## Random Numbers

# Anshul Sangrame CS21BTECH11004

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#### 1 Uniform Random Numbers

Let U be a uniform random variable between 0 and 1.

1.1 Generate  $10^6$  samples of U using a C program and save into a file called uni.dat .

#### **Solution:**

Download the following files,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution/1/ coeffs.h wget https://github.com/Anshul-Sangrame/

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /1/1\_1.c

Execute the above files using code,

1.2 Load the uni.dat file into python and plot the empirical CDF of *U* using the samples in uni.dat. The CDF is defined as

$$F_U(x) = \Pr\left(U \le x\right) \tag{1.1}$$

#### **Solution:**

Download the following files

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /1/1\_2.py

Execute the code using command

python3 1 2.py

Plot 1.1 is obtained,

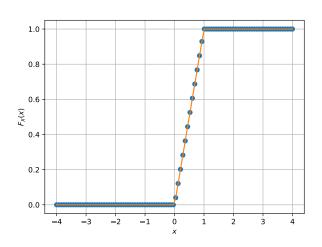


Fig. 1.1

1.3 Find a theoretical expression for  $F_U(x)$ .

#### **Solution:**

Pdf of Uniform distribution between [0,1] is given by,

$$f_U(x) = \begin{cases} 1, & x \in [0, 1] \\ 0, & \text{otherwise} \end{cases}$$
 (1.2)

$$F_U(x) = \int_{-\infty}^x f_U(x)dx \tag{1.3}$$

Case-1: x < 0,

$$F_U(x) = \int_{-\infty}^x 0 dx \tag{1.4}$$

$$=0 (1.5)$$

Case-2:  $x \in [0,1]$ ,

$$F_U(x) = \int_{-\infty}^0 0 dx + \int_0^x 1 dx$$
 (1.6)

$$= x \tag{1.7}$$

Case-3: x > 1,

$$F_U(x) = \int_{-\infty}^0 0 dx + \int_0^1 1 dx + \int_1^x 0 dx \quad (1.8)$$
  
= 1 (1.9)

Hence,

$$F_U(x) = \begin{cases} 0, & x < 0 \\ x, & x \in [0, 1] \\ 1, & x > 1 \end{cases}$$
 (1.10)

1.4 The mean of U is defined as

$$E[U] = \frac{1}{N} \sum_{i=1}^{N} U_i$$
 (1.11)

and its variance as

$$var[U] = E[U - E[U]]^2$$
 (1.12)

Write a C program to find the mean and variance of U.

#### **Solution:**

Download the following files,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution/1/ coeffs.h

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /1/1 4.c

Execute the above files using code,

The following result is obtained

$$E[U] = 0.500007$$
  
var $[U] = 0.083301$ 

1.5 Verify your result theoretically given that

$$E\left[U^{k}\right] = \int_{-\infty}^{\infty} x^{k} dF_{U}(x) dx \qquad (1.13)$$

**Solution:** 

 $F_U(x)$  for uniform distribution,

$$F_U(x) = \begin{cases} 0, & x < 0 \\ x, & x \in [0, 1] \\ 1, & x > 1 \end{cases}$$
 (1.14)

$$E[U] = \int_{-\infty}^{\infty} x dF_U(x)$$
 (1.15)

$$= \int_{-\infty}^{0} 0 dx + \int_{0}^{1} x dx + \int_{1}^{\infty} 0 dx \qquad (1.16)$$

$$=\frac{1}{2}$$
 (1.17)

$$E[U^2] = \int_{-\infty}^{\infty} x^2 dF_U(x) \tag{1.18}$$

$$= \int_{-\infty}^{0} 0 dx + \int_{0}^{1} x^{2} dx + \int_{1}^{\infty} 0 dx \quad (1.19)$$

$$=\frac{1}{3}$$
 (1.20)

$$E[U - E[U]]^{2} = E[U^{2}] - [E[U]]^{2}$$
 (1.21)

From Equation (1.17) and (1.20),

$$E[U - E[U]]^2 = \frac{1}{3} - \left(\frac{1}{2}\right)^2$$
 (1.22)

$$=\frac{1}{3}-\frac{1}{4}\tag{1.23}$$

$$= \frac{1}{12} \approx 0.083 \tag{1.24}$$

#### 2 Central Limit Theorem

2.1 Generate 10<sup>6</sup> samples of the random variable

$$X = \sum_{i=1}^{12} U_i - 6 \tag{2.1}$$

using a C program, where  $U_i$ , i = 1, 2, ..., 12 are a set of independent uniform random variables between 0 and 1 and save in a file called gau.dat

#### **Solution:**

Download the following files,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution/1/ coeffs.h

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /2/2 1.c Execute the above files using code,

2.2 Load gau.dat in python and plot the empirical CDF of *X* using the samples in gau.dat. What properties does a CDF have?

#### **Solution:**

Download the following files,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /2/2\_2.py

Execute the above file using code,

Plot 2.1 obtained is symmetric about (0,0.5)

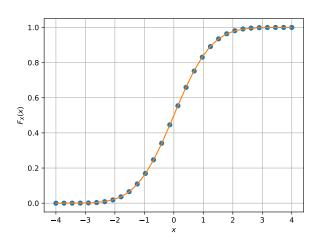


Fig. 2.1

2.3 Load gau.dat in python and plot the empirical PDF of *X* using the samples in gau.dat. The PDF of *X* is defined as

$$p_X(x) = \frac{d}{dx} F_X(x) \tag{2.2}$$

What properties does the PDF have?

#### **Solution:**

Download the following files,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /2/2\_3.py

Execute the above file using code,

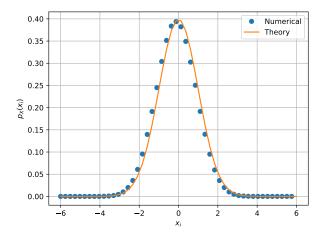


Fig. 2.2

Plot 2.2 obtained is symmetric about y-axis 2.4 Find the mean and variance of *X* by writing a C program.

#### **Solution:**

Download the following files,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/soltion/1/ coeffs.h

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /2/2\_4.c

Execute the above files using code,

The result is,

$$E[U] = 0.000294$$
 (2.3)

$$var[U] = 0.999561$$
 (2.4)

2.5 Given that

$$p_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right), -\infty < x < \infty, \quad (2.5)$$

repeat the above exercise theoretically.

**Solution:** 

$$E[U] = \int_{-\infty}^{\infty} u p_X(u) du \qquad (2.6)$$

 $up_X(u)$  is a odd function.

$$E[U] = 0 \tag{2.8}$$

(2.9)

$$E\left[U^{2}\right] = \int_{-\infty}^{\infty} \frac{u^{2}}{\sqrt{2\pi}} \exp\left(-\frac{u^{2}}{2}\right) du \qquad (2.10)$$
$$= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} u \left(u \exp\left(-\frac{u^{2}}{2}\right)\right) du \quad (2.11)$$

Using integration by parts,

$$E\left[U^{2}\right] = -u\frac{1}{\sqrt{2\pi}}\exp\left(-\frac{u^{2}}{2}\right)\Big|_{-\infty}^{\infty} + \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}}\exp\left(-\frac{u^{2}}{2}\right)$$
(2.12)

$$= 0 + \frac{1}{\sqrt{2\pi}} \sqrt{2\pi} \tag{2.13}$$

$$= 1 \tag{2.14}$$

Hence,

$$var[U] = E[U^{2}] - [E[U]]^{2}$$
 (2.15)  
= 1 (2.16)

#### 3 From Uniform to Other

#### 3.1 Generate samples of

$$V = -2\ln(1 - U) \tag{3.1}$$

and plot its CDF.

#### **Solution:**

Download the following files,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution/ coeffs.h

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /3/3\_1.c

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /3/3 1.py

Execute the above files using code,

#### Plot 3.1 is obtained

3.2 Find a theoretical expression for  $F_V(x)$ .

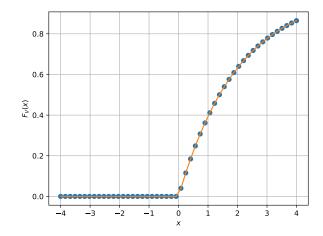


Fig. 3.1

#### **Solution:**

$$F_V(x) = Pr(V \le x) \tag{3.2}$$

$$= Pr(-2\ln(1-U) \le x) \tag{3.3}$$

$$= Pr\left(U \le 1 - e^{-\frac{x}{2}}\right) \tag{3.4}$$

$$F_V(x) = F_U \left( 1 - e^{-\frac{x}{2}} \right) \tag{3.5}$$

$$= \begin{cases} 0, & 1 - e^{-\frac{x}{2}} < 0\\ 1 - e^{-\frac{x}{2}}, & 1 - e^{-\frac{x}{2}} \in [0, 1]\\ 1, & 1 - e^{-\frac{x}{2}} > 1 \end{cases}$$
 (3.6)

Now,

$$1 - e^{-\frac{x}{2}} < 0 \tag{3.7}$$

$$\implies x < 0$$
 (3.8)

$$0 \le 1 - e^{-\frac{x}{2}} \le 1 \tag{3.9}$$

$$\implies x \ge 0$$
 (3.10)

$$1 < 1 - e^{-\frac{x}{2}} \tag{3.11}$$

$$\implies x \in \phi$$
 (3.12)

Hence,

$$F_V(x) = \begin{cases} 0, & x < 0 \\ 1 - e^{-\frac{x}{2}}, & x \ge 0 \end{cases}$$
 (3.13)

#### 4 Triangular Distribution

#### 4.1 Generate

$$T = U_1 + U_2 (4.1)$$

#### **Solution:**

Download the following files,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution/ coeffs.h

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /4/4\_1.c

Execute the above files using code,

4.2 Find the CDF of T.

#### **Solution:**

Download the following files,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /4/4 2.py

Execute the above files using code,

Plot 4.1 is obtained

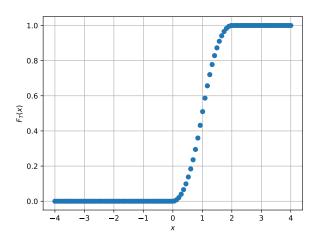


Fig. 4.1

#### 4.3 Find the PDF of T.

#### **Solution:**

Download the following files,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /4/4\_3.py

Execute the above files using code,

#### python3 4\_3.py

Plot 4.2 is obtained

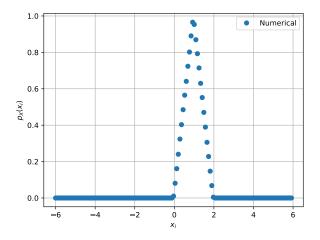


Fig. 4.2

#### 4.4 Find the theoretical PDF and CDF of T.

#### **Solution:**

The CDF of T is given by

$$F_T(t) = \Pr(T \le t) = \Pr(U_1 + U_2 \le t)$$
 (4.2)

Since  $U_1, U_2 \in [0, 1] \implies U_1 + U_2 \in [0, 2]$ Therefore, if  $t \ge 2$ , then  $U_1 + U_2 \le t$  is always true and if t < 0, then  $U_1 + U_2 \le t$  is always false.

Now, fix the value of  $U_1$  to be some x

$$x + U_2 \le t \implies U_2 \le t - x \tag{4.3}$$

If  $0 \le t \le 1$ , then x can take all values in [0, t]

$$F_T(t) = \int_0^t \Pr(U_2 \le t - x) \, p_{U_1}(x) \mathrm{d}x \qquad (4.4)$$

$$= \int_0^t F_{U_2}(t-x)p_{U_1}(x)\mathrm{d}x \tag{4.5}$$

$$0 \le x \le t \implies 0 \le t - x \le t \le 1$$
 (4.6)

$$\implies F_{U_2}(t-x) = t-x$$
 (4.7)

$$F_T(t) = \int_0^t (t - x) \cdot 1 \cdot \mathrm{d}x \tag{4.8}$$

$$= tx - \frac{x^2}{2} \bigg|_0^t \tag{4.9}$$

$$=\frac{t^2}{2}$$
 (4.10)

If 1 < t < 2, x can only take values in [0, 1] as  $U_1 \le 1$ 

$$F_T(t) = \int_0^1 F_{U_2}(t - x) \cdot 1 \cdot dx$$
 (4.11)

$$0 \le x \le t - 1 \implies 1 \le t - x \le t$$

$$t - 1 \le x \le 1 \implies 0 < t - 1 \le t - x \le 1$$

$$(4.12)$$

$$(4.13)$$

$$F_T(t) = \int_0^{t-1} 1 dx + \int_{t-1}^1 (t - x) dx$$
 (4.14)  

$$= t - 1 + t(1 - (t - 1)) - \frac{1}{2} + \frac{(t - 1)^2}{2}$$
 (4.15)  

$$= t - 1 + 2t - t^2 - \frac{1}{2} + \frac{t^2}{2} + \frac{1}{2} - t$$
 (4.16)  

$$= -\frac{t^2}{2} + 2t - 1$$
 (4.17)

Therefore,

$$F_T(t) = \begin{cases} 0 & t < 0 \\ \frac{t^2}{2} & 0 \le t \le 1 \\ 2t - \frac{t^2}{2} - 1 & 1 < t < 2 \\ 1 & t \ge 2 \end{cases}$$
(4.18)

The PDF of T is given by

$$p_T(t) = \frac{\mathrm{d}}{\mathrm{d}t} F_T(t) \tag{4.19}$$

$$\therefore p_T(t) = \begin{cases} 0 & t < 0 \\ t & 0 \le t \le 1 \\ 2 - t & 1 < t < 2 \\ 0 & t \ge 2 \end{cases}$$
 (4.20)

4.5 Verify your results through a plot.

#### **Solution:**

Download the following files,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /4/4 5 1.py

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /4/4 5 2.py

Execute the above files using code,

### python3 4 5 2.py

Plot 4.3 is obtained is pdf Plot 4.4 is obtained

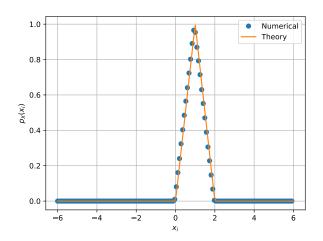


Fig. 4.3

is cdf

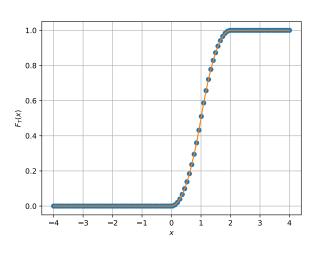


Fig. 4.4

#### 5 Maximum Likelihood

5.1 Generate equiprobable  $X \in \{1, -1\}$ .

#### **Solution:**

Download the following files,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /5/5.1.c

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution/ coeffs.h Execute the above files using code,

#### 5.2 Generate

$$Y = AX + N \tag{5.1}$$

where A = 5 dB,  $X \in \{1, -1\}$  is Bernoulli and  $N \sim \mathcal{N}(0, 1)$ .

#### **Solution:**

Download the following files,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /5/5.2.c

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution/ coeffs.h

Execute the above files using code,

5.3 Plot Y using a scatter plot.

#### **Solution:**

Download the following files,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /5/5.3.py

Execute the above files using code,

python3 5.3.py

Plot 5.1 is obtained

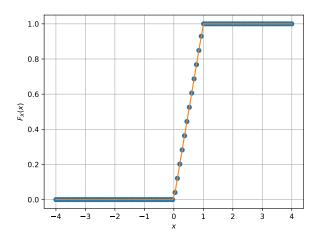


Fig. 5.1

5.4 Guess how to estimate *X* from *Y*.

#### **Solution:**

From the plot of Y, we see that the estimate model can be written as

$$\hat{X} = \begin{cases} 1 & Y > 0 \\ 0 & Y < 0 \end{cases}$$
 (5.2)

5.5 Find

$$P_{e|0} = \Pr(\hat{X} = -1|X = 1)$$
 (5.3)

and

$$P_{e|1} = \Pr(\hat{X} = 1|X = -1)$$
 (5.4)

#### **Solution:**

Letting X = 1 and X = -1 respectively, we see the number of mismatched data points to compute the error probabilities. The simulation is coded in

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /5/5.5.py

and can be run by typing

python3 5.5.py

The results are

$$P_{e|0} = 5.02 \times 10^{-4} \tag{5.5}$$

$$P_{e|1} = 5.24 \times 10^{-4} \tag{5.6}$$

5.6 Find  $P_e$  assuming that X has equiprobable symbols.

#### **Solution:**

Here, Pr(X = 1) = Pr(X = -1) = 0.5. Thus,

$$P_e = \Pr(X = 1) P_{e|1} + \Pr(X = -1) P_{e|0}$$
 (5.7)

$$= \frac{1}{2} \left( P_{e|0} + P_{e|1} \right) = 5.13 \times 10^{-4} \tag{5.8}$$

5.7 Verify by plotting the theoretical  $P_e$  wrt A from 0 dB to 10 dB.

#### **Solution:**

$$P_{e|0} = \Pr(\hat{X} = 1|X = -1)$$
 (5.9)

$$= \Pr(Y > 0 | X = -1) \tag{5.10}$$

$$= \Pr(AX + N > 0 | X = -1) \tag{5.11}$$

$$= \Pr(N > A) = Q(A)$$
 (5.12)

since X and N are independent. Writing a

similar expression for  $P_{e|1}$  and noting that

$$Pr(N < -A) = Pr(N > A) = Q(A)$$
 (5.13)

it follows that  $P_e = Q(A)$ . This is the idea used to plot the theoretical  $P_e$ . The plot is coded both in the rectangular axes and the semilog-y axes. Download the relevant codes using

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /5/5.7.py

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /5/5.7\_semilog.py

Execute the above files using code,

Plot 5.2 obtained in rectangular axis

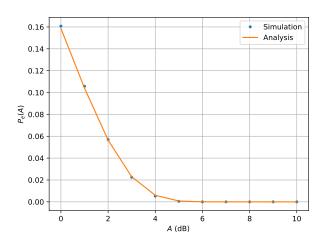


Fig. 5.2

Plot 5.3 obtained in semilog-y axis

5.8 Now, consider a threshold  $\delta$  while estimating X from Y. Find the value of  $\delta$  that maximizes the theoretical  $P_e$ .

#### **Solution:**

To estimate X from Y, we now consider the following:

$$X = \begin{cases} 1, & Y > \delta \\ -1, & Y < \delta \end{cases} \tag{5.14}$$

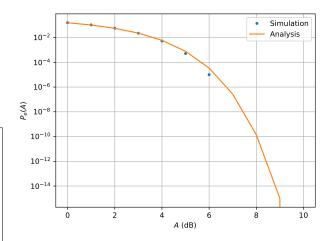


Fig. 5.3

Therefore,

$$P_{e} = \Pr(X = -1) Q (A + \delta)$$

$$+ \Pr(X = 1) Q (A - \delta)$$

$$= \frac{1}{2} (Q (A + \delta) + Q (A - \delta))$$
(5.15)

To minimise  $P_e$ , we differentiate the above equation wrt  $\delta$ :

$$0 = \frac{d}{d\delta} \left( \frac{1}{2} (Q_N(A - \delta) + Q_N(A + \delta)) \right)$$

$$= \frac{1}{2} \left( \frac{1}{\sqrt{2\pi}} e^{-\frac{(\delta - A)^2}{2}} - \frac{1}{\sqrt{2\pi}} e^{-\frac{(A + \delta)^2}{2}} \right)$$
(5.18)

Therefore,

$$(\delta - A)^2 = (\delta + A)^2$$

$$\implies \delta = 0$$
(5.19)
$$(5.20)$$

5.9 Repeat the above exercise when

$$p_{x}(0) = p$$

**Solution:** 

$$P_e = P_{e|0}p + P_{e|1}(1-p)$$

$$= pQ_N(A-\delta) + (1-p)Q_N(A+\delta)$$
 (5.22)

Differentiating as before, we get:

$$0 = p \frac{1}{\sqrt{2\pi}} e^{-\frac{(\delta - A)^2}{2}} - (1 - p) \frac{1}{\sqrt{2\pi}} e^{-\frac{(A + \delta)^2}{2}}$$
(5.23)

Taking In on both sides we have:

$$\ln p - \frac{(\delta - A)^2}{2} = \ln 1 - p - \frac{(\delta + A)^2}{2} \quad (5.24)$$

$$\implies 2\delta A = \ln \frac{1-p}{p} \tag{5.25}$$

$$\implies \delta = \frac{1}{2A} \ln \frac{1-p}{p} \tag{5.26}$$

5.10 Repeat the above exercise using the MAP criterion.

#### **Solution:**

Using theorem of conditional probabilities,

$$p_{X|Y}(x|y) = \frac{p_{Y|X}(y|x) \times p_X(x)}{p_Y(y)}$$
 (5.27)

$$p_{X|Y}(x|y) = \frac{p_X(x)Pr(Ax + N = y)}{p_Y(y)}$$
 (5.28)

$$p_{X|Y}(x|y) = \frac{p_X(x)Pr(N = y - Ax)}{p_Y(y)}$$
 (5.29)

$$p_{X|Y}(x|y) = \frac{p_Y(y)}{p_Y(y)}$$

$$p_{X|Y}(x|y) = \frac{p_X(x)p_N(y - Ax)}{p_Y(y)}$$
(5.29)

$$p_{X|Y}(x|y) = \frac{\frac{1}{\sqrt{2\pi}}p_X(x)e^{-\frac{(y-Ax)^2}{2}}}{p_Y(y)}$$
 (5.31)

(5.32)

When X = 1, we have:

$$p_{X|Y}(1|y) = \frac{p_{Y|X}(y|1) \times p_X(1)}{p_Y(y)}$$
 (5.33)

$$= \frac{(1-p)\frac{e^{-\frac{(y-A)^2}{2}}}{\sqrt{2\pi}}}{p\frac{e^{-\frac{(y+A)^2}{2}}}{\sqrt{2\pi}} + (1-p)\frac{e^{-\frac{(y-A)^2}{2}}}{\sqrt{2\pi}}}$$
(5.34)

$$= \frac{(1-p)e^{2yA}}{p+(1-p)e^{2yA}}$$
 (5.35)

Similarly when X = -1,

$$p_{X|Y}(-1|y) = \frac{p}{p + (1-p)e^{2yA}}$$
 (5.36)

Therefore, when  $p_{X|Y}(1|y) > p_{X|Y}(-1|y)$ , we

have:

$$\frac{(1-p)e^{2yA}}{p+(1-p)e^{2yA}} > \frac{p}{p+(1-p)e^{2yA}}$$
 (5.37)

$$e^{2yA} > \frac{p}{(1-p)}$$
 (5.38)

$$y > \frac{1}{2A} \ln \frac{p}{(1-p)}$$
 (5.39)

Therefore, when Eq. (5.39), we can assert that X = 1, and X = -1 otherwise.

#### 6 Gaussian to Other

6.1 Let  $X_1 \sim \mathcal{N}(0,1)$  and  $X_2 \sim \mathcal{N}(0,1)$ . Plot the CDF and PDF of

$$V = X_1^2 + X_2^2 \tag{6.1}$$

#### **Solution:**

Applying following transformation on  $X_1$  and

$$X_1 = R\cos\Theta \tag{6.2}$$

$$X_2 = R\sin\Theta \tag{6.3}$$

Where  $R \in [0, \infty), \Theta \in [0, 2\pi)$ . The Jacobian Matrix for this transformation is given by

$$\mathbf{J} = \begin{pmatrix} \frac{\partial X_1}{\partial R} & \frac{\partial X_2}{\partial R} \\ \frac{\partial X_1}{\partial \Theta} & \frac{\partial X_2}{\partial \Theta} \end{pmatrix} \tag{6.4}$$

$$= \begin{pmatrix} \cos \Theta & \sin \Theta \\ -R \sin \Theta & R \cos \Theta \end{pmatrix} \tag{6.5}$$

$$\implies |\mathbf{J}| = R \tag{6.6}$$

We also know that

$$|\mathbf{J}|p_{X_1,X_2}(x_1,x_2) = p_{R,\Theta}(r,\theta)$$
(6.7)

$$\implies p_{R,\Theta}(r,\theta) = Rp_{X_1}(x_1)p_{X_2}(x_2)$$
 (6.8)

$$= \frac{R}{2\pi} \exp\left(-\frac{X_1^2 + X_2^2}{2}\right) \quad (6.9)$$

$$= \frac{R}{2\pi} \exp\left(-\frac{R^2}{2}\right) \tag{6.10}$$

Where (6.8) follows as  $X_1, X_2$  are iid random variables. Thus,

$$p_R(r) = \int_0^{2\pi} p_{R,\Theta}(r,\theta) d\theta \qquad (6.11)$$

$$= R \exp\left(-\frac{R^2}{2}\right) \tag{6.12}$$

However,  $V = X_1^2 + X_2^2 = R^2 \ge 0$ , thus  $F_V(x) = 0$ . For x < 0,

$$F_V(x) = F_R(\sqrt{x}) \tag{6.13}$$

$$= \int_0^{\sqrt{x}} r \exp\left(-\frac{r^2}{2}\right) dr \tag{6.14}$$

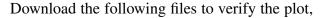
$$= \int_0^{\frac{x}{2}} e^{-t} dt = 1 - e^{-\frac{x}{2}}$$
 (6.15)

Where  $t = \frac{r^2}{2}$  for  $x \ge 0$ , thus

$$F_V(x) = \begin{cases} 1 - e^{-\frac{x}{2}} & x \ge 0\\ 0 & x < 0 \end{cases}$$
 (6.16)

$$p_V(x) = \frac{dF_V(x)}{dx} \tag{6.17}$$

$$= \begin{cases} 0, & x < 0 \\ \frac{1}{2}e^{-\frac{x}{2}}, & x \ge 0 \end{cases}$$
 (6.18)



wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /6/6.1.c

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution/ coeffs.h

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /6/6.1.cdf.py

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /6/6.1.pdf.py

Execute the programs using the code in terminal,

Plot 6.1 is CDF Plot 6.2 is PDF

6.2 If

$$F_V(x) = \begin{cases} 1 - e^{-\alpha x} & x \ge 0\\ 0 & x < 0 \end{cases}$$
 (6.19)

find  $\alpha$ .

#### **Solution:**

From (6.16),  $\alpha = 0.5$ 

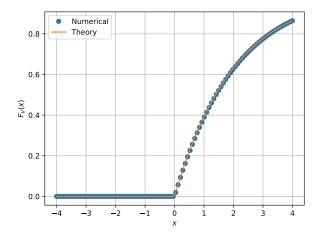


Fig. 6.1

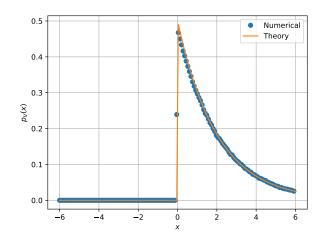


Fig. 6.2

6.3 Plot the CDF and PDF of

$$A = \sqrt{V} \tag{6.20}$$

#### **Solution:**

For  $x \ge 0$ ,

$$F_A(x) = \Pr\left(A \le x\right) \tag{6.21}$$

$$F_A(x) = \Pr\left(\sqrt{V} \le x\right) \tag{6.22}$$

$$F_A(x) = \Pr\left(V \le x^2\right) \tag{6.23}$$

$$F_A(x) = F_V(x^2)$$
 (6.24)

From equation (6.16),

$$F_A(x) = \begin{cases} 0, & x < 0 \\ 1 - e^{x^2/2}, & x \ge 0 \end{cases}$$
 (6.25)

$$p_A(x) = \frac{dF_V(x)}{dx} \tag{6.26}$$

$$= \begin{cases} 0, & x < 0 \\ xe^{-\frac{x^2}{2}}, & x \ge 0 \end{cases}$$
 (6.27)

Download the following files to verify the plot,

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /6/6.3.c

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution/ coeffs.h

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /6/6.3.cdf.py

wget https://github.com/Anshul-Sangrame/ AI1110/blob/main/Assignment/solution /6/6.3.pdf.py

Execute the programs using the code in terminal,

gcc 6.3.c -lm ./a.out python3 6.3.cdf.py python3 6.3.pdf.py

Plot 6.3 is CDF

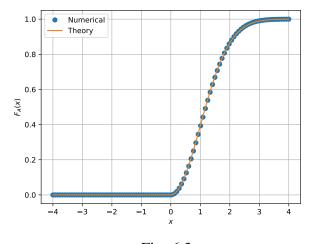


Fig. 6.3

Plot 6.4 is PDF

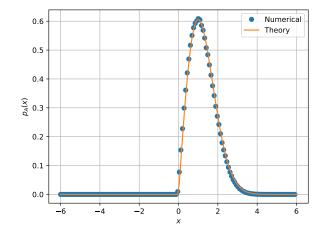


Fig. 6.4

#### 7 CONDITIONAL PROBABILITY

7.1 Plot

$$P_e = \Pr\left(\hat{X} = -1|X = 1\right) \tag{7.1}$$

for

$$Y = AX + N \tag{7.2}$$

where A is Rayleigh with  $E[A^2] = \gamma$ ,  $N \sim \mathcal{N}(0,1)$ ,  $X \in \{1,-1\}$  for  $0 \le \gamma \le 10$  dB.

#### **Solution:**

7.2 Assuming that N is a constant, find an expression for  $P_e$ . Call this  $P_e(N)$ .

#### **Solution:**

We rewrite the previous expression for  $P_e$  as

$$P_e(N) = \Pr(Y < 0|X = 1)$$
 (7.3)

$$= \Pr(A < -N) = F_A(-N)$$
 (7.4)

$$= \begin{cases} 1 - e^{-\frac{N^2}{\gamma}} & N \le 0\\ 0 & N > 0 \end{cases}$$
 (7.5)

7.3 For a function g,

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)p_X(x)dx \qquad (7.6)$$

Find  $P_e = E[P_e(N)]$ .

**Solution:** 

We write,

$$P_{e} = \int_{0}^{\infty} F_{A}(x) f_{N}(x) dx$$
 (7.7)  

$$= \int_{0}^{\infty} (1 - e^{-\frac{x^{2}}{\gamma}}) \frac{1}{\sqrt{2\pi}} e^{-\frac{x^{2}}{2}} dx$$
 (7.8)  

$$= \frac{1}{2} - \frac{1}{\sqrt{2\pi}} \int_{0}^{\infty} \exp\left(-x^{2} \left\{\frac{\gamma + 2}{2\gamma}\right\}\right) dx$$
 (7.9)  

$$= \frac{1}{2} \left(1 - \sqrt{\frac{\gamma}{\gamma + 2}}\right)$$
 (7.10)

Where  $f_N$  denotes PDF of standard normal distribution.