

The Fertility Race Between Technology and Social Norms*

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April 2025

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Abstract

This paper studies fertility as the outcome of a tug-of-war between technological progress and social norms governing the gender division of childcare. We first document that fertility declines more rapidly in economies experiencing rapid structural transformation, even when controlling for income growth, and this trend is more pronounced in societies with rigid social norms. To explain these findings, we develop a quantitative model of childcare bargaining, incorporating the novel feature of endogenously evolving social norms shaped by the collective opinions of all cohorts, which are influenced by past childcare choices. Calibrated to South Korean data, the model reveals that intense social pressure and resistance from older cohorts to adapt hinder equitable childcare adjustments, exacerbating fertility declines and reinforcing rigid norms. We predict a gradual fertility recovery as social norms slowly align with a new equilibrium. Policies promoting gender-equitable childcare could accelerate this transition.

JEL classification: J12, J13, O11

Keywords: Fertility, Structural Change, Social Norm

*We thank Xiao Ma, Uta Schönberg, David Weil, Junjian Yi, and Xiaodong Zhu for helpful comments and suggestions. Shixiong Ye and Lingrui Zhou provided excellent research assistance.

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1. Introduction

The drastic decline in fertility rates below the replacement level poses an imminent challenge for many economies over the coming decades, threatening fiscal sustainability and economic growth (Jones 2022). This challenge is especially conspicuous in nations “getting old before getting rich” where low fertility imperils economic stability and catching-up prospects. Uncovering the causes of low fertility, forecasting its path, and pinpointing effective policy solutions are thus critical questions for researchers and policymakers.

In this paper, we document that countries undergoing rapid structural change experience a more pronounced decline in fertility rates. This relationship is particularly evident in societies with stringent social norms. Specifically, our cross-country regression analysis reveals that a 1-percentage-point increase in the pace of service sector expansion per year is associated with an annual fertility decline of 0.1 children per woman. This correlation is 50% stronger in societies exhibiting above-median cultural rigidity, as measured by the tightness index from Uz (2015). The pattern persists even after accounting for GDP per capita growth and remains consistent when we substitute agricultural sector decline as a measure of structural change.

These facts motivate us to construct a quantitative model in which fertility emerges from the interplay between technological progress and evolving social norms on the gender division of childcare responsibilities. In this model, fertility decisions require mutual agreement of both parents, who negotiate childcare responsibilities and consumption within a bargaining framework with limited commitment, following Doepke and Kindermann (2019). Since parents cannot pre-commit to how consumption will be divided after childbirth, the allocation of childcare duties becomes pivotal in shaping each parent’s welfare and, consequently, their willingness to have children. Parents determine the childcare allocation by balancing the opportunity costs of childcare against the psychic costs of deviating from the prevailing social norm. For example, in a society with traditional gender expectations, assigning more childcare to men might lower costs if women earn higher wages, yet it could provoke social disapproval, adding psy-

chic costs. Thus, social norms create a gap between the actual childcare arrangement and the one that would minimize opportunity costs alone. As a result, these norms significantly influence fertility, labor supply, and gender welfare disparities in the model.

Different from prior studies that treat social norms as exogenously given, our key contribution is to explicitly model how social norm is endogenously formed. Inspired by the insights from the sociology literature (e.g., [Brooks and Bolzendahl 2004](#)), we hypothesize that the prevailing norm reflects a weighted average of the *opinions* expressed by the other members in the society. These opinions, in turn, are formed in a re-evaluation stage and depend on the current economic conditions as well as the childcare practices that these members adopted in the past. Therefore, the evolution of the social norm in the economy stems from two sources. On the one hand, when new cohorts enter the economy and adopt childcare practices that are different from the past, they shift future norms, reflecting *between-cohort effects* (also known as cohort replacement effects). On the other hand, older cohorts re-evaluate the current situation and weigh it against their past practices. As a result, changing economic conditions affect the opinions expressed by older cohorts, reflecting *within-cohort effects* (also known as social structural effects). Furthermore, the influence of the two effects depends on the population weight of the corresponding cohort, which in turn hinges on past fertility choices.

The model suggests that strong social pressure and older cohorts' reluctance to re-evaluate impose rigid constraints on young parents' choices, fostering cultural inertia. If technological advancements necessitate a shift in childcare allocation, this inertia prevents adjustment, increasing the burden on one side of the parents and making childbearing costlier for them. Given that fertility requires both parents' consent, this added cost can sharply reduce fertility rates, a pattern consistent with our empirical evidence. Our model also predicts a non-monotonic fertility trajectory: an initial drop followed by a recovery as norms adapt to economic realities, in line with the arguments in [Feyrer et al. \(2008\)](#) and [Doepke et al. \(2023\)](#). More importantly, it also offers quantitative predictions about the timing of this fertility recovery and the pace of norm evolution, driven by both within- and between-cohort effects.

To evaluate the model's quantitative predictions, we calibrate the model to South

Korean data from 1999 to 2014. Gender-specific wage paths are set to align with trends in total factor productivity and gender wage gaps. Preference parameters are adjusted to match fertility trends, while the substitutability of childcare between genders reflects the initial allocation. The strength of social pressure is calibrated to fit gender disparities in childcare time from the Korean Time Use Survey (KTUS), yielding estimates consistent with prior research (e.g., [Myong et al. 2021](#)). The weight that old cohorts put on their past childcare practices when they re-evaluate the current condition is calibrated to match the share of within-cohort changes in driving the overall social norm evolution, calculated using data from the Korean General Social Survey (KGSS).

Using the calibrated model, we conduct five counterfactual experiments. In the first, we adjust the pace of structural change in the economy. We find that faster structural change accelerates the fertility decline, aligning with empirical evidence. However, the transition paths of childcare and social norms remain largely unchanged, as strong social pressure limits shifts in childcare allocation during this transformation. In the second counterfactual, we vary the intensity of social pressure while holding the structural change constant. Here, economies with weaker social pressure exhibit a smaller fertility drop and a faster social norm transition toward a new steady state, consistent with our empirical findings. In the third counterfactual, we vary the weight that older cohorts put on their past childcare practices when they re-evaluate the current situation. We find that a smaller weight accelerates the social norm convergence towards the new steady state and raises fertility rates along the transition path. In the fourth counterfactual, we use the parameters calibrated to the U.S. economy and examine the path of fertility and social norms. Finally, we assess a hypothetical government policy that promotes a more egalitarian childcare norm. This policy also accelerates the social norm transition to the new steady state. Although its initial impact on fertility is modest initially, the pro-fertility effects strengthen over the transition, yielding significant differences in fertility along the transition path. This experiment underscores the need to account for dynamic effects in policy evaluations.

To summarize, this paper establishes that faster structural change or more rigid social norms exacerbates fertility decline, a fact we explain with a novel model of endoge-

nous norms and household bargaining. Calibrated to South Korea, it matches observed trends, delivers new predictions—non-monotonic fertility and policy-driven recovery—and enriches the fertility literature with dynamic theory and actionable insights for a pressing global issue.

Related Literature

This paper contributes to the growing body of literature examining the effects of technological change on fertility and female labor supply. The most closely related studies in this domain are [Goldin \(2024\)](#), [Doepke and Kindermann \(2019\)](#), and [Myong et al. \(2021\)](#).

The work by [Goldin \(2024\)](#) demonstrates that countries with “lowest low” fertility rates experienced rapid growth in GNP per capita alongside slowly evolving traditions. She proposes a theoretical model where both generational and gendered conflicts precipitate a sharp decline in the total fertility rate. Our paper aligns with this perspective, positing that fertility rates emerge from a dynamic interplay between technological advancements and social norms. However, unlike [Goldin \(2024\)](#), which directly assumes gender-specific desired fertility levels, we derive fertility choices and bargaining dynamics between women and men from microeconomic foundations. Furthermore, we propose a model of social norms incorporating both within- and between-cohort effects, enabling all economic agents to adapt to shifting economic conditions rather than adhering rigidly to traditions. Finally, while [Goldin \(2024\)](#) posits that men prioritize inherited traditions more than women, leading to a higher desired number of children, we assume homogeneous preferences across genders and explain gaps in desired fertility entirely through the negotiation of childcare responsibilities and initial technological conditions.

[Doepke and Kindermann \(2019\)](#) find that countries with more equitable childcare practices exhibit fewer discrepancies between men and women regarding desired fertility, resulting in higher birth rates. They explain this pattern with a bargaining model with limited commitment. We build on and extend their framework by incorporating the influence of social norms and modeling the norms’ endogenous evolution. This ad-

ditional element allows our model to address how economies adjust to technological changes over time through social norm adaptation.

The study by [Myong et al. \(2021\)](#) explores fertility and marriage patterns in East Asian societies, analyzing the effects of two social norms: one related to the stigma of out-of-wedlock births and another concerning the gender division of childcare. Estimating their model with South Korean data, they conclude that the latter norm exerts a stronger influence on fertility outcomes. This finding inspires our focus on the childcare social norm, abstracting from marriage norm considerations.¹ While [Myong et al. \(2021\)](#) treat the childcare social norm as exogenous and estimate it empirically, we micro-found the endogenous formation of such norms. Consequently, our approach enables predictions about future fertility trends and facilitates counterfactual analyses based on structural parameters of norm formation.

Our research complements studies on the evolution and transmission of social norms, such as [Bisin and Verdier \(2001\)](#) and [Bisin and Verdier \(2011\)](#). Notably, several influential papers have explored social learning as a key mechanism driving the increase in female labor force participation amid technological advancements (e.g., [Fernández and Fogli \(2009\)](#), [Fogli and Veldkamp \(2011\)](#), [Fernández \(2013\)](#)). These studies highlight how individuals learn about the effects of female labor force participation on child outcomes and how this knowledge is transmitted across generations (vertical socialization) and within neighborhoods or peer groups (horizontal socialization). Complementing their approach, we investigate the role of social norms surrounding childcare responsibilities, examining their impact on household decisions and their evolution through both within-cohort and between-cohort adjustments, ultimately driven by technological changes.

The remainder of the paper is structured as follows. Section 2 presents the motivating empirical evidence. Section 3 outlines the theoretical model. Section 4 details the calibration strategy and results. Section 5 conducts the primary counterfactual exercises. Section 6 evaluates the model’s robustness. Finally, Section 7 offers concluding

¹Similar to [Myong et al. \(2021\)](#), we also abstract away from status externalities in [Kim et al. \(2024\)](#) or the fertility norm in [De Silva and Tenreyro \(2020\)](#).

remarks.

2. Motivating Facts

This section presents cross-country evidence on the relationship between fertility, structural change, and social norms, drawing data from multiple sources. We document the following two empirical regularities:

Fact 1: Economies experiencing faster structural change exhibit more rapid fertility declines.

Fact 2: The relationship between structural change and fertility is stronger in societies with more stringent social norms.

To examine these relationships, we combine different sources of data. Data on total fertility rate is obtained from the United Nations Population Division. Sectoral employment data are from the Groningen Growth and Development Centre (GGDC) 10-Sector Database (Timmer et al., 2015). This database offers long-term, internationally comparable data on value added and employment across ten sectors for more than 40 economies, including both developing and developed ones.² Data on gross national product (GDP) is collected from the Penn World Table 10.01. Lastly, we follow the literature on sociology and cultural studies in measuring the stringency of social norms. In particular, we obtain the tightness/looseness index from Uz (2015), where the author measures the tightness of social norms using the dispersion of opinions. The motivation of this measure is that in a tight culture, people’s values, norms, and behavior are similar to each other because deviations are sanctioned. The merged sample includes 23 countries from all levels of economic development, with observations from the 1950s to 2010.

To construct the speed of fertility change for country i , we run the regression

$$\text{tfr}_{i,\text{year}} = \alpha_i^{\text{tfr}} + \text{speed_tfr}_i \times \text{year} + u_i \quad (1)$$

²We follow Herrendorf et al. (2014) in the sector assignments. In particular, Agriculture corresponds to the sum of International Standard Industrial Classification (ISIC) sections A–B, Manufacturing corresponds to the sum of ISIC sections C, D, F, and Services correspond to the sum of ISIC sections E, G–P.

where speed_tfr_i measures the average annual change of the total fertility rate for country i during the sample period.

Likewise, we measure the speed of structural change for country i using

$$\text{service share}_{i,\text{year}} = \alpha_i^{\text{ser}} + \text{speed_ser}_i \times \text{year} + v_i \quad (2)$$

where service share measures the fraction of service employment, and

$$\text{agriculture share}_{i,\text{year}} = \alpha_i^{\text{agr}} + \text{speed_agr}_i \times \text{year} + v_i \quad (3)$$

where agriculture share measures the fraction of agricultural employment. Therefore, speed_ser_i and speed_agr_i measures the average annual change of sectoral employment for country i during the sample period.

Figure 1a plots the correlation between speed_tfr_i and speed_ser_i . As can be seen, countries that have experienced faster expansion of the service sector, e.g., South Korea and Peru, also witnessed more rapid fertility declines. Similarly, Figure 1b indicates that countries that have experienced faster shrinkage of the agriculture sector, e.g., South Korea, China, and Mexico, also experienced more rapid fertility declines. The latter finding complements Ager et al. (2020) who document the relationship between agriculture decline and fertility transition in the American South.

To examine the relationship between structural change and fertility decline more systematically, we run a set of OLS regressions where we regress speed_tfr_i on speed_SC_i where SC (structural change) takes the value of speed_ser_i or speed_agr_i , after controlling for the growth rate of GDP per capita speed_gdp_i . We also interact the speed of structural change with a dummy variable “tight”, which takes the value of 1 if the country’s opinion dispersion, measured by Uz (2015), is below the median.

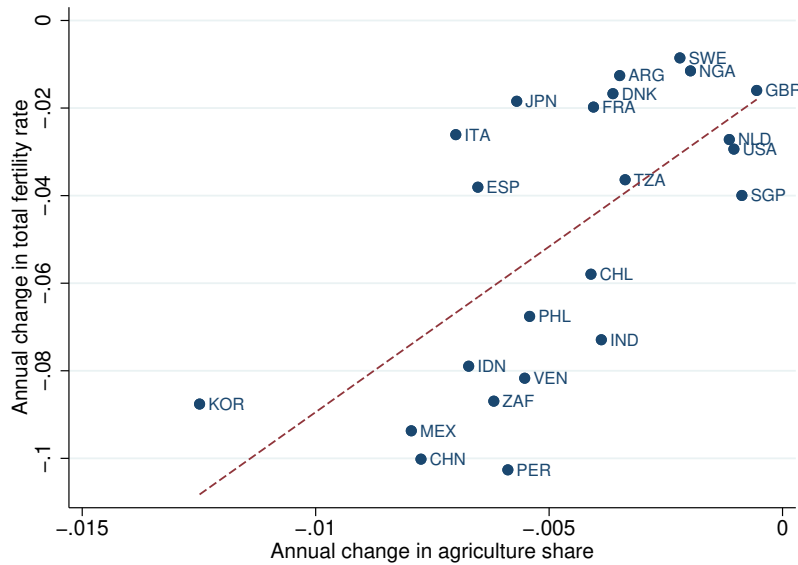
The regression results, presented in Table 1, reveal key insights. Column (1) shows that a 1-percentage-point increase in the speed of service sector expansion per year corresponds to an annual fertility decline of 0.1, a statistically and economically significant correlation. This relationship holds after controlling for GDP per capita growth in Column (2). In Column (3), we observe that the link between service sector expan-

Figure 1: Fertility Change and Structural Change

(a) Service Expansion and Fertility Decline



(b) Agriculture Shrinkage and Fertility Decline



Notes: “speed_tfr” is defined as the average annual percentage change in total fertility rate for each country in the sample period, using data from the United Nations Population Division. “speed_ser” and “speed_agr” are the average annual change in the fraction of employment in service and agriculture sectors respectively, using data from the Groningen Growth and Development Centre (GGDC).

sion and fertility decline is predominantly driven by countries with rigid social norms. Specifically, countries with above-median cultural tightness exhibit a coefficient on service sector expansion speed nearly 50% larger than those below the median. Column (4) confirms that this heterogeneity persists even after adjusting for GDP per capita growth. Columns (5) through (8) replicate the analysis using the decline in agricultural employment as an alternative measure, yielding findings consistent with those for service sector expansion.

Table 1: Regression Results: Cross-Sectional Data

	Dependent Variable: Fertility Change							
	Service				Agriculture			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
speed.SC	-10.44*** (3.38)	-11.82*** (3.78)	-5.48 (4.02)	-6.89 (4.27)	7.56*** (1.90)	8.39*** (2.06)	5.03** (2.38)	5.80** (2.41)
tight×speed.SC			-5.11* (2.56)	-5.23* (2.56)			3.23 (1.95)	3.51* (1.93)
speed.gdp		0.30 (0.35)		0.33 (0.33)		0.32 (0.31)		0.38 (0.30)
Observations	23	23	23	23	23	23	23	23
R-squared	0.31	0.34	0.43	0.46	0.43	0.46	0.50	0.54

Notes: “speed.tfr” is defined as the average annual change in total fertility rate for each country in the sample period, using data from the United Nations Population Division. “speed.SC” where SC ∈ {ser, agr} are the average annual change in the fraction of employment in service and agriculture sectors respectively, using data from the Groningen Growth and Development Centre (GGDC). “speed.gdp” is calculated as the annual change in GDP per capita in the sample period, using data from the Penn World Table. “tight” is a dummy variable that takes the value of 1 if the tightness index is above the sample median, reflecting stronger constraints imposed by social norms on individual behavior. The index is developed by [Uz \(2015\)](#).

In addition, panel data regressions were performed to examine the relationships between fertility change, structural change speed, and evolving social norms. Fertility and structural change speed were calculated as 10-year changes, based on the previ-

ously described datasets. Social norm (gender attitudes) changes were quantified using data from the 2002 and 2012 International Social Survey Programme (ISSP) Family and Changing Gender Roles modules. Specifically, we analyzed responses to the statement, “A man’s job is to earn money; a woman’s job is to look after the home and family.” Two measures were created: “Norm Change: Total” captures the difference in average scores between the 1980 and 1920 birth cohorts, while “Norm Change Recent” reflects the 10-year inter-cohort change in gender attitudes. For instance, the change between 1960-1969 is determined by comparing the 1940-1949 and 1930-1939 birth cohorts. GDP per capita level and growth were included as control variables.

The panel regression results, summarized in Table 2, corroborate our cross-sectional observations. We find that accelerated growth in service sector employment and a corresponding reduction in agricultural employment are significantly correlated with faster fertility declines, even after accounting for country fixed effects and year trends. Notably, a one-percentage-point increase in service sector share over a 10-year period is associated with a 0.07 decrease in the total fertility rate over the same time frame, a result consistent in magnitude with our cross-sectional analysis. Additionally, the interaction term between structural change speed and gender norm changes exhibits a positive sign for service sector growth, suggesting that the relationship between structural change and fertility is mitigated in societies with more adaptable social norms, consistent with the cross-sectional findings. The interaction term for agricultural share decline is negative, as the share of agriculture in the aggregate economy declines with economic development.

To summarize, we find that countries that have experienced faster structural change, i.e., service expansion or agriculture decline, witnessed more rapid decline in fertility. Furthermore, this correlation is robust to controlling for the growth in GDP per capita, and is stronger in societies with rigid social norms. In the next section, we propose a structural model to understand these relationship.

Table 2: Regression Results: Panel Data

	Dependent Variable: Fertility Change							
	Service				Agriculture			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Speed_SC	-6.64*** (0.70)	-7.32*** (0.74)	-10.40*** (1.57)	-7.23*** (0.91)	7.66*** (0.53)	8.91*** (0.62)	9.58*** (0.92)	9.61** (0.63)
Speed_SC×Norm Change Total			5.35** (2.40)				-1.94 (1.98)	
Speed_SC×Norm Change Recent				0.59 (0.38)				-0.49 (0.31)
Norm Change Recent				0.59 (0.38)				-19.42*** (4.09)
Controls	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Country FEs	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Year Trend	No	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	785	785	785	785	785	785	785	785
R-squared	0.26	0.38	0.39	0.39	0.35	0.45	0.45	0.47

Notes: Fertility change is defined as the 10-year change in total fertility rate for each country in the sample period, using data from the United Nations Population Division. Structural change speed ("Speed_SC") is measured as the 10-year change in the service and agricultural employment shares, respectively, calculated using the Groningen Growth and Development Centre (GGDC) 10-sector database for years after 1950 and [Mitchell \(2007\)](#) for years before 1950. The variables "Norm Change Total" and "Norm Change Recent" are constructed from data obtained from the International Social Survey Programme: Family and Changing Gender Roles III (ISSP) 2002 and 2012. Both variables are based on responses to the question, "To what extent do you agree or disagree: A man's job is to earn money; a woman's job is to look after the home and family. " Responses are coded as -1 if disagree, 1 if agree, and 0 if neutral. "Norm Change Total" represents the difference in average scores between the 1980 and 1920 birth cohorts. "Norm Change Recent" measures the change in gender attitude of each birth cohort in a country compared to the cohort from 10 years prior. For example, if we examine the change between 1960-1969, then we are looking at the change in attitudes of the 1940-1949 birth cohort compared to the 1930-1939 birth cohort.. Control variables include GDP per capita level and growth.

3. Model

This section introduces an overlapping generations model where parents bargain over fertility and childcare responsibilities under the influence of social norm. Moreover, the prevailing social norm in the economy is endogenously determined by older cohorts' opinions which depend on past childcare practices.

3.1 Household Problem

We analyze an overlapping generations economy where women and men live until age J . Each cohort consists of women and men of equal mass.³ To examine generational conflicts and fertility dynamics, we assume homogeneity among agents of the same gender within each cohort. To emphasize fertility and childcare decisions, we model individuals as consuming their labor income before and after age J_f , the period during which individuals form couples and engage in a bargaining problem.

A couple comprises a woman and a man, indexed by gender $g \in \{\varnothing, \sigma\}$. They choose individual consumption c^g , childcare contributions l^g , and the number of children n , which is a household-shared decision. We use t to denote time. Gender-specific market wages, denoted w_t^g , are exogenous and reflect technological changes impacting labor demand differentially by gender (Ngai and Petrongolo 2017).⁴

Individual's preference over fertility and consumption are given by:

$$u^g(c^g, n) = c^g + \gamma \cdot \frac{n^{1-\rho} - 1}{1 - \rho} \quad \rho > 0 \quad (4)$$

where c^g represents personal consumption and n the number of children. The parameters γ and ρ control the weight and curvature of fertility n in the utility function, respectively.

Raising each child incurs a time cost ϕ . Thus, to support n children, the couple must satisfy the childcare provision constraint:

$$n\phi = \left((l^\varnothing)^{\frac{\sigma-1}{\sigma}} + (l^\sigma)^{\frac{\sigma-1}{\sigma}} \right)^{\frac{\sigma}{\sigma-1}}, \quad \sigma > 1 \quad (5)$$

where l^\varnothing and l^σ denote childcare time from the woman and man, respectively, and σ governs the substitutability of their contributions.

Following Doepke and Kindermann (2019), we assume that spouses bargain over fertility and childcare under partial commitment. The decision-making timeline is out-

³The size of each incoming cohort may vary over time due to endogenous fertility decisions.

⁴Alternatively, we could model structural change via time-varying, gender-specific labor demand, with equilibrium wages determined by market clearing. However, as this paper focuses on household-level fertility and childcare responses, we adopt the simpler exogenous wage approach.

lined below.

First Stage: Childcare Arrangements

The couple first determines a plan for childcare allocation, conditional on having n children, and commits to this decision post-childbirth.⁵ For all n , the couple solves:

$$\min_{l_t^{\varnothing}, l_t^{\sigma}} w_t^{\varnothing} l_t^{\varnothing} + w_t^{\sigma} l_t^{\sigma} + \lambda \cdot w_t^{\sigma} \cdot \left(\frac{l_t^{\varnothing}}{l_t^{\sigma}} - \eta_t \right)^2, \quad (6)$$

subject to the childcare constraint (5). Here, $w_t^{\varnothing} l_t^{\varnothing} + w_t^{\sigma} l_t^{\sigma}$ represents the opportunity cost of childcare in the units of consumption goods, while $\lambda \cdot w_t^{\sigma} \cdot \left(\frac{l_t^{\varnothing}}{l_t^{\sigma}} - \eta_t \right)^2$ captures psychic costs from deviating from the prevailing social norm, η_t , which is taken as given by parents at time t . The parameter λ governs the intensity of these social pressure costs. We scale the psychic cost by w_t^{σ} so that what matters for the childcare allocation is the gender wage gap $w_t^{\varnothing}/w_t^{\sigma}$ rather than the universal wage levels. The solutions to this cost-minimization problem, denoted $l_t^{\varnothing}(n)$ and $l_t^{\sigma}(n)$, represent optimal childcare allocation under the influence of the social norm.

Second Stage: Bargaining over Fertility

Next, the couple negotiates the number of children, n_t . Only mutually agreed-upon fertility is realized, defined as:

$$n_t = \min\{n_t^{\varnothing}, n_t^{\sigma}\}, \quad (7)$$

where n_t^g is the fertility level that maximizes the ex-post utility of gender $g \in \{\varnothing, \sigma\}$ in the third stage.

Third Stage: Bargaining over Consumption

Post-childbirth, the couple implements the agreed childcare arrangement $l_t^g(n)$ and bargains over consumption $c^g(n)$. If the couple ends up having n children, each gender's

⁵Doepeke and Kindermann (2019) justifies this commitment by noting that childcare decisions involve significant switching costs and advance planning (e.g., securing daycare slots before birth) and interact with persistent choices like residential location, which affects childcare availability.

outside option in the non-cooperative case is:

$$\bar{u}^g(n) = w_t^g(1 - l_t^g(n)) + \gamma \cdot \frac{n^{1-\rho} - 1}{1 - \rho}, \quad \rho > 0, \quad (8)$$

reflecting consumption of residual labor income after childcare.

If they cooperate, the Nash bargaining problem is:

$$\max_{c^\varnothing, c^\sigma} \left(u^\varnothing(c^\varnothing, n) - \bar{u}^\varnothing(n) \right)^{1/2} \cdot \left(u^\sigma(c^\sigma, n) - \bar{u}^\sigma(n) \right)^{1/2}, \quad (9)$$

subject to the budget constraint:

$$c^\varnothing + c^\sigma = (1 + \alpha) \cdot [w_t^\varnothing(1 - l_t^\varnothing(n)) + w_t^\sigma(1 - l_t^\sigma(n))], \quad (10)$$

where α captures economies of scale from cooperation. Following [Doepeke and Kindermann \(2019\)](#), we assign each spouse equal bargaining power (one-half).

3.2 Endogenous Social Norm

A novel feature of this model is that the prevailing social norm is endogenously determined. Unlike exogenous social norm assumption used in the literature, we posit that the social norm at time t , denoted η_t , emerges as a weighted average of the *opinions* expressed by the other members in the society, which in turn depends on the childcare practices in the past. This dynamic process ties current norms to historical household decisions, capturing both social pressure and reevaluation by older cohorts.

Specifically, the prevailing social norm at time t is defined as:

$$\eta_t = \sum_{j=1}^{J-J_f} \phi_{J_f+j,t} \cdot \tilde{\eta}_{J_f+j}, \quad \sum_{j=1}^{J-J_f} \phi_{J_f+j,t} = 1, \quad (11)$$

where $\tilde{\eta}_{J_f+j}$ represents the opinions of childcare allocation expressed by the cohort at age $J_f + j$ at time t , and $\phi_{J_f+j,t}$ is the weight assigned to that cohort's influence.⁶ This

⁶By assuming away heterogeneity with cohorts, this paper abstracts away from the horizontal social-

weight reflects the share of households aged $J_f + j$ at time t among the subpopulation older than J_f , calculated as:

$$\phi_{j,t} = \frac{\pi_{j,t}}{\sum_{k=J_f+1}^J \pi_{k,t}}, \quad (12)$$

where $\pi_{j,t}$ denotes the population share of the cohort aged j at time t . These weights ensure that the influence of past decisions is proportional to the relative size of each cohort.

We assume that older cohorts form opinions by solving the following minimization problem:

$$\tilde{\eta}_{J_f+j} = \arg \min_{\eta} w_t^{\ominus} \cdot \eta + w_t^{\ominus} + \psi \cdot \left(\eta - \frac{l_{t-j}^{\ominus}}{l_{t-j}^{\ominus}} \right)^2. \quad (13)$$

where $\frac{l_{t-j}^{\ominus}}{l_{t-j}^{\ominus}}$ measures the childcare practice adopted by these agents j periods ago.

This minimization problem reflects changes in social norms within cohorts, where older generations re-evaluate the situation given the prevailing market wage, and choose the opinion they express to the current parents. Parameter ψ governs the utility costs of deviating from their own past choice, i.e., stubbornness. If they are perfectly altruistic towards the current parents, they would simply discard their own past decisions and minimize the opportunity costs of childcare, i.e., $\psi = 0$. However, if ψ is high, they will express opinions that are more in line with how they brought up their own children j periods ago.

3.3 Population Dynamics

The demographic structure of this economy, denoted $\{\pi_{j,t}\}_{j=1}^J$, evolves endogenously according to a law of motion driven by the fertility rate n_t .

Let $\boldsymbol{\pi}_t = (\pi_{1,t}, \dots, \pi_{J,t})^T$ represent the population distribution across age groups at time t , where $\pi_{j,t}$ is the share of individuals aged j . The population dynamics are governed by:

$$\boldsymbol{\pi}_{t+1} = \frac{\boldsymbol{\Pi}_t \cdot \boldsymbol{\pi}_t}{\|\boldsymbol{\Pi}_t \cdot \boldsymbol{\pi}_t\|_{L^2}}, \quad (14)$$

ization discussed in [Bisin and Verdier \(2011\)](#).

where Π_t is a $J \times J$ demographic transition matrix.

The matrix Π_t has two key features: (1) elements in the j -th row and $(j+1)$ -th column equal 1, reflecting aging from age j to $j+1$, and (2) the element in the first row and J_f -th column equals $n_t/2.1$, representing births from couples at age J_f , normalized by the replacement fertility rate (approximately 2.1 children per couple). The L^2 -norm ensures the population shares sum to 1, maintaining a normalized distribution over time.

3.4 Model Solution

We now outline the solution algorithms for the model's steady state and transition path.

Steady-State Solution The steady state is computed using an iterative method that converges to a social norm consistent with household decisions:

1. Guess an initial social norm η .
2. Given η , solve the childcare arrangement problem (6) subject to the childcare constraint (5), yielding optimal childcare functions $l^\varnothing(n)$ and $l^\sigma(n)$.
3. Using $l^\varnothing(n)$ and $l^\sigma(n)$, calculate each gender's outside option utility, $\bar{u}^g(n)$, for $g \in \{\varnothing, \sigma\}$.
4. With the outside options, solve the Nash bargaining problem for consumption $c^g(n)$ and utility $u^g(n)$, conditional on n .
5. Maximize each gender's utility, $n^g = \arg \max_n u^g(n)$, to determine desired fertility. The realized fertility is $n = \min\{n^\varnothing, n^\sigma\}$, with associated childcare arrangements $l^\varnothing(n)$ and $l^\sigma(n)$.
6. Compute the implied social norm, $\tilde{\eta} = l^\varnothing(n)/l^\sigma(n)$. Update η and iterate until $|\eta - \tilde{\eta}| < \epsilon$, for a small tolerance $\epsilon > 0$.

Transition Path To compute the economy's dynamics, starting from any initial state (not necessarily the steady state), we use forward iteration:

1. Given the current social norm η_t , solve the static household optimization problem to obtain n_t , $l^\varnothing(n_t)$, and $l^\sigma(n_t)$.
2. Using the current demographic structure π_t and fertility rate n_t , update the population distribution to π_{t+1} via Equation (14).
3. With π_{t+1} and the history of childcare arrangements $\{l_{t-j}^\varnothing, l_{t-j}^\sigma\}_{j=0}^{J-J_f}$, calculate tomorrow's social norm η_{t+1} using Equation (11) after solving the every older cohorts' minimization problem (13). Repeat from step 1 for the next period.

This iterative process captures the co-evolution of fertility, childcare norms, and population structure over time.

3.5 Mechanisms

The central prediction of this model is that fertility emerges from a dynamic tension between technological progress and evolving social norms. Technological change, by increasing women's relative wages, shifts the optimal childcare division toward greater gender equality for young couples making fertility decisions—absent social norm pressures. However, when social norms enforce traditional roles, requiring women to bear a disproportionate childcare burden, women prefer fewer children *ex ante*. This preference arises because raising children under such norms diminishes their utility in the outside option *ex post*, weakening their bargaining power over consumption within the household.

Conversely, social norms in this economy are endogenous, shaped by changes within- and between-cohort. Entering cohorts, influenced by improved labor market opportunities for women, may favor more equitable childcare arrangements. Older cohorts, aware of the changes in the gender wage gap, may also shift their positions and choose what opinions to express.

The evolution of the social norm thus stems from two sources. First, when new cohorts enter at age J_f and choose childcare contributions l^\varnothing and l^σ that diverge from established practices, they shift future norms toward their preferences. For instance, a

cohort opting for more egalitarian childcare ratios will incrementally reduce η_t in subsequent periods. This channel reflects between-cohort changes, also known as the cohort replacement effects.

Second, older cohorts choose the opinions they would like to express. This decision hinge on their past experiences, but also reflects their re-evaluation of the situation given the current conditions. This channel reflects within-cohort changes, also known as the social structural effects.

The influence of these two effects depends on the population weight of the corresponding cohort, $\phi_{j,t}$, which is itself an endogenous outcome of fertility decisions made by their parents' generation. Higher past fertility increases a cohort's size, amplifying its impact on η_t . This interplay between fertility, demographic structure, childcare choices, and opinions drives the norm's gradual adaptation over time.

Note that even though the social norm adjusts through both within- and between-cohort effects, the adaptation speed might still be slower than the rapid changes in technology. Therefore, the model predicts a non-monotonic fertility response to technological change that boosts women's wages. Initially, fertility declines sharply as the social norm lags, reflecting entrenched gender roles. Over time, as the social norm adjusts toward equality, fertility recovers. This transition highlights the interplay between short-term path-dependence and long-term adaptation.

The advantage of explicitly modeling the social norm is that we can quantify the above-mentioned channels and make numerical predictions. In the next section, we calibrate the model and conduct counterfactual analyses in the subsequent sections.

4. Calibration

This section outlines the calibration strategy and data sources, and presents the calibrated parameter values.

4.1 Identification Strategy

We calibrate the model to replicate key data moments from South Korea over the period 1999–2014. The parameters to be calibrated are:

$$\underbrace{J, J_f}_{\text{demographics}}, \quad \underbrace{\gamma, \rho, \psi, \lambda}_{\text{preferences}}, \quad \underbrace{\phi, \sigma, \alpha}_{\text{technologies}}.$$

Several parameters are directly sourced from the literature. The economies of scale from spousal cooperation, α , is set to 1.2, following [Doepke and Kindermann \(2019\)](#). The time cost per child, ϕ , is fixed at 0.15, based on [de La Croix and Doepke \(2003\)](#). We define each period as 5 years, setting $J = 16$ (total lifespan of 80 years) and $J_f = 6$ (child-bearing between 25 to 30) to align with South Korea’s life expectancy of approximately 80 years and a mean childbirth age of 25–30 during the study period.

The remaining parameters are calibrated to match the transition path of the South Korean economy from 1999 to 2014. In particular, we choose the exogenous time series for gender-specific wages, $w_t^{\text{♀}}$ and $w_t^{\text{♂}}$ to match the GDP per capita and gender wage gap statistics collected from the OECD database. Then, we simulate the path of the economy and ask the model to fit the observed trajectories of fertility, gender gaps in childcare responsibilities, and the share of within-cohort effects in social norm changes.

We collect fertility data from the United Nations. Gender-specific childcare time is computed using the micro-level data from the Korea Time Use Survey (KTUS) following the strategy of [Park \(2021\)](#). Lastly, we use the micro-level data from the Korean General Social Survey (KGSS) to compute the share of within-cohort effects in driving social norm changes. To be more specific, for any variable Y , we first compute the change in average over time, i.e., the gap between $\bar{Y}_{\text{start time}}$ and $\bar{Y}_{\text{end time}}$. To compute within-cohort changes, we identify the cohorts that have more than N observations both at the start time and the end time. Then, we restrict the samples to the individuals from these cohorts, and compute how much Y changes over time in this sample, i.e., the gap between $\bar{Y}_{\text{start time}}^{\text{subsample}}$ and $\bar{Y}_{\text{end time}}^{\text{subsample}}$. The ratio between these two gaps measures the importance of within-cohort effects, and is used as a targeted moment. In the KGSS data,

we evaluate this ratio for three variables as Y :

1. SEXROLE1: Opinions on gender roles: It is more important for a wife to help her husband's career than to pursue her own career,
2. SEXROLE2: Opinions on gender roles: A husband's job is to earn money; a wife's job is to look after the home and family, and
3. HBBYWK08: Agree or disagree: Husband's job is to earn money; wife's job is to look after the home and family.

For all three cases, we find that the ratio is close to 0.8, implying an important role for within-cohort effects in driving the broader social norm change in South Korea in the study period.

Although all parameters jointly influence the model's moments, certain moments provide stronger identification for specific parameters. Below, we outline the identification logic:

- The fertility weight, γ , is inferred from the initial fertility level in 1999, reflecting baseline preferences for children.
- The fertility curvature, ρ , governs the trade-off between consumption and fertility, identified by the fertility response to rising opportunity costs as wages increase for both genders.
- The childcare substitutability, σ , is determined by the initial gender gap in childcare time. Higher substitutability amplifies the initial gender gap by allowing greater specialization.
- The weight of individual's own experience in the formation of opinions, i.e., "stubbornness", ψ , is calibrated to match the share of between-cohort component in driving social norm changes. Smaller ψ leads to larger within-cohort effects.
- The social pressure parameter, λ , is calibrated to the persistence of gender gaps in childcare over time. A higher λ implies stronger resistance to change, as young couples face greater pressure from older cohorts' norms.

This strategy ensures the model captures both the levels and dynamics of fertility and childcare in South Korea over the calibration period.

4.2 Calibration Results

Table 3 presents the calibrated parameters alongside the corresponding data moments. The model closely replicates the observed trajectories of fertility and gender gaps in childcare responsibilities in South Korea from 1999 to 2014.

Table 3: Calibrated Parameters

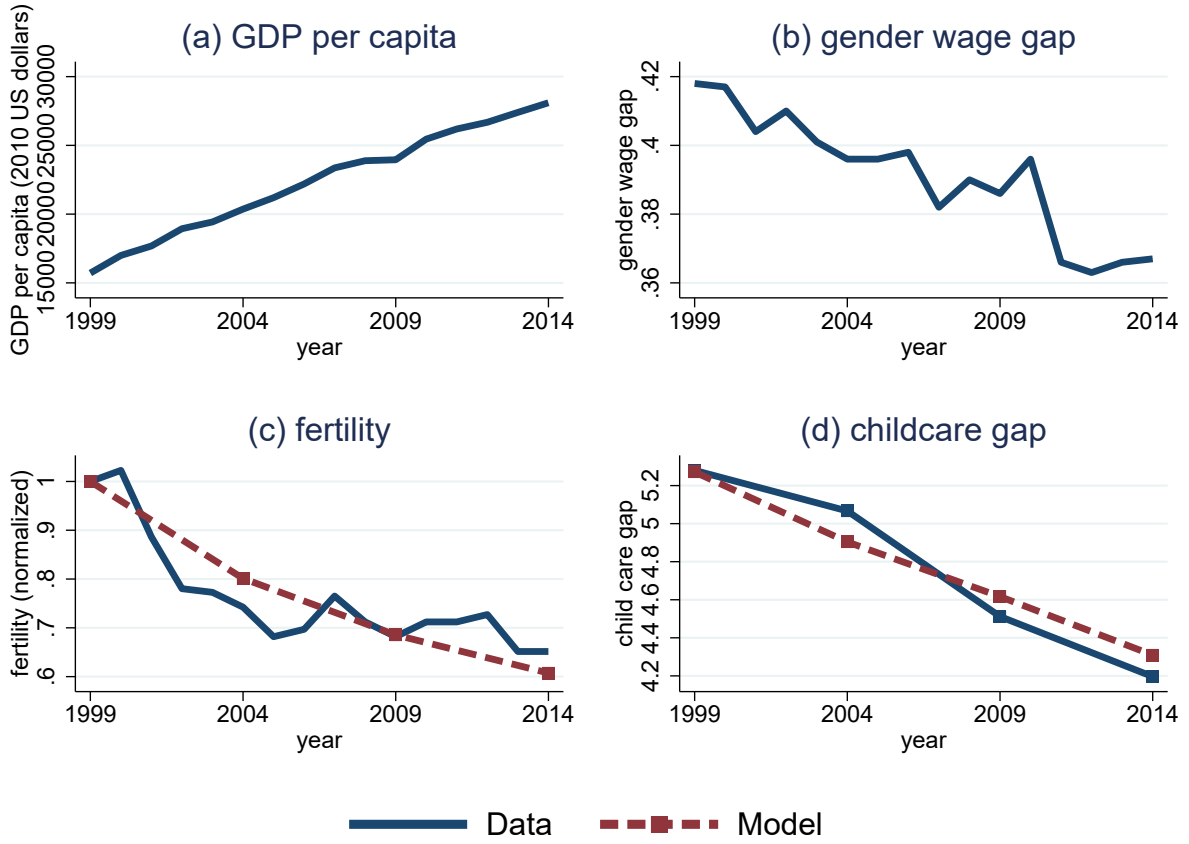
	Parameter	Value	Data moment	Source	Model fit
γ	Fertility weight	0.24	$n_{1999} = 1.42$	United Nations	1.42
σ	Childcare substitutability	3.05	$\eta_{1999} = 5.25$	Park (2021)	5.25
ρ	Fertility curvature	2.4	$n_{1999} \sim n_{2014}$	United Nations	See Figure 2
ψ	Stubbornness	3.0	Within-cohort effects	KGSS	80%
λ	Social pressure	0.0006	$\eta_{1999} \sim \eta_{2014}$	Park (2021)	See Figure 2
α	Economies of scale	1.2	Doepke and Kindermann (2019)		
ϕ	Time costs per child	0.15	de La Croix and Doepke (2003)		
J	Total number of periods	16	80 years	World Health Organization	
J_f	The fertile period	6	25 to 30 yo	Statista	

The calibrated childcare substitutability, $\sigma = 3.05$, aligns closely with the value of 3.03 reported by Knowles (2013). Similarly, the social pressure parameter, $\lambda = 0.0006$, yields psychic costs comparable to those in Myong et al. (2021) and Kim et al. (2024). To illustrate these parameters' implications and the model's mechanisms, Table 4 compares childcare gaps and fertility across three scenarios.

The first scenario “old tech. & old norm” depicts a steady-state economy with a large gender wage gap ($w^{\varphi}/w^{\sigma} = 0.58$). Here, women shoulder 5.25 times more childcare than men, consistent with the prevailing social norm ($\eta = 5.25$). Young couples' choices align with this norm, resulting in a fertility rate of $n = 1.43$.

The second scenario “new tech. & new norm” represents a new steady state with a narrower wage gap ($w^{\varphi}/w^{\sigma} = 0.74$). Reduced wage disparity lowers the opportunity

Figure 2: Calibration and Model Fit



cost of men’s childcare time, shifting the optimal allocation to a more equitable $l^{\varphi}/l^{\sigma} = 2.53$, which matches the updated norm ($\eta = 2.53$). Fertility dips slightly to $n = 1.37$, driven by the substitution effect between consumption and children, as parameterized by ρ .

The third scenario “new tech. & old norm” explores a transitional case where technology narrows the wage gap to $w^{\varphi}/w^{\sigma} = 0.74$, but the social norm remains fixed at $\eta = 5.25$. Young couples, constrained by this norm, settle on a childcare division of $l^{\varphi}/l^{\sigma} = 4.66$, balancing monetary and psychic costs. The unequal allocation worsens women’s outside option, leading them to favor fewer children. With fertility determined by mutual agreement, the result is $n = 1.32$ —notably lower than the long-run equilibrium in the second scenario. The psychic cost of adopting the egalitarian allocation ($l^{\varphi}/l^{\sigma} = 2.53$) under the old norm amounts to 3.48% of child-rearing costs, similar to

the 4.3% estimated by [Myong et al. \(2021\)](#) and [Kim et al. \(2024\)](#).

Table 4: Three Different Cases

	Old tech. & old norm	New tech. & new norm	New tech. & old norm
$w^{\varnothing}/w^{\sigma}$	0.58	0.74	0.74
η	5.25	2.53	5.25
$l^{\varnothing}/l^{\sigma}$	5.25	2.53	4.66
n	1.43	1.37	1.32

These results highlight the model’s ability to capture both steady-state outcomes and transitional dynamics driven by technology and social norms.

5. Counterfactual Analysis

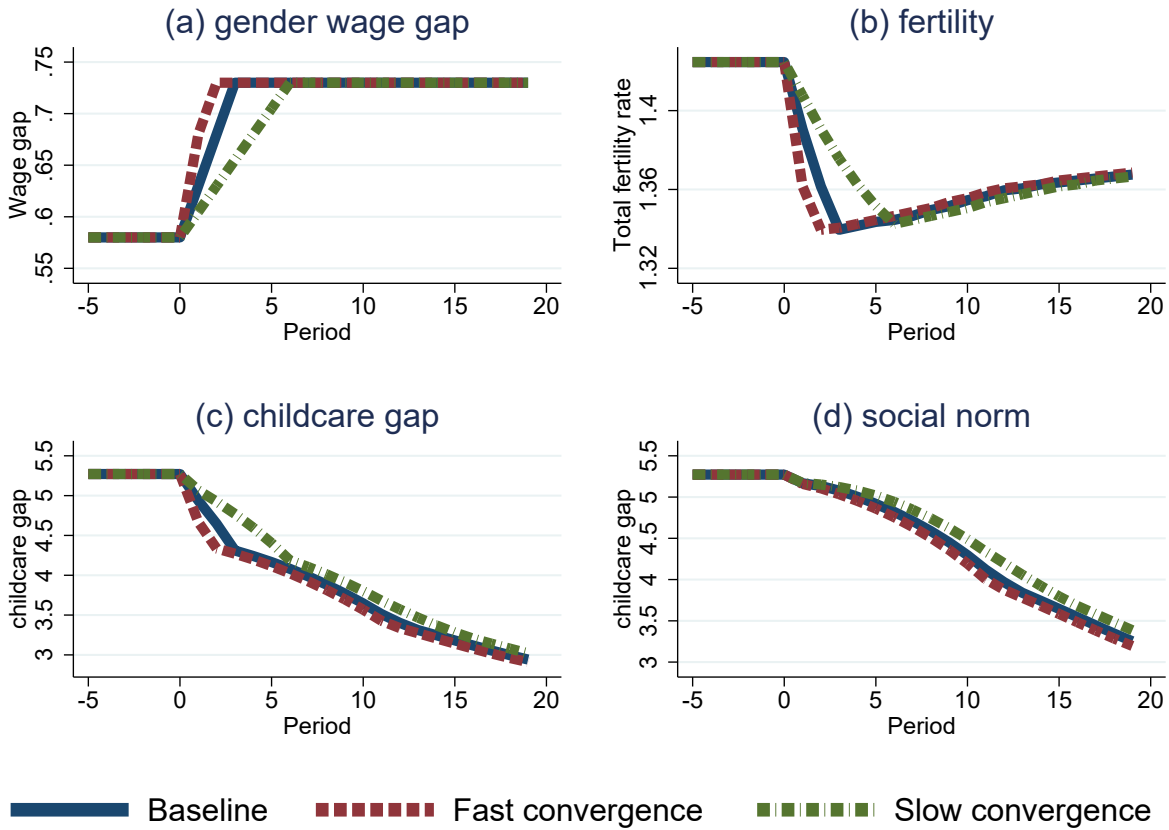
This section conducts counterfactual experiments to explore how technological change, social norms, and policy interventions affect fertility, childcare allocation, and social norms. To focus on technological impacts via the gender wage gap, we hold men’s wages, w_t^{σ} , constant, letting variations in w_t^{\varnothing} drive the results. We analyze five scenarios, with outcomes depicted in [Figures 3, 4, 5, 6, and 7](#), each plotting the trajectories of (a) gender wage gap, (b) fertility, (c) childcare gap, and (d) social norm.

5.1 Speed of Structural Change

We first examine how the pace of structural change influences the economy’s dynamics, comparing a baseline path to fast and slow convergence scenarios. Starting from the 1999 steady state (gender wage gap of 42% ($w_t^{\varnothing}/w_t^{\sigma} = 0.58$)), we simulate a reduction to 27% ($w_t^{\varnothing}/w_t^{\sigma} = 0.73$), as observed in South Korea by 2014. In the baseline, this convergence occurs over 15 years (1999–2014). The fast convergence scenario compresses it to 10 years (1999–2009), while the slow convergence scenario extends it to 25 years (1999–2024). [Figure 3](#) presents these results.

In the fast convergence case, the gender wage gap (panel a) declines sharply, outpacing the adjustment of the social norm (panel d), which remains near its initial value of 5.25 for several years due to cohort replacement lags and limited re-evaluation by older cohorts. This mismatch drives a rapid fertility decline (panel b), dropping from 1.43 to approximately 1.33 within a decade, as women, facing higher opportunity costs, prefer fewer children under a norm favoring unequal childcare (panel c, $l_t^{\varnothing}/l_t^{\sigma} \approx 5.0$). Over time, as entering cohorts adopt more equitable allocations, the childcare gap and norm converge toward 3, and fertility recovers to 1.37 by 2024. The slow convergence case shows a smoother transition: the gender wage gap declines gradually, allowing the social norm and childcare gap to adjust in tandem (e.g., η_t and $l_t^{\varnothing}/l_t^{\sigma}$ fall steadily to 5 and 4.3 by 2014 respectively), with fertility dipping only to 1.34 before stabilizing. The baseline lies between these, with fertility falling to 1.33 by 2014, consistent with calibration.

Figure 3: The Role of Structural Change



Notably, the paths of the childcare gap ($l_t^{\text{♀}}/l_t^{\text{♂}}$) and social norm (η_t) remain broadly similar across scenarios and to the calibrated baseline. This stability reflects the strong social pressure parameter ($\lambda = 0.0006$), which penalizes deviations from traditional practices, and limited re-evaluation by older cohorts ($\psi = 3$). Even as women’s wages rise, young couples adjust childcare allocations incrementally, constrained by the psychic costs of defying a norm shaped by older generations, who express opinions that are similar to their own past practices. The primary divergence thus lies in fertility, underscoring its sensitivity to the timing of structural change relative to norm evolution.

These results, depicted in Figure 3, highlight a non-monotonic fertility response consistent with the model’s mechanisms: rapid technological shifts exacerbate short-term fertility declines by amplifying the tension between economic opportunities and social constraints, while slower changes allow for smoother transitions in fertility.

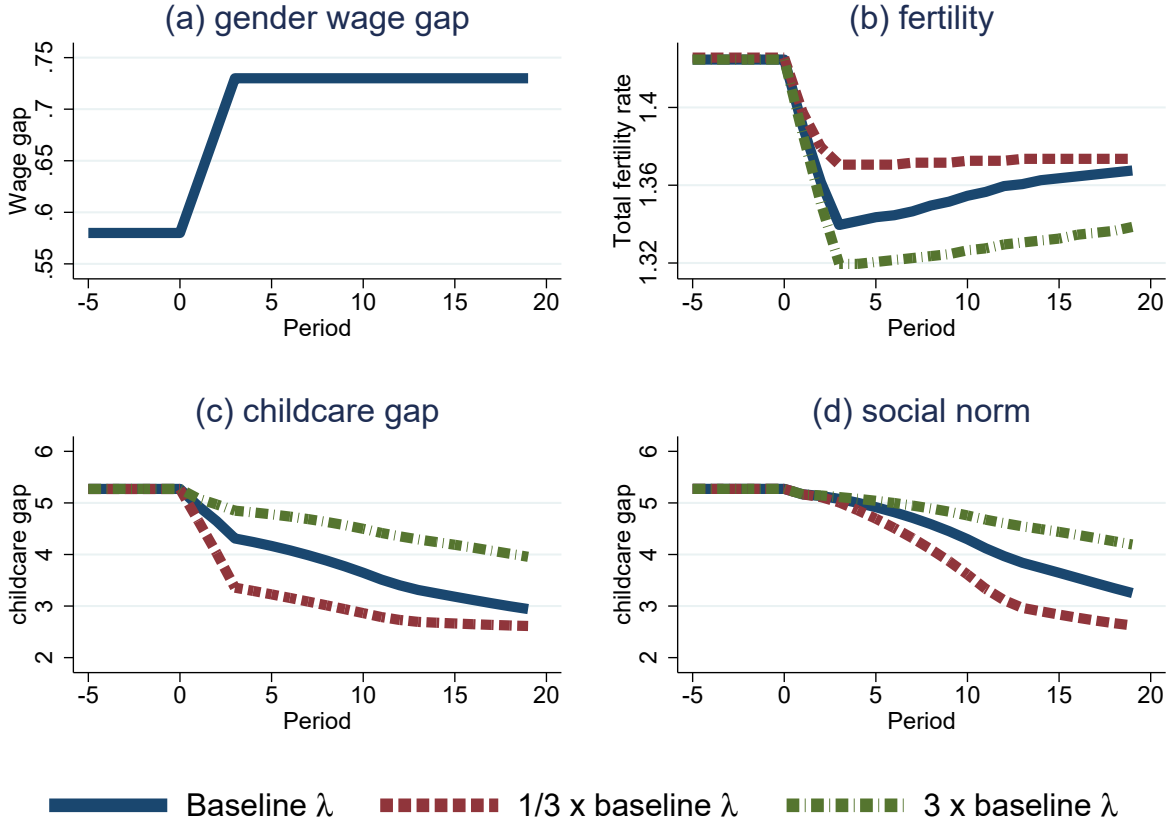
5.2 The Role of Social Pressure

Next, we assess how social pressure modulates the response to structural change. We simulate the baseline 15-year wage gap convergence (42% to 27%) under three levels of social pressure: the calibrated $\lambda = 0.0006$ (baseline), a lower $\lambda = 0.0002$ (weak pressure), and a higher $\lambda = 0.0018$ (strong pressure). Figure 4 displays the outcomes.

With weak social pressure, the gender wage gap’s decline (panel a) prompts a swift reduction in the childcare gap (panel c), falling from 5.25 to 3.3 within 15 years, as couples face lower psychic costs for equitable allocations. The social norm (panel d) follows suit, dropping to 4.8 by 2014, accelerating fertility convergence (panel b) to the long-run level of 1.37. Strong pressure, conversely, locks the childcare gap near 5.0 and the norm near 5.25 for longer, deepening the fertility drop to 1.32 by 2014, with minimal recovery thereafter. The baseline case shows intermediate dynamics: the childcare gap declines to 4.4 and the norm to 4.9 by 2014, with fertility at 1.33.

These findings, shown in Figure 4, underscore social pressure as a critical amplifier of technological shocks. Lower λ facilitates a smoother transition by reducing the tension between rising female wages and lagging norms, while higher λ intensifies short-term

Figure 4: The Role of Social Pressure



fertility losses and prolongs inequitable childcare burdens.

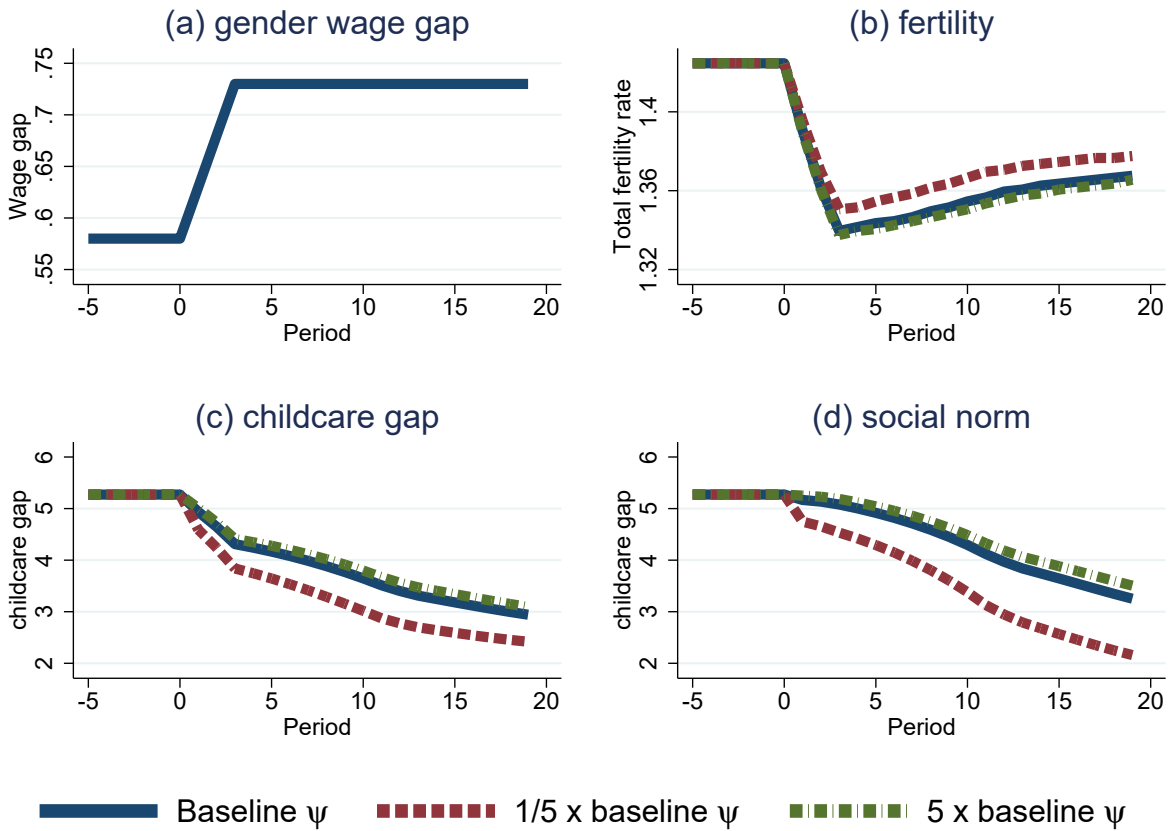
5.3 The Role of Older Cohorts' Re-evaluation

Next, we assess how the re-evaluation by older cohorts when they form social opinions affect the economy's response to structural change. We simulate the baseline 15-year wage gap convergence (42% to 27%) under three levels of "stubbornness": the calibrated $\psi = 3.0$ (baseline), a lower $\psi = 1.0$ (low "stubbornness"), and a higher $\psi = 9.0$ (high "stubbornness"). Figure 5 displays the outcomes.

With low "stubbornness", the gender wage gap's decline (panel a) prompts a faster reduction in the childcare gap (panel c), as older cohorts express more egalitarian opinions and exert lower psychic costs for young couples. The social norm (panel d) con-

verges faster to the new steady state, accelerating fertility convergence (panel b) to the long-run level of 1.37. High “stubbornness”, conversely, locks the childcare gap and the social norm at the original level for longer. The baseline case shows intermediate dynamics: the childcare gap declines to 4.4 and the norm to 4.9 by 2014, with fertility at 1.33.

Figure 5: The Role of Older Cohorts’ Reevaluation

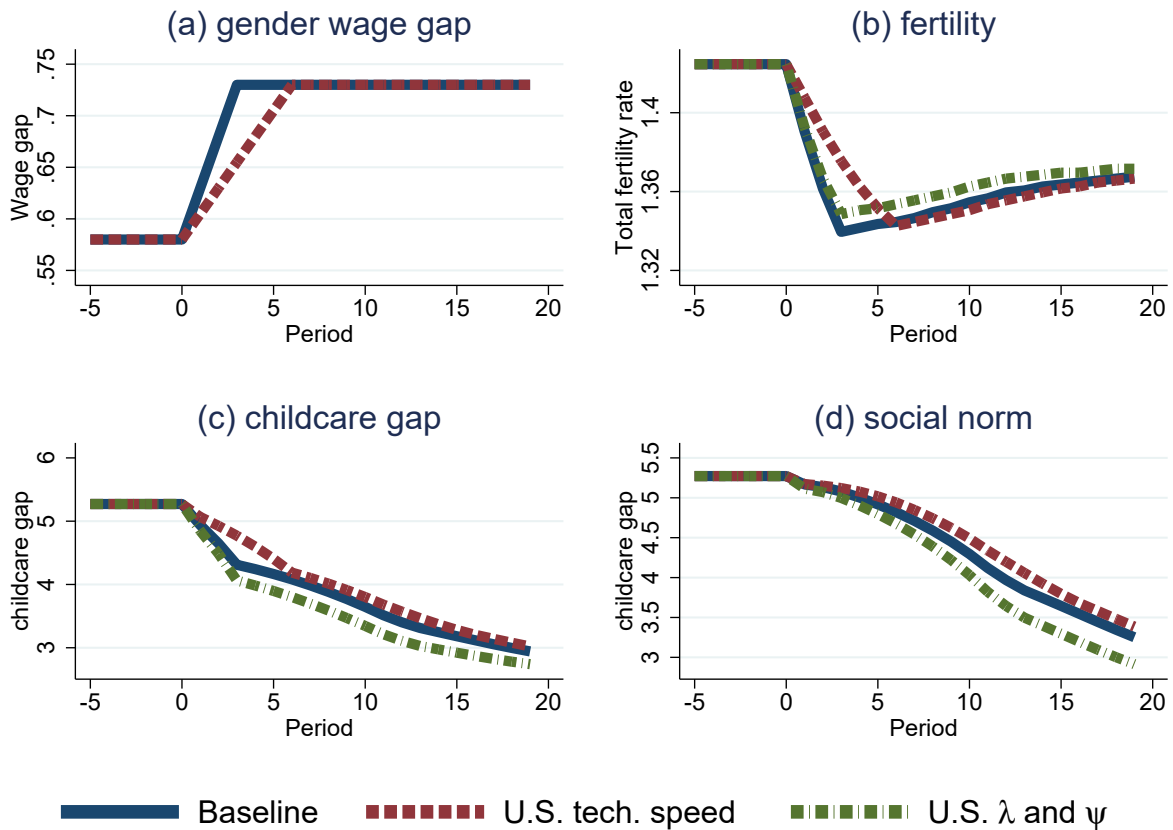


These findings, shown in Figure 5, underscore the re-evaluation by older cohorts as another important amplifier of technological shocks. Lower ψ facilitates a smoother transition by reducing the tension between rising female wages and lagging norms, while higher λ intensifies short-term fertility losses and prolongs inequitable childcare burdens.

5.4 U.S. Parameters

In this section, we examine the hypothetical fertility transition that would occur if we adjust the parameters calibrated for South Korea to those calibrated for the U.S. economy's transition from 1965 to 2015.⁷ We explore two counterfactual scenarios and compare them to the baseline case. In the first counterfactual, we adopt the speed of gender wage gap convergence observed in the U.S., which is half that of South Korea. In the second counterfactual, we modify the parameters to $\lambda = 0.0005$ and $\psi = 2.0$, reflecting lower social pressure and stubbornness, to align with U.S. estimates. The results are presented in Figure 5.

Figure 6: Using U.S. Parameters



On the one hand, with a slower convergence of the gender wage gap, the economy

⁷For details on the calibration process, refer to Section A.

experiences a delayed and more gradual transition in fertility, childcare gap, and social norms, aligning with the findings in Section 5.1.

On the other hand, with reduced social pressure (λ) and lower stubbornness among older cohorts (ψ), the economy avoids a sharp fertility decline—dropping to 1.35 rather than 1.33—while achieving a faster fertility recovery. This scenario also accelerates the transition of the childcare gap (panel c) and social norms (panel d) toward a new steady state, consistent with the findings in Sections 5.2 and 5.3.

5.5 Policy Evaluation

Finally, we evaluate a policy reducing the initial social norm by 1 unit (from 5.25 to 4.25 in 2004), simulating a campaign for egalitarian childcare, applied alongside the baseline 15-year wage convergence. Figure 7 compares this to the no-policy baseline.

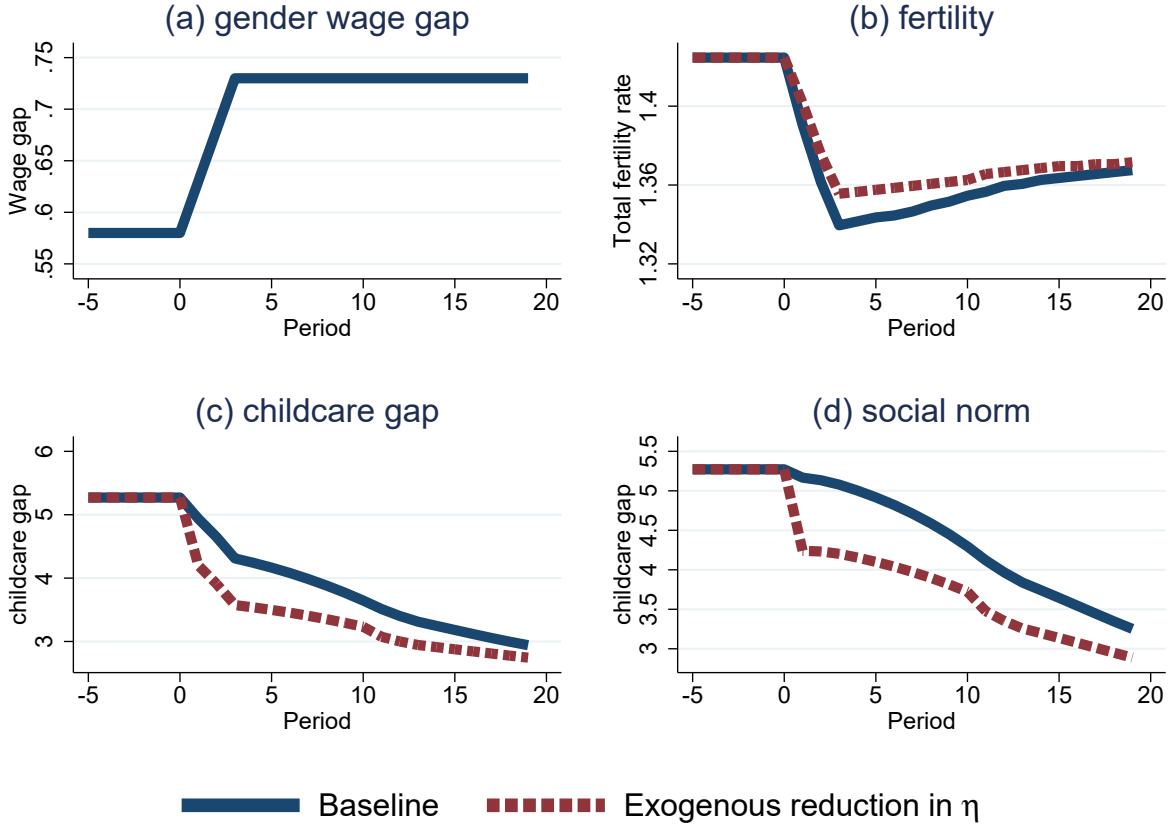
The policy lowers the social norm (panel d) immediately, easing the childcare gap's decline (panel c) to 4.2 by 2004 (vs. 5.0 without policy), as couples adjust allocations with reduced psychic costs. Fertility (panel b) is slightly higher in the short run (e.g., 1.40 vs. 1.39 by 2004). Over the medium term, the norm falls further (e.g., to 4.2 by 2014 vs. 5.1), boosting fertility to 1.36 (vs. 1.34 without the policy), reflecting larger cohorts' influence. These modest gains compound, amplifying long-term population differences.

Figure 7 illustrates these dynamics, highlighting the policy's role in breaking norm inertia. While short-run effects are modest due to persistent social norm, the medium- and long-run benefits, i.e., higher fertility and a larger population, suggest that norm-targeted interventions could effectively mitigate the adverse fertility consequences of rapid technological change. The difference between policy effects over time also highlights the importance of having a dynamic perspective in policy evaluation.

6. Robustness

This section discusses the robustness of our calibration and model specification.

Figure 7: Changing Initial Social Norm



6.1 Initial State

One limitation of our baseline calibration is the assumption that the initial gender gap in childcare responsibilities among young couples in 1999 (i.e., $l_{1999}^{\varnothing}/l_{1999}^{\sigma} = 5.25$) reflects the prevailing social norm, η_{1999} . Ideally, η_t should aggregate the opinions of old cohorts, as specified in Equation (11), which depends on historical childcare practices before 1999, but such data are unavailable to the best of our knowledge. To test the sensitivity of our results to this assumption, we experimented with initial norms more unequal than the observed 1999 value—e.g., setting $\eta_{1999} = 6.0$ or 6.5 , implying greater gender gaps among older generations. Alternatively, we have experimented a specification where we assume the economy was in the steady state in 1950 with a gender wage gap $w^{\varnothing}/w^{\sigma} = 0.3$, whereas the episode we observe between 1999 to 2014 is part of the

transition path. Recalibrating the model under these conditions, we find that the quantitative outcomes, including fertility paths (n_t), childcare allocations ($l_t^{\mathcal{Q}}/l_t^{\mathcal{S}}$), and norm evolution (η_t), remain consistent with the baseline. The counterfactual results also remain unaffected.

6.2 Rising Childcare Cost

A second concern is the potential rise in childcare costs over time, which could confound the effects of wage convergence on fertility. For instance, [Kim et al. \(2024\)](#) document increasing parental investment due to status competition, a trend that may elevate the time cost per child, ϕ . To explore this, we modify the model to allow ϕ to vary over the calibration period (1999–2014), calibrating its increase to match the rising parental childcare time reported by [Park \(2021\)](#).

Specifically, we assume ϕ_t grows linearly from 0.15 in 1999 to 0.20 by 2014, reflecting intensified child-rearing demands. Incorporating this time-varying ϕ_t requires adjustments to two parameters. First, the fertility curvature, ρ , rises from 2.4 to 3.4, as the observed fertility decline (e.g., from 1.42 to below 1.3) stems not only from productivity gains and wage gap convergence but also from higher childcare costs reducing the substitutability between consumption and children. Second, the fertility weight, γ , increases from 0.24 to 0.34 to maintain the initial fertility level ($n_{1999} = 1.42$) despite the elevated ϕ_{1999} . The value of other parameters—such as σ , ψ , and λ —remains unaffected. Critically, the model’s fit to the 1999–2014 fertility and childcare gap trajectories, as well as the counterfactual predictions (e.g., non-monotonic fertility responses), remains quantitatively similar to the baseline. The slightly higher ρ and γ amplify the fertility response to cost increases, but the core mechanisms—driven by wage gaps and norm dynamics—persist.

These robustness checks affirm that our key conclusions are not overly sensitive to the initial norm specification or the assumption of constant childcare costs. The model’s ability to accommodate these variations while preserving its predictive power underscores its reliability for analyzing the interplay of technology, norms, and fertility.

7. Conclusion

Fertility rates are declining globally, posing significant economic and demographic challenges. This paper documents an important pattern: fertility falls more precipitously in economies undergoing rapid structural change, with the decline amplified in societies where rigid social norms govern childcare responsibilities. To unpack this relationship, we develop a novel quantitative model of fertility bargaining, where equilibrium outcomes emerge from the interplay between technological progress—manifested as rising female wages—and endogenous social norms shaped by past cohorts’ opinions, which hinge on the historical childcare practices.

Calibrated to South Korea’s experience from 1999 to 2014, our model closely matches the observed trajectories of fertility and childcare gaps, capturing the tension between wage convergence and norm inertia. Counterfactual analyses highlight the pivotal role of adjustment speed: rapid technological change outpaces norm evolution, driving sharper fertility drops, while weaker social pressure, more re-evaluation by older cohorts, or norm-shifting policies mitigate these declines and expedite fertility recovery.

While the model is calibrated to match South Korea’s growth experience, we argue that its lessons offer valuable guidance for many developing countries navigating, or on the cusp of, rapid structural transformation. Reconciling technological progress with entrenched social norms is essential to tackling the “getting old before getting rich” dilemma these countries face.

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A. Calibration for the United States

In this section, we calibrate the model to match the transition path of fertility and the gender childcare gap observed in the United States from 1965 to 2015. The identification strategy follows from Section 4.1.

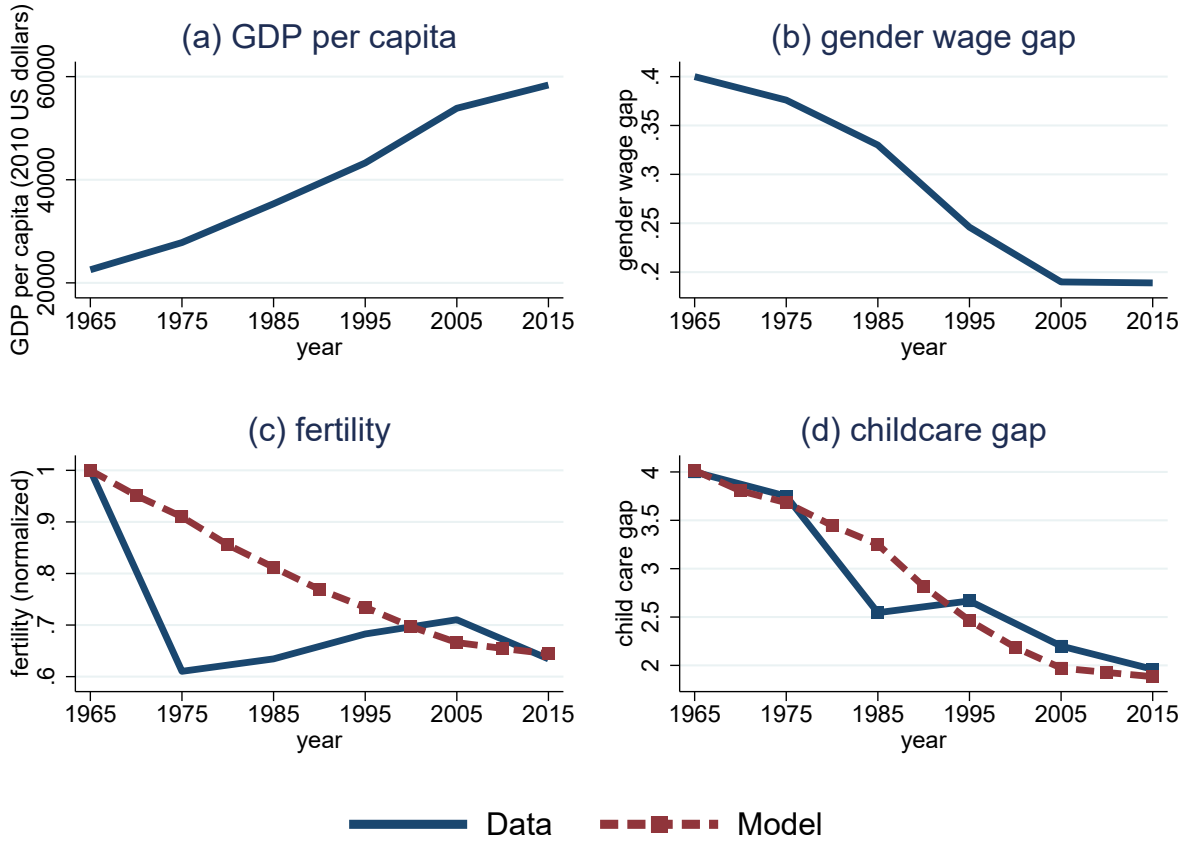
We collect fertility data from the United Nations. Gender-specific childcare time between 1965 and 1995 is collected from Egerton et al. (2005) who utilizes the micro-level data from the American Heritage Time Use Study (AHTUS). After 2003, gender-specific childcare time is calculated using the micro-level data from the American Time Use Survey (ATUS) following the strategy of Milkie et al. (2025). Lastly, we use the micro-level data from the General Social Survey (GSS) to compute the share of within-cohort effects in driving social norm changes. We adopt the same method as described in Section 4.1 and examine the variable “*fefam*: It is much better for everyone involved if the man is the achiever outside the home and the woman takes care of the home and family.” We find that the ratio of within-cohort change divided by total change is close to 0.3 over the course of 50 years, similar to the decomposition result by Brooks and Bolzendahl (2004).

Table 5: Calibrated Parameters - USA

	Parameter	Value	Data moment	Source	Model fit
γ	Fertility weight	1.27	$n_{1965} = 2.90$	United Nations	2.90
σ	Childcare substitutability	2.73	$\eta_{1965} = 4.0$	Egerton et al. (2005)	4.0
ρ	Fertility curvature	2.4	$n_{1965} \sim n_{2015}$	United Nations	See Figure 8
ψ	Stubbornness	2.0	Within-cohort effects	GSS	30%
λ	Social pressure	0.0005	$\eta_{1965} \sim \eta_{2015}$	Egerton et al. (2005)	See Figure 8
α	Economies of scale	1.2		Doepke and Kindermann (2019)	
ϕ	Time costs per child	0.15		de La Croix and Doepke (2003)	
J	Total number of periods	16	80 years	World Health Organization	
J_f	The fertile period	6	25 to 30 yo	Statista	

Figure 8 plots the time path of GDP per capita and gender wage gap as inputs into the transition path calibration. Panel (c) and (d) of 8 present the data and model-generated

Figure 8: Calibration and Model Fit - USA



path of fertility and childcare gap. As can be seen, while the model does not fully account for the decline and recovery of fertility after the baby boom, it fits the overall decline in fertility and the path of childcare gap in the study period very well.⁸

Table 5 displays the calibrated parameters for the U.S. economy. Except for the parameter γ which governs the level of fertility, parameters σ and ρ are very similar to the results in South Korea (see Table 3). The parameters governing the social norm evolution, however, are quite different. The value for λ in the U.S. is 17% smaller than that in South Korea, indicating a smaller magnitude of social pressure. In addition, the value for ψ in the U.S. is 33% smaller than that in South Korea, implying a smaller degree of stubbornness and hence greater re-evaluation by older cohorts.

⁸See Greenwood et al. (2005) for an in-depth study of the U.S. baby boom and its aftermath.