

# Special Probability Distributions

## DISCRETE

Discrete uniform

Binomial

Geometric

Negative binomial

Hypergeometric

Poisson

Multinomial

## CONTINUOUS

Uniform

Exponential

Gamma

Weibull

Pareto

Beta

Normal

Cauchy

Log-Normal

Reliability of Series & Parallel systems.

## PROBLEMS

6 lectures

## Rectangular or Uniform distribution

A RV is said to have continuous rectangular or uniform distribution over the interval  $(a, b)$   $(-\infty < a < b < \infty)$  if pdf is given by

$$f(x) = \begin{cases} \frac{1}{b-a} & \text{if } a < x < b \\ 0 & \text{otherwise} \end{cases}$$

→  $a$  &  $b$  are parameters

→ is written as  $X \sim U(a, b)$

The cumulative distribution fcn is .

$$F(x) = \begin{cases} 0 & \text{if } x \leq a \\ \frac{x-a}{b-a} & \text{if } a < x < b \\ 1 & \text{if } x \geq b. \end{cases}$$

$F(x)$  is continuous & not diff at  $x=a$  &  $b$

▣ Trains on a certain line run every half hour between midnight and six in the morning. What is the probability that a man entering the station at random time during this period will have to wait at least 20 min?

→  $X \sim U(0, 30)$ ,  $X \equiv$  waiting time

$$\Rightarrow P(X \geq 20) = \int_{20}^{30} f(x) dx = \frac{1}{30} \int_{20}^{30} dx = \frac{1}{3}.$$

$X \sim U(0, 1)$  Find pdf of  $Y = -2 \log X$ .

We know,  $F(y) = P(Y \leq y) = P(-2 \log x \leq y)$

$$= P(\log X \geq -\frac{y}{2})$$

$$= 1 - P(X \leq e^{-y/2})$$

$$= 1 - \int_0^{e^{-y/2}} dx = 1 - e^{-y/2}$$

$$\text{Hence pdf of } Y = \frac{d}{dy} F(y) = \frac{1}{2} e^{-y/2}$$

$$0 < y < \infty.$$

## Exponential dist.

A RV  $X$  is said to have exponential dist. with parameter  $\theta > 0$ , if its pdf is given by

$$f(x, \theta) = \begin{cases} \theta e^{-\theta x} & x \geq 0 \\ 0 & \text{otherwise} \end{cases}$$

cdf is

$$\begin{aligned} F(x) &= \int_0^x f(t) dt = \theta \int_0^x e^{-\theta t} dt \\ &= \begin{cases} 1 - e^{-\theta x} & x \geq 0 \\ 0 & \text{otherwise} \end{cases} \end{aligned}$$

## Moment generating f.

$$\begin{aligned} M_X(t) &= E(e^{tx}) = \theta \int_0^{\infty} e^{tx} \cdot e^{-\theta x} dx \\ &= \frac{\theta}{\theta - t} = \left(1 - \frac{t}{\theta}\right)^{-1} \\ &= \sum_{r=0}^{\infty} \left(\frac{t}{\theta}\right)^r, \quad \theta > t. \end{aligned}$$

## Moments

$$\begin{aligned}\mu_r' &= E(x^r) = \text{coefficient of } \frac{t^r}{r!} \text{ in } M_x(t) \\ &= \frac{r!}{\theta^r} \quad r = 1, 2, \dots\end{aligned}$$

$$\text{Mean} = \mu_1' = \frac{1}{\theta}$$

$$\text{Variance} = \mu_2' - \mu_1'^2 = \frac{2}{\theta^2} - \frac{1}{\theta^2} = \frac{1}{\theta^2}$$

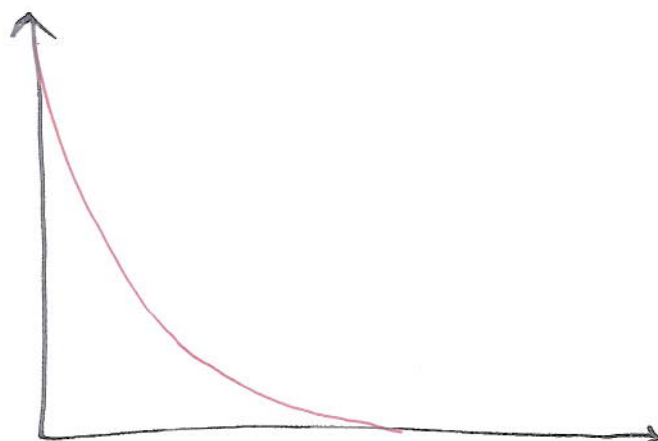
Note      Variance  $>$  Mean      if       $0 < \theta < 1$   
                  Variance  $=$  Mean      if       $\theta = 1$   
                  Variance  $<$  Mean      if       $\theta > 1$ .

$\text{Variance} = \frac{\text{Mean}}{\theta}$
--

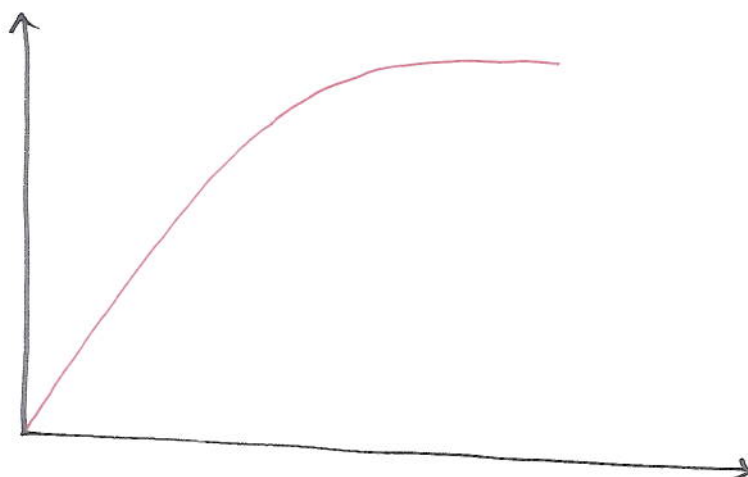
Exponential dist<sup>n</sup> has memory less property.


# Exponential dist.

pdf



cdf



 An electronic component is known to have a useful life represented by an exponential dist<sup>n</sup> with failure rate  $10^{-5}$  failure per hour. (i.e.  $\theta = 10^{-5}$ ). The mean time to failure is  $E(x)$  is thus  $10^5$  hours. Suppose we want to determine the fraction of such components that would fail before the mean life or expected life is

$$P(T < \frac{1}{\theta}) = \int_0^{\frac{1}{\theta}} \theta e^{-\theta x} dx = 1 - \frac{1}{e} = .63.$$



## Memoryless property of Exponential Distr.

$$P(X > s+t \mid X > t) = P(X > s)$$

$\Downarrow$   
 prob. that an item that  
 is functioning at time  
 t will continue to work  
 an additional time s.

$\Downarrow$   
 Prob. that item is  
 functioning s time  
 unit.

$\Rightarrow$  No need to remember the age of  
 functioning item As  $\downarrow$

as long as it is  
 functioning it is  
 as good as 'new'.

$$\begin{aligned}
 \underline{\text{LHS}} &= \frac{P(X > s+t, X > t)}{P(X > t)} \\
 &= \frac{P(X > s+t)}{P(X > t)} \\
 &= \frac{e^{-\lambda(s+t)}}{e^{-\lambda t}} = e^{-\lambda s} = P(X > s)
 \end{aligned}$$





Suppose that an amount of time one spends in a bank is exponentially distributed with mean 10 minutes.

(i) What is the prob. that a customer will spend more than 15 minutes in the bank?

(ii) What is the probability that the customer will spend more than 15 mins. in the bank given that he is still in the bank after 10 mins.

Ans.

$$(i) P(X > 15) = e^{-15\lambda} = e^{-3/2} = .2$$

$$\lambda = \frac{1}{10}.$$

$$(ii) P(X > 15 / X > 10)$$

$$= P(X > 5)$$

$$= e^{-5\lambda} = .6.$$

## Normal Dist.

41

$$X \sim N(\mu, \sigma^2) \quad \text{or} \quad N(\mu, \sigma)$$

$$\Rightarrow \text{pdf} \quad f(x) = \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{1}{2} \left( \frac{x-\mu}{\sigma} \right)^2}$$
$$-\infty < x < \infty$$

## Standard Normal variate

$$Z \sim N(0, 1)$$

$$\text{Where, } Z = \frac{X - \mu}{\sigma} \quad \& \quad X \sim N(\mu, \sigma)$$

$$\text{pdf} \quad \phi(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}, \quad -\infty < z < \infty$$

$$\text{cdf} \quad \Phi(z) = P(Z \leq z) = \int_{-\infty}^z \phi(z) dz$$

## Properties of $\Phi$

$$\rightarrow \Phi(-z) = 1 - \Phi(z)$$

$$\rightarrow P(a \leq \underset{\substack{\downarrow \\ \text{Normal}}}{X} \leq b) = \Phi\left(\frac{b-\mu}{\sigma}\right) - \Phi\left(\frac{a-\mu}{\sigma}\right)$$

## Properties of Normal distn.

42

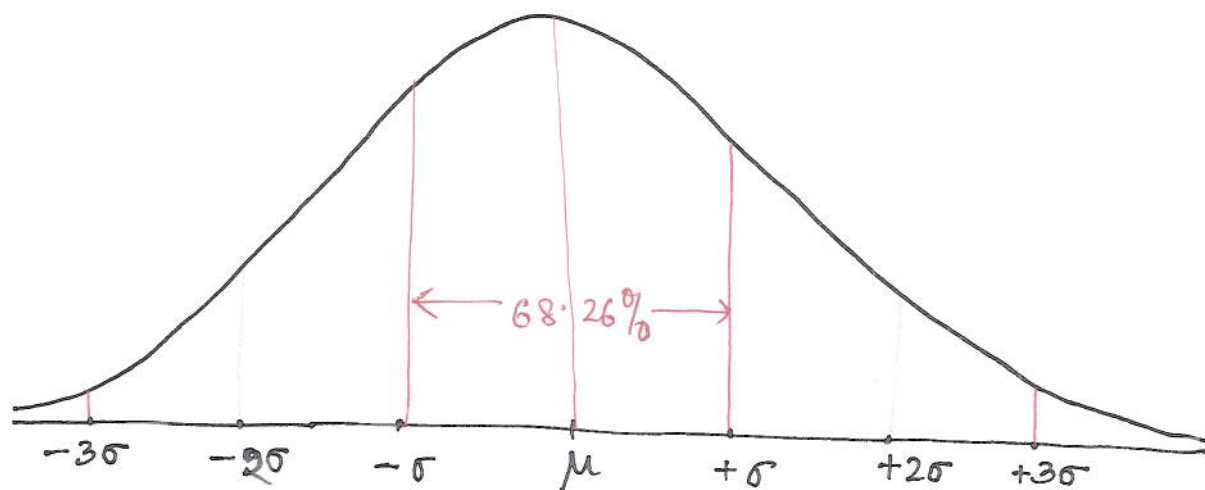
- The curve is bell-shaped and symmetrical about  $x = \mu$
- Mean is at the middle and divides the area into halves
- It is completely determined by  $\mu$  &  $\sigma$
- Mean, Mode and Median of distn. coincide
- Skewness = 0, Kurtosis =  $\frac{\mu_4}{\sigma^4} = \frac{E(x-\mu)^4}{[E(x-\mu)^2]^2} = 3$
- Points of inflexion are at  $x = \mu + \sigma$  &  $\mu - \sigma$
- Area property:

$$P(\mu - \sigma \leq X \leq \mu + \sigma) = 0.6826$$

$$P(\mu - 2\sigma \leq X \leq \mu + 2\sigma) = 0.9544$$

$$P(\mu - 3\sigma \leq X \leq \mu + 3\sigma) = 0.9972$$

- Total area under the curve is 1.



# Moment generating function

$$M_X(t) = \int_{-\infty}^{\infty} e^{tx} f(x) dx$$

$$= \int_{-\infty}^{\infty} e^{tx} \cdot \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx.$$

$$= \frac{1}{\sqrt{2\pi}} e^{\mu t} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(z^2 - 2tz\sigma)} dz$$

take  $z = \frac{x-\mu}{\sigma}$ .

$$= \frac{1}{\sqrt{2\pi}} e^{\mu t + \frac{\sigma^2 t^2}{2}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(z - \sigma t)^2} dz$$

$$= \left( \frac{1}{\sqrt{2\pi}} \right) e^{\mu t + \frac{\sigma^2 t^2}{2}}$$

$$\int_{-\infty}^{\infty} e^{-\frac{u^2}{2}} du$$

$$\text{mgf} = e^{\mu t + \frac{\sigma^2 t^2}{2}}$$

▣  $X \sim N(70, 4)$ . Find

$$(a) P(68 \leq X \leq 74) = .817$$

$$(b) P(72 \leq X \leq 75) = .1525$$

$$(c) P(63 \leq X \leq 68) = .1584$$

$$(d) P(X \geq 73) = .0668.$$

$$\begin{aligned} (a) P(68 \leq X \leq 74) &= P\left(\frac{68-70}{2} \leq \frac{X-70}{2} \leq \frac{74-70}{2}\right) \\ &= P(-1 \leq Z \leq 2) \\ &= \Phi(2) - \Phi(-1) \\ &= \Phi(2) + \Phi(1) - 1 \\ &= .97725 + .814134 - 1 \\ &= .8185 \end{aligned}$$

$$\begin{aligned} (b) P(72 \leq X \leq 75) &= P\left(\frac{72-70}{2} \leq \frac{X-70}{2} \leq \frac{75-70}{2}\right) \\ &= P(1 \leq Z \leq 2.5) \end{aligned}$$



$$f(x) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2}(x-\mu)^2}$$

$$f'(x) = -\frac{f(x)}{\sigma^2} (x-\mu)$$

$$f''(x) = -\frac{f(x)}{\sigma^2} \left[ 1 - \frac{(x-\mu)^2}{\sigma^2} \right]$$

At  $x = \mu$ ,  $f'(x) = 0$ ,  $f''(x) < 0$   
 $\Rightarrow$  Mode

At  $x = \mu \pm \sigma$ ,  $f''(x) = 0$ ,  $f'''(x) \neq 0$

$\Rightarrow$  Point of inflexion

▣ Let  $x_i \sim N(\mu_i, \sigma_i^2)$ , then for ~~an~~  $n$  independent  $x_i$ 's

$$\sum_{i=1}^n a_i x_i \sim N \left[ \sum_{i=1}^n a_i \mu_i, \sum_{i=1}^n a_i^2 \sigma_i^2 \right]$$

▣  $\bar{X} \sim N(\mu, \sigma^2/n)$ , where  $\bar{X} = \frac{1}{n} \sum x_i$   
 $\& x_i \sim N(\mu, \sigma^2)$



Proof:

$$\text{As } M_{x_i}(t) = e^{\mu_i t + \frac{1}{2} \sigma_i^2 t^2}$$

46

$$M_{\sum a_i x_i}(t) = \prod_{i=1}^n M_{a_i x_i}(t)$$

$$= M_{x_1}(a_1 t) M_{x_2}(a_2 t) \dots M_{x_n}(a_n t)$$

$$\left[ \text{as, } M_{cx}(t) = M_x(ct) \right]$$

$$= \prod_{i=1}^n e^{(a_i \mu_i t + \frac{1}{2} a_i^2 \sigma_i^2 t^2)}$$

$$= e^{(a_1 \mu_1 + a_2 \mu_2 + \dots + a_n \mu_n)}$$

X

$$e^{\frac{1}{2} t^2 (a_1^2 \sigma_1^2 + a_2^2 \sigma_2^2 + \dots + a_n^2 \sigma_n^2)}$$

$$= e^{\mu t + \frac{1}{2} \sigma t^2}$$

47

▣ Given  $X \sim N(12, 4)$ . Find probability of

(a) (i)  $(X \geq 20)$  (ii)  $(X \leq 20)$  (iii)  $\overset{12}{\cancel{12}} \leq X \leq \overset{18}{\cancel{18}}$

$\cdot 00003$                        $\cdot 99997$

(b) Find  $x'$  when  $\text{Prob}(X > x') = .24$

$13.42$

(c) Find  $x'_0$  and  $x'_1$  when  $\text{Prob}(x'_0 < X < x'_1) = .5$

$10.64.$

and  $\text{Prob}(X > x'_1) = .25.$

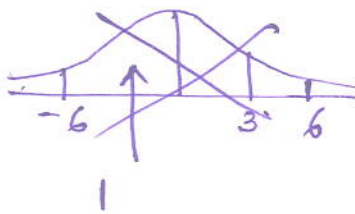
$13.36$

$$\begin{aligned}
 P(X > 20) &= 1 - P(X \leq 20) = 1 - P\left(\frac{X-12}{2} \leq \frac{20-12}{2}\right) \\
 &= 1 - P(Z \leq 4) \\
 &= 1 - .99997 = \boxed{.00003}
 \end{aligned}$$

48

$$P(X \leq 20) = \boxed{.99997}$$

$$\begin{aligned}
 P\left(\frac{12}{2} \leq X \leq 18\right) &= P\left(\frac{12-12}{2} \leq \frac{X-12}{2} \leq \frac{18-12}{2}\right) = P(-0 \leq Z \leq 3) \\
 &= .99865 - .5 \\
 &= .49865
 \end{aligned}$$



$$\begin{aligned}
 &= .99865 - .5 \\
 &= .49865
 \end{aligned}$$

$$\begin{aligned}
 P(X > x') &= .24 \Rightarrow P(X < x') = 1 - .24 = .76 \\
 &\Rightarrow P\left(Z < \frac{x'-12}{2}\right) = .76
 \end{aligned}$$

$$\Rightarrow \frac{x'-12}{2} = .71$$

$$\text{or } x' = 12 + 1.42 = \boxed{13.42}$$

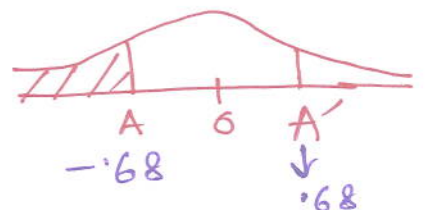
$$P(X > x'_1) = .25 \Rightarrow 1 - P(X < x'_1) = .25$$

$$\Rightarrow P(X < x'_1) = .75 \Rightarrow P\left(Z < \frac{x'_1-12}{2}\right) = .75$$

$$\Rightarrow \frac{x'_1-12}{2} = .68 \text{ or } \boxed{x'_1 = 13.36}$$

$$P(x'_0 < X < 13.36) = .5$$


$$P\left(\frac{x'_0-12}{2} < \frac{X-12}{2} \leq \frac{13.36-12}{2}\right) = .5$$



$$\text{or } P\left(\frac{x'_0-12}{2} < \frac{X-12}{2} \leq \frac{13.36-12}{2}\right) = .5$$

$$\text{or } .75 - \phi(\cdot) = .5 \text{ or } \phi\left(\frac{x'_0-12}{2}\right) = .25$$

$$\text{or } x'_0 - 12 = -1.36 \text{ or } \boxed{x'_0 = 10.64}$$

  $X$  is a normal variate with mean 30 and s.d. <sup>49</sup> 5.  
Find the probability that

(i)  $(26 \leq X \leq 40)$       (ii)  $(X \geq 45)$       (iii)  $|X - 30| > 5$ .

.7653

.00135

.68268

↓

$1 - P(25 \leq X \leq 35)$

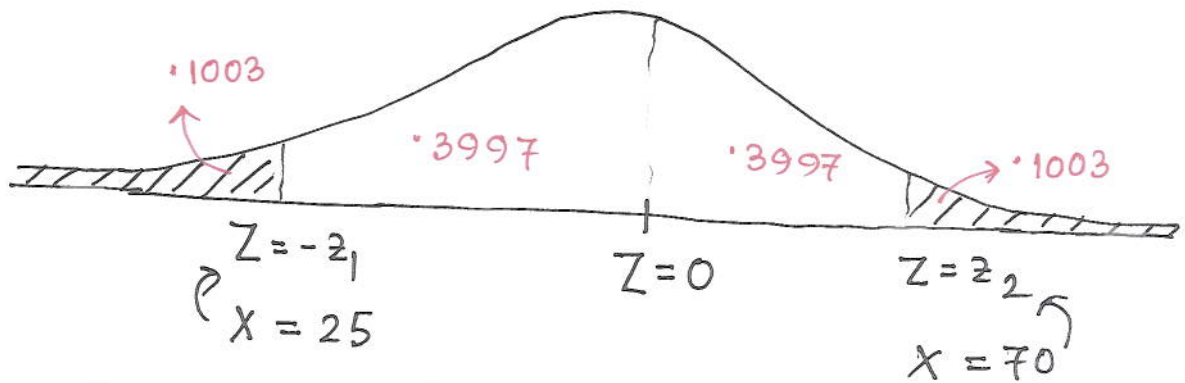
$$\begin{aligned}
 \text{(i)} \quad P(26 \leq X \leq 40) &= P\left(\frac{26-30}{5} \leq \frac{X-30}{5} \leq \frac{40-30}{5}\right) \\
 &= P(-.8 \leq Z \leq 2) = P(2) - P(-.8) = P(2) + P(.8) - 1 \\
 &= .97725 + .78814 - 1 \\
 &= \boxed{.76539}
 \end{aligned}$$

$$\begin{aligned}
 \text{(ii)} \quad P(X \geq 45) &= 1 - P(X \leq 45) = 1 - P\left(\frac{X-30}{5} \leq \frac{45-30}{5}\right) \\
 &= 1 - P(Z \leq 3) = 1 - .99865 \\
 &= \boxed{.00135}
 \end{aligned}$$

$$\begin{aligned}
 \text{(iii)} \quad P(|X-30| > 5) &= P(25 \leq X \leq 35) \\
 &= P(-1 \leq Z \leq 1) = \cancel{.97725} - \cancel{.24205} \\
 &= 2P(Z \leq 1) - 1 = 2 \times .84134 - 1 \\
 &= \boxed{.68268}
 \end{aligned}$$

// In a distribution, which is given normal,  $10.03\%$ <sup>51</sup> of the items are under 25 kg weight and  $89.97\%$  items are under 70 kg weight. What are the mean and s.d. of the distribution?

$$[ P(0 < Z < 1.28) = .3997 ]$$



$$\mu = 47.5 \quad \sigma = 17.578$$

$$\frac{25 - \mu}{\sigma} = \cancel{.0003} \quad 1.28$$

$$\frac{70 - \mu}{\sigma} = 1.28$$



## Mean deviation about mean

52

$$= \int_{-\infty}^{\infty} |x - \mu| f(x) dx = \frac{1}{\sqrt{2\pi} \sigma} \int_{-\infty}^{\infty} |x - \mu| e^{-\frac{1}{2} \left( \frac{x - \mu}{\sigma} \right)^2} dx$$

Take  $\frac{x - \mu}{\sigma} = z$

$$= \frac{\sigma}{\sqrt{2\pi}} \int_{-\infty}^{\infty} |z| e^{-z^2/2} dz$$

$$= \frac{2\sigma}{\sqrt{2\pi}} \int_0^{\infty} |z| e^{-z^2/2} dz$$

Take  $\frac{z^2}{2} = t$

$$= \frac{2\sigma}{\sqrt{2\pi}} \int_0^{\infty} e^{-t} dt$$

$$= \sqrt{\frac{2}{\pi}} \sigma \left[ -e^{-t} \right]_0^{\infty} = \sqrt{\frac{2}{\pi}} \sigma = \frac{4}{5} \sigma \text{ (Approx.)}$$

▣ Two independent RVs  $X$  and  $Y$  are both normally distributed with means 1 & 2 and s.d. 3 & 4 resp. If  $Z = X - Y$ , write the pdf of  $Z$ . And Find mean, median and s.d. of  $Z$ . Also find  $P(Z+1 \leq 0)$ .

$$Z = X - Y \sim N(1 - 2, 9 + 16)$$
$$f(Z) = \frac{1}{\sqrt{2\pi} \times 5} e^{-\frac{1}{2} \left( \frac{Z+1}{5} \right)^2}$$

$$\Rightarrow \text{Mean} = \text{Median} = -1 \text{ \& s.d. } 5.$$

$$P(Z \leq -1) = 0.5.$$

## Log-normal distribution

53

The positive RV  $X$  is said to have a log-normal distribution if  $\log_e X$  is normally distributed.

$$Y = \log_e X \sim N(\mu, \sigma^2) \text{ for } x > 0$$

OR

If  $X \sim N(\mu, \sigma^2)$ ,  $Y = e^X$  is log-normal.

$$F_X(x) = P(X \leq x) = P(\log_e X \leq \log_e x)$$

$$= P(Y \leq \log_e x)$$

$$= \frac{1}{\sqrt{2\pi} \sigma} \int_{-\infty}^{\log_e x} e^{-\frac{1}{2} \left( \frac{y - \mu}{\sigma} \right)^2} dy$$

$$= \frac{1}{\sqrt{2\pi} \sigma} \int_0^x e^{-\frac{1}{2} \left( \frac{\log u - \mu}{\sigma} \right)^2} \frac{du}{u} \left[ y = \log u \right]$$

$$\Rightarrow f(x) = \frac{1}{x \sqrt{2\pi} \sigma} e^{-\frac{1}{2} \left( \frac{\log x - \mu}{\sigma} \right)^2}, \quad x > 0.$$

$$\underline{r^{\text{th}} \text{ moment}} = \mu_r' = E(X^r) = E(e^{rY})$$

(since  $Y = \log X$ , or  $X = e^Y$ )

As  $Y \sim N(\mu, \sigma^2)$ ,  $X$  is log-normal

$$E(e^{rY}) = \text{mgf of } Y = e^{\mu r + \frac{1}{2} r^2 \sigma^2}$$

$$\underline{\text{Mean}} = \mu_1' = e^{\mu + \frac{1}{2} \sigma^2}$$

$$\mu_2' = e^{2\mu + 2\sigma^2}$$

$$\begin{aligned} \underline{\text{Variance}} &= \mu_2' - \mu_1'^2 = e^{2(\mu + \sigma^2)} - e^{2\mu + \sigma^2} \\ &= e^{2\mu + \sigma^2} (e^{\sigma^2} - 1) \\ &= \mu_1' (e^{\sigma^2} - 1) \end{aligned}$$