

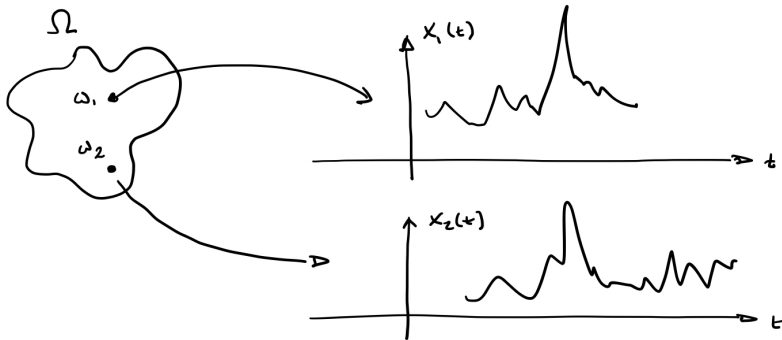


**LUND**  
UNIVERSITY

# **Stochastic processes**

Andreas Jakobsson

## Stochastic processes



A stochastic variable may have many dimensions, say  $N$  dimensions, so that

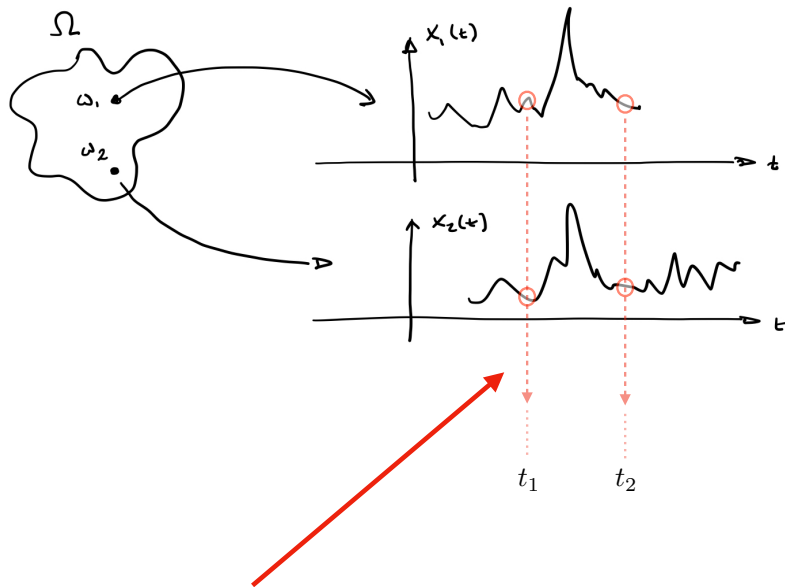
$$z = [x_1 \quad \dots \quad x_N]^T,$$

making the realisation a point in a  $N$ -D space. If we let  $N$  grow, making each realisation a function instead, we have a *stochastic process*.

From a practical perspective, you always only have limited amount of data, say  $N$  samples, which may be viewed as observing  $N$  samples of a stochastic process (or as an  $N$ -D stochastic variable).

As a stochastic process is a function along one (or more) variables, we often use the notation  $x_t$ , or  $x(t)$ , to indicate that it is a function of  $t$ . We will here use the same notation for the corresponding realisation.

# Stochastic processes



A stochastic variable may have many dimensions, say  $N$  dimensions, so that

$$z = [x_1 \quad \dots \quad x_N]^T,$$

making the realisation a point in a  $N$ -D space. If we let  $N$  grow, making each realisation a function instead, we have a *stochastic process*.

From a practical perspective, you always only have limited amount of data, say  $N$  samples, which may be viewed as observing  $N$  samples of a stochastic process (or as an  $N$ -D stochastic variable).

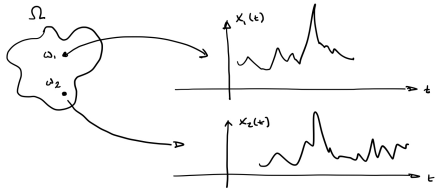
As a stochastic process is a function along one (or more) variables, we often use the notation  $x_t$ , or  $x(t)$ , to indicate that it is a function of  $t$ . We will here use the same notation for the corresponding realisation.

The mean function,  $m_x(t_1) = E\{x_{t_1}\}$ , is the mean value of the process at  $t_1$ .

The variance function,  $v_x(t_1) = V\{x_{t_1}\}$ , is the variance of the process at  $t_1$ .

The covariance function,  $r_x(t_1, t_2) = C\{x_{t_1}, x_{t_2}\}$ , is the dependence between realisations at different times.

# Stochastic processes

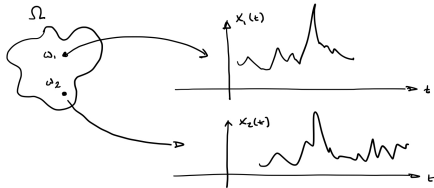


We will restrict our attention to *wide-sense stationary* (WSS) processes. For such processes, the statistical properties do not change over time. Furthermore,

- (i) The mean function is constant and finite,  $m_x(t) = m_x < \infty$ .
- (ii) The (auto-)covariance,  $C\{x_t, x_s\}$ , only depends on the difference  $(s - t)$  and *not* on the actual values of  $s$  and  $t$ .
- (iii) The variance of the process is finite, i.e.,  $E\{|x_t|^2\} < \infty$ .

Furthermore, we will assume that all considered processes are *ergodic*. Essentially, this implies that it is possible to estimate the characteristics of the process from a single realisation.

# Stochastic processes



We will restrict our attention to *wide-sense stationary* (WSS) processes. For such processes, the statistical properties do not change over time. Furthermore,

- (i) The mean function is constant and finite,  $m_x(t) = m_x < \infty$ .
- (ii) The (auto-)covariance,  $C\{x_t, x_s\}$ , only depends on the difference  $(s - t)$  and *not* on the actual values of  $s$  and  $t$ .
- (iii) The variance of the process is finite, i.e.,  $E\{|x_t|^2\} < \infty$ .

Furthermore, we will assume that all considered processes are *ergodic*. Essentially, this implies that it is possible to estimate the characteristics of the process from a single realisation.

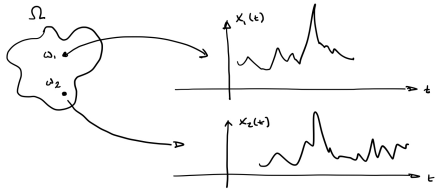
*Example:*

Let  $x_t = A \cos(2\pi ft + \phi)$ , where  $\phi$  is uniformly distributed on  $[0, 2\pi)$  and  $A$  and  $\phi$  are independent. Then,

$$\begin{aligned}
 E\{x_t\} &= E\{A \cos(2\pi ft + \phi)\} = E\{A\}E\{\cos(2\pi ft + \phi)\} \\
 &= m_A \int_{-\infty}^{\infty} \cos(2\pi ft + \phi) f_{\phi}(\phi) d\phi \\
 &= \frac{m_A}{2\pi} \int_0^{2\pi} \cos(2\pi ft + \phi) d\phi \\
 &= \frac{m_A}{2\pi} [\sin(2\pi ft + \phi)]_0^{2\pi} = 0
 \end{aligned}$$

The average value of this process is zero for all  $t$ . This satisfies (i). How about (ii) and (iii)?

# Stochastic processes



We will restrict our attention to *wide-sense stationary* (WSS) processes. For such processes, the statistical properties do not change over time. Furthermore,

- (i) The mean function is constant and finite,  $m_x(t) = m_x < \infty$ .
- (ii) The (auto-)covariance,  $C\{x_t, x_s\}$ , only depends on the difference  $(s - t)$  and *not* on the actual values of  $s$  and  $t$ .
- (iii) The variance of the process is finite, i.e.,  $E\{|x_t|^2\} < \infty$ .

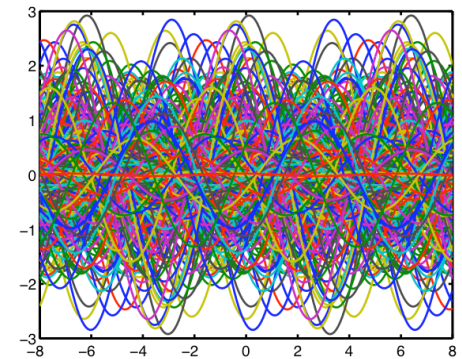
Furthermore, we will assume that all considered processes are *ergodic*. Essentially, this implies that it is possible to estimate the characteristics of the process from a single realisation.

*Example:*

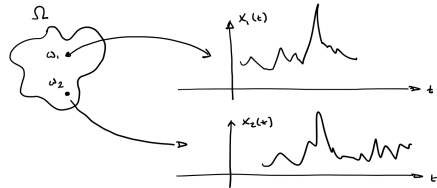
Let  $x_t = A \cos(2\pi ft + \phi)$ , where  $\phi$  is uniformly distributed on  $[0, 2\pi)$  and  $A$  and  $\phi$  are independent. Then,

$$\begin{aligned}
 E\{x_t\} &= E\{A \cos(2\pi ft + \phi)\} = E\{A\}E\{\cos(2\pi ft + \phi)\} \\
 &= m_A \int_{-\infty}^{\infty} \cos(2\pi ft + \phi) f_{\phi}(\phi) d\phi \\
 &= \frac{m_A}{2\pi} \int_0^{2\pi} \cos(2\pi ft + \phi) d\phi \\
 &= \frac{m_A}{2\pi} [\sin(2\pi ft + \phi)]_0^{\pi} = 0
 \end{aligned}$$

The average value of this process is zero for all  $t$ . This satisfies (i). How about (ii) and (iii)?



# Stochastic processes



*Example:*

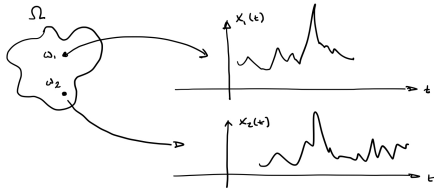
According to (ii), for a WSS process  $C\{x_s, x_t\}$  should only depend on  $(s - t)$ .  
Let's check!

$$\begin{aligned}
 C\{x_s, x_t\} &= E\{A^2 \cos(2\pi f s + \phi) \cos(2\pi f t + \phi)\} \\
 &= E\{A^2\} E\{\cos(2\pi f s + \phi) \cos(2\pi f t + \phi)\} \\
 &= \frac{E\{A^2\}}{2} E\{\cos(2\pi f(s + t) + 2\phi) + \cos(2\pi f(s - t))\} \\
 &= \frac{E\{A^2\}}{4\pi} \int_0^{2\pi} \cos(2\pi f(s + t) + 2\phi) + \cos(2\pi f(s - t)) d\phi \\
 &= \frac{E\{A^2\}}{4\pi} \cos(2\pi f(s - t))
 \end{aligned}$$

The covariance thus only depends on  $(s - t)$ , with  $C\{x_s, x_s\} < \infty$ . This is thus a WSS process!

It is worth noting that the covariance function of a sinusoid is itself a sinusoid, with the same frequency!

# Stochastic processes



The autocovariance function only depends on the difference between  $s$  and  $t$ . To simplify notation, we use  $r_x(t, t - k) = r_x(k)$ , i.e.,

$$r_x(k) = C\{x_t, x_{t-k}\} = E\{x_t x_{t-k}^*\} - m_x m_x^*$$

with the variance being  $V\{x_t\} = r_x(0)$ .

The autocovariance function of a WSS process satisfies

- (i) It is conjugate symmetric, i.e.,  $r_x(k) = r_x^*(-k)$ .
- (ii) The variance is always non-negative, i.e.,  $r_x(0) \geq 0$ .
- (iii) The function takes it largest values at lag 0, i.e.,

$$r_x(0) \geq |r_x(k)|, \quad \forall k$$

*Example:*

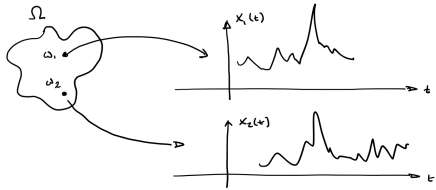
Returning to the sinusoidal process,  $x_t = A \cos(2\pi f t + \phi)$ , we note that

$$r_x(k) = C\{x_t, x_{t-k}\} = \frac{E\{A^2\}}{4\pi} \cos(2\pi f k)$$

As  $\cos(\cdot)$  is symmetric, with its maximum at  $k = 0$ , we satisfy (i) and (iii). Also,  $r_x(0) \geq 0$ , so (ii) is also satisfied. If  $r_x(k) = r_x(0)$  for any  $k \neq 0$ , the process is always periodic.



# Stochastic processes



The auto-correlation function (ACF) is defined as

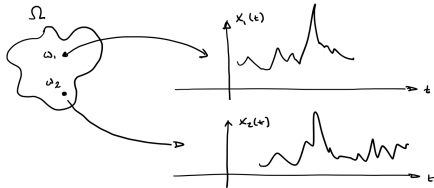
$$\rho_x(k) = \frac{r_x(k)}{r_x(0)}$$

which is thus bounded as  $|\rho_x(k)| \leq 1$ . Similarly, we define the cross-correlation function between  $x_t$  and  $y_t$  as

$$\rho_{x,y}(k) = \frac{r_{x,y}(k)}{\sqrt{r_x(0)}\sqrt{r_y(0)}}$$

where  $r_{x,y}(k) = C\{x_t, y_{t-k}\}$ . Show that  $|\rho_{x,y}(k)| \leq 1$ .

# Stochastic processes



The auto-correlation function (ACF) is defined as

$$\rho_x(k) = \frac{r_x(k)}{r_x(0)}$$

which is thus bounded as  $|\rho_x(k)| \leq 1$ . Similarly, we define the cross-correlation function between  $x_t$  and  $y_t$  as

$$\rho_{x,y}(k) = \frac{r_{x,y}(k)}{\sqrt{r_x(0)}\sqrt{r_y(0)}}$$

where  $r_{x,y}(k) = C\{x_t, y_{t-k}\}$ . Show that  $|\rho_{x,y}(k)| \leq 1$ .

Recall that for two random variables  $x$  and  $y$ , it holds that

$$V\{ax + by\} = |a|^2 V\{x\} + |b|^2 V\{y\} + (a^*b + ab^*)C\{x, y\}$$

As the variance is non-negative,  $V\{ax + by\} \geq 0$ . Thus, if  $a, b \in \mathbb{R}$ ,

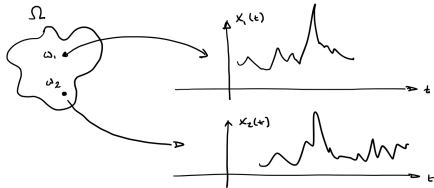
$$V\{ax_t + by_{t-k}\} = a^2 r_x(0) + b^2 r_y(0) + 2ab r_{x,y}(k) \geq 0$$

Let  $a = 1/\sqrt{r_x(0)}$  and  $b = \pm 1/\sqrt{r_y(0)}$ , then

$$1 + 1 \pm 2\rho_{x,y}(k) \geq 0$$

Thus,  $|\rho_{x,y}(k)| \leq 1$ .

# Stochastic processes



*Example:*

We want to determine the periodic heartbeat of a fetus in its mothers womb. As the mothers heart is also beating, we should expect our signal to be of the form

$$x_t = A_f \cos(2\pi f_f t + \phi_f) + A_m \cos(2\pi f_m t + \phi_m)$$

where  $A_f$  and  $f_f$  denote the amplitude and frequency of the fetus' heartbeat, whereas  $A_m$  and  $f_m$  are those of the mother. Clearly,  $A_m \gg A_f$ .

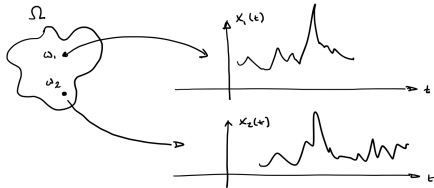
The autocovariance of  $x_t$  will thus be

$$r_x(k) = \frac{E\{A_f^2\}}{4\pi} \cos(2\pi f_f k) + \frac{E\{A_m^2\}}{4\pi} \cos(2\pi f_m k)$$

Thus, if we can estimate  $r_x(k)$ , we can use this to estimate the fetus' heartbeat!

In reality, the measured signal will not consist of only sinusoids; all forms of measurement are corrupted by noise. How can we deal with this?

# White noise



To model various forms of measurement noise, we make use of a particular form of process, termed a *white noise*. For such a process it holds that

$$r_x(k) = \begin{cases} \sigma_x^2 & \text{for } k = 0 \\ 0 & \text{for } k \neq 0 \end{cases}$$

where  $\sigma_x^2$  denotes the variance of the process. The process is thus such completely uncorrelated one sample to the next!

Often, one also assumes that each value of the process has a Gaussian distribution. If such a noise is assumed to be added to the relevant signal of interest, one then speaks of a *additive Gaussian white noise* (AGWN).

*Example:*

Returning to the fetus data, we will assume that the measured signal is

$$y_t = x_t + w_t$$

where  $w_t$  is an AWGN with variance  $\sigma_w^2$ . Then,

$$r_y(k) = r_x(k) + r_w(k) = r_x(k) + \sigma_w^2 \delta_K(k)$$

with  $\delta_K(k)$  denotes the Kronecker delta, i.e.,

$$\delta_K(k) = \begin{cases} 1 & \text{for } k = 0 \\ 0 & \text{for } k \neq 0 \end{cases}$$