

$$1a) \quad X_1, \dots, X_n | \theta, \sigma^2 \sim N(\theta, \sigma^2)$$

θ known

$$\sigma^2 \sim \text{Inv}\chi^2(\nu_0, \sigma_0^2)$$

EXERCISE
SET NO. 2
BAYESIAN
LEARNING

Posterior: (implicit conditioning on θ)

$$p(\sigma^2 | x_1, \dots, x_n) \propto p(x_1, \dots, x_n | \sigma^2) p(\sigma^2)$$

$$= \prod_{i=1}^n \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left(-\frac{1}{2\sigma^2}(x_i - \theta)^2\right) \cdot p(\sigma^2)$$

$$\left[\text{define } s^2 = \frac{\sum_{i=1}^n (x_i - \theta)^2}{n} \right]$$

$$\propto \frac{1}{(\sigma^2)^{n/2}} \exp\left(-\frac{ns^2}{2\sigma^2}\right) \exp\left(-\frac{\nu_0 \sigma_0^2}{2\sigma^2}\right) \frac{1}{(\sigma^2)^{\nu_0/2 + 1}}$$

$$= \frac{\exp\left(-\frac{ns^2 + \nu_0 \sigma_0^2}{2\sigma^2}\right)}{(\sigma^2)^{(n+\nu_0)/2 + 1}}$$

Density (pdf) of
 $\text{Inv}\chi^2(\nu_0, \sigma_0^2)$
prior.

$$\text{So } \sigma^2 | x_1, \dots, x_n, \theta \sim \text{Inv}\chi^2(\nu_n, \sigma_n^2)$$

$$\nu_n = \nu_0 + n$$

$$\sigma_n^2 = \frac{ns^2 + \nu_0 \sigma_0^2}{\nu_0 + n}$$

$$1b) \quad s^2 = \frac{\sum_{i=1}^3 (x_i - \theta)^2}{3} = \frac{(0.6-1)^2 + (3.2-1)^2 + (1.2-1)^2}{3} = 1.68$$

Non-informative : $\nu_0 \rightarrow 0$

Why is this non-informative?

Reason 1: ν_n becomes n

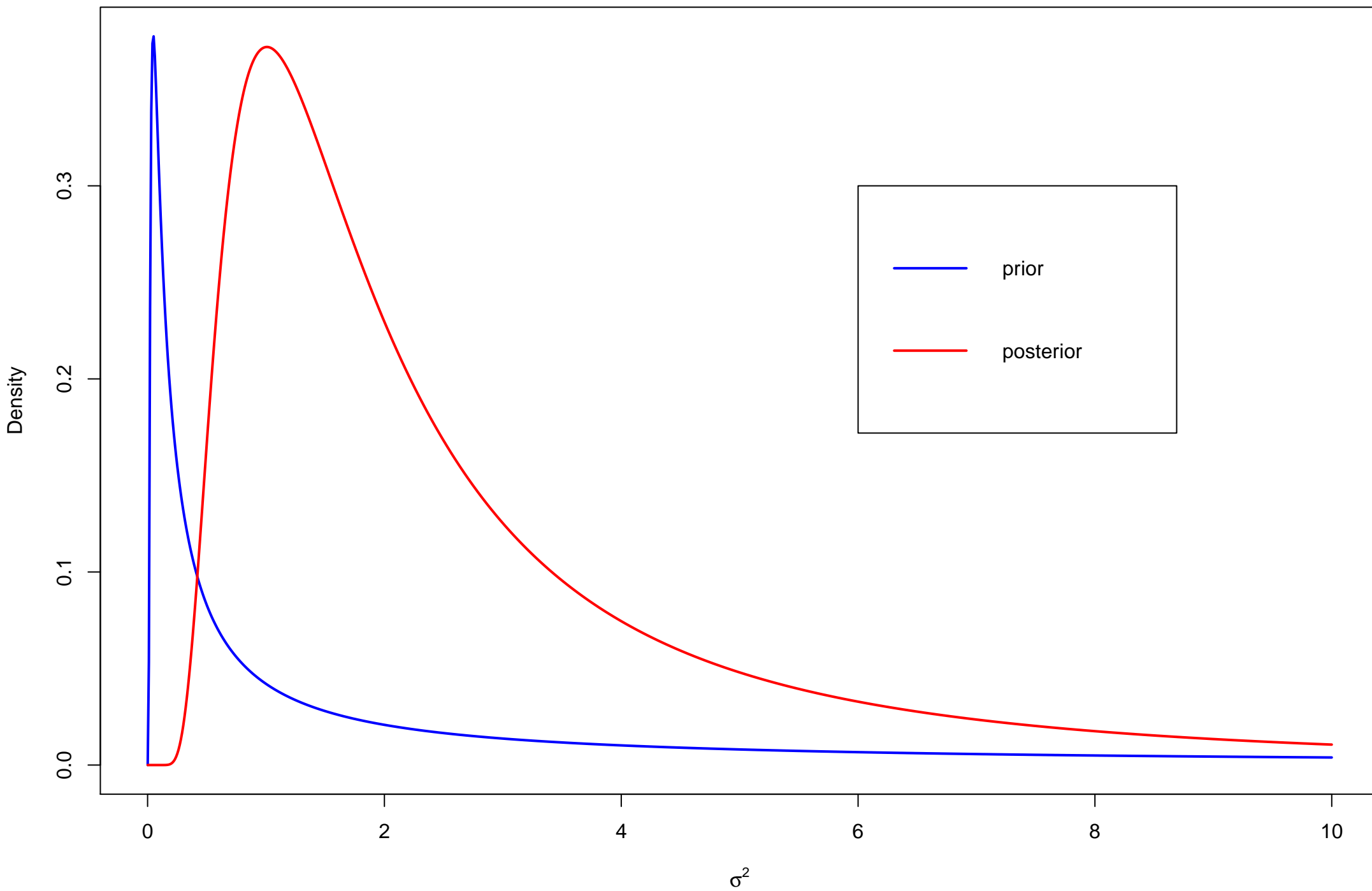
Reason 2: $\text{Inv}\chi^2(\nu_0, \sigma_0^2)$ becomes $\frac{1}{\sigma^2}$
when $\nu_0 \rightarrow 0$.

Note that as $\nu_0 \rightarrow 0$ the posterior approaches the $\text{Inv}\chi^2(n, s^2)$ density.

So,

$$\sigma^2 | x_1, x_2, x_3 \sim \text{Inv}\chi^2(3, 1.68)$$

Prior is InvChi(0.1,s2)



Problems of Chapter 6

6.1 Prediction of Bernoulli data

The predictive distribution of x_{n+1} given the first n trials ($x_{1:n}$) is

$$\begin{aligned}
 p(x_{n+1}|x_{1:n}) &= \int p(x_{n+1}|\theta)p(\theta|x_{1:n})d\theta && x_{n+1} \text{ is indep. of } x_{1:n} \text{ given } \theta \\
 &= \int \theta^{x_{n+1}}(1-\theta)^{1-x_{n+1}}p(\theta|x_{1:n})d\theta && \theta|x_{1:n} \sim \text{Beta}(\alpha+s, \beta+f) \\
 &= \int \theta^{x_{n+1}}(1-\theta)^{1-x_{n+1}} \frac{\Gamma(\alpha+\beta+n)}{\Gamma(\alpha+s)\Gamma(\beta+f)} \theta^{\alpha+s-1}(1-\theta)^{\beta+f-1}d\theta \\
 &= \frac{\Gamma(\alpha+\beta+n)}{\Gamma(\alpha+s)\Gamma(\beta+f)} \int \theta^{x_{n+1}+\alpha+s-1}(1-\theta)^{1-x_{n+1}+\beta+f-1}d\theta \\
 &= \frac{\Gamma(\alpha+\beta+n)}{\Gamma(\alpha+s)\Gamma(\beta+f)} \frac{\Gamma(x_{n+1}+\alpha+s)\Gamma(1-x_{n+1}+\beta+f)}{\Gamma(1+\alpha+\beta+n)} \\
 &= \frac{\Gamma(x_{n+1}+\alpha+s)\Gamma(1-x_{n+1}+\beta+f)}{\Gamma(\alpha+s)\Gamma(\beta+f)(\alpha+\beta+n)} && \text{using } \Gamma(y+1) = y\Gamma(y)
 \end{aligned}$$

So,

$$p(x_{n+1} = 1|x_{1:n}) = \frac{\Gamma(1+\alpha+s)}{\Gamma(\alpha+s)(\alpha+\beta+n)} = \frac{(\alpha+s)\Gamma(\alpha+s)}{\Gamma(\alpha+s)(\alpha+\beta+n)} = \frac{\alpha+s}{\alpha+\beta+n}$$

and therefore [since $p(x_{n+1} = 0|x_{1:n}) = 1 - p(x_{n+1} = 1|x_{1:n})$]

$$p(x_{n+1} = 0|x_{1:n}) = \frac{\beta+f}{\alpha+\beta+n}.$$

The predictive distribution is therefore

$$x_{n+1}|x_{1:n} \sim \text{Bern}\left(\frac{\alpha+s}{\alpha+\beta+n}\right).$$

7.1 Umbrella decision

(a) Let x_{11} be the binary variable indicating rain on the 11th day. From Problem 6.1, the predictive distribution for the $(n+1)$ th Bernoulli trial is

$$x_{n+1}|x_{1:n} \sim \text{Bern}\left(\frac{\alpha+s}{\alpha+\beta+n}\right).$$

and the predictive probability for rain is therefore here

$$\Pr(x_{11} = 1|x_{1:10}) = \frac{1+2}{1+1+10} = 0.25.$$

The expected utility from the decision to bring the umbrella is then

$$EU_{\text{bring}} = \Pr(\text{sunny}) \cdot U(\text{bring}, \text{sunny}) + \Pr(\text{rain}) \cdot U(\text{bring}, \text{rain}) = 0.75 \cdot 20 + 0.25 \cdot 10 = 17.5$$

and the expected utility of leaving the umbrella at home is

$$EU_{\text{leave}} = \Pr(\text{sunny}) \cdot U(\text{leave}, \text{sunny}) + \Pr(\text{rain}) \cdot U(\text{leave}, \text{rain}) = 0.75 \cdot 50 + 0.25 \cdot (-50) = 25.0.$$

The expected utility is therefore maximized by leaving the umbrella at home. This is the Bayesian decision.

(b) Figure 15.1 shows how the optimal Bayesian decision varies for different combinations of the prior hyperparameters.

(c) Figure 15.2 shows how the optimal Bayesian decision varies for different combinations of the prior hyperparameters when $s = 16$ and $f = 64$.

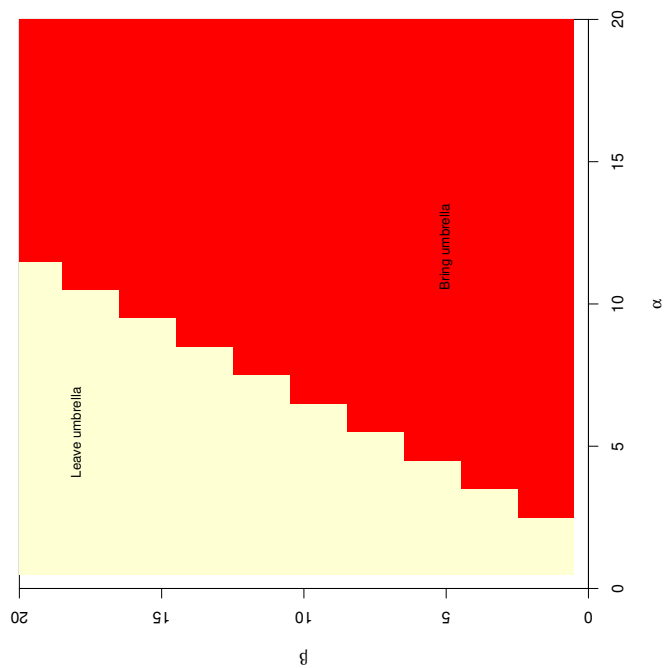


Fig. 15.1. How the Bayesian decision depends on the prior hyperparameters when $s = 2$ and $f = 8$

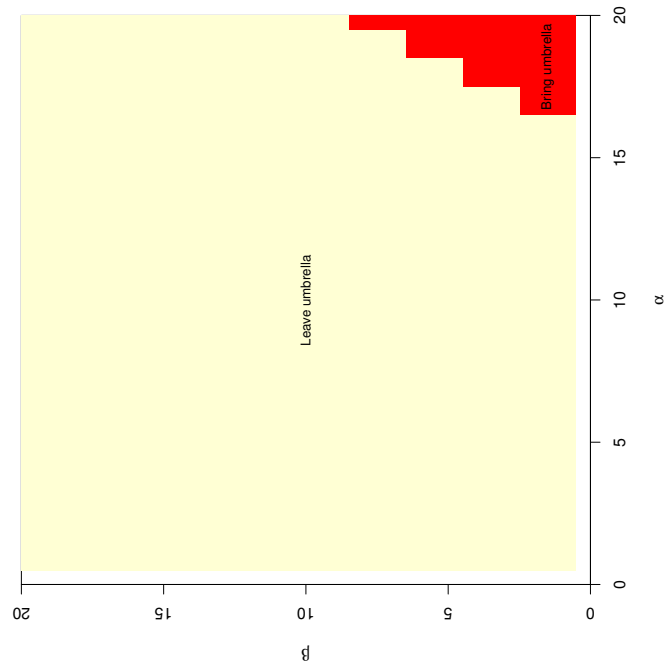


Fig. 15.2. How the Bayesian decision depends on the prior hyperparameters when $s = 16$ and $f = 64$

3a)

Predictive distribution of X_6 :

$$X_6 | X_{1:5} \sim N(\mu_n, \sigma^2 + \tau_n^2)$$

as shown in Lecture 4

Slide 6.

$$\text{Now, } \tau_n^2 = \frac{1}{\tau_0^2 + \frac{n}{\sigma^2}} = \frac{1}{\frac{1}{50^2} + \frac{5}{25^2}} \\ \approx 119$$

$$\mu_n = w\bar{x} + (1-w)\mu_0$$

$$\text{with } w = \frac{\frac{n}{\sigma^2}}{\frac{1}{\tau_0^2} + \frac{n}{\sigma^2}} = \frac{\frac{5}{25^2}}{\frac{1}{50^2} + \frac{5}{25^2}} = 0.95$$

So,

$$w = 0.95 \cdot 320.4 + 0.05 \cdot 200 \approx 314$$

$$X_6 | X_{1:5} \sim N(314, 25^2 + 119) \\ \sim N(314, 27.2^2)$$

$$3b) \quad U(a = \text{no campaign}) = (p-q)X_6 - 0 = 5X_6$$

$$\begin{aligned} U(a = \text{campaign}) &= (p-q)X_6 - 400 \\ &= 5X_6 - 400 \end{aligned}$$

Expected utility for $a = \text{no campaign}$:

$$EU_{\text{no campaign}} = E(5X_6 | X_{1:6}) = 5 \cdot 314 = 1570$$

$$\begin{aligned} EU_{\text{campaign}} &= E(5X_6 - 400 | X_{1:6}) \\ &= 5 \cdot (314 + 100) - 400 \\ &= 1670. \end{aligned}$$

So, we should do the campaign

Since it maximizes expected utility.

Note that the campaign shifted the whole predictive distribution $P(X_6 | X_{1:5})$ to the right by 100 units, that is why the predictive mean $E(X_6 | X_{1:5})$ also was 100 units larger under the campaign scenario.