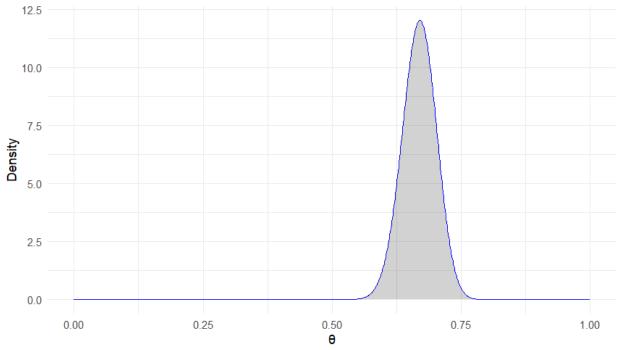
# CGS698 Assignment-3

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Part 1: Estimating the posterior distribution using different computational methods

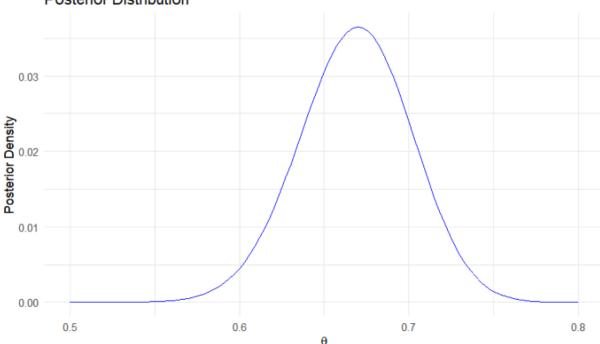
```
1.
> library(ggplot2)
> alpha <- 135
> beta <- 67
> theta_values <- seq(0, 1, length.out = 1000)
> posterior_density <- dbeta(theta_values,alpha,beta)
> data <- data.frame(theta = theta_values, density = posterior_density)
> ggplot(data, aes(x = theta, y = density)) +
+ geom_line(color = 'blue') +
+ geom_area( alpha = 0.2) +
+ labs(title = 'Posterior Distribution ', x = '0', y = 'Density') +
+ theme_minimal()
```

#### Posterior Distribution



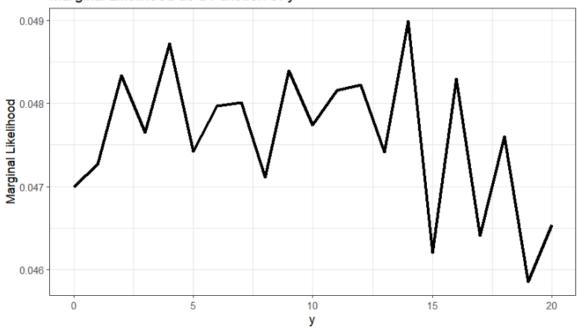
```
> y < -c(10, 15, 15, 14, 14, 14, 13, 11, 12, 16)
> theta grid <- seq(0.5, 0.8, length.out = 100)
> df.posterior <- data.frame(theta = theta grid, likelihood = NA, prior = NA,
posterior =
> for (i in 1:length(theta grid)) {
+ likelihood <- prod(dbinom(y, size = 20, prob = theta grid[i]))
+ prior <- dbeta(theta grid[i], shape1 = 1, shape2 = 1)
+ df.posterior$likelihood[i] <- likelihood</pre>
+ df.posterior$prior[i] <- prior</pre>
+ df.posterior$posterior[i] <- likelihood * prior</pre>
> df.posterior$posterior <- df.posterior$posterior /</pre>
sum(df.posterior$posterior)
> ggplot(df.posterior, aes(x = theta, y = posterior)) +
+ geom line(color = "blue") +
+ labs(title = "Posterior Distribution",
        x = expression(theta),
        y = "Posterior Density") +
+ theme minimal()
```

#### Posterior Distribution



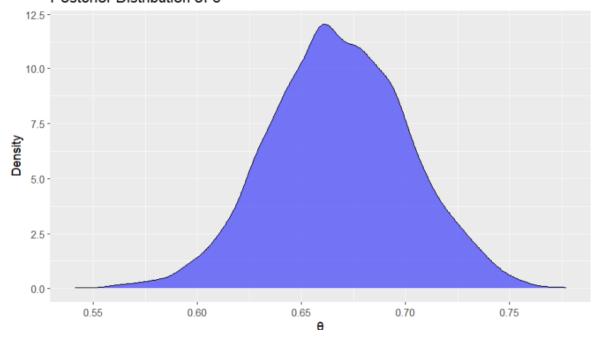
```
3.
> library(ggplot2)
> set.seed(123)
> y values <- 0:20
> num_samples <- 10000</pre>
> estimate marginal likelihood <- function(y) {</pre>
+ df.estimate <- data.frame(matrix(ncol = 2, nrow = num_samples))
 colnames(df.estimate) <- c("theta sample", "likelihood")</pre>
+ for(i in 1:num samples) {
     theta i <- rbeta(1, 1, 1)
    likelihood <- dbinom(y, size = 20, prob = theta i)</pre>
     df.estimate[i, ] <- c(theta i, likelihood)</pre>
  ML <- mean(df.estimate$likelihood)</pre>
  return(ML)
+ }
> ML_values <- sapply(y_values, estimate_marginal_likelihood)</pre>
> results <- data.frame(y = y values, ML = ML values)</pre>
> ggplot(results, aes(x = y, y = ML)) +
+ geom line(size = 1.2) +
+ theme bw() +
+ xlab("y") +
+ ylab("Marginal Likelihood") +
+ ggtitle("Marginal Likelihood as a Function of y")
```

#### Marginal Likelihood as a Function of y

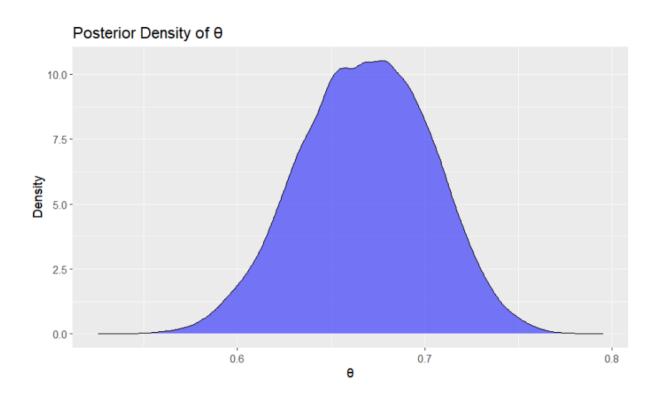


```
4.
> library(ggplot2)
> set.seed(123)
> y <- c(10, 15, 15, 14, 14, 14, 13, 11, 12, 16)
> n < -20
> num samples <- 10000
> df.estimate <- data.frame(matrix(ncol = 5, nrow = num samples))</pre>
> colnames(df.estimate) <- c("theta i", "likelihood", "prior", "proposal",</pre>
"posterior")
> for (i in 1:num samples) {
+ while (theta i < 0 \mid \mid theta i > 1) {
    theta i <- rnorm(1, 0.67, 0.2)
+ }
+ likelihood <- prod(dbinom(y, n, theta i))
+ prior <- dbeta(theta i, 1, 1)
+ posterior <- (likelihood * prior) / proposal
+ df.estimate[i, ] <- c(theta i, likelihood, prior, proposal, posterior)
+ for (i in 1:num samples) {
+ theta i < - rnorm(1, 0.67, 0.2)
+ while (theta i < 0 \mid \mid theta i > 1) {
     theta i <- rnorm(1, 0.67, 0.2)
+ likelihood <- prod(dbinom(y, n, theta i))
+ prior <- dbeta(theta i, 1, 1)
+ proposal <- dnorm(theta i, 0.67, 0.2)
+ posterior <- (likelihood * prior) / proposal
+ df.estimate[i, ] <- c(theta i, likelihood, prior, proposal, posterior)
+ }
+ df.estimate <- df.estimate[complete.cases(df.estimate), ]</pre>
+ df.estimate$posterior <- df.estimate$posterior / sum(df.estimate$posterior)
+ sample size <- num samples / 4
+ theta samples <- sample(df.estimate$theta i, size = sample size, prob =
                            df.estimate$posterior, replace = TRUE)
+ df.sample <- data.frame(theta samples)
+ ggplot(df.sample, aes(x = theta samples)) +
+ geom density(fill = "blue", alpha = 0.5) +
+ labs(title = "Posterior Distribution of \theta", x = "\theta", y = "Density")
+ hist(theta samples, breaks = 30, main = "Histogram of \theta Samples", xlab = "\theta",
col =
        "blue", border = "black")
```

#### Posterior Distribution of θ



```
5.
> library(ggplot2)
> set.seed(123)
> y <- c(10, 15, 15, 14, 14, 14, 13, 11, 12, 16)
> n < -20
> a <- 1
> b <- 1
> nsamp <- 50000
> theta_chain <- rep(NA, nsamp)</pre>
> theta_chain[1] <- rbeta(1, a, b)</pre>
> i <- 1
> step <- 0.08
> while (i < nsamp) {</pre>
+ proposal theta <- rnorm(1, theta chain[i], step)</pre>
  if (proposal theta > 0 & proposal theta < 1) {</pre>
     post new <- sum(dbinom(y, n, proposal theta, log = TRUE)) +</pre>
       dbeta(proposal theta, a, b, log = TRUE)
     post prev <- sum(dbinom(y, n, theta chain[i], log = TRUE)) +</pre>
       dbeta(theta chain[i], a, b, log = TRUE)
     Hastings_ratio <- exp(post_new + dnorm(theta_chain[i], proposal_theta,</pre>
step, log
+
                                               = TRUE) -
                               (post_prev + dnorm(proposal_theta, theta_chain[i],
step, log =
                                                      TRUE)))
```



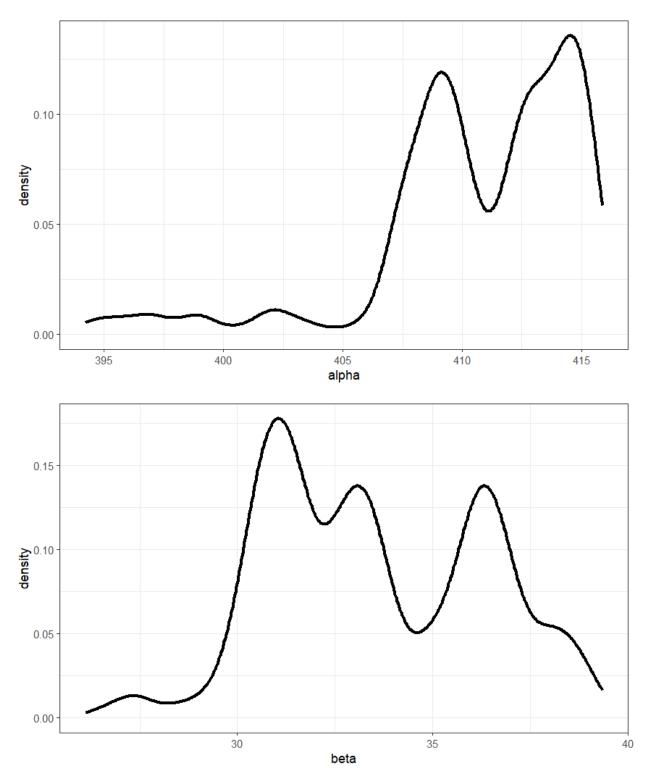
6. From above graphs, the posteriors from importance sampling and markov chain monte carlo have fatter tails as compared to analytical posterior

#### Part 2.5:

```
2.5.1:
> library(truncnorm)
> library(ggplot2)
> # Read the data from the corrected URL
"https://raw.githubusercontent.com/yadavhimanshu059/CGS698C/main/notes/Data/wor
d-recognition-times.csv"
> dat <- read.table(url, sep = ",", header = TRUE)[, -1]</pre>
> # Parameters
> nsamp <- 4000
> sigma <- 30
> step <- 0.2
> # Initialize chains
> alpha chain <- numeric(nsamp)</pre>
> beta chain <- numeric(nsamp)</pre>
> type chain <- integer(nsamp)</pre>
> # Initial values
> alpha chain[1] <- rnorm(1, 400, 50)</pre>
> beta chain[1] <- rtruncnorm(1, mean = 0, sd = 50, a = 0)
> type chain[1] <- sample(0:1, 1)</pre>
> # Metropolis-Hastings Algorithm
> for (i in 2:nsamp) {
+ # Sample from proposal distributions
+ proposal alpha <- rnorm(1, alpha chain[i - 1], step)
+ proposal beta <- rtruncnorm(1, mean = beta chain[i - 1], sd = step, a = 0)
+ proposal type <- sample(0:1, 1)
+ # Calculate mu
+ mu <- proposal alpha + proposal type * proposal beta
+ # Calculate log-posterior for the new and previous states
+ log posterior new <- sum(dnorm(dat$RT, mu, sigma, log = TRUE)) +
     dnorm(proposal alpha, 400, 50, log = TRUE) +
     proposal type * log(dtruncnorm(proposal beta, mean = 0, sd = 50, a = 0))
+ log posterior prev <- sum(dnorm(dat$RT, alpha chain[i - 1] + type chain[i -
1] * beta chain[i - 1], sigma, log = TRUE)) +
     dnorm(alpha_chain[i - 1], 400, 50, log = TRUE) +
     type chain[i - 1] * log(dtruncnorm(beta chain[i - 1], mean = 0, sd = 50, a
= 0))
+ # Calculate Hastings ratio and acceptance probability
```

+ hastings ratio <- exp(log posterior new - log posterior prev)

```
+ acceptance prob <- min(1, hastings ratio)
+ # Accept or reject the proposal
+ if (runif(1) < acceptance prob) {
   alpha chain[i] <- proposal alpha</pre>
   beta chain[i] <- proposal beta</pre>
    type chain[i] <- proposal type</pre>
+ } else {
    alpha chain[i] <- alpha chain[i - 1]</pre>
   beta chain[i] <- beta chain[i - 1]</pre>
    type chain[i] <- type chain[i - 1]</pre>
+ }
+ }
> # Quantiles for alpha and beta
> print(quantile(alpha chain, probs = c(0.025, 0.975)))
    2.5%
           97.5%
397.0015 415.5064
> print(quantile(beta chain, probs = c(0.025, 0.975)))
    2.5%
           97.5%
28.67353 38.62055
> # Histogram for beta chain
> hist(beta chain, main = "Histogram of Beta Chain", xlab = "Beta")
> # Posterior distributions
> posteriors <- data.frame(alpha = alpha chain, beta = beta chain)
> # Density plot for alpha
> ggplot(posteriors, aes(x = alpha)) +
+ geom_density(size = 1.2) +
+ theme bw() + xlab("alpha") +
+ theme(legend.title = element blank(), legend.position = "top")
Warning message:
Using `size` aesthetic for lines was deprecated in ggplot2 3.4.0.
i Please use `linewidth` instead.
This warning is displayed once every 8 hours.
Call `lifecycle::last lifecycle warnings()` to see where this warning was
generated.
> # Density plot for beta
> ggplot(posteriors, aes(x = beta)) +
+ geom density(size = 1.2) +
+ theme bw() + xlab("beta") +
+ theme(legend.title = element blank(), legend.position = "top")
```



2.5.2: For alpha,Quantile range is (395.2903 ,405.3474) And for beta,Quantile range is (31.53942,45.00786)

#### Part 3: Hamiltonian Monte Carlo sampler

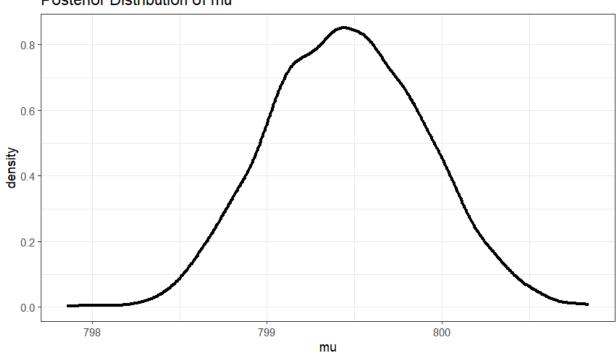
#### For below graphs:

```
> true mu <- 800
> true var <- 100 #sigma^2</pre>
> y <- rnorm(500, mean=true mu, sd=sqrt(true var))</pre>
> hist(v)
> #Gradient functions
> gradient <- function(mu, sigma, y, n, m, s, a, b) {</pre>
+ grad mu <- (((n*mu)-sum(y))/(sigma^2))+((mu-m)/(s^2))
+ qrad sigma <- (n/sigma) - (sum((y-mu)^2)/(sigma^3)) + ((sigma-a)/(b^2))
+ return(c(grad mu, grad sigma))
+ }
> #Potential energy function
> V <- function(mu, sigma, y, n, m, s, a, b) {</pre>
+ nlpd <-
-(sum(dnorm(y,mu,sigma,log=T))+dnorm(mu,m,s,log=T)+dnorm(sigma,a,b,log=T))
+ nlpd
+ }
> HMC <- function(y,n,m,s,a,b,step,L,initial q,nsamp,nburn){
+ mu chain <- rep(NA, nsamp)</pre>
+ sigma chain <- rep(NA, nsamp)
+ reject <- 0
+ #Initialization of Markov chain
+ mu chain[1] <- initial q[1]
+ sigma chain[1] <- initial q[2]
+ #Evolution of Markov chain
+ i <- 1
+ while(i < nsamp){</pre>
    q <- c(mu chain[i], sigma chain[i]) # Current position of the particle</pre>
     p <- rnorm(length(q),0,1) # Generate random momentum at the current
position
    current q <- q
     current p <- p
     current V = V(current q[1], current q[2], y, n, m, s, a, b) # Current potential
energy
    current T = sum(current p^2)/2 \# Current kinetic energy
    # Take L leapfrog steps
    for(1 in 1:L){
      # Change in momentum in 'step/2' time
       p \leftarrow p - ((step/2) * gradient(q[1], q[2], y, n, m, s, a, b))
      # Change in position in 'step' time
       q \leftarrow q + step*p
      # Change in momentum in 'step/2' time
       p \leftarrow p-((step/2)*gradient(q[1],q[2],y,n,m,s,a,b))
     proposed q <- q
```

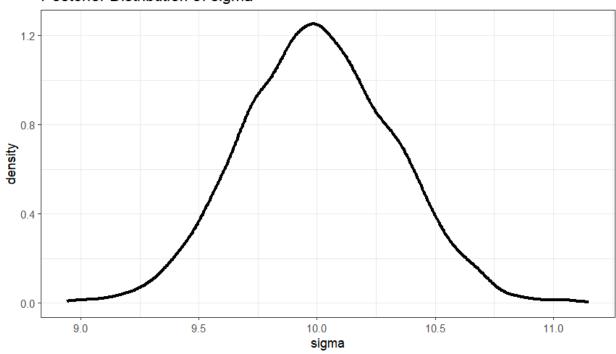
```
proposed p <- p
     proposed V = V(proposed q[1], proposed q[2], y, n, m, s, a, b) # Proposed
potential energy
     proposed T = sum(proposed p^2)/2 \# Proposed kinetic energy
     accept.prob <- min(1,exp(current V+current T-proposed V-proposed T))</pre>
     # Accept/reject the proposed position q
    if(accept.prob>runif(1,0,1)){
       mu chain[i+1] <- proposed q[1]</pre>
       sigma chain[i+1] <- proposed q[2]</pre>
      i <- i+1
    }else{
       reject <- reject+1</pre>
    }
+ posteriors <- data.frame(mu chain, sigma chain) [-(1:nburn),]
+ posteriors$sample id <- 1:nrow(posteriors)</pre>
+ posteriors
+ }
> df.posterior <- HMC(y=y,n=length(y), # data</pre>
                     m=1000, s=20, a=10, b=2, # priors
                      step=0.02, # step-size
                     L=12, # no. of leapfrog steps
                     initial q=c(1000,11), # Chain initialization
                     nsamp=6000, # total number of samples
                     nburn=2000) # number of burn-in samples
> ggplot(df.posterior, aes(x = mu chain)) +
+ geom_density(size = 1.2) +
+ theme bw() + xlab("mu") +
+ theme(legend.title = element blank(),
         legend.position = "top") +
+ ggtitle("Posterior Distribution of mu ")
> ggplot(df.posterior, aes(x = sigma chain)) +
+ geom density(size = 1.2) +
+ theme bw() + xlab("sigma") +
+ theme(legend.title = element blank(),
         legend.position = "top") +
+ ggtitle("Posterior Distribution of sigma")
```

## 3.1:

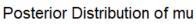


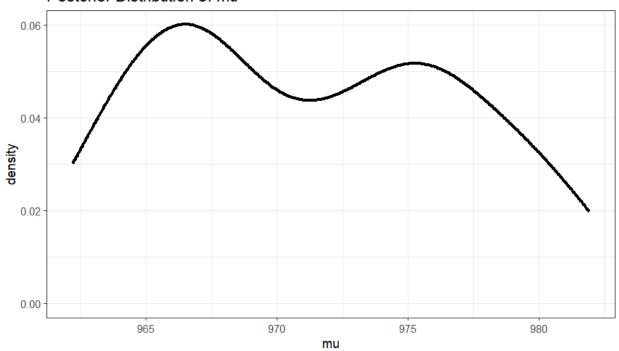


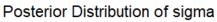
# Posterior Distribution of sigma

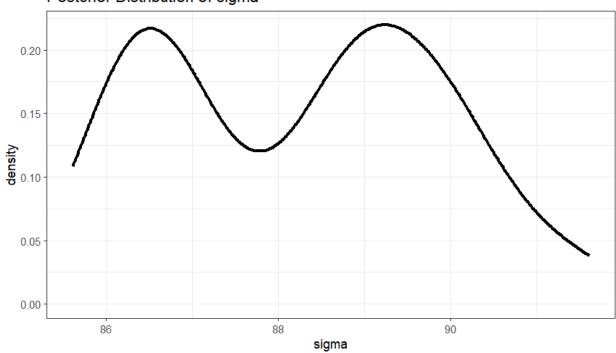


3.2: nsamp=100,nburn=100/3 = 33

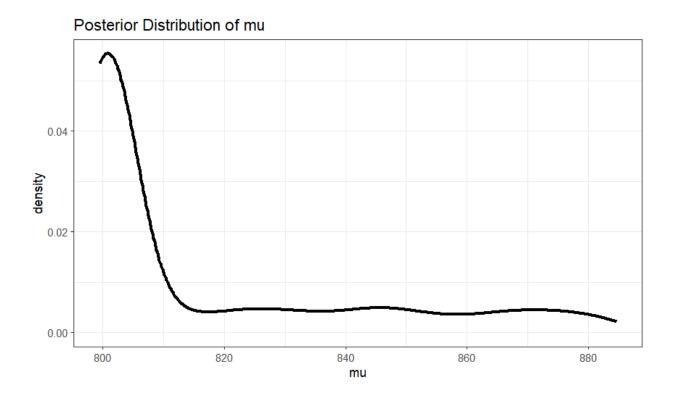


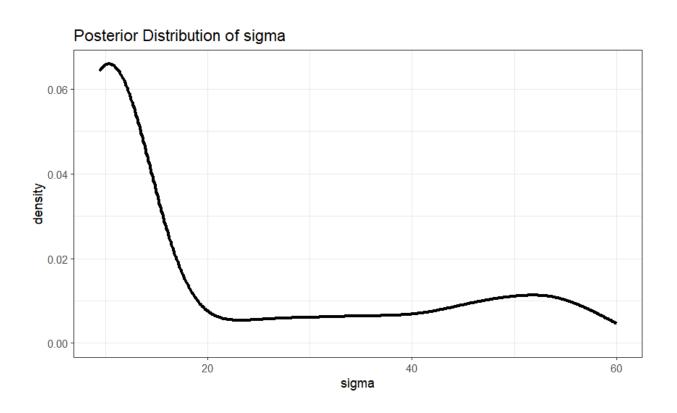






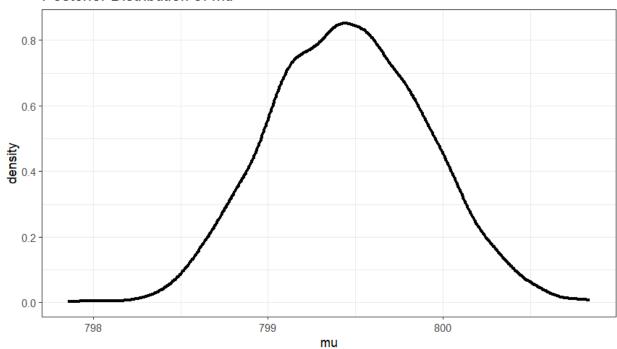
# nsamp=1000,nburn=1000/3 = 333



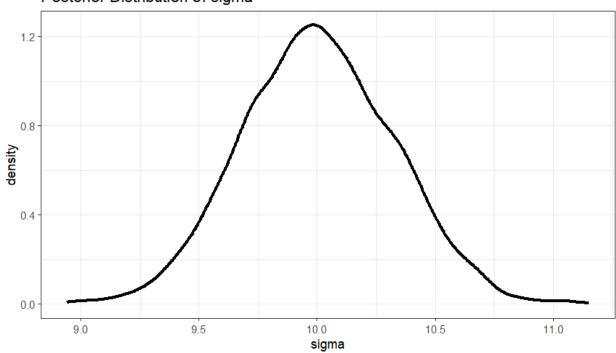


# nsamp=6000,nburn=6000/3=2000

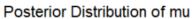


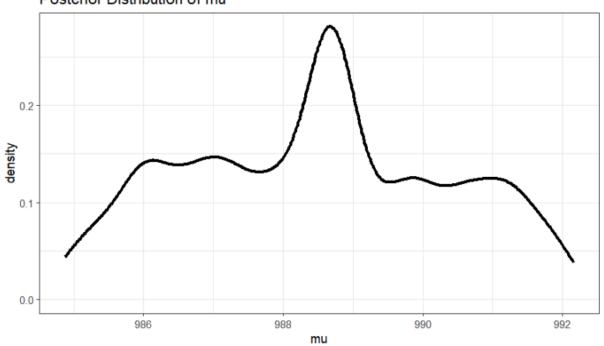


# Posterior Distribution of sigma

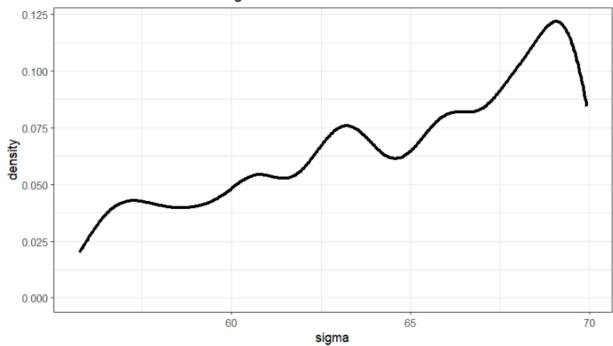


## step-size=0.01



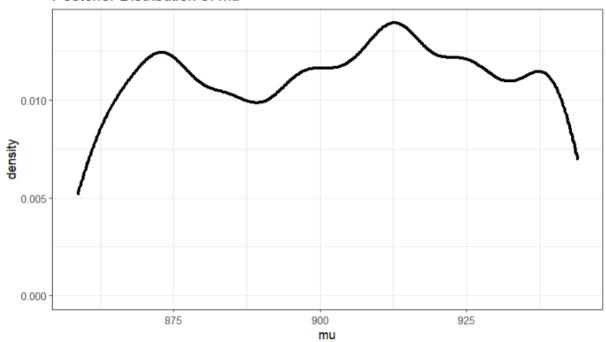


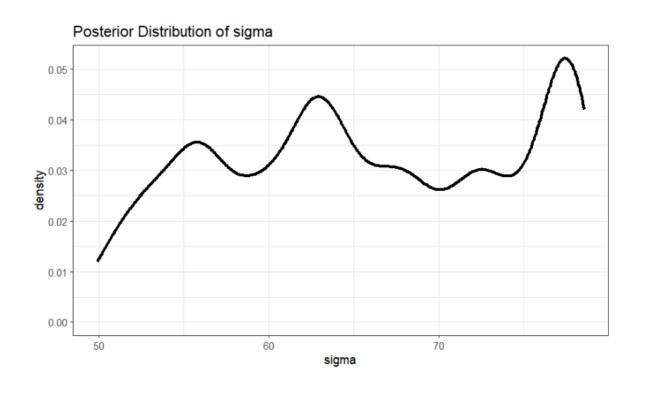




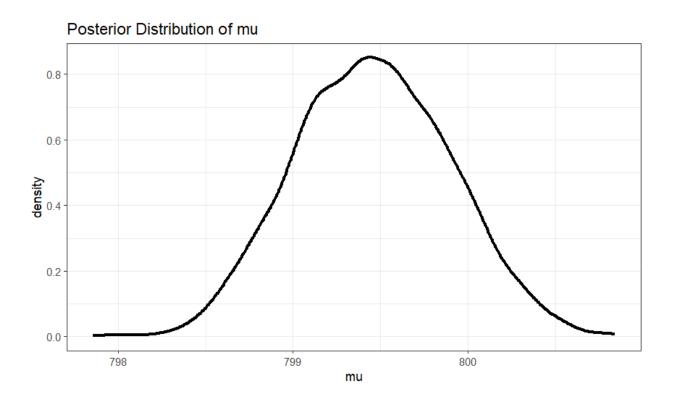
## step-size=0.005

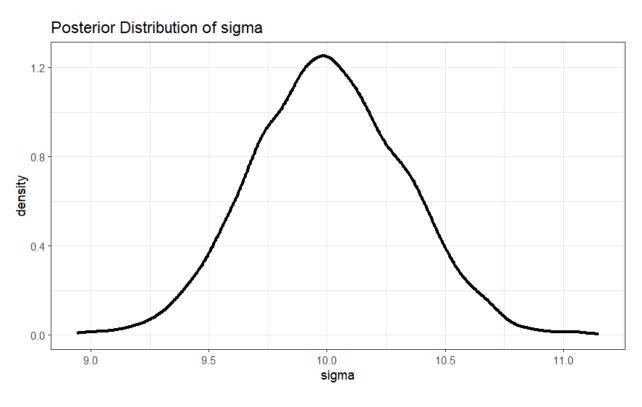






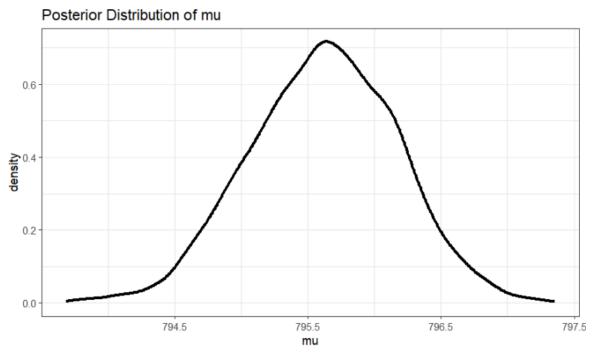
# step-size=0.02

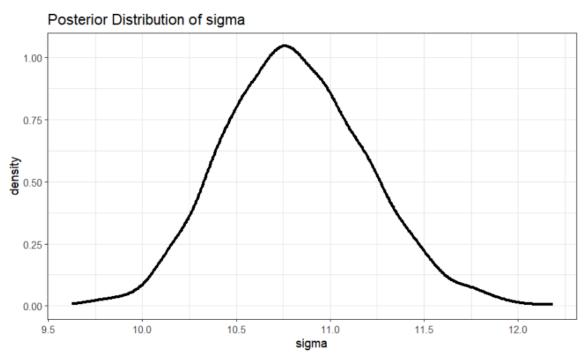




2.4: We can see that the above graphs the chain seems to move to slow and gets stuck around the mean.

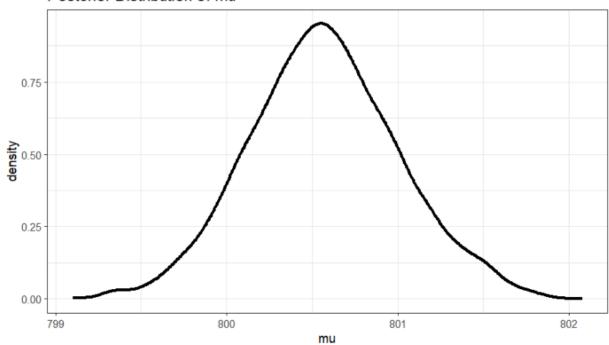
2.5:  $\mu \sim \text{Normal}(m = 400, s = 5)$ 



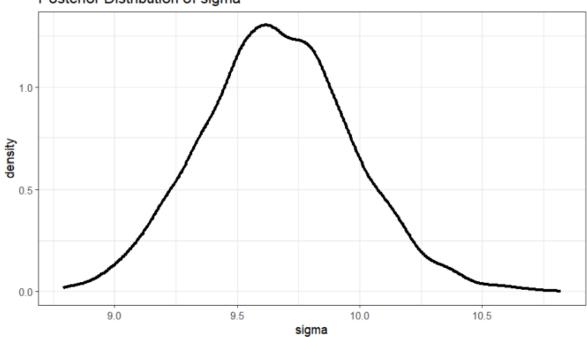


# $\mu \sim Normal(m = 400,s = 20)$

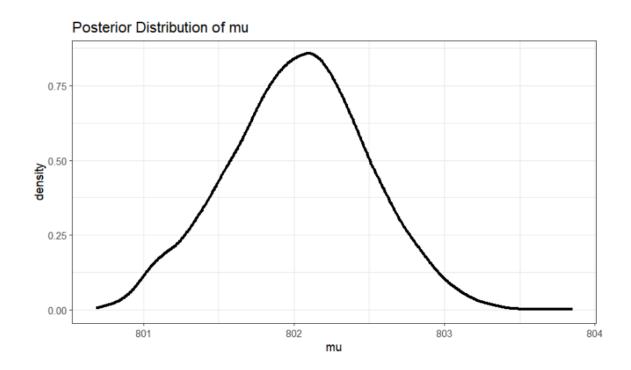
#### Posterior Distribution of mu

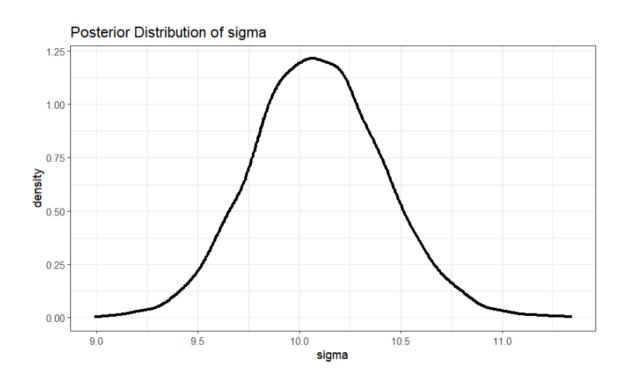




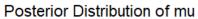


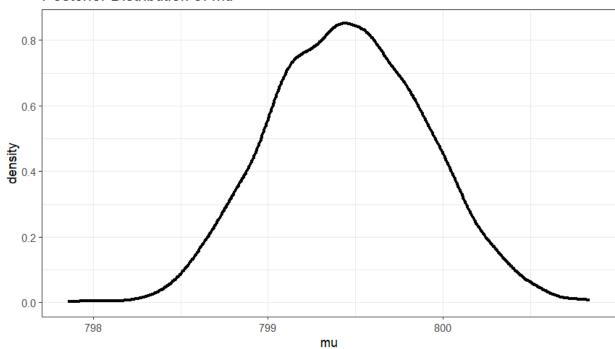
## $\mu \sim Normal(m = 1000,s = 5)$



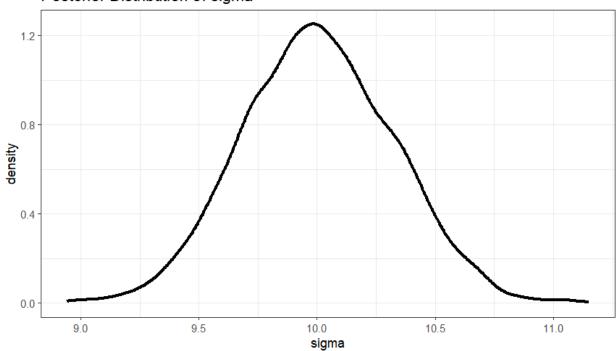


# $\mu \sim Normal(m = 1000,s = 20)$









# $\mu \sim Normal(m = 1000,s = 100)$

