

Bayesian Geostatistics

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Bayesian spatial random effect models

- Continuous data:

$$y(s_i) | \mu(s_i), \tau^2 \stackrel{\text{ind}}{\sim} N(\mu(s_i), \tau^2) ; \quad i = 1, 2, \dots, n ;$$

$$\mu(s_i) = \beta_0 + \beta_1 x_1(s_i) + \beta_2 x_2(s_i) + \dots + \beta_p x_p(s_i) + w(s_i) ;$$

$$\beta_j \stackrel{\text{ind}}{\sim} N(0, \sigma_\beta^2) ; \quad j = 0, 1, \dots, p ;$$

$$w = (w(s_1), w(s_2), \dots, w(s_n))^T \sim N(0, \sigma^2 R_w(\phi)) ;$$

$$1/\tau^2 \sim \text{Gamma}(a_\tau, b_\tau) ; \quad 1/\sigma^2 \sim \text{Gamma}(a_\sigma, b_\sigma) ;$$

$$\phi \sim \text{Unif}(a_\phi, b_\phi) .$$

- $R_w(\phi)$ is $n \times n$ spatial correlation matrix.

Bayesian spatial random effect models

- Count data:

$$y(s_i) \sim \text{Poi}(\lambda(s_i)) ; \quad i = 1, 2, \dots, n ;$$

$$\log \lambda(s_i) = \beta_0 + \beta_1 x_1(s_i) + \beta_2 x_2(s_i) + \dots + \beta_p x_p(s_i) + w(s_i) ;$$

$$\beta_j \overset{\text{ind}}{\sim} N(0, \sigma_\beta^2) ; \quad j = 0, 1, \dots, p ; \quad w \sim N(0, \sigma^2 R_w) ;$$

$$1/\sigma^2 \sim \text{Gamma}(a_\sigma, b_\sigma) ; \quad \phi \sim \text{Unif}(a_\phi, b_\phi) .$$

- $R_w(\phi)$ is $n \times n$ spatial correlation matrix.

Bayesian spatial random effect models

- Binary data:

$$y(s_i) \sim \text{Ber}(p(s_i)) ; \quad i = 1, 2, \dots, n ;$$

$$\log \left(\frac{p(s_i)}{1 - p(s_i)} \right) = \beta_0 + \beta_1 x_1(s_i) + \beta_2 x_2(s_i) + \dots + \beta_p x_p(s_i) + w(s_i) ;$$

$$\beta_j \stackrel{\text{ind}}{\sim} N(0, \sigma_\beta^2) ; \quad j = 0, 1, \dots, p ; \quad w \sim N(0, \sigma^2 R_w) ;$$

$$1/\sigma^2 \sim \text{Gamma}(a_\sigma, b_\sigma) ; \quad \phi \sim \text{Unif}(a_\phi, b_\phi) .$$

- $R_w(\phi)$ is $n \times n$ spatial correlation matrix.

Spatial random effects: Gaussian process

- We say that $w(s) \sim GP(0, \sigma^2 \rho(\cdot))$:

$$w = (w(s_1), w(s_2), \dots, w(s_n))^{\top} \sim N(0, \sigma^2 R_w) ;$$

- R_w is $n \times n$ spatial correlation matrix:

$$R_w[i, j] = \rho(s_i, s_j) .$$

- The correlation function is parametrized to capture strength of association as a function of distance. Practical choice (works well for a variety of situations):

$$\rho(s_i, s_j) = \exp(-\phi \|s_i - s_j\|) .$$

Bayesian inference for continuous spatial data

- Step-I: Estimate parameters (MCMC) by sampling from

$$[\beta, w, \tau^2, \sigma^2, \phi \mid y, X]$$

- Step-II: Estimate the latent process $w(s_0)$ at new location s_0 by sampling from

$$[w(s_0) \mid w, \sigma^2, \phi]$$

for each sampled value of w , σ^2 and ϕ obtained in Step-I.

- Step III: Obtain posterior samples of $\mu(s_0)$:

$$\mu(s_0) = \beta_0 + \beta_1 x_1(s_0) + \beta_2 x_2(s_0) + \cdots + \beta_p x_p(s_0) + w(s_0) .$$

- Step-IV: Predict $y(s_0)$ by drawing its value from

$$N(\mu(s_0), \tau^2)$$

for each sampled $\mu(s_0)$ (from Step-III) and τ^2 (from Step-I).

Bayesian inference for spatial count data

- Step-I: Estimate parameters (MCMC) by sampling from

$$[\beta, w, \sigma^2, \phi | y, X]$$

- Step-II: Estimate the latent process $w(s_0)$ at new location s_0 by drawing from

$$[w(s_0) | w, \sigma^2, \phi]$$

for each sampled value of w , σ^2 and ϕ obtained in Step-I.

- Step III: Obtain posterior samples of $\lambda(s_0)$:

$$\lambda(s_0) = \exp(\beta_0 + \beta_1 x_1(s_0) + \beta_2 x_2(s_0) + \cdots + \beta_p x_p(s_0) + w(s_0)) .$$

- Step-IV: Predict $y(s_0)$ by drawing its value from

$$Poi(\lambda(s_0))$$

for each sampled $\lambda(s_0)$ in Step-III.

Bayesian inference for binary count data

- Step-I: Estimate parameters (MCMC) by sampling from

$$[\beta, w, \sigma^2, \phi | y, X]$$

- Step-II: Estimate the latent process $w(s_0)$ at new location s_0 by drawing from

$$[w(s_0) | w, \sigma^2, \phi]$$

for each sampled value of w , σ^2 and ϕ obtained in Step-I.

- Step III: Obtain posterior samples of $p(s_0)$:

$$p(s_0) = \text{logit}^{-1}(\beta_0 + \beta_1 x_1(s_0) + \beta_2 x_2(s_0) + \cdots + \beta_p x_p(s_0) + w(s_0)) .$$

- Step-IV: Predict $y(s_0)$ by drawing its value from

$$\text{Ber}(p(s_0))$$

for each sampled $p(s_0)$ in Step-III.

Setting Priors

- For regression slopes we customarily assign non-informative priors.
- For the variance component σ^2 (partial sill), we customarily choose an Inverse-Gamma (or, equivalently, Gamma prior for $1/\sigma^2$)—the shape parameter is taken to be 2 and the scale parameter is chosen so that the prior mean is equal to the scale parameter. This value can be set from an exploratory variogram analysis. Strategy for τ^2 (nugget) is similar.
- For the range parameter ϕ , we usually set it so that the effective range (distance where spatial correlation drops to 0.05) is between some small number and does not exceed about 50% of the maximum inter-site distance. For example, with the exponential correlation function we solve $\rho(d; \phi) = 0.05$ and see that $\phi \approx 3/d$, where d is the effective spatial range. We bound $d \in (d_{\min}, d_{\max})$ and this suggests $\phi \sim \text{Unif}(3/d_{\max}, 3/d_{\min})$.