Statistical Learning Project: Analysis of Boston house price

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1 Introduction

Boston house price is one of the most famous statistic problems in the last century. So we try to analyze the house price through classical statistical methods and find out the major and minor factors that affects the price. According to our goal, linear regression is the best choice.

We collected enough data and did some preprocessing: we randomly selected 2/3 as training data and set the remaining as test data. The training data is used to train the linear model and the test data is used to assess the model. For problems with multiple variables such as Boston house price, there are several typical ways to optimize the model. We tried these optimization methods and found out the optimum fitting of the problem.

1.1 Datasets

Our data come from an old essay about Boston house price, *Hedonic prices and the demand for clean air* .

We randomly divided the data into two sets:

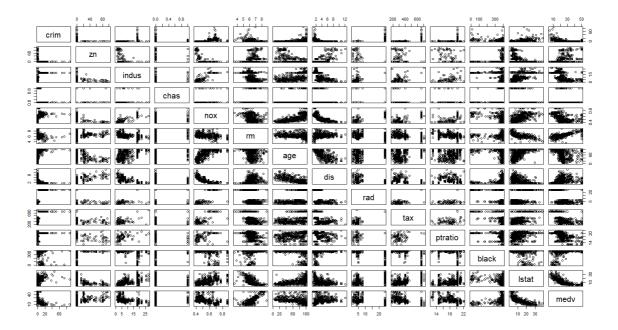
- Training set: randomly chosen 2/3 of origin sets,
- Test set: randomly chosen 1/3 of origin sets.

And we will use test setting to assess our model. Besides, the variable of "chas" is already dummied.

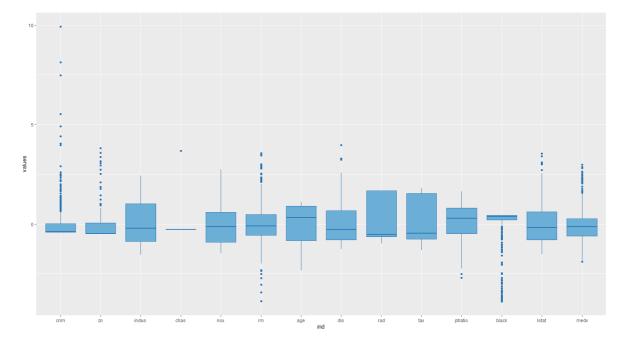
Variable	Definition	Source		
medv	Median Value of owner-occupied	1970 U. S. Census		
rm	Average number of rooms per dwelling.	1970 U. S. Census		
age	Proportion of owner-occupied units built prior to 1940.	1970 U. S. Census		
black	1000(Bk - 0.63)^2 where Bk is the proportion of blacks by own.	1970 U. S. Census		
Istat	Lower status of the population(percent).	1970 U. S. Census		
crim	Per crime rate by town.	FBI (1970)		
zn	Proportion of a town's residential land zoned fro lots greater that 25,000 square feet.	Metropolitan Area Planning Commission (1972)		
indus	Proportion non-retail business acres per town.	Vogt, Ivers, and Association		
tax	Full value property tax rate (\$/\$10,000).	Massachusetts Tax- payers Foundation(1970)		
ptratio	Pupil-teacher ratio by town school district	Massachusetts Dept. of Education (1971- 1972)		
chas	Charles River dummy := 1 if tract bounds the Charles River; 0 if otherwise.	1970 U. S. Census Tract maps		
dis	Weighted distances to five employment centers in the Boston region.	Schnare		
rad	Index of accessibility to radical highways.	MIT Boston Project		
nox	Nitrogen oxides concentration in pphm (annual aver-age concentration in parts per hundred million).	TASSIM		

1.2 Data Visualization

We drew a panel of scatterplot for a pair of variables whose identities are given by the corresponding row and column labels.



The box graph of all variables was shown to illustrate the data distribution.



1.3 Data Standardizing

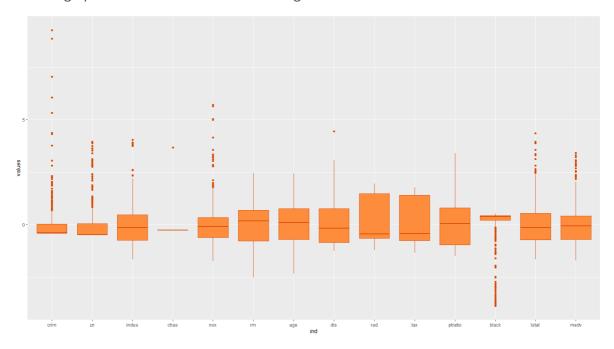
In the Lasso model, we need to use the formula below to standardizing the data:

$$\tilde{x}_{i,j} = \frac{x_{i,j}}{\sqrt{\frac{1}{n} \sum_{i=1}^{n} (\frac{x_{i,j}}{x_{i,j} - \bar{x}_{j}})^{2}}}$$
(8)

After standardization, the head of data is shown below:

zn	indus	chas	nox	rm	age	dis	rad	tax	ptratio	black	Istat	medv
0.808794	0.107039	0	0.024217	0.301427	3.285297	0.185721	0.04567	6.304818	0.727945	7.082784	0.231909	1.148223
0	0.34765	0	0.022368	0.312204	4.19072	0.239091	0.096855	5.145127	0.899769	7.060669	0.451788	1.09814
0	0.351682	0	0.022624	0.353389	3.272702	0.241842	0.097973	5.136803	0.910231	6.976837	0.201515	1.784631
0	0.110867	0	0.022595	0.351951	2.479367	0.301846	0.150278	4.680136	0.977038	6.96521	0.150288	1.754769
0	0.110153	0	0.022443	0.357083	2.930941	0.299832	0.149284	4.696069	0.971178	7.026242	0.270726	1.890694

The box graph of those data after standardizing is shown below:



1.3 Model assess

We use some statistical concept to assess our model:

1. Residual Standard Error

$$RSE = \sqrt{\frac{RSS}{n-2}} \tag{1}$$

Where RSS(residual sum of squares) is:

$$RSS = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2 \tag{2}$$

2. R² Statistic

$$R^2 = \frac{TSS - RSS}{TSS} = 1 - \frac{RSS}{TSS} \tag{3}$$

Where TSS(total sum of squares) is:

$$TSS = \sum_{i=1}^{n} (y_i - \bar{y}_i)^2$$
 (4)

3. Adjusted-R² statistic

$$R^{2} = \frac{TSS - RSS}{TSS} = 1 - \frac{RSS/(n - d - 1)}{TSS/(n - 1)}$$
 (5)

2 Model Selection and Implementation

Our goal is finding out the major and minor factors that affect the price. Although the linear model is more restrictive, linear model is of good explanatory, so the linear model is the best choice. Assuming that we don't choose a linear model, but choose other models with higher flexibility, we need to train the model with much bigger data than the training data used in linear model and the model we get is difficult to explain how any individual predictor is associated with the response.

2.1 Full Model

In full model, we take 13 variables into consideration. From the result of regression, it is shown that only few varibles play the major part in Boston house price. Therefore, it is urgent to improve the model. Statistics of full model are attached below.

	Estimate	Std.Error	t value	Pr(> t)
intercept	0.210528	0.252044	0.835	0.404177
crim	-0.048952	0.053972	-0.907	0.365087
zn	0.028821	0.014665	1.965	0.050246
indus	0.026266	0.063904	0.411	0.681325
chas	0.096118	0.031029	3.098	0.002122
nox	-8.215957	4.023916	-2.042	0.041985
rm	7.433033	0.455925	16.303	<0.0001
age	-0.028337	0.012587	-2.251	0.025042
dis	-1.195338	0.210403	-5.681	<0.0001
rad	0.345224	0.093474	3.693	0.000260
tax	-0.041890	0.012559	-3.335	0.000951
ptratio	-0.742733	0.125414	-5.922	<0.0001
black	0.031737	0.008901	3.565	0.000418
lstat	-0.285383	0.059129	-4.826	<0.0001

	RSE	R^2	Adjusted ${\cal R}^2$
training data	0.1508	0.8762	0.8712
test data	0.1106	0.8324	0.8172

2.2 Stepwise Selection

First, we sorted the variable through stepwise selection and we got the order of variables in order of importance. Therefore, we choose the first seven major variables, which are showed below.

```
crim zn indus chas nox rm age dis rad tax ptratio black lstat
            11 11
                                                                   11 11
2
  1 (1)
                                                                   11 11
                11 11 11 11
                               и и пап и и и и и и пап
3
  2 (1)
                                                                   пъп
4
5
                                                                   \Pi \not\cong \Pi
6
  5 (1)
                                                                   \Pi \not\cong \Pi
                                                              \Pi \not\simeq \Pi
7
   6 (1)
                                                                   пъп
                                                              пуп
8
  7 (1)
                                                              пжп
                                                                   пъп
9
  8 (1)
                                                                   11 ½ 11
  9 (1)
```

11	10	(1)""		11 % 11	11 * 11	11 % 11	11 % 11	11 % 11	11 % 11	11 % 11	11 1/2 11	11 % 11	11.7.11
12	11	(1)""	п*п п п	пжп	11 % 11	11 % 11	11 % 11	11 % 11	11 % 11	п*п	11.44.11	11 % 11	п*п
13	12	(1)"*"	п*п п п	пжп	11 % 11	11 % 11	11 % 11	11 % 11	11 % 11	п*п	11.44.11	11 % 11	п*п
14	13	(1)"*"	11×11 11×11	11 % 11	" * "	11 % 11	"*"	11 % 11	11 % 11	11 % 11	11 * 11	пУп	пУп

	Estimate	Std.Error	t value	Pr(> t)
intercept	-0.340526	0.061531	-5.534	<0.0001
rm	8.262658	0.279706	29.541	<0.0001
age	-0.040181	0.012491	-3.217	0.00142
dis	-0.920175	0.171070	-5.379	<0.0001
ptratio	-0.709611	0.105491	-6.727	<0.0001
black	0.035939	0.006729	5.341	<0.0001
lstat	-0.278504	0.055662	-5.004	<0.0001

	RSE	R^2	Adjusted ${\cal R}^2$
training data	0.1573	0.8624	0.8599
test data	0.1521	0.6830	0.6692

2.3 Lasso Model

The second term of the equation, $\lambda \sum_{j=1}^p |\beta_j|$, called the shrinkage penalty, has the effect of shrinking the estimates of β_j towars zero. The turning parameter λ serves to control the relative impact of these two terms on the regression coefficient estimates.

$$\sum_{i=1}^{n} (y_i - \beta_0 - \sum_{j=1}^{p} \beta_j x_{ij})^2 + \lambda \sum_{j=1}^{p} |\beta_j| = RSS + \lambda \sum_{j=1}^{p} |\beta_j|$$
 (6)

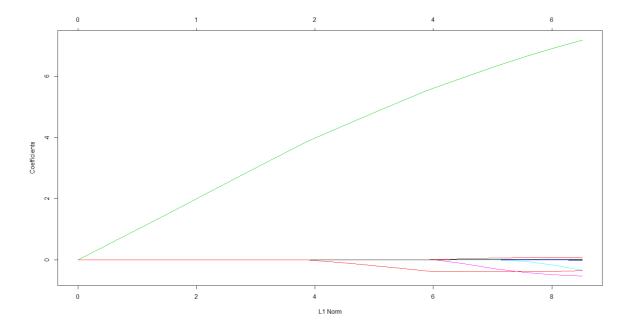
test

$$R^2 = 0.7823308$$
 $RSE = 0.1260907$ $Adjusted - R^2 = 0.7625427$

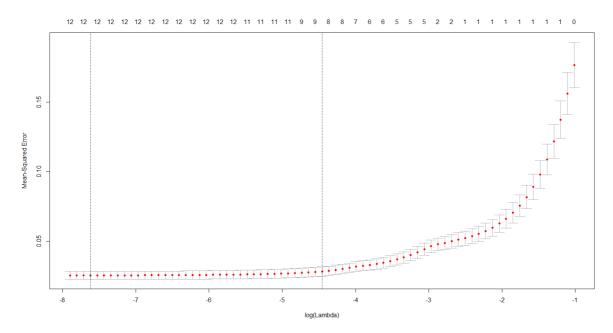
train

$$R^2 = 0.7941389 \hspace{1.5cm} RSE = 0.1914669 \hspace{1.5cm} Adjusted - R^2 = 0.7851884$$

Some of those coefficients go to zero denpending on the choice of tuning parameter.



We can also see the Mean-Value Error becomes bigger as λ becomes larger.



And the λ was chosen as:

$$\lambda = 0.0003384039 \tag{7}$$

	RSE	R^2	Adjusted ${\cal R}^2$
training data	0.1915	0.7941	0.7852
test data	0.1261	0.7823	0.7625

3 Assess the Model

We compare the effect of the three regression model and illustrate them.

stepwise 出现了过拟合现象。

full model 表现最好。



From the two graph above, we can draw two main conclusion: one is that the full model fits the data best, another is stepwise regression overfits the training data.

References

[1] Harrison, D. and Rubinfeld, D.L. (1978) Hedonic prices and the demand for clean air. J. Environ. Economics and Management 5, 81–102.

Appendix

codes for computing statistics

```
1
    RSE = function(y, x, beta){
 2
        a = sqrt(RSS(y,x,beta)/(length(y)-2))
 3
        print(a)
 4
        return(a)
 5
    }
 6
 7
    R_{square} = function(y, x, beta){
 8
       TSS = TSS(y)
 9
        RSS = RSS(y, x, beta)
10
        r = 1 - RSS/TSS
11
        print(r)
12
        return(r)
13
14
15
    Adjusted_R_square = function(y, x ,beta){
16
        TSS = TSS(y)
17
        RSS = RSS(y, x, beta)
18
        r = 1- RSS*(length(y)-1) / ((length(y)-1-length(beta))*TSS)
19
        print(r)
20
        return(r)
21
    }
22
23
    RSS = function(y, x, beta){
24
        y_hat = c()
25
        for (i in 1:length(y)){
26
            temp = 0
27
            for (j in 2: length(beta)){
                 temp = temp + x[i,j-1] * beta[j]
28
29
30
           y_hat[i] = temp + beta[1]
        }
31
32
        return(sum((y-y_hat)^2))
33
    }
34
35
    TSS = function(y){
        sum((y - median(y))^2)
36
```

```
37 | 3 | 38 |
```

codes for standardizing the data

```
1
    standardizing = function(x0){
 2
        chas = x0[,4]
 3
        x = x0[,-4]
 4
        n = dim(x)[1]
 5
        p = dim(x0)[2]
        X_ <- X
 6
 7
        for (j in 1:dim(x)[2]){
            x_bar = median(x[,j])
 9
             for (i in 1:dim(x)[1]){
                 sq = sum((x[i,]-x_bar)^2)
10
                 x_{[i,j]} = x[i,j]/sqrt(1/n*sq)
11
12
             }
13
14
        x_{new} = cbind(x_{1:3}, chas, x_{5:p-1})
    }
15
```

codes for data separating

```
1
    # The function was used to seperate data into training sets and test sets
 2
    set.seed(1)
 3
    sub<-sample(1:nrow(Boston), round(nrow(Boston)*2/3))</pre>
    length(sub)
    data_train<-Boston[sub,]# get 2/3 of data as training sets
    data_test<-Boston[-sub,]# get 1/3 of data as test sets
 7
    dim(data_train)
 8
    dim(data_test)
 9
    head(data_train)
10
    head(data_test)
11
    write.table(data_train, file = "data/data_train.csv", append = FALSE, quote
    = TRUE, sep = ",",
                eol = "\n", na = "NA", dec = ".", row.names = FALSE,
12
13
                col.names = TRUE, qmethod = c("escape", "double"),
                fileEncoding = "")
14
    write.table(data_test, file = "data/data_test.csv", append = FALSE, quote =
15
    TRUE, sep = ",",
                eol = "\n", na = "NA", dec = ".", row.names = FALSE,
16
                col.names = TRUE, qmethod = c("escape", "double"),
17
                fileEncoding = "")
18
19
```

codes for model fitting

```
# Loading standarding data
train_stand = read.table("data/train_stand.csv", header = T, na.string =
   "?", sep = ",")

test_stand = read.table("data/test_stand.csv", header = T, na.string =
   "?", sep = ",")

### Full Model
pairs(train_stand)
```

```
lm.full = lm(medv~., data = train_stand)
8
    summary(lm.full)
 9
10 | beta_full = coefficients(lm.full)
11 | R_2_full = R_square(test_stand[,ncol(test_stand)],
    test_stand[,1:ncol(test_stand)-1], beta_full)
12 | RSE_full = RSE(test_stand[,ncol(test_stand)],
    test_stand[,1:ncol(test_stand)-1], beta_full)
    Adjusted_R_2_full = Adjusted_R_square(test_stand[,ncol(test_stand)],
    test_stand[,1:ncol(test_stand)-1], beta_full)
14
15
16 ### Backward Model
17 regfit.bwd = regsubsets(medv~.,data=train_stand ,nvmax=19,
    method="backward")
18 | summary(regfit.bwd)
19
20 | lm.bwd = lm(medv~rm + age + dis + ptratio + black + lstat, data =
    train_stand)
21 summary(1m.bwd)
22
23 beta_bwd =coefficients(1m.bwd)
24 | test_stand_1 = test_stand[,-10][,-9][,-5][,-4][,-3][,-2][,-1]
25 | R_2_bwd = R_square(test_stand[,ncol(test_stand)], test_stand_1, beta_bwd)
   RSE_bwd = RSE(test_stand[,ncol(test_stand)], test_stand_1, beta_bwd)
27 | Adjusted_R_2_bwd = Adjusted_R_square(test_stand[,ncol(test_stand)],
    test_stand_1, beta_bwd)
28
29
30 ### Lasso adjusted
31 | library(Matrix)
32 library(foreach)
33 library(glmnet)
34 | x_train = model.matrix(medv~., train_stand)[,-1]
35 | x_test = model.matrix(medv~., test_stand)[,-1]
36 | y_train = train_stand$medv
37
   y_test = test_stand$medv
38
39 grid = 10^seq(10, -2, length = 100)
40
    lasso.mod=glmnet(x_train,y_train,alpha=1,lambda=grid)
41 plot(lasso.mod)
43 | set.seed(1)
44 | cv.out = cv.glmnet(x_train, y_train, alpha = 1)
45
   plot(cv.out)
46
47
    bestlam=cv.out$lambda.min
48
49
    lasso.pred=predict(lasso.mod,s=bestlam ,newx=x_test)
50
    plot(lasso.pred)
51
    mean((lasso.pred -y_test)^2)
    print(bestlam)
53
54 #coefficients
55
    Beta_lasso = predict(lasso.mod,type="coefficients",s=bestlam)
56
57
    R_2_lasso = R_square(y_test,x_test, beta_lasso)
```

```
59
     RSE_lasso = RSE(y_test,x_test, beta_lasso)
60
     Adjusted_R_2_Lasso = Adjusted_R_square(y_test, x_test, beta_lasso)
61
62
63 R_2_lasso = R_square(y_train, x_train, beta_lasso)
64 RSE_lasso = RSE(y_train, x_train, beta_lasso)
    Adjusted_R_2_Lasso = Adjusted_R_square(y_train, x_train, beta_lasso)
65
66
67 | ### Lasso + Stepwise
68
69 | x_train = model.matrix(medv~rm + age + dis + ptratio + black + lstat,
     train_stand)[,-ncol(train_stand)]
70 | x_test = model.matrix(medv~rm + age + dis + ptratio + black + lstat,
    test_stand)[,-ncol(test_stand)]
71 | y_train = train_stand$medv
72  y_test = test_stand$medv
73
74 | grid = 10^seq(10, -2, length = 100)
75 | lasso.mod=glmnet(x_train,y_train,alpha=1,lambda=grid)
76 plot(lasso.mod)
77
78 set.seed(1)
79 cv.out = cv.glmnet(x_train, y_train, alpha = 1)
80 plot(cv.out)
81
82 | bestlam=cv.out$lambda.min
83
84
    lasso.pred=predict(lasso.mod,s=bestlam ,newx=x_test)
85 mean((lasso.pred -y_test)^2)
86
     print(bestlam)
87
88
    #coefficients
89
     Beta_lasso = predict(lasso.mod,type="coefficients",s=bestlam)
90
91
92 R_2_lasso = R_square(test_stand, x_test, beta_lasso)
93
     RSE_lasso = RSE(y_test,x_test, beta_lasso)
94
    Adjusted_R_2_Lasso = Adjusted_R_square(y_test, x_test, beta_lasso)
95
96
     R_2_lasso = R_square(y_train, x_train, beta_lasso)
97
98
     RSE_lasso = RSE(y_train, x_train, beta_lasso)
99
     Adjusted_R_2_Lasso = Adjusted_R_square(y_train, x_train, beta_lasso)
100
```