

# MVA - Probabilistic Graphical Models

## DM1

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### 1 Learning in discrete graphical models

Consider the following model:  $z$  and  $x$  are discrete variables taking respectively  $M$  and  $K$  different values with  $p(z = m) = \pi_m$  and  $p(x = k|z = m) = \theta_{mk}$ .

Let  $\{(z_1, x_1), \dots, (z_n, x_n)\}$  be a sample of  $n$  observations. Since they are i.i.d, we have the likelihood function :

$$\begin{aligned} L(\pi, \theta) &= \prod_{i=1}^n p(z_i, x_i | \pi; \theta) \\ &= \prod_{i=1}^n p(x_i | z_i; \pi; \theta) p(z_i | \pi) \text{ using Bayes' rule} \\ &= \prod_{i=1}^n \theta_{z_i x_i} \pi_{z_i} \\ &= \prod_{i=1}^n \left( \prod_{m=1}^M \prod_{k=1}^K \theta_{mk}^{\mathbb{1}_{\{z_i=m\}} \mathbb{1}_{\{x_i=k\}}} \right) \left( \prod_{m=1}^M \pi_m^{\mathbb{1}_{\{z_i=m\}}} \right) \end{aligned}$$

We introduce the variables  $z_{im} = \mathbb{1}_{\{z_i=m\}}$  and  $x_{ik} = \mathbb{1}_{\{x_i=k\}}$  to simplify the notations.

When passing to the log, we then get the following log-likelihood function:

$$\begin{aligned} l(\pi, \theta) &= \sum_{i=1}^n \left( \sum_{m=1}^M \sum_{k=1}^K \log(\theta_{mk}^{z_{im} x_{ik}}) + \sum_{m=1}^M \log(\pi_m^{z_{im}}) \right) \\ &= \sum_{m=1}^M \sum_{k=1}^K \left( \sum_{i=1}^n z_{im} x_{ik} \right) \log(\theta_{mk}) + \sum_{m=1}^M \left( \sum_{i=1}^n z_{im} \right) \log(\pi_m) \end{aligned}$$

Our goal is to maximize this function  $l(\pi, \theta)$  while respecting the constraints on the probabilities :

- $\sum_{m=1}^M \pi_m = 1$
- $\forall m \in \{1, \dots, M\} \quad \sum_{k=1}^K \theta_{mk} = 1$

The two terms of the sum in  $l(\pi, \theta)$  are independant, therefore we can maximize them separately.

## MLE for $\pi$

We consider the problem:

$$\begin{aligned} \min_{\pi} & - \sum_{m=1}^M \left( \sum_{i=1}^n z_{im} \right) \log(\pi_m) \\ \text{s.t.} & \sum_{m=1}^M \pi_m = 1 \end{aligned}$$

The Langrangian of the problem is :

$$\mathcal{L}(\pi, \lambda) = - \sum_{m=1}^M \left( \sum_{i=1}^n z_{im} \right) \log(\pi_m) + \lambda \left( \sum_{m=1}^M \pi_m - 1 \right)$$

And the dual function is :

$$g(\lambda) = \min_{\pi} \mathcal{L}(\pi, \lambda)$$

Since  $\mathcal{L}(\pi, \lambda)$  is convex in  $\pi$ , we can find its minimum with respect to  $\pi$  by looking at the gradients with respect to the components of  $\pi$  :

$$\frac{\partial \mathcal{L}}{\partial \pi_m} = - \frac{\sum_{i=1}^n z_{im}}{\pi_m} + \lambda$$

Which is equal to 0 for  $\pi_m = \frac{\sum_{i=1}^n z_{im}}{\lambda}$ . To find  $\lambda$ , we look at the constraint that gives us :

$$\sum_{m=1}^M \frac{\sum_{i=1}^n z_{im}}{\lambda} = 1$$

Hence  $\lambda = n$  and the solution is  $\pi_m = \frac{\sum_{i=1}^n z_{im}}{n}$  with  $\sum_{i=1}^n z_{im}$  the number of observations of  $z$  that are equal to  $m$ .

## MLE for $\theta$

We consider the problem:

$$\begin{aligned} \min_{\theta} & - \sum_{m=1}^M \sum_{k=1}^K \left( \sum_{i=1}^n z_{im} x_{ik} \right) \log(\theta_{mk}) \\ \text{s.t.} & \forall m \in \{1, \dots, M\} \quad \sum_{k=1}^K \theta_{mk} = 1 \end{aligned}$$

The Langrangian of the problem is :

$$\mathcal{L}(\theta, \lambda) = - \sum_{m=1}^M \sum_{k=1}^K \left( \sum_{i=1}^n z_{im} x_{ik} \right) \log(\theta_{mk}) + \sum_{m=1}^M \lambda_m \left( \sum_{k=1}^K \theta_{mk} - 1 \right)$$

And the dual function is :

$$g(\lambda) = \min_{\theta} \mathcal{L}(\theta, \lambda)$$

Since  $\mathcal{L}(\pi, \lambda)$  is convex in  $\theta$ , we can find its minimum with respect to  $\theta$  by looking at the gradients with respect to the components of  $\theta$  :

$$\frac{\partial \mathcal{L}}{\partial \theta_{mk}} = -\frac{\sum_{i=1}^n z_{im} x_{ik}}{\theta_{mk}} + \lambda_m.$$

Which is equal to 0 for  $\theta_{mk} = \frac{\sum_{i=1}^n z_{im} x_{ik}}{\lambda_m}$ .

To find the  $\lambda_m$ , we look at the constraints that give us :

$$\sum_{k=1}^K \frac{\sum_{i=1}^n z_{im} x_{ik}}{\lambda_m} = 1$$

Hence  $\lambda_m = \sum_{i=1}^n z_{im}$  and the solution is  $\theta_{mk} = \frac{\sum_{i=1}^n z_{im} x_{ik}}{\sum_{i=1}^n z_{im}}$  with  $\sum_{i=1}^n z_{im}$  the number of observations of  $z$  that are equal to  $m$  and  $\sum_{i=1}^n z_{im} x_{ik}$  the number of observations where  $z$  is equal to  $m$  and  $x$  is equal to  $k$  simultaneously.

## 2 Linear classification

### 2.1 Generative model (LDA)

a. Let  $\{(x_1, y_1), \dots, (x_n, y_n)\}$  be a sample of  $n$  observations with the  $x_i$  in  $\mathbb{R}^2$  and the  $y_i$  in  $\{0, 1\}$ . Since they are i.i.d, we have the likelihood function :

$$\begin{aligned} L(\pi, \mu_0, \mu_1, \Sigma) &= \prod_{i=1}^n p(x_i, y_i | \pi, \mu_0, \mu_1, \Sigma) \\ &= \prod_{i=1}^n p(x_i | y_i; \pi, \mu_0, \mu_1, \Sigma) p(y_i | \pi) \text{ using Bayes' rule} \\ &= \prod_{i=1}^n \pi^{y_i} (1 - \pi)^{1-y_i} f_{\mu_{y_i}}(x_i) \end{aligned}$$

Where  $f_{\mu_{y_i}}(x_i) = \frac{1}{2\pi\sqrt{\det \Sigma}} \exp(-\frac{1}{2}(x - \mu_{y_i})^T \Sigma^{-1}(x - \mu_{y_i}))$ . To simplify we can write :

$$L(\pi, \mu_0, \mu_1, \Sigma) = \prod_{i=1}^n \pi^{y_i} (1 - \pi)^{1-y_i} f_{\mu_0}(x_i)^{1-y_i} f_{\mu_1}(x_i)^{y_i}$$

And we get the log-likelihood :

$$\begin{aligned} l(\pi, \mu_0, \mu_1, \Sigma) &= \sum_{i=1}^n (y_i \log \pi + (1 - y_i) \log(1 - \pi)) \\ &\quad + (1 - y_i)(f_{\mu_0}(x_i)) \\ &\quad + y_i(f_{\mu_1}(x_i)) \end{aligned}$$

$$\begin{aligned}
l(\pi, \mu_0, \mu_1, \Sigma) &= \sum_{i=1}^n (y_i \log \pi + (1 - y_i) \log(1 - \pi)) \\
&+ (1 - y_i) \left( -\log(2\Pi) - \frac{1}{2} \log(\det \Sigma) - \frac{1}{2} (x_i - \mu_0)^T \Sigma^{-1} (x_i - \mu_0) \right) \\
&+ y_i \left( -\log(2\Pi) - \frac{1}{2} \log(\det \Sigma) - \frac{1}{2} (x_i - \mu_1)^T \Sigma^{-1} (x_i - \mu_1) \right)
\end{aligned}$$

Our goal is to maximize this function, so we look at the gradients with respect to the parameters to find for which they are equal to 0:

**For  $\pi$ .**

$$\begin{aligned}
\frac{\partial l}{\partial \pi}(\pi, \mu_0, \mu_1, \Sigma) &= 0 \\
\Leftrightarrow \sum_{i=1}^n \frac{y_i}{\pi} - \frac{1 - y_i}{1 - \pi} &= 0 \\
\Leftrightarrow \frac{1}{\pi} \sum_{i=1}^n y_i &= \frac{1}{1 - \pi} \sum_{i=1}^n (1 - y_i) \\
\Leftrightarrow \left( \frac{1}{\pi} + \frac{1}{1 - \pi} \right) \sum_{i=1}^n y_i &= \frac{n}{1 - \pi} \\
\Leftrightarrow \left( \frac{1 - \pi}{\pi} + 1 \right) \sum_{i=1}^n y_i &= n \\
\Leftrightarrow \hat{\pi} &= \frac{\sum_{i=1}^n y_i}{n}
\end{aligned}$$

**For  $\mu_0$ .**

$$\begin{aligned}
\frac{\partial l}{\partial \mu_0}(\pi, \mu_0, \mu_1, \Sigma) &= 0 \\
\Leftrightarrow - \left( \sum_{i=1}^n (1 - y_i) \Sigma^{-1} (x_i - \mu_0) \right) &= 0 \\
\Leftrightarrow \sum_{i=1}^n x_i (1 - y_i) - \mu_0 \sum_{i=1}^n (1 - y_i) &= 0 \\
\Leftrightarrow \hat{\mu}_0 &= \frac{\sum_{i=1}^n x_i (1 - y_i)}{\sum_{i=1}^n (1 - y_i)}
\end{aligned}$$

**For  $\mu_1$ .**

$$\begin{aligned}
& \frac{\partial l}{\partial \mu_1}(\pi, \mu_0, \mu_1, \Sigma) = 0 \\
& \Leftrightarrow -\left(\sum_{i=1}^n (y_i) \Sigma^{-1} (x_i - \mu_1)\right) = 0 \\
& \Leftrightarrow \sum_{i=1}^n x_i y_i - \mu_1 y_i = 0 \\
& \Leftrightarrow \boxed{\hat{\mu}_1 = \frac{\sum_{i=1}^n x_i y_i}{\sum_{i=1}^n y_i}}
\end{aligned}$$

**For  $\Sigma$ .**

$$\begin{aligned}
& \frac{\partial l}{\partial \Sigma^{-1}}(\pi, \mu_0, \mu_1, \Sigma) = 0 \\
& \Leftrightarrow \frac{\partial}{\partial \Sigma^{-1}} \left( \sum_{i=1}^n \frac{(1-y_i)}{2} (\log(\det \Sigma^{-1}) - \text{Tr}((x_i - \mu_0)^T \Sigma^{-1} (x_i - \mu_0))) \right. \\
& \quad \left. + \frac{y_i}{2} (\log \det \Sigma^{-1} - \text{Tr}((x_i - \mu_1)^T \Sigma^{-1} (x_i - \mu_1))) \right) = 0 \\
& \Leftrightarrow \sum_{i=1}^n \frac{1-y_i}{2} (\Sigma - (x_i - \mu_0)(x_i - \mu_0)^T) + \frac{y_i}{2} (\Sigma - (x_i - \mu_1)(x_i - \mu_1)^T) = 0 \\
& \Leftrightarrow \boxed{\hat{\Sigma} = \frac{1}{n} \sum_{i=1}^n (1-y_i)(x_i - \mu_0)(x_i - \mu_0)^T + y_i(x_i - \mu_1)(x_i - \mu_1)^T}
\end{aligned}$$

**b.** We aim to determine the form of  $p(y = 1|x)$ . By applying Bayes' rule, we have :

$$\begin{aligned}
\mathbb{P}(Y = 1|X = x) &= \frac{\mathbb{P}(Y = 1, X = x)}{\mathbb{P}(X = x)} \\
&= \frac{\mathbb{P}(X = x|Y = 1)\mathbb{P}(Y = 1)}{\mathbb{P}(X = x)} \\
&= \frac{\mathbb{P}(X = x|Y = 1)\mathbb{P}(Y = 1)}{\mathbb{P}(X = x|Y = 1)\mathbb{P}(Y = 1) + \mathbb{P}(X = x|Y = 0)\mathbb{P}(Y = 0)} \\
&= \frac{f_{\mu_1}(x)\pi}{f_{\mu_1}(x)\pi + f_{\mu_0}(x)(1-\pi)} \\
&= \frac{1}{1 + \frac{f_{\mu_0}(x)(1-\pi)}{f_{\mu_1}(x)\pi}}
\end{aligned}$$

Let's look at  $\frac{f_{\mu_0}(x)(1-\pi)}{f_{\mu_1}(x)\pi}$ .

$$\begin{aligned}
\frac{f_{\mu_0}(x)(1-\pi)}{f_{\mu_1}(x)\pi} &= \frac{1-\pi}{\pi} \frac{\frac{1}{2\pi\sqrt{\det \Sigma}} \exp(-\frac{1}{2}(x-\mu_0)^T \Sigma^{-1}(x-\mu_0))}{\frac{1}{2\pi\sqrt{\det \Sigma}} \exp(-\frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1))} \\
&= \frac{1-\pi}{\pi} \exp(-\frac{1}{2}(x-\mu_0)^T \Sigma^{-1}(x-\mu_0) + \frac{1}{2}(x-\mu_1)^T \Sigma^{-1}(x-\mu_1)) \\
&= \frac{1-\pi}{\pi} \exp(-\frac{1}{2}(x^T \Sigma^{-1} x + \mu_0^T \Sigma^{-1} \mu_0) + \mu_0^T \Sigma^{-1} x + \frac{1}{2}(x^T \Sigma^{-1} x + \mu_1^T \Sigma^{-1} \mu_1) - \mu_1^T \Sigma^{-1} x) \\
&= \frac{1-\pi}{\pi} \exp((\mu_0 - \mu_1)^T \Sigma^{-1} x + \frac{1}{2}(\mu_1^T \Sigma^{-1} \mu_1 - \mu_0^T \Sigma^{-1} \mu_0)) \\
&= \exp((\mu_0 - \mu_1)^T \Sigma^{-1} x + \frac{1}{2}(\mu_1^T \Sigma^{-1} \mu_1 - \mu_0^T \Sigma^{-1} \mu_0) + \log(\frac{1-\pi}{\pi})) \\
&= \exp(-(\omega^T x + b))
\end{aligned}$$

Where  $\omega = \Sigma^{-1}(\mu_1 - \mu_0)$

And  $b = -\frac{1}{2}(\mu_1^T \Sigma^{-1} \mu_1 - \mu_0^T \Sigma^{-1} \mu_0) - \log(\frac{1-\pi}{\pi})$

Therefore we have :

$$\begin{aligned}
\mathbb{P}(Y = 1|X = x) &= \frac{1}{1 + \frac{f_{\mu_0}(x)(1-\pi)}{f_{\mu_1}(x)\pi}} \\
&= \frac{1}{1 + \exp(-(\omega^T x + b))} \\
&= \sigma(\omega^T x + b)
\end{aligned}$$

Which is similar to the form of the logistic regression.

c. The MLE has been implemented (cf. the Jupyter Notebook file *MVA DM1 Ariane ALIX Sacha BOZOU.ipynb*), applied to the datasets, and used to plot a decision boundary corresponding to  $p(y = 1|x) = 0.5$ :

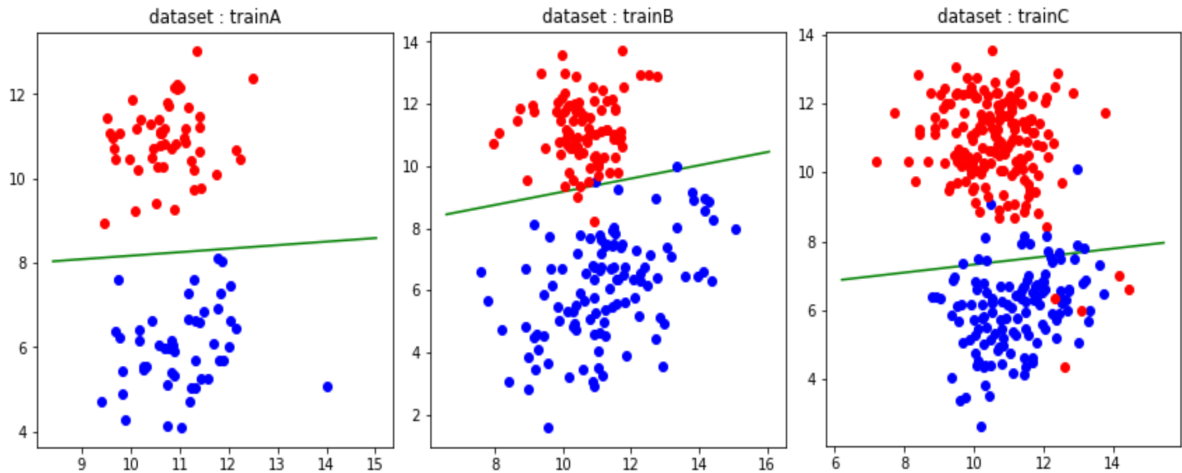


Figure 1 – Point cloud and decision boundary for the LDA

## 2.2 Logistic regression

In both logistic and linear regressions, we will use offset reparametrization.

a. For the logistic regression, we assume that :

$$\ln\left(\frac{\mathbb{P}(Y = 1|X = x)}{\mathbb{P}(Y = 0|X = x)}\right) = \omega^T x$$

Equivalently :

$$\mathbb{P}(Y = 1|X = x) = \sigma(\omega^T x)$$

with  $\sigma$  the sigmoid function

We iterate on the training data and we follow :

$$\omega^{new} = \omega^{old} + (X^T D_{\eta^{old}}^{-1} X^T (Y - \eta^{old}))$$

where :  $\eta_i = \sigma(\omega^T x_i)$

and :  $D_{\eta} = \text{Diag}(\eta_i(1 - \eta_i))$

The decision boundary is defined by  $\omega^T x = 0$

We give here after the numerical values for the parameters  $w$  and  $b$  learnt by the model on the different datasets.

- for trainA :  $w : 14.97, -59.05 ; b : 339.41$
- for train B :  $w : 1.84, -3.71 ; b : 13.43$
- for train C :  $w : -0.28, -1.91 ; b : 18.81$

b. The model has been implemented (cf. the Jupyter Notebook), applied to the datasets, and used to plot a decision boundary corresponding to  $p(y = 1|x) = 0.5$ :

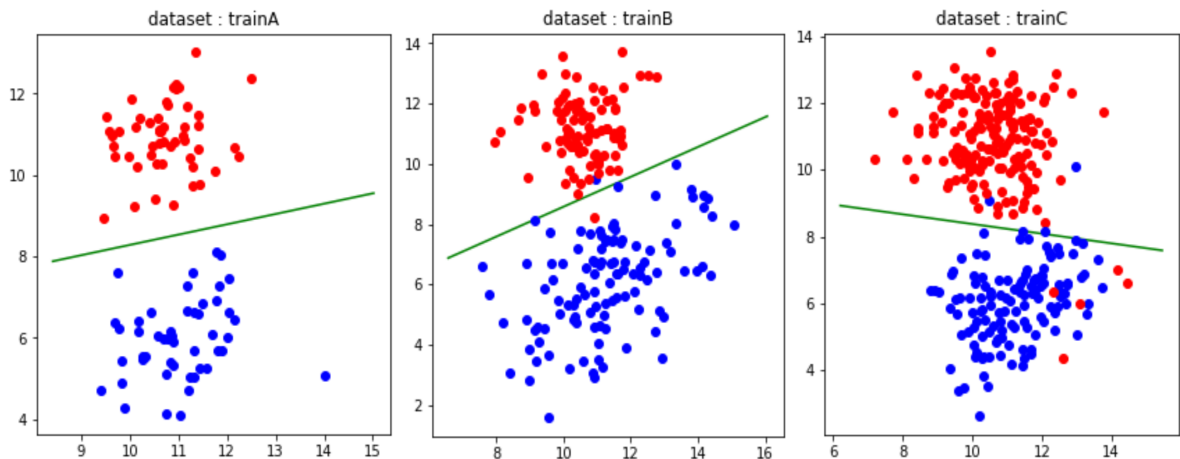


Figure 2 – Point cloud and decision boundary for the Logistic regression

## 2.3 Linear Regression

a. The probabilistic linear regression can be written with a noise  $\epsilon$  as :

$$Y = w^T X + \epsilon \quad \text{with} \quad \epsilon \sim \mathcal{N}(0, \sigma^2)$$

And we have the probability law:

$$p(y|x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(y - w^T x)^2}{2\sigma^2}\right)$$

The normal equation aims to minimize the cost defined by  $\frac{1}{2\sigma^2} \|Xw - y\|_2^2 = (Xw - y)^T (Xw - y)$  with regard to  $w$ . Looking at its derivative in  $w$ , we have :

$$\begin{aligned} \frac{\partial \text{cost}}{\partial w}(w) = 0 &\Leftrightarrow 2X^T Xw - 2X^T y = 0 \\ &\Leftrightarrow X^T Xw = X^T y \\ &\Leftrightarrow (X^T X)^{-1} X^T Xw = (X^T X)^{-1} X^T y \\ &\Leftrightarrow w = (X^T X)^{-1} X^T y \end{aligned}$$

Therefore  $\hat{w}_{MLE} = (X^T X)^{-1} X^T y$ .

Moreover,  $\hat{\sigma}_{MLE}^2$  should minimize the log-likelihood defined by :

$$\frac{1}{2} n \log(2\pi\sigma^2) + \frac{1}{2} \sum_{i=1}^n \frac{(y_i - w^T x_i)^2}{\sigma^2}$$

Hence,  $\hat{\sigma}_{MLE}^2 = \frac{1}{n} \sum_{i=1}^n (y_i - \hat{w}_{MLE}^T x_i)^2$

For the decision boundary, we have :

$$\begin{aligned} p(y = 1|x) = 0.5 &\Leftrightarrow \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(1 - w^T x)^2}{2\sigma^2}\right) = 0.5 \\ &\Leftrightarrow -\frac{1}{2} \log(2\pi\sigma^2) - \frac{(1 - w^T x)^2}{2\sigma^2} = \log\left(\frac{1}{2}\right) \\ &\Leftrightarrow \frac{(1 - w^T x)^2}{2\sigma^2} = \frac{1}{2} \log\left(\frac{2}{\pi\sigma^2}\right) \\ &\Leftrightarrow (1 - w^T x) = \sqrt{\sigma^2 \log\left(\frac{2}{\pi\sigma^2}\right)} \\ &\Leftrightarrow w^T x - 1 + \sqrt{\sigma^2 \log\left(\frac{2}{\pi\sigma^2}\right)} = 0 \end{aligned}$$



We give here after the numerical values for the parameters  $w$ ,  $b$  and  $\sigma^2$  (the variance of the noise) learnt by the model on the different datasets.

- for trainA :  $w : 0.06$  ,  $-0.18$  ;  $b : 1.38$  ;  $\sigma^2=0.03$
- for train B :  $w : 0.08$  ,  $-0.15$  ;  $b : 0.88$  ;  $\sigma^2=0.05$
- for train C :  $w : 0.02$  ,  $-0.16$  ;  $b : 1.64$  ;  $\sigma^2=0.06$

**b.** The model has been implemented (cf. the Jupyter Notebook), applied to the datasets, and used to plot a decision boundary corresponding to  $p(y = 1|x) = 0.5$ :

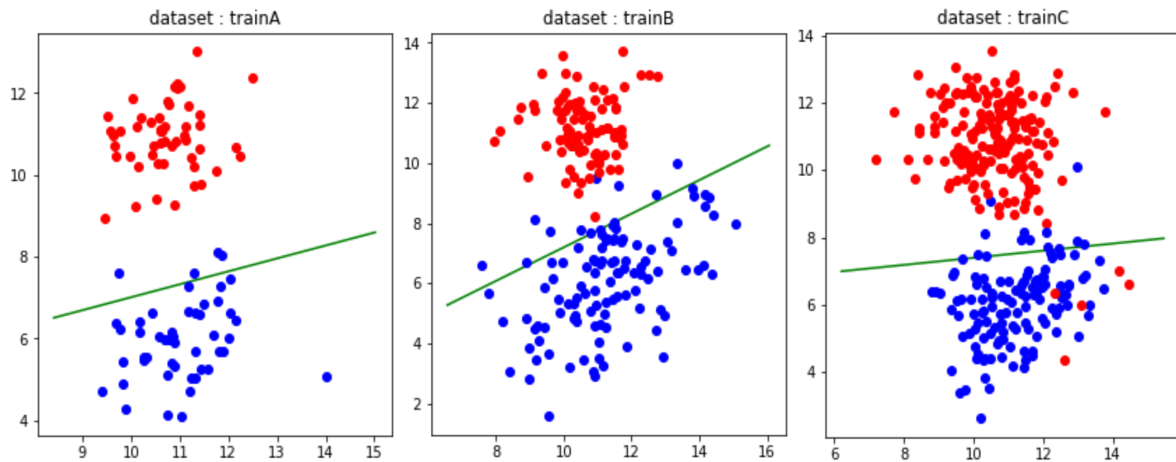


Figure 3 – Point cloud and decision boundary for the Linear regression

## 2.4 Application

**a.**

We have (cf computation on Jupyter notebook)

- MLE on Generative Model LDA  
 Error of classification for dataset trainA : 0.00%  
 Error of classification for dataset testA : 1.00%  
  
 Error of classification for dataset trainB : 2.00%  
 Error of classification for dataset testB : 4.00%  
  
 Error of classification for dataset trainC : 6.33%  
 Error of classification for dataset testC : 7.33%
- Logistic Regression

Error of classification for dataset trainA : 0.00%

Error of classification for dataset testA : 1.00%

Error of classification for dataset trainB : 1.00%

Error of classification for dataset testB : 3.50%

Error of classification for dataset trainC : 3.00%

Error of classification for dataset testC : 4.67%

- Linear Regression

## 2.5 Generative model (LDA)

The Maximum Likelihood Estimators for the parameters are the same as in the LDA, except for  $\Sigma$ .

We now have this log-likelihood:

$$\begin{aligned} l(\pi, \mu_0, \mu_1, \Sigma_0, \Sigma_1) = & \sum_{i=1}^n (y_i \log \pi + (1 - y_i) \log(1 - \pi)) \\ & + (1 - y_i)(f_{\mu_0}(x_i)) \\ & + y_i(f_{\mu_1}(x_i)) \end{aligned}$$

Where  $f_{\mu_{y_i}}(x_i) = \frac{1}{2\pi\sqrt{\det \Sigma_{y_i}}} \exp(-\frac{1}{2}(x - \mu_{y_i})^T \Sigma_{y_i}^{-1}(x - \mu_{y_i}))$ .

And we already have the MLEs :

$$\hat{\pi} = \frac{\sum_{i=1}^n y_i}{n}$$

$$\hat{\mu}_0 = \frac{\sum_{i=1}^n x_i(1 - y_i)}{\sum_{i=1}^n (1 - y_i)}$$

$$\hat{\mu}_1 = \frac{\sum_{i=1}^n x_i y_i}{\sum_{i=1}^n y_i}$$

We now have to find  $\hat{\Sigma}_0$  and  $\hat{\Sigma}_1$ .

$$\begin{aligned}
& \frac{\partial l}{\partial \Sigma_0^{-1}}(\pi, \mu_0, \mu_1, \Sigma_0, \Sigma_1) = 0 \\
& \Leftrightarrow \frac{\partial}{\partial \Sigma_0^{-1}} \left( \sum_{i=1}^n \frac{(1-y_i)}{2} (\log(\det \Sigma_0^{-1}) - \text{Tr}((x_i - \mu_0)^T \Sigma_0^{-1} (x_i - \mu_0))) \right) = 0 \\
& \Leftrightarrow \sum_{i=1}^n \frac{1-y_i}{2} (\Sigma_0 - (x_i - \mu_0)(x_i - \mu_0)^T) = 0 \\
& \Leftrightarrow \sum_{i=1}^n (1-y_i) \Sigma_0 = \sum_{i=1}^n (1-y_i) (x_i - \mu_0)(x_i - \mu_0)^T \\
& \Leftrightarrow \hat{\Sigma}_0 = \frac{1}{\sum_{i=1}^n (1-y_i)} \sum_{i=1}^n (1-y_i) (x_i - \mu_0)(x_i - \mu_0)^T
\end{aligned}$$

And with the same reasoning we get :

$$\hat{\Sigma}_1 = \frac{1}{\sum_{i=1}^n y_i} \sum_{i=1}^n y_i (x_i - \mu_1)(x_i - \mu_1)^T$$

Follwing the same logic as for the boundary of the LDA :

$$\mathbb{P}(Y = 1 | X = x) = \frac{1}{1 + \frac{f_{\mu_0}(x)(1-\pi)}{f_{\mu_1}(x)\pi}}$$

With :

$$\begin{aligned}
\frac{f_{\mu_0}(x)(1-\pi)}{f_{\mu_1}(x)\pi} &= \frac{1-\pi}{\pi} \frac{\frac{1}{2\Pi\sqrt{\det \Sigma_0}} \exp(-\frac{1}{2}(x - \mu_0)^T \Sigma_0^{-1} (x - \mu_0))}{\frac{1}{2\Pi\sqrt{\det \Sigma_1}} \exp(-\frac{1}{2}(x - \mu_1)^T \Sigma_1^{-1} (x - \mu_1))} \\
&= \frac{(1-\pi)\sqrt{\det \Sigma_1}}{\pi\sqrt{\det \Sigma_0}} \exp(-\frac{1}{2}(x - \mu_0)^T \Sigma_0^{-1} (x - \mu_0) + \frac{1}{2}(x - \mu_1)^T \Sigma_1^{-1} (x - \mu_1)) \\
&= \exp(-(x^T K x + \omega^T x + b))
\end{aligned}$$

Where :  $\omega = \Sigma_1^{-1} \mu_1 - \Sigma_0^{-1} \mu_0$

And  $b = -\frac{1}{2}(\mu_1^T \Sigma_1^{-1} \mu_1 - \mu_0^T \Sigma_0^{-1} \mu_0) - \log(\frac{(1-\pi)\sqrt{\det \Sigma_1}}{\pi\sqrt{\det \Sigma_0}})$

And  $K = \frac{1}{2}\Sigma_0^{-1} - \frac{1}{2}\Sigma_1^{-1}$

The decision boundary is defined by :

$$x^T K x + \omega^T x + b = 0$$

a. We implemented the new model in the Jupyter Notebook.

We give here after the numerical values for the parameters  $w$ ,  $b$  and  $K$  learnt by the model on the different datasets.

- for trainA :  $w : -9.38, -5.91 ; b : 79.95 ; K = \begin{vmatrix} 0.51 & -0.07 \\ -0.07 & 0.33 \end{vmatrix}$
- for train B :  $w : -5.07, -3.97 ; b : 46.86 ; K = \begin{vmatrix} 0.25 & -0.01 \\ -0.01 & 0.18 \end{vmatrix}$
- for train C :  $w : -5.07, -4.01 ; b : 46.38 ; K = \begin{vmatrix} 0.19 & 0.05 \\ 0.05 & 0.13 \end{vmatrix}$

c. We observe:

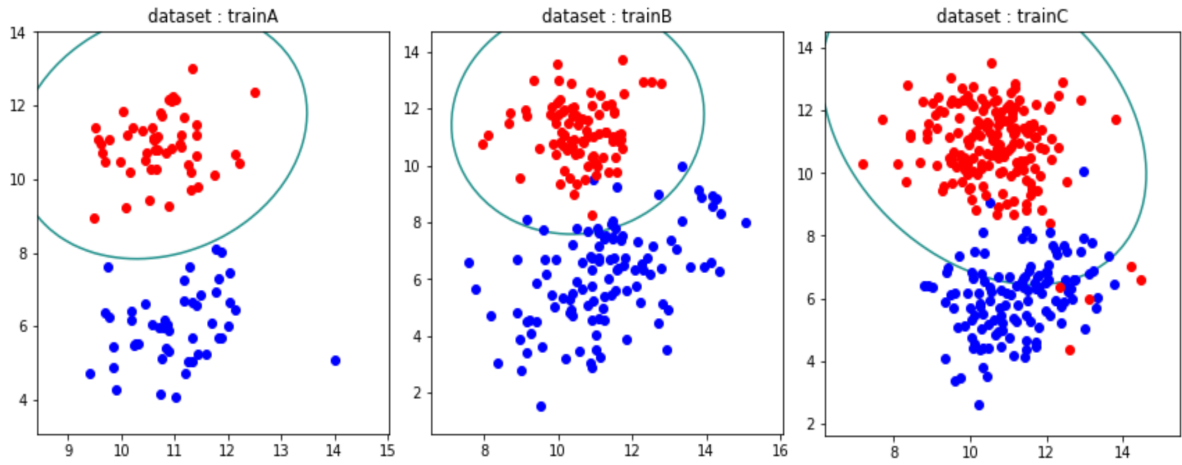


Figure 3 – Point cloud and decision boundary for the Linear regression

c. We get the following missclassification errors :

Error of classification for dataset trainA : 0.00%

Error of classification for dataset testA : 2.00%

Error of classification for dataset trainB : 5.50%

Error of classification for dataset testB : 7.00%

Error of classification for dataset trainC : 14.00%

Error of classification for dataset testC : 16.67%

d.