# 결측(缺測) 데이터와 그 대책

#### Chanseok Park (박찬석)

Applied Statistics Laboratory Department of Industrial Engineering Pusan National University

August 5, 2020

Hosted by SEC



#### Overview

- Missing · Incomplete Data
  - Types of Missing and Incomplete
  - Illustration of Missing Mechanism
  - Ad-hoc methods
  - Ad-hoc methods (deletion)
  - Ad-hoc methods (single imputation)
  - Which method can be used?
- Multiple Imputation
  - MLE
  - EM algorithm
  - MI algorithm

#### Overview

- Missing · Incomplete Data
  - Types of Missing and Incomplete
  - Illustration of Missing Mechanism
  - Ad-hoc methods
  - Ad-hoc methods (deletion)
  - Ad-hoc methods (single imputation)
  - Which method can be used?
- Multiple Imputation
  - MLE
    - EM algorithm
    - MI algorithm

### 1. Missing · Incomplete data

### Types of Missing and Incomplete

- Missing data: no value is observed. (Little and Rubin, 2002)
  - MCAR: Missing Completely AT Random if missingness does not depends on the data,  $Y = (Y_{\text{obs}}, Y_{\text{mis}})$ .  $f_{\theta}(M|Y) = f_{\theta}(M)$ , where M is a missing indicator.
  - MAR: Missing At Random if missingness depends only on the observed data  $Y_{\text{obs}}$ .  $f_{\theta}(M|Y) = f_{\theta}(M|Y_{\text{obs}})$ .
  - MNAR: Missing Not At Random if missingness depends on the data  $Y = (Y_{\text{obs}}, Y_{\text{mis}})$ .  $f_{\theta}(M|Y) = f_{\theta}(M|Y_{\text{obs}}, Y_{\text{mis}})$
- Incomplete data: value is partially observed.
  - Truncation
  - Censoring
  - Grouping
  - Masking

### 1. Missing · Incomplete data: Illustration

### Illustration A: Missing Mechanism (IQ is asked and GPA is known.)

Complete Data		MC	MCAR			MAR			MNAR		
GPA	IQ	GPA	IQ		GPA	IQ		GPA	IQ		
2.0	93	2.0	?	-	2.0	?		2.0	?		
2.2	115	2.2	115		2.2	?		2.2	115		
2.4	96	2.4	96		2.4	?		2.4	?		
2.6	116	2.6	?		2.6	?		2.6	116		
2.8	94	2.8	94		2.8	?		2.8	?		
3.0	106	3.0	106		3.0	106		3.0	106		
3.2	98	3.2	?		3.2	98		3.2	?		
3.4	103	3.4	103		3.4	103		3.4	103		
3.6	95	3.6	95		3.6	95		3.6	?		
3.8	112	3.8	?		3.8	112		3.8	112		
4.0	100	4.0	100		4.0	100		4.0	100		
4.2	120	4.2	?		4.2	120		4.2	120		

### 1. Missing · Incomplete data: Ad-hoc methods

### Deletion – Terminology

- Complete-case analysis (listwise deletion, casewise deletion)
- Available-case analysis (pairwise deletion)

#### Single imputation

- Mean substitution
- Regression imputation
- Hot-deck, Cold-deck

In the next illustration, we assume that GPA, IQ, and Hours of study in a week are asked. We are interested in relation between GPA and Hours.

# 1. Missing · Incomplete data: Ad-hoc methods (deletion)

### Illustration B: Complete and Available cases (GPA, IQ, Hr are asked)

Original Data			Complete case				Available case			
GPA	IQ	Hours	GPA	IQ	Hours		GPA	IQ	Hours	
2.0	93	NA	2.4	96	26	_	2.4	96	26	
2.2	115	NA	2.6	116	28		2.6	116	28	
2.4	96	26	3.8	112	40		3.2	NA	34	
2.6	116	28	4.0	100	42		3.8	112	40	
NA	NA	30	4.2	120	44		4.0	100	42	
NA	NA	32				_	4.2	120	44	
3.2	NA	34					2.0	93	NA	
NA	103	36					2.2	115	NA	
3.6	95	NA					NA	NA	30	
3.8	112	40					NA	NA	32	
4.0	100	42					NA	103	36	
4.2	120	44					3.6	95	NA	
Λ 'Ι Ι					/1.1	_	/ \/ \/	١.	1.\	

Available case depends on an estimate. (Here,  $Cov(X_1,X_3)$  is assumed.)

# 1. Missing · Incomplete data: Ad-hoc methods (deletion)

### Complete case (listwise/casewise deletion)

- MCAR: unbiased.
- MAR: biased.
- Popular in regression data.
- Loss in power. (small sample size).

### Available case (pairwise deletion)

- MCAR: biased.
- MAR: biased.
- Correlation estimate (small value case is OK).
  - Refer to Talk-R.r at 2020/Talk-R.
  - For available case, the correlation can be **even greater than 1.0.** (note: complete case (**MCAR**) is OK).

Illustration for available case continues on the next page.

# 1. Missing · Incomplete data: Example (Available Case)

mean/sd	mean/sd	Covariance between $X_1 \& X_3$					
$GPA\ (X_1)$	Hours $(X_3)$	$GPA\left(X_{1}\right)$	Hours $(X_3)$				
2.0	26	2.4	26				
2.2	28	2.6	28				
2.4	30	3.2	34				
2.6	32	3.8	40				
3.2	34	4.0	42				
3.6	36	4.2	44				
3.8	40	$Cov(X_1)$	$(X_3) = 5.67$				
4.0	42						
4.2	44						
$\bar{X}_1 = 3.11$	$\bar{X}_3 = 34.67$						
$S_1 = 0.83$	$S_3 = 6.32$						

Thus, 
$$r_{13} = \frac{\text{Cov}(X_1, X_3)}{S_1 \cdot S_3} = \frac{5.67}{0.83 \times 6.32} = 1.08 > 1$$
, which does not make sense at all. Refer to Talk-R.r at  $^{2020/\text{Talk-R}}$ 

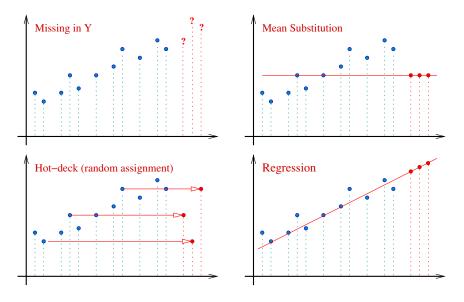
Chanseok Park (Applied Stat Lab. IE PNU)

### 1. Missing · Incomplete data: Ad-hoc (single imputation)

### Single imputation

- Mean substitution (location estimate)
   Median, mode, HL, etc. are also OK instead of mean.
   Distorted variance/covariance. Easily biased.
- Regression imputation
   Distorted variance/covariance
   . can be unbiased.
   A random error can be added to avoid distortion of variance/covariance. Note: if a dummy variable is used for a predictor, it can include mean substitution.
- Hot-deck / Cold-deck (similar to bootstrap/jackknife).
   Hot-deck: random sample from similar hot responding values. (here, "hot" means current source.)
   Cold-deck: random sample from similar cold responding values. (here, "cold" means previous/external source.)
   Both can be easily biased.

# 1. Missing · Incomplete data: Ad-hoc (single imputation)



### 1. Missing · Incomplete data: Which method can be used?

### Recall "Complete case" versus "Available case"

### Complete case (listwise/casewise deletion)

- MCAR: unbiased ← It looks OK although wasteful with deletion.
- MAR: biased.

### Available case (pairwise deletion)

- MCAR: biased.
- MAR: biased.

It is possible to test MCAR vs. MAR, but impossible to test MNAR.

Section 2.2.4 of van Buuren (2018) stated: several tests have been proposed to test MCAR vs. MAR. These tests are not widely used, and their practical value is unclear. ... It is **not** possible to test MAR versus MNAR since the information that is needed for such a test is missing.

### 1. Missing · Incomplete data: Which method can be used?

#### Which method can be used?

- Listwise deletion (complete case) is OK (in a sense of unbiasedness) under MCAR although it is wasteful.
  - Note: listwise deletion is **not** robust to violations of the MCAR assumption. Also, the MCAR (very strong condition) is often **unrealistic** in practice.
- There are several ad-hoc methods under MAR, but these are not robust to violation of MAR assumption.
- We can think of imputation.
   However single imputation also has several drawbacks.

Multiple Imputation (MI) can handle both MAR and MNAR. Some research papers show that listwise deletion method can outperform the MI method, but it is extremely rare in practice.

# 2. Multiple Imputation: MLE

Multiple imputation has a very similar mechanism as EM algorithm. We look at MLE and EM first, and then multiple imputation later.

### Likelihood and log-likelihood

$$L(\boldsymbol{\theta}|\mathbf{x}) = \prod_{i=1}^n f(x_i)$$
 and  $\ell(\boldsymbol{\theta}|\mathbf{x}) = \log L(\boldsymbol{\theta}|\mathbf{x}) = \sum_{i=1}^n \log f(x_i)$ ,

where  $\mathbf{x} = (x_1, x_2, \dots, x_n)$  and  $\boldsymbol{\theta} = (\theta_1, \theta_2, \dots, \theta_p)$ .

### MLE (maximum likelihood estimate/estimator)

$$\hat{\pmb{\theta}} = \arg\max_{\pmb{\theta}} L(\pmb{\theta}|\mathbf{x}) \ \ \mathrm{or} \ \ \hat{\pmb{\theta}} = \arg\max_{\pmb{\theta}} \ell(\pmb{\theta}|\mathbf{x}).$$

In many practical cases, the MLE is obtained in a closed form.



We hope to have full observations,  $\mathbf{x} = (x_1, x_2, \dots, x_m, x_{m+1}, \dots, x_n)$ , but  $\mathbf{x}$  can be split into complete observations  $(\mathbf{y})$  and incomplete observations  $(\mathbf{z})$  so that  $\mathbf{x} = (\mathbf{y}, \mathbf{z})$ .

That is, what if  $\mathbf{y} = (x_1, x_2, \dots, x_m)$  are complete and  $\mathbf{z} = (x_{m+1}, x_{m+2}, \dots, x_n)$  are incomplete. Say,  $a_j \le x_j \le b_j$ , where  $j = m+1, m+2, \dots, n$ . Then we have

$$L(\theta|\mathbf{y}, \mathbf{z}) = \prod_{i=1}^{m} f(x_i) \prod_{j=m+1}^{n} \{F(b_j) - F(a_j)\}$$

$$\ell(\theta|\mathbf{y}, \mathbf{z}) = \sum_{i=1}^{m} \log f(x_i) + \sum_{j=m+1}^{n} \log \{F(b_j) - F(a_j)\}.$$

In general, the MLE can not be obtained in a closed form.

Treat incomplete part as random variable with an appropriate distribution. In this case, we can set up  $\mathbf{z} = (z_{m+1}, z_{m+2}, \dots, z_n)$  where  $z_j$  has the pdf  $p(\mathbf{z}|\mathbf{y}, \boldsymbol{\theta})$ . Then the complete likelihood is given by

$$L^{c}(\boldsymbol{\theta}|\mathbf{y},\mathbf{z}) = \prod_{i=1}^{m} f(x_i) \prod_{j=m+1}^{n} \underline{f(z_j)},$$

where  $\ell^c(\boldsymbol{\theta}|\mathbf{y},\mathbf{z}) = \log L^c(\boldsymbol{\theta}|\mathbf{y},\mathbf{z})$ ,  $z_j$  has a distribution whose value is between  $a_j < z_j < b_j$  (we need to take an expectation w.r.t.  $z_j$  since  $z_j$  is like an random variable).

**Recall**: compare the above complete likelihood with the previous original likelihood function:

$$L(\boldsymbol{\theta}|\mathbf{y},\mathbf{z}) = \prod_{i=1}^{m} f(x_i) \prod_{j=m+1}^{n} \left[ \frac{\{F(b_j) - F(a_j)\}}{\{F(b_j) - F(a_j)\}} \right]$$

Start with an initial value (say, t=0). Repeat E-step and M-step below. (Note: we take an expectation w.r.t.  $z_j$ ).

#### **EM Algorithm**

- E-step:  $Q(\theta|\theta^{(t)}) = \int \log L^c(\theta|\mathbf{y}, \mathbf{z}) p(\mathbf{z}|\mathbf{y}, \theta^{(t)}) d\mathbf{z}$
- M-step:  $m{ heta}^{(t+1)} = rg \max_{m{ heta}} Q(m{ heta}|m{ heta}^{(t)})$
- In the M-step, we need to obtain the closed-form maximizer.
   (in general, using EM, this M-step helps one to obtain the closed-form).
- This iterative method provides clear benefit over the Newton-Raphson method. (If the likelihood is unimodal, finding the MLE is guaranteed).
- The issue is how to obtain the closed-form maximizer in the M-step. To relax this, **MC-EM** and **Q-EM** are developed.

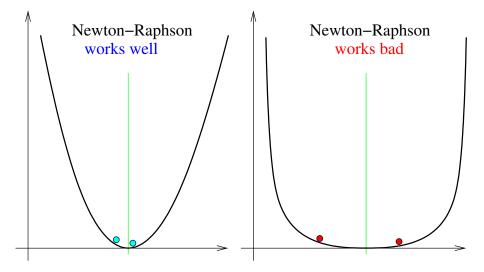
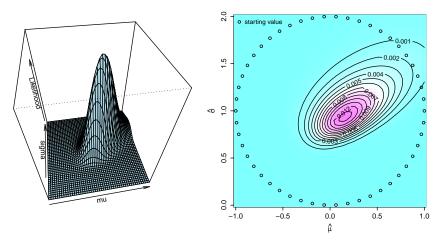
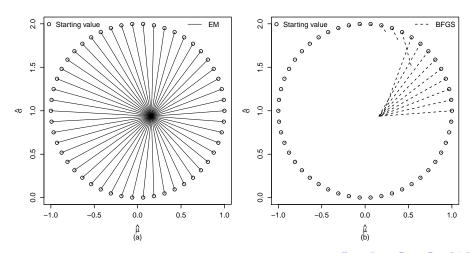


Illustration: Likelihood function with two parameters. Lognormal example from (Park, 2013).



# Illustration: Convergence of BFGS (improved Newton-Raphson) method (Park, 2013)



MI is similar to EM which solves an incomplete-data problem by repeatedly solving the complete-data version. In MI, the unknown missing data  $Y_{\rm mis}$  are replaced by simulated values  $Y_{\rm mis}^{(1)}, Y_{\rm mis}^{(2)}, \ldots, Y_{\rm mis}^{(m)}$ .

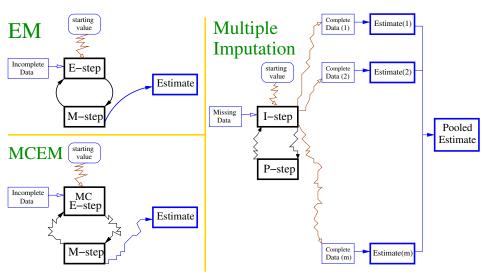
### Multiple Imputation (MI) Algorithm

- I-step:  $Y_{\mathrm{mis}}^{(t+1)} \sim p(Y_{\mathrm{mis}}|Y_{\mathrm{obs}}, \boldsymbol{\theta}^{(t)})$
- P-step:  $oldsymbol{ heta}^{(t+1)} \sim p(oldsymbol{ heta}|Y_{
  m obs},Y_{
  m mis}^{(t+1)})$

Thus, we obtain m completed data sets. With each of data sets, we analyze it by a **standard method** with a completed data set. We will have m different results. By pooling them (summarizing them), we can obtain a result along with the uncertainty due to missing.

#### R Packages for MI

- NORM Package: https://cran.r-project.org/web/packages/norm
- MICE Package by van Buuren and Groothuis-Oudshoorn (2011) https://cran.r-project.org/web/packages/mice



### 2. Multiple Imputation: Summary

#### Recall

- MCAR: unbiased with listwise deletion (but, wasteful).
- MAR: biased with either listwise or pairwise deletion.
- MNAR: biased. Impossible to test MNAR.

#### When does MI work well?

- MAR with <u>distinctness</u>: <u>unbiased</u>. The parameters  $\theta$  and  $\psi$  are <u>distinct</u> if  $g(\theta, \psi) = g_1(\theta) \cdot g_2(\psi)$ . See Definition 6.4 of Little and Rubin (2002).
- MI method is very robust to MNAR.
   See Section 6.2 of van Buuren (2018).

NOTE: with MI, we can generate complete-data sets (not parameter estimate).

#### References

- Little, R. J. A. and Rubin, D. B. (2002). Statistical Analysis with Missing Data. John Wiley & Sons, New York, 2nd edition.
- Park, C. (2013). Parameter estimation from load-sharing system data using the expectation-maximization algorithm. <a href="IETransactions">IETransactions</a>, 45:147–163.
- Park, C. (2018). A quantile variant of the Expectation-Maximization algorithm and its application to parameter estimation with interval data. Journal of Algorithms & Computational Technology, 12:253–272.
- Schafer, J. L. (1997). <u>Analysis of Incomplete Multivariate Data</u>. Chapman & Hall, Boca Raton, FL.
- van Buuren, S. (2018). <u>Flexible Imputation of Missing Data</u>. Chapman & Hall/CRC, Boca Raton, second edition.
- van Buuren, S. and Groothuis-Oudshoorn, K. (2011). mice: Multivariate imputation by chained equations in R. <u>Journal of Statistical Software</u>, 45(3):1–67.

#### References

- Wei, G. C. G. and Tanner, M. A. (1990a). A Monte Carlo implementation of the EM algorithm and the poor man's data augmentation algorithm. Journal of the American Statistical Association, 85:699–704.
- Wei, G. C. G. and Tanner, M. A. (1990b). Posterior computations for censored regression data. <u>Journal of the American Statistical</u> Association, 85:829–839.