

# BAYESIAN STATISTICS - LECTURE 5

LECTURE 5: REGRESSION. REGULARIZATION PRIORS.

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- **Normal model** with conjugate prior
- The **linear regression** model
- **Non-linear regression**
- **Regularization priors**

## ■ Model

$$y_1, \dots, y_n | \theta, \sigma^2 \stackrel{iid}{\sim} N(\theta, \sigma^2)$$

## ■ Conjugate prior

$$\begin{aligned}\theta | \sigma^2 &\sim N\left(\mu_0, \frac{\sigma^2}{\kappa_0}\right) \\ \sigma^2 &\sim \text{Inv-}\chi^2(\nu_0, \sigma_0^2)\end{aligned}$$

## ■ Posterior

$$\begin{aligned}\theta|y, \sigma^2 &\sim N\left(\mu_n, \frac{\sigma^2}{\kappa_n}\right) \\ \sigma^2|y &\sim \text{Inv-}\chi^2(\nu_n, \sigma_n^2).\end{aligned}$$

where

$$\begin{aligned}\mu_n &= \frac{\kappa_0}{\kappa_0 + n}\mu_0 + \frac{n}{\kappa_0 + n}\bar{y} \\ \kappa_n &= \kappa_0 + n \\ \nu_n &= \nu_0 + n \\ \nu_n\sigma_n^2 &= \nu_0\sigma_0^2 + (n-1)s^2 + \frac{\kappa_0 n}{\kappa_0 + n}(\bar{y} - \mu_0)^2.\end{aligned}$$

## ■ Marginal posterior

$$\theta \sim t_{\nu_n}(\mu_n, \sigma_n^2/\kappa_n)$$

# THE LINEAR REGRESSION MODEL

- The ordinary **linear regression** model:

$$y_i = \beta_1 x_{i1} + \beta_2 x_{i2} + \dots + \beta_k x_{ik} + \varepsilon_i$$
$$\varepsilon_i \stackrel{iid}{\sim} N(0, \sigma^2).$$

- Parameters  $\theta = (\beta_1, \beta_2, \dots, \beta_k, \sigma^2)$ .

- **Assumptions:**

- $E(y_i) = \beta_1 x_{i1} + \beta_2 x_{i2} + \dots + \beta_k x_{ik}$  (linear function)
- $Var(y_i) = \sigma^2$  (homoscedasticity)
- $Corr(y_i, y_j | X, \beta, \sigma^2) = 0, i \neq j$ .
- Normality of  $\varepsilon_i$ .
- The  $x$ 's are assumed known (non-random).

# LINEAR REGRESSION IN MATRIX FORM

- The linear regression model in **matrix form**

$$\underset{(n \times 1)}{\mathbf{y}} = \underset{(n \times k)(k \times 1)}{\mathbf{X}\beta} + \underset{(n \times 1)}{\varepsilon}$$

$$\mathbf{y} = \begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}, \beta = \begin{pmatrix} \beta_1 \\ \vdots \\ \beta_k \end{pmatrix}, \varepsilon = \begin{pmatrix} \varepsilon_1 \\ \vdots \\ \varepsilon_n \end{pmatrix}$$

$$\mathbf{X} = \begin{pmatrix} \mathbf{x}'_1 \\ \vdots \\ \mathbf{x}'_n \end{pmatrix} = \begin{pmatrix} x_{11} & \cdots & x_{1k} \\ \vdots & \ddots & \vdots \\ x_{n1} & \cdots & x_{nk} \end{pmatrix}$$

- Usually  $x_{i1} = 1$ , for all  $i$ .  $\beta_1$  is the intercept.

- **Likelihood**

$$\mathbf{y} | \beta, \sigma^2, \mathbf{X} \sim N(\mathbf{X}\beta, \sigma^2 I_n)$$

- Standard **non-informative prior**: uniform on  $(\beta, \log \sigma^2)$

$$p(\beta, \sigma^2) \propto \sigma^{-2}$$

- **Joint posterior** of  $\beta$  and  $\sigma^2$ :

$$\begin{aligned}\beta | \sigma^2, \mathbf{y} &\sim N[\hat{\beta}, \sigma^2 (\mathbf{X}'\mathbf{X})^{-1}] \\ \sigma^2 | \mathbf{y} &\sim \text{Inv-}\chi^2(n-k, s^2)\end{aligned}$$

where  $\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y}$  and  $s^2 = \frac{1}{n-k}(\mathbf{y} - \mathbf{X}\hat{\beta})'(\mathbf{y} - \mathbf{X}\hat{\beta})$ .

- **Simulate** from the joint posterior by simulating from

- $p(\sigma^2 | \mathbf{y})$
- $p(\beta | \sigma^2, \mathbf{y})$

- **Marginal posterior** of  $\beta$ :

$$\beta | \mathbf{y} \sim t_{n-k}[\hat{\beta}, s^2 (\mathbf{X}'\mathbf{X})^{-1}]$$

## ■ **Joint prior** for $\beta$ and $\sigma^2$

$$\begin{aligned}\beta | \sigma^2 &\sim N(\mu_0, \sigma^2 \Omega_0^{-1}) \\ \sigma^2 &\sim \text{Inv} - \chi^2(\nu_0, \sigma_0^2)\end{aligned}$$

## ■ **Posterior**

$$\begin{aligned}\beta | \sigma^2, \mathbf{y} &\sim N[\mu_n, \sigma^2 \Omega_n^{-1}] \\ \sigma^2 | \mathbf{y} &\sim \text{Inv} - \chi^2(\nu_n, \sigma_n^2)\end{aligned}$$

$$\mu_n = (\mathbf{X}'\mathbf{X} + \Omega_0)^{-1} (\mathbf{X}'\mathbf{X}\hat{\beta} + \Omega_0\mu_0)$$

$$\Omega_n = \mathbf{X}'\mathbf{X} + \Omega_0$$

$$\nu_n = \nu_0 + n$$

$$\nu_n \sigma_n^2 = \nu_0 \sigma_0^2 + (\mathbf{y}'\mathbf{y} + \mu_0' \Omega_0 \mu_0 - \mu_n' \Omega_n \mu_n)$$



# POLYNOMIAL REGRESSION

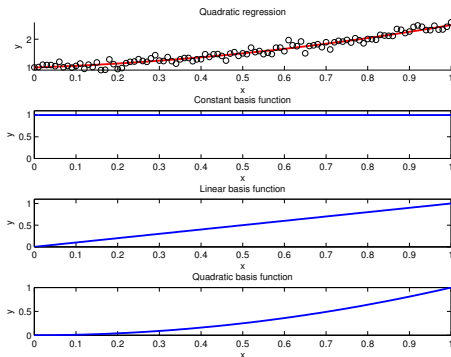
## ■ Polynomial regression

$$f(x_i) = \beta_0 + \beta_1 x_i + \beta_2 x_i^2 + \dots + \beta_k x_i^k.$$

$$\mathbf{y} = \mathbf{X}_p \boldsymbol{\beta} + \boldsymbol{\varepsilon},$$

where

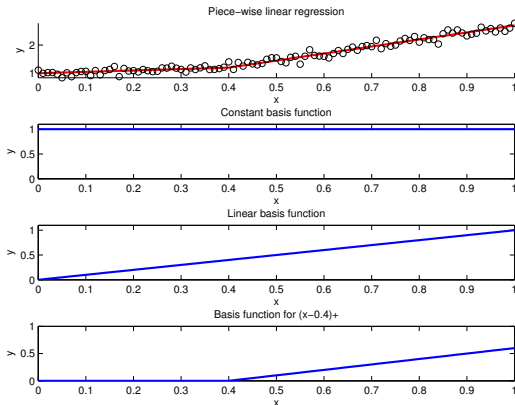
$$\mathbf{X}_p = (1, x, x^2, \dots, x^k).$$



# SPLINE REGRESSION

- Polynomials are too global. Need more local basis functions.
- Truncated power splines given knot locations  $k_1, \dots, k_m$

$$b_{ij} = \begin{cases} (x_i - k_j)^p & \text{if } x_i > k_j \\ 0 & \text{otherwise} \end{cases}$$



- **Spline regression is linear** in the  $m$  'dummy variables'  $b_j$

$$\mathbf{y} = \mathbf{X}_b \boldsymbol{\beta} + \varepsilon,$$

where  $\mathbf{X}_b$  is the **basis matrix**

$$\mathbf{X}_b = (b_1, \dots, b_m).$$

- Adding intercept and linear term

$$\mathbf{X}_b = (1, x, b_1, \dots, b_m).$$

# SMOOTHNESS PRIOR FOR SPLINES

- Problem: too many knots leads to **over-fitting**.
- **Smoothness/shrinkage/regularization prior**

$$\beta_i | \sigma^2 \stackrel{iid}{\sim} N\left(0, \frac{\sigma^2}{\lambda}\right)$$

- Larger  $\lambda$  gives smoother fit. Note:  $\Omega_0 = \lambda I$ .
- Equivalent to **penalized likelihood**:

$$-2 \cdot \log p(\beta | \sigma^2, \mathbf{y}, \mathbf{X}) \propto \text{RSS}(\beta) + \lambda \beta' \beta$$

- Posterior mean gives **ridge regression** estimator

$$\tilde{\beta} = (\mathbf{X}'\mathbf{X} + \lambda I)^{-1} \mathbf{X}'\mathbf{y}$$

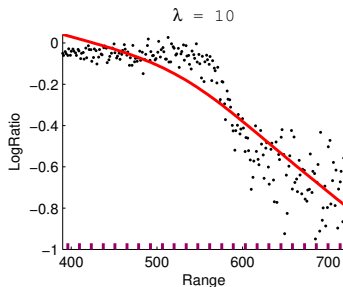
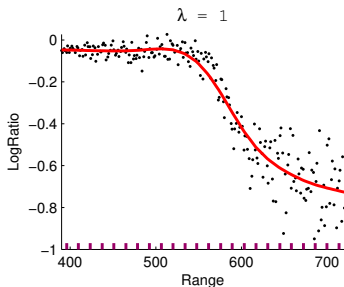
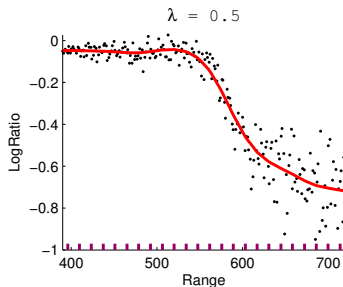
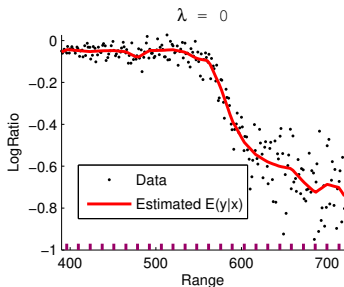
- **Shrinkage** toward zero

$$\text{As } \lambda \rightarrow \infty, \tilde{\beta} \rightarrow 0$$

- When  $\mathbf{X}'\mathbf{X} = I$

$$\tilde{\beta} = \frac{1}{1 + \lambda} \hat{\beta}_{OLS}$$

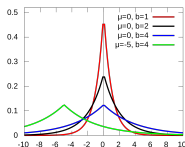
# BAYESIAN SPLINE WITH SMOOTHNESS PRIOR



# SMOOTHNESS PRIOR FOR SPLINES, CONT.

- **Lasso** is equivalent to posterior mode under Laplace prior

$$\beta_i | \sigma^2 \stackrel{iid}{\sim} \text{Laplace} \left( 0, \frac{\sigma^2}{\lambda} \right)$$



- The **Bayesian shrinkage** prior is **interpretable**. **Not ad hoc**.
- Laplace distribution have heavy tails.
- **Laplace prior**: many  $\beta_i$  close to zero, but some  $\beta_i$  very large.
- Normal distribution have light tails.
- **Normal prior**: all  $\beta_i$ 's are similar in magnitude.

# ESTIMATING THE SHRINKAGE

- Cross-validation is often used to determine the degree of smoothness,  $\lambda$ .
- Bayesian:  $\lambda$  is **unknown**  $\Rightarrow$  **use a prior** for  $\lambda$ .
- $\lambda \sim \text{Gamma}\left(\frac{\eta_0}{2}, \frac{\eta_0}{2\lambda_0}\right)$ . The user specifies  $\eta_0$  and  $\lambda_0$ .
- Hierarchical setup:

$$\begin{aligned}\mathbf{y}|\beta, \mathbf{X} &\sim N(\mathbf{X}\beta, \sigma^2 I_n) \\ \beta|\sigma^2, \lambda &\sim N(\mathbf{0}, \sigma^2 \lambda^{-1} I_m) \\ \sigma^2 &\sim \text{Inv} - \chi^2(\nu_0, \sigma_0^2) \\ \lambda &\sim \text{Gamma}\left(\frac{\eta_0}{2}, \frac{\eta_0}{2\lambda_0}\right)\end{aligned}$$

$$\text{so } \Omega_0 = \lambda I_m.$$

- The **joint posterior** of  $\beta$ ,  $\sigma^2$  and  $\lambda$  is

$$\beta | \sigma^2, \lambda, \mathbf{y} \sim N(\mu_n, \Omega_n^{-1})$$

$$\sigma^2 | \lambda, \mathbf{y} \sim \text{Inv} - \chi^2(\nu_n, \sigma_n^2)$$

$$p(\lambda | \mathbf{y}) \propto \sqrt{\frac{|\Omega_0|}{|\mathbf{X}^T \mathbf{X} + \Omega_0|}} \left( \frac{\nu_n \sigma_n^2}{2} \right)^{-\nu_n/2} \cdot p(\lambda)$$

where  $\Omega_0 = \lambda I_m$ , and  $p(\lambda)$  is the prior for  $\lambda$ , and

$$\mu_n = (\mathbf{X}^T \mathbf{X} + \Omega_0)^{-1} \mathbf{X}^T \mathbf{y}$$

$$\Omega_n = \mathbf{X}^T \mathbf{X} + \Omega_0$$

$$\nu_n = \nu_0 + n$$

$$\nu_n \sigma_n^2 = \nu_0 \sigma_0^2 + \mathbf{y}^T \mathbf{y} - \mu_n^T \Omega_n \mu_n$$



- The **location of the knots** can be unknown. Joint posterior:

$$p(\beta, \sigma^2, \lambda, k_1, \dots, k_m | \mathbf{y}, \mathbf{X})$$

- The marginal posterior for  $\lambda, k_1, \dots, k_m$  is a nightmare.
- Simulate from joint posterior by MCMC. Li and Villani (2013).
- The basic spline model can be extended with:
  - **Heteroscedastic errors** (also modelled with a spline)
  - **Non-normal errors** (student-t or mixture distributions)
  - **Autocorrelated/dependent errors** (AR process for the errors)