

7.1

Derive the estimate of in-sample error (7.24).

It suffices to show that

$$\sum_{i=1}^N \text{Cov}(\hat{y}_i, y_i) = d\sigma_\epsilon^2.$$

Note that for a linear fit, we have $\hat{\mathbf{y}} = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$, so

$$\begin{aligned} \text{Cov}(\hat{\mathbf{y}}, \mathbf{y}) &= \text{Cov}(\mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}, \mathbf{y}) \\ &= \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \text{Cov}(\mathbf{y}, \mathbf{y}) \\ &= \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \sigma_\epsilon^2. \end{aligned}$$

Therefore, by *cyclic property* of trace operator,

$$\begin{aligned} \sum_{i=1}^N \text{Cov}(\hat{y}_i, y_i) &= \text{trace}(\mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T) \sigma_\epsilon^2 \\ &= \text{trace}(\mathbf{X}^T \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1}) \sigma_\epsilon^2 \\ &= \text{trace}(\mathbf{I}_d) \sigma_\epsilon^2 \\ &= d\sigma_\epsilon^2. \end{aligned}$$

7.3

Let $\hat{\mathbf{f}} = \mathbf{S}\mathbf{y}$ be a linear smoothing of \mathbf{y} .

(a)

If S_{ii} is the i th diagonal element of \mathbf{S} , show that for \mathbf{S} arising from least squares projections and cubic smoothing splines, the cross-validated residual can be written as

$$y_i - \hat{f}^{-i}(x_i) = \frac{y_i - \hat{f}(x_i)}{1 - S_{ii}}. \quad (1)$$

(b)

Use this result to show that $|y_i - \hat{f}^{-i}(x_i)| \geq |y_i - \hat{f}(x_i)|$.

(c)

Find general conditions on any smoother \mathbf{S} to make result (1) hold.

Without loss of generality, we assume

$$\mathbf{S} = \mathbf{X}(\mathbf{X}^T \mathbf{X} + \lambda \mathbf{\Omega})^{-1} \mathbf{X}^T.$$

For least squares we have $\lambda = 0$, and for cubic smoothing we have $\lambda \geq 0$. See Chapters 3 & 5 in the text for more details.

(a)

We have

$$\begin{aligned} S_{ii} &= \mathbf{x}_i^T (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{\Omega})^{-1} \mathbf{x}_i, \\ \hat{f}(x_i) &= \mathbf{x}_i^T (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{\Omega})^{-1} \mathbf{X}^T \mathbf{y}. \end{aligned}$$

Let \mathbf{X}_{-i} and \mathbf{y}_{-i} be the corresponding results with x_i removed, then we have

$$\begin{aligned} \hat{f}^{-i}(x_i) &= \mathbf{x}_i^T (\mathbf{X}_{-i}^T \mathbf{X}_{-i} + \lambda \mathbf{\Omega})^{-1} \mathbf{X}_{-i}^T \mathbf{y}_{-i} \\ &= \mathbf{x}_i^T (\mathbf{X}^T \mathbf{X} - \mathbf{x}_i \mathbf{x}_i^T + \lambda \mathbf{\Omega})^{-1} (\mathbf{X}^T \mathbf{y} - \mathbf{x}_i y_i). \quad (2) \end{aligned}$$

Let $\mathbf{A} = (\mathbf{X}^T \mathbf{X} + \lambda \mathbf{\Omega})$, by [Woodbury matrix identity](#), we have

$$(\mathbf{A} - \mathbf{x}_i \mathbf{x}_i^T)^{-1} = \mathbf{A}^{-1} + \frac{\mathbf{A}^{-1} \mathbf{x}_i \mathbf{x}_i^T \mathbf{A}^{-1}}{1 - \mathbf{x}_i^T \mathbf{A}^{-1} \mathbf{x}_i}.$$

Therefore, (2) becomes

$$\begin{aligned} \hat{f}^{-i}(x_i) &= \mathbf{x}_i^T \left(\mathbf{A}^{-1} + \frac{\mathbf{A}^{-1} \mathbf{x}_i \mathbf{x}_i^T \mathbf{A}^{-1}}{1 - \mathbf{x}_i^T \mathbf{A}^{-1} \mathbf{x}_i} \right) (\mathbf{X}^T \mathbf{y} - \mathbf{x}_i y_i) \\ &= \left(\mathbf{x}_i^T \mathbf{A}^{-1} + \frac{S_{ii} \mathbf{x}_i^T \mathbf{A}^{-1}}{1 - S_{ii}} \right) (\mathbf{X}^T \mathbf{y} - \mathbf{x}_i y_i) \\ &= \mathbf{x}_i^T \mathbf{A}^{-1} \mathbf{X}^T \mathbf{y} - \mathbf{x}_i^T \mathbf{A}^{-1} \mathbf{x}_i y_i + \frac{S_{ii} \mathbf{x}_i^T \mathbf{A}^{-1} \mathbf{X}^T \mathbf{y}}{1 - S_{ii}} - \frac{S_{ii} \mathbf{x}_i^T \mathbf{A}^{-1} \mathbf{x}_i y_i}{1 - S_{ii}} \\ &= \hat{f}(x_i) - y_i S_{ii} + \frac{S_{ii} \hat{f}(x_i)}{1 - S_{ii}} - \frac{y_i S_{ii}^2}{1 - S_{ii}} \\ &= \frac{\hat{f}(x_i) - y_i S_{ii}}{1 - S_{ii}}. \end{aligned}$$

Therefore by simple algebra we have (1).

(b)

Note that $\mathbf{S} = \mathbf{X}(\mathbf{X}^T \mathbf{X} + \lambda \mathbf{\Omega})^{-1} \mathbf{X}^T$ is positive-semidefinite and has eigen-decomposition

$$\mathbf{S} = \sum_{k=1}^N \rho_k(\lambda) \mathbf{u}_k \mathbf{u}_k^T.$$

See Section 5.4.1 in the text for more details. Therefore, we know that $\mathbf{S}\mathbf{S} \preceq \mathbf{S}$, so that

$$0 \leq (S^2)_{ii} = \sum_{k \neq i} S_{ik}^2 + S_{ii}^2 \leq S_{ii}$$

from which we know $0 \leq S_{ii} \leq 1$.

By (1) we have $|y_i - \hat{f}^{-i}(x_i)| \geq |y_i - \hat{f}(x_i)|$.

(c)

For general linear smoother $\hat{\mathbf{f}} = \mathbf{S}\mathbf{y}$, if \mathbf{S} only depends on \mathbf{X} and other tuning parameters (i.e., independent of \mathbf{y}), (1) still holds.

To see that, note that if we replace y_i with $\hat{f}^{-i}(x_i)$ (obtained by (2)) in \mathbf{y} and denote the new vector by \mathbf{y}' , \mathbf{S} is not changed. Thus we have

$$\begin{aligned} \hat{f}^{-i}(x_i) &= (\mathbf{S}\mathbf{y}')_i \\ &= \sum_{j \neq i} S_{ij} y'_j + S_{ii} \hat{f}^{-i}(x_i) \\ &= \hat{f}(x_i) - S_{ii} y_i + S_{ii} \hat{f}^{-i}(x_i), \end{aligned}$$

therefore we obtain (1).

7.4

Consider the in-sample prediction error (7.18) and the training error $\overline{\text{err}}$ in the case of squared-error loss:

$$\begin{aligned} \text{Err}_{\text{in}} &= \frac{1}{N} \sum_{i=1}^N E_{Y^0} (Y_i^0 - \hat{f}(x_i))^2 \\ \overline{\text{err}} &= \frac{1}{N} \sum_{i=1}^N (y_i - \hat{f}(x_i))^2. \end{aligned}$$

Add and subtract $f(x_i)$ and $E\hat{f}(x_i)$ in each expression and expand. Hence establish that the average optimism in the training error is

$$\frac{2}{N} \sum_{i=1}^N \text{Cov}(\hat{y}_i, y_i),$$

as given in (7.21).

We start with Err_{in} . Let's denote $\hat{y}_i = \hat{f}(x_i)$ and write

$$Y_i^0 - \hat{f}(x_i) = Y_i^0 - f(x_i) + f(x_i) - E\hat{y}_i + E\hat{y}_i - \hat{y}_i$$

so that

$$\begin{aligned}\text{Err}_{\text{in}} &= \frac{1}{N} \sum_{i=1}^N E_{Y^0} (Y_i^0 - f(x_i) + f(x_i) - E\hat{y}_i + E\hat{y}_i - \hat{y}_i)^2 \\ &= \frac{1}{N} \sum_{i=1}^N A_i + B_i + C_i + D_i + E_i + F_i,\end{aligned}$$

where

$$\begin{aligned}A_i &= E_{Y^0} (Y_i^0 - f(x_i))^2 \\ B_i &= E_{Y^0} (f(x_i) - E\hat{y}_i)^2 = (f(x_i) - E\hat{y}_i)^2 \\ C_i &= E_{Y^0} (E\hat{y}_i - \hat{y}_i)^2 = (E\hat{y}_i - \hat{y}_i)^2 \\ D_i &= 2E_{Y^0} (Y_i^0 - f(x_i))(f(x_i) - E\hat{y}_i) \\ E_i &= 2E_{Y^0} (Y_i^0 - f(x_i))(E\hat{y}_i - \hat{y}_i) \\ F_i &= 2E_{Y^0} (f(x_i) - E\hat{y}_i)(E\hat{y}_i - \hat{y}_i) = 2(f(x_i) - E\hat{y}_i)(E\hat{y}_i - \hat{y}_i)\end{aligned}$$

Similarly for $\overline{\text{err}}$ we have

$$y_i - \hat{f}(x_i) = y_i - f(x_i) + f(x_i) - E\hat{y}_i + E\hat{y}_i - \hat{y}_i$$

and

$$\begin{aligned}\overline{\text{err}} &= \frac{1}{N} \sum_{i=1}^N (y_i - f(x_i) + f(x_i) - E\hat{y}_i + E\hat{y}_i - \hat{y}_i)^2 \\ &= \frac{1}{N} \sum_{i=1}^N G_i + B_i + C_i + H_i + J_i + F_i,\end{aligned}$$

where

$$\begin{aligned}G_i &= (y_i - f(x_i))^2 \\ H_i &= 2(y_i - f(x_i))(f(x_i) - E\hat{y}_i) \\ J_i &= 2(y_i - f(x_i))(E\hat{y}_i - \hat{y}_i).\end{aligned}$$

Therefore, we have

$$\begin{aligned} E_{\mathbf{y}}(\text{op}) &= E_{\mathbf{y}}(\text{Err}_{\text{in}} - \overline{\text{err}}) \\ &= \frac{1}{N} \sum_{i=1}^N E_{\mathbf{y}}[(A_i - G_i) + (D_i - H_i) + (E_i - J_i)]. \end{aligned}$$

For A_i and G_i , the expectation over \mathbf{y} captures unpredictable error and thus $E_{\mathbf{y}}(A_i - G_i) = 0$. Similarly we have $E_{\mathbf{y}}D_i = E_{\mathbf{y}}H_i = E_{\mathbf{y}}E_i = 0$, and thus

$$\begin{aligned} E_{\mathbf{y}}(\text{op}) &= -\frac{2}{N} \sum_{i=1}^N J_i \\ &= -\frac{2}{N} \sum_{i=1}^N E_{\mathbf{y}}(y_i - f(x_i))(E\hat{y}_i - \hat{y}_i) \\ &= \frac{2}{N} \sum_{i=1}^N [E_{\mathbf{y}}(y_i \hat{y}_i) - E_{\mathbf{y}}y_i E_{\mathbf{y}}\hat{y}_i] \\ &= 2\text{Cov}(y_i, \hat{y}_i). \end{aligned}$$

7.5

For a linear smoother $\hat{\mathbf{y}} = \mathbf{S}\mathbf{y}$, show that

$$\sum_{i=1}^N \text{Cov}(\hat{y}_i, y_i) = \text{trace}(\mathbf{S})\sigma_{\epsilon}^2,$$

which justifies its use as the effective number of parameters.

$$\begin{aligned} \sum_{i=1}^N \text{Cov}(\hat{y}_i, y_i) &= \text{trace}(\text{Cov}(\hat{\mathbf{y}}, \mathbf{y})) \\ &= \text{trace}(\text{Cov}(\mathbf{S}\mathbf{y}, \mathbf{y})) \\ &= \text{trace}(\mathbf{S}\text{Cov}(\mathbf{y}, \mathbf{y})) \\ &= \text{trace}(\mathbf{S}\text{Var}(\mathbf{y})) \\ &= \text{trace}(\mathbf{S})\sigma_{\epsilon}^2. \end{aligned}$$