

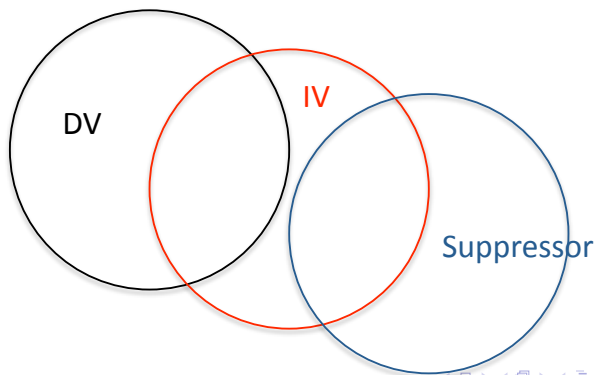
# Suppressor Effects in Linear Models

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## Thinking Graphically About Suppressors

- A suppressor effect occurs when a variable has a non-zero  $\beta$  weight but a zero structure coefficient.
- The inclusion of a suppressor in a regression equation removes the unwanted variance from the predictor variable, thus enhancing the relationship between the other independent variable and the dependent variable.



# Illustrating Suppressors in R

```
> library(MASS)
> correlation <- matrix(c(1, 0.5, 0, 0.5, 1, 0.5,
+   0, 0.5, 1), ncol = 3, nrow = 3, dimnames = list(c("dv",
+   "iv1", "iv2"), c("dv", "iv1", "iv2")))
> correlation

      dv iv1 iv2
dv  1.0 0.5 0.0
iv1 0.5 1.0 0.5
iv2 0.0 0.5 1.0

> set.seed(12346)
> suppressor.set <- data.frame(mvrnorm(n = 1000,
+   rep(10, 3), correlation))
```

## Illustration cont.

```
> m1 <- lm(dv ~ iv1 + iv2, suppressor.set)
> summary(m1)
```

Call:

```
lm(formula = dv ~ iv1 + iv2, data = suppressor.set)
```

Residuals:

	Min	1Q	Median	3Q	Max
	-2.75190	-0.54804	0.03108	0.53789	2.93339

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	6.00115	0.29836	20.114	<2e-16
iv1	0.67842	0.03044	22.285	<2e-16
iv2	-0.28112	0.03015	-9.323	<2e-16

Residual standard error: 0.8217 on 997 degrees of freedom

Multiple R-squared: 0.3356, Adjusted R-squared: 0.3343

F-statistic: 251.8 on 2 and 997 DF, p-value: < 2.2e-16

# Illustration cont.

```
> library(yhat)
> regr(m1)$Beta_Weights
```

```
      dv
iv1  0.673360
iv2 -0.281694
```

```
> regr(m1)$Structure_Coefficients
```

```
      iv1      iv2
0.9096263 0.1178144
```

```
> regr(m1)$Commonality_Data$CC
```

	Coefficient	% Total
Unique to iv1	0.3310	98.61
Unique to iv2	0.0579	17.26
Common to iv1, and iv2	-0.0533	-15.87
Total	0.3356	100.00

## Types of Effect Size Corrections

$$\text{Wherry-1:} = 1 - \frac{N-1}{N-p-1}(1-R^2)$$

$$\text{Claudy3:} = 1 - \frac{(N-4)(1-R^2)}{N-p-1} \left[ 1 + \frac{2(1-R^2)}{N-p+1} \right]$$

$$\text{Smith-1:} = 1 - \frac{N}{N-p}(1-R^2)$$

$$\text{Wherry-2:} = 1 - \frac{N-1}{N-p}(1-R^2)$$

$$\text{Olkin\&Pratt:} = 1 - \frac{(N-3)(1-R^2)}{N-p-1} \left[ 1 + \frac{2(1-R^2)}{N-p+1} \right]$$

$$\text{Pratt:} = 1 - \frac{(N-3)(1-R^2)}{N-p-1} \left[ 1 + \frac{2(1-R^2)}{N-p-2.3} \right]$$

## Yin and Fan, 2001

- Yin and Fan (2002) showed how the multiple  $R^2$  is really an upward bound for the actual  $R^2$  in the population.
- Perhaps most important, they noted that the Wherry formula (basis for the default adjusted  $R^2$  in commonly used statistical software) has been found to be an ineffective analytical formula when sample sizes are small.
- What `regr` does is compute the multiple  $R^2$  correction that has the least amount of bias given the simulations done by Yin and Fan (2002).

## Homework

- Find a journal article in a journal in your field that has used multiple regression.
- Make sure that the one you identify has used both multiple regression *and* has the correlation matrix attached in a table.
- Reproduce their results by simulating their dataset.
- Try other competing models against their full model and see if you can find a different, better (or at least as good as) model.
- Use the `anova` function to compare the models.

```
> m0 <- lm(dv ~ iv1, suppressor.set)
> anova(m0, m1)
```

### Analysis of Variance Table

Model 1: dv ~ iv1

Model 2: dv ~ iv1 + iv2

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	998	731.86				
2	997	673.17	1	58.685	86.915	< 2.2e-16