

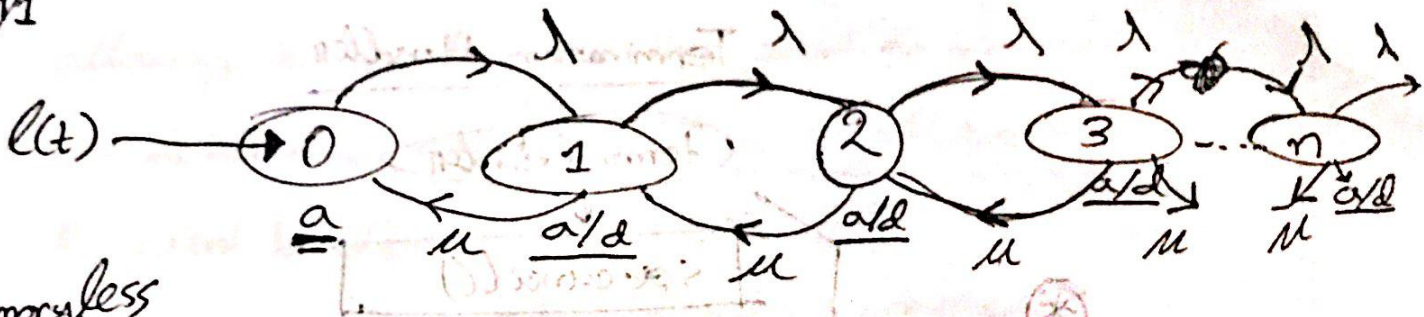
Book.

# Ross: Introduction to Probability Models

Topic → Continuous time Markov Chain

$\lambda \rightarrow$  arrival  
 $\mu \rightarrow$  departure

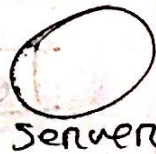
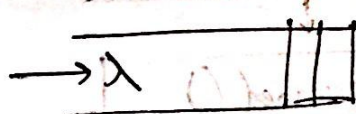
M/M/1



memoryless

State Transition Diagram

Arrival Rate



$\mu$  (Departure Rate)

(Poisson)

Interarrival time  
(exponential)



Interdeparture time  
(exponential time)

The given state transition diagram is a form of birth-death process.

$\lambda_i \rightarrow$  arrival rate for  $i^{\text{th}}$  state

$\mu_i \rightarrow$  departure rate for  $i^{\text{th}}$  state



## Steady State Scenario

$P_i \triangleq$  steady state prob. that system is in state  $i$ .

If we run system for 1 hour and it stays at 0 state for 6 minutes then.

$$P_0 = \frac{6 \text{ min}}{60 \text{ min}} = 10\%$$

(i) Average no. of customers

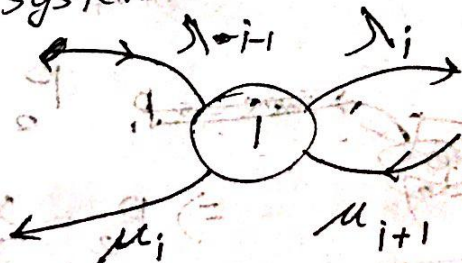
$$P_N(i) = \begin{cases} P_i & i=0,1,2,3 \dots \\ 0 & \text{otherwise} \end{cases}$$

$$E[N] = \sum_{i=0}^{\infty} i \cdot P_i$$

Steady state (\*) thing in terms of that state only)

(i) Rate at which system enters state  $i$ .

(ii) Rate at which system leaves state  $i$ .



For each state,

the rate will be  $P_{i-1} \lambda_{i-1} + P_{i+1} \mu_{i+1}$  for entering the state

and  $P_i \mu_i + P_i \lambda_i$  for leaving the state.

→ As it is a steady state system, these rates will be equal.



## Condition of Stable System

State	Rate of leaving = Rate of entering	
0	$\lambda_0 P_0 = \mu_1 P_1$	(i)
1	$(\mu_1 + \lambda_1) P_1 = \mu_2 P_2 + \lambda_0 P_0$	(ii)
2	$(\mu_2 + \lambda_2) P_2 = \mu_3 P_3 + \lambda_1 P_1$	(iii)
...	...	
n	$(\mu_n + \lambda_n) P_n = \mu_{n+1} P_{n+1} + \lambda_{n-1} P_{n-1}$	(n)

Adding, the equations with its previous ones i.e (i with ii, ii with iii, and so on), we get,

~~$\lambda_0 P_0 + \mu_1 P_1 + \lambda_1 P_1 = \mu_2 P_2 + \lambda_0 P_0 + \mu_1 P_1$~~   
 $\Rightarrow \lambda_1 P_1 = \mu_2 P_2$  (Visudree)  
 Similarly,  $\lambda_2 P_2 = \mu_3 P_3$  and so on.

$\therefore \lambda_n P_n = \mu_{n+1} P_{n+1}$

$\Rightarrow P_1 = \frac{\lambda_0 P_0}{\mu_1}$

~~$\Rightarrow P_2 = \frac{\lambda_1}{\mu_2} \times \frac{\lambda_0}{\mu_1} \times P_0$~~   
 $\Rightarrow P_2 = \frac{\lambda_1}{\mu_2} \times \frac{\lambda_0}{\mu_1} \times P_0$

$\therefore P_n = \frac{\lambda_{n-1}}{\mu_n} P_{n-1} = \frac{\lambda_{n-1} \lambda_{n-2} \dots \lambda_1 \lambda_0}{\mu_n \mu_{n-1} \dots \mu_2 \mu_1} \times P_0$   
 $\sum_{n=0}^{\infty} P_n = 1$

$P_0 + P_0 \sum_{n=1}^{\infty} \frac{\lambda_{n-1} \lambda_{n-2} \dots \lambda_1 \lambda_0}{\mu_n \mu_{n-1} \dots \mu_2 \mu_1} = 1$

$P_0 = \frac{1}{1 + \sum_{n=1}^{\infty} \frac{\lambda_{n-1} \lambda_{n-2} \dots \lambda_1 \lambda_0}{\mu_n \mu_{n-1} \dots \mu_2 \mu_1}} = 1$

This sum can't be  $\infty$

If  $\lambda_i = \lambda$ ,  $\mu_i = \mu$ , we get,

$P_n = \left(\frac{\lambda}{\mu}\right)^n \cdot P_0$

$P_0 = \frac{1}{1 + \sum_{n=1}^{\infty} \left(\frac{\lambda}{\mu}\right)^n}$

(continued)



#9  
Continued

## Analytical Solution of SSR

$$B) \quad P_0 = \frac{1}{1 + \sum_{n=1}^{\infty} \left(\frac{\lambda}{\mu}\right)^n}$$

$$P_n = \left(\frac{\lambda}{\mu}\right)^n \times P_0$$

$$P_0 = \frac{1}{\left(\frac{\lambda}{\mu}\right)^0 + \left[\left(\frac{\lambda}{\mu}\right)^1 + \left(\frac{\lambda}{\mu}\right)^2 + \dots + \left(\frac{\lambda}{\mu}\right)^{\infty}\right]}$$

So, the necessary condition is

$$\frac{\lambda}{\mu} < 1$$

$$\Rightarrow \lambda < \mu$$

$$P_0 = \frac{1}{\frac{1}{1 - \frac{\lambda}{\mu}}} = 1 - \frac{\lambda}{\mu}$$

$$\rho \triangleq \frac{\lambda}{\mu} \rightarrow \text{Traffic intensity}$$

$$\therefore P_0 = 1 - \rho$$

$$(*) \quad P_n = \rho^n (1 - \rho)$$

(\*) Expected value,  $\bar{N} = \sum_{n=0}^{\infty} n \cdot P_n$

$$= \sum_{n=0}^{\infty} n \rho^n (1 - \rho)$$

$$= \frac{\rho}{1 - \rho}$$

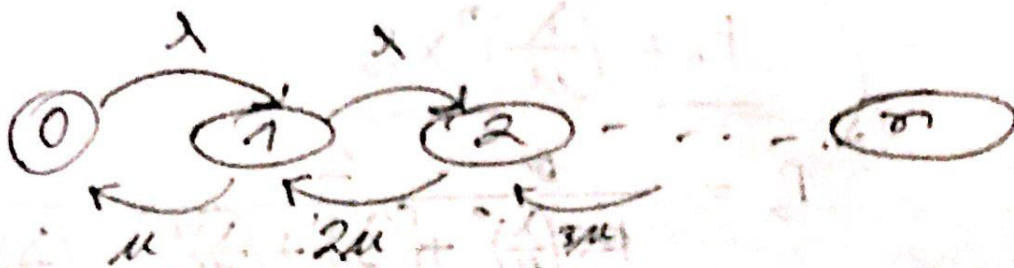
(\*) System Time,  $T$  Using Little's formula  $\lambda/\mu$

$$\lambda = \frac{\bar{N}}{T}$$

$$T = \frac{1}{\lambda} \bar{N} = \frac{1}{\lambda} \cdot \frac{\rho}{1 - \rho} = \frac{1}{\mu(1 - \rho)}$$

Rate

M/M/∞



$$P_n = P_0 \frac{\lambda^n \lambda \dots \lambda}{(n\mu)(n-1)\mu \dots \mu}$$

$$= P_0 \prod_{i=0}^{n-1} \frac{\lambda}{(i+1)\mu}$$

$$P_n = P_0 \left( \frac{\lambda}{\mu} \right)^n \left( \frac{1}{n!} \right)$$

(new term)

$$P_0 = \frac{1}{1 + \sum_{n=1}^{\infty} \frac{\lambda^n}{\mu^n n!}}$$

$$\left( \frac{\lambda^0}{\mu^0 0!} \right) = \frac{1}{\sum_{n=0}^{\infty} \frac{\lambda^n}{\mu^n n!}} = \frac{1}{e^{\lambda/\mu}} = e^{-\lambda/\mu} = e^{-\rho}$$

$$\therefore P_0 = e^{-\rho}$$

$$P_n = e^{-\rho} \frac{\rho^n}{n!}$$



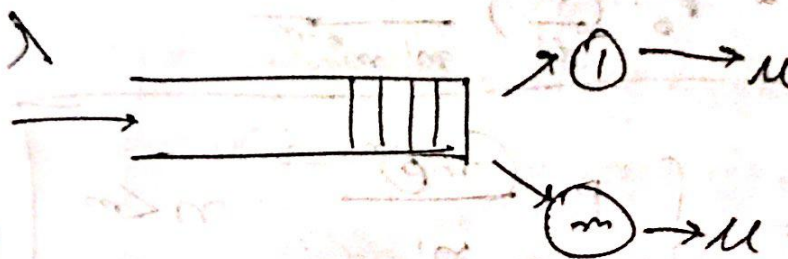
$$P_n = \left(\frac{\lambda}{\mu}\right)^n \cdot \frac{e^{-\lambda\mu}}{n!}$$

Similar to Poisson distribution

$$\bar{N} = \frac{\lambda}{\mu} \text{ for Poisson.}$$

$$\therefore T = \frac{1}{\mu}$$

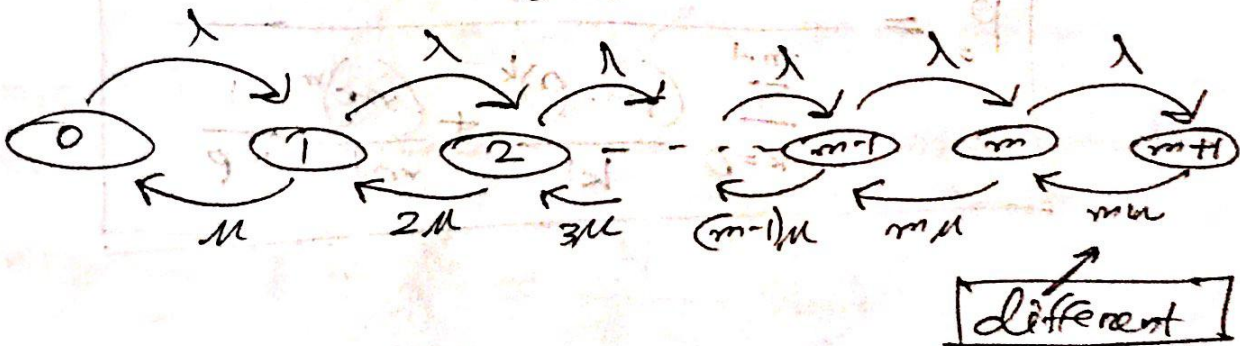
M/M/m → server



$$\lambda_n = \lambda, \quad n = 0, 1, 2, \dots$$

$$\mu_n = \min(n\mu, m\mu) \quad \left( \begin{array}{l} \text{Bottlenecked by} \\ \text{number of} \\ \text{people or servers} \end{array} \right)$$

$$= \begin{cases} n\mu, & n < m \\ m\mu, & n \geq m \end{cases}$$



(i) For  $n < m$ ,  $P_n = P_0 \prod_{i=0}^{n-1} \frac{\lambda}{(i+1)\mu}$

follows

$$(M/M/\infty) = P_0 \left(\frac{\lambda}{\mu}\right)^n \left(\frac{1}{n!}\right)$$

(ii) For  $n \geq m$ ,  $P_n = P_0$

$$P_n = P_0 \left[ \prod_{i=0}^{m-1} \left( \frac{\lambda}{i+1} \right) \frac{1}{\mu n!} \right] \cdot \left[ \prod_{i=m}^n \frac{\lambda}{m! \mu} \right]$$

Infinite Series      Finite Series

$$= P_0 \left( \frac{\lambda}{\mu} \right)^n \cdot \frac{1}{m! m^{n-m}}$$

$$= P_0 \left( \frac{\lambda}{\mu} \right)^n \frac{1}{m! m^{n-m}}$$

$$P_n = \begin{cases} P_0 \frac{(m\rho)^n}{n!} & n < m \\ P_0 \frac{m^m \rho^n}{m!} & n \geq m \end{cases}$$

$$\left[ \rho = \frac{\lambda}{m\mu} < 1 \right]$$

$$P_0 = \frac{1}{\sum_{k=0}^{m-1} \frac{(m\rho)^k}{k!} + \frac{(m\rho)^m}{m!} \cdot \frac{1}{1-\rho}}$$