

Assessing the Impact of the 2017 PPPs on the International Poverty Line and Global Poverty

Dean Jolliffe

Daniel Gerszon Mahler

Christoph Lakner

Aziz Atamanov

Samuel Kofi Tetteh-Baah



WORLD BANK GROUP

Development Data Group &
Poverty and Equity Global Practice
February 2022

Abstract

Purchasing power parity exchange rates (PPPs) are used to estimate the international poverty line (IPL) in a common currency and account for relative price differences across countries when measuring global poverty. This paper assesses the impact of the 2017 PPPs on the nominal value of the IPL and global poverty. The analysis indicates that updating the \$1.90 IPL in 2011 PPP dollars to 2017 PPP dollars results in an IPL of approximately \$2.15—a finding that is robust to various methods and assumptions. Based on an

updated IPL of \$2.15, the global extreme poverty rate in 2017 falls from the previously estimated 9.3 to 9.1 percent, reducing the count of people who are poor by 15 million. This is a modest change compared with previous updates of PPP data. The paper also assesses the methodological stability between the 2011 and 2017 PPPs, scrutinizes large changes at the country level, and analyzes higher poverty lines with the 2017 PPPs.

This paper is a product of the Development Data Group, Development Economics and the Poverty and Equity Global Practice. It is part of a larger effort by the World Bank to provide open access to its research and make a contribution to development policy discussions around the world. Policy Research Working Papers are also posted on the Web at <http://www.worldbank.org/prwp>. The authors may be contacted at djolliffe@worldbank.org, dmahler@worldbank.org, clakner@worldbank.org, aatamanov@worldbank.org, and stettehbaah@worldbank.org.

The Policy Research Working Paper Series disseminates the findings of work in progress to encourage the exchange of ideas about development issues. An objective of the series is to get the findings out quickly, even if the presentations are less than fully polished. The papers carry the names of the authors and should be cited accordingly. The findings, interpretations, and conclusions expressed in this paper are entirely those of the authors. They do not necessarily represent the views of the International Bank for Reconstruction and Development/World Bank and its affiliated organizations, or those of the Executive Directors of the World Bank or the governments they represent.

Assessing the Impact of the 2017 PPPs on the International Poverty Line and Global Poverty¹

Dean Jolliffe (r) *Daniel Gerszon Mahler* (r) *Christoph Lakner* (r)

Aziz Atamanov (r) *Samuel Kofi Tetteh-Baah*

JEL code: I32 - Measurement and Analysis of Poverty

Keywords: 2017 PPPs, global poverty, the international poverty line

¹ The author ordering was constructed through American Economic Association's randomization tool (confirmation code: R-TVbwJ99ttd). Corresponding author: Christoph Lakner (clakner@worldbank.org). All authors are with the World Bank. We would like to thank the staff of the International Comparison Program Global Office at the World Bank, particularly Nada Hamadeh, Marko Rissanen, William Vigil Oliver, Maurice Nsabimana and Mizuki Yamanaka for providing the PPP data and metadata, patiently answering our technical questions, and for their feedback on earlier drafts. We are thankful for helpful comments and guidance from Laurence Chandy, Stefan Dercon, Francisco Ferreira, Deon Filmer, Haishan Fu, Ana Revenga, Carolina Sanchez-Paramo, and Umar Serajuddin. Many thanks are also due to several World Bank poverty economists and members of the World Bank's Global Poverty Working Group (GPWG) for their feedback, as well as to Judy Yang who was involved in the earlier stages of this project. We would also like to thank Angus Deaton, Bob Allen, and members of the International Comparison Program's Technical Advisory Group for comments. Finally, we would like to thank the audience at the ECINEQ 2021 conference. This paper is an expanded version of Atamanov et al. (2020), in which we analyzed the impact of the revised 2011 PPPs and 2017 PPPs on global poverty. In this paper, we expand the analysis of the impact of the 2017 PPPs on the derivation of global poverty lines, analyze whether the ICP methodology has stabilized between the 2011 and 2017 ICP rounds, and examine country-level changes in greater detail. We gratefully acknowledge financial support from the Knowledge for Change program and from the UK government through the Data and Evidence for Tackling Extreme Poverty (DEEP) Research Programme. The findings, interpretations, and conclusions expressed in this paper are entirely those of the authors. They do not necessarily represent the views of the World Bank and its affiliated organizations, or those of the Executive Directors of the World Bank or the governments they represent.

1. Introduction

The ambition to eliminate extreme poverty around the world by 2030 is at the forefront of development policy. It is the first of the Sustainable Development Goals (SDGs), as well as the World Bank’s Twin Goals. Given this emphasis on poverty within the development community, it is important to accurately measure the number of poor people in the world and know where they live. Some of the most challenging issues regarding the estimation of global poverty concern the construction of the international poverty line (IPL)—which is derived from the national poverty lines of some of the poorest countries in the world (Ferreira et al. 2016; Ravallion et al. 2009)—and the comparability of living standards across countries. The principal data that allow for the comparability of national poverty lines and different countries’ living standards are purchasing power parities (PPPs). For this reason, the use of PPPs for the measurement of global poverty, and the revisions to PPPs over time, have been the subject of considerable debate (Deaton and Heston 2010; Ravallion 2014; Kakwani and Son 2016; Deaton and Aten 2017; Inklaar and Rao 2017).

In May 2020, the International Comparison Program (ICP) published new 2017 PPPs.² The 2017 PPPs reflect the most recent relative price differences across a wide range of countries around the world. Differences between PPP rounds may not only reflect changes in real cost-of-living differences across countries over time but also ICP methodological changes across rounds. For example, the 2005 ICP round overstated price levels in most countries in the developing world, particularly owing to its approach of linking regional prices through “ring” countries (Deaton and Aten 2017). This contributed to increasing the estimated number of extremely poor people in the world by 400 million compared to the earlier 1993 PPPs (Chen and Ravallion 2010).

An important question is whether the new 2017 PPP revisions reflect new information on relative price differences around the world or are driven by changes in ICP’s methodology. With the background of large revisions in earlier ICP rounds being driven by methodological changes, the *Atkinson Commission on Global Poverty* (World Bank 2017) raised concerns that adopting future ICP rounds for global poverty monitoring might “shift the goalposts” and undermine the credibility of the World Bank’s efforts in monitoring global poverty until 2030, the target date for the poverty elimination goal. In its response to the Commission, the World Bank left open the possibility that the 2017 PPPs would be adopted if the ICP methods and price data collection had stabilized (World Bank 2016).

² The ICP also published revised 2011 PPPs in May 2020. This was the first time the ICP revised PPP estimates it had already published. The revisions to the 2011 PPPs reflect changes made to national accounts expenditures that are used as weights to aggregate elementary PPPs. The revised 2011 PPPs had only a small effect on the global poverty rates, relative to the original 2011 PPPs (see Atamanov et al. (2020) and Appendix A for more details). In the rest of this paper, we use the revised 2011 PPPs as the baseline, since they are currently being used for the global poverty measures reported by the World Bank.

With this background, we analyze the 2017 PPPs to make three contributions to the literature on global poverty measurement. First, we provide empirical evidence that is consistent with stability between the 2011 and 2017 rounds when methodological stability is understood at the global level (see Deaton and Schreyer (2022) for further evidence to that end). Yet for certain countries, we find that the 2017 PPPs yield large differences to estimated poverty. In several cases, this is because of country-level improvements in the data collection.

Second, we derive global poverty lines with the 2017 PPPs. We follow Jolliffe and Prydz (2016) by calculating the IPL as the median national poverty line of low-income countries, which yields a line of \$2.15 in 2017 PPP, compared with \$1.90 in 2011 PPP. We show that this value is robust to varying the set of countries and set of poverty lines, and hence reliably reflects the typical standard by which the poorest countries of the world judge their citizens to be impoverished.

In keeping with Atkinson's concerns about not shifting the goalpost, we show that the \$2.15 IPL in 2017 PPP largely keeps the global poverty rate constant in 2010, 2012, and 2015, compared to the 2011 PPPs. These years contained the latest available global poverty estimates when the World Bank's goal of ending extreme poverty was set, when the United Nations agreed on the SDGs, and when Atkinson was writing his report, respectively.

Beyond the IPL, we also compute with the 2017 PPPs the higher poverty lines the World Bank uses to monitor poverty in countries with a low incidence of extreme poverty—\$3.20 and \$5.50 in 2011 PPPs. These higher poverty lines reflect the median national poverty line of lower-middle-income countries and upper-middle-income countries, respectively. With the 2017 PPPs, we compute these lines to be approximately \$3.65 and \$6.85 per person per day. We also update the Bank's societal poverty line (SPL) originally defined as $\max(\$1.90, \$1 + 50\% \text{ of median consumption})$ in 2011 PPPs to $\max(\$2.15, \$1.15 + 50\% \text{ of median consumption})$ in 2017 PPPs.

Third, we analyze the impact on the global and regional poverty estimates of the World Bank of using these global poverty lines and the 2017 PPPs. The 2017 PPPs would slightly increase historical estimates of extreme poverty and slightly decrease extreme poverty since 2014, compared with the 2011 PPPs. Extreme poverty would decrease marginally by 0.2 percentage points (pp) in 2017 with the 2017 PPPs. Compared to past adoptions of new PPP rounds, this is a minor change. The global count of extreme poor reduces by 15 million, which is largely driven by 34 million fewer poor people in Sub-Saharan Africa, while poverty increases slightly in all other regions. However, Sub-Saharan Africa remains the region with the largest share of extreme poor in 2017; from 62% with the revised 2011 PPPs to 58% with the 2017 PPPs. Nigeria accounts for almost half of the reduction in extreme poverty in Sub-Saharan Africa.

There is also a limited change in global poverty with the revised poverty line of \$3.65 in 2017 PPP, as declines in Sub-Saharan Africa and the Middle East & North Africa are offset by increases in East Asia & the Pacific and Europe & Central Asia. The revision of the \$5.50 line (2011 PPP)

to \$6.85 (2017 PPP) increases poverty markedly globally by 4.3pp. This increase is seen in all regions with about two-thirds of the added poor living in China, India, and Indonesia alone. With the updated parameters of the SPL, about 17.6 million more people would be living in societal poverty in the world in 2017. Societal poverty slightly increases in all regions, except Sub-Saharan Africa and the Middle East & North Africa.

The rest of the paper is organized as follows. Section 2 describes the data we use. Section 3 assesses the stability of the 2017 PPPs, relative to the 2005 and 2011 PPPs. Section 4 proposes updates to the IPL, the higher global poverty lines, and the societal poverty line in terms of the 2017 PPPs. Section 5 documents the changes observed in global, regional, and country-level poverty estimates with the 2017 PPPs. Section 6 discusses the drivers of the changes observed in the global poverty estimates as well as the global poverty lines. Section 7 concludes.

2. Data

We utilize different types of data that are typically used in estimating global poverty. The data can be categorized into five different headings: (a) consumption or income distributions, (b) national poverty lines and national poverty rates, (c) CPIs, (d) PPPs and market exchange rates, and (e) GDP per capita.

(a) Consumption or income distributions: We use micro or grouped data from 1,971 income distributions in PovcalNet, which is the World Bank's database for global poverty estimates.³ For most countries, the distributions use micro data from nationally representative household surveys. For a few countries, including China, the distributions are constructed from grouped data (i.e., quantiles of the distribution such as percentiles). The distributions cover 168 economies that represent over 97% of the world's population. Some economies have more surveys than others depending on how often surveys are conducted. There is an average of 10 surveys for all economies spanning the period between 1991 and 2017, but more advanced economies are more likely to have annual distributions. For economies without annual data, annual poverty estimates are produced by extrapolating forward or backward from surveys, or interpolating between surveys, using national accounts growth rates and assuming no distributional changes (Prydz et al. 2019). While this is a strong assumption to impose, it performs relatively well (Mahler et al. 2021) and has the advantage that the same set of countries can be compared over time.

(b) National poverty lines and poverty rates: We estimate global poverty lines using national poverty lines and national poverty rates covering a wide range of countries. We obtain a total of 1,438 country-year observations of national poverty rates from the World Bank's Poverty and

³ PovcalNet is available at <http://iresearch.worldbank.org/PovcalNet/home.aspx>. PovcalNet includes surveys that measure consumption or income. For simplicity, we refer to them interchangeably as income distributions. This paper is based on the June 2021 vintage of PovcalNet.

Equity Database and the Organization for Economic Development and Co-operation (OECD) (see Table B.1 for more details). In addition, we use the national poverty lines of 15 poor countries that Ferreira et al. (2016) estimated, which Ravallion et al. (2009) originally derived.

(c) Consumer price indices (CPIs): CPI data are required to adjust income distributions to obtain a measure of household welfare in real terms. CPI data are also required to inflate national poverty lines to the ICP reference year. We use the CPI data used in PovcalNet (Lakner et al. 2018). The main source of CPI data is the International Monetary Fund's International Financial Statistics (IFS).⁴ For a few countries, including Bangladesh, Ghana, the Lao People's Democratic Republic, Malawi, and Tajikistan, PovcalNet uses inflation rates imputed from household survey data, because they are deemed to be more reliable than CPIs from standard sources (Lakner et al. 2018).

(d) Purchasing power parity conversion factors (PPPs) and market exchange rates: When estimating global poverty, both income distributions and national poverty lines are converted into a common, internationally comparable currency unit using exchange rates. PPP conversion factors are preferred to market exchange rates since market exchange rates only equilibrate the relative prices of tradable goods across countries. PPP conversion factors incorporate the relative prices of non-tradable services (e.g., getting a haircut) across countries, which tend to be less expensive in developing countries, where labor costs tend to be lower (a phenomenon known as the Balassa-Samuelson-Penn effect). PPP conversion factors are price indices that express the currency units of one country in terms of the currency units of a reference country, typically the United States. In other words, PPPs measure how much it costs to purchase a basket of goods and services in one country compared to how much it costs to purchase the same basket of goods and services in the United States. We obtain estimates of PPP and market exchange rates from the ICP. We use PPP estimates from the 2005, 2011 (original and revised), and 2017 rounds. We use the PPPs for final household consumption expenditure, which includes non-profit institutions serving households (NPISHs).

For the 2011 PPP estimates, we do not use the official PPP estimates for six countries, namely the Arab Republic of Egypt, Iraq, Jordan, Lao PDR, Myanmar, and the Republic of Yemen, following Ferreira et al. (2016) and PovcalNet. There have been concerns over the official PPP estimates for these countries due to large differences in the price changes implied by the 2005 and 2011 PPPs and domestic CPI inflation (Ferreira et al. 2016; Atamanov et al. 2018).⁵ A seemingly unrelated regressions (SUR) model—similar to what the ICP uses for countries without a PPP (so called non-benchmark economies)—was then used to impute 2011 PPPs for these countries.⁶ We re-estimate the imputed PPPs for these six countries, using revised 2011 PPPs and updated input data (Atamanov et al. 2020) (see Table B.2 for more details). For 2017, we use the official 2017 PPPs

⁴ For this paper, we use monthly CPI series from the November 2020 IFS vintage averaged into annual CPI series.

⁵ We have verified this information with the revised 2011 PPPs (Atamanov et al. 2020).

⁶ Countries that collect prices to compute PPPs as part of the ICP exercise are referred to as benchmark countries. For non-benchmark countries, there is no price collection and PPPs are imputed using a cross-country relationship.

for most economies, as published by the ICP. We use alternative 2017 PPPs for 14 exception countries (see Section 3.B below and Table B.3 for more details).

Citing urban bias in the ICP price data collection, previous studies have imputed subnational (rural/urban) PPPs for the three most populous countries in the developing world, namely China, India, and Indonesia, when estimating global poverty (Chen and Ravallion 2008; 2010; Ferreira et al. 2016). We follow these studies and impute rural and urban revised 2011 PPPs and 2017 PPPs for these three countries, using the share of price collection outlets in urban areas, the ratio of urban to rural poverty lines in the ICP reference year, and the national PPP as inputs (see Table B.4 for more details).

(e) Gross domestic product (GDP) per capita: We use data on GDP per capita, denominated in 2011 PPP, 2017 PPP, and current USD, for the analyses done in this paper. We obtained these data from the June 2021 vintage of the World Development Indicators (WDI).

3. Assessing the 2017 PPPs

A. Examining stability between the 2011 and 2017 PPPs at the global level

The ICP has estimated purchasing power parity (PPP) conversion factors for different reference years—so far, 1970, 1973, 1975, 1980, 1985, 1993, 2005, 2011, and 2017. The ICP methodology has evolved, resulting in improvements in the coverage and quality of PPP estimates.

Broadly speaking, there are five main components of the ICP methodology by which stability between the 2011 and 2017 rounds can be assessed (more details are provided in Appendix C):

- First, *the adoption of a common classification of expenditures across countries*. For this, the ICP relies on the final expenditures classification on GDP from the System of National Accounts (SNA). While the 2005 and original 2011 rounds were based on the 1993 SNA, the revised 2011 and 2017 rounds are based on the 2008 SNA. This means that there has not been any drastic change in the classification of expenditures from the revised 2011 round to the 2017 round.
- Second, *the selection of the basket of goods and services that will be priced*. This basket is based on the classification of expenditures mentioned above. Though the 2008 SNA were used for both the 2011 and 2017 rounds, a few changes have been introduced into the ICP expenditure classification at the basic-heading level.⁷ Since only aggregate PPPs (e.g., household final consumption or GDP) are relevant for our purposes, these changes at the

⁷ Basic headings in the ICP nomenclature refer to detailed expenditure categories containing similar item varieties. For example, “Rice” basic heading contains several rice varieties. The changes at the basic-heading level are more structural than substantive (e.g., “opening value of inventories” and “closing value of inventories” in the 2011 round have been merged as “changes in inventories” in the 2017 round).

sub-component level do not impact the PPPs we use, and hence this step does not influence the stability of the 2011 and 2017 PPPs for poverty monitoring purposes.

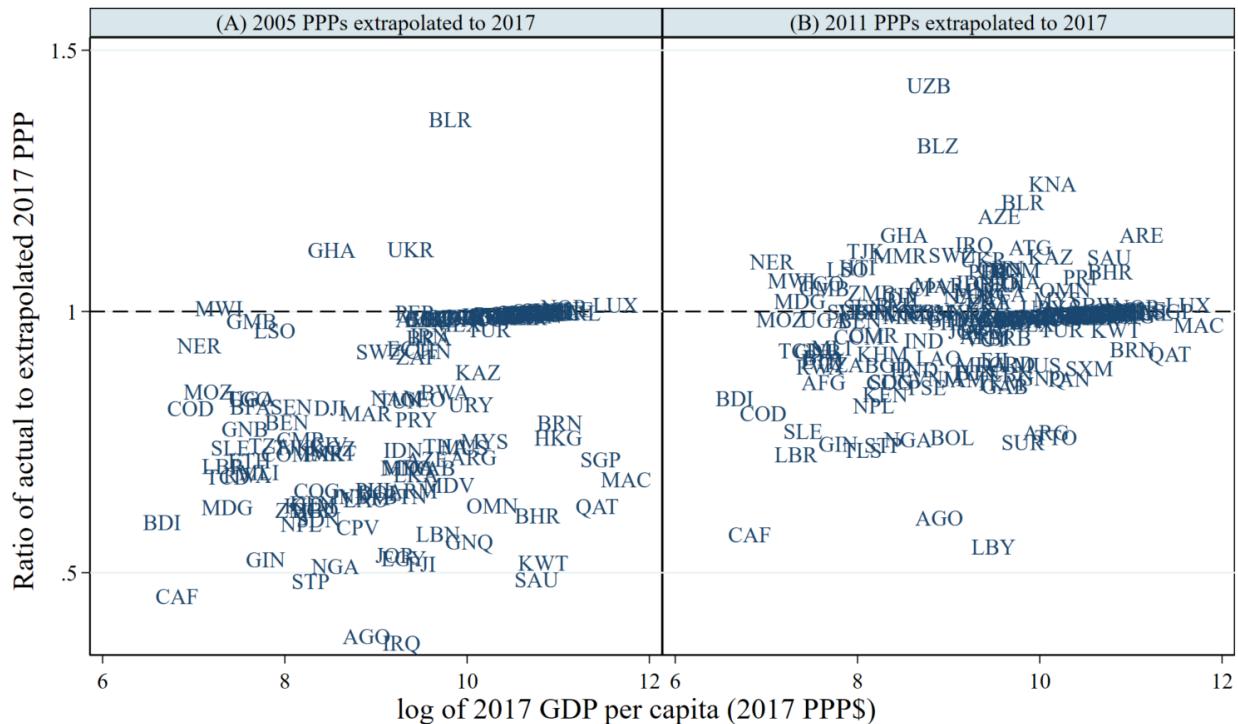
- Third, *the computation of the PPPs*. While there was a major change in the method used in linking prices within and between regions from the “ring” method in 2005 to the “Global Core List” method in 2011, the computational steps followed to produce the final global PPPs have remained unchanged since the 2011 round. The Country Product Dummy (CPD) method is used to estimate PPPs at the basic-heading level, while the Gini, Eltetö and Köves and Szulc (GEKS) method is used to estimate aggregated PPPs, which are both transitive and base-country invariant.
- Fourth, *the participating countries*. The calculation of a country’s PPPs partially depends on other countries’ PPPs. The number of participating countries declined from 199 to 176 between the 2011 and 2017 rounds, which is mostly explained by the dropping out of several small Pacific Islands. However, since they were linked to the other participating economies via three bridge-countries (i.e., selected countries which additionally collect prices on a range of items from another region)—Australia (Eurostat-OECD), Fiji (Asia and the Pacific), and New Zealand (Eurostat-OECD)—the inclusion of these islands did not impact the global results.
- Fifth, *the spatial coverage of ICP price collection*. The spatial coverage may vary over time within countries. In the 2005 round, for instance, price survey outlets were predominantly in capital cities and urban areas (World Bank 2008). In India, 74% of prices were collected from urban areas in the 2011 round, while 64% of prices were collected from urban areas in the 2017 round. In OECD countries prices have been largely collected from urban areas in both rounds. These differences do not necessarily imply overestimation or underestimation of national prices, as the ICP requests countries to submit nationally representative prices.⁸ Under the assumption that countries submit nationally representative prices, changes to this component do not affect the stability of the 2011 and 2017 PPPs. The absence of documentation of what countries do makes it difficult to assess whether this is indeed the case.

In addition to examining the ICP methodology by these five components, we can use the *ratio of actual to extrapolated PPPs* to assess stability empirically (Deaton and Heston 2010; Deaton and Aten 2017). PPPs are extrapolated from the past ICP rounds using a factor equal to domestic inflation relative to US inflation, and compared to the actual PPP estimate. Deviations are expected from such a comparison, because of conceptual and estimation differences between CPIs and PPPs. Substantial, systematic deviations could, however, suggest issues with the PPPs, as shown with previous rounds. For example, Deaton and Aten (2017) use this to show that the ICP overstated prices in the 2005 round, especially in developing countries, which was reversed in the 2011 round.

⁸ National statistical offices apply temporal and spatial price indices to estimate nationally representative average prices, although it is impossible to know whether these adjustments result in prices that are truly nationally representative.

Figure 1 shows that the ratio of actual 2017 PPPs to 2017 PPPs extrapolated from the 2005 round is systematically lower than 1, while the ratio of actual 2017 PPPs to 2017 PPPs extrapolated from the revised 2011 round is more uniformly distributed around a mean of almost 1. The revised 2011 PPPs can better predict the 2017 PPPs. This is consistent with stability between the 2011 and 2017 PPPs, especially when compared with the 2005 PPPs.

Figure 1: Ratio of actual to extrapolated PPPs



Notes: This figure shows the ratio of actual to extrapolated PPPs as a function of GDP per capita. The actual PPPs are household final consumption PPPs from the 2017 round. The extrapolations are based on official PPPs from the 2005 or revised 2011 rounds together with CPIs mainly from the IMF International Financial Statistics (IFS). Eurostat PPPs have been used for OECD countries. Eurostat-OECD countries have a ratio of close to one, as they publish annual PPP numbers, and therefore see only minimal revisions at new rounds of the ICP. Panel (A) shows 2005 PPPs extrapolated to 2017. For the full sample of 143 observations, the ratio of actual to extrapolated PPPs has a mean of 0.822 with a standard deviation of 0.18. Panel (B) shows revised 2011 PPPs extrapolated to 2017. For the full sample of 187 observations, the ratio of actual to extrapolated PPPs has a mean of 0.975 with a standard deviation of 0.11. The difference in means between panels (A) and (B) is statistically significant.

In sum, we think that both when evaluating changes to the five components of how the PPPs are calculated, and when evaluating the ratio of actual to extrapolated PPPs, the evidence speaks in favor of the 2017 PPPs vis-à-vis the 2011 PPPs having stabilized at the global level.

B. Examining stability between the 2011 and 2017 PPPs at the country level

Though the ICP methodology may have stabilized and the changes between 2011 and 2017 are smaller than earlier rounds, the changes in poverty at the country level could still be large. This might result from different factors causing large changes in PPPs and/or CPIs. For example,

changes in PPPs can arise from: (i) geographical changes in ICP price collection (with an imperfect adjustment to national average prices); or (ii) changes in the share of items in the ICP price survey for which prices were actually collected (for example, Angola priced 18% of items in the 2011 round which almost doubled to 35% in the 2017 round). Furthermore, the 2017 ICP round has benefited from technological advances that have reportedly improved data quality assurance across all regions (World Bank 2020b). It is also important to note that the extrapolation of different PPPs to a common year requires the use of national CPIs, which may have measurement issues, such as following non-standard consumption classifications, or using expenditure weights and baskets that are outdated or not nationally representative (Berry et al. 2019).

Ferreira et al. (2016) summarized the changes in PPPs and CPIs using the *delta* ratio. The delta ratio is the ratio of the relative change in CPI between 2011 and 2017 to the relative change in PPP between 2011 and 2017. It equals the ratio of the PPP-denominated mean income levels, and is thus closely related to the change in poverty.⁹ See Appendix E, section 9 for the derivation of the delta ratio. The delta ratio is given by

$$\delta = \frac{CPI2017}{CPI2011} / \frac{PPP2017}{PPP2011} = \frac{Y_{PPP2017}}{Y_{PPP2011}}$$

where *CPI2017* is the CPI in 2017, *CPI2011* is the CPI in 2011, *PPP2017* is the 2017 PPP, and *PPP2011* is the (revised) 2011 PPP. *Y_{PPP2017}* is income in 2017 PPPs and *Y_{PPP2011}* is income in (revised) 2011 PPPs.

Since the PPPs are expressed in the US dollar (i.e., the numeraire), for the United States the delta ratio boils down to the relative change in CPI between 2011 and 2017 (i.e., $\delta = 1.0897$ or an inflation rate of 8.97%). For Angola, with a delta ratio of 1.80, the same bundle of goods and services in Angola that would cost \$1.00 in 2011 (converted by 2011 PPP into LCU) would cost \$1.80 in 2017 (converted by 2017 PPP into LCU). Thus, the Angolan Kwanza has appreciated in PPP terms relative to the US dollar, and it is less expensive to buy the same bundle of goods and services in Angola, hence poverty will reduce.

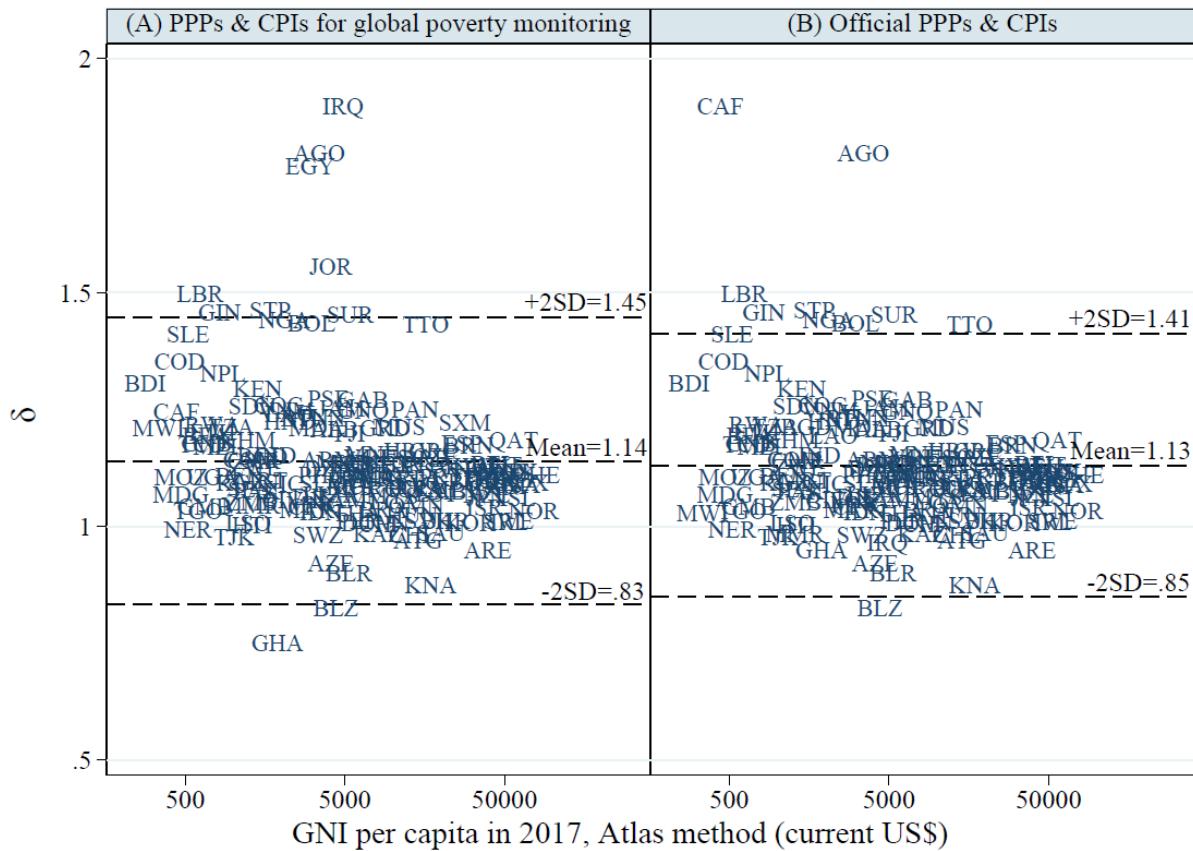
Large differences in the movements in PPPs and CPIs—outlier delta ratios—indicate large changes in poverty when moving to the 2017 PPPs. Ferreira et al. (2016) identified countries with delta ratios of two standard deviations from the mean as outliers. If the delta ratio for a country is an outlier, we do not know if this is caused by the 2011 PPP, 2017 PPP and/or the CPI. Put differently, a large change in the poverty rate between the revised 2011 and 2017 PPP does not necessarily mean the 2017 PPP-based poverty rate is wrong and the 2011 PPP-based poverty rate is more appropriate.

⁹ The delta ratio can also be expressed as the inverse of the actual to extrapolated PPPs multiplied by the US inflation between the two rounds.

Even if the data collection and methodology stabilized between any two rounds, the delta ratio would still deviate from the U.S inflation rate since CPIs and PPPs do not measure the same thing (Krijnse Locker and Faerber 1984; Dalgaard and Sørensen 2002; McCarthy 2013).

Figure 2 shows the delta ratio between 2011 and 2017 PPPs and indicates the mean and cut-offs outside of which a country is considered as an outlier. The mean delta ratio is 1.13, which incidentally is also the relative change in the IPL, which we derive in a later part of the paper (i.e., \$2.15 in 2017 PPP compared with \$1.90 in 2011 PPP). This means that extreme poverty decreases in countries with a delta ratio higher than 1.13 and increases in countries with a delta ratio less than 1.13.

Figure 2: Real changes in PPP-adjusted dollars between 2011 and 2017 PPPs (delta ratio)



Notes: The horizontal axis is on a log scale. In contrast to Figure 1, this graph includes only benchmark countries in both 2011 and 2017 ICP rounds. Panel (A) uses PPPs and CPIs currently used for global poverty monitoring, and panel (B) uses official PPPs (from the ICP) and CPIs (from IFS). Official 2017 PPPs are used in both cases.

Panels (A) and (B) use slightly different PPPs and CPIs, which results in somewhat different outliers to be identified. We use PPPs and CPIs for global poverty monitoring in panel (A), while

panel (B) uses those published by the ICP and IFS.¹⁰ The delta ratio is slightly higher on average with greater dispersion in panel (A) than in panel (B). Taken together, there are 14 outlier countries: Angola, Belize, Bolivia, Central African Republic, Egypt, Ghana, Guinea, Iraq, Jordan, Liberia, Nigeria, São Tomé and Príncipe, Suriname, and Trinidad and Tobago.

As done with the 2011 PPPs (Ferreira et al. 2016), for these outlier countries we have conducted a series of additional checks to decide whether to use the official PPPs or imputed PPPs for global poverty monitoring.¹¹ Apart from the delta criterion, we have also included a PPP residual criterion to identify outlier countries. We define the PPP residual as the log difference between published and predicted PPP (using the ICP's model for non-benchmark country PPP estimation), and select as an outlier any country whose residual is more than two standard deviations from the mean. The residual criterion is added to address the concern that the delta criterion fails to unambiguously identify whether there are issues with PPPs or CPIs. We select Egypt, Suriname, Central African Republic, Djibouti, and Sudan based on the residual criterion (three of which were already selected based on the delta criterion). We also include as outliers the countries whose revised 2011 PPPs were imputed by Ferreira et al. (2016) (Egypt, Iraq, Jordan, Lao PDR, and Myanmar).

For all countries selected as outliers, we consider a range of indicators that could provide some evidence on whether official or imputed 2017 PPPs are more appropriate for these countries.¹² To examine whether the quality of CPIs rather than PPPs might be the reason why a country is an outlier, we assess the CPI expenditure classification used, the reference year of the CPI weights, and the spatial coverage of the expenditure data used in deriving CPI weights (Berry et al. 2019). To examine the quality of the PPPs, we look at the changes in the share of items priced between the 2011 and 2017 rounds.¹³ We also compare 2017-PPP-based poverty rates with correlates of poverty that do not rely on PPPs, such as the multidimensional poverty index (MPI) from the Oxford Poverty and Human Development Initiative (OPHI) (Alkire et al. 2020) and age dependency ratio from the World Development Indicators (WDI). Last but not the least, we incorporate the assessments of the country poverty economists—World Bank staff specialized in the measurement of poverty in a country—who provide country-specific nuances. Appendix D has the details of all the steps we followed to identify and deal with outlier countries.

¹⁰ The differences are explained by the revised 2011 PPPs which are imputed for exception countries (Egypt, Iraq, Jordan, Lao PDR, and Myanmar) and survey-based CPIs used for some countries (such as Ghana). The Republic of Yemen, which also has an imputed 2011 PPP, drops from the analysis since it is not included in the 2017 ICP.

¹¹ Deaton (2001) similarly argues that with a new ICP round, the poverty estimates should be subjected to detailed, local scrutiny.

¹² We impute 2017 PPPs using the seemingly unrelated regressions (SUR) model the ICP uses to predict PPPs for non-benchmark economies, while excluding from the regression those economies which are subjected to greater scrutiny. With the revised 2011 PPPs, a slightly different model was used, which particularly addressed concerns in the Middle East & North Africa region (e.g., by including a conflict dummy in the model) (Atamanov et al. 2018). Since this region does not stand out this time around, we decided to use the ICP model without any adjustments.

¹³ Unfortunately, we cannot assess the changes in the geographical coverage of ICP price data collection between rounds, as we lack (reliable) metadata from the 2011 round for all regions.

Given this analysis, we decided to deviate from the official PPPs for eight countries (Belize, Egypt, Guinea, Iraq, Nigeria, São Tomé and Príncipe, Sudan, and Trinidad and Tobago). For these countries, we use the geometric average of official and imputed 2017 PPPs. In the absence of a definitive conclusion on which PPP is appropriate, we believe using the averages of the official and imputed PPPs for these eight countries provides the most robust PPPs.¹⁴ We use extrapolated 2017 PPPs for a few other exception countries that do not have official 2017 PPPs, including Kiribati, Nauru, the Syrian Arab Republic, Tuvalu, the República Bolivariana de Venezuela, and the Republic of Yemen.¹⁵

We conduct sensitivity analysis to understand the extent to which the alternative PPPs influence the IPL and extreme poverty estimates. The IPL remains robust to alternative PPPs and global and regional trends are virtually unaffected by the alternative PPPs (see Figure D.4). Global extreme poverty would decline by 0.4pp (instead of 0.2pp) in 2017 if the official 2017 PPPs were used throughout (see Table D.7). This change is largely driven by Sub-Saharan Africa (particularly Nigeria) and the Middle East & North Africa (particularly Egypt).

4. Derivation of global poverty lines

National poverty lines form the basis of the construction of the World Bank's global poverty lines. A national poverty line typically expresses the minimum amount of expenditure expected to cover the basic needs of a person, usually including food, clothing, and other non-food items. By relying on national poverty lines to guide the determination of global poverty lines, we rely on countries' own judgments of what it means to be poor. National statistical offices are responsible for deriving the national poverty lines in most low- and middle-income countries.¹⁶ National poverty lines are expressed in domestic currencies, and PPPs are used to convert different currencies into a common, comparable unit. PPPs are therefore instrumental in setting global poverty lines.

The IPL has historically been derived as a summary measure (e.g., mean or median) of the national poverty lines of some of the poorest countries in the world. For example, the IPL of \$1.25 per person per day was the mean of PPP-adjusted national poverty lines of 15 of the poorest countries in the world, as judged by countries' household final consumption expenditure per capita around 2008 when the 2005 PPPs were released (Ravallion et al. 2009). When the 2011 PPPs were released in 2014, the same 15 national poverty lines were used, but now converted to 2011 PPPs

¹⁴ Using the geometric average is similar to the ICP's approach for interpolating between ICP rounds, where forward and backward extrapolations are averaged (Inklaar and Rao 2020). Furthermore, invoking a Bayesian interpretation, Chen and Ravallion (2010) take the geometric average between mean consumption in surveys and national accounts. Recently, Ravallion suggested taking a weighted average of the revised 2011 and 2017 PPPs (Ravallion 2020).

¹⁵ We extrapolate from the revised 2011 PPPs used for global poverty monitoring, meaning that for the Republic of Yemen we extrapolate from imputed revised 2011 PPP.

¹⁶ Low-capacity national statistical offices often do so in collaboration with the World Bank. In such cases, the goal of the Bank is to assure that the national poverty lines reflect countries' own normative judgments and are based on the latest advances in poverty estimation.

yielding an IPL of \$1.88, rounded to \$1.90 (Ferreira et al. 2016; Ferreira et al. 2015). In this paper, we maintain the principle that the IPL should be constructed as the typical poverty line of the poorest countries in the world but, following Jolliffe and Prydz (2016), make four changes to the methodology, addressing four issues with the current derivation of the IPL.

First, the national poverty lines of the 15 poor countries were reported in different units, e.g., some are in adult equivalent terms and others in per capita terms, creating incomparability in the final PPP-adjusted national poverty lines (Jolliffe and Prydz 2016). Second, the 15 countries were selected based on data showing that, for countries with per capita household final consumption expenditure at the level of these countries, national poverty lines were uncorrelated to per capita household final consumption expenditure (Ravallion et al. 2009). But later evidence with more data found that national poverty lines are positively correlated to per capita household final consumption expenditure at all levels (Jolliffe and Prydz 2016). Hence, increasing the sample to more than 15 countries might induce greater statistical support. Third, using national poverty lines constructed primarily in the 1990s for setting global poverty lines for measuring poverty today would rely heavily on CPI series over a long period of time from countries with typically weak statistical capacity. This would make the IPL sensitive to revisions made to historical CPI data from the 15 poor countries. Finally, using the mean of the national poverty lines might make the final IPL vulnerable to outliers, compared to the median.

In an attempt to address these four issues, we follow the ‘harmonized poverty line’ approach of Jolliffe and Prydz (2016) to derive the IPL in 2017 PPPs. Rather than converting national poverty lines using PPPs and CPIs, the harmonized poverty line approach matches national poverty rates with consumption/income distributions in PovcalNet, which are already expressed in the same units; per capita PPP terms.¹⁷ For each country, the percentile of the PovcalNet distribution that corresponds to the reported national poverty rate is found (in other words, we query the inverse cumulative distribution function). This yields national poverty lines expressed in per capita PPP terms for all countries. Jolliffe and Prydz (2016) selected surveys conducted as close as possible to the 2011 ICP reference year, such that only one survey per country is used for deriving the IPL. Here we will do the same with 2017 as the reference year, thus minimizing the reliance on CPI series to move national poverty lines to the 2017 ICP reference year. We again follow Jolliffe and Prydz (2016) and use all low-income countries rather than 15 countries, thus increasing statistical support, and take the median poverty line instead of the mean to make the IPL robust to outliers.

In what follows, we describe the results of using this methodology to update the IPL. We show that the IPL is robust to choosing a broader set of countries, a broader set of national poverty lines, and using an equivalent poverty line—a line that retains the global poverty rate in a pre-specified

¹⁷ There are thus a practical and a conceptual reason for this approach. First, the national poverty rates are readily available in WDI while the national poverty lines are not. Second, the national poverty lines may be in different units and using the harmonized lines avoids this problem.

reference year (Kakwani and Son, 2016).¹⁸ We also update the two higher lines Jolliffe and Prydz (2016) derived (see Section 4.A below for more details).

A. Global poverty lines based on harmonized national poverty lines

Jolliffe and Prydz (2016) estimate a median harmonized national poverty line of \$1.91 in 2011 PPPs from 33 low-income countries. This finding provided evidence in support of the IPL of \$1.90 per person per day. They also estimated \$3.21 and \$5.47 as the median harmonized national poverty lines for 32 lower- and 32 upper-middle-income countries, respectively. These lines, rounded to \$3.20 and \$5.50, have been adopted by the World Bank as additional global poverty lines. Using the same approach to update the global poverty lines with the 2017 PPPs would ensure consistency in the Bank's methodology in monitoring global poverty. When we update the analysis of Jolliffe and Prydz (2016) with the revised 2011 PPPs, the median lines become \$1.85, \$3.21, and \$5.65 (see Table 1).¹⁹ The World Bank decided to keep the global poverty lines unchanged with the revised 2011 PPPs, including the \$5.50 line (World Bank 2020a), and we follow that approach here when we use the revised 2011 PPPs.

Table 1: Updating global poverty lines with harmonized national poverty lines

Income classification	(A) Original 2011 PPP			(B) Revised 2011 PPP			(C) 2017 PPP		
	Median	Mean	N	Median	Mean	N	Median	Mean	N
Low	1.91	2.23	33	1.85	2.22	33	2.15	2.42	28
Lower-middle	3.21	3.88	32	3.21	3.89	32	3.63	3.95	54
Upper-middle	5.47	5.62	32	5.65	5.64	32	6.85	7.05	37
High income	21.70	21.19	29	21.70	21.31	29	24.36	23.36	38
Observations			126			126			157

Notes: In panels (A) and (B), we use the harmonized poverty lines Jolliffe and Prydz (2016) derived. Panel (A) replicates the results of Jolliffe and Prydz (2016). Panel (B) uses revised 2011 PPPs to update the analysis. Panel (C) is based on new survey data together with changes in PPPs to 2017 PPPs and changes in CPIs between 2011 and 2017 (see Appendix E, section 5). Countries are categorized according to the World Bank's income classification in the years the surveys were conducted. In panel (C), the line that is closest to 2017 is selected, one for each country. If harmonized national poverty lines are available for 2016 and 2018 but not 2017, 2018 is selected.

For the 2017 PPPs, we use new harmonized poverty lines to estimate global poverty lines. We select for each country one survey that was conducted in 2017 or the closest year. Eighty-four percent of selected surveys were conducted after 2011 and 43% were conducted in 2017. More than three-quarters of selected surveys were conducted within three years from 2017. Compared with Jolliffe and Prydz (2016), new countries have been included in the analysis of the harmonized national poverty lines (e.g., Algeria, Bangladesh, and Egypt), and some of the countries in the low-

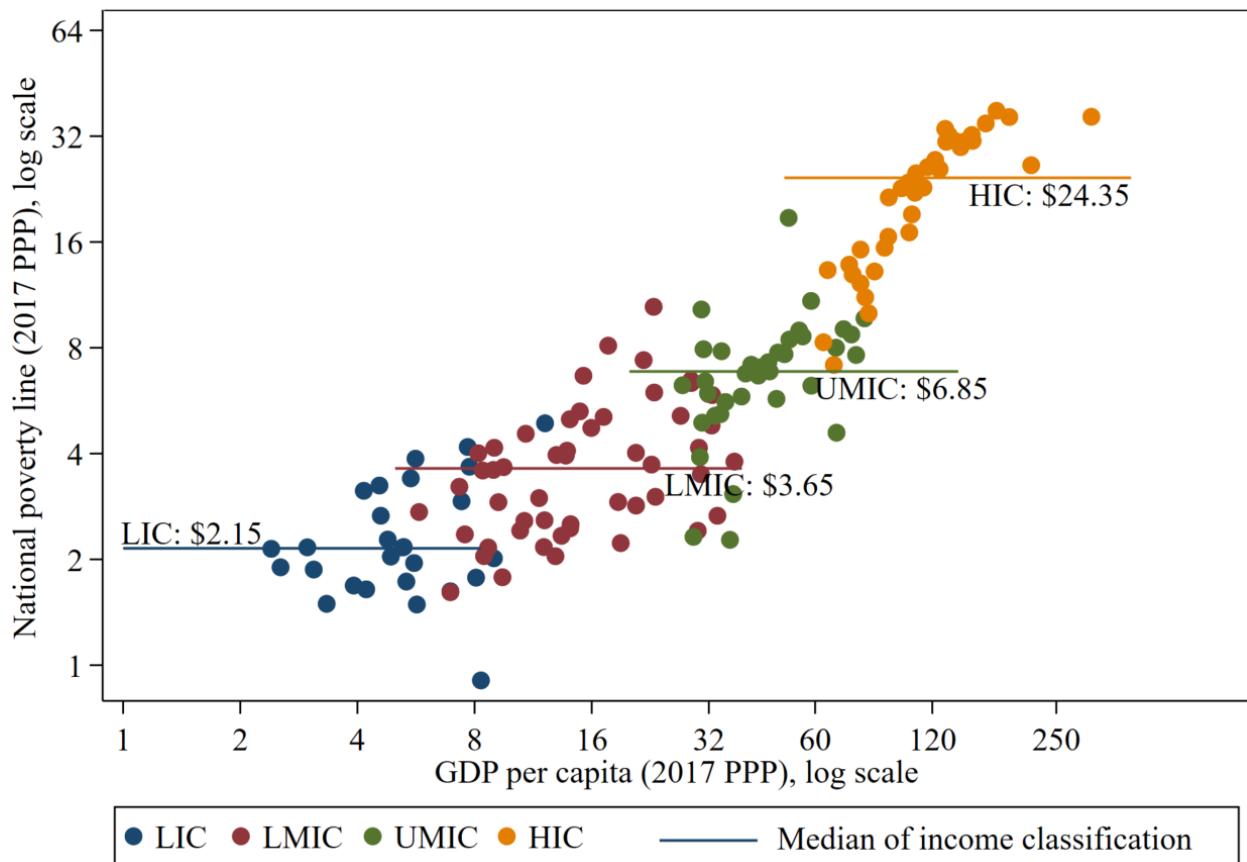
¹⁸ There are alternative methods that could be used to derive global poverty lines than the ones we use. Global poverty lines, for example, could be estimated using PPPs that equilibrate the least-cost diets across countries (Bai and Masters 2020; Allen 2017).

¹⁹ See Appendix E, section 3 for the derivations of the harmonized national poverty lines, section 4 for how these lines are updated with the revised 2011 PPPs, and section 5 for the updating to the 2017 PPPs.

income category moved up to the lower-middle-income category (e.g., Ghana, Kenya, Pakistan, São Tomé and Príncipe, and Tajikistan). We categorize countries into low, lower-middle, upper-middle, and high-income countries based on the World Bank's income classification in the years the surveys were conducted.

Based on this approach, the median poverty line for low-income countries, or the international poverty line (IPL), is estimated to be \$2.15 (2017 PPP). The higher lines using data from lower-middle-income countries and upper-middle-income countries are estimated to be \$3.65 and \$6.85, respectively, in 2017 PPP. We use these lines with the 2017 PPPs in the rest of the paper. Figure 3 illustrates the global poverty lines graphically (updating a figure in Ferreira and Sánchez-Páramo (2017)).

Figure 3: Global poverty lines using harmonized national poverty lines



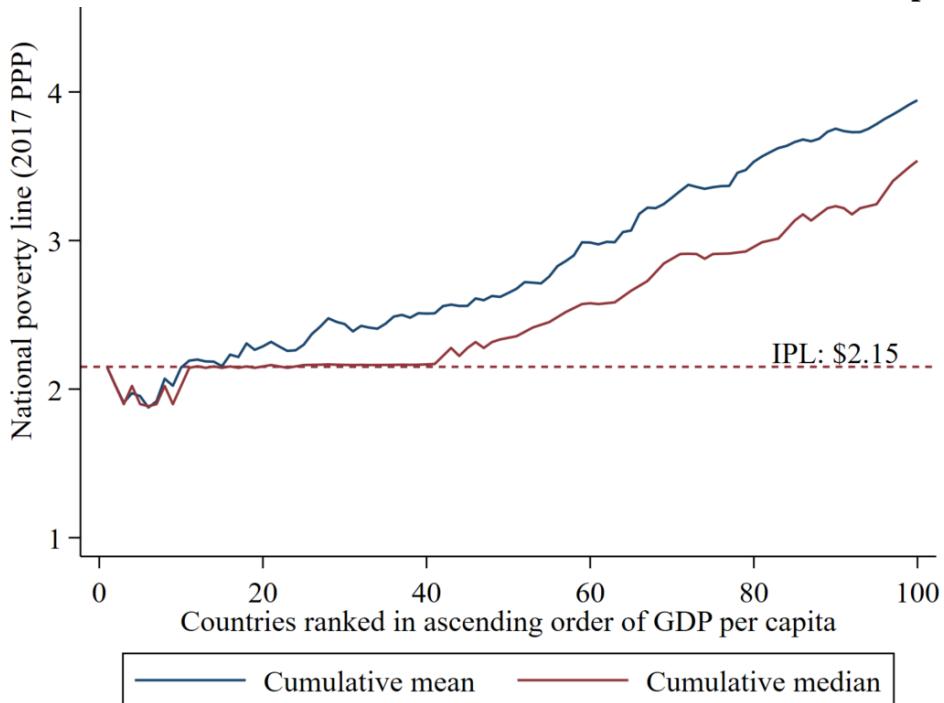
Notes: This chart is based on the data summarized in Table 1, panel (C). The national poverty lines are the harmonized national poverty lines, expressed in 2017 PPP per person per day. GDP is also given in 2017 PPP per person per day. The global poverty lines are the median poverty lines for low-income countries (LIC), lower-middle-income countries (LMIC), upper-middle-income countries (UMIC), and high-income countries (HIC), rounded to the nearest 5 cents. Income classifications are determined using a slightly different measure—GNI per capita with the Atlas method (see the footnote in the main text)—hence GDP per capita does not perfectly order countries according to their income classification. There are fewer observations of GNI per capita, which is why GDP per capita is used here.

In line with one recommendation of the *Atkinson Commission on Global Poverty*—to express the global poverty lines in local prices—in Appendix F we list the four global poverty lines with the 2011 PPPs and 2017 PPPs in 2020 local currency units for each country with available data.

B. Robustness check: Using different sets of poorest countries

One issue with the approach above is that it is not clear that the group of low-income countries is the appropriate set of poor countries. Given the positive slope between national income and national poverty lines, one faces a trade-off between including more countries in the calculation—and hence increasing statistical support—and focusing on only the poorest countries. To check the robustness of the \$2.15 IPL derived above, we order countries by GDP per capita in 2017 PPP and plot the cumulative median of harmonized poverty lines in 2017 PPPs, sequentially increasing the pool of countries starting with the country with the lowest GDP per person in 2017 (see Table B.5 for details). Figure 4 shows that using anywhere between 11 and 41 poorest countries, the median poverty line rounds to \$2.15. This suggests that using the median national poverty line of low-income countries is robust to using a broad range of poorest countries.

Figure 4: Cumulative mean and cumulative median of harmonized national poverty lines



Notes: The harmonized national poverty line is in 2017 PPP-adjusted dollars per person per day. The number of countries indicates the count of countries ranked by GDP per capita in 2017 PPP-adjusted dollars from the poorest to the richest. For example, 20 indicates the poorest 20 countries. This chart is based on the data summarized in Table 1, panel (C) but the number of countries is capped at 100 to better illustrate the IPL.

C. Robustness check: Using different sets of harmonized national poverty lines

The global poverty lines we derive are different from the \$1.90, \$3.20, and \$5.50 lines because the PPPs have changed, but also because of changes to national poverty lines, the set of countries with national poverty lines, and the income classification of these countries. In Appendix B, we show that when we use the same harmonized national poverty lines as Jolliffe and Prydz (2016), we still get an estimated international poverty line of \$2.15 (see Table B.6). When we pool together the over 1,300 harmonized national poverty lines that we newly derived, we obtain an IPL of \$2.16 (Table B.6). One might worry about using old harmonized national poverty lines given that the statistical capacity of most countries is increasing over time and given that using old poverty lines relies on many years of CPI data. If we restrict our sample to poverty lines at most 10 years old with respect to the PPP year, we arrive at an IPLs \$2.14 or \$2.16 depending on whether we use the pooled sample or only the poverty line closest to 2017 for each country (see Tables B.6 and B.7).

D. Robustness check: Using equivalent poverty lines

We use equivalent poverty lines, a concept introduced by Kakwani and Son (2016) at the country level, as an alternative method to derive the IPL. The idea behind this method is to estimate the IPL with the 2017 PPPs by choosing a line that retains the global poverty rate in a pre-specified reference year. For example, the equivalent poverty line could be set to keep constant the global poverty rate in 2010, the year that had the most recent global poverty estimate when the World Bank's goal of ending extreme poverty was set. This method is attractive in light of the *Atkinson Commission Report*, which was concerned with future ICP rounds shifting the goalposts (World Bank, 2017). With this approach, the parameter of choice is then the year in which poverty should be held constant, for which several options appear reasonable. For example, this could be 2012, which is the year with the most recent global poverty estimate when the UN agreed on the SDGs. It could also be 2015, the year that had the most recent global poverty estimate when Atkinson wrote his report. Finally, we include the 2017 reference year in the analysis, which is the year with the latest global poverty numbers at the time of writing. See Appendix E, section 6 for the derivation of equivalent poverty lines.

Table 2 presents the results of the analysis described above.²⁰ The equivalent IPL that retains the global poverty rate rounds to \$2.15 (2017 PPP) regardless of which of the four reference years is

²⁰ The poverty rates in Table 2 are based on the revised 2011 PPPs, even though the World Bank and UN goals to eradicate extreme poverty were set with reference to the \$1.25 IPL in 2005 PPP. We use the revised 2011 PPPs for two reasons. First, the problems with the 2005 ICP round are well documented (Deaton and Aten 2017; Inklaar and Rao 2017). Second, Atkinson wrote his report based on the \$1.90 IPL in 2011 PPP with the acknowledgment that the adoption of the 2011 ICP round might have already shifted the goalposts.

selected.²¹ Our preferred year for evaluating the equivalent IPL is 2017, which is both the latest ICP reference year and the year with the latest global poverty data. The equivalent IPL for 2017 is \$2.173.

Table 2: Equivalent international poverty line, 2017 PPP

Region	Year	Headcount, % (Revised 2011 PPP)	Equivalent IPL (2017 PPP)
World	2010	16.02	2.119
World	2012	12.89	2.128
World	2015	10.14	2.161
World	2017	9.27	2.173

Note: This table shows equivalent international poverty lines that keep global poverty headcounts constant in pre-specified reference years.

E. Robustness check: Using the approach of Ravallion et al. (2009)

For completeness, let us also state that if we were to use the 15 national poverty lines selected by Ravallion et al. (2009) and further applied by Ferreira et al. (2016), convert them to 2017 PPPs, and take the average, we would end up with an IPL of \$2.08 (see Table B.9. in Appendix B for details). As argued above, this is not our preferred approach to estimate the IPL.

F. Updating the societal poverty line

Just as the value of the international poverty line and the higher lines reflecting typical assessments of basic needs in middle-income countries are expressed in PPP dollar terms, so too is the World Bank's societal poverty line (SPL). The SPL was introduced by the World Bank (2018) in response to the *Atkinson Commission on Global Poverty* (2017, recommendation 16) recommendation to report a measure of poverty that combines fixed and relative elements of poverty. The SPL, estimated for each country separately, was initially parameterized as: $\max(IPL, \$1 + 50\% \text{ of median consumption})$, with the fixed parameters (i.e., the IPL and the intercept) expressed in 2011 PPPs (Jolliffe and Prydz 2021). Of the two parameters in PPP terms, the IPL is updated to \$2.15 in 2017 PPPs, leaving only the intercept of \$1 in need of an update. We examine two methodological approaches to consider how best to update this value. First, we follow an approach similar to how we have updated the higher lines by replicating the methodology used to determine these values except now using the 2017 PPPs and more recent data. As with the IPL and the higher poverty lines, we consider different samples and approaches.

Following Jolliffe and Prydz (2021) we regress harmonized national poverty lines on median values of consumption (or income, depending on the welfare aggregate used in the country).

²¹ Table B.8 reports a slight variant of the equivalent poverty lines. The estimated international poverty line in this case is derived as the weighted sum of the equivalent regional poverty lines that hold constant regional poverty rates in a particular reference year. The weights correspond to the regional shares of the global poor in the reference year. This approach suggests an IPL between \$2.16 and \$2.24 with the 2017 PPPs.

Jolliffe and Prydz subsample by country type to try to exclude countries with relative poverty lines. As a preferred alternative to this, we have collected metadata on type of poverty line from most countries and subsample on all countries that have absolute poverty lines. (The original objective of selecting SPL parameters was to demonstrate an empirical relationship between absolute poverty lines and level of economic wellbeing in a country.) Table 3 shows the estimated intercept and slope coefficients when considering three such specifications. We first consider a sample ($n=119$) where we have one observation for each country with an absolute poverty line. For each country, we select the observation which is closest in time to 2017 (and refer to this as our circa-2017 sample). We then consider the subsample of all ($n=815$) country-year observations, allowing multiple observations per country capturing some within-country variation over time in their values. We use two different weighting schemes, where in both cases we ensure that all countries are weighted equally, but in one specification, we weight more heavily the data points that are closer in time to 2017 (see table notes for details). Across these three specifications, the intercept ranges from 1.15 to 1.17 (in 2017 PPPs). While that variation is not substantial, it's noteworthy that these are very imprecisely estimated parameters with 95 percent confidence intervals that are almost \$1 in width (e.g., $\$1.15 \pm \0.49 for the circa-2017 sample). For the circa-2017 sample, the standard error is \$0.25 but this drops only a few cents when considering sample sizes that are nearly seven-fold larger.²²

Table 3: Updating the societal poverty line with new survey data and 2017 PPPs

Description	Intercept	Slope	N	Delta ratio
One line per country – circa-2017 sample	1.15 (0.249)	0.47 (0.043)	119	1.15 (0.014)
Pool all lines – uniform weights	1.16 (0.213)	0.52 (0.043)	815	
Pool all lines – higher weights for lines closer to 2017	1.17 (0.235)	0.49 (0.044)	815	

Notes: Using the same model specification in Jolliffe and Prydz (2021) and the new data used in updating the absolute poverty lines in this paper (Table 1, panel C), the parameters are updated with the 2017 PPPs. This table excludes all national poverty lines that are defined as relative lines (which is typical of high-income countries). For the circa-2017 sample, every country with an absolute poverty line is included but only for the year that is closest to the 2017 ICP benchmark year. In the larger samples ($n=815$), the weighting schemes differ. In one scheme, every observation has a weight equal to the inverse of the total number of observations for each country, so that each country has a sum of weights equal to 1. In the other scheme, we weight the observations to give greater importance to those close to 2017 and to give equal weight to each country: In the first step, we assign a survey conducted in 2017 a weight of 1, a survey conducted in 2016 or 2018 a weight of 1/2, a survey conducted in 2015 or 2019 a weight of 1/3, and so on. Next, we rescale the weights to ensure that the sum of weights for each country is again equal to 1. For example, if a country has three harmonized poverty lines for 2016, 2017, and 2018, the corresponding weights will be 0.25, 0.5, and 0.25. All parameter estimates are statistically significant at the 1% level. Standard errors of the estimates are in parentheses.

The magnitude of the confidence intervals is concerning and leads us to consider a complementary approach to updating the intercept. We use this complementary approach to assess the sensitivity

²² This relatively minor decline in standard errors is due to the significant intra-country correlation of the error term (and correcting for this with the sandwich variance estimator).

of the estimated intercept. Part of the reason why the confidence intervals around the intercept are so large is because the intercept is an out-of-sample prediction. The intercept essentially represents a minimum cost of basic needs under which no person could survive, no matter how poor a particular country is. No country provides a direct estimate of this value; we can only infer it if we impose modelling assumptions. This is conceptually very different from how we approached updating the IPL which reflects a median value over a large sample of national poverty lines and is a well-supported, within-sample estimate. The alternative method considered is to update the intercept based on the delta ratio, first discussed in Section 3.B. The delta ratio, unique to each country, directly converts 2011 PPP values for each country into 2017 PPP values. For the subsample of countries that use absolute poverty lines (i.e., those countries used in this analysis), Table 3 reports an average value of the delta ratio of 1.15. If using this estimation approach to update the intercept, the revised intercept would be \$1.15 in 2017 PPPs.

Based on these findings, we assert that an appropriate updating of the SPL in 2017 PPPs is parameterized as: $\max(IPL, \$1.15 + 50\% \text{ of median consumption})$. We base this inference on three main points coming from the two approaches. The first is that both methods result in estimated intercepts that fall within the narrow range of (\$1.15, \$1.17), with two of the four estimates being \$1.15. Second, while the confidence intervals from the estimates based on replicating the methodology of Jolliffe-Prydz are massive, the confidence interval based on the delta method is very tight ($\$1.15 \pm \0.03) indicating significantly greater precision from this estimation approach. And, finally, we note that all of the estimated intercepts, when rounded to the nearest five cents (as we do for the other poverty lines), are equal to \$1.15.

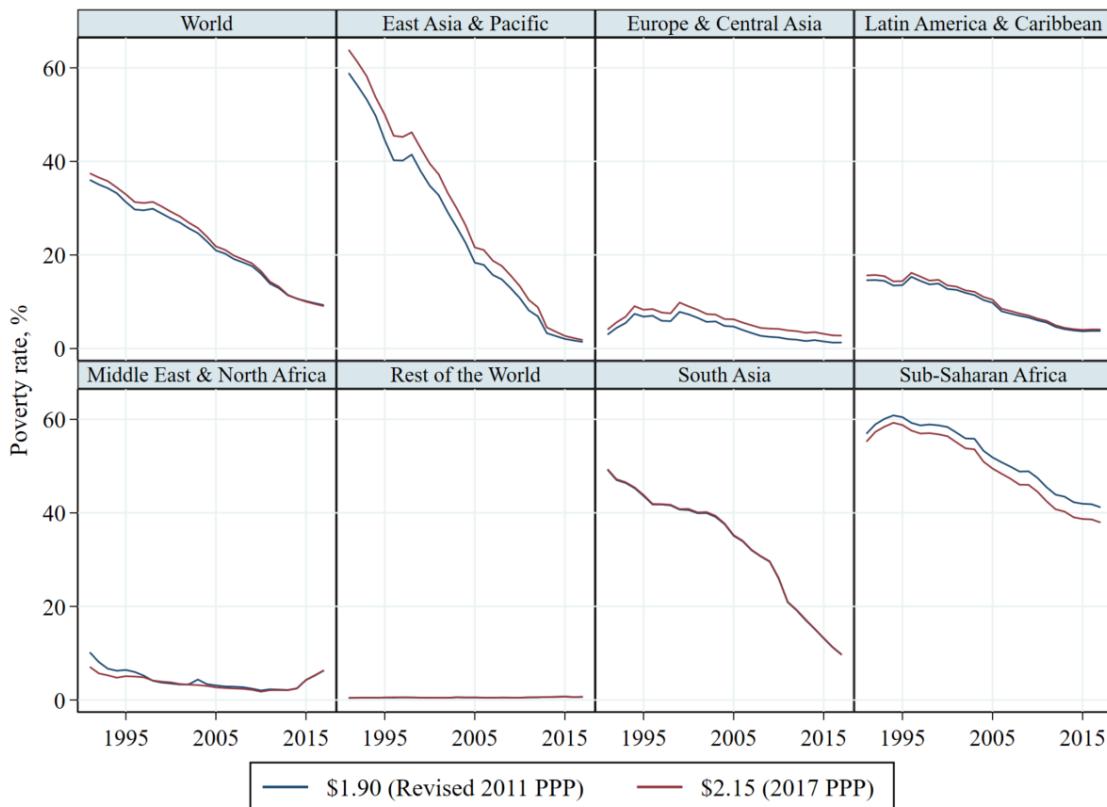
5. Global, regional, and country-level poverty estimates

In this section, we compare global, regional, and country-level poverty estimates over time with the 2017 PPPs, relative to the revised 2011 PPPs. We analyze how using the 2017 PPPs would change the incidence and geographic distribution of poverty at the IPL and at the higher global poverty lines.

Figure 5 shows regional and global trends in extreme poverty between 1991 and 2017 with the two sets of PPPs. Extreme poverty is measured as the share of the population living on less than \$1.90 or \$2.15 a day expressed in revised 2011 PPP or 2017 PPP, respectively. At the global level, the revised 2011 PPPs and 2017 PPPs induce relatively small changes to extreme poverty.

Between 1991 to 2017, extreme poverty in the world falls from 36.05% to 9.27% with the revised 2011 PPPs, and from 37.46% to 9.07% with the 2017 PPPs. The 2017 PPPs slightly increase historical estimates of extreme poverty and slightly decrease extreme poverty since 2014. Regional trends in extreme poverty are similar for both revised 2011 and 2017 PPPs. There is progress against poverty across regions, except for the Middle East & North Africa where conflict and fragility reverse the progress made (World Bank 2020a).

Figure 5: Global and regional trends in extreme poverty



Note: This figure shows the trends in extreme poverty at the global and regional levels. The international poverty lines used are \$1.90/day (revised 2011 PPP) and \$2.15/day (2017 PPP).

Table 4 reports the changes in global and regional poverty levels in 2017, the most recent year. Overall, extreme poverty in the world decreases marginally by 0.2pp with the 2017 PPPs. Extreme poverty more than doubles in Europe & Central Asia (but from a low level), and slightly increases in all other regions, except Sub-Saharan Africa. Extreme poverty reduces substantially in Sub-Saharan Africa (34 million fewer poor people), driving down the global count of the extreme poor by 15 million. Sub-Saharan Africa still has the largest share of millions of extreme poor, though it falls from 62% with the revised 2011 PPPs to 58% with the 2017 PPPs.

To understand the regional trends further, Table 5 shows the ten countries with the largest absolute changes in extreme poverty.²³ The change observed in Europe & Central Asia is driven by Uzbekistan, where poverty increases by 22pp (equivalent to 7 million more poor people).²⁴ About

²³ We use the 2017 estimates for all countries. As explained above, when there is no survey in 2017, these numbers are estimated using growth rates from national accounts and assuming no distributional changes.

²⁴ This is the first time the ICP provides a PPP estimate for Uzbekistan based on actual prices. The price data collection was based on an experimental approach, meaning that prices were collected for only *actual individual consumption* and *individual consumption expenditure by households*, but not for all the components included in GDP. The original and revised 2011 PPPs were imputed by the ICP like other non-benchmark countries. The 2017 PPP estimate is deemed more credible, but potentially not comparable to estimates from previous rounds. This might explain why Uzbekistan has large changes in poverty.

half of the change in extreme poverty in Sub-Saharan Africa is driven by Nigeria, where poverty falls by 8pp (equivalent to 15 million fewer poor people).

Table 4: Changes in extreme poverty in 2017 at the global and regional levels

Region	Poverty rate, % (2011 PPP)	Poverty rate, % (2017 PPP)	Change in poverty, pp	Change in poverty, %	Change in millions of poor
World	9.27	9.07	-0.20	-2.21	-15.36
East Asia & Pacific	1.41	1.82	0.41	28.98	8.45
Europe & Central Asia	1.30	2.77	1.48	113.93	7.26
Latin America & Caribbean	3.77	4.07	0.31	8.15	1.93
Middle East & North Africa	6.34	6.36	0.02	0.36	0.09
Rest of the World	0.68	0.68	0.00	0.12	0.01
South Asia	9.65	9.71	0.06	0.58	1.01
Sub-Saharan Africa	41.18	37.94	-3.25	-7.88	-34.10

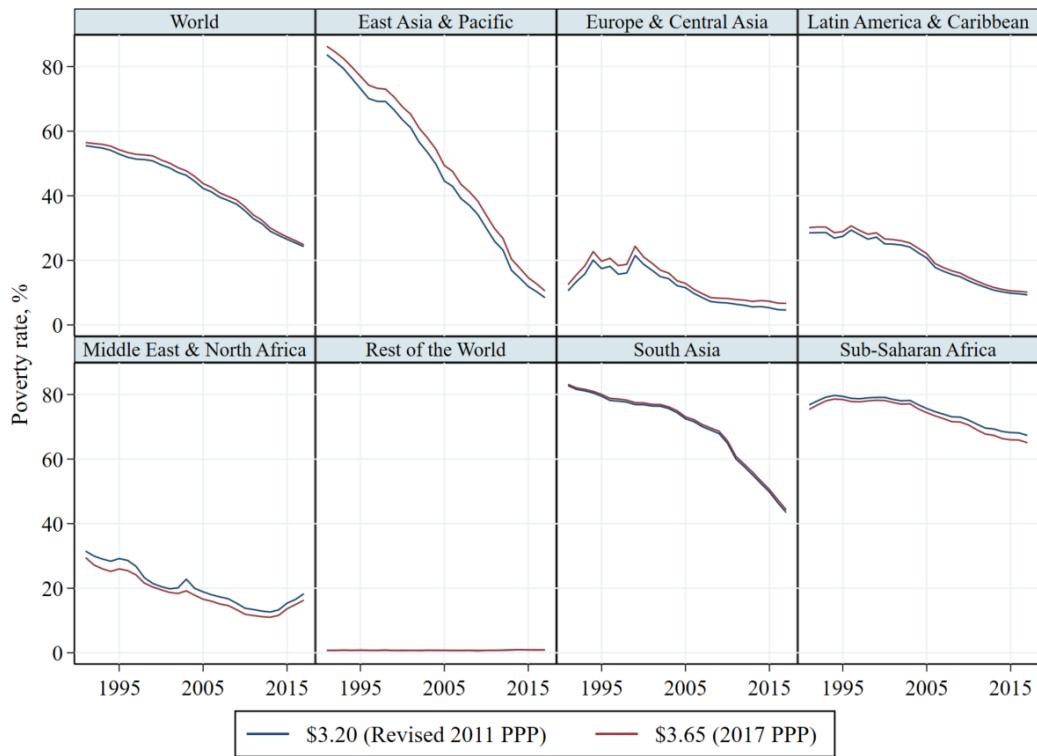
Note: This table compares extreme poverty between the revised 2011 PPPs and 2017 PPPs at the global and regional levels.

Table 5: Changes in extreme poverty in 2017 at the country level, top 10 countries

Country	Region	Poverty rate, % (2011 PPP)	Poverty rate, % (2017 PPP)	Change in poverty, pp	Change in poverty, %	Change in millions of poor
Nigeria	SSA	41.36	33.41	-7.96	-19.23	-15.19
Uzbekistan	ECA	12.70	34.35	21.65	170.53	7.01
Congo, Dem. Rep.	SSA	72.65	64.42	-8.24	-11.34	-6.70
Indonesia	EAP	4.46	6.55	2.09	47.01	5.54
Angola	SSA	45.07	26.50	-18.57	-41.20	-5.54
Kenya	SSA	35.11	27.50	-7.61	-21.67	-3.82
Ethiopia	SSA	24.66	21.16	-3.50	-14.18	-3.72
Ghana	SSA	12.34	24.68	12.33	99.91	3.59
Pakistan	SAS	4.02	5.44	1.42	35.35	2.95
China	EAP	0.35	0.55	0.19	55.05	2.69

Notes: This table compares extreme poverty between the revised 2011 PPPs and 2017 PPPs at the country level. The top 10 countries with the largest absolute changes in extreme poverty are shown in this table. Countries are ranked in descending order of absolute changes in millions of poor. When a country does not have a survey in 2017, recent surveys are extrapolated or interpolated using growth rates in national accounts data, particularly GDP or household final consumption expenditure (HFCE).

Figure 6: Global and regional trends in poverty at \$3.20



Note: This figure shows the trends in poverty at the global and regional levels at the following poverty lines: \$3.20/day (revised 2011 PPP) and \$3.65/day (2017 PPP).

Figure 7: Global and regional trends in poverty at \$5.50

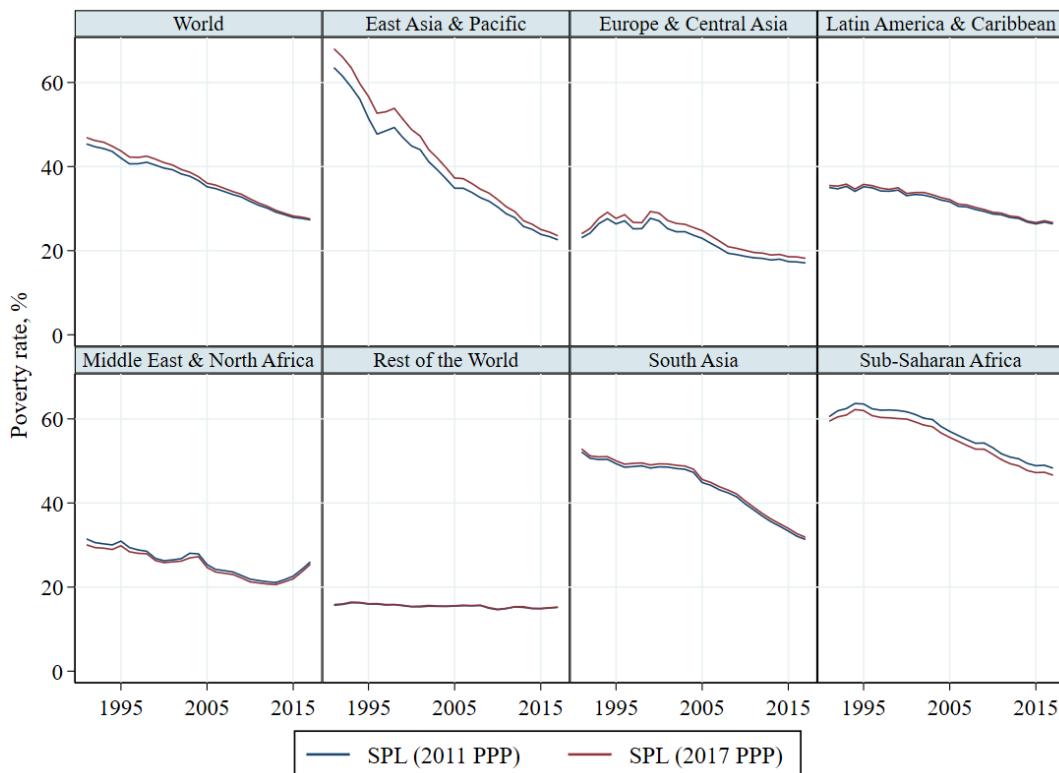


Notes: This figure shows the trends in poverty at the global and regional levels at the highest global poverty line. The lines used are \$5.50/day (revised 2011 PPP) and \$6.85/day (2017 PPP).

Changes in poverty at the global level are also relatively small when the \$3.65 line is used, as regional changes offset each other (see Figure 6). Poverty increases in East Asia & the Pacific, and Europe & Central Asia, and Latin America & the Caribbean, while poverty decreases in the Middle East & North Africa and Sub-Saharan Africa. In 2017, global poverty changes by 0.6pp, or 43 million more poor people, with the 2017 PPPs. The largest changes in millions of poor in 2017 are observed in East Asia & the Pacific (43 million more poor people), and Sub-Saharan Africa (24 million fewer poor people). The change in East Asia & the Pacific is mainly driven by China and Indonesia, while the change in Sub-Saharan Africa is mainly driven by Nigeria. See Tables B.10 and B.11 for a full set of results.

Using a \$6.85 line (2017 PPP) relative to a \$5.50 line (2011 PPP) would lead to increases in poverty markedly in all regions, including Sub-Saharan Africa (see Figure 7). With the 2017 PPPs, about 321 million more people in the world would be considered poor in 2017 by the standards of upper-middle-income countries. About 129 million of them live in China and 64% of them live in China, India, and Indonesia alone. See Tables B.12 and B.13 for a full set of results.

Figure 8: Global and regional trends in societal poverty



Notes: This figure shows the trends in societal poverty at the global and regional levels. With the revised 2011 PPP, the societal poverty line (SPL) used is given as $\max(\$1.90, \$1.00 + 0.5 * \text{median consumption/income})$. With the 2017 PPP, the societal poverty line (SPL) used is given as $\max(\$2.15, \$1.15 + 0.5 * \text{median consumption/income})$.

If the societal poverty line were strongly relative (without any fixed or absolute elements), as in the OECD where the poverty line is 60% of the median, revisions to PPP data would not change global and regional poverty trends. The global welfare distribution would change by the same factor as the societal poverty line. Unlike the OECD poverty line, the SPL has two terms expressed in PPPs (i.e., the IPL and the intercept) whose revision drives the changes observed in societal poverty trends when moving to the 2017 PPPs (see Figure 8). In fact, the changes in societal poverty trends with the 2017 PPPs are small and similar to the changes observed in extreme poverty with the 2017 PPPs. Societal poverty slightly increases globally and, in all regions, except Sub-Saharan Africa and the Middle East & North Africa. About 17.6 million more people would be living in societal poverty in the world in 2017 when moving from the revised 2011 to 2017 PPPs. A few large countries drive this relatively small change, including China and Indonesia from East Asia & the Pacific, Nigeria and the Democratic Republic of Congo from Sub-Saharan Africa, India and Pakistan from South Asia, Iraq and the Arab Republic of Egypt from the Middle East & North Africa, and Uzbekistan from Europe and Central Asia (see Tables B.14 and B.15).

6. Discussion

A. Poverty in poorer versus richer countries

Compared to the 2011 PPPs, the 2017 PPPs imply lower price levels in relatively poor countries and higher price levels in relatively rich countries. In fact, in 2017, the share of the world's population living in poverty in low-income countries systematically decreases while the share of the world's poor living in upper-middle-income countries systematically increases at all three global poverty lines (see Figure B.1). The largest change occurs at the highest line, where the share of people in upper-middle-income countries considered as poor by the standards of these countries increases from 15% to 19%. Some of this change is driven by upward revisions in the national poverty lines of these countries (see Section 6.C below for more details).

We can make sense of the global changes in poverty and the shifting shares of poverty across income groups by investigating (1) the changes in poverty at the country level, and (2) the aggregation of country-level poverty estimates. The global pattern we observe could be explained by the former if the real changes in welfare when moving to the 2017 PPPs (i.e., the delta ratio discussed in Section 3.B) are systematically higher for low-income countries than for middle-income countries. As the IPL/LIC line increases by 13% from \$1.90 to \$2.15, extreme poverty increases in countries whose delta ratios fall short of 1.13 and decreases in countries whose delta ratios exceed 1.13. Figure B.2 shows that the average delta ratio by income group is decreasing in income, and the delta ratio for low-income countries on average exceeds 1.13. These results suggest that low-income countries are slightly richer with the 2017 PPPs, hence slightly shifting the concentration of extreme poverty away from the poorest countries.

The global pattern we observe could also be driven by the fact that populous countries carry a larger weight when aggregating country-level poverty estimates. Figure B.2 also plots delta ratios that are population weighted. For low-income countries, the delta ratio is even higher when population weighted, while the reverse holds for upper-middle-income countries. This suggests that the observation that poverty is shifting to relatively rich countries is *partly* driven by a few populous countries in the world. With their extreme poverty levels decreasing with the 2017 PPPs, Ethiopia and the Democratic Republic of Congo are top populous countries driving the change in extreme poverty in low-income countries. At the LMIC line, the changes in poverty in the populous countries offset each other; for example, the poverty counts increase in Indonesia by 13 million and decrease in Nigeria by 14 million. At the highest line, poverty increases in China, Brazil, the Russian Federation, and Mexico with the 2017 PPPs, driving down the delta ratio for upper-middle-income countries markedly when population weighted (see Figure B.2).

B. Extreme poverty in Sub-Saharan Africa

This sub-section presents additional analysis to better understand the changes in Sub-Saharan Africa, which shows the largest changes in extreme poverty as a result of the 2017 PPPs. Since Sub-Saharan Africa accounts for nearly two-thirds of the global extreme poor, it is not surprising that the region drives the global changes. At the same time, it is important to understand if there are any systematic issues with the new (or old) round of PPPs in the region.

On the whole, the 2017 ICP round has benefited from improved data quality assurance across all regions (World Bank 2020b). We investigate the quality of PPPs by comparing the residuals between published and predicted estimates of price level indices (PLIs) across regions. We predict PLIs from country-level characteristics (e.g., GDP per capita, export and import shares, and dependency ratio), following the official ICP model used to estimate PPPs for non-benchmark economies (for the model specification and more details, see Appendix D). Sub-Saharan Africa does not have systematically different residuals (Figure B.3), suggesting that there is no evidence of a worse PPP quality in the region.

Additionally, we consider the quality of CPIs for Sub-Saharan African countries relative to other regions. We draw on metadata covering 196 countries to assess the quality and coverage of official CPI data used in the global poverty estimates (Berry et al. 2019). We identify and analyze three relevant indicators, including the CPI expenditure classification, the reference year for the weights, and the spatial coverage of the expenditure weights (also see Section 3.B and Appendix D). Most countries in the world (77%) follow the standard *Classification of Individual Consumption According to Purpose* (COICOP) classification, with the Sub-Saharan Africa region not being very different (71%). On average, the CPI weights reference year for countries in Sub-Saharan Africa

is 2010.96, which is close to the world's average of 2011.85. CPI expenditure weights with national coverage in Sub-Saharan Africa (73%) is also similar to the world (79%).²⁵

In summary, it does not appear that the substantial reduction in extreme poverty in Sub-Saharan Africa is related to PPP or CPI quality issues. While there are obviously changes in rankings within the region, African countries are actually quite evenly split between increasing and decreasing poverty: Extreme poverty is lower with the 2017 PPPs in 24 countries and higher in 21 countries, with most countries being quite stable (see Figure B.4). However, most of the large countries see a decline with the 2017 PPPs. In particular, Nigeria, the Democratic Republic of Congo, Angola, Kenya, and Ethiopia are driving the regional result. Nigeria alone accounts for about half of the change in the millions of extremely poor people in the region.

C. Drivers of changes in the global poverty lines

In this sub-section, we try to understand the drivers of the revisions to the global poverty lines.²⁶ We pay particular attention to the large change at the highest line from \$5.50 to \$6.85.

The observed changes in the global poverty lines are not only due to the change in PPPs but also three other factors: (a) the national poverty lines used, (b) income group classifications, and (c) the set of countries with national poverty lines and welfare aggregates available. The global poverty lines are obviously also a function of the method used to derive them. Though we have changed the method by which we derived the IPL—we now use the harmonized national poverty line approach to update the IPL, similar to what is used for the higher lines—we can abstract from this methodological change since this approach also resulted in a \$1.90 poverty line with the 2011 PPPs (Jolliffe and Prydz 2016).

When deriving the IPL we already showed that it was robust to using the exact same poverty lines as Jolliffe and Prydz (2016). In Table B.14 we try to gauge whether any of the other three factors are driving changes to the poverty lines. To isolate the various factors, we update the harmonized national poverty lines previously used by Jolliffe and Prydz (2016) with the 2017 PPPs while keeping everything else constant, and sequentially account for each of the three factors above.²⁷ First we revisit our result of maintaining the exact same sample of national poverty lines but update the PPPs from the 2011 round to the 2017 round. As we argued in Section 4.A, this generates an IPL of \$2.15, meaning that our revision to the IPL can entirely be explained by the new PPPs, regardless of the other factors. For the poverty line typical of lower-middle-income countries, we arrive at a poverty line of \$3.68, quite close to the final one when accounting for all changes

²⁵ These differences are not statistically significant. These indicators are available in Appendix D, Table D.2 for the outlier countries identified in Section 3 of the paper.

²⁶ We are thankful to Francisco Ferreira and Benoît Decerf for encouraging us to delve deeper into this point.

²⁷ This decomposition is path dependent, meaning that the order in which we account for the various factors matters. Our results remain qualitatively the same regardless of which order we use.

(\\$3.63), meaning that nearly all of the updates to that line can be explained by the changes to PPPs. For the line typical of upper-middle-income countries, revising only the PPPs brings the line to \\$6.32, quite far from the final line (\\$6.85) meaning that the PPP changes cannot explain all of the increase of that line.

Next, we account for the fact that the countries Jolliffe and Prydz (2016) used for their analysis might since have published more recent national poverty lines. This has small impacts on the IPL and the lower-middle-income line, but increases the upper-middle-income line notably, from \\$6.32 to \\$7.15. Thus, the relatively high increase in the upper-middle-income line is partially driven by real upward shifts in the national poverty lines of upper-middle-income countries. Part of this can be explained by some of these countries now being high-income countries, but even when removing those countries, the line still sits at \\$6.95 (see Table B.16 for more details). This might be because poverty in upper-middle-income countries is more likely to reflect the concept of relative poverty, which changes over time with consumption patterns and incomes, rather than prices, which are more relevant for absolute poverty (Smeeding 2016). The fact that we now have national poverty lines and welfare aggregates for more countries than Jolliffe and Prydz (2016) only has a minimal impact on the value of the global lines.

Even for countries whose national poverty lines are based on concepts of absolute poverty, we know that as countries get wealthier, they tend to increase the real value of their national poverty line (Ravallion 1998; Jolliffe and Prydz 2016). Between the release of the 2011 PPPs and 2017 PPPs, on average countries got wealthier. This means that we would expect the national poverty lines close to 2017 to be higher in real terms than the national poverty lines close to 2011. How is this consistent with our finding that the real value of the IPL and lower-middle-income line is purely driven by PPPs and not by changes to the real value of national poverty lines? The answer lies in the use of income groups to define the global poverty lines. Suppose a poor country grew between the release of the 2011 PPPs and 2017 PPPs and as a consequence adopted a higher national poverty line. Suppose further that the country graduated from being a low-income country to a lower-middle income country as a result of this growth. Then it is not obvious that the increase in the real value of their national poverty line will increase the median national poverty line of low-income countries or lower-middle income countries. As long as the median income level *within each income group* has not changed between 2011 and 2017, then—even if countries' national poverty lines increase as they grow—it does not need to be the case that the global poverty lines we derive on expectation increase in real value.

7. Conclusion

In this paper, we analyze the impact of the 2017 PPPs on three important issues related to global poverty monitoring.

First, we analyze the stability of the 2017 PPPs relative to the 2011 PPPs, the previous round of PPPs. The adoption of previous ICP rounds has resulted in large revisions to global poverty estimates, especially the 2005 round that added 400 million more people to the estimated number of extremely poor people in the developing world (Chen and Ravallion 2010). Against this background, the Atkinson Commission on global poverty argued that future ICP rounds should not be adopted until 2030, the target date for the SDGs (World Bank 2017, also see Deaton 2001 and Klasen et al. 2016 for a similar argument made earlier). We show in this paper that the ICP methodology has remained fairly stable between the 2011 and 2017 rounds. While changes at the country level could still be important, we find no evidence of substantial broad-based changes which characterized earlier rounds and motivated the Atkinson recommendation.

Second, we estimate global poverty lines with the 2017 PPPs. The international poverty line (IPL) we derive to measure extreme poverty is \$2.15 per person per day in 2017 PPP. We employ the harmonized national poverty lines technique (Jolliffe and Prydz 2016) to arrive at this line, and show that it is consistent with an equivalent poverty lines approach (Kakwani and Son 2016). As such, the IPL we derive keeps the global poverty rate largely at the levels which were observed when the SDGs were set. Therefore, the 2017 PPPs do not significantly shift the goalposts. We estimate the line that is more typical of lower-middle-income countries (i.e., \$3.20 in 2011 PPP) to be \$3.65 in 2017 PPP, and the line common in upper-middle-income countries (i.e., \$5.50 in 2011 PPP) to be \$6.85 in 2017 PPP. The Bank's societal poverty line, which was originally defined as $\max(\$1.90, \$1 + 50\% \text{ of median consumption})$ with the 2011 PPPs, is also updated to $\max(\$2.15, \$1.15 + 50\% \text{ of median consumption})$ with the 2017 PPPs.

Third, we analyze the impact of these global poverty lines and the 2017 PPPs on the global poverty counts. While the changes at the upper-middle-income line are larger, the 2017 PPPs have small implications for extreme poverty and poverty at the lower-middle-income line. Between 1991 to 2017, extreme poverty in the world falls from 36.05% to 9.27% with the revised 2011 PPPs, and from 37.46% to 9.07% with the 2017 PPPs. This is equivalent to a decrease in the estimated number of poor people in the world by 15 million in 2017 (or 0.2pp). This is a very small change compared to prior PPP rounds, which is consistent with the methodological stability over the last two ICP rounds.

References

- Alkire, S., U. Kanagaratnam, and N. Suppa. 2020. “The Global Multidimensional Poverty Index (MPI).” 49. OPHI MPI Methodological Notes. Oxford: Oxford Poverty and Human Development Initiative, University of Oxford.
- Allen, Robert C. 2017. “Absolute Poverty: When Necessity Displaces Desire.” *American Economic Review* 107 (12): 3690–3721. <https://doi.org/10.1257/aer.20161080>.
- Atamanov, Aziz, R Andres Castaneda Aguilar, Tony H M J Fujs, Reno Dewina, Carolina Diaz-bonilla, Daniel Gerszon Mahler, Dean Jolliffe, et al. 2020. “March 2020 PovcalNet Update: What’s New.” Global Poverty Monitoring Technical Note 11.
- Atamanov, Aziz, Dean Jolliffe, Christoph Lakner, and Espen Beer Prydz. 2018. “Purchasing Power Parities Used in Global Poverty Measurement.” Global Poverty Monitoring Technical Note 5.
- Atamanov, Aziz, Christoph Lakner, Daniel Gerszon Mahler, Samuel Kofi Tetteh-Baah, and Judy Yang. 2020. “The Effect of New PPP Estimates on Global Poverty: A First Look,” Global Poverty Monitoring Technical Note 12.
- Bai, Yan, and William A. Masters. 2020. “Retail Food Prices at Purchasing Power Parity Exchange Rates: A First Look at Aggregate ICP 2017 Data.” World Bank Data Blog. 2020. <https://blogs.worldbank.org/opendata/retail-food-prices-purchasing-power-parity-exchange-rates-first-look-aggregate-icp-2017>.
- Berry, Francien, Brian Graf, Michael Stanger, and Mari Ylä-Jarkko. 2019. “Price Statistics Compilation in 196 Economies.” IMF Working Papers 19 (163). <https://doi.org/10.5089/9781513508313.001>.
- Chen, Shaohua, and Martin Ravallion. 2008. “China Is Poorer than We Thought, but No Less Successful in the Fight against Poverty.” World Bank Policy Research Working Paper 4621. Washington, DC.
- . 2010. “The Developing World Is Poorer than We Thought, but No Less Successful in the Fight against Poverty.” *The Quarterly Journal of Economics*, no. November: 1577–1625.
- Dalgaard, Esben, and Henrik Sejerbo Sørensen. 2002. “Consistency between PPP Benchmarks and National Price and Volume Indices.” In *OECD Meeting of National Accounts Experts*. Stockholm, Sweden.
- Deaton, Angus. 2001. “Counting the World’s Poor: Problems and Possible Solutions.” *World Bank Research Observer* 16 (2): 125–47. <https://doi.org/10.1093/wbro/16.2.125>.
- Deaton, Angus, and Bettina Aten. 2017. “Trying to Understand the PPPs in ICP 2011: Why Are the Results so Different.” *American Economic Journal: Macroeconomics* 9 (1): 243–64. <https://doi.org/10.1257/mac.20150153>.
- Deaton, Angus, and Alan Heston. 2010. “Understanding PPPs and PPP-Based National Accounts.” *American Economic Journal: Macroeconomics* 2 (4): 1–35.
- Deaton, Angus, and Paul Schreyer. 2022. “GDP, Wellbeing, and Health: Thoughts on the 2017 Round of the International Comparison Program.” *Review of Income and Wealth* 68(1): 1–15. <https://doi.org/10.1111/roiw.12520>.
- Ferreira, Francisco H. G., and Carolina Sánchez-Páramo. 2017. “A Richer Array of International Poverty Lines.” World Bank Blog: Let’s Talk Development. 2017. <https://blogs.worldbank.org/developmenttalk/richer-array-international-poverty-lines>.
- Ferreira, Francisco H.G., Shaohua Chen, Andrew Dabalen, Yuri Dikhanov, Nada Hamadeh, Dean Jolliffe, Ambar Narayan, et al. 2016. “A Global Count of the Extreme Poor in 2012: Data

- Issues, Methodology and Initial Results.” *Journal of Economic Inequality* 14 (2): 141–72. <https://doi.org/10.1007/s10888-016-9326-6>.
- Ferreira, Francisco H.G., Dean Jolliffe, and Espen Beer Prydz. 2015. “The International Poverty Line Has Just Been Raised to \$1.90 a Day, but Global Poverty Is Basically Unchanged. How Is That Even Possible?” World Bank Blog: Let’s Talk Development. 2015. <https://blogs.worldbank.org/developmenttalk/international-poverty-line-has-just-been-raised-190-day-global-poverty-basically-unchanged-how-even>.
- Inklaar, Robert, and D S Prasada Rao. 2017. “Cross-Country Income Levels over Time: Did the Developing World Suddenly Become Much Richer?” *American Economic Journal: Macroeconomics* 9 (1): 265–90.
- Inklaar, Robert, and Prasada Rao. 2020. “ICP PPP Time Series Implementation.” Washington, DC. <https://thedocs.worldbank.org/en/doc/f32e966db6ec404699b3381b32e2e589-0050022021/original/2-01-RA-Item-01-ICP-PPP-Time-Series-Implementation-Rao-and-Inklaar.pdf>.
- Jolliffe, Dean, and Espen Beer Prydz. 2016. “Estimating International Poverty Lines from Comparable National Thresholds.” *Journal of Economic Inequality* 14 (2): 185–98. <https://doi.org/10.1007/s10888-016-9327-5>.
- . 2021. Jolliffe, Dean, and Espen Beer Prydz. “Societal Poverty: A Relative and Relevant Measure.” *The World Bank Economic Review* 35, no. 1 (2021): 180–206. <https://doi.org/10.1093/wber/lhz018>.
- Kakwani, Nanak, and Hyun H. Son. 2016. “Global Poverty Estimates Based on 2011 Purchasing Power Parity: Where Should the New Poverty Line Be Drawn?” *Journal of Economic Inequality* 14 (2): 173–84. <https://doi.org/10.1007/s10888-016-9322-x>.
- Klasen, Stephan, Tatyana Krivobokova, Friederike Greb, Rahul Lahoti, Syamsul Hidayat, and Pasaribu Manuel. 2016. “International Income Poverty Measurement: Which Way Now?” *The Journal of Economic Inequality*, 199–225. <https://doi.org/10.1007/s10888-016-9324-8>.
- Krijnse Locker, H., and H. D. Faerber. 1984. “Space and Time Comparisons of Purchasing Power Parities and Real Values.” *Review of Income and Wealth* 30 (1): 53–83. <https://doi.org/10.1111/j.1475-4991.1984.tb00477.x>.
- Lakner, Christoph, Daniel Gerszon Mahler, Minh C Nguyen, Joao Pedro Azevedo, Shaohua Chen, and Dean Jolliffe. 2018. “Consumer Price Indices Used in Global Poverty Measurement.” Global Poverty Monitoring Technical Note 8.
- Mahler, Daniel Gerszon, R. Andres Castaneda Aguilar, and David Newhouse. 2021. “Nowcasting Global Poverty.” Policy Research Working Paper 9860. World Bank, Washington, DC.
- McCarthy, Paul. 2013. “Extrapolating PPPs and Comparing ICP Benchmark Results.” In *Measuring the Real Size of the World Economy*, 473–505. https://doi.org/10.1596/9780821397282_ch18.
- Prydz, Espen Beer, Dean Jolliffe, Christoph Lakner, Daniel Gerszon Mahler, and Prem Sangraula. 2019. “National Accounts Data Used in Global Poverty Measurement.” Global Poverty Monitoring Technical Note 8.
- Ravallion, Martin. 1998. “Poverty Lines in Theory and Practice.” LSMS Working Paper 133. The World Bank, Washington, DC.
- Ravallion, Martin. 2014. “An Exploration of the International Comparison Program’s New Global Economic Landscape.” National Bureau of Economic Research Working Paper 20338. <https://doi.org/10.3386/w20338>.
- . 2020. “Book Review of Measuring Poverty around the World.” *The Journal of Economic*

- Inequality* 18 (1): 131–36. <https://doi.org/10.1007/s10888-019-09430-w>.
- Ravallion, Martin, Shaohua Chen, and Prem Sangraula. 2009. “Dollar a Day Revisited.” *World Bank Economic Review* 23 (2): 163–84. <https://doi.org/10.1093/wber/lhp007>.
- Smeeding, Timothy M. 2016. “Poverty Measurement.” In *The Oxford Handbook of the Social Science of Poverty*, edited by David Brady and Linda M. Burton. Oxford.
- World Bank. 2008. *Global Purchasing Power Parities and Real Expenditures: 2005 International Comparison Program*. Washington, DC: World Bank.
- . 2015. *Purchasing Power Parities and the Real Size of World Economies: A Comprehensive Report of the 2011 International Comparison Program*. Washington, DC: World Bank.
- . 2016. “Monitoring Global Poverty: A Cover Note to the Report of the Commission on Global Poverty, Chaired by Prof. Sir Anthony B. Atkinson.” Washington, DC: World Bank. <http://pubdocs.worldbank.org/en/733161476724983858/MonitoringGlobalPovertyCoverNote.pdf>.
- . 2017. *Monitoring Global Poverty: Report of the Commission on Global Poverty*. Washington, DC: Washington, DC: World Bank.
- . 2018. *Poverty and Shared Prosperity Report 2018: Piecing Together the Poverty Puzzle*. Washington, DC: Washington, DC: World Bank.
- . 2020a. *Poverty and Shared Prosperity Report 2020: Reversing Reversals of Fortune*. Washington, DC: Washington, DC: World Bank.
- . 2020b. *Purchasing Power Parities and the Size of World Economies: Results from the 2017 International Comparison Program*. Washington, DC: World Bank. <https://doi.org/10.1596/978-1-4648-1530-0>.

The Price of Poverty: The Impact of the 2017 PPPs on the International Poverty Line and Global Poverty²⁸

Dean Jolliffe (r) Daniel Gerszon Mahler (r) Christoph Lakner (r)
Aziz Atamanov (r) Samuel Kofi Tetteh-Baah

Online Appendix

Appendix A. Revised 2011 PPPs

In May 2020, the ICP released new PPP estimates for the 2017 reference year, and also revised PPP estimates originally published for the 2011 round. The 2011 PPPs were revised primarily to account for revisions to national accounts expenditures. The ICP expenditure classification for the original 2011 round was based on the 1993 *System of National Accounts*. For the 2017 round, the ICP adopted the new 2008 *System of National Accounts*, which was also used to revise the original 2011 PPPs. The underlying price data were largely unchanged in the revised 2011 round. For non-benchmark economies—countries that do not collect prices to compute PPPs as part of the ICP exercise—new input data were used to impute PPPs. For details on the revisions, see World Bank (2020) and Tetteh-Baah et al. (2020). Using the harmonized poverty line approach (Jolliffe and Prydz 2016), the revised 2011 PPPs do not change the international poverty line of \$1.90 or the \$3.20 line (the \$5.50 line slightly increases by \$0.15) (Atamanov et al. 2020). Table A.1 updates the IPL with the revised 2011 PPPs using the approach by Ravallion et al. (2009). Table A.1 shows how revisions to the historic CPIs affect the IPL; when we take the national poverty lines denominated in the original 2011 PPPs in Ferreira et al. (2016) as given and consider only the revisions to the 2011 PPPs, the IPL remains at \$1.90.

Compared to the original 2011 PPPs, the revised 2011 PPPs increase extreme poverty in the world in 2017 by 0.24pp, representing 18.0 million more poor people (see Tables A.2). This result is mainly driven by the two poorest regions—Sub-Saharan Africa (11 million more poor people) and South Asia (11 million more poor people). These regional numbers are in turn largely driven by Nigeria (5 million poorer) and India (11 million poorer). Table A.3 shows the top 10 countries with the largest absolute changes in extreme poverty in 2017.

²⁸ The author ordering was constructed through American Economic Association's randomization tool (confirmation code: R-TVbwJ99ttd). Corresponding author: Christoph Lakner (clakner@worldbank.org). All authors are with the World Bank. We gratefully acknowledge financial support from the Knowledge for Change program and from the UK government through the Data and Evidence for Tackling Extreme Poverty (DEEP) Research Programme. The findings, interpretations, and conclusions expressed in this paper are entirely those of the authors. They do not necessarily represent the views of the World Bank and its affiliated organizations, or those of the Executive Directors of the World Bank or the governments they represent.

With a poverty line of \$3.20 a day, global poverty increases by 0.29pp, or 21.5 million. Poverty falls in East Asia & the Pacific (10 million), which is offset by large increases in South Asia (21 million) and Sub-Saharan Africa (10 million). These changes are largely driven by India, Indonesia, and Nigeria. See Tables A.4 and A.5 for the full set of results.

With a poverty line of \$5.50 a day, global poverty slightly increases by 0.07pp, or 5 million. Poverty again falls in East Asia & the Pacific (12 million), which is offset by a similar increase in poverty in South Asia (12 million). These changes are largely driven by Indonesia and India, respectively. See Tables A.6 and A.7 for the full set of results.

Table A.1: Updating the \$1.90 IPL with revised 2011 PPPs

Country	Poverty Line, Original 2011 PPP	Poverty Line, Update PPP only	Poverty Line, Update CPI only	Poverty Line, Update both PPP and CPI
	(1)	(2)	(3)	(4)
Ethiopia	2.03	1.98	2.02	1.97
Ghana	3.07	3.11	3.01	3.04
Gambia, The	1.82	1.81	1.82	1.81
Guinea-Bissau	2.16	2.08	2.16	2.08
Mali	2.15	2.13	2.15	2.13
Mozambique	1.26	1.24	1.33	1.30
Malawi	1.34	1.33	1.29	1.28
Niger	1.49	1.48	1.49	1.48
Nepal	1.47	1.47	1.38	1.38
Rwanda	1.50	1.47	1.51	1.49
Sierra Leone	2.73	2.64	2.10	2.04
Chad	1.28	1.29	1.36	1.36
Tajikistan	3.18	3.35	3.18	3.35
Tanzania	0.88	0.88	0.88	0.88
Uganda	1.77	1.77	1.71	1.72
Mean	1.88	1.87	1.83	1.82

Notes: Column (1) reproduces the national poverty lines from Ferreira et al. (2016). When only PPPs are updated, the IPL still rounds to \$1.87 (see column (2)). Since the derivation by Ferreira et al. (2016), some of the CPIs used to convert the national poverty lines to 2011 prices have been updated. When CPIs are updated, or when both CPIs and PPPs are updated, the IPL rounds to \$1.80 (columns (3) and (4)). This is mainly driven by CPI revisions in Ghana, Malawi, Sierra Leone, and Tajikistan. See Appendix E, section 1 for the derivations.

Table A.2: Global and regional changes in poverty in 2017 at \$1.90/day

<i>Region</i>	<i>Poverty rate, % (original)</i>	<i>Poverty rate, % (revised)</i>	<i>Change in poverty, pp</i>	<i>Change in poverty, %</i>	<i>Change in millions of poor</i>
World	9.04	9.27	0.24	2.65	17.98
East Asia & Pacific	1.61	1.41	-0.20	-12.27	-4.08
Europe & Central Asia	1.36	1.30	-0.06	-4.74	-0.32
Latin America & Caribbean	3.73	3.77	0.03	0.94	0.22
Middle East & North Africa	6.43	6.34	-0.09	-1.45	-0.36
Rest of the World	0.68	0.68	0.00	0.00	0.00
South Asia	9.02	9.65	0.63	7.00	11.33
Sub-Saharan Africa	40.12	41.18	1.06	2.65	11.18

Notes: This table compares extreme poverty between the original and revised 2011 PPPs at the global and regional levels. Extreme poverty is measured as the share of the population living below the international poverty line, which is \$1.90/day.

Table A.3: Countries with the largest changes in millions of poor in 2017 at \$1.90/day

<i>Country</i>	<i>Poverty rate, % (original)</i>	<i>Poverty rate, % (revised)</i>	<i>Change in poverty, pp</i>	<i>Change in poverty, %</i>	<i>Change in millions of poor</i>
India	9.71	10.55	0.85	8.71	11.32
Nigeria	38.65	41.36	2.71	7.02	5.18
Indonesia	5.71	4.46	-1.25	-21.92	-3.31
Ethiopia	23.14	24.66	1.52	6.56	1.62
Angola	41.18	45.07	3.89	9.44	1.16
Congo, Dem. Rep.	72.00	72.65	0.65	0.91	0.53
Cameroon	21.73	23.66	1.93	8.87	0.47
Egypt, Arab Rep.	2.58	2.97	0.39	15.03	0.37
Côte d'Ivoire	24.02	25.49	1.47	6.14	0.36
Myanmar	2.01	1.36	-0.65	-32.21	-0.35

Notes: This table compares extreme poverty in 2017 estimated with the original and revised 2011 PPPs at the country level. The global poverty line used is \$1.90/day. The top 10 countries with the largest absolute changes in millions of poor are shown here. Countries are ranked in descending order of absolute changes in millions of poor.

Table A.4: Global and regional changes in poverty in 2017 at \$3.20/day

<i>Region</i>	<i>Poverty rate, % (original)</i>	<i>Poverty rate, % (revised)</i>	<i>Change in poverty, pp</i>	<i>Change in poverty, %</i>	<i>Change in millions of poor</i>
World	23.96	24.25	0.29	1.20	21.52
East Asia & Pacific	8.92	8.43	-0.49	-5.53	-10.19
Europe & Central Asia	4.87	4.63	-0.23	-4.78	-1.14
Latin America & Caribbean	9.26	9.31	0.05	0.56	0.33
Middle East & North Africa	17.99	18.30	0.32	1.76	1.21
Rest of the World	0.89	0.89	0.00	0.00	0.00
South Asia	42.26	43.43	1.16	2.75	20.85
Sub-Saharan Africa	66.34	67.34	1.00	1.50	10.47

Notes: This table compares poverty in 2017 estimated with the original and revised 2011 PPPs at the global and regional levels. The poverty line used is \$3.20/day.

Table A.5: Countries with the largest change in millions of poor in 2017 at \$3.20/day

<i>Country</i>	<i>Poverty rate, % (original)</i>	<i>Poverty rate, % (revised)</i>	<i>Change in poverty, pp</i>	<i>Change in poverty, %</i>	<i>Change in millions of poor</i>
India	43.65	45.17	1.52	3.48	20.34
Indonesia	27.25	24.64	-2.62	-9.60	-6.93
Nigeria	69.34	71.74	2.41	3.47	4.59
Egypt, Arab Rep.	22.22	24.79	2.57	11.56	2.48
Ethiopia	59.11	61.33	2.22	3.76	2.36
Myanmar	19.35	14.95	-4.40	-22.72	-2.35
Angola	63.88	68.03	4.15	6.50	1.24
Iraq	15.12	12.19	-2.94	-19.41	-1.10
Cameroon	42.15	44.58	2.43	5.76	0.60
Bangladesh	47.85	47.49	-0.36	-0.75	-0.58

Notes: This table compares extreme poverty in 2017 estimated with the original and revised 2011 PPPs at the country level. The poverty line used is \$3.20/day. The top 10 countries with the largest absolute changes in millions of poor are shown here. Countries are ranked in descending order of absolute changes in millions of poor.

Table A.6: Global and regional changes in poverty in 2017 at \$5.50/day

<i>Region</i>	<i>Poverty rate, % (original)</i>	<i>Poverty rate, % (revised)</i>	<i>Change in poverty, pp</i>	<i>Change in poverty, %</i>	<i>Change in millions of poor</i>
World	43.47	43.53	0.07	0.15	4.93
East Asia & Pacific	28.21	27.62	-0.59	-2.10	-12.27
Europe & Central Asia	12.89	12.63	-0.27	-2.07	-1.31
Latin America & Caribbean	22.87	22.97	0.09	0.41	0.60
Middle East & North Africa	42.85	43.08	0.23	0.54	0.88
Rest of the World	1.29	1.29	0.00	0.27	0.04
South Asia	77.80	78.48	0.68	0.87	12.20
Sub-Saharan Africa	85.77	86.23	0.46	0.53	4.80

Notes: This table compares poverty in 2017 estimated with the original and revised 2011 PPPs at the global and regional levels. The poverty line used is \$5.50/day.

Table A.7: Countries with the largest change in millions of poor in 2017 at \$5.50/day

<i>Country</i>	<i>Poverty rate, % (original)</i>	<i>Poverty rate, % (revised)</i>	<i>Change in poverty, pp</i>	<i>Change in poverty, %</i>	<i>Change in millions of poor</i>
India	78.64	79.52	0.87	1.11	11.69
Indonesia	58.66	55.79	-2.87	-4.90	-7.60
Myanmar	60.81	54.30	-6.52	-10.71	-3.48
Nigeria	90.61	91.80	1.19	1.31	2.27
Egypt, Arab Rep.	66.76	69.11	2.34	3.51	2.26
Iraq	53.11	47.59	-5.52	-10.40	-2.07
Russian Federation	2.53	3.82	1.29	51.11	1.87
Iran, Islamic Rep.	11.30	13.24	1.94	17.16	1.56
Argentina	9.07	10.96	1.89	20.83	0.83
Angola	83.56	86.29	2.73	3.27	0.81

Notes: This table compares poverty in 2017 estimated with the original and revised 2011 PPPs at the country level. The poverty lines used is \$5.50/day. The top 10 countries with the largest absolute changes in millions of poor are shown here. Countries are ranked in descending order of absolute changes in millions of poor.

Appendix B. Additional tables and figures

Table B.1: Country-year observations of national poverty rates

Source	All	Matched into PovcalNet
Poverty and Equity database	1,315	1,315
OECD	123	66
Total	1,438	1,381

Notes: 1. Most of the national poverty rates were obtained from the World Bank's Poverty and Equity database. These data can be found in the following series from <https://databank.worldbank.org/source/poverty-and-equity>: *Poverty headcount ratio at national poverty lines (% of population), including noncomparable values (SI.POV.NAHC.NC)*. We downloaded the series on July 11, 2021.

2. The remaining data on national poverty rates were obtained from the Organization for Economic Co-operation and Development (OECD). We downloaded the series *PVT6A: Poverty rate after taxes and transfers, Poverty line 60% (i.e., share of population living on below 60% of median disposable income)* on December 19, 2020 from the OECD website: <https://stats.oecd.org/Index.aspx?DataSetCode=IDD#>. We use the OECD data only for countries with missing data in the Poverty and Equity database, or where the OECD time series is longer.

Table B.2: Exception 2011 PPPs for global poverty monitoring

Country	Original (official)	Original (imputed)	Revised (official)	Revised (imputed)
1. Egypt, Arab Rep.	1.80	2.78	1.71	2.87
2. Iraq	573.42	1003.80	477.56	939.22
3. Jordan	0.32	0.45	0.33	0.44
4. Lao PDR	2914.85	3325.20	3124.08	3248.44
5. Myanmar	275.83	320.60	278.39	296.14
6. Yemen, Rep.	82.09	111.30	76.77	109.53

Source: Atamanov et al. (2020), Table A.1.

Notes: See main text, Atamanov et al. (2020), Atamanov et al. (2018), and Appendix C for details. Whenever the (revised) 2011 PPPs are used in the main text, the PPPs in the last column are used.

Table B.3: Exception 2017 PPPs for global poverty monitoring

Country	Revised 2011 PPP	2017 PPP (official)	2017 PPP (imputed)	2017 PPP (exception)	Note
	(1)	(2)	(3)	(4)	(5)
1. Belize	1.17	1.48	1.16	1.31	Average
2. Egypt, Arab Rep.	2.87	3.41	6.88	4.84	Average
3. Guinea	2599.89	3213.98	3696.76	3446.93	Average
4. Iraq	939.22	555.39	619.35	586.50	Average
5. Nigeria	83.58	112.10	136.44	123.67	Average
6. São Tomé and Príncipe	10.49	10.76	12.14	11.43	Average
7. Sudan	1.46	5.38	7.77	6.46	Average
8. Trinidad and Tobago	4.52	4.21	5.16	4.66	Average
9. Kiribati	1.07			0.98	Extrapolated
10. Nauru	1.21			1.38	Extrapolated
11. Syrian Arab Republic	22.26			151.78	Extrapolated
12. Tuvalu	1.17			1.25	Extrapolated
13. Venezuela, RB	2.94			315.91	Extrapolated
14. Yemen, Rep.	109.53			255.68	Extrapolated

Notes: The PPPs in columns (1) and (4) are used in this paper. The geometric averages of official and imputed 2017 PPPs (column 4) are used for the exception countries. For countries without official 2017 PPPs, extrapolated PPPs are used: PPPs are extrapolated from the revised 2011 PPP currently used for global poverty monitoring, together with domestic and US inflation between 2011 and 2017. See Section 3.B of the main paper for details.

Table B.4: Rural/Urban PPPs used in global poverty monitoring

<i>Country, variable</i>	<i>Original 2011 PPP</i>	<i>Revised 2011 PPP</i>	<i>2017 PPP</i>
A) China			
Ratio of urban to rural poverty line (ω)	1.29	1.29	1.24
ICP urban share of outlets (λ)	0.76	0.76	0.79
Rural PPP	3.038	3.039	3.495
Urban PPP	3.904	3.905	4.318
National PPP	3.696	3.698	4.147
B) India			
Ratio of urban to rural poverty line (ω)	1.22	1.22	1.22
ICP urban share of outlets (λ)	0.74	0.74	0.64
Rural PPP	12.908	13.173	17.092
Urban PPP	15.695	16.018	20.787
National PPP	14.975	15.283	19.469
C) Indonesia			
Ratio of urban to rural poverty line (ω)	1.18	1.18	1.07
ICP urban share of outlets (λ)	0.61	0.61	0.98
Rural PPP	3678.414	3498.876	4746.852
Urban PPP	4352.751	4140.299	5098.196
National PPP	4091.939	3892.218	5089.686

Source: Adapted from Atamanov et al. (2020), Table A.2.

Note: See main text for details. The urban and rural PPPs are computed as follows: $Rural\ PPP = \frac{National\ PPP}{\omega \times \lambda + (1 - \lambda)}$

$Urban\ PPP = \omega * Rural\ PPP$. See online appendix of Ferreira et al. (2016) for details.

Table B.5: Cumulative mean and median of poverty lines of the poorest countries

No.	Country	Year	Harmonized national poverty lines, 2017 PPP		
			Country-specific	Cumulative mean	Cumulative median
1	Burundi	2013	2.14	2.14	2.14
2	Congo, Dem. Rep.	2012	1.90	2.02	2.02
3	Malawi	2016	1.68	1.91	1.90
4	Central African Republic	2008	2.16	1.97	2.02
5	Niger	2014	1.87	1.95	1.90
6	Mozambique	2014	1.49	1.88	1.89
7	Togo	2015	2.17	1.92	1.90
8	Liberia	2016	3.13	2.07	2.02
9	Madagascar	2012	1.64	2.02	1.90
10	Sierra Leone	2018	3.24	2.14	2.02
11	Chad	2011	2.66	2.19	2.14
12	Guinea-Bissau	2010	2.28	2.20	2.15
13	Ethiopia	2015	2.04	2.19	2.14
14	Burkina Faso	2014	2.16	2.18	2.15
15	Rwanda	2016	1.73	2.15	2.14
16	Guinea	2012	3.40	2.23	2.15
17	Mali	2009	1.95	2.22	2.14
18	Gambia, The	2015	3.87	2.31	2.15
19	Uganda	2016	1.49	2.26	2.14
20	Kiribati	2006	2.73	2.29	2.15
21	Nepal	2010	2.93	2.32	2.16
22	Tanzania	2017	1.62	2.29	2.15
23	Solomon Islands	2012	1.61	2.26	2.14
24	Senegal	2011	2.36	2.26	2.15
25	Lesotho	2017	3.22	2.30	2.16
26	Comoros	2004	4.18	2.37	2.16
27	Tajikistan	2015	4.01	2.43	2.16
28	Haiti	2012	3.67	2.48	2.17
29	Benin	2015	1.77	2.45	2.16
30	Zimbabwe	2017	2.01	2.44	2.16
31	Uzbekistan	2003	0.91	2.39	2.16
32	Papua New Guinea	2009	2.16	2.38	2.16
33	Vanuatu	2010	2.04	2.37	2.16
34	Timor-Leste	2014	3.59	2.41	2.16
35	Tuvalu	2010	4.16	2.46	2.16
36	Sudan	2009	3.58	2.49	2.16
37	Cameroon	2014	2.91	2.50	2.16
38	Zambia	2015	1.78	2.48	2.16
39	Micronesia, Federated States of	2013	3.66	2.51	2.16
40	Kenya	2015	2.41	2.51	2.17

Notes: Countries are shown in ascending order of GDP per capita (2017 PPP), starting with the poorest country. The full sample consists of 157 countries, excluding South Sudan, the Syrian Arab Republic, the Republic of Yemen (whose GDP per capita data are missing). Only the first 40 countries are listed here due to space constraints.

Table B.6: Updating global poverty lines using harmonized poverty lines: different options

Income classification	(A) Harmonized lines (Jolliffe and Prydz 2016)			(B) Own harmonized lines (pooled)			(C) Own harmonized lines (pooled, since 2007)		
	Median	Mean	N	Median	Mean	N	Median	Mean	N
Low	2.15	2.53	33	2.16	2.43	177	2.14	2.35	59
Lower-middle	3.68	4.49	32	3.80	4.20	394	3.59	4.04	214
Upper-middle	6.32	6.18	32	6.76	7.06	343	6.76	7.04	257
High income	23.65	23.34	29	24.15	23.19	460	23.97	23.08	359
Observations			126			1,374			889

Notes: Panel (A) uses the same implicit national poverty lines originally derived by Jolliffe and Prydz (2016). Panel (B) and (C) pool together the implicit national poverty lines that we have derived. Our total sample of implicit poverty lines is 1,381, but 7 lines from the 1980s drop out because they could not be matched into the World Bank's historical income classification, which begins from 1988. Panels (B) and (C) provide weighted medians and means for country-year observations based on the income classification in the year in which the surveys were conducted. We weight the observations to give greater importance to those close to 2017 and to give equal weight to each country: In the first step, we assign a survey conducted in 2017 a weight of 1, a survey conducted in 2016 or 2018 a weight of 1/2, a survey conducted in 2015 or 2019 a weight of 1/3, and so on. Secondly, we rescale the weights such that each country has a total weight of 1. For example, if a country has three harmonized poverty lines for 2016, 2017, and 2018, the corresponding weights will be 0.25, 0.5, and 0.25. The weights are based on the relevant data set ($N = 1,374$) or subsample in question ($N = 889$). The sample in panel (A) has one harmonized national poverty for each country, so each country has a weight of 1 (there are 126 countries in panel (A), 157 in panel (B), and 150 in panel (C)). The oldest lines ($n = 2$) in panel (A) are from 2001 (i.e., 10 years before the 2011 ICP reference year). Panel (C) includes only lines since 2007 (i.e., 10 years before the 2017 ICP). When one line is selected for each country from the most recent lines (since 2007), the proposed international poverty line would be \$2.14 (see Table B.7 below).

Table B.7: Updating global poverty lines using harmonized poverty lines: full sample vs. subsample

Income classification	(A) Full sample			(B) Subsample (since 2007)		
	Median	Mean	N	Median	Mean	N
Low	2.15	2.42	28	2.14	2.31	25
Lower-middle	3.63	3.95	54	3.59	3.93	51
Upper-middle	6.85	7.05	37	6.95	7.08	36
High income	24.36	23.36	38	24.36	23.36	38
Observations			157			150

Notes: Panel (A) replicates Panel (C) from Table 1 in the main text. The line that is closest to 2017 is selected, one for each country. If harmonized national poverty lines are available for 2016 and 2018 but not 2017, 2018 is selected. Panel (B) excludes 7 lines derived from surveys conducted more than 10 years before the 2017 ICP reference year: AZE2001, COM2004, JAM2004, KIR2006, SYR2003, UZB2003, VEN2006.

Table B.8: Equivalent international poverty line, 2017 PPP: different options

Region	Year	Headcount, % (Revised 2011 PPP)	Equivalent IPL (2017 PPP)	Share of global poor	Weighted Equivalent IPL
	(1)	(2)	(3)	(4)	(5)
World	2010	16.02	2.119	1.00	2.16
East Asia & Pacific	2010	10.79	1.973	0.19	0.38
Europe & Central Asia	2010	2.37	1.801	0.01	0.02
Latin America & Caribbean	2010	6.03	2.058	0.03	0.07
Middle East & North Africa	2010	2.04	2.215	0.01	0.01
Rest of the World	2010	0.50	2.076	0.00	0.01
South Asia	2010	25.95	2.146	0.38	0.82
Sub-Saharan Africa	2010	47.47	2.274	0.37	0.85
World	2012	12.89	2.128	1.00	2.19
East Asia & Pacific	2012	6.88	1.979	0.15	0.30
Europe & Central Asia	2012	1.88	1.801	0.01	0.03
Latin America & Caribbean	2012	4.63	2.072	0.03	0.06
Middle East & North Africa	2012	2.23	2.185	0.01	0.02
Rest of the World	2012	0.57	2.114	0.01	0.01
South Asia	2012	19.15	2.147	0.35	0.76
Sub-Saharan Africa	2012	43.90	2.284	0.44	1.01
World	2015	10.14	2.161	1.00	2.22
East Asia & Pacific	2015	2.06	2.012	0.06	0.11
Europe & Central Asia	2015	1.51	1.801	0.01	0.02
Latin America & Caribbean	2015	3.69	2.067	0.03	0.06
Middle East & North Africa	2015	4.28	2.143	0.02	0.05
Rest of the World	2015	0.73	2.132	0.01	0.02
South Asia	2015	13.18	2.148	0.31	0.67
Sub-Saharan Africa	2015	41.95	2.294	0.56	1.29
World	2017	9.27	2.173	1.00	2.23
East Asia & Pacific	2017	1.41	2.023	0.04	0.08
Europe & Central Asia	2017	1.30	1.801	0.01	0.02
Latin America & Caribbean	2017	3.77	2.055	0.03	0.07
Middle East & North Africa	2017	6.34	2.147	0.03	0.07
Rest of the World	2017	0.68	2.143	0.01	0.02
South Asia	2017	9.65	2.147	0.25	0.53
Sub-Saharan Africa	2017	41.18	2.297	0.62	1.43

Notes: This table shows equivalent poverty lines in column (3) that keep the global and regional poverty headcount ratios in column (2) constant in a pre-specified reference year (year given in column (1)). The results for the world in column (3) are presented in the main text in Table 2. Column (5) provides an alternative approach, where the equivalent international poverty line for the world is calculated as the weighted average of the regional equivalent poverty lines (in column (3)), with the weights being the regions' respective shares of total millions of poor in column (4). The weights are based on the revised 2011 PPPs (with a \$1.90 poverty line).

Table B.9: Updating the \$1.90 IPL with 2017 PPPs

<i>Country</i>	<i>Poverty Line Original 2011 PPP</i>	<i>Update PPP & CPI (<u>update</u> CPI in \$1.90)</i>	<i>Update PPP & CPI II (<u>do not update</u> CPI in \$1.90)</i>	<i>Take \$1.90 as given, update PPP & CPI</i>
	(1)	(2)	(3)	(4)
Ethiopia	2.03	2.37	2.38	2.23
Ghana	3.07	2.28	2.33	1.44
Gambia, The	1.82	1.88	1.88	1.97
Guinea-Bissau	2.16	2.45	2.45	2.15
Mali	2.15	2.49	2.49	2.20
Mozambique	1.26	1.44	1.37	2.06
Malawi	1.34	1.55	1.61	2.28
Niger	1.49	1.47	1.47	1.88
Nepal	1.47	1.83	1.95	2.52
Rwanda	1.5	1.81	1.79	2.27
Sierra Leone	2.73	2.88	3.73	2.60
Chad	1.28	1.60	1.51	2.24
Tajikistan	3.18	3.27	3.27	1.95
Tanzania	0.88	1.06	1.06	2.29
Uganda	1.77	1.90	1.96	2.11
Mean	1.88	2.02	2.08	2.15

Notes: Column (1) reproduces national poverty lines from Ferreira et al. (2016). Column (2) updates both PPPs and CPIs to the 2017 ICP reference year, while incorporating the revisions to the CPI data used in deriving the \$1.90 line. Column (3) updates both PPPs and CPIs to the 2017 ICP reference year, without incorporating the revisions to CPI data used in deriving the \$1.90 line. Column (4) takes the \$1.90 line as given for all countries and updates the \$1.90 line with both PPPs and updated CPIs to the 2017 reference year. See Appendix E, section 2 for the derivations. If we updated the IPL with US inflation between 2011 and 2017 (i.e., approximately 9%), the new IPL would be \$2.07.

Table B.10: Global and regional changes in poverty in 2017 at \$3.20/day

<i>Region</i>	<i>Poverty rate, % (2011 PPP)</i>	<i>Poverty rate, % (2017 PPP)</i>	<i>Change in poverty, pp</i>	<i>Change in poverty, %</i>	<i>Change in millions of poor</i>
World	24.25	24.82	0.57	2.35	42.79
East Asia & Pacific	8.43	10.52	2.09	24.85	43.30
Europe & Central Asia	4.63	6.70	2.06	44.51	10.13
Latin America & Caribbean	9.31	10.14	0.83	8.87	5.21
Middle East & North Africa	18.30	16.32	-1.99	-10.85	-7.57
Rest of the World	0.89	0.91	0.02	2.31	0.23
South Asia	43.43	44.29	0.86	1.98	15.43
Sub-Saharan Africa	67.34	65.06	-2.28	-3.39	-23.94

Notes: This table compares poverty in 2017 estimated with the revised 2011 and 2017 PPPs at the global and regional levels. The poverty lines used are \$3.20/day (2011 PPP) and \$3.65/day (2017 PPP).

Table B.11: Countries with the largest changes in millions of poor in 2017 at \$3.20/day

<i>Country</i>	<i>Poverty rate, % (2011 PPP)</i>	<i>Poverty rate, % (2017 PPP)</i>	<i>Change in poverty, pp</i>	<i>Change in poverty, %</i>	<i>Change in millions of poor</i>
China	3.80	5.72	1.92	50.41	26.56
Nigeria	71.74	64.67	-7.08	-9.86	-13.50
Indonesia	24.64	29.59	4.95	20.10	13.11
Pakistan	33.95	38.86	4.91	14.47	10.21
Uzbekistan	43.66	71.46	27.80	63.68	9.00
India	45.17	45.81	0.64	1.41	8.53
Angola	68.03	47.81	-20.23	-29.73	-6.03
Egypt, Arab Rep.	24.79	18.58	-6.21	-25.05	-5.99
Ghana	28.72	48.13	19.41	67.56	5.65
Ethiopia	61.33	56.50	-4.83	-7.88	-5.14

Notes: This table compares poverty in 2017 estimated with the revised 2011 and 2017 PPPs at the country level. The poverty lines used are \$3.20/day (2011 PPP) and \$3.65/day (2017 PPP). The top 10 countries with the largest absolute changes in millions of poor are shown here. Countries are ranked in descending order of absolute changes in millions of poor.

Table B.12: Global and regional changes in poverty in 2017 at \$5.50/day

<i>Region</i>	<i>Poverty rate, % (2011 PPP)</i>	<i>Poverty rate, % (2017 PPP)</i>	<i>Change in poverty, pp</i>	<i>Change in poverty, %</i>	<i>Change in millions of poor</i>
World	43.53	47.81	4.28	9.83	321.22
East Asia & Pacific	27.62	36.21	8.60	31.13	177.79
Europe & Central Asia	12.63	16.43	3.81	30.15	18.70
Latin America & Caribbean	22.97	27.59	4.63	20.15	29.16
Middle East & North Africa	43.08	44.09	1.00	2.33	3.83
Rest of the World	1.29	1.49	0.20	15.56	2.20
South Asia	78.48	82.66	4.18	5.33	74.96
Sub-Saharan Africa	86.23	87.61	1.39	1.61	14.58

Notes: This table compares poverty in 2017 estimated with the revised 2011 and 2017 PPPs at the global and regional levels. The poverty lines used are \$5.50/day (2011 PPP) and \$6.85/day (2017 PPP).

Table B.13: Countries with the largest changes in millions of poor in 2017 at \$5.50/day

<i>Country</i>	<i>Poverty rate, % (2011 PPP)</i>	<i>Poverty rate, % (2017 PPP)</i>	<i>Change in poverty, pp</i>	<i>Change in poverty, %</i>	<i>Change in millions of poor</i>
China	19.88	29.18	9.30	46.77	128.91
India	79.52	83.37	3.85	4.85	51.59
Indonesia	55.79	65.17	9.38	16.81	24.82
Pakistan	74.52	82.05	7.53	10.10	15.65
Brazil	20.23	25.93	5.70	28.15	11.84
Iraq	47.59	21.04	-26.55	-55.79	-9.97
Myanmar	54.3	68.2	13.91	25.61	7.42
Ghana	54.34	77.94	23.60	43.44	6.87
Philippines	60.02	66.25	6.23	10.38	6.55
Iran, Islamic Rep.	13.24	20.94	7.70	58.13	6.21

Notes: This table compares poverty in 2017 estimated with the revised 2011 and 2017 PPPs at the country level. The global poverty lines used are \$5.50/day (2011 PPP) and \$6.85/day (2017 PPP). The top 10 countries with the largest absolute changes in millions of poor are shown here. Countries are ranked in descending order of absolute changes in millions of poor.

Table B.14: Global and regional changes in societal poverty in 2017

<i>Region</i>	<i>Poverty rate, % (2011 PPP)</i>	<i>Poverty rate, % (2017 PPP)</i>	<i>Change in poverty, pp</i>	<i>Change in poverty, %</i>	<i>Change in millions of poor</i>
World	27.31	27.55	0.23	0.86	17.59
East Asia & Pacific	22.58	23.55	0.97	4.29	20.04
Europe & Central Asia	17.10	18.18	1.09	6.35	5.33
Latin America & Caribbean	26.36	26.67	0.31	1.18	1.96
Middle East & North Africa	26.01	25.47	-0.54	-2.09	-2.07
Rest of the World	15.19	15.25	0.06	0.39	0.66
South Asia	31.32	31.84	0.52	1.65	9.29
Sub-Saharan Africa	48.28	46.60	-1.68	-3.48	-17.62

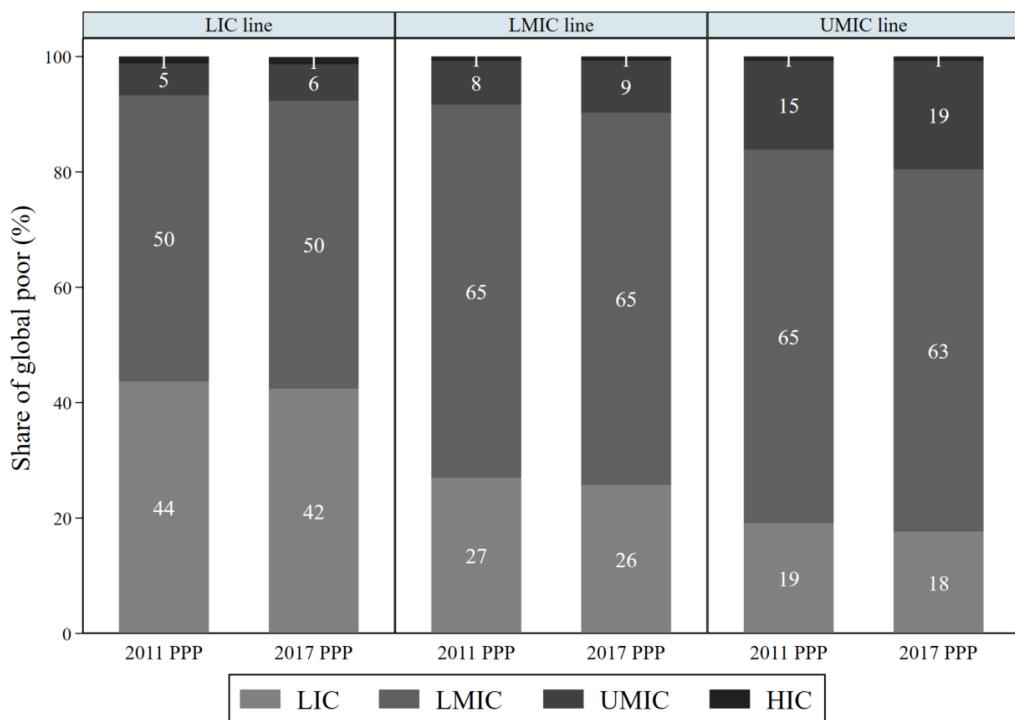
Notes: This table compares societal poverty in 2017 estimated with the revised 2011 and 2017 PPPs at the global and regional levels. With the revised 2011 PPP, the societal poverty line (SPL) used is given as $\max(\$1.90, \$1.00 + 0.5 * \text{median consumption/income})$. With the 2017 PPP, the societal poverty line (SPL) used is given as $\max(\$2.15, \$1.15 + 0.5 * \text{median consumption/income})$.

Table B.15: Countries with the largest changes in millions living in societal poverty in 2017

<i>Country</i>	<i>Poverty rate, % (2011 PPP)</i>	<i>Poverty rate, % (2017 PPP)</i>	<i>Change in poverty, pp</i>	<i>Change in poverty, %</i>	<i>Change in millions of poor</i>
China	21.04	22.07	1.03	4.88	14.25
Nigeria	41.36	37.57	-3.79	-9.16	-7.23
Congo, Dem. Rep.	72.65	64.42	-8.24	-11.34	-6.7
India	31.43	31.87	0.45	1.42	5.98
Uzbekistan	32.18	44.93	12.75	39.62	4.13
Indonesia	29.36	30.86	1.49	5.09	3.95
Pakistan	30.88	32.69	1.81	5.87	3.77
Angola	51.81	43.58	-8.24	-15.9	-2.46
Iraq	21.31	15.82	-5.49	-25.78	-2.06
Egypt, Arab Rep.	30.53	28.54	-1.99	-6.52	-1.92

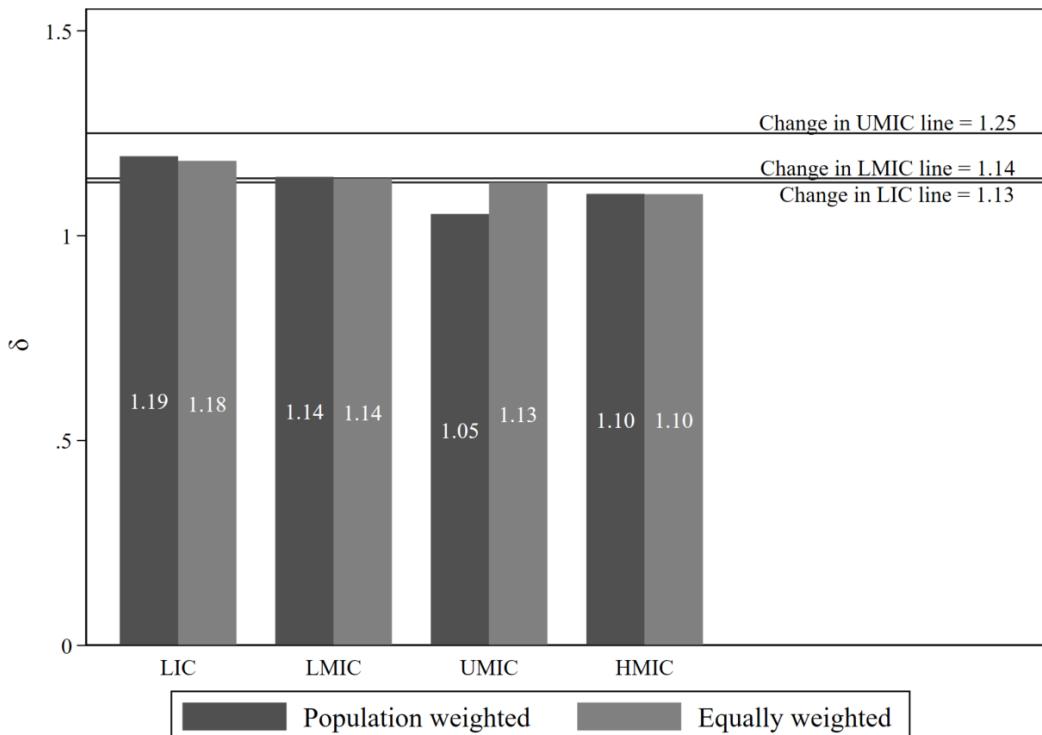
Notes: This table compares societal poverty in 2017 estimated with the revised 2011 and 2017 PPPs at the country level. With the revised 2011 PPP, the societal poverty line (SPL) used is given as $\max(\$1.90, \$1.00 + 0.5 * \text{median consumption/income})$. With the 2017 PPP, the societal poverty line (SPL) used is given as $\max(\$2.15, \$1.15 + 0.5 * \text{median consumption/income})$. The top 10 countries with the largest absolute changes in millions of poor are shown here. Countries are ranked in descending order of absolute changes in millions of poor.

Figure B.1: Share of global poor by income group in 2017



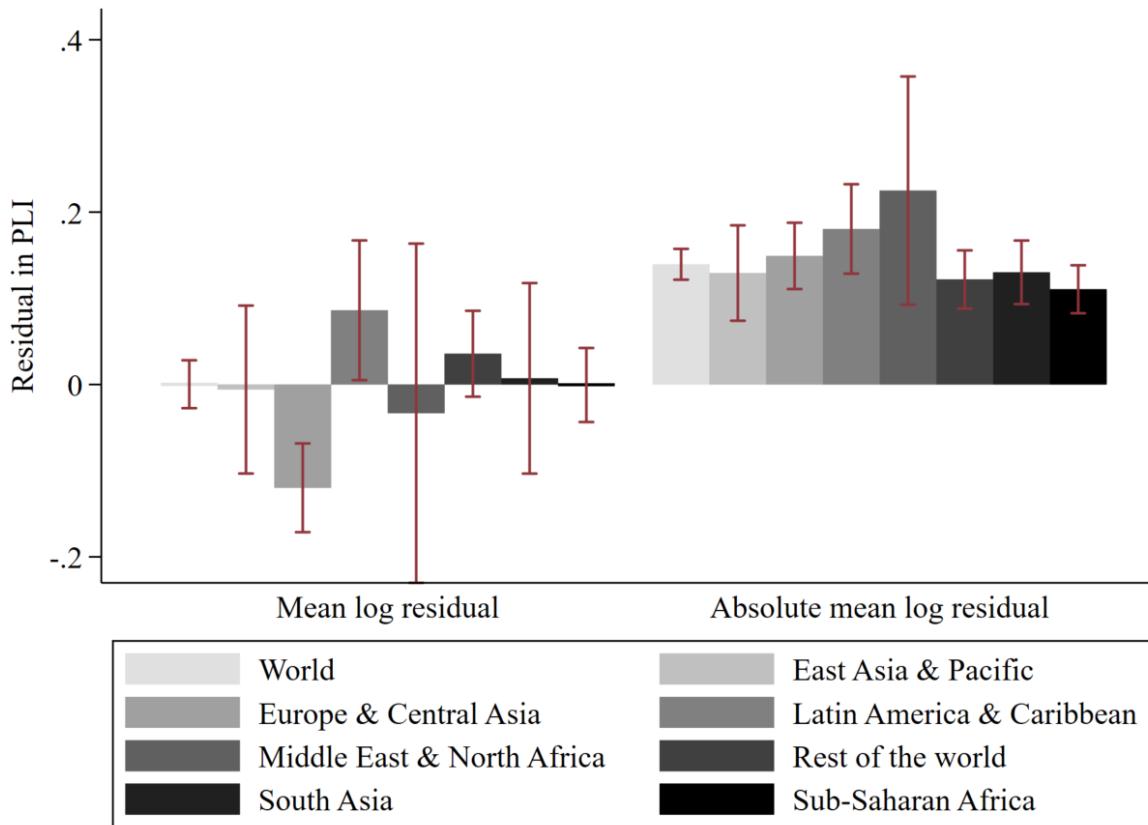
Notes: This chart shows the share of global poor in 2017 in lower-income countries (LIC), lower-middle-income countries (LMIC), upper-middle-income countries (UMIC) and high-income countries (HIC) based on the LIC, LMIC, and UMIC global poverty lines.

Figure B.2: Delta ratio by income group



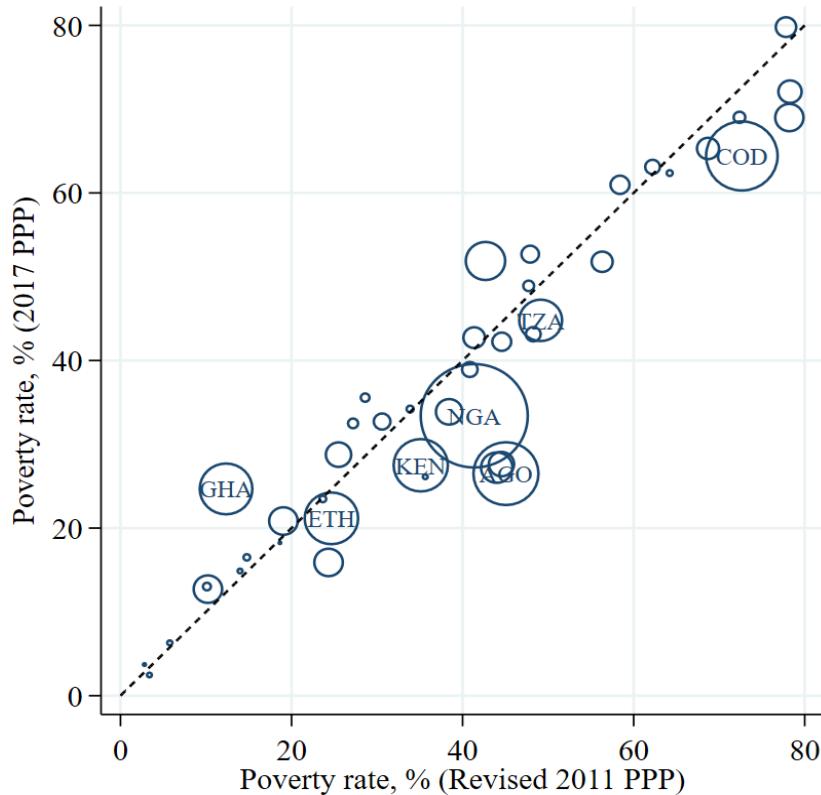
Notes: This chart shows the mean delta ratio for each income group based on the World Bank's income classification for 2017 (see more details on the delta ratio in Section 3 of the paper). The delta ratios are either weighted by countries' population sizes in 2017 or weighted equally. The weights are population sizes of countries in 2017. The relative changes in global poverty lines derived with data from low-income countries (LIC), lower-middle-income countries (LMIC) and upper-middle-income countries (UMIC) when moving to the 2017 PPPs are marked in the chart.

Figure B.3: Comparison of residuals in price level indices (PLIs) across regions



Notes: This graph shows the mean log residuals and absolute mean log residuals in price level indices (PLIs) for the 2017 PPP for household consumption expenditure across regions. The residual is defined as the published PLI (log) minus predicted PLI (log). Error bars represent 95% confidence intervals.

Figure B.4: Changes in extreme poverty in Sub-Saharan Africa, 2017



Notes: This chart shows estimates of extreme poverty in 2017 for countries in Sub-Saharan Africa. For countries, without a survey in 2017, the estimates are based on extrapolations or interpolations from recent surveys. Extreme poverty is measured using the IPL of \$1.9 (revised 2011 PPP) or \$2.15 (2017 PPP). Marker size is proportional to absolute change in millions of poor. The dotted line is a 45-degree line.

Table B.16: Decomposition of changes to global poverty lines

	<i>Jolliffe and Prydz (original 2011 PPP)</i> (1)	<i>Update with revised 2011 PPP</i> (2)	<i>Update with 2017 PPP</i> (3)	<i>Update national poverty lines</i> (4)	<i>Obs</i>	<i>Update income groups</i> (5)	<i>Obs</i>	<i>Update country availa- bility</i> (6)	
Low	1.91	1.85	2.15	2.17	33	2.16	25	2.15	28
Lower-middle	3.21	3.21	3.68	3.76	32	3.58	37	3.63	54
Upper-middle	5.47	5.65	6.32	7.15	32	6.95	30	6.85	37
High-income	21.70	21.70	23.65	27.39	29	24.31	34	24.36	38

Notes: The table shows what is driving the changes to the global poverty lines, starting from the global lines replicating Jolliffe and Prydz (2016) in column (1), to the final global poverty lines derived in this paper in column (6). The intermediate columns all make one change to the derivation of the harmonized national poverty lines, allowing for a (path dependent) decomposition of what is driving the final derivation. Columns (2) and (3) update the harmonized poverty lines from Jolliffe and Prydz (2016) with the revised 2011 PPPs and 2017 PPPs, respectively. Column (4) updates with new harmonized poverty lines derived from new survey data for the exact same countries that Jolliffe and Prydz (2016) used. Column (5) updates the income classification for the countries that with new data have changed income groups. Column (6) adds the countries for which national poverty lines have been made available after Jolliffe and Prydz (2016) made their analysis and removes the countries for which the national poverty lines or welfare aggregates have been removed due to quality concerns.

Appendix C. More details on the stability of the ICP methodology

The *System of National Accounts* (SNA) is the most widely used framework for measuring economic activities across countries. The conceptual framework of the ICP methodology is centered around the SNA's definition of GDP from the expenditure side; the sum of final expenditures on consumption, gross capital formation, and net exports (World Bank 2013; 2020). The most recent ICP rounds were based on different versions of the SNA: the 2005 and original 2011 rounds were based on the 1993 SNA, whereas the revised 2011 and 2017 rounds were based on the 2008 SNA. Even though both versions share the same fundamental structure, the 2008 SNA introduced some changes to the 1993 SNA (OECD 2013; World Bank 2020). These changes have implications for GDP, either by re-allocating components and/or changing the level of GDP. For example, research and development (R&D) has been capitalized for the first time, which has two implications: (1) it re-classifies R&D from *government final consumption expenditure* to *government gross fixed capital formation*, and (2) it adds to the level of GDP via the cost of depreciation that has to be imputed and included (OECD 2013).²⁹

A few changes have been introduced into the ICP expenditure classification in the 2017 round at different levels of aggregation, mainly in light of the SNA changes. The ICP classifies expenditure and price data as main aggregates, categories, groups, classes, and basic headings, in descending order.³⁰ For example, *gross capital formation* is a new main aggregate of GDP that now combines two main aggregates in the (original) 2011 round (i.e., *gross fixed capital formation* and *changes in inventories and acquisitions less disposal of valuables*). Even though there are still 155 basic headings in the 2017 round as in the 2011 round, some basic headings have been merged (e.g., *opening value of inventories* and *closing value of inventories* in the 2011 round have been merged as *changes in inventories* in the 2017 round) and new ones have been added (e.g., *actual and imputed rentals for housing* was a basic heading in the 2011 round but is now separated as *actual rentals for housing* and *imputed rentals for housing* under the housing rentals category). A key point to note is that aggregation is done from the level of basic headings up to the desired aggregate (e.g., GDP, household final consumption, or other levels) so that all these changes at the level of basic headings do not have a significant impact on the aggregates. Apart from changes in national accounts structures caused by SNA revisions, national accounts expenditure data of countries get routinely revised or rebased to incorporate new information. This primarily explains the revisions to the original 2011 PPPs.

²⁹ As an example, OECD (2013) shows that the capitalization of R&D increases Australia's GDP by 1.25-1.5%. Other SNA changes relate to the classification and/or valuation of weapon systems and ammunitions, computer software and databases, financial intermediation services indirectly measured (FISIM), output of central banks, and output for own use, among others. A more detailed description of the SNA changes can be found in OECD (2013).

³⁰ A basic heading consists of a well-defined group of goods or services at the lowest level of ICP expenditure classification. *Rice and bread* are examples of basic headings under the class of *bread and cereals*, which forms a part of the group called *food*. *Food* is within the *Food and nonalcoholic beverages* category, under the main aggregate of *individual consumption by households*. PPPs are first estimated at the level of basic headings and are aggregated along higher levels of aggregation up to the level of GDP, which sums up the main aggregates of expenditure.

The overall methodology for selecting the basket of goods and services the ICP collects prices on has remained quite stable. The ICP relies on sampling theory to determine the number of products to be priced in each basic heading, the number of survey outlets, and the number of times surveys will be conducted in the reference year. For example, only 10-15 items may be priced for the rice basic heading while 70-100 items may be priced for the more heterogeneous garments basic heading to obtain a similar level of precision in the estimates for the basic-headings PPPs (World Bank 2013). The exact items that are priced, as well as where or when surveys are exactly conducted, are subject to expert advice *ex-ante* and validation *ex-post*. The list of items also requires at least partial updates for each comparison (e.g., IT equipment, cars, home entertainment, etc.). In the 2005 round, there was a separate price and expenditure survey with a “ring” list of 1,000 items that were priced in 18 selected “ring” countries across all regions, to link prices within and across regions. The ring list of items has been found to have a bias for traded goods, which are often unrepresentative, scarce, and expensive in developing countries (Deaton and Aten 2017). As a result, the 2005 round overstated price levels in most developing countries. The lessons learned from the 2005 round brought about significant improvements in ICP product specification. In subsequent rounds, the list of items in the price surveys conducted in all participating economies has been carefully chosen to ensure that the items are both representative within countries and comparable across countries, as best as possible (World Bank 2013; 2015; 2020). A global core list (GCL) of items is priced as part of the price and expenditure surveys, for linking prices within and across regions. The GCL items are marked as “important” or “less important” and prices are weighted accordingly, so that price levels in developing countries are not overstated. Each country decides its own sets of important and less important items, depending on their national consumption patterns. Also, importance indicators apply only within basic headings, not across. While the general approach of obtaining the price and expenditure data remains similar, especially between the 2011 and 2017 rounds, data quality assurance has been improved with the latest available technology, particularly in the 2017 round (World Bank 2020).

As described in more detail in World Bank (2013; 2015; 2020), the computational steps the ICP follows to produce the final global PPPs have also remained largely unchanged. Since the major change involving the linking of regional prices—namely, the 2005 “ring” method vs. 2011 or 2017 “Global Core List” (GCL) method—the common elements of PPP estimation have not changed. At the global level, the Country Product Dummy (CPD) method is used to estimate PPPs. This method estimates a regression model that produces in one step PPPs that are both transitive and base-country invariant. The resulting PPPs are multilateral (i.e., the basic-heading PPPs for one economy depend on those of all other participating economies). PPPs are calculated for countries within regions, first at the basic heading level using the Country Product Dummy (CPD) method and then at the aggregate level using the Gini, Eltetö and Köves, and Szulc (GEKS) method, using prices in the GCL and a regional list, and expressed in the currency of a reference country in the region. Finally, the PPPs are linked across regions solely using items in the GCL to produce the

global PPPs, expressed in US dollars, in a way that ensures regional fixity (i.e., the ratio of real expenditures between any pair of participating economies within a region remains the same after linking prices across regions).

The ICP data coverage varies across rounds, which potentially impacts the global PPPs used for measuring global poverty. Given that the PPPs are multilateral, PPPs would, in principle, be affected if the number of economies changes between any two ICP rounds. In practice, however, the total number of countries participating is less of an issue due to the regional fixity principle. In particular, the differences in country coverage between the 2011 and 2017 rounds mainly come from small Pacific Islands (e.g., Cook Islands and Kiribati) that participated in the 2011 round. However, since these countries were linked via three bridge-countries that participated from other regions—Australia (Eurostat-OECD), Fiji (Asia and the Pacific), and New Zealand (Eurostat-OECD)—the inclusion of these islands did not impact the rest of the countries in any manner. Overall, comparability of different ICP rounds may still be an issue in cases where countries move regions (e.g., Colombia and Costa Rica have been moved from Latin America and the Caribbean to OECD in the 2017 round) or where countries join or drop from the comparison altogether.

Within countries, the price data coverage also varies between rounds, although with the currently available metadata it is not possible to assess this everywhere for past ICP rounds. In the 2005 round, for instance, price survey outlets were predominantly in capital cities and urban areas (World Bank 2008). In the 2017 round, this is still the case in some countries, including many high-income countries. Survey frames do not necessarily imply overestimation or underestimation of national prices, as the ICP requests countries to submit nationally representative annual prices. However, the ICP acknowledges that national average prices are often hard to achieve in large economies with settlements and populations that are substantially rural (World Bank 2015), and ultimately the ICP does not know whether prices have been adjusted appropriately.

Appendix D. Deciding when to impute 2017 PPPs for exception countries

While we argue that the ICP methodology has largely stabilized, the change from 2011 to 2017 PPPs can still lead to erratic changes in poverty estimates at the country level. In this part of the Appendix, we describe the analysis we have done to identify the countries where we have decided to deviate from the official 2017 PPPs.

Step 1: Select all countries where it is worthwhile to take a deeper look

Countries that fulfill any of the following conditions are selected for more detailed analysis:

- a) *Is there a large change in the real value of welfare when moving to the 2017 PPPs?*

As described in the main text and following Ferreira et al. (2016), we select all countries with delta ratios more than two standard deviations from the mean using the CPIs and revised 2011 PPPs used for global poverty monitoring. From Figure 2A (in the main text), we select Angola, Belize, the Arab Republic of Egypt, Ghana, Guinea, Iraq, Jordan, Liberia, São Tomé and Príncipe, and Suriname.

- b) *Is there a big difference in the movements of CPIs and PPPs?*

We select all countries with delta ratios more than two standard deviations from the mean using *IFS* CPIs and *official* 2011 PPPs. This is very similar to the previous condition, except for the slightly different CPIs and PPPs (see the main text for details). From Figure 2B (in the main text), we select Angola, Belize, Bolivia, Central African Republic, Guinea, Liberia, Nigeria, São Tomé and Príncipe, Suriname, and Trinidad and Tobago.

- c) *Are the 2017 PPPs very different from what we would expect based on the country's characteristics?*

We select all countries whose log difference in the PPP residual is more than two standard deviations from the mean. The PPP residual is defined as the difference in the logs of the published PPP and predicted PPP. PPPs are predicted using the seemingly unrelated regressions (SUR) model the ICP uses to predict PPPs for non-benchmark countries.

The ICP model for predicting 2017 PPPs for non-benchmark countries is given as:

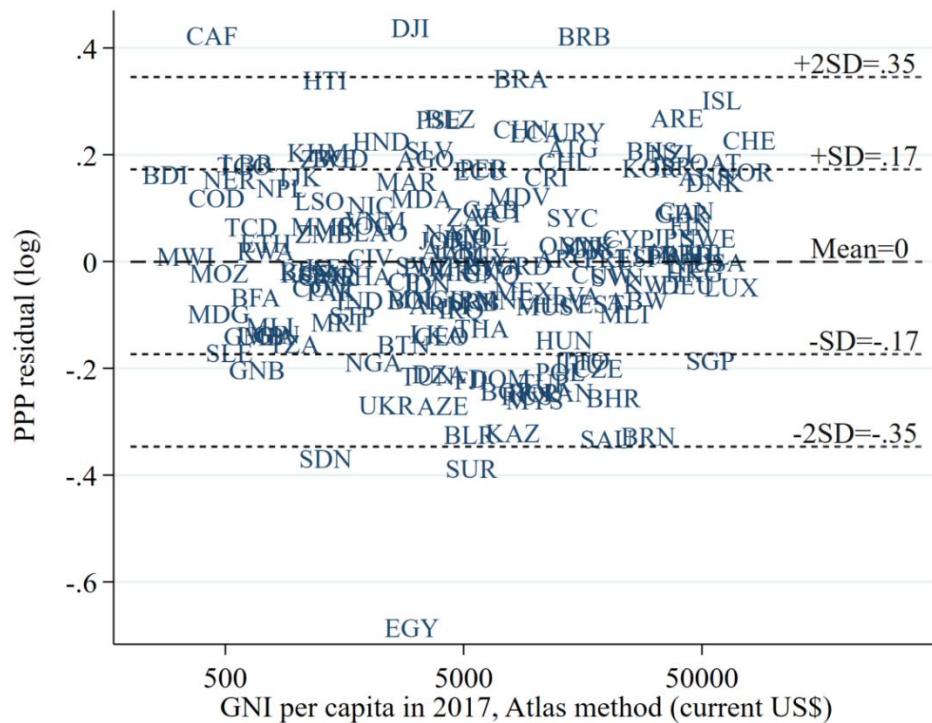
$$PLI_i - PLI_{USA} = b * (X_i - X_{USA}) + e_i$$

where PLI_i is the price level index of country i , calculated as the ratio of the PPP conversion factor to the market exchange rate, and X is a vector of explanatory variables: GDP per capita in U.S. dollars (based on market exchange rates), imports as a share of GDP, exports as a share of GDP, age dependency ratio, and dummies for Sub-Saharan

Africa, the EU, island economies, and landlocked developing economies. Interaction terms between GDP per capita and the country-group-dummy variables are also included.³¹

From Figure D.1, we select Barbados, Central African Republic, Egypt, Djibouti, Sudan, and Suriname. Barbados has no surveys in PovcalNet, so we drop it from the set of countries.

Figure D.1: Outlier countries based on the residual criterion



d) Are revised 2011 PPPs currently imputed for global poverty monitoring?

We also include countries whose revised 2011 PPPs are currently imputed (Ferreira et al. 2016; Atamanov et al. 2018; 2020). This adds Egypt, Iraq, Jordan, the Lao People's Democratic Republic, Myanmar, and the Republic of Yemen. The Republic of Yemen is not included in the rest of the investigation because it has no official 2017 PPP.

In summary, in step 1 of the analysis we select the following countries for a more detailed analysis: Angola, Belize, Bolivia, Central African Republic, Djibouti, Egypt, Ghana, Guinea, Iraq, Jordan,

³¹ Note that this model is not targeted towards predicting the price levels of the poor, for which other variables may be more relevant. This is a general issue with PPPs and CPIs used for global poverty monitoring. We thank Stefan Dercon for raising this issue. We nonetheless use the model by ICP to stay consistent with the PPPs used for benchmark countries, which likewise are not tailored towards the prices faced by the poor.

Lao PDR, Liberia, Myanmar, Nigeria, São Tomé and Príncipe, Sudan, Suriname, Trinidad and Tobago.

Step 2: Remove countries where imputing PPPs would lead to even larger changes

For the countries identified in steps 1a and 1b, we predict PPPs using the ICP model (while excluding all the countries that have been identified in step 1). If the predicted PPP makes the delta ratio more extreme, the country is removed from the list of countries worthwhile a deeper look. The reason for this is that our proposed solution to dealing with countries selected as outliers will be to partially rely on imputed PPPs. If imputed PPPs make the country an even stronger outlier, our proposed solution will not improve the situation. From Table D.1, using imputed PPPs would not improve the situation for Angola, Central African Republic, Ghana, Jordan, and Liberia, so they are excluded from the subsequent analysis.

Table D.1: Delta ratios with official and imputed 2017 PPP

Country	Delta ratio		Delta ratio increases/decreases further?
	Official PPP	Imputed PPP	
1. Angola	1.80	2.14	Yes
2. Belize	0.83	1.05	No
3. Bolivia	1.43	1.33	No
4. Central African Republic	1.24	1.93	Yes
5. Egypt, Arab Rep.	1.77	0.88	No
6. Ghana	0.75	0.73	Yes
7. Guinea	1.46	1.27	No
8. Iraq	1.90	1.70	No
9. Jordan	1.53	1.57	Yes
10. Liberia	1.50	1.81	Yes
11. Nigeria	1.44	1.18	No
12. São Tomé and Príncipe	1.46	1.30	No
13. Suriname	1.45	0.96	No
14. Trinidad and Tobago	1.43	1.17	No

Note: The countries whose delta ratios become more extreme with the imputed PPP are in **bold**. When the delta ratio increases further, poverty decreases further. When the delta ratio decreases further, poverty increases further. Also see Figure D.3, panel (A) below.

3: Criteria to suggest that the 2017 PPPs may be problematic

We use four criteria to evaluate whether the 2017 PPP for a country may be problematic. Given that no information can cast definite doubt on the 2017 PPPs—if so, the ICP team would have

dealt with the issue already—each of these criteria provides a signal of evidence on whether the 2017 PPPs are fit for global poverty monitoring.

a) *Are the CPIs fit for purpose?*

As described in the main text, an extreme delta ratio could arise from issues with either PPP or the CPI. We evaluate the quality of the CPI by whether it classifies expenditures using COICOP, and whether the weights have national coverage and are no older than 10 years from 2017. If none of these issues applies, we do not have evidence that there is a problem with the CPIs, which may suggest that the problem could be with the PPPs, creating an argument for imputing PPPs (see Table D.2).

Table D.2: CPI quality indicators

<i>Country</i>	<i>CPI classification system</i>	<i>CPI expenditure weights coverage</i>	<i>CPI weights reference year</i>	<i>Flag</i>
1. Belize	COICOP	National	2009	Yes
2. Bolivia	Other	National	2016	No
3. Djibouti	Other	Capital City	2013	No
4. Egypt, Arab Rep.	COICOP	National	2010	Yes
5. Guinea	Other	Capital City	2002	No
6. Iraq	COICOP	National	2012	Yes
7. Lao PDR	Other	National	2010	No
8. Myanmar	Other	National	2012	No
9. Nigeria	Other	National	2004	No
10. São Tomé and Príncipe	COICOP	National	2014	Yes
11. Sudan	Other	National	2007	No
12. Suriname	COICOP	National	2014	Yes
13. Trinidad and Tobago	COICOP	National	2009	Yes
14. <i>Angola</i>	<i>COICOP</i>	<i>National</i>	2009	Yes
15. <i>Central African Republic</i>	Other	<i>National</i>	2005	No
16. <i>Ghana</i>	<i>Survey</i>	<i>National</i>	2005.67	No
17. <i>Jordan</i>	Other	<i>National</i>	2010	No
18. <i>Liberia</i>	<i>COICOP</i>	<i>National</i>	2016	Yes

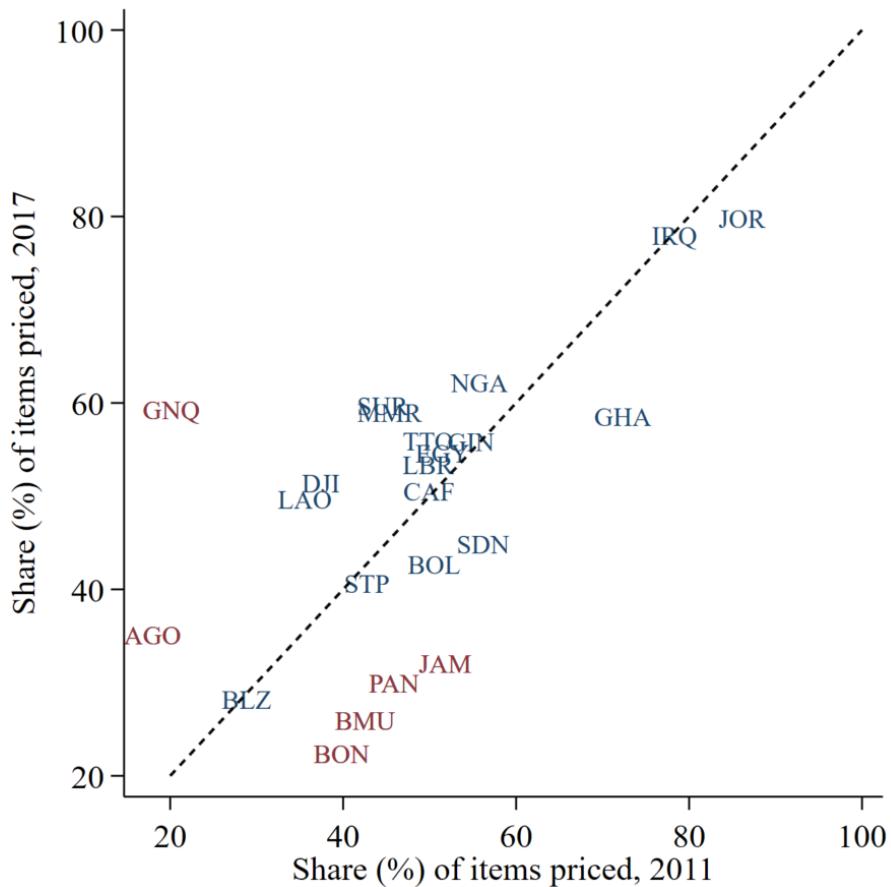
Note: Countries with a “Yes” have no obvious problems with their CPIs. Countries dropped in step 2 above are still added (in italics) for completeness.

b) *Is there a large decline in the share of items priced?*

The share of items priced is defined as the ratio of the sum of global core list and regional items priced to the sum of all global and regional items. Figure D.2 plots the share of items

priced in the 2017 round against the 2011 round. For countries in red, the log difference in the share of items priced in 2011 and 2017 is more than two standard deviations from the mean (evaluated over all countries, not just the ones plotted). Among the countries identified above, only Angola shows up as an outlier by this criterion. In Angola, the share of items priced almost doubles between the two rounds, which is an improvement. None of the countries under consideration has a large decline in the share of items priced.

Figure D.2: Share of items priced: 2011 vs. 2017 rounds



Note: Countries with a marked change in the share of items priced are shown in red, including Angola (AGO), Bermuda (BMU), Bonaire (BON), Equatorial Guinea (GNQ), Jamaica (JAM), and Panama (PAN). Outlier countries under consideration are shown in blue, including those that should be dropped based on step 2. All other countries that participated in the 2011 and 2017 rounds are suppressed. The dotted line is a 45-degree line.

- c) Does the imputed PPP make the poverty rate more consistent with related indicators that do not rely on PPPs, such as the multidimensional poverty index (MPI) and the age dependency ratio (ADR)?

We regress the monetary poverty rate (in 2017 PPPs with a \$2.15 IPL) on the related indicator (MPI or ADR) with regional dummies and interactions. The countries identified in step 1 are excluded from the regression. Using the estimated parameters, we then predict the monetary poverty rate for these countries. Finally, we compute the difference between the predicted poverty rate and (a) the poverty rate with official 2017 PPP and (b) the poverty rate with imputed 2017 PPP. We create a flag for a country if the former difference is greater than the latter difference (see Tables D.3 and D.4). Such a case would imply that the poverty rate with imputed 2017 PPPs is closer to our MPI- or ADR-based priors than the poverty rate with official 2017 PPPs.

Table D.3: Multidimensional poverty index and monetary poverty

<i>Country</i>	<i>Year</i>	<i>Log difference between predicted poverty from MPI and poverty with official 2017 PPP</i>	<i>Log difference between predicted poverty from MPI and poverty with imputed 2017 PPP</i>	<i>Flag</i>
1. Belize	2015	1.79	1.45	Yes
2. Bolivia	2008	0.12	0.06	Yes
3. Djibouti	2006	1.37	0.56	Yes
4. Egypt, Arab Rep.	2014	2.21	2.43	No
5. Guinea	2018	1.36	0.96	Yes
6. Iraq	2018	3.41	2.17	Yes
7. Lao PDR	2017	0.57	0.77	No
8. Myanmar	2015	1.37	1.54	No
9. Nigeria	2018	0.26	0.08	Yes
10. São Tomé and Príncipe	2014	0.21	0.43	No
11. Sudan	2014	1.42	0.37	Yes
12. Suriname	2018	1.84	2.05	No
13. Trinidad and Tobago	2011	1.82	1.46	Yes
<i>14. Angola</i>	<i>2015</i>	<i>0.55</i>	<i>0.86</i>	<i>No</i>
<i>15. Central African Republic</i>	<i>2010</i>	<i>0.07</i>	<i>0.37</i>	<i>No</i>
<i>16. Ghana</i>	<i>2014</i>	<i>0.18</i>	<i>0.22</i>	<i>No</i>
17. Jordan	2017	0.41	0.13	Yes
<i>18. Liberia</i>	<i>2013</i>	<i>0.46</i>	<i>0.95</i>	<i>No</i>

Notes: We use the multidimensional poverty index (MPI) from the Oxford Poverty and Human Development Initiative (OPHI) for the latest available year, which is indicated in the table (Alkire et al. 2020). These estimates are matched with the lined-up, annual estimates of poverty in PovcalNet. For countries with a “Yes”, the poverty rate with imputed PPP is more in tune with what we would expect from the MPI than with the official PPP. Countries that are dropped in step 2 above are still added (in italics) for completeness.

Table D.4: Age dependency ratio and monetary poverty

<i>Country</i>	<i>Year</i>	<i>Log difference between predicted poverty from ADR and poverty with official 2017 PPP</i>	<i>Log difference between predicted poverty from ADR and poverty with imputed 2017 PPP</i>	<i>Flag</i>
1. Belize	1999	0.40	0.09	Yes
2. Bolivia	2019	0.76	0.59	Yes
3. Djibouti	2017	3.22	2.38	Yes
4. Egypt, Arab Rep.	2017	1.07	2.28	No
5. Guinea	2012	1.01	0.69	Yes
6. Iraq	2012	3.90	2.98	Yes
7. Lao PDR	2018	0.90	0.66	Yes
8. Myanmar	2017	0.66	0.38	Yes
9. Nigeria	2018	0.47	0.12	Yes
10. São Tomé and Príncipe	2017	0.43	0.18	Yes
11. Sudan	2014	1.23	0.19	Yes
12. Suriname	1999	1.25	1.53	No
13. Trinidad and Tobago	1992	1.20	0.55	Yes
<i>14. Angola</i>	<i>2018</i>	<i>0.72</i>	<i>0.94</i>	<i>No</i>
<i>15. Central African Republic</i>	<i>2008</i>	<i>0.36</i>	<i>0.06</i>	<i>Yes</i>
<i>16. Ghana</i>	<i>2016</i>	<i>0.47</i>	<i>0.51</i>	<i>No</i>
<i>17. Jordan</i>	<i>2010</i>	<i>3.75</i>	<i>3.75</i>	<i>No</i>
<i>18. Liberia</i>	<i>2016</i>	<i>0.09</i>	<i>0.52</i>	<i>No</i>

Notes: We use estimates of poverty for the latest survey year available in PovcalNet, which is indicated in the table. These estimates are matched with annual estimates of ADR in WDI. For countries with a “Yes”, the poverty rate with imputed PPP is more in tune with what we would expect from ADR than with the official PPP. We use “Age dependency ratio, young (% of working-age population)”, the ratio of younger dependents (i.e., people younger than 15) to the working-age population (i.e., those aged 15-64) from WDI for the analysis reported here. The flags are the same when we use “Age dependency ratio (% of working-age population)”, the ratio of all dependents (i.e., people younger than 15 or older than 64) to the working-age population (i.e., those aged 15-64). Countries that should be dropped based on step 2 above are still added (in italics) for completeness.

d) Does the poverty economist find the official 2017 PPP less appropriate?

Following Ferreira et al. (2016), we consulted with country poverty economists (PE)—World Bank staff specialized in the measurement of poverty in a country—to assess whether the imputed PPP seems more appropriate for the countries identified in step 1. Our partial reliance on poverty economists reflects our view that the prior three steps might have missed relevant country-specific information and nuances. There is, however, a trade-off between using country-specific information and preserving comparability across countries. For that reason, we consider the judgment of poverty economists as one input alongside the other indicators. Poverty economists could not influence the set of countries

we considered outliers nor the PPPs we use for countries not considered outliers. Their responses are provided in Table D.5.

Table D.5: Summary of responses

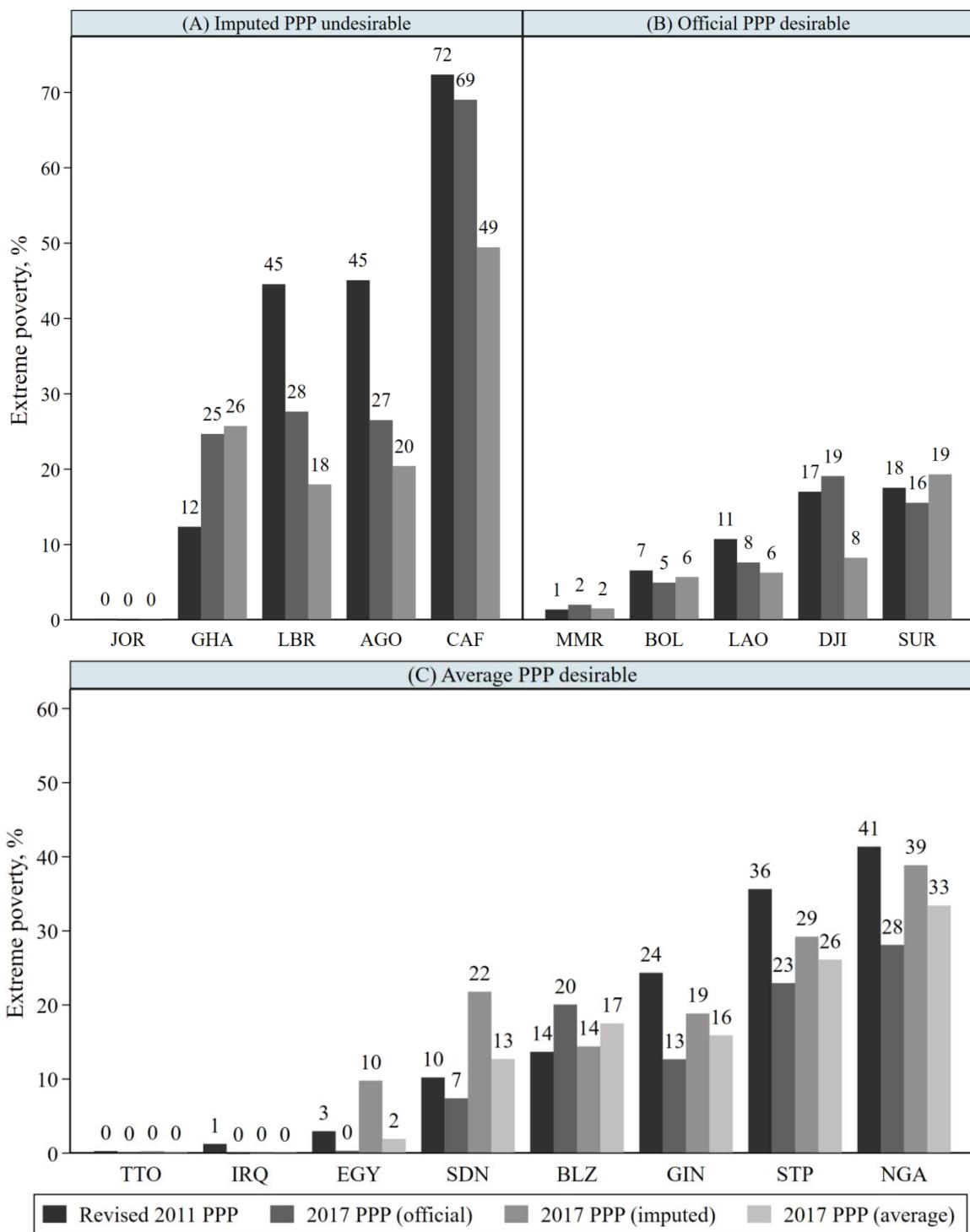
Country	Panel (A)					Panel (B)			Flags (sum of panel A Yes's)
	CPIs	Items priced	MPI	ADR	PE	Delta ratio outlier	PPP residual outlier	Proble m with 2011 PPP	
1. Belize	Yes	No	Yes	Yes	Yes	Yes	No	No	4
2. Bolivia	No	No	Yes	Yes	No	Yes	No	No	2
3. Djibouti	No	No	Yes	Yes	No	No	Yes	No	2
4. Egypt, Arab Rep.	Yes	No	No	No	Yes	Yes	Yes	Yes	2
5. Guinea	No	No	Yes	Yes	Yes	Yes	No	No	3
6. Iraq	Yes	No	Yes	Yes	Yes	Yes	No	Yes	4
7. Lao PDR	No	No	No	Yes	Yes	Yes	No	Yes	2
8. Myanmar	No	No	No	Yes	No	Yes	No	Yes	1
9. Nigeria	No	No	Yes	Yes	Yes	Yes	No	No	3
10. São Tomé and Príncipe	Yes	No	No	Yes	Yes	Yes	No	No	3
11. Sudan	No	No	Yes	Yes	Yes	No	Yes	No	3
12. Suriname	Yes	No	No	No	N/A	Yes	Yes	No	1
13. Trinidad and Tobago	Yes	No	Yes	Yes	Yes	Yes	No	No	4
<i>14. Angola</i>	<i>Yes</i>	<i>No</i>	<i>No</i>	<i>No</i>	<i>No</i>	Yes	No	No	1
<i>15. Central African Republic</i>	<i>No</i>	<i>No</i>	<i>No</i>	Yes	<i>No</i>	Yes	Yes	No	1
<i>16. Ghana</i>	<i>No</i>	<i>No</i>	<i>No</i>	<i>No</i>	<i>No</i>	Yes	No	No	0
<i>17. Jordan</i>	<i>No</i>	<i>No</i>	Yes	<i>No</i>	<i>No</i>	Yes	No	Yes	1
<i>18. Liberia</i>	Yes	No	No	No	No	Yes	No	No	1

Note: In panel (A), we have the factors identified in Step 3 providing suggestive evidence as to whether there might be issues with the official 2017. In panel (B), we include the questions raised in Step 1 with which we selected the 18 problem countries. There is no poverty economist for Suriname.

Step 4: Decide on when to use imputed PPPs

- a) For the 18 countries identified, we consider two options: either the official PPP, or the average between the imputed PPP (using the ICP model) and the official PPP. We prefer the average of official and imputed, instead of the imputed PPP by itself, because we think both PPPs represent signals that are worthy accounting for. The choice between these options is guided by the analysis in step 3 (see Table D.5 above). The more flags a country has, the more likely we are to decide not to use the official PPP.

Figure D.3: Poverty estimates in 2017 with different PPPs



Notes: This chart shows extreme poverty estimates in 2017 using revised 2011 and 2017 PPPs. Panel (A) shows the poverty estimates for countries dropped in step 2 because using an imputed PPP would make the delta ratio even more extreme. Panel (B) shows outlier countries where there is little evidence against using the official PPP to estimate poverty. Panel (C) shows outlier countries where we propose to use the average of official and imputed PPPs. This figure uses the international poverty lines of \$1.90 (2011 PPP) and \$2.15 (2017 PPP) for estimating extreme poverty. See Table D.6 below for the poverty estimates and associated delta ratios.

In the end, we have decided to use the average of official and imputed PPPs for Belize, Egypt, Guinea, Iraq, Nigeria, São Tomé and Príncipe, Sudan, and Trinidad and Tobago, while for the remaining outlier countries we use the official 2017 PPPs. This is equivalent to all countries with at least three flags and Egypt. According to the official PPP, Egypt has the lowest price level index in the world (World Bank 2020) and Egypt thus represents an extreme outlier. Egypt is also the only country that is identified as possibly having a problematic 2017 PPP estimate based on the delta criterion, PPP residual criterion, and countries with exceptional 2011 PPPs (see Table B.5, panel (B)).

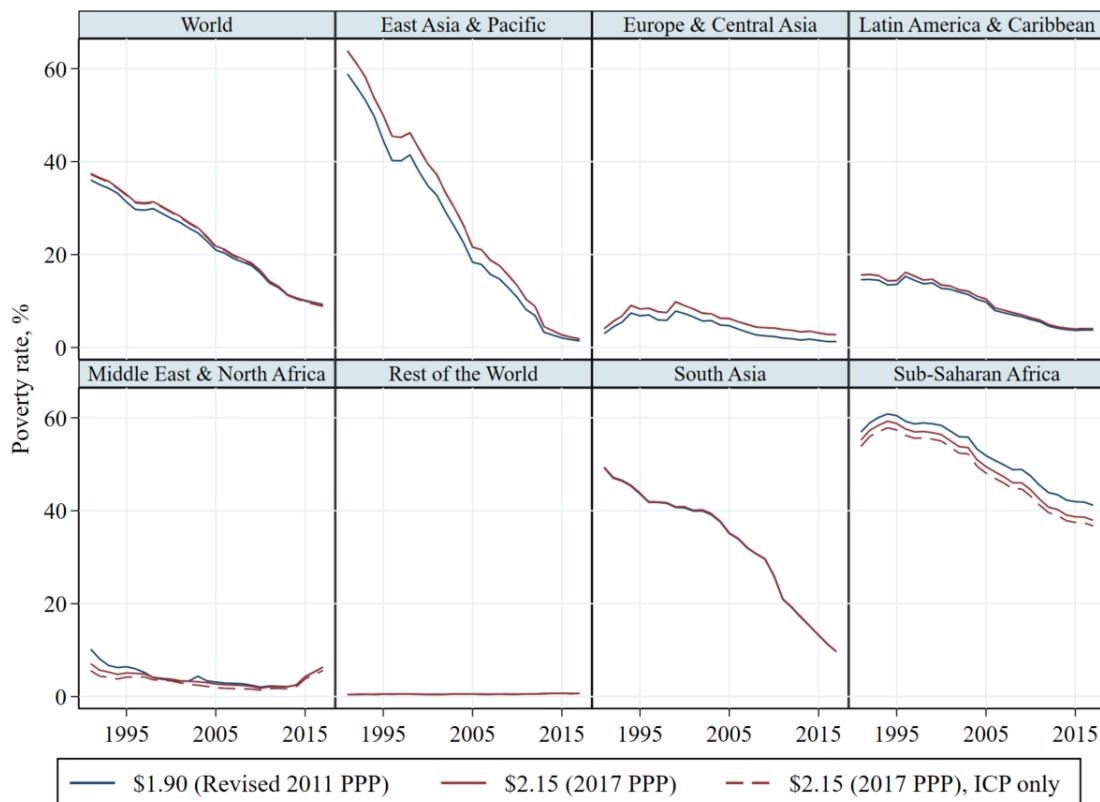
See Figure D.3 and Table D.6 for the poverty estimates for all 18 countries under review for the year 2017 with different PPPs.

Table D.6: Delta ratios and poverty changes

Country	2011 PPP Poverty	2017 PPP (official)			2017 PPP (imputed)			2017 PPP (average)		
		Poverty	Change	Delta	Poverty	Change	Delta	Poverty	Change	Delta
<i>(A) Use official 2017 PPP: Imputed PPP leads to even larger changes</i>										
Jordan	0.10	0.04	-0.06	1.53	0.03	-0.07	1.57			
Ghana	12.34	24.68	12.33	0.75	25.72	13.38	0.73			
Liberia	44.55	27.64	-16.91	1.50	17.97	-26.58	1.81			
Angola	45.07	26.50	-18.57	1.80	20.41	-24.66	2.14			
Central African Republic	72.37	69.03	-3.35	1.24	49.44	-22.93	1.93			
<i>(B) Use official 2017 PPP: Official PPP seems appropriate</i>										
Myanmar	1.36	1.99	0.63	1.05	1.51	0.14	1.10			
Bolivia	6.55	4.93	-1.61	1.43	5.68	-0.87	1.33			
Lao PDR	10.72	7.61	-3.11	1.24	6.26	-4.46	1.31			
Djibouti	17.01	19.08	2.07	1.06	8.24	-8.77	1.63			
Suriname	17.53	15.53	-2.01	1.45	19.30	1.77	0.96			
<i>(C) Use average 2017 PPP</i>										
Trinidad and Tobago	0.28	0.17	-0.10	1.43	0.26	-0.02	1.17	0.21	-0.07	1.29
Iraq	1.24	0.02	-1.22	1.90	0.09	-1.15	1.70	0.05	-1.19	1.80
Egypt, Arab Rep.	2.97	0.31	-2.65	1.77	9.78	6.81	0.88	1.93	-1.04	1.24
Sudan	10.22	7.39	-2.83	1.26	21.78	11.56	0.87	12.72	2.50	1.05
Belize	13.66	20.05	6.39	0.83	14.39	0.73	1.05	17.50	3.83	0.93
Guinea	24.33	12.67	-11.66	1.46	18.84	-5.49	1.27	15.90	-8.43	1.36
São Tomé and Príncipe	35.64	22.94	-12.70	1.46	29.20	-6.44	1.30	26.12	-9.52	1.38
Nigeria	41.36	28.09	-13.27	1.44	38.86	-2.50	1.18	33.41	-7.96	1.31

Notes: Poverty is in percentages (%), and poverty change is in percentage points (pp). Poverty estimates with the average of official and imputed PPPs are not applicable for panels (A) and (B). Countries are sorted in ascending order of poverty rate, as measured by the revised 2011 PPP.

Figure D.4: Sensitivity of extreme poverty trends to alternative PPPs



Notes: This figure shows the trends in extreme poverty at the global and regional levels using the revised 2011 PPPs and 2017 PPPs. The solid red line uses 2017 PPPs officially published by the ICP, except for 14 countries where alternative PPPs are used (see Table B.3), while the dashed red line uses only 2017 PPPs officially published by the ICP.

Table D.7: Sensitivity of changes in extreme poverty in 2017 to alternative PPPs

Region	A. 2017 PPP		B. 2017 PPP (ICP)	
	Change in poverty, pp	Change in millions of poor	Change in poverty, pp	Change in millions of poor
World	-0.20	-15.36	-0.41	-30.96
East Asia & Pacific	0.41	8.45	0.41	8.45
Europe & Central Asia	1.48	7.26	1.48	7.26
Latin America & Caribbean	0.31	1.93	0.31	1.94
Middle East & North Africa	0.02	0.09	-0.72	-2.76
Rest of the World	0.00	0.01	0.00	0.01
South Asia	0.06	1.01	0.06	1.01
Sub-Saharan Africa	-3.25	-34.10	-4.46	-46.86

Notes: This table shows the changes to extreme poverty estimates and millions of poor when moving from the revised 2011 PPPs to the 2017 PPPs. Extreme poverty is estimated using \$1.90 (2011 PPP) and \$2.15 (2017 PPP). Panel (A) uses 2017 PPPs officially published by the ICP, except for 14 countries where alternative PPPs are used (see Table B.3), while panel (B) uses only 2017 PPPs officially published by the ICP.

Appendix E. Methodological details

Section 1: Update \$1.90 with revised 2011 PPPs

Equation 1 expresses how the national poverty lines of the 15 poor countries denominated in the original 2011 PPPs were originally derived by Ferreira et al. (2016) from the national poverty lines estimated in Ravallion et al. (2009).

$$PL2011_i = PL2005_i * \frac{CPI2011_i}{CPI2005_i} * \frac{PPP2005_i}{PPP2011_i} \quad (1)$$

where:

- $PL2011$ is the national poverty line in the original 2011 PPP
- $PL2005$ is national poverty line Ravallion et al. (2009) estimated in 2005 PPP
- $CPI2011$ is consumer price index (CPI) for 2011 as used by Ferreira et al. (2016)
- $CPI2005$ is consumer price index (CPI) for 2005 as used by Ferreira et al. (2016)
- $PPP2011$ is the original 2011 PPP
- $PPP2005$ is the 2005 PPP

The international poverty line (IPL), denominated in the original 2011 PPP, is given as:

$$IPL2011 = \frac{1}{15} * \sum_{i=1}^{15} PL2011_i \quad (2)$$

We use the following equation to update the IPL in (2) with the revised 2011 PPPs only.

$$PL2011_i = PL2005_i * \frac{CPI2011_i}{CPI2005_i} * \frac{PPPr2005_i}{PPPr2011_i} \quad (3)$$

where:

- $PPPr2011$ is the revised 2011 PPP

We use the following equation to update the IPL in (2) with CPI revisions only.

$$PL2011_i = PL2005_i * \frac{CPIr2011_i}{CPIr2005_i} * \frac{PPP2005_i}{PPP2011_i} \quad (4)$$

where:

- $CPIr2011$ is the updated consumer price index (CPI) for 2011
- $CPIr2005$ is the updated consumer price index (CPI) for 2005

We use the following equation to update the IPL in (2) with both PPP and CPI revisions.

$$PL2011_i = PL2005_i * \frac{CPIr2011_i}{CPIr2005_i} * \frac{PPPr2005_i}{PPPr2011_i} \quad (5)$$

Section 2: Update \$1.90 with 2017 PPPs

Using equation (5) above, the following expression updates the 15 national poverty lines selected by Ravallion et al. (2009) to 2017, while updating the CPIs used by Ferreira et al. (2016)

$$PL2017_i = PL2005_i * \frac{CPIr2011_i}{CPIr2005_i} * \frac{CPI2017_i}{CPIr2011_i} * \frac{PPP2005_i}{PPPr2011_i} * \frac{PPPr2011_i}{PPP2017_i} \quad (6)$$

Here $PL2017$ is the national poverty line in 2017 PPP. Similarly, the following equation updates the poverty lines to 2017 while using the CPIs that Ferreira et al. (2016) had originally used,

$$PL2017_i = PL2005_i * \frac{CPI2011_i}{CPI2005_i} * \frac{CPI2017_i}{CPIr2011_i} * \frac{PPP2005_i}{PPPr2011_i} * \frac{PPPr2011_i}{PPP2017_i} \quad (7)$$

Lastly, when we take \$1.90 as given for all 15 countries, and update both PPP and CPI to 2017 we use the following expression:

$$PL2017_i = 1.9 * \frac{PPPr2011_i}{PPP2017_i} * \frac{CPI2017_i}{CPIr2011_i} \quad (8)$$

In all these three cases, the international poverty line (IPL), denominated in 2017 PPP, is given as:

$$IPL2017 = \frac{1}{15} * \sum_{i=1}^{15} PL2017_i \quad (9)$$

Section 3: Derive harmonized national poverty lines

The harmonized national poverty line is derived from the following equation.

$$F(z) = \int_0^z f(y) dy = P \quad (10)$$

where:

- F is a cumulative distribution function of income, denominated in original or revised 2011 PPP per person per day (as in PovcalNet)
- $f(y)$ is a probability distribution function of consumption or income, denominated in original or revised 2011 PPP per person per day (as in PovcalNet)
- P is the national poverty headcount ratio reported in WDI
- z is the harmonized national poverty line that yields a poverty headcount ratio of P

Section 4: Update harmonized poverty lines with revised 2011 PPPs

Let $z2011$ be z in equation (10) if the income distribution is denominated in the original 2011 PPP, and let $zr2011$ be z in equation (10) above if the income distribution is denominated in the revised 2011 PPP. Then

$$zr2011 = z2011 * \frac{PPP2011}{PPPr2011} \quad (11)$$

Section 5: Update harmonized poverty lines with 2017 PPPs

The harmonized poverty line, denominated in 2017 PPP, is derived as:

$$z2017 = zr2011 * \frac{PPPr2011}{PPP2017} * \frac{CPI2017}{CPI2011} \quad (12)$$

Section 6: Derive equivalent poverty lines

Let the global poverty rate in a reference year (e.g., 2010) be given as:

$$F(1.9) = \int_0^{1.9} f(y(PPP2011)) dy = P^* \quad (13)$$

with an international poverty line of \$1.90 per person per day in revised 2011 PPP and a global welfare (probability distribution) function, $f(y(\cdot))$ expressing daily income per person in revised 2011 PPP. P^* is the global poverty rate.

The equivalent poverty line in 2017 PPP, \tilde{z} , is derived by solving the following equation for a given reference year.

$$F(\tilde{z}) = \int_0^{\tilde{z}} f(y(PPP2017)) dy = P^* \quad (14)$$

with a global welfare (probability distribution) function, $f(y(\cdot))$ expressing daily income per person in 2017 PPP.

Section 7: Convert PovcalNet distributions from revised 2011 PPPs to original 2011 PPPs

The income distribution, denominated in original 2011 PPP, is derived as:

$$y2011 = y_r2011 * \frac{PPPr2011}{PPP2011} \quad (15)$$

where:

- y_r2011 is the income distribution, denominated in revised 2011 PPP

Section 8: Convert PovcalNet distributions from revised 2011 PPPs to 2017 PPPs

The income distribution, denominated in 2017 PPP, is derived as:

$$y2017 = y_r2011 * \frac{PPPr2011}{PPP2017} * \frac{CPI2017}{CPIr2011} \quad (16)$$

Section 9: Explain changes in poverty with adopting the 2017 PPPs

The change in poverty is a function of the relative change in income per capita (given a poverty line and the distribution). Re-arranging (16),

$$\frac{y2017}{y_r2011} = \frac{CPI2017}{CPIr2011} / \frac{PPP2017}{PPPr2011} = \delta \quad (17)$$

Thus, the rate of change in income per capita depends on the rates of change in CPI and PPP. If, for example, the PPP increases faster than the CPI between the 2011 and 2017 ICP reference years, income per capita decreases and poverty increases at a given poverty line.

Appendix F. Global poverty lines in local currency units

The following table expresses the four global poverty lines typical of low-income countries (LIC), lower-middle-income countries (LMIC), upper-middle-income countries (UMIC), and high-income countries (HIC) in local currency units in 2020 prices. Global poverty lines determined from both revised 2011 and 2017 PPP rounds are provided in the table.

For example, the 2011 PPP exchange rate for Albania is 54.65 Albanian lek per US dollar, and the inflation in Albania from 2011 to 2020 was 16.97%. This means that the \$1.90 line in 2011 USD equates $1.90\$ * 54.65 \frac{\text{lek}}{\$} * 1.1697 = 121.46$ lek in 2020 prices, which is the first entry in the table below.

Table F.1: Global poverty lines in local currency units in 2020 prices

Economy	PPP	LIC	LMIC	UMIC	HIC
Albania	2011	121.46	204.56	351.59	1387.19
Albania	2017	113.84	193.26	362.70	1289.82
Algeria	2011	88.81	149.58	257.09	1014.33
Algeria	2017	89.44	151.84	284.96	1013.37
Angola	2011	627.76	1057.28	1817.20	7169.70
Angola	2017	394.99	670.56	1258.45	4475.30
Argentina	2011	85.10	143.33	246.35	971.96
Argentina	2017	68.04	115.51	216.78	770.91
Armenia	2011	382.20	643.70	1106.36	4365.09
Armenia	2017	378.64	642.81	1206.37	4290.11
Australia	2011	3.39	5.71	9.81	38.72
Australia	2017	3.43	5.83	10.94	38.90
Austria	2011	1.87	3.16	5.43	21.41
Austria	2017	1.88	3.18	5.98	21.25
Azerbaijan	2011	0.92	1.55	2.66	10.49
Azerbaijan	2017	1.13	1.92	3.60	12.79
Bangladesh	2011	76.02	128.04	220.06	868.25
Bangladesh	2017	74.74	126.89	238.13	846.85
Belarus	2011	1.23	2.08	3.57	14.08
Belarus	2017	1.55	2.63	4.94	17.55
Belgium	2011	1.90	3.20	5.50	21.70
Belgium	2017	1.90	3.22	6.04	21.47
Belize	2011	2.33	3.93	6.75	26.64
Belize	2017	2.83	4.81	9.02	32.07
Benin	2011	475.67	801.12	1376.93	5432.61
Benin	2017	485.82	824.77	1547.85	5504.48
Bhutan	2011	53.19	89.58	153.96	607.45
Bhutan	2017	48.86	82.95	155.68	553.63

Table F.1: Global poverty lines in local currency units in 2020 prices (continued)

<i>Economy</i>	<i>PPP</i>	<i>LIC</i>	<i>LMIC</i>	<i>UMIC</i>	<i>HIC</i>
Bolivia	2011	7.48	12.60	21.66	85.47
Bolivia	2017	5.90	10.02	18.81	66.89
Bosnia and Herzegovina	2011	1.63	2.74	4.71	18.56
Bosnia and Herzegovina	2017	1.70	2.88	5.40	19.22
Botswana	2011	11.39	19.19	32.98	130.11
Botswana	2017	11.64	19.76	37.09	131.90
Brazil	2011	5.08	8.56	14.71	58.05
Brazil	2017	5.55	9.43	17.69	62.92
Bulgaria	2011	1.68	2.84	4.87	19.23
Bulgaria	2017	1.70	2.89	5.43	19.31
Burkina Faso	2011	455.35	766.91	1318.13	5200.63
Burkina Faso	2017	431.67	732.84	1375.32	4890.93
Burundi	2011	1681.06	2831.25	4866.22	19199.44
Burundi	2017	1457.43	2474.25	4643.45	16513.07
Cabo Verde	2011	97.16	163.64	281.26	1109.69
Cabo Verde	2017	105.43	178.99	335.92	1194.60
Cameroon	2011	544.41	916.90	1575.91	6217.70
Cameroon	2017	540.32	917.28	1721.48	6121.92
Canada	2011	2.78	4.69	8.06	31.80
Canada	2017	2.90	4.93	9.26	32.91
Central African Republic	2011	774.54	1304.49	2242.10	8846.10
Central African Republic	2017	704.13	1195.38	2243.38	7977.92
Chad	2011	572.98	965.02	1658.63	6544.07
Chad	2017	551.18	935.72	1756.08	6244.98
Chile	2011	966.73	1628.17	2798.42	11041.05
Chile	2017	1114.85	1892.64	3551.95	12631.46
China	2011	8.53	14.36	24.68	97.38
China	2017	9.59	16.28	30.56	108.66
Colombia	2011	3215.11	5414.92	9306.89	36719.89
Colombia	2017	3343.95	5676.94	10653.98	37887.74
Comoros	2011	493.11	830.50	1427.42	5631.81
Comoros	2017	488.03	828.52	1554.89	5529.52
Congo, Dem. Rep.	2011	2440.70	4110.65	7065.17	27875.32
Congo, Dem. Rep.	2017	2043.47	3469.15	6510.59	23152.99
Congo, Rep.	2011	720.26	1213.07	2084.96	8226.10
Congo, Rep.	2017	647.20	1098.73	2062.00	7332.88
Costa Rica	2011	821.50	1383.58	2378.03	9382.40
Costa Rica	2017	842.26	1429.89	2683.49	9543.05
Côte d'Ivoire	2011	504.03	848.89	1459.02	5756.51
Côte d'Ivoire	2017	540.13	916.97	1720.89	6119.85
Croatia	2011	8.63	14.54	24.99	98.59
Croatia	2017	8.40	14.26	26.77	95.19
Cyprus	2011	1.42	2.39	4.12	16.24
Cyprus	2017	1.45	2.46	4.62	16.45

Table F.1: Global poverty lines in local currency units in 2020 prices (continued)

<i>Economy</i>	<i>PPP</i>	<i>LIC</i>	<i>LMIC</i>	<i>UMIC</i>	<i>HIC</i>
Czech Republic	2011	32.34	54.47	93.63	369.41
Czech Republic	2017	31.81	54.00	101.35	360.41
Denmark	2011	17.08	28.77	49.45	195.11
Denmark	2017	17.31	29.38	55.14	196.08
Djibouti	2011	222.70	375.08	644.66	2543.48
Djibouti	2017	237.14	402.59	755.55	2686.89
Dominican Republic	2011	51.15	86.15	148.07	584.22
Dominican Republic	2017	57.38	97.42	182.82	650.16
Ecuador	2011	1.23	2.08	3.57	14.08
Ecuador	2017	1.23	2.09	3.92	13.95
Egypt, Arab Rep.	2011	15.03	25.31	43.50	171.62
Egypt, Arab Rep.	2017	13.66	23.19	43.52	154.77
El Salvador	2011	1.08	1.81	3.11	12.29
El Salvador	2017	1.12	1.89	3.56	12.64
Estonia	2011	1.28	2.15	3.70	14.58
Estonia	2017	1.35	2.29	4.30	15.30
Eswatini	2011	12.76	21.49	36.94	145.73
Eswatini	2017	14.69	24.94	46.81	166.47
Ethiopia	2011	30.81	51.88	89.17	351.84
Ethiopia	2017	28.99	49.22	92.36	328.47
Fiji	2011	2.32	3.91	6.71	26.49
Fiji	2017	2.19	3.71	6.97	24.78
Finland	2011	2.01	3.39	5.83	23.00
Finland	2017	2.06	3.49	6.55	23.28
France	2011	1.81	3.05	5.24	20.67
France	2017	1.86	3.15	5.92	21.06
Gabon	2011	836.53	1408.90	2421.54	9554.08
Gabon	2017	745.16	1265.05	2374.13	8442.89
Gambia, The	2011	36.20	60.97	104.79	413.43
Gambia, The	2017	39.28	66.68	125.14	445.02
Georgia	2011	2.06	3.46	5.95	23.47
Georgia	2017	2.25	3.82	7.17	25.50
Germany	2011	1.75	2.94	5.05	19.94
Germany	2017	1.76	2.98	5.59	19.89
Ghana	2011	2.95	4.97	8.55	33.72
Ghana	2017	4.45	7.56	14.19	50.47
Greece	2011	1.42	2.39	4.12	16.24
Greece	2017	1.38	2.35	4.41	15.68

Table F.1: Global poverty lines in local currency units in 2020 prices (continued)

<i>Economy</i>	<i>PPP</i>	<i>LIC</i>	<i>LMIC</i>	<i>UMIC</i>	<i>HIC</i>
Guatemala	2011	10.27	17.30	29.73	117.31
Guatemala	2017	10.51	17.85	33.50	119.12
Guinea	2011	11838.66	19938.80	34269.82	135210.00
Guinea	2017	9854.55	16729.82	31397.07	111654.39
Guinea-Bissau	2011	534.37	899.99	1546.86	6103.08
Guinea-Bissau	2017	509.72	865.34	1624.00	5775.28
Guyana	2011	255.94	431.06	740.88	2923.11
Guyana	2017	258.90	439.52	824.86	2933.38
Haiti	2011	97.28	163.85	281.61	1111.08
Haiti	2017	109.64	186.13	349.32	1242.25
Honduras	2011	28.41	47.86	82.25	324.53
Honduras	2017	26.26	44.58	83.66	297.51
Hungary	2011	315.21	530.88	912.45	3600.02
Hungary	2017	347.97	590.75	1108.66	3942.64
Iceland	2011	341.30	574.82	987.98	3898.03
Iceland	2017	364.06	618.05	1159.91	4124.87
India	2011	49.17	82.81	142.33	561.56
India	2017	48.12	81.69	153.31	545.20
Indonesia	2011	10815.32	18215.28	31307.52	123522.39
Indonesia	2017	11858.56	20131.97	37781.92	134360.22
Iran, Islamic Rep.	2011	60201.79	101392.49	174268.34	687567.82
Iran, Islamic Rep.	2017	67542.39	114664.99	215193.20	765271.01
Iraq	2011	2019.23	3400.80	5845.13	23061.70
Iraq	2017	1270.34	2156.62	4047.36	14393.24
Ireland	2011	1.88	3.17	5.46	21.53
Ireland	2017	2.11	3.58	6.72	23.91
Israel	2011	8.40	14.15	24.31	95.93
Israel	2017	9.14	15.52	29.12	103.57
Italy	2011	1.66	2.80	4.82	19.00
Italy	2017	1.68	2.85	5.34	18.99
Jamaica	2011	189.92	319.86	549.76	2169.06
Jamaica	2017	171.65	291.41	546.90	1944.88
Japan	2011	231.41	389.75	669.88	2642.98
Japan	2017	246.49	418.46	785.34	2792.81
Jordan	2011	1.01	1.71	2.93	11.58
Jordan	2017	0.75	1.27	2.38	8.47
Kazakhstan	2011	279.37	470.52	808.71	3190.71
Kazakhstan	2017	320.79	544.59	1022.04	3634.57
Kenya	2011	118.85	200.17	344.05	1357.43
Kenya	2017	103.95	176.47	331.19	1177.77

Table F.1: Global poverty lines in local currency units in 2020 prices (continued)

<i>Economy</i>	<i>PPP</i>	<i>LIC</i>	<i>LMIC</i>	<i>UMIC</i>	<i>HIC</i>
Kiribati	2011	2.04	3.44	5.91	23.30
Kiribati	2017	2.12	3.60	6.75	24.01
Korea, Rep.	2011	1910.14	3217.08	5529.35	21815.79
Korea, Rep.	2017	2145.06	3641.61	6834.25	24304.00
Kosovo	2011	0.77	1.29	2.22	8.77
Kosovo	2017	0.80	1.35	2.54	9.03
Kyrgyz Republic	2011	43.94	74.01	127.20	501.85
Kyrgyz Republic	2017	45.48	77.21	144.90	515.29
Lao PDR	2011	8193.58	13799.72	23718.27	93579.36
Lao PDR	2017	7466.18	12675.14	23787.58	84593.50
Latvia	2011	1.21	2.04	3.50	13.80
Latvia	2017	1.26	2.14	4.01	14.25
Lebanon	2011	3613.81	6086.42	10461.04	41273.54
Lebanon	2017	3308.51	5616.78	10541.07	37486.20
Lesotho	2011	11.62	19.57	33.63	132.70
Lesotho	2017	13.04	22.14	41.55	147.76
Liberia	2011	250.73	422.27	725.78	2863.55
Liberia	2017	189.34	321.44	603.25	2145.28
Lithuania	2011	1.12	1.88	3.23	12.75
Lithuania	2017	1.12	1.90	3.56	12.65
Luxembourg	2011	2.09	3.52	6.05	23.85
Luxembourg	2017	2.16	3.67	6.89	24.50
Madagascar	2011	2333.90	3930.79	6756.04	26655.65
Madagascar	2017	2474.20	4200.39	7882.93	28033.31
Malawi	2011	743.51	1252.22	2152.25	8491.62
Malawi	2017	694.72	1179.40	2213.39	7871.28
Malaysia	2011	3.51	5.91	10.15	40.06
Malaysia	2017	3.73	6.33	11.88	42.24
Maldives	2011	20.26	34.12	58.65	231.40
Maldives	2017	20.79	35.29	66.22	235.51
Mali	2011	452.14	761.50	1308.83	5163.95
Mali	2017	437.23	742.27	1393.02	4953.88
Malta	2011	1.33	2.25	3.86	15.23
Malta	2017	1.40	2.37	4.45	15.83
Mauritania	2011	280.46	472.35	811.85	3203.12
Mauritania	2017	288.80	490.29	920.14	3272.20
Mauritius	2011	44.11	74.28	127.68	503.74
Mauritius	2017	41.16	69.88	131.15	466.39

Table F.1: Global poverty lines in local currency units in 2020 prices (continued)

<i>Economy</i>	<i>PPP</i>	<i>LIC</i>	<i>LMIC</i>	<i>UMIC</i>	<i>HIC</i>
Mexico	2011	23.90	40.24	69.17	272.91
Mexico	2017	23.83	40.46	75.93	270.03
Micronesia	2011	2.10	3.54	6.09	24.03
Micronesia	2017	2.18	3.70	6.95	24.72
Moldova	2011	16.46	27.72	47.65	188.01
Moldova	2017	15.38	26.11	49.01	174.29
Mongolia	2011	2065.69	3479.06	5979.63	23592.35
Mongolia	2017	2232.10	3789.38	7111.57	25290.20
Montenegro	2011	0.94	1.59	2.73	10.77
Montenegro	2017	0.92	1.57	2.95	10.48
Morocco	2011	8.69	14.63	25.15	99.23
Morocco	2017	9.49	16.11	30.23	107.49
Mozambique	2011	50.86	85.67	147.24	580.92
Mozambique	2017	52.08	88.41	165.92	590.03
Myanmar	2011	952.66	1604.49	2757.71	10880.43
Myanmar	2017	1030.63	1749.67	3283.62	11677.24
Namibia	2011	15.32	25.80	44.35	174.98
Namibia	2017	16.40	27.84	52.24	185.79
Nauru	2011	3.01	5.07	8.72	34.40
Nauru	2017	3.13	5.31	9.96	35.44
Nepal	2011	88.78	149.53	257.01	1014.01
Nepal	2017	75.71	128.53	241.21	857.81
Netherlands	2011	1.91	3.21	5.52	21.77
Netherlands	2017	1.93	3.28	6.15	21.88
Nicaragua	2011	27.44	46.21	79.43	313.38
Nicaragua	2017	28.42	48.25	90.55	322.00
Niger	2011	478.41	805.74	1384.86	5463.90
Niger	2017	544.56	924.49	1735.00	6170.01
Nigeria	2011	434.05	731.04	1256.47	4957.36
Nigeria	2017	376.01	638.34	1197.97	4260.23
North Macedonia	2011	49.54	83.44	143.41	565.80
North Macedonia	2017	49.31	83.71	157.10	558.68
Norway	2011	22.43	37.78	64.93	256.19
Norway	2017	24.56	41.70	78.26	278.29
Pakistan	2011	86.60	145.86	250.69	989.08
Pakistan	2017	91.16	154.75	290.43	1032.82
Panama	2011	1.19	2.01	3.45	13.63
Panama	2017	1.08	1.83	3.44	12.23

Table F.1: Global poverty lines in local currency units in 2020 prices (continued)

<i>Economy</i>	<i>PPP</i>	<i>LIC</i>	<i>LMIC</i>	<i>UMIC</i>	<i>HIC</i>
Papua New Guinea	2011	6.36	10.71	18.41	72.63
Papua New Guinea	2017	6.60	11.21	21.04	74.82
Paraguay	2011	5759.64	9700.45	16672.65	65781.19
Paraguay	2017	5937.75	10080.37	18917.96	67276.12
Peru	2011	3.84	6.46	11.10	43.81
Peru	2017	4.29	7.28	13.67	48.61
Philippines	2011	45.31	76.31	131.16	517.48
Philippines	2017	46.14	78.33	147.00	522.77
Poland	2011	4.07	6.86	11.79	46.53
Poland	2017	4.26	7.23	13.58	48.28
Portugal	2011	1.40	2.36	4.05	15.99
Portugal	2017	1.43	2.43	4.56	16.23
Romania	2011	4.13	6.96	11.96	47.18
Romania	2017	4.33	7.35	13.78	49.02
Russian Federation	2011	61.08	102.87	176.80	697.57
Russian Federation	2017	60.24	102.27	191.94	682.58
Rwanda	2011	768.69	1294.63	2225.15	8779.23
Rwanda	2017	714.66	1213.26	2276.94	8097.26
Samoa	2011	3.97	6.69	11.49	45.34
Samoa	2017	4.12	7.00	13.13	46.71
São Tomé and Príncipe	2011	38.16	64.27	110.47	435.84
São Tomé and Príncipe	2017	31.35	53.22	99.89	355.22
Senegal	2011	513.16	864.27	1485.47	5860.84
Senegal	2017	537.73	912.88	1713.22	6092.57
Serbia	2011	109.10	183.74	315.81	1246.01
Serbia	2017	108.84	184.77	346.77	1233.17
Seychelles	2011	20.06	33.79	58.07	229.11
Seychelles	2017	20.74	35.20	66.07	234.95
Sierra Leone	2011	8628.24	14531.77	24976.48	98543.55
Sierra Leone	2017	6915.81	11740.79	22034.09	78357.73
Slovak Republic	2011	1.19	2.01	3.46	13.64
Slovak Republic	2017	1.34	2.27	4.26	15.14
Slovenia	2011	1.40	2.36	4.06	16.03
Slovenia	2017	1.42	2.41	4.52	16.08
Solomon Islands	2011	17.29	29.12	50.06	197.50
Solomon Islands	2017	17.96	30.49	57.21	203.46
South Africa	2011	14.93	25.15	43.23	170.55
South Africa	2017	15.82	26.85	50.39	179.20

Table F.1: Global poverty lines in local currency units in 2020 prices (continued)

<i>Economy</i>	<i>PPP</i>	<i>LIC</i>	<i>LMIC</i>	<i>UMIC</i>	<i>HIC</i>
South Sudan	2011	429.05	722.62	1241.99	4900.23
South Sudan	2017	334.51	567.89	1065.78	3790.12
Spain	2011	1.61	2.70	4.65	18.33
Spain	2017	1.54	2.62	4.91	17.47
Sri Lanka	2011	126.36	212.82	365.78	1443.19
Sri Lanka	2017	133.94	227.39	426.74	1517.57
St. Lucia	2011	4.19	7.06	12.13	47.87
St. Lucia	2017	4.50	7.64	14.34	51.00
Sudan	2011	79.29	133.55	229.54	905.63
Sudan	2017	85.78	145.62	273.29	971.87
Suriname	2011	12.19	20.53	35.29	139.25
Suriname	2017	9.50	16.13	30.27	107.64
Sweden	2011	18.49	31.14	53.52	211.15
Sweden	2017	20.68	35.11	65.89	234.31
Switzerland	2011	2.88	4.86	8.35	32.94
Switzerland	2017	2.94	4.99	9.36	33.28
Syrian Arab Republic	2011	759.41	1279	2198.28	8673.21
Syrian Arab Republic	2017	788.58	1338.75	2512.45	8934.79
Taiwan, China	2011	32.98	55.55	95.48	376.71
Taiwan, China	2017	36.35	61.72	115.82	411.89
Tajikistan	2011	5.85	9.85	16.93	66.82
Tajikistan	2017	6.78	11.51	21.60	76.83
Tanzania	2011	1921.85	3236.81	5563.26	21949.60
Tanzania	2017	1794.46	3046.41	5717.24	20331.69
Thailand	2011	26.23	44.17	75.92	299.53
Thailand	2017	28.83	48.94	91.85	326.65
Timor-Leste	2011	1.23	2.08	3.57	14.10
Timor-Leste	2017	0.94	1.60	3.01	10.70
Togo	2011	500.81	843.47	1449.71	5719.77
Togo	2017	548.62	931.38	1747.93	6215.99
Tonga	2011	3.93	6.62	11.37	44.88
Tonga	2017	4.08	6.93	13.00	46.23
Trinidad and Tobago	2011	11.77	19.82	34.07	134.42
Trinidad and Tobago	2017	10.29	17.46	32.78	116.56
Tunisia	2011	2.10	3.54	6.09	24.03
Tunisia	2017	1.93	3.28	6.15	21.87
Turkey	2011	5.33	8.98	15.43	60.87
Turkey	2017	5.33	9.04	16.97	60.35
Turkmenistan	2011	5.11	8.60	14.78	58.31
Turkmenistan	2017	4.58	7.78	14.59	51.89

Table F.1: Global poverty lines in local currency units in 2020 prices (continued)

<i>Economy</i>	<i>PPP</i>	<i>LIC</i>	<i>LMIC</i>	<i>UMIC</i>	<i>HIC</i>
Tuvalu	2011	2.78	4.68	8.05	31.77
Tuvalu	2017	2.89	4.90	9.20	32.72
Uganda	2011	2812.38	4736.65	8141.11	32120.40
Uganda	2017	2876.69	4883.68	9165.27	32593.57
Ukraine	2011	16.20	27.29	46.90	185.06
Ukraine	2017	18.52	31.45	59.01	209.87
United Arab Emirates	2011	5.06	8.53	14.65	57.82
United Arab Emirates	2017	6.03	10.24	19.22	68.35
United Kingdom	2011	1.72	2.89	4.97	19.60
United Kingdom	2017	1.76	2.99	5.61	19.96
United States	2011	2.19	3.68	6.33	24.97
United States	2017	2.27	3.85	7.23	25.72
Uruguay	2011	65.02	109.51	188.21	742.59
Uruguay	2017	68.09	115.59	216.93	771.43
Uzbekistan	2011	3337.31	5620.74	9660.64	38115.62
Uzbekistan	2017	4968.84	8435.47	15830.96	56298.13
Vanuatu	2011	270.05	454.81	781.71	3084.20
Vanuatu	2017	280.42	476.06	893.43	3177.22
Venezuela, RB	2011	2103457343.59	3542664999.72	6088955468.27	24023697029.37
Venezuela, RB	2017	2184265599.02	3708171830.90	6959171792.23	24748237205.67
Vietnam	2011	20343.08	34262.03	58887.87	232339.41
Vietnam	2017	18441.94	31308.41	58756.87	208951.45
West Bank and Gaza	2011	5.03	8.47	14.56	57.43
West Bank and Gaza	2017	4.47	7.59	14.24	50.66
Yemen, Rep.	2011	931.33	1568.56	2695.96	10636.77
Yemen, Rep.	2017	967.11	1641.84	3081.26	10957.58
Zambia	2011	11.46	19.29	33.16	130.83
Zambia	2017	12.33	20.93	39.29	139.72
Zimbabwe	2011	12.95	21.82	37.50	147.94
Zimbabwe	2017	13.02	22.11	41.49	147.54

Notes: Poverty increases (decreases) when the global poverty line in local currency units increases (decreases) with the 2017 PPP, relative to the revised 2011 PPP. These results are mechanically similar to the results from the analysis of the delta ratio. For example, a relatively high delta ratio of 1.80 for Angola indicates a reduction in extreme poverty (see Table D.6), as the local currency units of the IPL in 2017 PPP reduce relative to the revised 2011 PPP.

Appendix references

- Alkire, S., U. Kanagaratnam, and N. Suppa. 2020. “The Global Multidimensional Poverty Index (MPI).” 49. OPHI MPI Methodological Notes. Oxford: Oxford Poverty and Human Development Initiative, University of Oxford.
- Allen, Robert C. 2017. “Absolute Poverty: When Necessity Displaces Desire.” *American Economic Review* 107 (12): 3690–3721. <https://doi.org/10.1257/aer.20161080>.
- Atamanov, Aziz, Dean Jolliffe, Christoph Lakner, and Espen Beer Prydz. 2018. “Purchasing Power Parities Used in Global Poverty Measurement.” *Global Poverty Monitoring Technical Note 5* (September).
- Atamanov, Aziz, Christoph Lakner, Daniel Gerszon Mahler, Samuel Kofi Tetteh-Baah, and Judy Yang. 2020. “The Effect of New PPP Estimates on Global Poverty: A First Look,” no. May.
- Bai, Yan, and William A. Masters. 2020. “Retail Food Prices at Purchasing Power Parity Exchange Rates: A First Look at Aggregate ICP 2017 Data.” World Bank Data Blog. 2020. <https://blogs.worldbank.org/opendata/retail-food-prices-purchasing-power-parity-exchange-rates-first-look-aggregate-icp-2017>.
- Chen, Shaohua, and Martin Ravallion. 2010. “The Developing World Is Poorer than We Thought, but No Less Successful in the Fight against Poverty.” *The Quarterly Journal of Economics*, no. November: 1577–1625.
- Deaton, Angus. 2001. “Counting the World’s Poor: Problems and Possible Solutions.” *World Bank Research Observer* 16 (2): 125–47. <https://doi.org/10.1093/wbro/16.2.125>.
- Deaton, Angus, and Bettina Aten. 2017. “Trying to Understand the PPPs in ICP 2011: Why Are the Results so Different.” *American Economic Journal: Macroeconomics* 9 (1): 243–64. <https://doi.org/10.1257/mac.20150153>.
- Deaton, Angus, and Paul Schreyer. 2022. “GDP, Wellbeing, and Health: Thoughts on the 2017 Round of the International Comparison Program.” *Review of Income and Wealth* 68 (1): 1–15. <https://doi.org/10.1111/roiw.12520>.
- Ferreira, Francisco H.G., Shaohua Chen, Andrew Dabalen, Yuri Dikhanov, Nada Hamadeh, Dean Jolliffe, Ambar Narayan, et al. 2016. “A Global Count of the Extreme Poor in 2012: Data Issues, Methodology and Initial Results.” *Journal of Economic Inequality* 14 (2): 141–72. <https://doi.org/10.1007/s10888-016-9326-6>.
- Inklaar, Robert, and D S Prasada Rao. 2017. “Cross-Country Income Levels over Time: Did the Developing World Suddenly Become Much Richer?” *American Economic Journal: Macroeconomics* 9 (1): 265–90.
- Inklaar, Robert, and Prasada Rao. 2020. “ICP PPP Time Series Implementation.” Washington, DC. <https://thedocs.worldbank.org/en/doc/f32e966db6ec404699b3381b32e2e589-0050022021/original/2-01-RA-Item-01-ICP-PPP-Time-Series-Implementation-Rao-and-Inklaar.pdf>.
- Jolliffe, Dean, and Espen Beer Prydz. 2016. “Estimating International Poverty Lines from Comparable National Thresholds.” *Journal of Economic Inequality* 14 (2): 185–98. <https://doi.org/10.1007/s10888-016-9327-5>.
- OECD. 2013. “The 2008 SNA – Changes from the 1993 SNA.” In *National Accounts at a Glance 2013*, 90–92. Paris: OECD Publishing. https://doi.org/10.1787/na_glance-2010-28-en.
- Ravallion, Martin. 2020. “Book Review of Measuring Poverty around the World.” *The Journal of Economic Inequality* 18 (1): 131–36. <https://doi.org/10.1007/s10888-019-09430-w>.
- Ravallion, Martin, Shaohua Chen, and Prem Sangraula. 2009. “Dollar a Day Revisited.” *World*

- Bank Economic Review* 23 (2): 163–84. <https://doi.org/10.1093/wber/lhp007>.
- Tetteh-Baah, Samuel Kofi, Aziz Atamanov, Christoph Lakner, Daniel Gerszon Mahler, Marko Rissanen, William Vigil-Oliver, Mizuki Yamanaka, and Judy Yang. 2020. “Why Have the 2011 PPPs Been Revised and What Does It Mean for Estimates of Poverty?” *World Bank Data Blog*, 2020. <https://blogs.worldbank.org/opendata/why-have-2011-ppps-been-revised-and-what-does-it-mean-estimates-poverty>.
- World Bank. 2008. *Global Purchasing Power Parities and Real Expenditures: 2005 International Comparison Program*. Washington, DC: World Bank.
- . 2013. *Measuring the Real Size of the World Economy: The Framework, Methodology, and Results of the International Comparison Program (ICP)*. Washington, DC: World Bank. <https://doi.org/10.5860/choice.51-0989>.
- . 2015. *Purchasing Power Parities and the Real Size of World Economies: A Comprehensive Report of the 2011 International Comparison Program*. Washington, DC: World Bank.
- . 2016. “Monitoring Global Poverty: A Cover Note to the Report of the Commission on Global Poverty, Chaired by Prof. Sir Anthony B. Atkinson.” Washington, DC: World Bank. <http://pubdocs.worldbank.org/en/733161476724983858/MonitoringGlobalPovertyCoverNote.pdf>.
- . 2020. *Purchasing Power Parities and the Size of World Economies: Results from the 2017 International Comparison Program*. Washington, DC: World Bank. <https://doi.org/10.1596/978-1-4648-1530-0>.