



Parametric estimation of line spectra

Linear time series (part 2) TSIA202b







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 - pairs or triplets of strings in a piano, plus coupling of the vertical and horizontal vibration modes







Part I

Parametric signal model





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- The observed signal x_t is modeled as the signal s_t plus a complex Gaussian white noise b_t of variance σ^2 (sequence of complex IID

r.v. of PDF
$$p(b) = \frac{1}{\pi \sigma^2} e^{-\frac{|b|^2}{\sigma^2}}$$





Peak detection in the Fourier transform





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 - widening of the peak in case of exponential damping





直送影響 Resolution problems

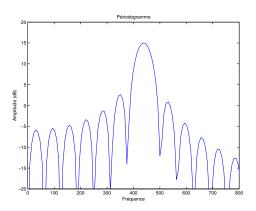
Test signal:

- Sampling frequency : 8000 Hz
- First sinusoid : 440 Hz (A)
- Second sinusoid : 415,3 Hz (G#)
- No damping, all amplitudes equal to 1
- Length of the rectangular window : N = 128 (16 ms)
- Length of the transform : 1024 samples





Resolution problems

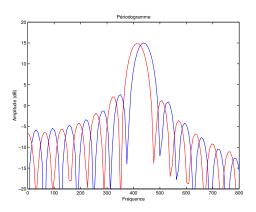








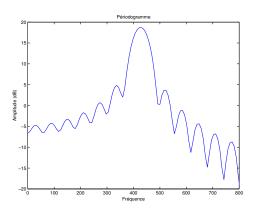
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Part II

Maximum Likelihood Method







直接 Definitions

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- $V^N = [v(z_0), ..., v(z_{K-1})]$ is a Vandermonde matrix :

$$\boldsymbol{V}^{N} = \begin{bmatrix} 1 & 1 & \dots & 1 \\ z_{0} & z_{1} & \dots & z_{K-1} \\ \vdots & \vdots & \vdots & \vdots \\ z_{0}^{N-1} & z_{1}^{N-1} & \dots & z_{K-1}^{N-1} \end{bmatrix}$$



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- $\mathbf{x}(t) = [x_{t-l+1}, \dots, x_{t+n-1}]^{\top} = \mathbf{s}(t) + \mathbf{b}(t)$







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- Probability density function (PDF) of the random vector $\mathbf{x}(t) = \mathbf{s}(t) + \mathbf{b}(t)$, where $\mathbf{s}(t) = \mathbf{V}^N \alpha(t)$ is deterministic and $\mathbf{b}(t) \sim \mathcal{N}(\mathbf{0}, \mathbf{R}_{bb})$:

$$p(\mathbf{x}(t)) = \frac{1}{\pi^N \det(\mathbf{R}_{bb})} e^{-(\mathbf{x}(t) - \mathbf{s}(t))^H \mathbf{R}_{bb}^{-1}(\mathbf{x}(t) - \mathbf{s}(t))}$$



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Log-likelihood of the observations :

$$L(\sigma^2, z_0 \dots z_{K-1}, \alpha(t)) = -N \ln(\pi \sigma^2) - \frac{1}{\sigma^2} g(z_0 \dots z_{K-1}, \alpha(t))$$
where $g(z_0 \dots z_{K-1}, \alpha(t)) = \left(\mathbf{x}(t) - \mathbf{V}^N \alpha(t)\right)^H \left(\mathbf{x}(t) - \mathbf{V}^N \alpha(t)\right)$





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Maximization w.r.t. $\sigma^2 : \sigma^2 = \frac{1}{N} \| \boldsymbol{x}(t) - \boldsymbol{V}^N \boldsymbol{\alpha}(t) \|^2$





Maximum likelihood method

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- We finally have to minimize function

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$$+ \left(\alpha(t) - \left(\mathbf{V}^{NH} \mathbf{V}^N \right)^{-1} \mathbf{V}^{NH} \mathbf{x}(t) \right)^H \left(\mathbf{V}^{NH} \mathbf{V}^N \right) \left(\alpha(t) - \left(\mathbf{V}^{NH} \mathbf{V}^N \right)^{-1} \mathbf{V}^{NH} \mathbf{x}(t) \right)$$

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Minimization w.r.t. $\alpha(t)$: $\alpha(t) = (\boldsymbol{V}^{NH} \boldsymbol{V}^{N})^{-1} \boldsymbol{V}^{NH} \boldsymbol{x}(t)$

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- This optimization problem has to be solved numerically







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- Difficulties of the first step :





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 - computational complexity





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where
$$X(f_0) = \mathbf{v}(e^{i2\pi f_0})^H \mathbf{x}(t) = \sum_{\tau=0}^{N-1} x_{t-l+1+\tau} e^{-i2\pi f_0 \tau}$$



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- The maximum likelihood principle results in detecting the frequency at which the periodogram reaches its maximum.
- The complex amplitude is proportional to the value of the TFD.
- The noise variance is the residual power.





TENTAL SECTION Fourier resolution

■ If
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Table 1 Fourier resolution

- If $K \ge 1$, we assume that $N >> \frac{1}{\min\limits_{k_1 \ne k_2} |f_{k_2} f_{k_1}|}$
- We have $\{\boldsymbol{V}^{NH}\boldsymbol{V}^N\}_{(k_1, k_2)} = \sum_{\tau=0}^{N-1} (z_{k_1}^* z_{k_2})^{\tau}$.



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- Then $\frac{1}{N} \boldsymbol{V}^{NH} \boldsymbol{V}^{N} = \boldsymbol{I}_{K} + O\left(\frac{1}{N}\right)$, thus $\left(\boldsymbol{V}^{NH} \boldsymbol{V}^{N}\right)^{-1} = \frac{1}{N} \boldsymbol{I}_{K} + O\left(\frac{1}{N^{2}}\right)$

Roland Badeau

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多数型 Fourier resolution

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- $\sigma^2 = \frac{1}{N} \left(\| \mathbf{x}(t) \|^2 \sum_{k=0}^{K-1} \widehat{S}_{P,xx}(f_k) \right) + O\left(\frac{1}{N^2}\right).$
- Find the K greatest values of the periodogram
- Limit of Fourier analysis : $\min_{k_1 \neq k_2} |f_{k_2} f_{k_1}| >> \frac{1}{N}$











Part III

High resolution methods based on linear prediction

Jean-Baptiste Joseph Fourier (1768-1830)







Gaspard-Marie Riche of Prony (1755-1839)







Linear prediction methods

■ Principle : any signal such that $s_t - z_0 s_{t-1} = 0$ is of the form $s_t = \alpha_0 z_0^t$

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■ General case : let $P[z] \triangleq \prod_{k=0}^{K-1} (z - z_k) = \sum_{\tau=0}^{K} p_{\tau} z^{K-\tau}$.



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- A discrete signal $\{s_t\}_{t \in \mathbb{Z}}$ is solution of the recursion $\sum_{\tau=0}^K p_\tau \, s_{t-\tau} = 0$ if and only if it is of the form $s_t = \sum_{k=0}^{K-1} \alpha_k \, z_k^{\ t}$



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 - Estimate polynomial P[z] by means of linear prediction
 - Extract the roots of this polynomial





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- Drawback: mediocre performance in presence of noise





直接影響 Prony method

■ Let
$$\varepsilon_t \triangleq \sum_{\tau=0}^K p_\tau x_{t-\tau} = \sum_{\tau=0}^K p_\tau b_{t-\tau}$$
 be the prediction error

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- Let $\varepsilon_t \triangleq \sum_{\tau=0}^K p_\tau x_{t-\tau} = \sum_{\tau=0}^K p_\tau b_{t-\tau}$ be the prediction error
- We let n = K + 1, and we assume that $l \ge K + 1$



超過WProny method

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- We let n = K + 1, and we assume that $l \ge K + 1$

Then
$$\boldsymbol{p}^H \boldsymbol{X}(t) = \varepsilon(t)^H$$
 where $\boldsymbol{p} = [p_K, p_{K-1}, \dots, p_0]^H$,
$$\varepsilon(t) = [\varepsilon_{t-l+K+1}, \varepsilon_{t-l+K+2}, \dots, \varepsilon_{t+K}]^H \text{ and}$$

$$\boldsymbol{X}(t) = \begin{bmatrix} x_{t-l+1} & \dots & x_{t-1} & x_t \\ x_{t-l+2} & \dots & x_t & x_{t+1} \\ \vdots & \dots & \vdots & \vdots \\ x_{t-l+K+1} & \dots & x_{t+K-1} & x_{t+K} \end{bmatrix}$$

图图 Prony method

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■ We minimize $\frac{1}{7} \|\varepsilon\|^2$ w.r.t. \boldsymbol{p} , under the constraint $p_0 = 1$





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- However $\frac{1}{7} \|\varepsilon\|^2 = \boldsymbol{p}^H \widehat{\boldsymbol{R}}_{XX}(t) \, \boldsymbol{p}$, where $\widehat{\boldsymbol{R}}_{XX}(t) = \frac{1}{7} \, \boldsymbol{X}(t) \, \boldsymbol{X}(t)^H$





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- The solution is $\boldsymbol{p} = \frac{1}{\boldsymbol{e}_{1}^{H} \widehat{\boldsymbol{R}}_{xx}(t)^{-1} \boldsymbol{e}_{1}} \widehat{\boldsymbol{R}}_{xx}(t)^{-1} \boldsymbol{e}_{1}$ where $\boldsymbol{e}_{1} \triangleq [1, 0 \dots 0]^{\top}$









直光影 Prony and Pisarenko methods

Prony method :

- Construct matrix $\boldsymbol{X}(t)$ and compute $\hat{\boldsymbol{R}}_{xx}(t)$
- Compute $\boldsymbol{p} = \frac{1}{\boldsymbol{e}_{x}^{H} \widehat{\boldsymbol{R}}_{xx}(t)^{-1} \boldsymbol{e}_{1}} \widehat{\boldsymbol{R}}_{xx}(t)^{-1} \boldsymbol{e}_{1}$
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Prony and Pisarenko methods

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直光图M Prony and Pisarenko methods

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- Method of Pisarenko
 - Minimize $\frac{1}{7} \|\varepsilon\|^2 = \boldsymbol{p}^H \widehat{\boldsymbol{R}}_{xx}(t) \, \boldsymbol{p}$ under the constraint $\|\boldsymbol{p}\|_2 = 1$





Prony and Pisarenko methods

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Prony and Pisarenko methods

Prony method:

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Method of Pisarenko

- Minimize $\frac{1}{T} ||\varepsilon||^2 = \boldsymbol{p}^H \hat{\boldsymbol{R}}_{xx}(t) \boldsymbol{p}$ under the constraint $||\boldsymbol{p}||_2 = 1$
- Solution : \mathbf{p} = eigenvector of $\hat{\mathbf{R}}_{xx}(t)$ of lowest eigenvalue
- Pisarenko method :
 - Construct the matrix X(t) and compute $\hat{R}_{xx}(t)$
 - Diagonalize $\hat{\mathbf{R}}_{xx}(t)$
 - \mathbf{p} = eigenvector of $\hat{\mathbf{R}}_{xx}(t)$ of lowest eigenvalue
 - Determine the z_k 's as the roots of $P[z] = \sum_{k=0}^{K} p_k z^{K-k}$







Part IV

Subspace-based HR methods





Matrix representation of the signal

■ Observation horizon : $t \in \{0 \dots N-1\}$, where N > 2K

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- Data matrix (n > K, I > K and N = n + I 1):

$$oldsymbol{S} = \left[egin{array}{ccccc} oldsymbol{s}_0 & oldsymbol{s}_1 & \dots & oldsymbol{s}_{l-1} \ oldsymbol{s}_1 & oldsymbol{s}_2 & \dots & oldsymbol{s}_l \ dots & dots & dots & dots \ oldsymbol{s}_{n-1} & oldsymbol{s}_n & \dots & oldsymbol{s}_{N-1} \ \end{array}
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$$oldsymbol{S} = \left[egin{array}{cccc} oldsymbol{s}_0 & oldsymbol{s}_1 & \dots & oldsymbol{s}_{l-1} \ oldsymbol{s}_1 & oldsymbol{s}_2 & \dots & oldsymbol{s}_l \ dots & dots & dots & dots \ oldsymbol{s}_{n-1} & oldsymbol{s}_n & \dots & oldsymbol{s}_{N-1} \ \end{array}
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$$\boldsymbol{V}^{n} = \begin{bmatrix} 1 & 1 & \dots & 1 \\ z_{0} & z_{1} & \dots & z_{K-1} \\ z_{0}^{2} & z_{1}^{2} & \dots & z_{K-1}^{2} \\ \vdots & \vdots & \vdots & \vdots \\ z_{0}^{n-1} & z_{1}^{n-1} & \dots & z_{K-1}^{n-1} \end{bmatrix}$$





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直接影響 Empirical covariance matrix

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- In the same way, let $\hat{\mathbf{R}}_{xx} = \frac{1}{7} \mathbf{X} \mathbf{X}^H$ and $\mathbf{R}_{xx} = \mathbb{E} \left[\hat{\mathbf{R}}_{xx} \right]$.
- Then $\mathbf{R}_{xx} = \mathbf{R}_{ss} + \sigma^2 \mathbf{I}_n$





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- In the same way, $Span(\mathbf{W}_{\perp})$ is referred to as the noise subspace
- The poles $\{z_k\}_{k \in \{0...K-1\}}$ are the solutions of equation $\| \boldsymbol{W}_{\perp}^{H} \boldsymbol{v}(z) \|^{2} = 0$, where $\boldsymbol{v}(z) = [1, z, \dots, z^{n-1}]$



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- The MUSIC method consists in solving this equation
- \blacksquare The Spectral-MUSIC method consists in detecting the K highest peaks in function $z \mapsto \frac{1}{\|\mathbf{W}_{\perp}^H \mathbf{v}(z)\|^2}$.



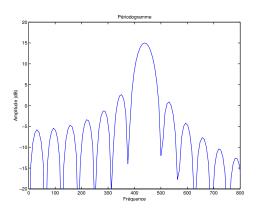
Test signal:

- Sampling frequency: 8000 Hz
- First sinusoid : 440 Hz (A)
- Second sinusoid : 415,3 Hz (G#)
- No damping, all amplitudes equal to 1
- Length of the rectangular window : N = 128 (16 ms)
- Length of the transform : 1024 samples





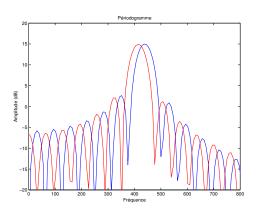
直接影響 Spectral MUSIC method







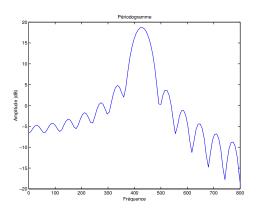
国選擇Man Spectral MUSIC method







直接影響 Spectral MUSIC method

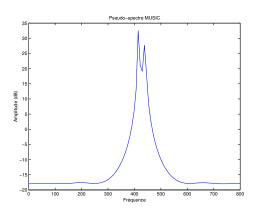








三選記MI Spectral MUSIC method







■選択 ESPRIT method

Rotational invariance property of Vⁿ:

$$\begin{bmatrix}
1 & \dots & 1 \\
z_0 & \dots & z_{K-1} \\
\vdots & \dots & \vdots \\
z_0^{n-2} \dots z_{K-1}^{n-2} \\
z_0^{n-1} \dots z_{K-1}^{n-1}
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直接影 ESPRIT method

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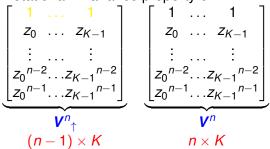
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$$V^n_{\uparrow}$$

$$(n-1) \times K$$

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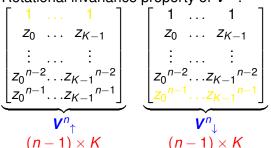


Roland Badeau



多数 ESPRIT method

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\vdots & \dots & \vdots \\
(0) & z_{K-1}
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Part V

Estimation of the other parameters

■ 透園間 Estimation of the modeling order

Information Theoretic Criteria (ITC): we minimize

$$ITC(p) = -(n-p) I \ln \left(\frac{\left(\prod\limits_{q=p+1}^{n} \sigma_q^2\right)^{\frac{1}{n-p}}}{\frac{1}{n-p} \sum\limits_{q=p+1}^{n} \sigma_q^2} \right) + p (2n-p) C(I)$$

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圖選擇MEstimation of the modeling order

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 Maximization of the ratio between the geometrical mean of the eigenvalues of the noise subspace and their arithmetical mean.





直接影欄 Estimation of the modeling order

Information Theoretic Criteria (ITC): we minimize

$$ITC(p) = -(n-p) I \ln \left(\frac{\left(\prod\limits_{q=p+1}^{n} \sigma_q^2 \right)^{\frac{1}{n-p}}}{\frac{1}{n-p} \sum\limits_{q=p+1}^{n} \sigma_q^2} \right) + p (2n-p) C(I)$$

where σ_q^2 are the eigenvalues of $\hat{\boldsymbol{R}}_{xx}$ by decreasing order

- The criterion AIC is given by C(I) = 1, and MDL by $C(I) = \frac{1}{2} \ln(I)$
- The criteria EDC are such that $\lim_{l \to +\infty} \frac{C(l)}{l} = 0$ and

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- Maximization of the ratio between the geometrical mean of the eigenvalues of the noise subspace and their arithmetical mean.
- The maximum is reached when all these eigenvalues are equal. The criterion thus measures the noise whiteness.





超影響 Estimation of the modeling order

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- The criterion AIC is given by C(I) = 1, and MDL by $C(I) = \frac{1}{2} \ln(I)$
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$$\lim_{l \to +\infty} \frac{C(l)}{\ln(\ln(l))} = +\infty$$

- Maximization of the ratio between the geometrical mean of the eigenvalues of the noise subspace and their arithmetical mean.
- The maximum is reached when all these eigenvalues are equal. The criterion thus measures the noise whiteness.
- The penalty term C(I) avoids over-estimating p.







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- We finally get $\hat{a}_k = |\hat{\alpha}_k|$ and $\hat{\phi}_k = \arg(\hat{\alpha}_k)$







Part VI

Performance of the estimators





■祭園間 Cramér-Rao bounds

- Regular statistical model
 - Consider a statistical model $p(\mathbf{x}; \theta)$ parameterized by θ
 - Score function : $I(\mathbf{x}; \theta) \triangleq \nabla_{\theta} \ln p(\mathbf{x}; \theta) \mathbf{1}_{p(\mathbf{x}; \theta) > 0}$ The parameterization is said *regular* if :
 - - 1. $p(x; \theta)$ is continuously differentiable w.r.t. θ .
 - 2. $\mathbf{F}(\theta) \triangleq \int_{\mathcal{H}} \mathbf{I}(\mathbf{x}; \, \theta) \, \mathbf{I}(\mathbf{x}; \, \theta)^{\top} \, p(\mathbf{x}; \, \theta) \, d\mathbf{x}$ (Fisher information matrix) is positive definite for all θ and continuous w.r.t. θ





■ ※ I Cramér-Rao bounds

Regular statistical model

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- Cramér-Rao bounds
 - Consider a regular statistic model parameterized by θ
 - Let $\widehat{\theta}$ be an unbiased estimator of θ ($\forall \theta \in \Theta$, $\mathbb{E}_{\theta}[\widehat{\theta}] = \theta$)
 - Then the dispersion matrix $\mathbf{\mathcal{D}}(\theta,\widehat{\theta}) \triangleq \mathbb{E}_{\theta} \left[\left(\widehat{\theta} \theta \right) \left(\widehat{\theta} \theta \right)^{\top} \right]$ is such that matrix $\mathbf{D}(\theta, \widehat{\theta}) - \mathbf{F}(\theta)^{-1}$ is positive semidefinite.





■終實**聞** Cramér-Rao bounds

For a family of complex Gaussian distributions of covariance $\mathbf{R}_{bb}(\theta) \in \mathcal{C}^1(\Theta, \mathbb{C}^{N \times N})$ and of mean $\mathbf{s}(\theta) \in \mathcal{C}^1(\Theta, \mathbb{C}^N)$,

$$\mathbf{F}_{(i,j)}(\theta) \in \mathcal{C}^{1}(\Theta, \mathbb{C}^{N \times N})$$
 and of mean $\mathbf{s}(\theta) \in \mathcal{C}^{1}(\Theta, \mathbb{C}^{N})$,
$$\mathbf{F}_{(i,j)}(\theta) = \operatorname{trace}\left(\mathbf{R}_{bb}^{-1} \frac{\partial \mathbf{R}_{bb}(\theta)}{\partial \theta_{i}} \mathbf{R}_{bb}^{-1} \frac{\partial \mathbf{R}_{bb}(\theta)}{\partial \theta_{j}}\right) + 2\mathcal{R}e\left(\frac{\partial \mathbf{s}(\theta)}{\partial \theta_{i}}^{H} \mathbf{R}_{bb}^{-1} \frac{\partial \mathbf{s}(\theta)}{\partial \theta_{j}}\right)$$



Roland Badeau

直接影响 Cramér-Rao bounds

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$$m{F}_{(i,j)}(heta) = ext{trace}\left(m{R}_{bb}^{-1} rac{\partial m{R}_{bb}(heta)}{\partial heta_i} m{R}_{bb}^{-1} rac{\partial m{R}_{bb}(heta)}{\partial heta_j}
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■ The Cramér-Rao bounds for the parameters $(\phi_k, \, \delta_k, \, f_k)$ are independent of $a_{k'}$ for all $k' \neq k$, and proportional to $1/a_k^2$

一般意識 Cramér-Rao bounds

For a family of complex Gaussian distributions of covariance $\mathbf{R}_{bb}(\theta) \in \mathcal{C}^1(\Theta, \mathbb{C}^{N \times N})$ and of mean $\mathbf{s}(\theta) \in \mathcal{C}^1(\Theta, \mathbb{C}^N)$.

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- The Cramér-Rao bounds for the parameters (ϕ_k, δ_k, f_k) are independent of $a_{k'}$ for all $k' \neq k$, and proportional to $1/a_k^2$
- **The bound for parameter a_k is independent of all a_{k'}**



Tamér-Rao bounds

■ For a family of complex Gaussian distributions of covariance $\mathbf{R}_{bb}(\theta) \in \mathcal{C}^1(\Theta, \mathbb{C}^{N \times N})$ and of mean $\mathbf{s}(\theta) \in \mathcal{C}^1(\Theta, \mathbb{C}^N)$, $\mathbf{F}_{(i,j)}(\theta) = \operatorname{trace}\left(\mathbf{R}_{bb}^{-1} \frac{\partial \mathbf{R}_{bb}(\theta)}{\partial \theta_i} \mathbf{R}_{bb}^{-1} \frac{\partial \mathbf{R}_{bb}(\theta)}{\partial \theta_j}\right) + 2\mathcal{R}e\left(\frac{\partial \mathbf{s}(\theta)}{\partial \theta_i}^H \mathbf{R}_{bb}^{-1} \frac{\partial \mathbf{s}(\theta)}{\partial \theta_j}\right)$

- The Cramér-Rao bounds for the parameters (ϕ_k, δ_k, f_k) are independent of $a_{k'}$ for all $k' \neq k$, and proportional to $1/a_k^2$
- The bound for parameter a_k is independent of all $a_{k'}$
- All bounds are independent of the phases ϕ_k and are unchanged by any translation of the set of frequencies f_k





■ Cramér-Rao bounds

For a family of complex Gaussian distributions of covariance $\mathbf{R}_{bb}(\theta) \in \mathcal{C}^1(\Theta, \mathbb{C}^{N \times N})$ and of mean $\mathbf{s}(\theta) \in \mathcal{C}^1(\Theta, \mathbb{C}^N)$,

$$m{R}_{bb}(heta) \in \mathcal{C}^{1}(\Theta, \mathbb{C}^{N \times N}) ext{ and of mean } m{s}(heta) \in \mathcal{C}^{1}(\Theta, \mathbb{C}^{N}),$$
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- The Cramér-Rao bounds for the parameters (ϕ_k, δ_k, f_k) are independent of $a_{k'}$ for all $k' \neq k$, and proportional to $1/a_k^2$
- **The bound for parameter a_k is independent of all a_{k'}**
- All bounds are independent of the phases ϕ_k and are unchanged by any translation of the set of frequencies f_k
- If $\forall k, \, \delta_k = 0$ and if $N \to +\infty$, then :

• CRB
$$\{\sigma\} = \frac{\sigma^2}{4N} + O\left(\frac{1}{N^2}\right)$$

• CRB
$$\{f_k\} = \frac{6\sigma^2}{4\pi^2 N^3 a_k^2} + O(\frac{1}{N^4})$$

• CRB
$$\{a_k\} = \frac{2\sigma^2}{N} + O\left(\frac{1}{N^2}\right)$$

CRB
$$\{\phi_k\} \equiv \frac{2\sigma^2}{Na^2} + O\left(\frac{1}{N^2}\right)$$





直接影响 Performance of HR methods

- Performance of an estimator
 - Performance expressed in terms of bias and variance
 - Efficiency: ratio between variance and Cramér-Rao bound
 - An estimator is said efficient if its efficiency is equal to 1.



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Performance of HR methods

- Performance of an estimator
 - Performance expressed in terms of bias and variance
 - Efficiency: ratio between variance and Cramér-Rao bound
 - An estimator is said efficient if its efficiency is equal to 1.
- Maximum likelihood : unbiased and asymptotically efficient $(N \to +\infty)$
- HR methods : results based on the perturbation theory
 - Assumptions : $N \to +\infty$ or SNR $\to +\infty$
 - All the HR methods are asymptotically unbiased
 - The Prony and Pisarenko methods are very inefficient: their variances are significantly greater than the Cramér-Rao bounds.
 - MUSIC and ESPRIT have an asymptotic efficiency close to 1







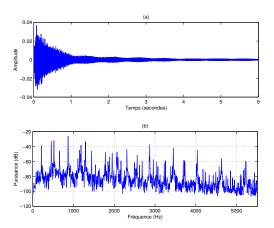
Part VII

Signals to be processed





Bell sound



(a) Signal waveform (b) Power spectral density





