

*“Begin at the beginning,” the King said gravely,
“and go on till you come to the end: then stop.”*

— Lewis Carroll, *Alice in Wonderland*

3

Using rare variants to detect haplotype sharing and identity by descent

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3.1 Introduction

Identity by descent (IBD) is a fundamental concept in genetics that describes the genealogical relation between individuals (Malécot, 1948). Two chromosomes are said to be identical by descent, or rather to share a haplotype by descent, if they have inherited the same genetic material from a common ancestor (*e.g.*, see Browning and Browning, 2012; Thompson, 2013). Over generations, the length of an ancestral haplotype is broken down through meiotic recombination, as the genetic material is blended with haplotypes that derive from different ancestral lineages. Consequently, any random sample of two

different chromosomes carries a unique pattern of relatedness, with different ancestries at different loci, arising as the result of historical recombination events. The underlying structure of pairwise relatedness can be thought of as a mosaic of segments at which two chromosomes share a haplotype by descent, but where each of these IBD segments traces back to a different most recent common ancestor (MRCA).

In general, knowledge about relatedness, haplotype sharing by descent, or the recombination history of a sample is of importance in a variety of statistical operations that are used in both population and medical genetics research (Milligan, 2003; Albrechtsen *et al.*, 2009; Gusev *et al.*, 2009); for example, to provide insights into the demographic history of a population (Harris and Nielsen, 2013), to inform methods for genotype phasing and imputation (Kong *et al.*, 2008), to map disease loci using linkage analysis (Purcell *et al.*, 2007; Albrechtsen *et al.*, 2009), as well as to reveal patterns of population stratification and to identify unreported relatedness among individuals in disease association analysis (Freedman *et al.*, 2004; Price *et al.*, 2006; Choi *et al.*, 2009; Mathieson and McVean, 2012).

The entire IBD structure of a sample can be represented by the ancestral recombination graph (ARG) (Griffiths, 1991; Griffiths and Marjoram, 1996, 1997b), which is straightforward to generate in coalescent simulations, but inference from observed data is limited (Rasmussen *et al.*, 2014). This is because even complete data is unlikely to provide sufficient information to explicitly infer the ARG, in addition to the problem that inference becomes computationally expensive for larger sample sizes. Most methods for IBD discovery operate on summary statistics to make inference computationally tractable.

In practice, IBD discovery is largely dependent on the length of a shared haplotype and the genetic similarity between compared sequences. Co-inherited haplotypes that are separated by only a few meioses are expected to cover relatively long tracts, because recombination had less time to break down the length of the region shared between the two chromosomes (Thompson, 2008, 2013). Likewise, as mutations are accumulated along different genealogical lineages, the similarity between shared segments is expected to

decrease over time. Thus, for most purposes, the detection of *recent* IBD is of primary interest (Browning and Browning, 2010).

Numerous approaches for the detection of IBD segments have been proposed, most of which attempt to infer IBD based on measures of genetic similarity or through use of statistical models to determine salient patterns of linkage disequilibrium (LD). Commonly employed tools are PLINK (Purcell *et al.*, 2007), GERMLINE (Gusev *et al.*, 2009), fastIBD (Browning and Browning, 2011a), and Refined IBD (Browning and Browning, 2013), to name a few. The methodological diversity of existing approaches emphasises the central role of IBD in genetics, but also indicates that there is a need for an accurate as well as efficient method to detect IBD in larger samples of purportedly unrelated individuals.

Due to the growing magnitude of available genomic datasets, IBD discovery is becoming more computationally expensive. Note that alternate approaches exist, for example methods to perform long range phasing (LRP) implicitly harness long IBD regions among related individuals (Kong *et al.*, 2008; Palin *et al.*, 2011; Loh *et al.*, 2016a), which employ computationally efficient methods to match relatively long (*e.g.* >10 cM) haplotypes even in very large datasets. But in a general context, as IBD describes a pairwise relationship between two haplotypes, a search algorithm may visit each of the possible pairs of chromosomes in a sample to determine IBD status from patterns of shared genetic variation observed along the full length of the chromosome. For instance, in a sample of n chromosomes, there are $\binom{n}{2} = n(n-1)/2$ possible pairs that need to be scanned to resolve IBD status if done in an exhaustive manner. To reduce this search space, it would be convenient if a pairwise approach could be targeted to regions and individuals for whom it is more likely to find recent haplotype sharing by descent.

In this chapter, I present a non-probabilistic method to detect IBD segments in pairs of diploid individuals, which utilises rare variants as indicators of recent relatedness. The computational burden of IBD detection is thereby reduced due to the relative low number of individuals that share a given rare or low-frequency allele. In each pair, the regions to each side of a focal allele are scanned, so as to infer the “breakpoints” of historical recombination events that delimit the underlying IBD segment. The inference

of recombination is based on the *four-gamete test* by Hudson and Kaplan (1985), for which haplotype information is required, but which is extended, following Mathieson and McVean (2014), such that recombination breakpoints can be inferred in genotype data.

In the following section, I highlight the genealogical properties of rare variants which make them useful for the inference of recent and relatively long haplotypes by descent. I then describe the method by which IBD segments are detected, conditional on variation observed at a focal rare variant. In addition, I present a simple approach to infer the shared haplotype sequence from genotype data. For the evaluation of the methodology presented, I generated a large dataset using coalescent simulations, so as to measure the accuracy of the IBD detection method in comparison to the true IBD structure (determined from simulation records). These results are also compared to IBD detected using an alternate method. Lastly, I apply the method presented in this chapter to data from the 1000 Genomes Project (1000G).

3.2 Rare variants as indicators of haplotype sharing by descent

One of the properties of rare variants is their presumed young age, as a low frequency is indicative of a recent origin through mutation; *i.e.* the frequency of an allele is assumed to be a proxy to its age (Kimura and Ota, 1973; Griffiths and Tavaré, 1998). Individuals that share a rare allele are therefore likely to have a relatively long chromosomal segment co-inherited from a common ancestor. For example, genetic markers tend to be in high LD with alleles at lower frequencies, because the alleles near a rare variant site are likely to segregate together on the same haplotype (Kruglyak, 1999; Slatkin, 2008b).

To explain the relation between IBD length and age, consider a focal site at which two haplotypes are shared by descent. The length of the IBD segment is defined by the nearest ancestral recombination events that have occurred to either side of the focal position; *i.e.* haplotype sharing is broken down by recombination on both sides independently. The expected length of the IBD segment is determined by the number of meioses that separate two haplotypes in relation to the MRCA who lived t generations in the past;

hence, the pair is separated by $2t$ meioses. In each meiosis, recombination is modelled as a Poisson process with rate of 1 per unit of genetic distance (*Morgan*). It follows that the recombination process over $2t$ meioses is Poisson distributed with rate equal to $2t$. The expected length can be expressed as the sum of two independent random variables that are exponentially distributed, and which describes the distance to either side of the focal position (see Wakeley and Wilton, 2016); *i.e.* the length, L , is gamma-distributed with shape 2 and rate $2t$, namely $L \propto \Gamma(2, 2t)$.

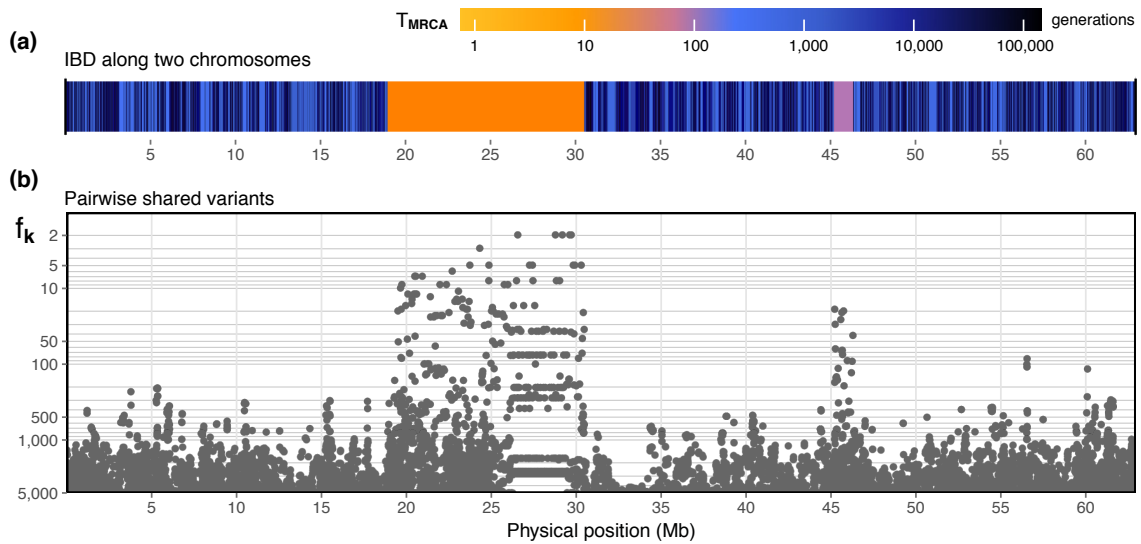


Figure 3.1: IBD structure and pairwise variant sharing. A dataset of $N = 5,000$ haplotypes was simulated under the coalescent using *msprime* (Kelleher *et al.*, 2016). IBD status was determined from simulated genealogies for a pair of chromosomes selected at random from the set of chromosomes that shared a rare allele (frequency $\leq 0.5\%$). Panel (a) shows the “mosaic” of IBD segments along the full length of the simulated region for the two selected chromosomes. The length of a given IBD segment is defined by the chromosomal interval over which the MRCA of the selected pair does not change. The colour of each segment indicates the time to the most recent common ancestor (T_{MRCA}) for the selected pair. Panel (b) shows the physical position of f_k variants shared by the two chromosomes, ranging from very low allele frequency at the top (f_2) to very high frequency at the bottom (*e.g.* $f_{>500}$). Note that the simulation was carried out under variable recombination rates using the genetic map for human chromosome 20 from the International HapMap Project (HapMap) Phase II Build 37. The pattern of extended shared variation seen at positions around 25–30 Megabase (Mb) arises from a low recombination rate at the region of the centromere.

Given this exponential “decay” of IBD length over time, rare or low-frequency variants are useful for identifying genomic regions in which individuals are likely to share recent

and relatively long IBD tracts. For example, Mathieson and McVean (2014) selected doubletons (alleles that are present only twice in a sample), which they refer to as f_2 variants, to identify the shared haplotype in the two individuals sharing the allele. To borrow from this notation, henceforth, f_k is used to denote a variant at which k allele copies are found in a sample.

To emphasise the utility of rare variants, see the example shown in Figure 3.1. Using coalescent simulations, a sample of $N = 5,000$ chromosomes was generated.* A rare variant was randomly selected (frequency $\leq 0.5\%$), as well as two of the chromosomes which share the focal allele. The underlying IBD structure for the given pair of chromosomes was determined from simulation records and shown in Figure 3.1a. IBD segments are distinguished by the time to the most recent common ancestor (T_{MRCA}) at each position along the sequence. To illustrate pairwise allele sharing, the frequency of each allele shared by the two haplotypes is shown by chromosomal position in alignment with the IBD structure above; see Figure 3.1b. As suggested in the figure, the majority of low-frequency variants align with IBD segments that are more recent.

The majority of variants observed in the human genome are low in frequency or rare. For example, there are 84.7 million single-nucleotide polymorphisms (SNP) in the final release dataset of the 1000 Genomes Project (1000G) Phase III ($N = 2,504$), of which 71.9 % are below 1% allele frequency and 64.2 % are below 0.5% (after removing singletons and monomorphic sites), suggesting that there are ample opportunities to find rare allele sharing. This is illustrated in Figure 3.2 (next page), which indicates the number of alleles shared between each pair in the dataset (chromosomes 1–22), at allele frequency $\leq 0.5\%$. Notably, the sharing pattern highlights population structure, as the number of shared alleles is generally larger within a sub-population.

3.3 IBD detection around rare variants

In the following sections, I describe the methodology by which IBD segments are detected around rare variant sites. I then describe the implementation of each of the two tests for

* See Section 3.5.1 (page 102) for a description of how data were simulated.

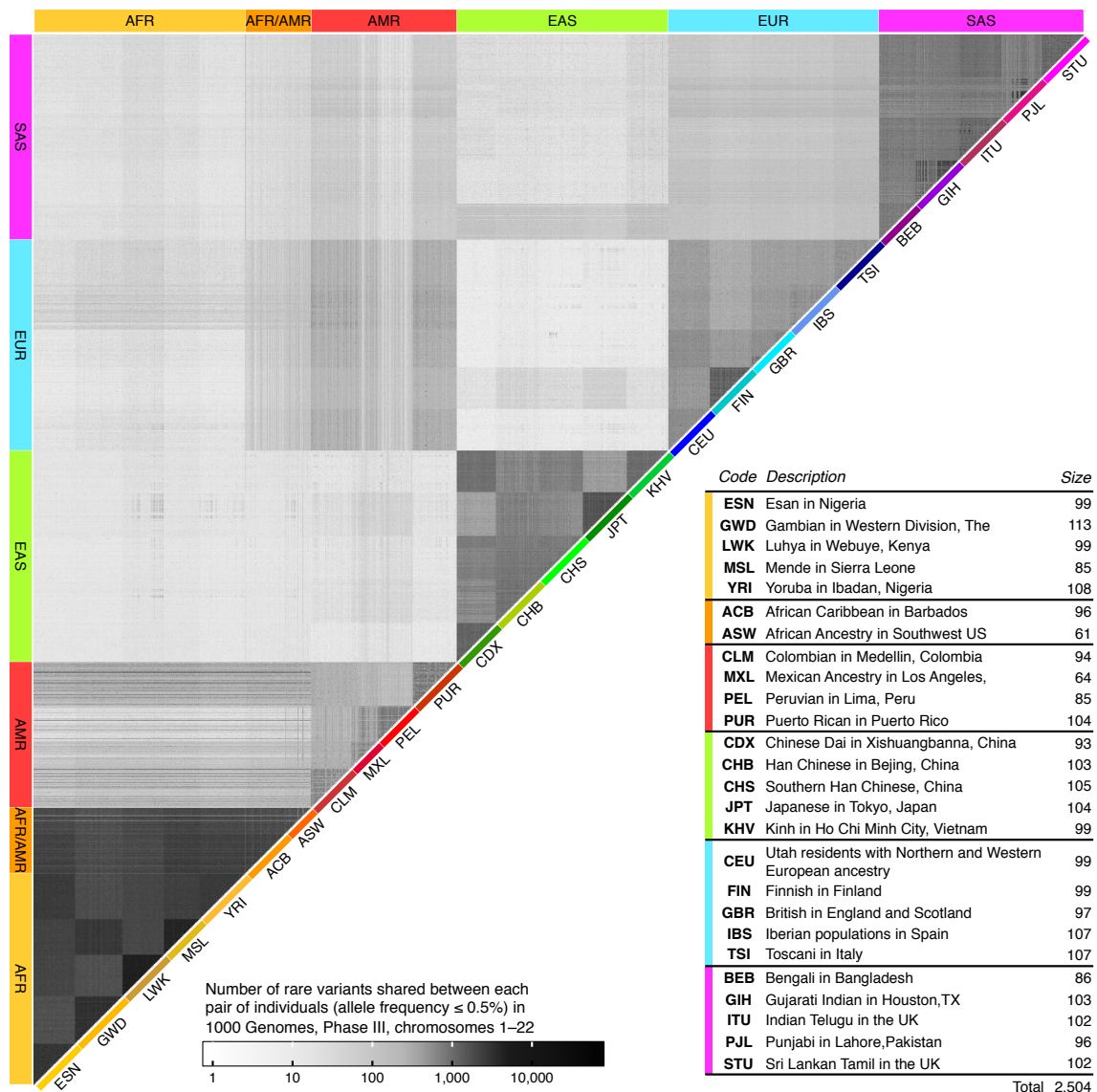


Figure 3.2: Rare variant sharing in the 1000 Genomes dataset. The plot shows the upper triangle of a pairwise sharing matrix in which the number of variants shared in each pair of individuals is indicated by tones of grey (log-scaled), ranging from *light* (low number) to *dark* (high number); see legend. Pairwise rare variant sharing was determined for all shared alleles observed at frequency $\leq 0.5\%$, across chromosomes 1–22, and in each pair of the 2,504 individuals present in the final release dataset of the 1000 Genomes Project Phase III. The dataset comprises sample data from six continental populations (or *super-populations*) which are further subdivided in 26 populations of different ethnic background. Each group is abbreviated using a three-letter code. The six continental populations are defined as follows; African (AFR), African-American (AFR/AMR), American (AMR), East Asian (EAS), European (EUR), and South Asian (SAS). The table in the lower right corner shows the code and description of each population sample, as well as the number of individuals in each group.

the detection of IBD segments in large sample data. Lastly, I conclude this section by highlighting certain caveats of the implemented method before its evaluation using simulated data.

3.3.1 Inference of historical recombination events

Two approaches for a non-probabilistic inference of recombination events are described below; these are the *four-gamete test* (Hudson and Kaplan, 1985), which requires haplotype information, and the criterion of *inconsistent homozygote genotypes* (see Mathieson and McVean, 2014), which requires genotype data; henceforth referred to as the *discordant genotype test*.

Four-gamete test (FGT). Given four haplotypes in two diploid individuals, a recombination event is inferred between two loci if all four possible gametes are observed. This holds true under the infinite sites model (Kimura, 1969), where mutation events may only occur once per site in the history of a sample, such that at most two allelic states can be observed at a given site. It follows that for a pair of sites there are four possible allelic state configurations; (0, 0), (0, 1), (1, 0), and (1, 1), where 0 and 1 denote the ancestral and derived type, respectively. If all four configurations are observed, genealogies at the two sites are incompatible and the observation can only be explained by a recombination event that occurred in the history of the sample. Because recurring mutations or back mutations are assumed to have zero probability, at least one recombination event must have occurred in the interval between the two sites. An example configuration is shown in Figure 3.3 (next page). Notably, private or *de novo* mutations appearing as singletons in the sample cannot lead to the observation of the four required configurations. Although the exact location of chromosomal crossover cannot be retrieved from the data, the FGT can be used to find the smallest interval in which recombination occurred. In the following, the term *breakpoint* is used for either of the two sites that together delimit the interval.

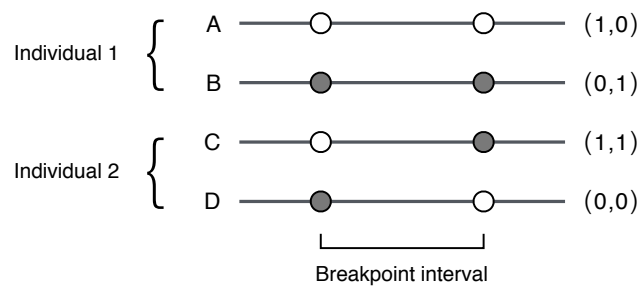


Figure 3.3: Breakpoint detection using the four-gamete test (FGT). The four haplotypes (gametes) in a pair of two diploid individuals are shown (*horizontal lines*). A breakpoint interval is detected if all four possible allelic state configurations are observed at two variant sites along the sequence. The interval delimits the region in which at least one recombination event must have occurred in the history of the sample (given the assumptions of the infinite sites model). The four allelic state configurations are shown on the *right*. The alleles are shown at the two breakpoint sites; indicated as ancestral (*hollow circle*) and derived state (*solid*). Note that the order of gametes is ignored.

Discordant genotype test (DGT). In absence of haplotype information, data are represented as genotypes, where genotypic states are encoded as 0, 1, and 2, for variants that are homozygous for the ancestral allele, heterozygous, and homozygous for the derived allele, respectively. Given the genotype sequences of two diploid individuals, recombination is inferred between two sites; one being heterozygous in both individuals (*i.e.* the genotypes 1 and 1) and another with opposite homozygous genotypes (0 and 2). In the latter case, it follows that the two individuals cannot share a haplotype at that locus.

The DGT is a special case of the FGT, as the same composition of alleles is implied. For example, if the allelic configurations (0, 1) and (0, 0) are seen in individual 1, and configurations (1, 0) and (1, 1) in individual 2, the corresponding genotypic configurations are (0, 1) and (2, 1), respectively, which satisfies the breakpoint condition in both the FGT and DGT. However, because genotype data result from haplotype occurrence in individuals, not all breakpoints detectable under the FGT can be found using the DGT. At sites where the FGT detects a breakpoint interval, the DGT cannot break if both sites are heterozygous in the same individual. As a consequence, it can be expected that the DGT is more restrictive than the FGT; *e.g.* if breakpoints are found, they are likely to sit farther apart. This is further exemplified in Figure 3.4 (next page), which highlights the difference to the FGT by comparison to the example shown in Figure 3.3 (this page).

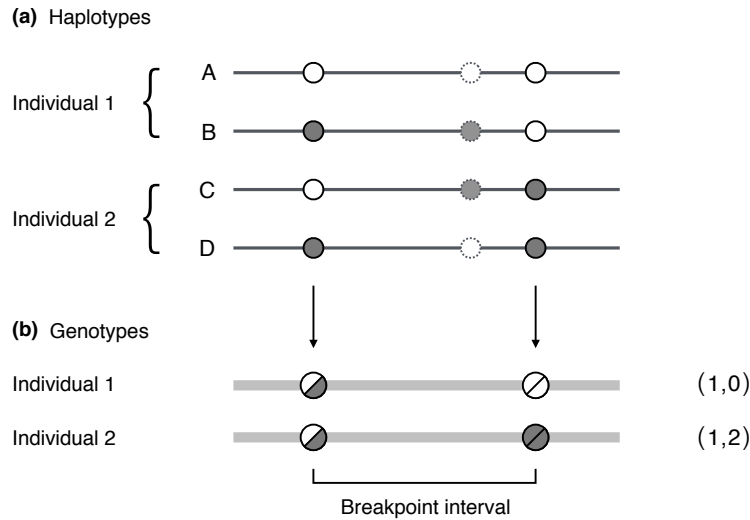


Figure 3.4: Breakpoint detection using the discordant genotype test (DGT). Unlike the FGT, which requires haplotype information, the DGT identifies a breakpoint interval using genotype data. For comparison, Panel (a) shows the four gametes of the two individuals involved; see Figure 3.3 (page 93). To highlight the difference to the FGT, an additional variant is shown in between both sites; alleles indicated by a *dotted* edge. This site would satisfy the breakpoint condition under the FGT, but is missed under the DGT. Panel (b) shows the two genotype sequences per individual (*thick* horizontal lines) from which a breakpoint interval is inferred using the DGT. The genotypic states of the breakpoint sites are given on the *right*. Genotypes can either be homozygous for the ancestral allele (*hollow* circle), heterozygous (*semi-solid*), or homozygous for the derived allele (*solid*).

3.3.2 Description of the algorithm

The FGT and DGT provide the means for non-probabilistic inference of recombination breakpoints from either haplotype or genotype data, respectively. This methodology is implemented such that the full length of an IBD segment can be found around a given target site in a pair of diploid individuals. The allele at a target site serves as an indicator for haplotype sharing by descent; hence, to detect recent IBD, rare variants are used as primary targets. The aim of this method is to infer breakpoint intervals independently on both sides of the target position along the sequence, so as to infer the two recombination events that delimit the underlying IBD segment. As such, the target variant is set as the *focal* breakpoint. The algorithm is described below; a more intuitive example is illustrated in Figure 3.5 (next page).

Let M be the number of variant sites observed in a sample of N diploid individuals. At the target site, b_i , where $i \in \{1, 2, \dots, M\}$, the subset of individuals sharing the derived

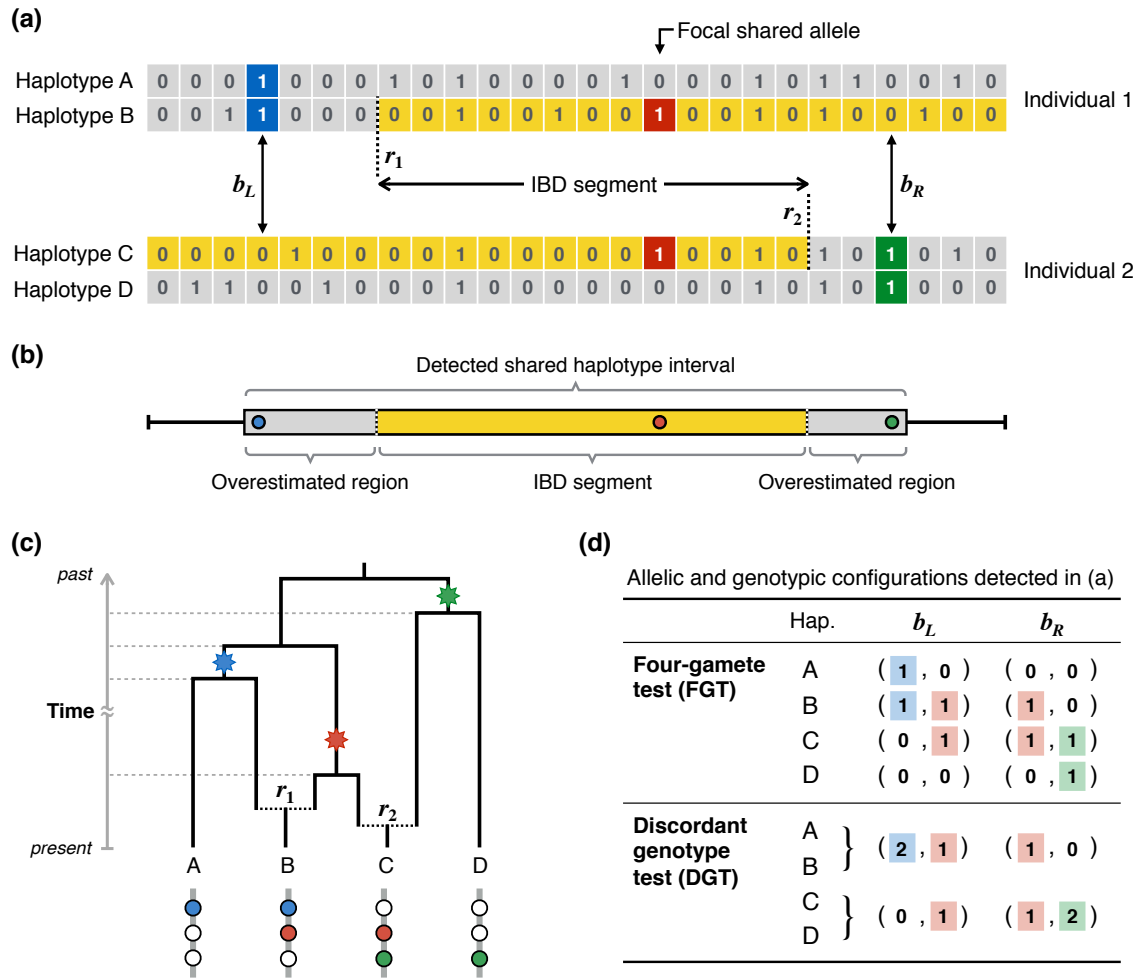


Figure 3.5: Illustration of shared haplotype detection in a pair of diploid individuals. Panel (a) shows two individuals composed of haplotypes A and B, and haplotypes C and D, respectively. Each haplotype is represented as a sequence of observed allelic states, where 0 and 1 denote the ancestral and derived allele, respectively. Breakpoints are detected by independently scanning to the left and right-hand side from the target position. The two individuals share a haplotype by descent (highlighted in yellow) which is tagged by the focal allele (red), for which the two individuals are heterozygous. Two sites (blue and green) mark the first sites at which a breakpoint condition is satisfied, such that b_L and b_R are detected. The IBD segment shared by both individuals is indicated by r_1 and r_2 (dashed lines). Panel (b) shows the detected breakpoint interval, delimited by b_L and b_R (inclusive). Note that detected breakpoints are only the first indication of recombination found distal to the focal site, but may not mark the points of the actual crossover events; thus, it is expected that the length of the detected segment is overestimated, dependent on available data. Panel (c) represents the history of the sample as an ancestral recombination graph (ARG). Mutation events are indicated on the tree (stars) and gave rise to the alleles highlighted in (a); blue, red, and green. The dotted grey lines indicate the time of coalescent events in the history of the sample; dotted black lines indicate recombination events. Panel (d) provides a table outlining the configurations of allelic and genotypic states at breakpoint sites as considered in the FGT and DGT, respectively. Notably, in the example shown, both the FGT and DGT detect breakpoints at indicated sites. But, for example, if individual 1 was composed of haplotypes A and C, and individual 2 of haplotypes B and D, the breakpoints would be detected as shown under the FGT, but not the DGT.

allele is identified and compared in a pairwise fashion. Importantly, the allele at this site is used as an identifier for haplotype sharing, on which inference is conditioned in either the FGT or DGT. Thus, individuals are only considered if they are heterozygous for the focal allele, as the breakpoint condition in either test cannot be satisfied otherwise. However, note that this restriction arises from the variant-centric focus on a given rare allele; *e.g.* the condition of the FGT could be satisfied for individuals homozygous for a given allele, but without that the allele is shared by the other individual (hence, defying the purpose of this implementation). In each pair, chromosomes are scanned to the left and right-hand side from the target site until the first site is found that, together with the allelic or genotypic states observed at b_i , satisfies the breakpoint condition, which is done independently on each side. Detected breakpoints are labelled as b_L and b_R on the left and right-hand side, respectively, such that the intervals $[b_L, b_i]$ and $[b_i, b_R]$ delimit the chromosomal regions in which recombination events occurred, respectively; where $L, R \in \{1, 2, \dots, M\}$. Hence, the underlying IBD segment is enclosed in $[b_L, b_R]$.

The allelic or genotypic states at b_L or b_R provide only the first indication of recombination found along the sequence on either side of the focal allele, but may not mark the points of the actual crossover events. The detected interval is therefore inclusive of the breakpoints such that the full length of the underlying IBD segment is covered. In cases where the end of a chromosome is reached without detecting any evidence of recombination, the terminal site is recorded to capture the length of the segment; this is hereafter referred to as a *boundary case*. Further, note that a recombination event can occur with chromosomes outside the sub-tree of the lineages deriving from the focal mutation, such that a breakpoint may be falsely inferred; however, such cases are assumed to be sufficiently rare and treated as being negligible.

3.3.3 Anticipated limitations

As noted by Hudson and Kaplan (1985), not all recombination events in the history of a sample are found by the FGT, and are therefore also missed by the DGT. In the implementation presented, a breakpoint is found by performing a scan along the sequence

away from a target position. Provided that the neighbouring haplotype regions derive from different ancestral lineages, in the general case, it is likely that a breakpoint will be found eventually (or the boundary of the chromosome is reached).

The main limitation to the accuracy of the detected breakpoints is the overestimation of the interval, in relation to the underlying true IBD length; as shown in Figure 3.5b. While the underlying IBD segment is enclosed in the interval, it can be expected that breakpoints are detected at sites some distance away from where recombination occurred, thus overestimating the true length of the underlying IBD tract. The extent of overestimation is dependent on the number and density of observed variant sites in the sample. This suggests that the method may become more accurate with larger samples, as the number of segregating sites observed in the sample is expected to increase with larger sample size. Likewise, because the rate of mutation is directly proportional to the expected number of segregating sites (Watterson, 1975), a higher mutation rate can generally be expected to decrease the overestimation of segment length.

Conversely, it is also possible that segment length is underestimated. Given the four chromosomes required, there are $\binom{4}{2} = 6$ possible pairs of chromosomes which may share extended regions by descent. For example, if two individuals share multiple IBD tracts at different pairs of their chromosomes, these tracts may overlap. In such cases, breakpoint detection cannot distinguish between overlapping segments. This is illustrated in Figure 3.6 (next page), which shows two examples generated using coalescent simulations.

In Example 1 (3.6a), a rare allele target site was randomly selected, as well as the two individuals sharing the this allele. The true IBD structure was determined from simulation records for each pair of the four chromosomes in the two individuals. Each pair is represented by a mosaic of IBD segments along the sequence, where each segment is distinguished by time to the most recent common ancestor (T_{MRCA}). Both the FGT and DGT were applied, but where all consecutive breakpoints after the first detection were also recorded along the sequence on both sides of the focal variant. The innermost interval delimits the detected shared haplotype segment around the target

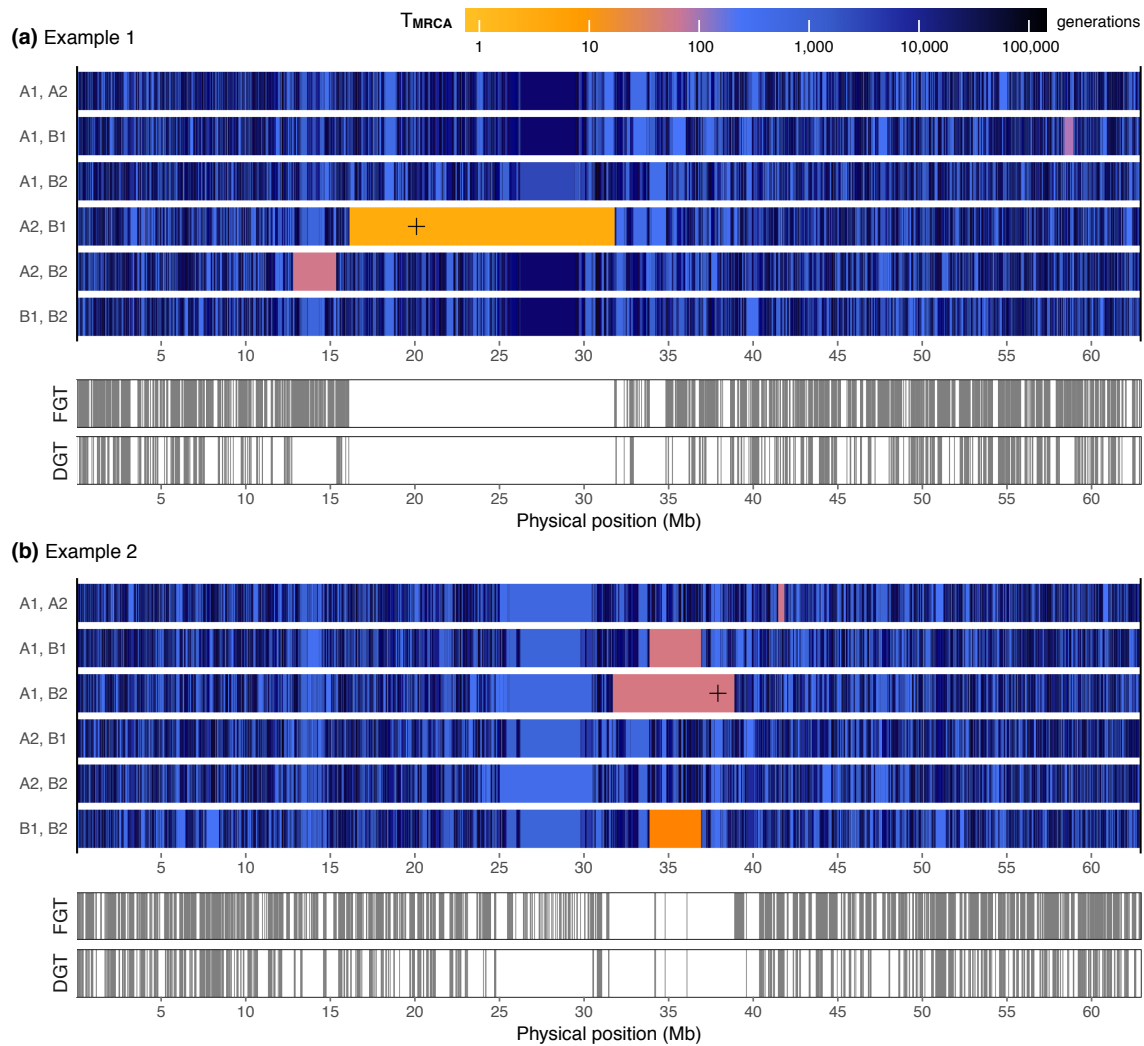


Figure 3.6: Examples of the underlying IBD structure in each pair of four chromosomes.

The true, underlying IBD structure is shown for each possible pair among four chromosomes in two diploid individuals; two examples are shown. Each chromosome is labelled by its occurrence in individuals A and B, where chromosomes 1 and 2 are distinguished. The “mosaic” of IBD segments per pair was determined from coalescent records produced in simulations using *msprime* (Kelleher *et al.*, 2016); see Section 3.5.1 (page 102). Each segment defines the region that was co-inherited from a most recent common ancestor (MRCA), and is colour-coded by the number of generations separating the two chromosomes from their shared MRCA in that region. The *cross* marks the position of the focal allele in the pair that shares it. Below, all breakpoints detected relative to the focal variant along the simulated region are indicated, using the FGT (*top*) and DGT (*bottom*). Panel (a) shows that the innermost breakpoint intervals (relative to the target position) detected in the FGT or DGT align closely with the true termini of the IBD segment. The extent of overestimation appears to be negligible in relation to the length of the detected segment. Panel (b) shows that the innermost intervals are underestimated, due to an overlap of recently co-inherited haplotypes on different chromosome pairs.

site. Example 1 illustrates the general case in which a rare allele identifies the underlying co-inherited haplotype segment, which may stand out as being much younger due to recent shared ancestry.

The same was done in Example 2 (3.6b), but here the target site and the pair of individuals was chosen because it was found that the length of the detected IBD segment was underestimated. As can be seen, this underestimation is due to an overlap of multiple pairwise shared IBD tracts in other chromosome pairs within the same pair of individuals. Such a result may be expected in cases of inbreeding, where the maternal and paternal chromosomes in an individual are more closely related to each other than to other chromosomes in the population. Note that in the simulations conducted, the generated haplotypes were randomly paired to form diploid individuals.

3.4 Genotype phasing by inference of the shared haplotype

In this section, I extend the IBD detection method such that the allelic sequence of the shared haplotype can be deduced. Since it is assumed that a breakpoint interval covers a region in which two individuals share a haplotype recently co-inherited from a common ancestor, in principle it is possible to infer the shared haplotype sequence based on genealogical constraints. By knowing the sequence of the shared haplotype, it follows that the sequences of the “unshared” haplotypes in both individuals can be derived. The approach presented can therefore be seen as an application to genotype *phasing*, *i.e.* the inference of haplotypes from genotype data, which has become a fundamental problem in genetic research (Browning and Browning, 2011b).

Currently, the most accurate and widely employed phasing algorithms are designed to infer haplotype sharing using Hidden Markov Models (HMMs), which typically operate under a probabilistic model of coalescence with recombination (Stephens *et al.*, 2001; Delaneau *et al.*, 2008). Notable examples include computational tools such as SHAPEIT (Delaneau *et al.*, 2011, 2013) and EAGLE (Loh *et al.*, 2016a,b). The basic phasing process can be described in context of the influential Li and Stephens (2003) model, where haplotypes

are reconstructed from genotype data as “imperfect mosaics” of other haplotypes (*e.g.* using a reference panel). Existing phasing methods typically show high levels of accuracy. However, the phase at low-frequency variants is often problematic to resolve because patterns of allelic variation or linkage disequilibrium (LD) may be too rare to allow correlation with available haplotype information.

Here, I explore the feasibility of phasing genotypes at sites within a given breakpoint interval using the IBD detection method presented in the previous section; haplotypes are therefore inferred *locally* and are limited to the region covered by the interval. In the following, I explain the genealogical constraints that determine the possible haplotype allocation of alleles in genotype data. I then describe the general phasing approach as implemented here.

3.4.1 Genealogical constraints arising from IBD

At a given site, the genotypic states observed at two diploid individuals represent the sum of alleles, where each allele derives from a mutation event on some branch of the genealogical tree of the sample. By knowing that two haplotypes were co-inherited from a recent MRCA, a part of this unseen genealogy is resolved, such that certain genotypes can only be formed if mutation events occurred on certain branches of the tree. This can be seen as being analogous to having partial pedigree information; *i.e.* the two individuals are implicitly treated as the “offspring” of their ancestral “parent”. By following the line of descent, the ancestral haplotype can be inferred if consistent with observed genotypes. However, here, the assumed pedigree is incomplete, because ancestral relations of the two unshared haplotypes are not known from available IBD information. There is nonetheless sufficient information to deduce the allelic states given certain genotypic arrangements.

A set of rules can be formulated which define the state of the shared allele if a certain genotype pair is observed in the two individuals considered. As before, the infinite sites model is assumed. If both genotypes are homozygous for the same allele, the deduction of the shared allele is trivial. If they are homozygous for different alleles, they cannot share an allele. Importantly, because this is seen at breakpoint sites detected using the

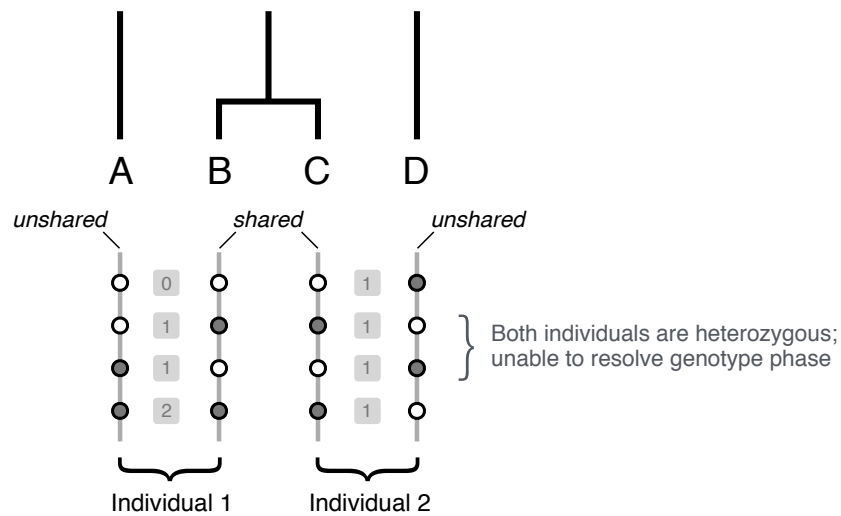


Figure 3.7: Genealogical constraints from haplotype sharing. The top graph indicates the genealogy of four haplotypes (A, B, C, and D) as seen in two individuals (1 and 2), which can be resolved partially if two haplotypes are shared by descent. Below, the allelic states are shown at four positions along the sequence for each haplotype; ancestral and derived states are indicated by *hollow* and *solid* circles. Only the genotypic state is seen in each individual as indicated between individual haplotypes; *i.e.* homozygous for the ancestral allele (0), heterozygous (1), and homozygous for the derived allele (2). The phase of a genotype can be resolved at sites with one homozygous genotype (phasing is redundant if both genotypes are homozygous). Genotype phase cannot be determined if both individuals are heterozygous.

DGT, which mark the limits of the underlying IBD segment, breakpoints are excluded from this analysis; recall that the FGT cannot be used with genotype data. A caveat of this analysis, however, is that haplotypes cannot be distinguished if both individuals are heterozygous, as it is not known whether the allele is shared or not. This conflict arises because an unshared haplotype may be more closely related to either the shared or the other unshared haplotype. Thus, haplotype inference is restricted to sites at which *heterozygous-homozygous* genotype pairs are found. This is further illustrated in Figure 3.7.

A complete definition of these rules is given in Table 3.1. The table gives the inferred allelic state of the shared haplotype for each of the possible combinations of two genotypes. Note that the alleles at *heterozygous-heterozygous* sites are likely to be shared if three conditions are met; first, if the allele frequency at that site is lower than the frequency at the target position, second, if the allele only segregates within the subsample sharing the focal allele and, third, if the genealogy does not change along the sequence within an inferred interval. An allele that is lower in frequency is likely to be younger than the focal

Table 3.1: Shared haplotype inference from genotype pairs. Given the two genotypes observed at a pair of individuals (A and B), there are nine possible ordered combinations of alleles. The corresponding haplotypes are shown for both individuals (unordered). Note that genotype phase cannot be resolved for individuals that are both heterozygous. If a pair is homozygous for different alleles, they do not share a haplotype.

Genotype Individual A	Genotype Individual B	Haplotypes Individual A	Haplotypes Individual B	Shared haplotype
0	0	(0, 0)	(0, 0)	0
0	1	(0, 0)	(0, 1)	0
0	2	(0, 0)	(1, 1)	–
1	0	(0, 1)	(0, 0)	0
1	1	(0, 1)	(0, 1)	–
1	2	(0, 1)	(1, 1)	1
2	0	(1, 1)	(0, 0)	–
2	1	(1, 1)	(0, 1)	1
2	2	(1, 1)	(1, 1)	1

allele and, thus, to **be shared by some of the haplotypes within the sub-tree of the focal mutation**. However, this may only be the case under the infinite sites model, *e.g.* because recurrent or back mutation may otherwise suggest a different genealogical order of descent.

3.5 Evaluation

The IBD detection method presented in this chapter was evaluated using simulated data. This allowed assessment of the accuracy of detected breakpoint intervals in relation to the known genealogy of the simulated sample. For comparison, an alternate IBD detection method was applied to the same data. Lastly, the method presented was applied to data from the 1000 Genomes Project.

3.5.1 Data generation

The coalescent simulator used to generate data was *msprime* (version 0.4.0), which simulates the exact coalescent with recombination, and where mutations are generated under the infinite sites model (Kelleher *et al.*, 2016).^{*} The software is a reimplementation of the classic *ms* algorithm by Hudson (2002), but allows efficient simulation of extended chromosomal regions for very large sample sizes, where the entire history of the simulated

^{*} Coalescent simulator *msprime*: <https://github.com/jeromekelleher/msprime> [Date accessed: 2016-11-12]

sample can be stored and queried for further analysis. Notably, `msprime` allows simulation under variable recombination rates, for example by using established recombination maps of the human genome.

Demographic model

A demographic model was defined following Gutenkunst *et al.* (2009), who used intergenic data from four global populations to estimate parameters from diffusion approximations of expected allele frequency spectra. Accordingly, here, data were simulated with an ancestral population size of $N_e = 7,300$ (denoted by N_A in the model) and under the assumption of a generation time of 25 years. The mutation rate was set to a constant $\mu = 2.35 \times 10^{-8}$ per site per generation, which was estimated from the human-chimp divergence in Gutenkunst *et al.* (2009). Note that recent studies have estimated the human mutation rate to be slightly lower; for example, Scally and Durbin (2012) have estimated $\mu \approx 1.2 \times 10^{-8}$ from analyses of genome-wide *de novo* mutations using recent sequencing technologies, but which is in the same order as the mutation rate used here.

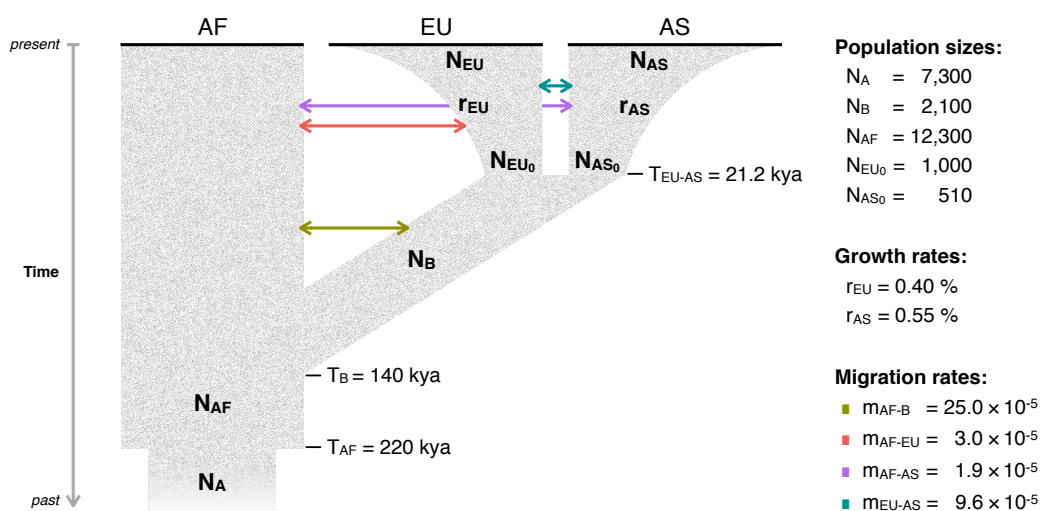


Figure 3.8: Demographic model used in simulations. Three populations were modelled, African (AF), European (EU), and Asian (AS), which derive from an ancestral population (A). Both EU and AS experienced a bottleneck with subsequent exponential growth following the out-of-Africa expansion of a founder population (B) that split from the ancestral population. Modified from Gutenkunst *et al.* (2009), Figure 2 (see [doi:10.1371/journal.pgen.1000695.g002](https://doi.org/10.1371/journal.pgen.1000695.g002)), with parameter values taken from Table 1 (see [doi:10.1371/journal.pgen.1000695.t001](https://doi.org/10.1371/journal.pgen.1000695.t001)).

The demographic history as defined in the simulation model is illustrated in Figure 3.8 (page 103); parameter values of the model are specified therein. The model recapitulates the human expansion out of Africa, for which three populations were considered; African (AF), European (EU), and Asian (AS). The African population was included with a constant population size, while EU and AS experienced exponential growth after divergence and split from an ancestral African population. Population sizes of EU and AS were calculated as $N = N_0/e^{-rt}$, where N is the size at present, N_0 the initial size at EU-AS divergence, r the growth rate, and t the time since divergence (in years).

Simulated dataset

A sample of 5,000 haplotypes was simulated, where the set of generated chromosomes represented a sample taken from the EU population. To reproduce realistic distributions of recombination variability along the simulated sequence, the simulation was performed using recombination rates from human chromosome 20, as provided in Build 37 of the International HapMap Project (HapMap) Phase II (International HapMap Consortium *et al.*, 2007; International HapMap 3 Consortium *et al.*, 2010).^{*} The resulting dataset consisted of 0.673 million segregating sites observed over a chromosomal length of 62.949 Mb (108.267 cM). The history of the simulated sample was stored separately to derive genealogical information in subsequent analyses.

The simulated chromosomes were used to generate three datasets. In the first, haplotypes were randomly paired to construct a sample of 2,500 diploid individuals. From this, second, a corresponding genotype dataset was generated by forming genotypes (encoded as 0, 1, and 2) as the sum of alleles (encoded as 0 and 1) along the sequence in each individual. This dataset was then used to generate a third dataset in which haplotypes were estimated from genotype data; *i.e.* resulting data consisted of phased haplotypes. Phasing was conducted using SHAPEIT version 2 (Delaneau *et al.*, 2008, 2013), using default parameters without a reference panel.[†]

^{*} HapMap recombination map: ftp://ftp.ncbi.nlm.nih.gov/hapmap/recombination/2011-01_phaseII_B37/genetic_map_HapMapII_GRCh37.tar.gz [Date accessed: 2016-11-12]

[†] Phasing software SHAPEIT: https://mathgen.stats.ox.ac.uk/genetics_software/shapeit/shapeit.html [Date accessed: 2016-11-12]

3.5.2 Accuracy analysis

The detection of IBD was evaluated in relation to the underlying true IBD structure of the sample, which was determined from the stored simulation records. Given a target site and the two haplotypes sharing the focal allele, the genealogy was scanned along the sequence of variant sites observed in the sample, in both directions from the target position. The MRCA of the pair was identified at each variant site and a breakpoint was defined as the first site at which a different MRCA was found. This returned the smallest interval detectable from available data around the nearest recombination events that delimit an IBD segment.

Accuracy was measured in terms of the distance between a given breakpoint site and the focal position of the segment. Two measurements were considered; the squared Pearson correlation coefficient, r^2 , which measures the strength of the linear relation between detected and true distance, and the root mean squared logarithmic error (RMSLE);

$$\text{RMSLE} = \sqrt{\frac{1}{n} \sum_{i=1}^n \left[\log_{10} \left(\frac{\hat{d}_i + 1}{d_i + 1} \right) \right]^2} \quad (3.1)$$

where d_i and \hat{d}_i are the distances of the true and detected breakpoints, respectively, and n is the overall number of comparisons. The RMSLE is similar to the root mean squared error (RMSE), which measures the variance and bias in the set of compared values, and is equal to the standard deviation when there is no bias. As such, the RMSLE can be interpreted as a score metric for the magnitude of error. Here, this is useful because larger departures from the actual values are penalised more than smaller ones. A lower score value indicates a lower magnitude of error, where $\text{RMSLE} = 0$ indicates that true and inferred values are identical. Also, note that the RMSLE is usually defined using the natural logarithm; here, \log_{10} was used as a more intuitive representation of error magnitude.

The performance of the proposed IBD detection method was assessed for the described haplotype and genotype-based tests on the three datasets derived from coalescent simulations. The following approaches were distinguished:

- (a) FGT on haplotype data as simulated; *i.e.* *true* haplotypes,
- (b) FGT on *phased* haplotypes, and
- (c) DGT on genotype data.

3.6 Results

IBD detection was carried out on a large set of target sites, for which all f_k variants found at $k \in \{2, \dots, 25\}$ were selected, *i.e.* alleles shared at frequency $\leq 0.5\%$. This threshold was chosen arbitrarily, but such that the considered frequency range was expected to be sufficiently low to identify recent IBD given the size of the sample. The set of target sites comprised 0.317 million SNPs that were heterozygous in the individuals sharing a focal allele. This resulted in 11.598 million pairwise analyses and an equal number of IBD segments detected using the FGT on the true and phased haplotypes in Approaches (a) and (b), respectively, and the DGT on genotype data in Approach (c).

Because the same IBD segment may be inferred from multiple target sites, the analysis was reduced to the set of uniquely detected breakpoint intervals per pair of individuals in each approach. Duplicate segments were removed after the corresponding target variants were sorted by allele frequency, with lower frequencies at the top and removing duplicate segments below. Detected segments were thereby tagged by the presumably youngest shared alleles within the intervals. The number of uniquely identified segments differed slightly in Approaches (a), (b), and (c); 2.983 million (25.723 %), 3.091 million (26.654 %), and 2.978 million (25.679 %), respectively. For the corresponding true IBD segments, the number of unique segments was 3.001 million (25.876 %). These data were further reduced to the intersection of retained target sites across approaches, so as to enable direct comparisons on the same set of targets, which resulted in 2.978 million (25.679 %) unique intervals. The results obtained from these data are summarised in Table 3.2 (next page).

The proportion of breakpoints that were overestimated (in relation to the corresponding true IBD breakpoints) was noticeably high overall; using the FGT, 97.390 % and 95.666 % were overestimated in Approaches (a) and (b), respectively. However,

Table 3.2: Accuracy of detected breakpoints per f_k category. The accuracy of detected IBD breakpoints was measured using the squared Pearson correlation coefficient, r^2 , and the RMSLE in relation to the true IBD segments determined from simulation records; measured in terms of the distance between breakpoint site and the corresponding focal position per segment. The analysis included of 317,020 target sites around which IBD was detected in Approaches (a), (b), and (c). In each, accuracy was computed after data were reduced to identical sets of unique IBD segments ($n = 2,978,220$). The table specifies the allele frequency (%) corresponding to each f_k category, as well as the number of target sites identified.

f_k	Freq. %	Targets	r^2			RMSLE		
			FGT*	FGT**	DGT [†]	FGT*	FGT**	DGT [†]
2	0.04	76,515	0.998	0.895	0.995	0.219	0.598	0.317
3	0.06	46,138	0.989	0.957	0.978	0.243	0.516	0.359
4	0.08	31,658	0.963	0.959	0.941	0.256	0.463	0.379
5	0.10	23,581	0.975	0.963	0.929	0.276	0.429	0.408
6	0.12	19,241	0.954	0.938	0.904	0.281	0.409	0.421
7	0.14	15,869	0.955	0.944	0.892	0.298	0.403	0.447
8	0.16	13,175	0.898	0.918	0.813	0.320	0.398	0.469
9	0.18	10,966	0.932	0.927	0.827	0.314	0.375	0.467
10	0.20	11,142	0.879	0.887	0.773	0.332	0.387	0.494
11	0.22	9,392	0.895	0.892	0.758	0.344	0.401	0.513
12	0.24	7,751	0.835	0.848	0.733	0.358	0.398	0.526
13	0.26	6,933	0.842	0.835	0.721	0.361	0.405	0.532
14	0.28	5,767	0.816	0.816	0.679	0.367	0.391	0.540
15	0.30	5,062	0.871	0.860	0.712	0.381	0.406	0.556
16	0.32	4,711	0.839	0.830	0.701	0.373	0.395	0.546
17	0.34	4,210	0.829	0.832	0.681	0.387	0.410	0.566
18	0.36	3,913	0.813	0.832	0.670	0.380	0.397	0.561
19	0.38	3,684	0.801	0.798	0.642	0.381	0.401	0.566
20	0.40	3,214	0.831	0.837	0.685	0.401	0.416	0.587
21	0.42	3,333	0.773	0.778	0.603	0.399	0.413	0.584
22	0.44	2,863	0.753	0.795	0.571	0.399	0.406	0.586
23	0.46	2,595	0.732	0.745	0.596	0.414	0.425	0.599
24	0.48	2,653	0.784	0.780	0.581	0.396	0.406	0.583
25	0.50	2,654	0.701	0.730	0.560	0.400	0.408	0.585

* Approach (a), using the FGT with true haplotypes

** Approach (b), using the FGT with phased haplotypes

[†] Approach (c), using the DGT with genotype data

overestimation was highest when the DGT was used (98.362 %) in Approach (c). Conversely, the proportion of underestimated breakpoints was lowest in (c), 1.543 %, and highest when haplotypes were phased in (b), 4.147 %; in (a), 2.418 % of breakpoints were underestimated. The proportion of detected breakpoints that coincided with the corresponding true breakpoints was 0.192 %, 0.188 %, and 0.095 % in (a), (b), and (c), respectively.

The highest overall accuracy was found for the FGT on true haplotypes, followed by the analysis on phased haplotypes, which had $r^2 = 0.926$ and $r^2 = 0.892$ in Approaches (a) and (b), respectively. The accuracy achieved by the DGT was lower, but still considerably high with $r^2 = 0.847$ in Approach (c). This was also reflected in the measured magnitude of error (RMSLE), which was 0.400, 0.434, and 0.569 in (a), (b), and (c), respectively. The measurement of accuracy was further broken down by the allele frequency of target variants (f_k category); results are shown in Table 3.2 (page 107). Accuracy decreased towards higher allele frequency in each approach. For example, for f_2 variants, r^2 was 0.998, 0.985, and 0.995 in (a), (b), and (c), respectively, which was reduced for f_{25} variants where $r^2 = 0.701$ in (a) and $r^2 = 0.730$ in (b), but where (c) was seen to decrease more rapidly by comparison ($r^2 = 0.560$). Notably, when haplotype data were phased in (b), accuracy was highest at f_5 variants ($r^2 = 0.963$), indicating that accuracy was decreased at lower frequencies. Similarly, RMSLE scores reflected the same general pattern, but where the magnitude of error in (b) was at a maximum at f_2 variants.

A more intuitive representation of results is provided in Figure 3.9 (next page), which compares true and detected breakpoint distances in two ways. First, in Figure 3.9A, breakpoint densities are shown in separate scatterplots for breakpoints detected on the left and right-hand side of focal positions. For example, a clear difference in the proportion of underestimated breakpoints can be seen between Approaches (a) and (b), *i.e.* where the FGT was used on true and phased haplotypes, respectively. In Approach (c), where the DGT was used on genotype data, breakpoint densities indicate a higher proportion of overestimated distances compared to (a) or (b). Second, in Figure 3.9B, the relative distance was calculated as $x = \hat{d}_i/d_i$, where \hat{d} and d denote detected and true distances, respectively. By doing so, detected breakpoint distances were “mapped” relative to the corresponding true distances, such that $0 < x < 1$ indicates underestimation and $x > 1$ indicates overestimation. The cumulative distribution function (CDF) of the relative distance is shown separately per f_k category. For example, it can be seen that a larger proportion of f_2 variants (15.215 %) contributed to the overall underestimation found in Approach (b), *e.g.* compared to f_5 (5.544 %) and f_{25} variants (0.755 %).

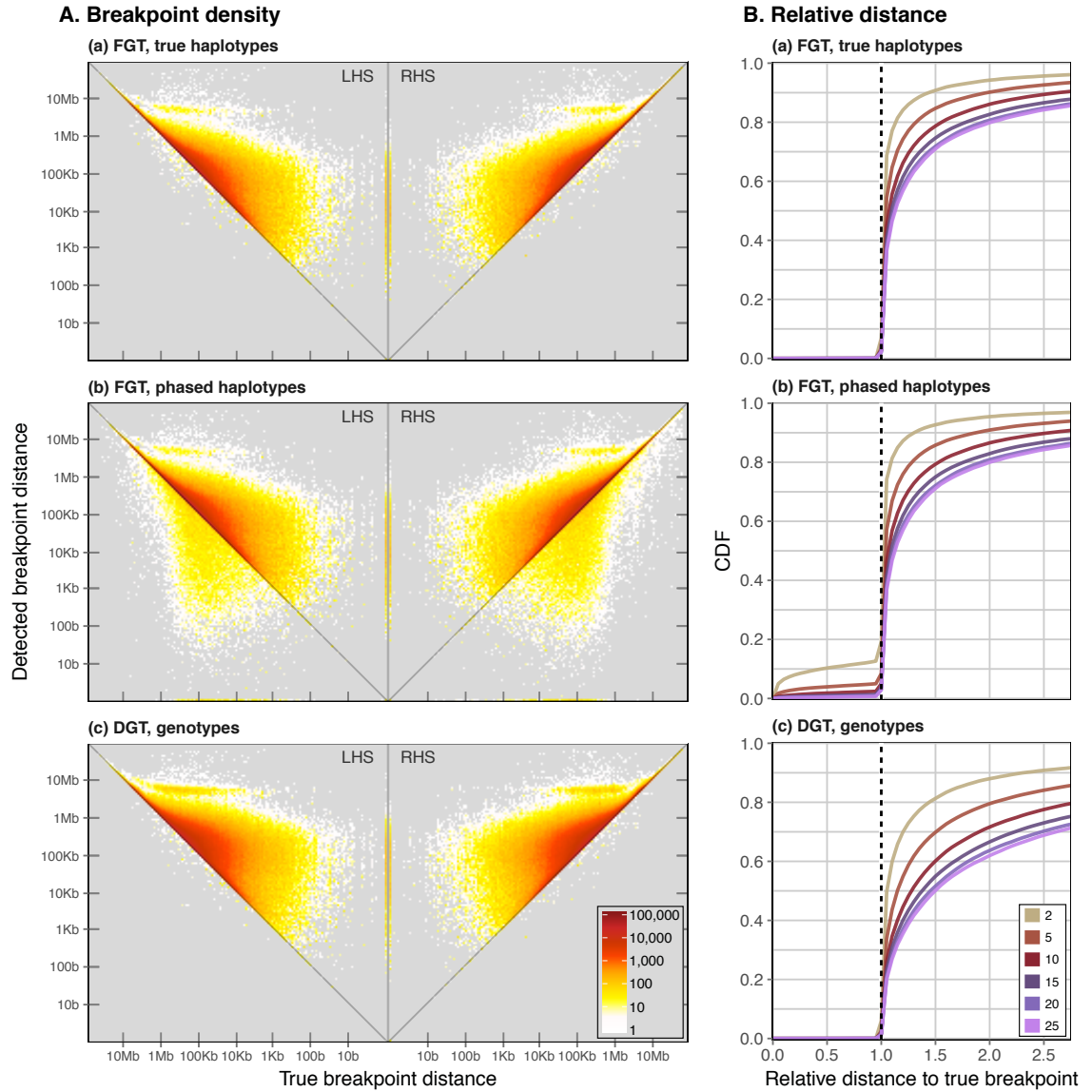


Figure 3.9: Accuracy of breakpoint detection in simulated data. Breakpoints detected in f_k pairs at $k \in \{2, \dots, 25\}$ are compared to true IBD breakpoint sites, after removing boundary cases in either the detected or true dataset. Segments were inferred using the FGT on true haplotypes (a) and phased haplotypes (b), as well as the DGT on genotype data (c). Panel (A) shows the density of detected breakpoints in relative distance to the focal and true breakpoint sites. The physical distance between detected breakpoint and focal position was divided by the distance between true breakpoint and focal position, such that values < 1 indicate underestimation and > 1 overestimation relative to the true distance (*dashed line*). Panel (B) illustrates the relationship between each detected breakpoint and the corresponding true breakpoint, measured as the physical distance to the focal site. Along each axis, distances were pooled into 200 bins (on log scale) and cells in the resulting 200^2 grid were colour-coded for the number of intersecting true and detected breakpoints, where grey indicates zero. Segment breakpoints to the left (*LHS*) and right-and side (*RHS*) of the focal position are shown separately.

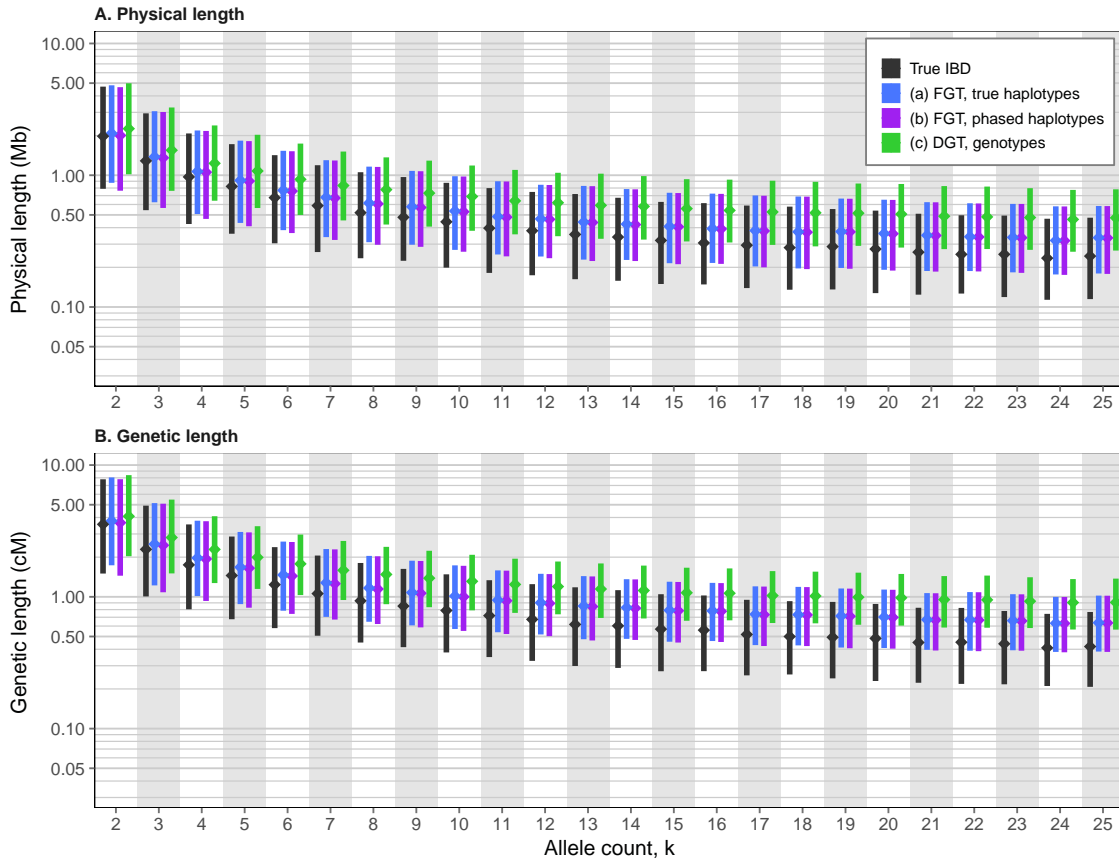


Figure 3.10: IBD segment lengths inferred in simulated data. The distribution of median physical and genetic length of detected IBD segments is shown by allele frequency of the focal variant ($f_{[2,25]}$). IBD detection was performed using the FGT on true and phased haplotypes, as well as the DGT on genotype data; Approaches (a), (b), and (c), respectively. The true IBD length is shown for comparison. Bottom and top of each bar indicate 1st and 3rd quartiles, respectively, between which the median (2nd quartile) is marked (*diamonds*).

The distribution of physical and genetic IBD length is shown in Figure 3.10 (this page). These results were obtained after boundary cases were removed in each approach (*i.e.* discarding segments where the end of a chromosome was reached without detecting a breakpoint), so as to ensure that observed IBD length was delimited by recombination on both sides of a segment; 1.449 %, 1.400 %, and 1.637 % was removed in (a), (b), and (c), respectively, and 1.340 % in the set of true IBD segments. Data were then intersected again to retain the same set of target sites in each approach; as a result, 2.929 million unique segments were retained (98.363 %).

Median physical length (and median genetic length) over the set of retained segments was computed for each approach. A small difference was seen for the FGT, where

median length was 0.417 Mb (0.800 cM) on true haplotypes in Approach (a), and 0.413 Mb (0.791 cM) on phased haplotypes in Approach (b). For the DGT on genotype data, median length was longer by comparison, 0.570 Mb (1.094 cM). The median of true IBD length was 0.328 Mb (1.573 cM), which was shorter than detected in each approach. But as seen in Figure 3.10, the distribution of IBD lengths in (a), (b), and (c) closely followed the true lengths along the allele frequency range. However, the gap between true and detected lengths increased towards higher allele frequencies. For example, for f_2 variants, median length of true IBD segments was 1.978 Mb (3.551 cM), which is only marginally shorter compared to 2.079 Mb (3.773 cM), 2.005 Mb (3.652 cM), and 2.256 Mb (4.085 cM) in (a), (b), and (c), respectively. For f_{25} variants the difference was more pronounced; *i.e.* median length of true IBD segments was 0.243 Mb (0.419 cM), compared to 0.336 Mb (0.636 cM), 0.335 Mb (0.634 cM), and 0.475 Mb (0.907 cM) in (a), (b), and (c), respectively.

In summary, the FGT on true haplotype data in Approach (a) overall achieved the highest levels of accuracy while maintaining low error. This was particularly seen in comparison to Approach (b), which differed only in the additionally included phasing step. Since genomic datasets were typically composed of phased haplotypes, Approach (b) can be seen as being the realistic approach. However, the higher rate of error at lower frequency variants may pose a problem for analysis for rare variants. As an alternative, the DGT on genotype data, Approach (c), can be used to detect IBD with high accuracy and comparatively low error rates. However, the larger proportion of overestimated IBD breakpoints may result in additional error, *e.g.* if it is assumed that the genealogy is consistent along the sequence of inferred IBD segments.

IBD detection using the *Refined IBD* algorithm

Simulated data were additionally analysed using the Refined IBD algorithm implemented in Beagle version 4.1 (Browning and Browning, 2013).^{*} The method is based on the non-probabilistic GERMLINE algorithm (Gusev *et al.*, 2009), which identifies putative IBD segments from short exact matches between haplotype pairs; candidate segments are

^{*} Beagle 4.1: <https://faculty.washington.edu/browning/beagle/beagle.html> [Date accessed: 2016-11-22]

found by extending identified regions to longer inexact matches. In Refined IBD, an additional probabilistic approach is included to assess candidate segments conditional on the likelihood ratio (LR) of the data, calculated under IBD and non-IBD models. A logarithm of odds (LOD) score is calculated as $\log_{10}(\text{LR})$, and segments are reported as IBD if the LOD score is above a specified threshold. This approach has been found to achieve greater accuracy than GERMLINE alone or fastIBD, which is a non-probabilistic method that detects IBD based on haplotype frequency (Browning and Browning, 2011a, 2013).

The analysis was performed using default parameters in Refined IBD (retaining candidate segments at $\text{LOD} > 3.0$) and after conversion of simulated data into Variant Call Format (VCF)*. Note that haplotype data are required; therefore only the FGT was evaluated using true and phased haplotype data in Approaches (a) and (b), respectively. The analysis returned 13.689 million IBD segments in (a), of which 0.248 % were duplicated, and 13.647 million in (b), of which 0.249 % were duplicated. The median length of all detected segments was 0.191 Mb and 0.191 Mb in (a) and (b), respectively (after removal of duplicates).

The accuracy of detected IBD segments was measured in relation to the true IBD intervals, which were already determined for the set of previously analysed target sites; *i.e.* all $f_{[2,25]}$ variants found in the data (allele frequency $\leq 0.5\%$). Note that the detection approach employed by Refined IBD reports all segments inferred for a given pair of haplotypes, such that detected and true intervals cannot be matched by direct reference to a particular target site. Hence, for a given pair of haplotypes, true and detected segments were matched if the focal allele associated with a true segment fell within the interval of the detected segment, which was discarded if none of the pairwise shared target alleles were found within the detected interval. This matching process resulted in a set of 2.166 million unique segments in the analysis conducted on true haplotypes, Approach (a). For the phased dataset, Approach (b), 0.528 million unique segments were matched. The lower number of segments retained in Approach (b) is likely to be the consequence of mismatched haplotypes in original and phased data.

* Variant Call Format: <http://vcftools.sourceforge.net/VCF-poster.pdf> [Date accessed: 2016-11-22]

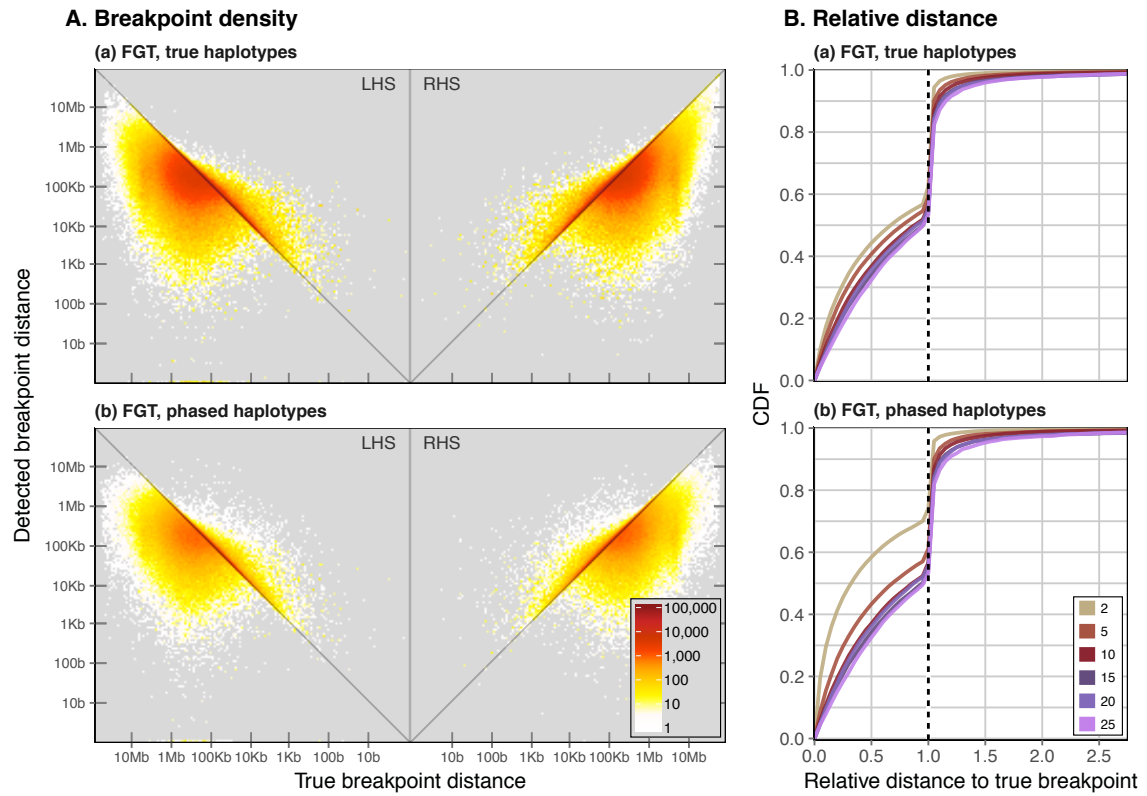


Figure 3.11: Accuracy of breakpoint detection in simulated data using Refined IBD in Beagle 4.1. Results are shown for one million randomly selected shared haplotype segments, after removing boundary cases in either the detected or true segments. Segments were inferred using the FGT on true haplotypes (a) and phased haplotypes (b), as well as the DGT on genotype data (c). Panel (A) shows the density of detected breakpoints in relative distance to the focal and true breakpoint sites. For each detected breakpoint, its physical distance to the focal site was divided by the distance between the corresponding true breakpoint and the focal site, such that values < 1 indicate underestimation and > 1 overestimation of the true distance (dashed line). Panel (B) provides a heatmap representation of a scatter plot, comparing physical distances between focal site and true breakpoint (x-axis) and detected breakpoint (y-axis). Along each axis, distances were pooled into 200 bins (on log scale) and the resulting 200^2 squares were colour-coded for the number of intersecting true and detected breakpoints, where grey indicates zero. Segment breakpoints to the left (LHS) and right-hand side (RHS) of the focal position are shown separately.

After the removal of segments at which haplotype pairs did not share any of the alleles in $f_{[2,25]}$, median lengths were longer in comparison to the original datasets; 0.255 Mb and 0.256 Mb in (a) and (b), respectively. This can be seen as the result of both the removal of falsely identified segments, as well as segments that were older and thereby expected to be shorter.

The majority of retained breakpoints was underestimated in both Approaches (a) and (b); 55.356 % and 56.606 %, respectively. In Approach (a), 44.385 % were overestimated and 0.259 % coincided with true breakpoint positions. This was similar in Approach (b), where 43.155 % were overestimated and 0.238 % coincided. Accuracy was measured in terms of the physical distance between breakpoint position and the corresponding focal site. Because the latter was not specified in the results obtained from Refined IBD, the same focal position as associated with the matched true IBD segment was assumed. Data were not reduced to the intersection of segments retained across datasets, due to the low number of consistent matches. These results are illustrated in Figure 3.11 (page 113).

The distance distribution of true and detected breakpoints (Figure 3.11A) suggests that detected breakpoints were closely distributed around the corresponding true breakpoints. However, overall accuracy was low in both (a) and (b), reaching $r^2 = 0.287$ and $r^2 = 0.171$, respectively. The magnitude of error, RMSLE, was lower in (a) compared to (b); 0.595 and 0.616, respectively. When true haplotypes were analysed, Approach (a), accuracy decreased steadily towards higher allele frequencies. For example, accuracy was highest for f_2 variants ($r^2 = 0.346$) but lowest for f_{25} variants ($r^2 = 0.074$). However, the magnitude of error was highest for f_2 variants (RMSLE = 0.750) and lowest for f_{25} variants (RMSLE = 0.543). When haplotypes were phased, Approach (b), error was further increased at f_2 variants (RMSLE = 0.999) in comparison to f_{25} variants (RMSLE = 0.546). The higher error at lower allele frequencies was also reflected in r^2 values; *e.g.* accuracy was low at f_2 ($r^2 = 0.087$), higher at f_5 ($r^2 = 0.132$), but lowest at f_{25} ($r^2 = 0.048$). The difference between true and phased datasets is further highlighted in Figure 3.11B, where a higher proportion of f_2 variants is seen to be underestimated in (b).

The distribution of physical and genetic lengths for the segments retained in Approaches (a) and (b) are shown in Figure 3.12 (next page), in relation to the true IBD lengths at each f_k category. Because (a) and (b) were compared on different sets of detected segments, the reported lengths of true segments were computed from the set matched

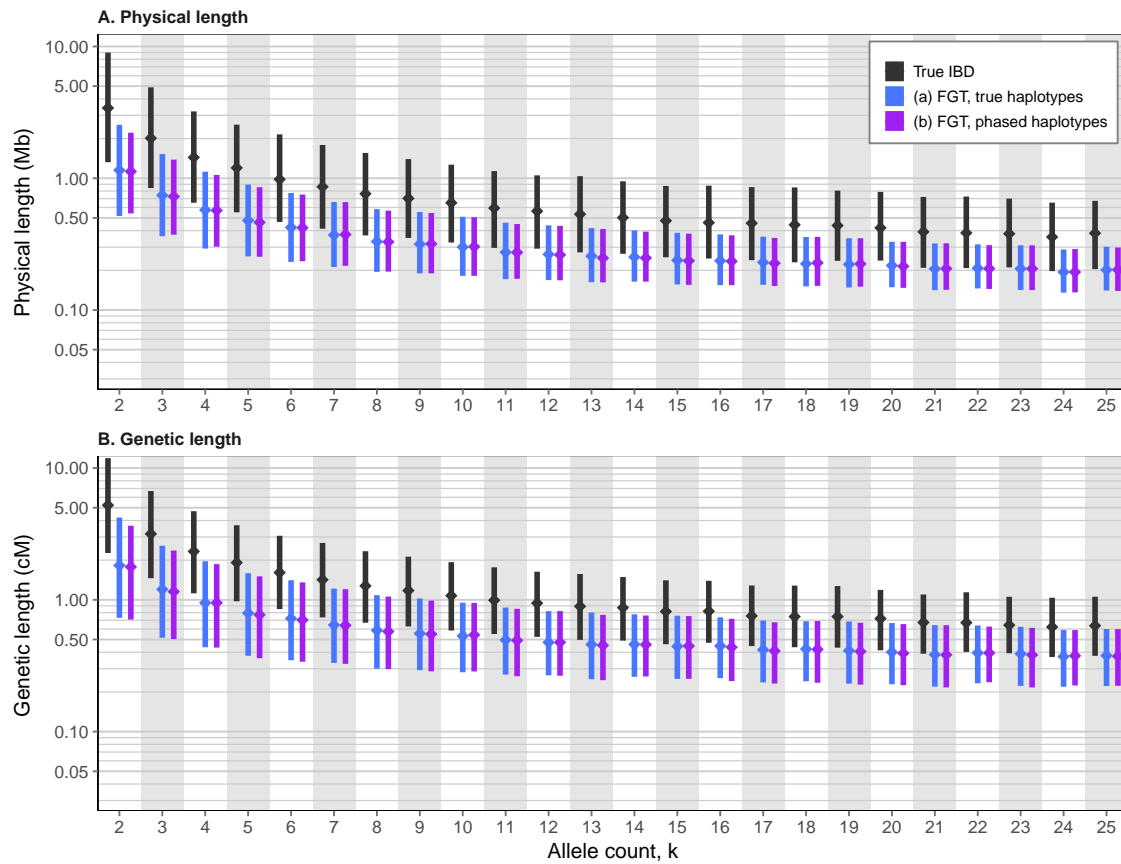


Figure 3.12: IBD segment lengths inferred using *Refined IBD* in Beagle 4.1. The distribution of median physical and genetic segment length is shown by allele count (f_k category). IBD segments were estimated using the *Refined IBD* algorithm implemented in Beagle 4.1. Because the method requires haplotype data, only the FGT was evaluated using true haplotypes, Approach (a), and phased haplotypes, Approach (b). These are compared to the true lengths of corresponding IBD tracts. Bottom and top of each bar indicate 1st and 3rd quartiles, respectively, between which the median (2nd quartile) is marked (*diamonds*).

to (a). Boundary cases were removed to avoid potential bias in length comparisons; 1.012% and 0.893% in (a) and (b), respectively.

Overall median physical length (and median genetic length) was 0.185 Mb (0.621 cM) in (a) and 0.192 Mb (0.429 cM) in (b), but both were shorter in comparison to true IBD segments at 0.260 Mb (0.912 cM). At f_2 variants, the median length of IBD segments inferred in (a) was 1.150 Mb (1.825 cM), which was longer compared to (b), where median length was 1.126 Mb (1.777 cM). However, both were considerably shorter in comparison to the true segments, 3.413 Mb (5.229 cM). This difference persisted towards higher allele frequencies; *e.g.* for f_{25} variants, where the median of true lengths was

0.383 Mb (0.636 cM), which was longer compared to 0.202 Mb (0.378 cM) in (a) and 0.202 Mb (0.373 cM) in (b).

While this analysis does not permit to make statements about falsely identified IBD segments, as these were among the segments removed in the matching process, the results presented showed that the lengths of inferred breakpoint intervals are likely to be shorter than the underlying haplotype region shared by descent. Thus, the Refined IBD algorithm is less accurate with regard to the inference of the recombination events that delimit the underlying IBD tract.

IBD detection in real data: 1000 Genomes, chromosome 20

The IBD detection method presented was applied to the final release dataset of the 1000 Genomes Project Phase III (1000 Genomes Project Consortium *et al.*, 2012, 2015), which included $N = 2,504$ individuals. IBD detection was performed for each autosome (chromosomes 1–22), where selected target sites comprised all shared rare variants at allele frequency $\leq 0.5\%$; *i.e.* f_k where $k \in \{2, \dots, 25\}$. However, to enable a closer comparison to the results obtained on the simulated dataset (which simulated variable recombination rates as inferred for chromosome 20), the following results are presented for chromosome 20 only. A summary of the IBD detection results for chromosomes 1–22 is given in Table 3.3 (next page).

Data were available as phased haplotypes, which enabled the analysis using both the FGT and DGT; *i.e.* the results produced can therefore be seen as being analogous to Approach (b) and Approach (c), respectively. In each, 18.0 million IBD segments were inferred, of which 43.2 % were unique for the FGT, and 39.4 % for the DGT. After removal of boundary cases (0.194 % and 0.285 % for the FGT and DGT, respectively), data were intersected to retain a common set of target sites in the analysis, which retained 7.069 million unique segments.

As there is no “truth” dataset that could serve as a reference to measure accuracy, the following analysis was limited to the quantitative description of the inferred IBD lengths. These results are shown in Figure 3.13 (page 118). Median physical length

Table 3.3: Inferred IBD length per chromosome in 1000 Genomes. Shared haplotype segments in 1000G Phase III were inferred using the FGT and DGT, on data from 2,504 individuals across all autosomes. Pairwise shared segments were identified from rare variants at allele frequency $\leq 0.5\%$ ($f_{[2,25]}$). Median genetic and physical lengths over all inferred segments were calculated per chromosome, after removing boundary cases and retaining unique segments only.

Chr.	SNPs	Targets	Segments	Unique (%)		Length (Mb)		Length (cM)	
				FGT*	DGT**	FGT*	DGT**	FGT*	DGT**
1	6,196,151	2,126,720	64,449,399	40.3	35.9	0.125	0.237	0.150	0.300
2	6,786,300	2,323,889	70,274,554	38.1	33.8	0.136	0.248	0.143	0.280
3	5,584,397	1,893,872	57,220,884	37.1	33.2	0.138	0.243	0.154	0.290
4	5,480,936	1,847,521	57,598,118	36.6	32.8	0.138	0.247	0.150	0.283
5	5,037,955	1,716,580	53,055,802	36.4	32.8	0.139	0.245	0.158	0.293
6	4,800,101	1,625,828	50,544,859	37.0	33.0	0.133	0.238	0.148	0.280
7	4,517,734	1,546,940	47,303,666	39.2	34.8	0.119	0.218	0.139	0.270
8	4,417,368	1,519,028	46,250,487	37.3	33.4	0.119	0.212	0.140	0.268
9	3,414,848	1,171,960	35,718,922	40.6	36.6	0.110	0.203	0.156	0.296
10	3,823,786	1,313,699	40,488,078	39.6	35.3	0.114	0.210	0.154	0.299
11	3,877,543	1,318,559	39,668,383	38.3	34.2	0.128	0.228	0.148	0.283
12	3,698,098	1,255,880	38,116,079	39.4	35.3	0.124	0.221	0.164	0.311
13	2,727,881	919,222	28,252,993	38.9	35.2	0.126	0.222	0.166	0.305
14	2,539,149	861,549	25,955,712	39.5	35.6	0.119	0.214	0.157	0.299
15	2,320,474	795,882	23,977,630	42.6	38.2	0.100	0.183	0.153	0.304
16	2,596,072	901,185	26,907,909	43.5	38.3	0.081	0.153	0.140	0.286
17	2,227,080	775,133	22,914,233	44.5	39.8	0.096	0.175	0.150	0.300
18	2,171,378	739,822	22,405,301	41.5	37.7	0.109	0.193	0.169	0.311
19	1,751,878	607,451	18,033,860	46.1	41.3	0.079	0.146	0.147	0.293
20	1,739,315	599,065	18,040,053	43.2	39.4	0.102	0.180	0.182	0.339
21	1,054,447	365,330	11,051,666	44.7	40.4	0.090	0.172	0.162	0.312
22	1,055,454	363,748	10,748,355	47.2	42.5	0.070	0.133	0.145	0.291
<i>Total</i>		77,818,345	26,588,863	808,976,943					

* 1000G data are available as phased haplotypes; hence, results are analogous to Approach (b).

** Conducted on genotype data; hence, results are analogous to Approach (c).

(and median genetic length) over the whole set of retained IBD segments was 0.101 Mb (0.188 cM) using the FGT and 0.180 Mb (0.339 cM) using the DGT. As was seen in the analysis of simulated data, the DGT generally is more likely to overestimate breakpoint distance, leading to the discovery of longer intervals. This discrepancy in length was more pronounced for f_2 variants, for which median length was 0.124 Mb (0.195 cM) using the FGT and 0.253 Mb (0.428 cM) using the DGT. Notably, IBD lengths were more than twice as long in half of the detected segments using the DGT, compared to the FGT. The length of segments identified at lower frequencies was longer in comparison to higher frequencies; *e.g.* for f_{25} variants, median length was 0.084 Mb (0.165 cM) and 0.149 Mb

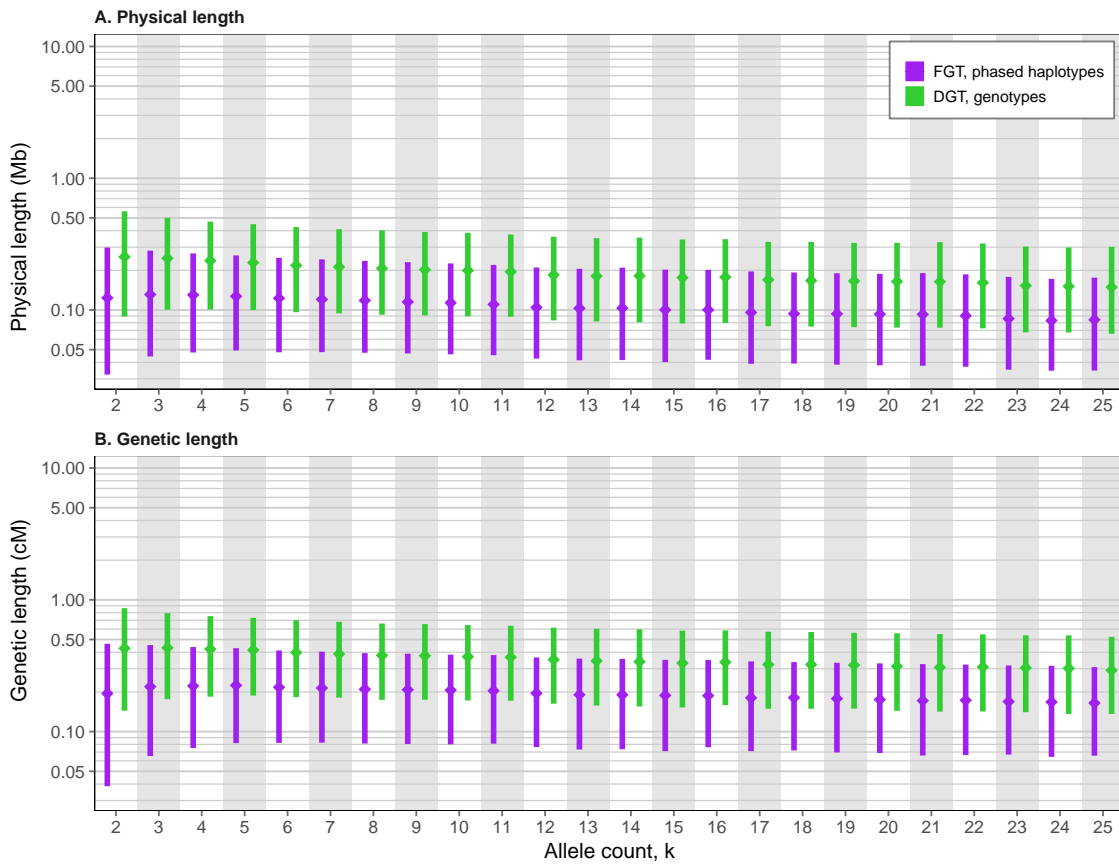


Figure 3.13: Distribution of inferred IBD lengths in 1000 Genomes data, chromosome 20.

Results are shown for the detected physical and genetic lengths of shared haplotype segments by f_k , using chromosome 20 in the final release dataset of 1000 Genomes Project Phase III, including $N = 2,504$ individuals. IBD segments were detected using the FGT (on phased haplotypes) and the DGT (on genotype data). Bottom and top of each bar represent the 1st and 3rd quartile, respectively, between which the median (2nd quartile) is marked (*diamonds*).

(0.292 cM) using the FGT and DGT, respectively. However, the IBD lengths were highest at $f_{[3,5]}$ when the FGT was used, but which was not the case for the DGT.

Inference of the shared haplotype sequence

The IBD-based phasing concept described in Section 3.4 (page 99) was explored in this section. Given the genealogical constraints that follow from the assumption that the genealogy does not change along the sequence within a given breakpoint interval, it is straightforward to derive the sequence of the shared haplotype. This applies to sites where exactly one genotype is heterozygous and one is homozygous in a pair of individuals. In the following, such heterozygous-homozygous sites are referred to as being *informative*, as

opposed to sites where both genotypes are heterozygous, which are referred to as being *indeterminate*. At sites where both genotypes are homozygous, phasing is trivial. Hence, phasing is attempted at heterozygous sites along the sequence.

The analysis was performed on the full set of IBD segments detected using the DGT in the simulated dataset; recall that the FGT requires (phased) haplotypes and cannot be used on genotype data. Also, note that breakpoint sites delimit the interval enclosing a haplotype that is shared by descent, but they themselves represent the first positions along the sequence at which haplotype sharing was broken through recombination; hence, breakpoint sites were excluded from the inference. The purpose of this evaluation was to determine the proportions of correctly and incorrectly phased sites, as well as the proportion of indeterminate sites. This was then compared to the same haplotypes inferred using the set of true IBD segments.

In the following, two approaches were explored; first, haplotype inference was performed separately per IBD segment and, second, the segments identified by the same shared allele were grouped and a majority-rule was applied to determine the shared haplotype sequence (referred to as *consensus* approach). The latter represents an attempt to increase the number of informative sites, as these may differ if more pairs of individuals are considered. Since they share the same focal allele, it is assumed that this allele identifies the same haplotype in all pairs.

A random subset of 10,000 IBD segments was drawn for each f_k category from the full set of detected intervals using the DGT in which the shared haplotype was inferred. Similarly, for the consensus approach, target sites were randomly drawn for the set of identified targets for each f_k category, such that the total number of identified segments did not exceed 10,000 per f_k category. Haplotypes were then inferred and combined in each focal group by applying a majority-rule to estimate the sequence of the shared haplotype. For each segment, the proportions of correctly inferred, incorrectly inferred, or indeterminate alleles was then determined per f_k category. Additionally, to compare these results to the “truth”, the same subsets of IBD segments was analysed in each approach, but where estimated haplotype length was delimited by the corresponding true IBD segments.

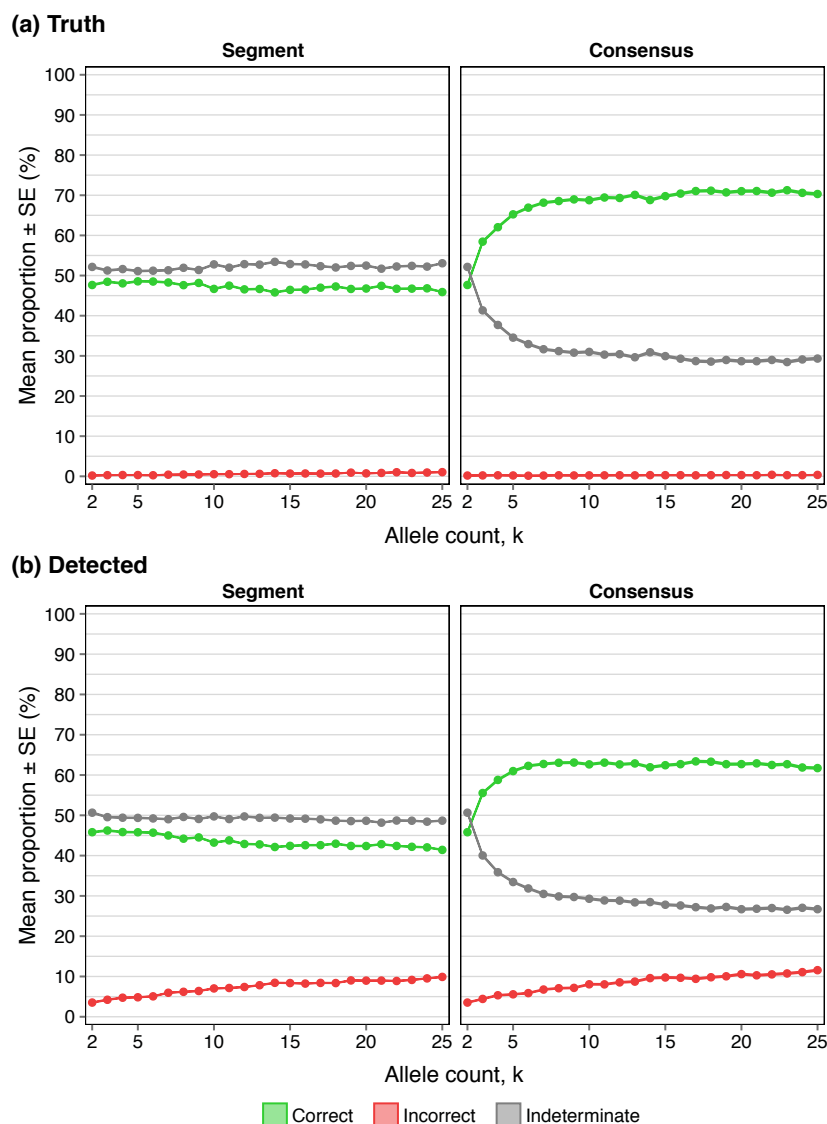


Figure 3.14: Accuracy of alleles inferred through IBD-based phasing by focal allele frequency. IBD segments were randomly selected; 10,000 per f_k category. Genotypes within each segment were phased and the relative proportions of correct, incorrect, and indeterminate alleles were recorded and averaged (\pm SE) per f_k category. This was done separately per segment (*left*) and using the consensus approach (*right*). The results shown in Panel (a) correspond to the “truth”, where the true IBD segments were used to delimit the extent of the inferred shared haplotype sequence. This is compared to Panel (b), where the DGT was used to detect IBD breakpoints in genotype data.

The same was done for both approaches, but using the set of true IBD segments to delimit the haplotype sequence. The results of this analysis are shown in Figure 3.14 (this page).

When true IBD segments were used, Figure 3.14a, the proportion of incorrectly phased sites was consistently low along the allele frequency range; however, note that

this proportion was not equal to zero. In comparison, the proportion of error increased towards higher allele frequencies when IBD was detected using the DGT, Figure 3.14b, which was seen in both the segment-based approach and the consensus approach. Notably, the proportion of indeterminate sites was consistently close to 50% in the segment-based approach using both true and detected data. However, the proportion of correctly phased genotypes decreased towards higher allele frequencies; from 46.173 % at f_2 variants to 42.330 % at f_{25} variants. In the consensus approach, the proportion of correctly phased genotypes showed a rapid increase but then reached a plateau around 70% in when true IBD segments were used, and between 60% and 65% when detected segments were used.

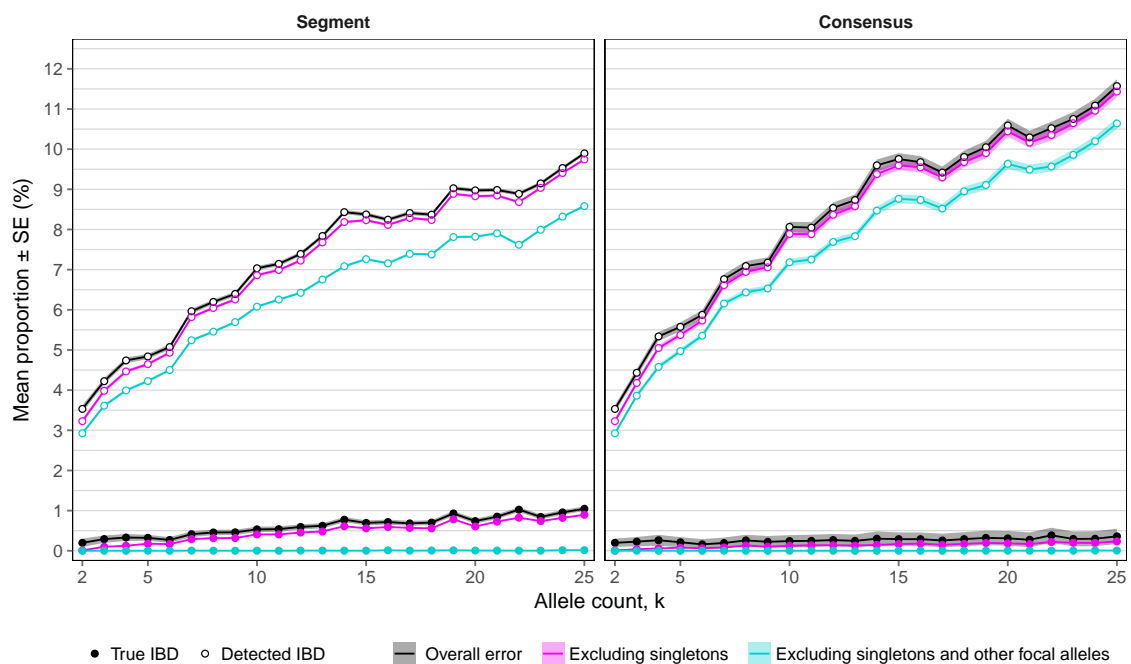


Figure 3.15: Error distribution of alleles inferred through IBD-based phasing. The same results are shown as in Figure 3.14, but with focus on incorrectly inferred alleles only; *i.e.* the proportion of overall error corresponds to the proportion of incorrect alleles in Figure 3.14. The source of error was further

The distribution of incorrectly phased genotypes was further analysed to distinguish error due to singletons and other focal alleles that fell within the intervals of the segments analysed. These results are shown in Figure 3.15 (this page). Note that in both, the exclusion of singletons did not markedly reduce the overall proportion of

incorrectly phased genotypes. In both the segment-based approach and the consensus approach, error was noticeably reduced if genotypes at other focal variants were excluded. Importantly, by excluding singletons and other focal alleles that fell within the interval of the segments analysed, error was reduced to zero when true IBD segments were used. In fact, incorrectly phased genotypes only occurred within the overestimated regions of detected IBD segments; *i.e.* using the DGT.

This finding is important as it suggests that haplotypes could be determined correctly if overestimation in the detection of breakpoints could be reduced or excluded. For example, instead of considering the full length of the segments detected, genotype phase could be determined from a narrow interval around a given target site.

Phasing coverage of pairwise shared haplotypes per individual

Although the presented IBD-based phasing approach attempted to only determine genotype phase locally, it may nonetheless be of interest to explore the possibility to phase individuals “globally”; that is, to distinguish haplotypes from genotype data along the full length of a chromosome per individual. To this end, it is relevant to determine the coverage of shared haplotypes along a chromosome per individual. Since simulated data may not be suited to derive expectations that would apply to reality, the following analysis was conducted on data from the 1000 Genomes Project Phase III, chromosome 20.

I randomly selected 500 of the 2,504 individuals and, for each individual, I identified all other individuals (in the whole sample of 2,504 individuals) which shared any rare allele at $f_{[2,25]}$ with a given individual. The average number of rare alleles shared between a given individual and any other individual was 15200 (± 455 SE), and the average number of unique individuals who shared any rare allele with a given individual was 1400 (± 9.65 SE), which ranged between 888 and 1,978. The DGT was used to detect segment breakpoints per pairwise shared focal allele; for comparison, the same was done using the FGT, although only the DGT would be applied in the context of genotype phasing.

Each set of detected segments was then reduced to include only unique IBD intervals; again after sorting by focal allele frequency such that the focal allele is the one with

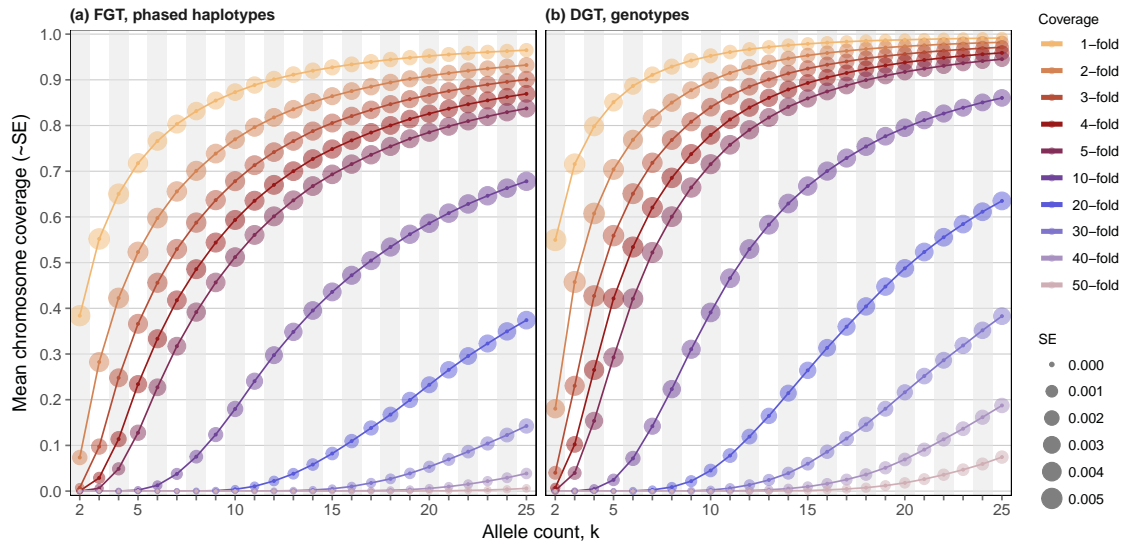


Figure 3.16: Cumulative shared haplotype coverage by focal allele count in 1000 Genomes data. Results are shown for 500 randomly selected individuals, among $N = 2,504$ in the 1000G dataset. Segments detected using the FGT (a, left) are compared to the segments detected using the DGT (b, right). Coverage was defined as the proportion of chromosome length covered by inferred and aligned shared haplotypes, where n -fold indicates that the chromosome was covered n -times by all segments up to a given f_k category (hence, cumulative). Standard error (SE) is indicated by the size of dots (see figure legend).

lowest frequency within a given segment. Recall that duplicate segments refer to IBD intervals that are tagged by multiple f_k variants, which are assumed to sit on the same underlying shared haplotype. The average number of unique segments retained per individual was $5070 (\pm 80.0 \text{ SE})$ for the DGT and $5460 (\pm 85.3 \text{ SE})$ for the FGT; note that haplotype data were phased. Retained haplotype segments were aligned by position along the chromosome to measure coverage, *i.e.* the proportion of the chromosome covered. Mean coverage across all 500 randomly selected individuals is given in Figure 3.16 (this page), which indicates the n -fold cumulative coverage by f_k category; shown for true IBD segments (Figure 3.16a) and segments detected using the DGT (Figure 3.16b).

Coverage was overall higher for the DGT compared to the FGT, due to overestimating IBD lengths. Notably, neither the DGT nor the FGT reached full coverage on average. However, a steady increase was seen when f_k variants at higher frequency were included. Considering f_2 variants alone, the DGT reached $> 50\%$ 1-fold coverage, which was higher compared to the FGT with $< 40\%$ 1-fold coverage. Considering focal alleles occurring at

higher frequencies, the DGT was at $\approx 99\%$ 1-fold cumulative coverage at $f_{\leq 20}$, whereas the FGT slowly approached $\approx 95\%$ 1-fold cumulative coverage at $f_{\leq 25}$.

These results indicate that a “global” phasing approach using inferred IBD information alone would not be able to fully determine haplotypes in sample data of unrelated individuals. However, note that similar approaches have been applied to sets of related individuals, for which it can be expected that there are more and longer pairwise shared haplotypes; *i.e.* long range phasing (LRP) methods (Kong *et al.*, 2008; Palin *et al.*, 2011; Loh *et al.*, 2016a).

3.7 Discussion

In this chapter, I presented a novel IBD detection method which is able to infer recombination events in both haplotype and genotype data. To be able to apply this method on a larger scale, I implemented the IBD detection algorithm described in this chapter as a computational tool written in C++; called **tidy** (targeted IBD detection done thoroughly).*

Although the FGT showed overall high levels of accuracy, phasing error was identified as a problem. Current phasing methods such as SHAPEIT2 typically show very low error rates (O’Connell *et al.*, 2014). However, occasionally, alleles are placed on the wrong haplotype. This may happen at single loci (*flip errors*) or such that longer haplotype stretches are exchanged (*switch errors*). Both types of error can affect breakpoint detection under the FGT as both flip and switch errors may change the configuration of alleles observed in relation to a given focal variant. As an alternate solution to using phased haplotypes, the DGT can be used on genotype data, which is therefore not affected by phasing error. However, the lengths of detected IBD segments tend to be overestimated.

Notably, the IBD results obtained from analysis of the 1000G dataset suggests that the FGT was similarly affected by phasing error as seen in the simulation analysis. However, the DGT was also affected by other sources of error. One consideration is that both the FGT

* Targeted IBD detection done thoroughly, tidy: <https://github.com/pkalbers/tidy>

and DGT assume the infinite sites model, but this is only an approximation of conditions observable in nature. In particular, back mutations and recurrent mutations are excluded in the model, but these are prevalent in the (human) genome. For instance, recurrent mutations can produce patterns of variation that would otherwise only be observable if recombination had occurred (McVean *et al.*, 2002). Thus, false positive breakpoints may be inferred, such that IBD length is underestimated. Nonetheless, the infinite sites model is usually seen as a reasonable approximation to reality, as the number of variant sites in a sample is typically much smaller than the number of nucleotides in the chromosomal sequence (Hein *et al.*, 2004). However, the presence of genotype error cannot be ruled out and therefore requires further consideration.

Lastly, the IBD-based phasing concept showed that it is possible to derive the shared allelic state sequence in pairs of individuals that share a haplotype by descent. This approach, in its current implementation, is only able to phase genotype data over locally shared segments. However, a potential utility of such a local phasing approach may arise considering that rare variants can be used to identify recent relatedness patterns in large datasets. Because a given rare allele identifies the underlying shared haplotype, the inferred sequence is likely to correctly distinguish the haplotypes of an individual within a shorter region around the target position. For example, this information may serve as a *post-hoc* correction to haplotype data produced using existing phasing methods.

The key test for an acronym is to ask whether it helps or hurts communication.

— Elon Musk

Abbreviations

1000G	1000 Genomes Project
ARG	Ancestral recombination graph
CDF	Cumulative distribution function
DGT	Discordant genotype test
FGT	Four-gamete test
HapMap	International HapMap Project
HMM	Hidden Markov Model
IBD	Identity by descent
LD	Linkage disequilibrium
LOD	Logarithm of odds
LR	Likelihood ratio
LRP	Long range phasing
Mb	Megabase
MRCA	Most recent common ancestor
RMSLE	Root mean squared logarithmic error
SE	Standard error
SNP	Single-nucleotide polymorphism
T_{MRCA}	Time to the most recent common ancestor
VCF	Variant Call Format

*My definition of a scientist is that you
can complete the following sentence:
'he or she has shown that ...'*

— E. O. Wilson

Bibliography

- 1000 Genomes Project Consortium, Abecasis, G. R., Auton, A., Brooks, L. D., DePristo, M. A., Durbin, R. M., Handsaker, R. E., Kang, H. M., Marth, G. T., and McVean, G. A. (2012). An integrated map of genetic variation from 1,092 human genomes. *Nature*, **491**(7422), 56–65.
- 1000 Genomes Project Consortium, Auton, A., Brooks, L. D., Durbin, R. M., Garrison, E. P., Kang, H. M., Korbel, J. O., Marchini, J. L., McCarthy, S., McVean, G. A., and Abecasis, G. R. (2015). A global reference for human genetic variation. *Nature*, **526**(7571), 68–74.
- Abecasis, G. R., Cherny, S. S., and Cardon, L. R. (2001). The impact of genotyping error on family-based analysis of quantitative traits. *European Journal of Human Genetics*, **9**(2), 130–134.
- Albrechtsen, A., Korneliussen, T. S., Moltke, I., Hansen, T. v. O., Nielsen, F. C., and Nielsen, R. (2009). Relatedness Mapping and Tracts of Relatedness for Genome-Wide Data in the Presence of Linkage Disequilibrium. *Genetic Epidemiology*, **33**(3), 266–274.
- Altshuler, D., Daly, M. J., and Lander, E. S. (2008). Genetic Mapping in Human Disease. *Science*, **322**(5903), 881–888.
- Altshuler, D. L., Bentley, D. R., Chakravarti, A., Collins, F. S., Donnelly, P., Gabriel, S. B., Knoppers, B. M., Lander, E. S., Lehrach, H., Mardis, E. R., Nickerson, D. A., Peltonen, L., Wang, J., Wilson, R. K., Gibbs, R. A., Deiros, D., Metzker, M., Wheeler, D., Li, J., Jian, M., Li, G., Li, R., Liang, H., Tian, G., Wang, B., Wang, W., Yang, H., Zhang, X., Zheng, H., Ambrogio, L., Bloom, T., Sougnez, C. L., Gormley, N., Humphray, S., Kingsbury, Z., Koko-Gonzales, P., Stone, J., McKernan, K. J., Costa, G. L., Ichikawa, J. K., Lee, C. C., Borodina, T. A., Dahl, A., Davydov, A. N., Marquardt, P., Mertes, F., Nietfeld, W., ROSENSTIEL, P., Schreiber, S., Soldatov, A. V., Timmermann, B., Tolzmann, M., Egholm, M., Affourtit, J., Ashworth, D., Attiya, S., Bachorski, M., Buglione, E., Burke, A., Caprio, A., Celone, C., Clark, S., Connors, D., Gu, L., Guccione, L., Kao, K., Keibel, A., Knowlton, J., Labrecque, M., McDade, L., Mealmaker, C., Minderman, M., Nawrocki, A., Niazi, F., Pareja, K., Ramenani, R., Riches, D., Song, W., Turcotte, C., Wang, S., Fulton, L., Weinstock, G., Burton, J., Carter, D. M., Churcher, C., Coffey, A., Cox, A., Quail, M., Skelly, T., Stalker, J., Swerdlow, H. P., Turner, D., De Witte, A., Giles, S., Sabo, A., Yu, J., Fang, X., Guo, X., Li, Y., Luo, R., Tai, S., Wu, H., Zheng, H., Zheng, X., Zhou, Y., Wang, J., Huang, W., Indap, A., Kural, D., Lee, W.-P., Stromberg, M. P., Ward, A. N., Lee, C., Mills, R. E., Shi, X., Daly, M. J., DePristo, M. A., Ball, A. D., Browning, B. L., Cibulskis, K., Garimella, K. V., Grossman, S. R., Hanna, M., Jaffe, D. B., Kernysky, A. M., Li, H., Maguire, J. R., McCarroll, S. A., McKenna, A., Nemesh, J. C., Philippakis, A. A., Poplin, R. E., Price, A., Rivas, M. A., Sabeti, P. C., Schaffner, S. F., Shefler, E., Shlyakhter, I. A., Cooper, D. N., Ball, E. V., Mort, M., Phillips, A. D., Stenson, P. D., Sebat, J., Makarov, V., Ye, K., Yoon, S. C., Clark, A. G., Boyko, A., Degenhardt, J., Gutenkunst, R. N., Kaganovich, M., Keinan, A., Lacroute, P., Ma, X., Reynolds, A., Clarke, L., Cunningham, F., Herrero, J., Keenen, S., Kulesha, E., Leinonen, R., McLaren, W. M., Radhakrishnan, R., Smith, R. E., Zalunin, V., Zheng-Bradley, X., Korbel, J. O., Stütz, A. M., Bauer, M., Keira Cheetham, R., Cox, T., Eberle, M., James, T., Kahn, S., Murray, L., Ye, K., De La Vega, F. M., Fu, Y., Hyland, F. C. L., Manning, J. M., McLaughlin, S. F., Peckham, H. E., Sakarya, O., Sun, Y. A., Tsung, E. F., Batzer, M. A., Konkel, M. K., Walker, J. A., Sudbrak, R., Albrecht, M. W., Amstislavskiy, V. S., Herwig, R., Parkhomchuk, D. V., Agarwala, R., Khouri, H. M., Morgulis, A. O., Paschall, J. E., Phan, L. D., Rotmistrovsky, K. E., Sanders, R. D., Shumway, M. F., Xiao, C., Lunter, G., Marchini, J. L., Moutsianas, L., Myers, S., Tumian, A., Desany, B., Knight, J., Winer, R., Craig, D. W., Beckstrom-Sternberg, S. M., Christoforides, A., Kurdoglu, A. A., Pearson, J. V., Sinari, S. A., Tembe, W. D., Haussler, D., Hinrichs, A. S., Katzman, S. J., Kern, A., Kuhn, R. M., Przeworski, M., Hernandez, R. D., Howie, B., Kelley, J. L., Cord Melton, S., Abecasis, G. R., Li, Y., Anderson, P., Blackwell, T., Chen, W., Cookson, W. O., Ding, J., Min Kang, H., Lathrop, M., Liang, L., and Moffatt, M. F. a. (2010). A map of human genome variation from population-scale sequencing. *Nature*, **467**(7319), 1061–1073.
- Bodmer, W. and Bonilla, C. (2008). Common and rare variants in multifactorial susceptibility to common diseases. *Nature Genetics*, **40**(6), 695–701.
- Boehnke, M. and Cox, N. J. (1997). Accurate Inference of Relationships in Sib-Pair Linkage Studies. *The American Journal of Human Genetics*, **61**(2), 423–429.

- Boomsma, D. I., Wijmenga, C., Slagboom, E. P., Swertz, M. A., Karssen, L. C., Abdellaoui, A., Ye, K., Guryev, V., Vermaat, M., van Dijk, F., Francioli, L. C., Hottenga, J. J., Laros, J. F. J., Li, Q., Li, Y., Cao, H., Chen, R., Du, Y., Li, N., Cao, S., van Setten, J., Menelaou, A., Pulit, S. L., Hehir-Kwa, J. Y., Beekman, M., Elbers, C. C., Byelas, H., de Craen, A. J. M., Deelen, P., Dijkstra, M., den Dunnen, J. T., de Knijff, P., Houwing-Duistermaat, J., Koval, V., Estrada, K., Hofman, A., Kanterakis, A., van Enkevort, D., Mai, H., Kattenberg, M., van Leeuwen, E. M., Neerincx, P. B. T., Oostra, B., Rivadeneira, F., Suchiman, E. H. D., Uitterlinden, A. G., Willemsen, G., Wolffenbuttel, B. H., Wang, J., de Bakker, P. I. W., van Ommen, G.-J., and van Duijn, C. M. (2013). The Genome of the Netherlands: design, and project goals. *European Journal of Human Genetics*, **22**(2), 221–227.
- Borodulin, K., Vartiainen, E., Peltonen, M., Jousilahti, P., Juolevi, A., Laatikainen, T., Mannisto, S., Salomaa, V., Sundvall, J., and Puska, P. (2015). Forty-year trends in cardiovascular risk factors in Finland. *European Journal of Public Health*, **25**(3), 539–546.
- Botstein, D. and Risch, N. (2003). Discovering genotypes underlying human phenotypes: past successes for mendelian disease, future approaches for complex disease. *Nature Genetics*, **33**(3s), 228–237.
- Botstein, D., White, R. L., Skolnick, M., and Davis, R. W. (1980). Construction of a genetic linkage map in man using restriction fragment length polymorphisms. *American journal of human genetics*, **32**(3), 314–331.
- Browning, B. L. and Browning, S. R. (2011a). A fast, powerful method for detecting identity by descent. *American journal of human genetics*, **88**(2), 173–182.
- Browning, B. L. and Browning, S. R. (2013). Improving the accuracy and efficiency of identity-by-descent detection in population data. *Genetics*, **194**(2), 459–471.
- Browning, B. L. and Browning, S. R. (2016). Genotype Imputation with Millions of Reference Samples. *American journal of human genetics*, **98**(1), 116–126.
- Browning, S. R. (2008). Estimation of pairwise identity by descent from dense genetic marker data in a population sample of haplotypes. *Genetics*, **178**(4), 2123–2132.
- Browning, S. R. and Browning, B. L. (2010). High-Resolution Detection of Identity by Descent in Unrelated Individuals. *The American Journal of Human Genetics*, **86**(4), 526–539.
- Browning, S. R. and Browning, B. L. (2011b). Haplotype phasing: existing methods and new developments. **12**(10), 703–714.
- Browning, S. R. and Browning, B. L. (2012). Identity by Descent Between Distant Relatives: Detection and Applications. *Annual Review of Genetics*, **46**(1), 617–633.
- Buetow, K. H. (1991). Influence of aberrant observations on high-resolution linkage analysis outcomes. *The American Journal of Human Genetics*, **49**(5), 985–994.
- Burdett, T., Hall, P., Hasting, E., Hindorff, L., Junkins, H., Klemm, A., MacArthur, J., Manolio, T., Morales, J., Parkinson, H., et al. (2016). The nhgri-ebi catalog of published genome-wide association studies. Available at: www.ebi.ac.uk/gwas. Accessed 2017-01-20, version 1.0.
- Bustamante, C. D., Burchard, E. G., and De La Vega, F. M. (2011). Genomics for the world. *Nature*, **475**(7355), 163–165.
- Cai, L., Fisher, A. L., Huang, H., and Xie, Z. (2016). CRISPR-mediated genome editing and human diseases. *Genes & Diseases*, **3**(4), 244–251.
- Chen, J., Zhang, J.-G., Li, J., Pei, Y.-F., and Deng, H.-W. (2013). On Combining Reference Data to Improve Imputation Accuracy. *PloS one*, **8**(1).
- Choi, Y., Wijsman, E. M., and Weir, B. S. (2009). Case-control association testing in the presence of unknown relationships. *Genetic Epidemiology*, **33**(8), 668–678.
- Cirulli, E. T. and Goldstein, D. B. (2010). Uncovering the roles of rare variants in common disease through whole-genome sequencing. *Nature Publishing Group*, **11**(6), 415–425.
- Colombo, R. (2007). Dating mutations. *eLS*.

- Correns, K. F. J. (1899). Untersuchungen über die Xenien bei *Zea mays*. *Berichte der Deutschen Botanischen Gesellschaft*, **17**, 410–418.
- Cotterman, C. W. (1940). *A calculus for statistico-genetics*. Ph.D. thesis, The Ohio State University.
- Coventry, A., Bull-Otterson, L. M., Liu, X., Clark, A. G., Maxwell, T. J., Crosby, J., Hixson, J. E., Rea, T. J., Muzny, D. M., Lewis, L. R., Wheeler, D. A., Sabo, A., Lusk, C., Weiss, K. G., Akbar, H., Cree, A., Hawes, A. C., Newsham, I., Varghese, R. T., Villasana, D., Gross, S., Joshi, V., Santibanez, J., Morgan, M., Chang, K., Hale IV, W., Templeton, A. R., Boerwinkle, E., Gibbs, R., and Sing, C. F. (2010). Deep resequencing reveals excess rare recent variants consistent with explosive population growth. *Nature communications*, **1**(8), 131–6.
- Cox, D. G. and Kraft, P. (2006). Quantification of the Power of Hardy-Weinberg Equilibrium Testing to Detect Genotyping Error. *Human heredity*, **61**(1), 10–14.
- Crow, J. F. (1954). Breeding structure of populations. ii. effective population number. *Statistics and mathematics in biology*, **543**, 556.
- Daly, M. J., Rioux, J. D., Schaffner, S. F., Hudson, T. J., and Lander, E. S. (2001). High-resolution haplotype structure in the human genome. *Nature Genetics*, **29**(2), 229–232.
- de Bakker, P. I. W., Yelensky, R., Pe'er, I., Gabriel, S. B., Daly, M. J., and Altshuler, D. (2005). Efficiency and power in genetic association studies. *Nature Genetics*, **37**(11), 1217–1223.
- De Vries, H. M. (1900). Sur la loi de disjonction des hybrides. *Comptes rendus de l'Académie des Sciences*, **130**, 845–847.
- Deelen, P., Kanterakis, A., van Dijk, F., Medina-Gomez, C., Kreiner-Møller, E., Rivadeneira, F., Gutierrez-Achury, J., van Enckevort, D., Dijkstra, M., Byelas, H., Genome of Netherlands Consortium, de Bakker, P. I. W., and Swertz, M. A. (2014). Improved imputation quality of low-frequency and rare variants in European samples using the 'Genome of The Netherlands'. *European Journal of Human Genetics*, **22**(11), 1321–1326.
- Delaneau, O., Coulonges, C., and Zagury, J.-F. (2008). Shape-IT: new rapid and accurate algorithm for haplotype inference. **9**(1), 540.
- Delaneau, O., Marchini, J., and Zagury, J.-F. (2011). A linear complexity phasing method for thousands of genomes. *Nature methods*, **9**(2), 179–181.
- Delaneau, O., Zagury, J.-F., and Marchini, J. (2013). Improved whole-chromosome phasing for disease and population genetic studies. *Nature methods*, **10**(1), 5–6.
- Devlin, B., Roeder, K., and Wasserman, L. (2001). Genomic control, a new approach to genetic-based association studies. *Theoretical population biology*, **60**(3), 155–166.
- Donnelly, K. P. (1983). The probability that related individuals share some section of genome identical by descent. *Theoretical population biology*, **23**(1), 34–63.
- Douglas, J. A., Boehnke, M., and Lange, K. (2000). A Multipoint Method for Detecting Genotyping Errors and Mutations in Sibling-Pair Linkage Data. *The American Journal of Human Genetics*, **66**(4), 1287–1297.
- Douglas, J. A., Skol, A. D., and Boehnke, M. (2002). Probability of Detection of Genotyping Errors and Mutations as Inheritance Inconsistencies in Nuclear-Family Data. *The American Journal of Human Genetics*, **70**(2), 487–495.
- Eberle, M. A., Fritzilas, E., Krusche, P., Kallberg, M., Moore, B. L., Bekritsky, M. A., Iqbal, Z., Chuang, H.-Y., Humphray, S. J., Halpern, A. L., Kruglyak, S., Margulies, E. H., McVean, G., and Bentley, D. R. (2016). A reference data set of 5.4 million phased human variants validated by genetic inheritance from sequencing a three-generation 17-member pedigree. *Genome Research*, **27**(1), 1–9.
- Ewens, W. J. (2012). *Mathematical Population Genetics 1: Theoretical Introduction*, volume 27. Springer Science & Business Media.
- Fisher, R. A. (1930). *The genetical theory of natural selection*. Oxford University Press, Oxford.

- Fisher, R. A. (1949). The theory of inbreeding. *The theory of inbreeding*.
- Fisher, R. A. (1954). A fuller theory of “junctions” in inbreeding. *Heredity*, **8**(2), 187–197.
- Forney, G. D. (1973). The Viterbi Algorithm. In *Proceedings of the IEEE*, pages 268–278.
- Frazer, K. A., Murray, S. S., Schork, N. J., and Topol, E. J. (2009). Human genetic variation and its contribution to complex traits. *Nature Reviews Genetics*, **10**(4), 241–251.
- Freedman, M. L., Reich, D., Penney, K. L., McDonald, G. J., Mignault, A. A., Patterson, N., Gabriel, S. B., Topol, E. J., Smoller, J. W., Pato, C. N., Pato, M. T., Petryshen, T. L., Kolonel, L. N., Lander, E. S., Sklar, P., Henderson, B., Hirschhorn, J. N., and Altshuler, D. (2004). Assessing the impact of population stratification on genetic association studies. *Nature Genetics*, **36**(4), 388–393.
- Fu, W., O’Connor, T. D., Jun, G., Kang, H. M., Abecasis, G., Leal, S. M., Gabriel, S., Altshuler, D., Shendure, J., Nickerson, D. A., Bamshad, M. J., Project, N. E. S., and Akey, J. M. (2012). Analysis of 6,515 exomes reveals the recent origin of most human protein-coding variants. *Nature*, **493**(7431), 216–220.
- Fu, Y.-X. and Li, W.-H. (1999). Coalescing into the 21st Century: An Overview and Prospects of Coalescent Theory. *Theoretical population biology*, **56**(1), 1–10.
- Fuchsberger, C., Flannick, J., Teslovich, T. M., Mahajan, A., Agarwala, V., Gaulton, K. J., Ma, C., Fontanillas, P., Moutsianas, L., McCarthy, D. J., Rivas, M. A., Perry, J. R. B., Sim, X., Blackwell, T. W., Robertson, N. R., Rayner, N. W., Cingolani, P., Locke, A. E., Tajcs, J. F., Highland, H. M., Dupuis, J., Chines, P. S., Lindgren, C. M., Hartl, C., Jackson, A. U., Chen, H., Huyghe, J. R., van de Bunt, M., Pearson, R. D., Kumar, A., Müller-Nurasyid, M., Grarup, N., Stringham, H. M., Gamazon, E. R., Lee, J., Chen, Y., Scott, R. A., Below, J. E., Chen, P., Huang, J., Go, M. J., Stitzel, M. L., Pasko, D., Parker, S. C. J., Varga, T. V., Green, T., Beer, N. L., Day-Williams, A. G., Ferreira, T., Fingerlin, T., Horikoshi, M., Hu, C., Huh, I., Ikram, M. K., Kim, B.-J., Kim, Y., Kim, Y. J., Kwon, M.-S., Lee, J., Lee, S., Lin, K.-H., Maxwell, T. J., Nagai, Y., Wang, X., Welch, R. P., Yoon, J., Zhang, W., Barzilai, N., Voight, B. F., Han, B.-G., Jenkinson, C. P., Kuulasmaa, T., Kuusisto, J., Manning, A., Ng, M. C. Y., Palmer, N. D., Balkau, B., áková, A. S., Abboud, H. E., Boeing, H., Giedraitis, V., Prabhakaran, D., Gottesman, O., Scott, J., Carey, J., Kwan, P., Grant, G., Smith, J. D., Neale, B. M., Purcell, S., Butterworth, A. S., Howson, J. M. M., Lee, H. M., Lu, Y., Kwak, S.-H., Zhao, W., Danesh, J., Lam, V. K. L., Park, K. S., Saleheen, D., So, W. Y., Tam, C. H. T., Afzal, U., Aguilar, D., Arya, R., Aung, T., Chan, E., Navarro, C., Cheng, C.-Y., Palli, D., Correa, A., Curran, J. E., Rybin, D., Farook, V. S., Fowler, S. P., Freedman, B. I., Griswold, M., Hale, D. E., Hicks, P. J., Khor, C.-C., Kumar, S., Lehne, B., Thuillier, D., Lim, W. Y., Liu, J., van der Schouw, Y. T., Loh, M., Musani, S. K., Puppala, S., Scott, W. R., Yengo, L., Tan, S.-T., Taylor, H. A., Thameem, F., Wilson, G., Wong, T. Y., Njølstad, P. R., Levy, J. C., Mangino, M., Bonnycastle, L. L., Schwarzmayr, T., Fadista, J., Surdulescu, G. L., Herder, C., Groves, C. J., Wieland, T., Bork-Jensen, J., Brandslund, I., Christensen, C., Koistinen, H. A., Doney, A. S. F., Kinnunen, L., Esko, T., Farmer, A. J., Hakaste, L., Hodgkiss, D., Kravic, J., Lyssenko, V., Hollensted, M., Jørgensen, M. E., Jørgensen, T., Ladenvall, C., Justesen, J. M., Käräjämäki, A., Kriebel, J., Rathmann, W., Lannfelt, L., Lauritzen, T., Narisu, N., Linneberg, A., Melander, O., Milani, L., Neville, M., Orho-Melander, M., Qi, L., Qi, Q., Roden, M., Rolandsson, O., Swift, A., Rosengren, A. H., Stirrups, K., Wood, A. R., Mihailov, E., Blanche, C., Carneiro, M. O., Maguire, J., Poplin, R., Shakir, K., Fennell, T., DePristo, M., de Angelis, M. H., Deloukas, P., Gjesing, A. P., Jun, G., Nilsson, P., Murphy, J., Onofrio, R., Thorand, B., Hansen, T., Meisinger, C., Hu, F. B., Isomaa, B., Karpe, F., Liang, L., Peters, A., Huth, C., O’Rahilly, S. P., Palmer, C. N. A., Pedersen, O., Rauramaa, R., Tuomilehto, J., Salomaa, V., Watanabe, R. M., Syvänen, A.-C., Bergman, R. N., Bharadwaj, D., Bottinger, E. P., Cho, Y. S., Chandak, G. R., Chan, J. C. N., Chia, K. S., Daly, M. J., Ebrahim, S. B., Langenberg, C., Elliott, P., Jablonski, K. A., Lehman, D. M., Jia, W., and Ma, R. (2016). The genetic architecture of type 2 diabetes. *Nature*, **536**(7614), 41–47.
- Genome of the Netherlands Consortium (2014). Whole-genome sequence variation, population structure and demographic history of the Dutch population. *Nature Genetics*, **46**(8), 818–825.
- Gordon, D., Heath, S. C., Liu, X., and Ott, J. (2001). A Transmission/Disequilibrium Test That Allows for Genotyping Errors in the Analysis of Single-Nucleotide Polymorphism Data. *The American Journal of Human Genetics*, **69**(2), 371–380.
- Gordon, D., Finch, S. J., Nothnagel, M., and Ott, J. u. r. (2002). Power and Sample Size Calculations for Case-Control Genetic Association Tests when Errors Are Present: Application to Single Nucleotide Polymorphisms. *Human heredity*, **54**(1), 22–33.

- Gore, A., Li, Z., Fung, H.-L., Young, J. E., Agarwal, S., Antosiewicz-Bourget, J., Canto, I., Giorgetti, A., Israel, M. A., Kiskinis, E., Lee, J.-H., Loh, Y.-H., Manos, P. D., Montserrat, N., Panopoulos, A. D., Ruiz, S., Wilbert, M. L., Yu, J., Kirkness, E. F., Izpisua Belmonte, J. C., Rossi, D. J., Thomson, J. A., Eggan, K., Daley, G. Q., Goldstein, L. S. B., and Zhang, K. (2011). Somatic coding mutations in human induced pluripotent stem cells. *Nature*, **471**(7336), 63–67.
- Griffiths, R. C. (1991). The Two-Locus Ancestral Graph. In *Selected Proceedings of the Sheffield Symposium on Applied Probability*, pages 100–117. Institute of Mathematical Statistics, Hayward, CA.
- Griffiths, R. C. and Marjoram, P. (1996). Ancestral inference from samples of DNA sequences with recombination. *Journal of Computational Biology*, **3**(4), 479–502.
- Griffiths, R. C. and Marjoram, P. (1997a). An ancestral recombination graph. *Institute for Mathematics and its Applications*, **87**, 257.
- Griffiths, R. C. and Marjoram, P. (1997b). Progress in population genetics and human evolution.
- Griffiths, R. C. and Tavaré, S. (1998). The age of a mutation in a general coalescent tree. *Communications in Statistics. Stochastic Models*, **14**(1-2), 273–295.
- Gusev, A., Lowe, J. K., Stoffel, M., Daly, M. J., Altshuler, D., Breslow, J. L., Friedman, J. M., and Pe'er, I. (2009). Whole population, genome-wide mapping of hidden relatedness. *Genome Research*, **19**(2), 318–326.
- Gutenkunst, R. N., Hernandez, R. D., Williamson, S. H., and Bustamante, C. D. (2009). Inferring the Joint Demographic History of Multiple Populations from Multidimensional SNP Frequency Data. *PLoS Genetics*, **5**(10), e1000695–11.
- Haldane, J. (1919). The combination of linkage values and the calculation of distances between the loci of linked factors. *J Genet*, **8**(29), 299–309.
- Hardy, G. H. (1908). Mendelian Proportions in a Mixed Population. *Science*, **28**(706), 49–50.
- Harris, K. and Nielsen, R. (2013). Inferring Demographic History from a Spectrum of Shared Haplotype Lengths. *PLoS Genetics*, **9**(6).
- Hayes, B. J., Visscher, P. M., McPartlan, H. C., and Goddard, M. E. (2003). Novel multilocus measure of linkage disequilibrium to estimate past effective population size. *Genome Research*, **13**(4), 635–643.
- Hein, J., Schierup, M., and Wiuf, C. (2004). *Gene genealogies, variation and evolution: a primer in coalescent theory*. Oxford University Press, Oxford.
- Henn, B. M., Gutenkunst, R. N., Indap, A. R., Marth, G. T., Clark, A. G., Yu, F., Gibbs, R. A., 1000 Genomes Project, Durbin, R. M., Flicek, P., Gabriel, S. B., Lander, E. S., Wheeler, D., Cibulskis, K., Fennell, T. J., Jaffe, D. B., Shefler, E., Egholm, M., Fulton, R., Bainbridge, M., Challis, D., Sabo, A., Wang, J., Lee, C., Korn, J. M., Sudbrak, R., Auton, A., Iqbal, Z., Desany, B., Dooling, D., Hurles, M. E., MacArthur, D. G., Abyzov, A., Zhang, Z., Garrison, E. P., Banks, E., Handsaker, R. E., Hartl, C., De La Vega, F. M., Alkan, C., Snyder, M., Muzny, D., Reid, J., Quinlan, A. R., Stewart, C., Wu, J., Gravel, S., Sherry, S. T., McVean, G. A., Abecasis, G. R., Koboldt, D. C., Palotie, A., Bustamante, C. D., Schafer, A. J., and Brooks, L. D. (2011). Demographic history and rare allele sharing among human populations. *Proceedings of the National Academy of Sciences*, **108**(29), 11983–11988.
- Howie, B., Marchini, J., and Stephens, M. (2011a). Genotype imputation with thousands of genomes. *G3 (Bethesda, Md.)*, **1**(6), 457–470.
- Howie, B., Marchini, J., and Stephens, M. (2011b). Genotype Imputation with Thousands of Genomes. *G3 (Bethesda, Md.)*, **1**(6), 457–470.
- Howie, B. N., Donnelly, P., and Marchini, J. (2009). A Flexible and Accurate Genotype Imputation Method for the Next Generation of Genome-Wide Association Studies. *PLoS Genetics*, **5**(6), e1000529.
- Hudson, R. (1990). Gene genealogies and the coalescent process. *Oxford surveys in evolutionary biology*, **7**(1), 44.
- Hudson, R. and Kaplan, N. L. (1985). Statistical Properties of the Number of Recombination Events in the History of a Sample of Dna-Sequences. *Genetics*, **111**(1), 147–164.

- Hudson, R. R. (1983). Properties of a neutral allele model with intragenic recombination. *Theoretical population biology*, **23**(2), 183–201.
- Hudson, R. R. (2001). Two-locus sampling distributions and their application. *Genetics*, **159**(4), 1805–1817.
- Hudson, R. R. (2002). Generating samples under a Wright-Fisher neutral model of genetic variation. *Bioinformatics*, **18**(2), 337–338.
- International HapMap 3 Consortium, Schaffner, S. F., Peltonen, L., Dermitzakis, E., Bonnen, P. E., Altshuler, D. M., de Bakker, P. I. W., Deloukas, P., Gabriel, S. B., Gwilliam, R., Hunt, S., Inouye, M., Jia, X., Palotie, A., Parkin, M., Whittaker, P., Yu, F., Chang, K., Hawes, A., Lewis, L. R., Ren, Y., Wheeler, D., Gibbs, R. A., Muzny, D. M., Barnes, C., Darvishi, K., Hurles, M., Korn, J. M., Kristiansson, K., Lee, C., McCarroll, S. A., Nemesh, J., Keinan, A., Montgomery, S. B., Pollack, S., Soranzo, N., Gonzaga-Jauregui, C., Anttila, V., Brodeur, W., Daly, M. J., Leslie, S., McVean, G., Moutsianas, L., Nguyen, H., Zhang, Q., Ghorri, M. J. R., McGinnis, R., McLaren, W., Price, A. L., Takeuchi, F., Grossman, S. R., Shlyakhter, I., Hostetter, E. B., Sabeti, P. C., Adebamowo, C. A., Foster, M. W., Gordon, D. R., Licinio, J., Manca, M. C., Marshall, P. A., Matsuda, I., Ngare, D., Wang, V. O., Reddy, D., Rotimi, C. N., Royal, C. D., Sharp, R. R., Zeng, C., Brooks, L. D., and McEwen, J. E. (2010). Integrating common and rare genetic variation in diverse human populations. *Nature*, **467**(7311), 52–58.
- International HapMap Consortium (2003). The International HapMap Project. *Nature*, **426**(6968), 789–796.
- International HapMap Consortium, Frazer, K. A., Ballinger, D. G., Cox, D. R., Hinds, D. A., Stuve, L. L., Gibbs, R. A., Belmont, J. W., Boudreau, A., Hardenbol, P., Leal, S. M., Pasternak, S., Wheeler, D. A., Willis, T. D., Yu, F., Yang, H., Zeng, C., Gao, Y., Hu, H., Hu, W., Li, C., Lin, W., Liu, S., Pan, H., Tang, X., Wang, J., Wang, W., Yu, J., Zhang, B., Zhang, Q., Zhao, H., Zhao, H., Zhou, J., Gabriel, S. B., Barry, R., Blumenstiel, B., Camargo, A., Defelice, M., Faggart, M., Goyette, M., Gupta, S., Moore, J., Nguyen, H., Parkin, M., Roy, J., Stahl, E., Winchester, E., Ziaugra, L., Shen, Y., Yao, Z., Huang, W., Chu, X., He, Y., Jin, L., Liu, Y., Shen, Y., Sun, W., Wang, H., Wang, Y., Wang, Y., Xiong, X., Xu, L., Waye, M. M. Y., Tsui, S. K. W., Xue, H., Wong, J. T.-F., Galver, L. M., Fan, J.-B., Gunderson, K., Murray, S. S., Oliphant, A. R., Chee, M. S., Montpetit, A., Chagnon, F., Ferretti, V., Leboeuf, M., Olivier, J.-F., Phillips, M. S., Roumy, S., Sallée, C., Verner, A., Hudson, T. J., Kwok, P.-Y., Cai, D., Koboldt, D. C., Miller, R. D., Pawlikowska, L., Taillon-Miller, P., Xiao, M., Tsui, L.-C., Mak, W., Song, Y. Q., Tam, P. K. H., Nakamura, Y., Kawaguchi, T., Kitamoto, T., Morizono, T., Nagashima, A., Ohnishi, Y., Sekine, A., Tanaka, T., Tsunoda, T., Deloukas, P., Bird, C. P., Delgado, M., Dermitzakis, E. T., Gwilliam, R., Hunt, S., Morrison, J., Powell, D., Stranger, B. E., Whittaker, P., Bentley, D. R., de Bakker, P. I. W., Barrett, J., Chretien, Y. R., Maller, J., McCarroll, S., Patterson, N., Pe'er, I., Price, A., Purcell, S., Sabeti, P., Saxena, R., Schaffner, S. F., Sham, P. C., Varilly, P., Altshuler, D., Stein, L. D., Krishnan, L., Smith, A. V., Tello-Ruiz, M. K., Thorisson, G. A., Chakravarti, A., Chen, P. E., Cutler, D. J., Kashuk, C. S., Lin, S., Abecasis, G. R., Guan, W., Li, Y., Munro, H. M., Qin, Z. S., Thomas, D. J., McVean, G., Auton, A., Bottolo, L., Cardin, N., Eyheramendy, S., Freeman, C., Marchini, J., Myers, S., Spencer, C., Stephens, M., Donnelly, P., Cardon, L. R., Clarke, G., Evans, D. M., Morris, A. P., Weir, B. S., Mullikin, J. C., Sherry, S. T., Feolo, M., Skol, A., Zhang, H., Matsuda, I., Fukushima, Y., Macer, D. R., Suda, E., Rotimi, C. N., Adebamowo, C. A., Ajayi, I., Aniagwu, T., Marshall, P. A., Nkwodimmah, C., Royal, C. D. M., Leppert, M. F., Dixon, M., Peiffer, A., Qiu, R., Kent, A., Kato, K., Niikawa, N., Adewole, I. F., Knoppers, B. M., Foster, M. W., Clayton, E. W., Watkin, J., Muzny, D., Nazareth, L., Sodergren, E., Weinstock, G. M., Yakub, I., Onofrio, R. C., Richter, D. J., Birren, B. W., Daly, M. J., Wilson, R. K., Fulton, L. L., Rogers, J., Burton, J., Carter, N. P., Clee, C. M., Griffiths, M., Jones, M. C., McLay, K., Plumb, R. W., Ross, M. T., Sims, S. K., Willey, D. L., Chen, Z., Han, H., Kang, L., Godbout, M., Wallenburg, J. C., L'Archevêque, P., Bellemare, G., Saeki, K., Wang, H., An, D., Fu, H., Li, Q., Wang, Z., Wang, R., Holden, A. L., Brooks, L. D., McEwen, J. E., Guyer, M. S., Wang, V. O., Peterson, J. L., Shi, M., Spiegel, J., Sung, L. M., Zacharia, L. F., Collins, F. S., Kennedy, K., Jamieson, R., and Stewart, J. (2007). A second generation human haplotype map of over 3.1 million SNPs. *Nature*, **449**(7164), 851–861.
- International Human Genome Sequencing Consortium (2001). Initial sequencing and analysis of the human genome. *Nature*, **409**(6822), 860–921.
- International Human Genome Sequencing Consortium (2004). Finishing the euchromatic sequence of the human genome. *Nature*, **431**(7011), 931–945.
- Kaiser, J. (2008). A plan to capture human diversity in 1000 genomes (Science (395)). *Science*, **319**(5868), 1336.
- Kang, S. J., Gordon, D., and Finch, S. J. (2004). What SNP genotyping errors are most costly for genetic association studies? *Genetic Epidemiology*, **26**(2), 132–141.

- Keinan, A. and Clark, A. G. (2012). Recent explosive human population growth has resulted in an excess of rare genetic variants. *Science*, **336**(6082), 740–743.
- Kelleher, J., Etheridge, A. M., and McVean, G. (2016). Efficient Coalescent Simulation and Genealogical Analysis for Large Sample Sizes. *PLoS computational biology*, **12**(5), e1004842–22.
- Kimura, M. (1969). The number of heterozygous nucleotide sites maintained in a finite population due to steady flux of mutations. *Genetics*, **61**(4), 893–903.
- Kimura, M. and Ota, T. (1973). The age of a neutral mutant persisting in a finite population. *Genetics*, **75**(1), 199–212.
- Kingman, J. F. C. (1982a). Exchangeability and the evolution of large populations. *Exchangeability in Probability and Statistics*, pages 97—112.
- Kingman, J. F. C. (1982b). On the genealogy of large populations. *Journal of Applied Probability*, **19**(A), 27–43.
- Kingman, J. F. C. (1982c). The coalescent. *Stochastic processes and their applications*, **13**(3), 235–248.
- Klein, R. J., Zeiss, C., Chew, E. Y., Tsai, J.-Y., Sackler, R. S., Haynes, C., Henning, A. K., SanGiovanni, J. P., Mane, S. M., Mayne, S. T., Bracken, M. B., Ferris, F. L., Ott, J., Barnstable, C., and Hoh, J. (2005). Complement Factor H Polymorphism in Age-Related Macular Degeneration. *Science*, **308**(5720), 385–389.
- Kong, A., Masson, G., Frigge, M. L., Gylfason, A., Zusmanovich, P., Thorleifsson, G., Olason, P. I., Ingason, A., Steinberg, S., Rafnar, T., Sulem, P., Mouy, M., Jonsson, F., Thorsteinsdottir, U., Gudbjartsson, D. F., Stefansson, H., and Stefansson, K. (2008). Detection of sharing by descent, long-range phasing and haplotype imputation. *Nature Genetics*, **40**(9), 1068–1075.
- Kruglyak, L. (1999). Prospects for whole-genome linkage disequilibrium mapping of common disease genes. *Nature Genetics*, **22**(2), 139–144.
- Kryukov, G. V., Pennacchio, L. A., and Sunyaev, S. R. (2007). Most rare missense alleles are deleterious in humans: implications for complex disease and association studies. *The American Journal of Human Genetics*, **80**(4), 727–739.
- Lawson, D. J., Hellenthal, G., Myers, S., and Falush, D. (2012). Inference of population structure using dense haplotype data. *PLoS Genetics*, **8**(1), e1002453.
- Leutenegger, A.-L., Prum, B., Génin, E., Verny, C., Lemaître, A., Clerget-Darpoux, F., and Thompson, E. A. (2003). Estimation of the Inbreeding Coefficient through Use of Genomic Data. *American journal of human genetics*, **73**(3), 516–523.
- Li, N. and Stephens, M. (2003). Modeling linkage disequilibrium and identifying recombination hotspots using single-nucleotide polymorphism data. *Genetics*, **165**(4), 2213–2233.
- Li, W. H. (1975). The first arrival time and mean age of a deleterious mutant gene in a finite population. *The American Journal of Human Genetics*, **27**(3), 274–286.
- Li, Y., Willer, C., Sanna, S., and Abecasis, G. (2009). Genotype imputation. *Annual review of genomics and human genetics*, **10**, 387–406.
- Li, Y., Willer, C. J., Ding, J., Scheet, P., and Abecasis, G. R. (2010). MaCH: using sequence and genotype data to estimate haplotypes and unobserved genotypes. *Genetic Epidemiology*, **34**(8), 816–834.
- Lim, E. T., Würtz, P., Havulinna, A. S., Palta, P., Tukiainen, T., Rehnström, K., Esko, T., Mägi, R., Inouye, M., Lappalainen, T., Chan, Y., Salem, R. M., Lek, M., Flannick, J., Sim, X., Manning, A., Ladvall, C., Bumpstead, S., Hämäläinen, E., Aalto, K., Maksimow, M., Salmi, M., Blankenberg, S., Ardisson, D., Shah, S., Horne, B., McPherson, R., Hovingh, G. K., Reilly, M. P., Watkins, H., Goel, A., Farrall, M., Girelli, D., Reiner, A. P., Stitzel, N. O., Kathiresan, S., Gabriel, S., Barrett, J. C., Lehtimäki, T., Laakso, M., Groop, L., Kaprio, J., Perola, M., McCarthy, M. I., Boehnke, M., Altshuler, D. M., Lindgren, C. M., Hirschhorn, J. N., Metspalu, A., Freimer, N. B., Zeller, T., Jalkanen, S., Koskinen, S., Raitakari, O., Durbin, R., MacArthur, D. G., Salomaa, V., Ripatti, S., Daly, M. J., Palotie, A., and for the Sequencing Initiative Suomi (SISu) Project (2014). Distribution and Medical Impact of Loss-of-Function Variants in the Finnish Founder Population. *PLoS Genetics*, **10**(7), e1004494–12.

- Liu, L., Li, Y., Li, S., Hu, N., He, Y., Pong, R., Lin, D., Lu, L., and Law, M. (2012). Comparison of Next-Generation Sequencing Systems. *Journal of biomedicine & biotechnology*, **2012**(7), 1–11.
- Loh, P.-R., Palamara, P. F., and Price, A. L. (2016a). Fast and accurate long-range phasing in a uk biobank cohort. *Nature genetics*.
- Loh, P.-R., Danecek, P., Palamara, P. F., Fuchsberger, C., Reshef, Y. A., Finucane, H. K., Schoenherr, S., Forer, L., McCarthy, S., Abecasis, G. R., Durbin, R., and Price, A. L. (2016b). Reference-based phasing using the Haplotype Reference Consortium panel. Technical report.
- Malécot, G. (1948). Mathematics of heredity. *Les mathématiques de l'hérédité*.
- Manolio, T. A., Collins, F. S., Cox, N. J., Goldstein, D. B., Hindorf, L. A., Hunter, D. J., McCarthy, M. I., Ramos, E. M., Cardon, L. R., Chakravarti, A., Cho, J. H., Guttmacher, A. E., Kong, A., Kruglyak, L., Mardis, E., Rotimi, C. N., Slatkin, M., Valle, D., Whittemore, A. S., Boehnke, M., Clark, A. G., Eichler, E. E., Gibson, G., Haines, J. L., Mackay, T. F. C., McCarroll, S. A., and Visscher, P. M. (2009). Finding the missing heritability of complex diseases. *Nature*, **461**(7265), 747–753.
- Marchini, J. and Howie, B. (2010). Genotype imputation for genome-wide association studies. **11**(7), 499–511.
- Marchini, J., Howie, B., Myers, S., McVean, G., and Donnelly, P. (2007). A new multipoint method for genome-wide association studies by imputation of genotypes. *Nature Genetics*, **39**(7), 906–913.
- Mardis, E. R. (2006). Anticipating the 1,000 dollar genome. *Genome biology*, **7**(7), 112.
- Mardis, E. R. (2017). DNA sequencing technologies: 2006–2016. *Nature Protocols*, **12**(2), 213–218.
- Marth, G. T., Yu, F., Indap, A. R., Garimella, K., Gravel, S., Leong, W. F., Tyler-Smith, C., Bainbridge, M., Blackwell, T., Zheng-Bradley, X., Chen, Y., Challis, D., Clarke, L., Ball, E. V., Cibulskis, K., Cooper, D. N., Fulton, B., Hartl, C., Koboldt, D., Muzny, D., Smith, R., Sougnez, C., Stewart, C., Ward, A., Yu, J., Xue, Y., Altshuler, D., Bustamante, C. D., Clark, A. G., Daly, M., DePristo, M., Flicek, P., Gabriel, S., Mardis, E., Palotie, A., Gibbs, R., and 1000 Genomes Project (2011). The functional spectrum of low-frequency coding variation. *Genome biology*, **12**(9), R84.
- Maruyama, T. (1974). The age of an allele in a finite population. *Genetical research*, **23**(2), 137–143.
- Mathieson, I. and McVean, G. (2012). Differential confounding of rare and common variants in spatially structured populations. *Nature Genetics*, **44**(3), 243–U29.
- Mathieson, I. and McVean, G. (2014). Demography and the Age of Rare Variants. *PLoS Genetics*, **10**(8), e1004528.
- Maynard Smith, J. (1989). *Evolutionary genetics*. Oxford University Press.
- McCarroll, S. A., Kuruvilla, F. G., Korn, J. M., Cawley, S., Nemesh, J., Wysoker, A., Shapero, M. H., de Bakker, P. I. W., Maller, J. B., Kirby, A., Elliott, A. L., Parkin, M., Hubbell, E., Webster, T., Mei, R., Veitch, J., Collins, P. J., Handsaker, R., Lincoln, S., Nizzari, M., Blume, J., Jones, K. W., Rava, R., Daly, M. J., Gabriel, S. B., and Altshuler, D. (2008). Integrated detection and population-genetic analysis of SNPs and copy number variation. *Nature Genetics*, **40**(10), 1166–1174.
- McCarthy, M. I., Abecasis, G. R., Cardon, L. R., Goldstein, D. B., Little, J., Ioannidis, J. P. A., and Hirschhorn, J. N. (2008). Genome-wide association studies for complex traits: consensus, uncertainty and challenges. *Nature Publishing Group*, **9**(5), 356–369.
- McCarthy, S., Das, S., Kretschmar, W., Delaneau, O., Wood, A. R., Teumer, A., Kang, H. M., Fuchsberger, C., Danecek, P., Sharp, K., Luo, Y., Sidore, C., Kwong, A., Timpson, N., Koskinen, S., Vrieze, S., Scott, L. J., Zhang, H., Mahajan, A., Veldink, J., Peters, U., Pato, C., van Duijn, C. M., Gillies, C. E., Gandini, I., Mezzavilla, M., Gilly, A., Cocca, M., Traglia, M., Angius, A., Barrett, J. C., Boomsma, D., Branham, K., Breen, G., Brummett, C. M., Busonero, F., Campbell, H., Chan, A., Chen, S., Chew, E., Collins, F. S., Corbin, L. J., Smith, G. D., Dedoussis, G., Dorr, M., Farmaki, A.-E., Ferrucci, L., Forer, L., Fraser, R. M., Gabriel, S., Levy, S., Groop, L., Harrison, T., Hattersley, A., Holmen, O. L., Hveem, K., Kretzler, M., Lee, J. C., McGue, M., Meitinger, T., Melzer, D., Min, J. L., Mohlke, K. L., Vincent, J. B., Nauck, M., Nickerson, D., Palotie, A., Pato, M., Pirastu, N., McInnis, M., Richards, J. B., Sala, C., Salomaa, V., Schlessinger, D., Schoenherr, S., Slagboom, P. E., Small, K., Spector, T., Stambolian, D., Tuke, M., Tuomilehto, J., Van den Berg, L. H.,

- Van Rheezen, W., Volker, U., Wijmenga, C., Toniolo, D., Zeggini, E., Gasparini, P., Sampson, M. G., Wilson, J. F., Frayling, T., de Bakker, P. I. W., Swertz, M. A., McCarroll, S., Kooperberg, C., Dekker, A., Altshuler, D., Willer, C., Iacono, W., Ripatti, S., Soranzo, N., Walter, K., Swaroop, A., Cucca, F., Anderson, C. A., Myers, R. M., Boehnke, M., McCarthy, M. I., Durbin, R., Abecasis, G., and Marchini, J. (2016). A reference panel of 64,976 haplotypes for genotype imputation. *Nature Genetics*, **48**(10), 1279–1283.
- McClellan, J. and King, M.-C. (2010). Genetic Heterogeneity in Human Disease. *Cell*, **141**(2), 210–217.
- McLaren, W., Gil, L., Hunt, S. E., Riat, H. S., Ritchie, G. R. S., Thormann, A., Flicek, P., and Cunningham, F. (2016). The Ensembl Variant Effect Predictor. *Genome biology*, pages 1–14.
- McQuillan, R., Leutenegger, A.-L., Abdel-Rahman, R., Franklin, C. S., Pericic, M., Barac-Lauc, L., Smolej-Narancic, N., Janicijevic, B., Polasek, O., Tenesa, A., MacLeod, A. K., Farrington, S. M., Rudan, P., Hayward, C., Vitart, V., Rudan, I., Wild, S. H., Dunlop, M. G., Wright, A. F., Campbell, H., and Wilson, J. F. (2008). Runs of Homozygosity in European Populations. *The American Journal of Human Genetics*, **83**(3), 359–372.
- McVean, G., Awadalla, P., and Fearnhead, P. (2002). A coalescent-based method for detecting and estimating recombination from gene sequences. *Genetics*.
- McVean, G. A. T., Myers, S. R., Hunt, S., Deloukas, P., Bentley, D. R., and Donnelly, P. (2004). The fine-scale structure of recombination rate variation in the human genome. *Science*, **304**(5670), 581–584.
- Mendel, G. (1866). Versuche über Pflanzenhybriden. *Verhandlungen des naturforschenden Vereines in Brünn*, **4**, 3–47.
- Metzker, M. L. (2009). Sequencing technologies — the next generation. *Nature Publishing Group*, **11**(1), 31–46.
- Milligan, B. G. (2003). Maximum-Likelihood Estimation of Relatedness. *Genetics*, **163**(3), 1153–1167.
- Moore, G. E. (1965). Cramming more components onto integrated circuits. *Electronics*, **38**(8).
- Morgan, T. H. (1911). Random segregation versus coupling in Mendelian inheritance. *Science*, **34**(873), 384–384.
- Morral, N., Bertranpetit, J., Estivill, X., and Nunes, V. (1994). The origin of the major cystic fibrosis mutation (DF508) in European populations. *Nature*.
- Morris, A. and Cardon, L. (2007). Whole genome association. *Handbook of Statistical Genetics, Third Edition*, pages 1238–1263.
- Moskvina, V. and Schmidt, K. M. (2006). Susceptibility of Biallelic Haplotype and Genotype Frequencies to Genotyping Error. *Biometrics*, **62**(4), 1116–1123.
- Moskvina, V., Craddock, N., Holmans, P., Owen, M., and O'Donovan, M. (2005). Minor genotyping error can result in substantial elevation in type I error rate in haplotype based case control analysis. In *American Journal of Medical Genetics Part B-Neuropsychiatric Genetics*, volume 138, pages 19–19.
- Moutsianas, L., Agarwala, V., Fuchsberger, C., Flannick, J., Rivas, M. A., Gaulton, K. J., Albers, P. K., GoT2D Consortium, McVean, G., Boehnke, M., Altshuler, D., and McCarthy, M. I. (2015). The Power of Gene-Based Rare Variant Methods to Detect Disease-Associated Variation and Test Hypotheses About Complex Disease. *PLoS Genetics*, **11**(4), e1005165–24.
- Myers, S., Bottolo, L., Freeman, C., McVean, G., and Donnelly, P. (2005). A fine-scale map of recombination rates and hotspots across the human genome. *Science*.
- Naidoo, N., Pawitan, Y., Soong, R., Cooper, D. N., and Ku, C.-S. (2011). Human genetics and genomics a decade after the release of the draft sequence of the human genome. *Human Genomics*, **5**(6), 577–622.
- Neuhauser, C. (2001). Mathematical models in population genetics. *Handbook of statistical genetics*.
- Nordborg, M. (2001). Coalescent theory. *Handbook of statistical genetics*.
- O'Connell, J., Gurdasani, D., Delaneau, O., Pirastu, N., Ulivi, S., Cocca, M., Traglia, M., Huang, J., Huffman, J. E., Rudan, I., McQuillan, R., Fraser, R. M., Campbell, H., Polasek, O., Asiki, G., Ekoru, K., Hayward, C., Wright, A. F., Vitart, V., Navarro, P., Zagury, J.-F., Wilson, J. F., Toniolo, D., Gasparini, P., Soranzo, N., Sandhu, M. S., and Marchini, J. (2014). A General Approach for Haplotype Phasing across the Full Spectrum of Relatedness. *PLoS Genetics*, **10**(4), e1004234–21.

- Ott, J. (1999). *Analysis of human genetic linkage*. JHU Press.
- Pajunen, P., Rissanen, H., Härkänen, T., Jula, A., Reunanen, A., and Salomaa, V. (2010). The metabolic syndrome as a predictor of incident diabetes and cardiovascular events in the Health 2000 Study. *Diabetes and Metabolism*, **36**(5), 395–401.
- Palamara, P. F. and Pe'er, I. (2013). Inference of historical migration rates via haplotype sharing. *Bioinformatics*, **29**(13), i180–i188.
- Palamara, P. F., Lencz, T., Darvasi, A., and Pe'er, I. (2012). Length Distributions of Identity by Descent Reveal Fine-Scale Demographic History. *The American Journal of Human Genetics*, **91**(5), 809–822.
- Palin, K., Campbell, H., Wright, A. F., Wilson, J. F., and Durbin, R. (2011). Identity-by-descent-based phasing and imputation in founder populations using graphical models. *Genetic Epidemiology*, **35**(8), 853–860.
- Pasaniuc, B., Zaitlen, N., Shi, H., Bhatia, G., Gusev, A., Pickrell, J., Hirschhorn, J., Strachan, D. P., Patterson, N., and Price, A. L. (2014). Fast and accurate imputation of summary statistics enhances evidence of functional enrichment. *Bioinformatics (Oxford, England)*, **30**(20), 2906–2914.
- Pe'er, I., de Bakker, P. I. W., Maller, J., Yelensky, R., Altshuler, D., and Daly, M. J. (2006). Evaluating and improving power in whole-genome association studies using fixed marker sets. *Nature Genetics*, **38**(6), 663–667.
- Pennisi, E. (2007). Human Genetic Variation. *Science*, **318**(5858), 1842–1843.
- Powell, J. E., Visscher, P. M., and Goddard, M. E. (2010). Reconciling the analysis of IBD and IBS in complex trait studies. *Nature Publishing Group*, **11**(11), 800–805.
- Price, A. L., Patterson, N. J., Plenge, R. M., Weinblatt, M. E., Shadick, N. A., and Reich, D. (2006). Principal components analysis corrects for stratification in genome-wide association studies. *Nature Genetics*, **38**(8), 904–909.
- Pritchard, J. K. (2001). Are rare variants responsible for susceptibility to complex diseases? *The American Journal of Human Genetics*, **69**(1), 124–137.
- Purcell, S., Neale, B., Todd-Brown, K., Thomas, L., Ferreira, M. A. R., Bender, D., Maller, J., Sklar, P., de Bakker, P. I. W., Daly, M. J., and Sham, P. C. (2007). PLINK: a tool set for whole-genome association and population-based linkage analyses. *The American Journal of Human Genetics*, **81**(3), 559–575.
- Rabiner, L. R. (1989). A tutorial on hidden Markov models and selected applications in speech recognition. In *Proceedings of the IEEE*, pages 257–286. Alcatel-Lucent Bell Labs, Murray, United States.
- Rasmussen, M. D., Hubisz, M. J., Gronau, I., and Siepel, A. (2014). Genome-Wide Inference of Ancestral Recombination Graphs. *PLoS Genetics*, **10**(5), e1004342–27.
- Reich, D. E., Cargill, M., Bolk, S., Ireland, J., Sabeti, P. C., Richter, D. J., Lavery, T., Kouyoumjian, R., Farhadian, S. F., Ward, R., and Lander, E. S. (2001). Linkage disequilibrium in the human genome. *Nature*, **411**(6834), 199–204.
- Risch, N. and Merikangas, K. (1996). The future of genetic studies of complex human diseases. *Science*, **273**(5281), 1516–1517.
- Risch, N., de Leon, D., Ozelius, L., Kramer, P., Almas, L., Singer, B., Fahn, S., Breakefield, X., and Bressman, S. (1995). Genetic analysis of idiopathic torsion dystonia in Ashkenazi Jews and their recent descent from a small founder population. *Nature Genetics*, **9**(2), 152–159.
- Risch, N. J. (2000). Searching for genetic determinants in the new millennium. *Nature*, **405**(6788), 847–856.
- Roach, J. C., Glusman, G., Smit, A. F. A., Huff, C. D., Hubley, R., Shannon, P. T., Rowen, L., Pant, K. P., Goodman, N., Bamshad, M., Shendure, J., Drmanac, R., Jorde, L. B., Hood, L., and Galas, D. J. (2010). Analysis of Genetic Inheritance in a Family Quartet by Whole-Genome Sequencing. *Science*, **328**(5978), 636–639.
- Roshyara, N. R. and Scholz, M. (2015). Impact of genetic similarity on imputation accuracy. *BMC Genetics*, **16**(1), 1–16.

- Rousset, F. (2002). Inbreeding and relatedness coefficients: What do they measure? *Heredity*, **88**(5), 371–380.
- Sanger, F. and Coulson, A. R. (1975). A rapid method for determining sequences in DNA by primed synthesis with DNA polymerase. *Journal of molecular biology*, **94**(3), 441–448.
- Sanger, F., Nicklen, S., and Coulson, A. R. (1977). DNA sequencing with chain-terminating inhibitors. *Proceedings of the National Academy of Sciences of the United States of America*, **74**(12), 5463–5467.
- Scally, A. and Durbin, R. (2012). Revising the human mutation rate: implications for understanding human evolution. *Nature Reviews Genetics*, **13**(10), 745–753.
- Schork, N. J., Murray, S. S., Frazer, K. A., and Topol, E. J. (2009). Common vs. rare allele hypotheses for complex diseases. *Current opinion in genetics & development*, **19**(3), 212–219.
- Schroff, M. H. (2016). *Genealogical properties of rare variation and their implications for demographic inference*. Ph.D. thesis, University of Oxford.
- Serre, J. L., Simon-Bouy, B., Mornet, E., Jaume-Roig, B., Balassopoulou, A., Schwartz, M., Taillandier, A., Boue, J., and Boue, A. (1990). Studies of RFLP closely linked to the cystic fibrosis locus throughout Europe lead to new considerations in populations genetics. *Human genetics*, **84**(5), 449–454.
- Shields, D. C., Collins, A., Buetow, K. H., and Morton, N. E. (1991). Error filtration, interference, and the human linkage map. *Proceedings of the National Academy of Sciences of the United States of America*, **88**(15), 6501–6505.
- Simons, Y. B., Turchin, M. C., Pritchard, J. K., and Sella, G. (2014). The deleterious mutation load is insensitive to recent population history. *Nature Publishing Group*, **46**(3), 220–224.
- Slack, J. (2014). *Genes: A Very Short Introduction*. Oxford University Press, Oxford.
- Slatkin, M. (2000). Allele age and a test for selection on rare alleles. *Philosophical Transactions of the Royal Society B: Biological Sciences*, **355**(1403), 1663–1668.
- Slatkin, M. (2008a). Inbreeding coefficients and coalescence times. *doi.org*, pages 1–9.
- Slatkin, M. (2008b). Linkage disequilibrium – understanding the evolutionary past and mapping the medical future. *Nature Reviews Genetics*, **9**(6), 477–485.
- Slatkin, M. and Bertorelle, G. (2001). The use of intraallelic variability for testing neutrality and estimating population growth rate. *Genetics*, **158**(2), 865–874.
- Slatkin, M. and Rannala, B. (2000). Estimating allele age. *Annual review of genomics and human genetics*, **1**(1), 225–249.
- Sobel, E., Papp, J. C., and Lange, K. (2002). Detection and integration of genotyping errors in statistical genetics. *The American Journal of Human Genetics*, **70**(2), 496–508.
- Stam, P. (1980). The distribution of the fraction of the genome identical by descent in finite random mating populations. *Genetical research*, **35**(02), 131.
- Stephens, M., Smith, N. J., and Donnelly, P. (2001). A new statistical method for haplotype reconstruction from population data. *The American Journal of Human Genetics*, **68**(4), 978–989.
- Stone, M. (1961). The Opinion Pool. *The Annals of Mathematical Statistics*, **32**(4), 1339–1342.
- Sturtevant, A. H. (1913). The linear arrangement of six sex-linked factors in drosophila, as shown by their mode of association. *Journal of Experimental Zoology Part A: Ecological Genetics and Physiology*, **14**(1), 43–59.
- Su, Z., Marchini, J., and Donnelly, P. (2011). HAPGEN2: simulation of multiple disease SNPs. *Bioinformatics*, **27**(16), 2304–2305.
- Tajima, F. (1983). Evolutionary relationship of DNA sequences in finite populations. *Genetics*, **105**(2), 437–460.
- Tajima, F. (1993). Measurement of dna polymorphism. *Mechanisms of molecular evolution*, pages 37–59.

- Takahata, N. (1993). Allelic genealogy and human evolution. *Molecular Biology and Evolution*, **10**(1), 2–22.
- Tavaré, S. (2004). Part I: Ancestral inference in population genetics. In *Lectures on probability theory and statistics*, pages 1–188. Springer.
- Tavaré, S., Balding, D. J., Griffiths, R. C., and Donnelly, P. (1997). Inferring coalescence times from DNA sequence data. *Genetics*, **145**(2), 505–518.
- Tennessen, J. A., Bigham, A. W., O'Connor, T. D., Fu, W., Kenny, E. E., Gravel, S., McGee, S., Do, R., Liu, X., Jun, G., Kang, H. M., Jordan, D., Leal, S. M., Gabriel, S., Rieder, M. J., Abecasis, G., Altshuler, D., Nickerson, D. A., Boerwinkle, E., Sunyaev, S., Bustamante, C. D., Bamshad, M. J., Akey, J. M., GO, B., GO, S., and Project, N. E. S. (2012). Evolution and Functional Impact of Rare Coding Variation from Deep Sequencing of Human Exomes. *Science*, **337**(6090), 64–69.
- Thompson, E. A. (1974). Gene Identities and Multiple Relationships. *Biometrics*, **30**(4), 667.
- Thompson, E. A. (1975). The estimation of pairwise relationships. *Annals of human genetics*, **39**(2), 173–188.
- Thompson, E. A. (1976). Estimation of age and rate of increase of rare variants. *The American Journal of Human Genetics*, **28**(5), 442–452.
- Thompson, E. A. (2008). The IBD process along four chromosomes. *Theoretical population biology*, **73**(3), 369–373.
- Thompson, E. A. (2013). Identity by descent: variation in meiosis, across genomes, and in populations. *Genetics*, **194**(2), 301–326.
- Tschermak, E. (1900). Über künstliche Kreuzung bei *Pisum sativum*. *Berichte der Deutschen Botanischen Gesellschaft*, **18**, 232–239.
- UK10K Consortium, Walter, K., Crooks, L., Memari, Y., Perry, J. R. B., Xu, C., Futema, M., Lawson, D., Iotchkova, V., Schiffels, S., Hendricks, A. E., Li, R., Floyd, J., Wain, L. V., Humphries, S. E., Barrett, J. C., Plagnol, V., Richards, J. B., Greenwood, C. M. T., Timpson, N. J., Soranzo, N., Danecsek, P., Barroso, I., McCarthy, S., Tachmazidou, I., Durbin, R., Hurles, M. E., Kennedy, K., Palotie, A., Zeggini, E., Cocca, M., Huang, J., and Min, J. L. (2015). The UK10K project identifies rare variants in health and disease. *Nature*, **526**(7571), 82–90.
- Vartiainen, E., Laatikainen, T., Peltonen, M., Juolevi, A., Mannisto, S., Sundvall, J., Jousilahti, P., Salomaa, V., Valsta, L., and Puska, P. (2010). Thirty-five-year trends in cardiovascular risk factors in Finland. *International Journal of Epidemiology*, **39**(2), 504–518.
- Venter, J. C., Adams, M. D., Myers, E. W., Li, P. W., Mural, R. J., Sutton, G. G., Smith, H. O., Yandell, M., Evans, C. A., Holt, R. A., Gocayne, J. D., Amanatides, P., Ballew, R. M., Huson, D. H., Wortman, J. R., Zhang, Q., Kodira, C. D., Zheng, X. H., Chen, L., Skupski, M., Subramanian, G., Thomas, P. D., Zhang, J., Miklos, G. L. G., Nelson, C., Broder, C., Clark, A. G., Nadeau, J., McKusick, V. A., Zinder, N., Levine, A. J., Roberts, R. J., Simon, M., Slayman, C., Hunkapiller, M., Bolanos, R., Delcher, A., Dew, I., Fasulo, D., Flanigan, M., Florea, L., Halpern, A., Hannenhalli, S., Kravitz, S., Levy, S., Mobarry, C., Reinert, K., Remington, K., Abu-Threideh, J., Beasley, E., Biddick, K., Bonazzi, V., Brandon, R., Cargill, M., Chandramouliswaran, I., Charlab, R., Chaturvedi, K., Deng, Z., Di Francesco, V., Dunn, P., Eilbeck, K., Evangelista, C., Gabrielian, A. E., Gan, W., Ge, W., Gong, F., Gu, Z., Guan, P., Heiman, T. J., Higgins, M. E., Ji, R.-R., Ke, Z., Ketchum, K. A., Lai, Z., Lei, Y., Li, Z., Li, J., Liang, Y., Lin, X., Lu, F., Merkulov, G. V., Milshina, N., Moore, H. M., Naik, A. K., Narayan, V. A., Neelam, B., Nusskern, D., Rusch, D. B., Salzberg, S., Shao, W., Shue, B., Sun, J., Wang, Z. Y., Wang, A., Wang, X., Wang, J., Wei, M.-H., Wides, R., Xiao, C., Yan, C., Yao, A., Ye, J., Zhan, M., Zhang, W., Zhang, H., Zhao, Q., Zheng, L., Zhong, F., Zhong, W., Zhu, S. C., Zhao, S., Gilbert, D., Baumhueter, S., Spier, G., Carter, C., Cravchik, A., Woodage, T., Ali, F., An, H., Awe, A., Baldwin, D., Baden, H., Barnstead, M., Barrow, I., Beeson, K., Busam, D., Carver, A., Angela Center, Cheng, M. L., Curry, L., Danaher, S., Davenport, L., Desilets, R., Dietz, S., Dodson, K., Doup, L., Ferriera, S., Garg, N., Gluecksmann, A., Hart, B., Haynes, J., Haynes, C., Heiner, C., Hladun, S., Hostin, D., Houck, J., Howland, T., Ibegwam, C., Johnson, J., Kalush, F., Kline, L., Koduru, S., Love, A., Mann, F., May, D., McCawley, S., McIntosh, T., McMullen, I., Moy, M., Moy, L., Murphy, B., Nelson, K., Pfannkoch, C., Pratts, E., Puri, V., Qureshi, H., Reardon, M., Rodriguez, R., Rogers, Y.-H., Romblad, D., Ruhfel, B., Scott, R., Sitter, C., Smallwood, M., Stewart, E., Strong, R., Suh, E., Thomas, R., Tint, N. N., Tse, S., Vech, C., Wang, G., Wetter, J., Williams, S., Williams, M., Windsor, S., Winn-Deen, E., Wolfe, K., Zaveri, J., Zaveri, K., Abril, J. F., Guigo, R., Campbell, M. J., Sjolander, K. V., Karlak, B., Kejariwal, A., Mi, H., Lazareva, B., Hatton, T., Narechania, A., Diemer, K.,

- Muruganujan, A., Guo, N., Sato, S., Bafna, V., Istrail, S., Lippert, R., Schwartz, R., Walenz, B., Yooseph, S., Allen, D., Basu, A., Baxendale, J., Blick, L., Caminha, M., Carnes-Stine, J., Caulk, P., Chiang, Y.-H., Coyne, M., Dahlke, C., Mays, A. D., Dombroski, M., Donnelly, M., Ely, D., Esparham, S., Fosler, C., Gire, H., Glanowski, S., Glasser, K., Glodek, A., Gorokhov, M., Graham, K., Gropman, B., Harris, M., Heil, J., Henderson, S., Hoover, J., Jennings, D., Jordan, C., Jordan, J., Kasha, J., Kagan, L., Kraft, C., Levitsky, A., Lewis, M., Liu, X., Lopez, J., Ma, D., and Majoros... (2001). The Sequence of the Human Genome. *Science*, **291**(5507), 1304–1351.
- Viterbi, A. J. (1967). Error Bounds for Convolutional Codes and an Asymptotically Optimum Decoding Algorithm. *IEEE Transactions on Information Theory*, **13**(2), 260–269.
- Voight, B. F. and Pritchard, J. K. (2005). Confounding from Cryptic Relatedness in Case-Control Association Studies. *PLoS Genetics*, **1**(3), e32–10.
- Wakeley, J. (2008). *Coalescent Theory: An Introduction*. W. H. Freeman.
- Wakeley, J. and Wilton, P. (2016). Coalescent and models of identity by descent. In R. M. Kliman, editor, *Encyclopedia of Evolutionary Biology*, pages 287 – 292. Academic Press, Oxford.
- Watson, J. D. and Crick, F. H. (1953). Molecular structure of nucleic acids; a structure for deoxyribose nucleic acid. *Nature*, **171**(4356), 737–738.
- Watterson, G. (1996). Motoo Kimura's Use of Diffusion Theory in Population Genetics. *Theoretical population biology*, **49**(2), 154–188.
- Watterson, G. A. (1975). On the number of segregating sites in genetical models without recombination. *Theoretical population biology*, **7**(2), 256–276.
- Watterson, G. A. (1976). Reversibility and the age of an allele. i. moran's infinitely many neutral alleles model. *Theoretical Population Biology*, **10**(3), 239–253.
- Weinberg, W. (1908). Über den Nachweis der Vererbung beim Menschen. *Jahreshefte des Vereins für vaterländische Naturkunde in Württemberg*, **64**, 368–382.
- Weissenbach, J. (1993). A second generation linkage map of the human genome based on highly informative microsatellite loci. *Gene*, **135**(1-2), 275–278.
- Wheeler, D. A., Srinivasan, M., Egholm, M., Shen, Y., Chen, L., McGuire, A., He, W., Chen, Y.-J., Makhijani, V., Roth, G. T., Gomes, X., Tartaro, K., Niazi, F., Turcotte, C. L., Irzyk, G. P., Lupski, J. R., Chinault, C., Song, X.-z., Liu, Y., Yuan, Y., Nazareth, L., Qin, X., Muzny, D. M., Margulies, M., Weinstock, G. M., Gibbs, R. A., and Rothberg, J. M. (2008). The complete genome of an individual by massively parallel DNA sequencing. *Nature*, **452**(7189), 872–876.
- Winkler, H. (1920). *Verbreitung und Ursache der Parthenogenesis im Pflanzen- und Tierreich*. Verlag G. Fischer, Jena.
- Wiuf, C. and Hein, J. (1997). On the number of ancestors to a DNA sequence. *Genetics*, **147**(3), 1459–1468.
- Wright, S. (1921). Systems of Mating. I. the Biometric Relations between Parent and Offspring. *Genetics*, **6**(2), 111–123.
- Wright, S. (1922). Coefficients of Inbreeding and Relationship. *The American Naturalist*, **56**(645), 330–338.
- Wright, S. (1931). Evolution in Mendelian Populations. *Genetics*, **16**(2), 97–159.
- Yu, N., Zhao, Z., Fu, Y. X., Sambuughin, N., Ramsay, M., Jenkins, T., Leskinen, E., Patthy, L., Jorde, L. B., Kuromori, T., and Li, W. H. (2001). Global patterns of human DNA sequence variation in a 10-kb region on chromosome 1. *Molecular Biology and Evolution*, **18**(2), 214–222.

1. *I have told you more than I know [...].*
2. *What I have told you is subject to change without notice.*
3. *I hope I raised more questions than I have given answers.*
4. *In any case, as usual, a lot more work is necessary.*

– Fuller Albright