. For each integer $i \in \mathbb{Z}$, let $k_i = \left\lceil (1+\epsilon)^i \right\rceil$, and let p_i be the probability that an ideal hash function $h_i : [n] \to [k_i]$ hashes none of the d distinct elements to 1. Suppose that for each i, we had a value \widetilde{p}_i such that

$$(1 - c\epsilon)p_i \le \widetilde{p}_i \le (1 + c\epsilon)p_i$$

for some conveniently small constant c > 0. Show how to use these values $\{\widetilde{p}_i, i \in \mathbb{Z}\}$ to identify a value i such that $(1+\epsilon)^{i-2} \leq d \leq (1+\epsilon)^{i+2}$.

- 3. Show how to extend these ideas to obtain an $(1 \pm \epsilon)$ -multiplicative approximation for F_0 with high probability in $O(\log^2 n/\epsilon^3)$ space in the incremental setting (not including space for the ideal hash functions).¹
- 4. Let us now return to the strict turnstile settings, whether the Δ_j 's might be negative, and the goal is to estimate the number of coordinates with strictly positive value at the end of the stream. In the previous two parts, we analyze an algorithm that simply counted the number of elements from the stream to hash to the first bucket. For the turnstile setting this won't work, because an element that is added and then deleted can still be hashed to the first bucket and effect our estimate. Modify the algorithm to work for turnstile streams, using the same amount of space.

¹In a full implementation, the ideal hash functions can be replaced with one with k-wise independence for $k = O(\log(n)/\epsilon^2)$.

Solution (sketch). For part 1(a), use the inequality $(1 - x/d)^d \le e^{-x}$ and the Taylor series expansion of e to obtain

$$p = (1 - 1/k)^d \le (1 - (1 + \epsilon)/d)^d \le e^{-1}e^{\epsilon} \le e^{-1}(1 - \epsilon + \epsilon^2/2) \le e^{-1}(1 - \epsilon/2).$$

For part 1(b), use other exponential identities and Taylor expansions.

For part 2, there exists i such that $d/(1+\epsilon) \le k_i \le d(1+\epsilon)$. For such i, we know that $(1-c\epsilon)e^{-1} \le p_i \le (1+c\epsilon)e^{-1}$, and thus $(1-c\epsilon)^2e^{-1} \le \tilde{p}_i \le (1+c\epsilon)^2e^{-1}$. The number possible values of i is at most $\log_{1+\epsilon} n = O(\log n/\epsilon)$, so we can binary search to find such \tilde{p}_i .

Thus to do part 3, it suffices to compute $(1 \pm c\epsilon)$ approximations for p_i for each i. For each i, let Z_i be the indicator random variable for whether or not something hashes to 1 under some h_i . We know that outputting $\tilde{p}_i = Z_i$ is an unbiased estimator. For large enough p, we can use Chebyshev's inequality to show that taking the average of $O(1/\epsilon^2)$ copies, each with an independently chosen hash function, outputs an approximately correct answer with constant probability. Then take the median of $O(\log n)$ copies and apply a Chernoff bound to obtain the same result with high probability. The space used is $O(\log n/\epsilon^2)$ for each i, for a total of $O(\log^2 n/\epsilon^3)$ space, as desired.

For part 4, instead of keeping track of whether or not something hashes into 1, we keep track of $\sum_{j:h_i(j)=1} \Delta_j$ and return whether or not this sum is positive. Then we can apply the analysis from the previous part.