

# ACS : Information Theory

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Remark: I will not solve the easy questions

## Exercise 2

$$\text{Bayes' formula : } \begin{cases} P_{X|Y}(x | y) = \frac{P_{X,Y}(x,y)}{P_Y(y)} \\ P_{Y|X}(y | x) = \frac{P_{X,Y}(x,y)}{P_X(x)} \\ P_Y(y) = \sum_{x' \in \mathcal{X}} P_{X,Y}(x', y) \end{cases} \implies P_{X|Y}(x | y) = \frac{P_X(x) P_{Y|X}(y|x)}{\sum_{x'} P_X(x') P_{Y|X}(y|x')}$$

## Exercise 3

Given  $P_{X,Y}$  of the  $BSC(p)$  channel

Assume that  $\mathcal{X} = \{0, 1\}$  and  $\mathcal{Y} = \{0, 1, E\}$  and that  $P_{X,Y}$  is given in the table

$Y, X$	0	1
0	$\frac{1-e}{2}$	0
1	0	$\frac{1-e}{2}$
$E$	$\frac{e}{2}$	$\frac{e}{2}$

 $\implies$ 

$Y X$	0	1
0	$1-e$	0
1	0	$1-e$
$E$	$e$	$e$

$e$  is called an erasure probability.

This channel model can also be written in an additive form,

$$Y = X \oplus W$$

where  $X$  is a Bern(0.5) random variable,  $W$  is a Bern( $e$ ) random variable on  $\{E, 1\}$ ,  $X$  and  $W$  are independent, and  $\oplus$  is the binary XOR operation.

## Exercise 5

The  $KL$  divergence verifies a certain set of properties :

1. Asymmetry :  $D_{KL}(P_X \| Q_X) \neq D_{KL}(Q_X \| P_X)$ .
2. Null element :  $D_{KL}(P_X \| P_X) = 0$ .
3. Positivity:  $D_{KL}(P_X \| Q_X) \geq 0$ , for all laws  $(P_X, Q_X) \neq 0$

**Gibbs' inequality:**

Suppose that

$$P = \{p_1, \dots, p_n\}$$

is a discrete **probability distribution**. Then for any other probability distribution

$$Q = \{q_1, \dots, q_n\}$$

the following inequality between positive quantities (since  $P$  and  $Q$  are between zero and one)

$$-\sum_{i=1}^n p_i \log p_i \leq -\sum_{i=1}^n p_i \log q_i$$

with equality if and only if

$$p_i = q_i$$

for all  $i$ . Put in words, the **information entropy** of a distribution  $P$  is less than or equal to its **cross entropy** with any other distribution  $Q$ .

The difference between the two quantities is the **Kullback–Leibler divergence** or relative entropy

$$D_{\text{KL}}(P\|Q) \equiv \sum_{i=1}^n p_i \log \frac{p_i}{q_i} \geq 0.$$

**Proof :**

**First method :** Let  $I$  denote the set of all  $i$  for which  $p_i$  is non-zero. Then, since  $\ln x \leq x - 1$  for all  $x > 0$ , with equality if and only if  $x = 1$ , we have:

$$\begin{aligned} & - \sum_{i \in I} p_i \ln \frac{q_i}{p_i} \geq - \sum_{i \in I} p_i \left( \frac{q_i}{p_i} - 1 \right) \\ = & - \sum_{i \in I} q_i + \sum_{i \in I} p_i = - \sum_{i \in I} q_i + 1 \geq 0 \end{aligned}$$

The last inequality is a consequence of the  $p_i$  and  $q_i$  being part of a probability distribution. Specifically, the sum of all non-zero values is 1. Some non-zero  $q_i$ , however, may have been excluded since the choice of indices is conditioned upon the  $p_i$  being non-zero. Therefore the sum of the  $q_i$  may be less than 1.

So far, over the index set  $I$ , we have:

$$- \sum_{i \in I} p_i \ln \frac{q_i}{p_i} \geq 0$$

or equivalently

$$- \sum_{i \in I} p_i \ln q_i \geq - \sum_{i \in I} p_i \ln p_i$$

Both sums can be extended to all  $i = 1, \dots, n$ , i.e. including  $p_i = 0$ , by recalling that the expression  $p \ln p$  tends to 0 as  $p$  tends to 0, and  $(-\ln q)$  tends to  $\infty$  as  $q$  tends to 0. We arrive at

$$- \sum_{i=1}^n p_i \ln q_i \geq - \sum_{i=1}^n p_i \ln p_i$$

**Second method :** Below we give a proof based on Jensen's inequality:

Because  $\log$  is a concave function, we have that:

$$\sum_i p_i \log \frac{q_i}{p_i} \leq \log \sum_i p_i \frac{q_i}{p_i} = \log \sum_i q_i \leq 0$$

The first inequality is due to Jensen's inequality, and the last is due to the same reason given in the above proof.

**Third method :**

**Log sum inequality**

==Statement==

Let  $a_1, \dots, a_n$  and  $b_1, \dots, b_n$  be nonnegative numbers. Denote the sum of all  $a_i$ s by  $a$  and the sum of all  $b_i$ s by  $b$ . The log sum inequality states that

$$\sum_{i=1}^n a_i \log \frac{a_i}{b_i} \geq a \log \frac{a}{b},$$

==Proof==

Notice that after setting  $f(x) = x \log x$  we have

$$\begin{aligned} \sum_{i=1}^n a_i \log \frac{a_i}{b_i} &= \sum_{i=1}^n b_i f\left(\frac{a_i}{b_i}\right) = b \sum_{i=1}^n \frac{b_i}{b} f\left(\frac{a_i}{b_i}\right) \\ &\geq b f\left(\sum_{i=1}^n \frac{b_i}{b} \frac{a_i}{b_i}\right) = b f\left(\frac{1}{b} \sum_{i=1}^n a_i\right) = b f\left(\frac{a}{b}\right) \\ &= a \log \frac{a}{b}, \end{aligned}$$

where the inequality follows from **Jensen's inequality** since  $\frac{b_i}{b} \geq 0$ ,  $\sum_{i=1}^n \frac{b_i}{b} = 1$ , and  $f$  is convex.

—Let  $P = (p_i)_{i \in N}$  and  $Q = (q_i)_{i \in N}$  be pmfs. In the log sum inequality, substitute  $n = \infty$ ,  $a_i = p_i$  and  $b_i = q_i$  to get

$$D_{\text{KL}}(P \| Q) \equiv \sum_i p_i \log_2 \frac{p_i}{q_i} \geq 1 \log \frac{1}{1} = 0$$

**Fourth method : As particular case of Bregman divergence**  
**Exercise 6**

Here are a few special cases of entropy:

1. The entropy of a constant random variable,  $X = c$  with probability 1, is given by

$$H(X) = 0.$$

2. The entropy of a random variable  $X$  uniformly distributed over  $\mathcal{X}$ , c.à.d,  $P_X(x) = \frac{1}{|\mathcal{X}|}$  for all  $x \in \mathcal{X}$ , is given by

$$H(X) = \log_2(|\mathcal{X}|).$$

3. Let  $X$  be a binary Bernoulli random variable  $\text{Bern}(p)$  where  $p = P_X(1)$ . The entropy of  $X$  is given by

$$H(X) = H_2(p) - p \log_2(p) - (1-p) \log_2(1-p) = H_2(p).$$

$H_2(p)$  is commonly known as the binary entropy function. It is maximal when  $p = 0.5$  (equals 1 bit), and is minimal when  $p = 0$  or  $p = 1$ , (equals 0 bits). It is symmetric around  $p = 0.5$ .

**Exercise 7**

Discrete entropy  $H(X)$  satisfies the following properties :

1. Minimum :  $H(X) \geq 0$  for all discrete distributions  $P_X$ , and the minimum is achieved for a degenerate distribution, i.e.,  $X$  is a constant.
2. Maximum:  $H(X) \leq \log_2(|\mathcal{X}|)$  for all discrete distributions  $P_X$ , and the maximum is achieved for the uniform distribution over  $\mathcal{X}$ .
3. Data Processing Inequality (DPI) : the entropy of a function  $f$  of  $X$ , is no greater than the entropy of  $X$ , i.e.,

$$H(f(X)) \leq H(X)$$

with equality iff  $f$  is a one-to-one function.

**Corollary of Gibbs' inequality**

Information entropy of  $P$  is bounded by:

$$H(p_1, \dots, p_n) \leq \log n.$$

The proof is trivial – simply set  $q_i = 1/n$  for all "i".

**Exercise 9**

Conditional entropy  $H(X | Y)$  satisfies the following properties

1. Asymmetry :  $H(X | Y) \neq H(Y | X)$
2. Minimum:  $H(X | Y) \geq 0$  for all  $P_{X,Y}$ , and the minimum is achieved when  $P_{X|Y}$  is degenerate, i.e.,  $X$  is a function of  $Y$ .
3. Upper bound: conditional entropy is always smaller than the individual entropy

$$H(X | Y) \leq H(X)$$

with equality iff  $X$  and  $Y$  are independent :  $P_{X,Y} = P_X P_Y$ .

4. Maximum:  $H(X | Y) \leq \log_2(|\mathcal{X}|)$  for all  $P_{X,Y}$  and the maximum is achieved when  $X$  and  $Y$  are independent, and  $X$  is uniform, i.e.,

$$\left\{ \begin{array}{l} \text{Independence : } \forall(x,y) \quad P_{X,Y}(x,y) = P_X(x)P_Y(y) \\ P_{X|Y}(x | y) = \frac{P_{X,Y}(x,y)}{P_Y(y)}, P_{Y|X}(y | x) = \frac{P_{X,Y}(x,y)}{P_X(x)} \\ H(X | Y) = \sum_{(x,y) \in \mathcal{X} \times \mathcal{Y}} P_{X,Y}(x,y) \log_2 (P_{X|Y}(x | y)) \\ \quad = \sum_{y \in \mathcal{Y}} P_Y(y) H(X | Y = y) \\ H(X | Y = y) = \sum_{x \in \mathcal{X}} P_{X|Y}(x | y) \log_2 (P_{X|Y}(x | y)) \end{array} \right. \implies H(X | Y) = H(X)$$

*Propertie 2 :* When  $X$  is a function of  $Y$

*Hint :*

$$\log_2 (P_{X|Y}(x | y)) = 0$$

### Exercise 10

Compute the conditional entropy  $H(Y | X)$  where  $Y = X \oplus W$ , where  $X$  follows a Bern(1/2), independent from  $W$  which follows a Bern( $p$ ) and  $\oplus$  is the binary *XOR* operation. (Hint :  $P_{X,Y}$  and  $P_{Y|X}$  were given previously in example 2)

**Propertie :** The individual entropies  $H(X)$  and  $H(Y)$ , the joint entropy  $H(X, Y)$  and the conditional entropies  $H(X | Y)$  et  $H(Y | X)$  can be related as follows

$$\begin{aligned} H(X, Y) &= H(X) + H(Y | X) \\ &= H(Y) + H(X | Y) \end{aligned}$$

### Exercise 13

The differential entropy of a real-valued Gaussian random variable  $X_G \sim \mathcal{N}(\mu, \sigma^2)$  is given by

$$h(X_G) = \frac{1}{2} \log_2 (2\pi e \sigma^2)$$

### Exercise 14

Mutual information is positive.

Hint : proof is similar to the proof of positivity of the *KL* divergence.

Proof :

$$I(X; Y) = \sum_{(x,y) \in \mathcal{X} \times \mathcal{Y}} P_{X,Y}(x,y) \log_2 \left( \frac{P_{X,Y}(x,y)}{P_X(x)P_Y(y)} \right)$$

If we use this definition and the exercise 5, we deduct simply many versions of the proof.

### Exercise 18

Shannon formula for the AWGN :

$$\log_2(1 + SNR)$$

Let  $X \sim \mathcal{N}(0, P)$  and  $W \sim \mathcal{N}(0, \sigma^2)$  be two independent Gaussian random variables. Assume that we observe  $Y$  given by

$$Y = X + W.$$

This model describes the so called Additive White Gaussian Noise (AWGN) channel with input signal  $X$ , additive noise  $W$ , and output signal  $Y$ .

The mutual information between  $X$  and  $Y$  is given by :

$$I(X; Y) = \frac{1}{2} \log_2 \left( 1 + \frac{P}{\sigma^2} \right).$$

*Proof* We have :

$$\begin{aligned} I(X; Y) &= h(X) - h(Y | X) \\ &= h(X) - h(Y | X) \\ &= h(X) + h(Y) - h(X; Y) \end{aligned}$$

Because  $X \perp\!\!\!\perp W$  then

$$h(Y) = \frac{1}{2} \log(2\pi \exp(\sigma^2 + P))$$

Or

$$h(Y|X) = h(W) = \frac{1}{2} \log(2\pi \exp \sigma^2)$$

Finally :

$$I(X; Y) = h(X) - h(Y | X) = \frac{1}{2} \log(2\pi \exp(\sigma^2 + P)) - \frac{1}{2} \log(2\pi \exp \sigma^2) = \frac{1}{2} \log_2 \left( 1 + \frac{P}{\sigma^2} \right)$$