

# Statistical Techniques and Time Series Coursework

## Task 1. Obtain the time series.

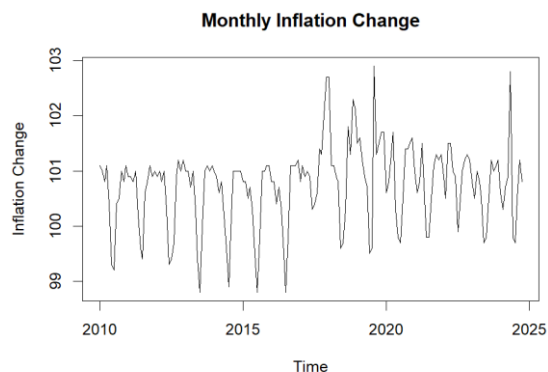


Figure 1. Monthly Inflation Change in Uzbekistan (Compared to Previous Month).

Figure 1 shows monthly inflation change in Uzbekistan compared to previous month in percentage. The figure was obtained using the following code:

```
plot(ts_data, main = "Monthly Inflation Change in Uzb", ylab = "Inflation Change", xlab = "Time")
```

The time series span from January 2010 to October 2024. There is a noticeable change which occurred in around 2017, where the inflation rose to approximately 102.5% and kept varying within 2% over the remaining period.

## Task 2. Detrend the series.

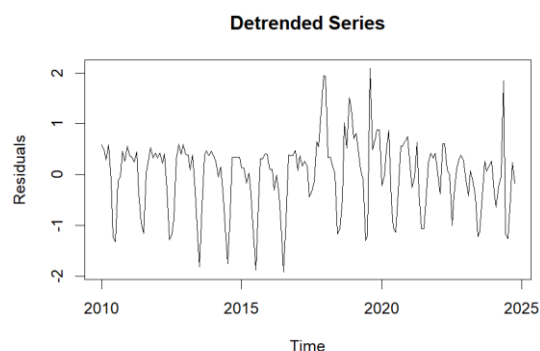


Figure 2. Detrended Time Series.

Figure 2 demonstrates the detrended time series (DTS) from Figure 1 which was obtained in two steps:

1. Identify the trend using a linear regression model (which has a formula of  $y_t = \beta_0 + \beta_1 * t + \epsilon_t$ ).
2. Subtract the trend from the series.

The line graph was plot with the following code:

```
plot(detrended_series, main = "Detrended Series", ylab = "Residuals", xlab = "Time")
```

It is worth noting that there is a noticeable seasonality throughout the period.

## Task 3. Create ACF & PACF. Is it a unit root?

The Figure 3 demonstrates the AutoCorrelation Function (ACF) and Partial AutoCorrelation Function (PACF). Because the DTS and its lagged version plummet at around 0.1 lag, the time series seem to be stationary.

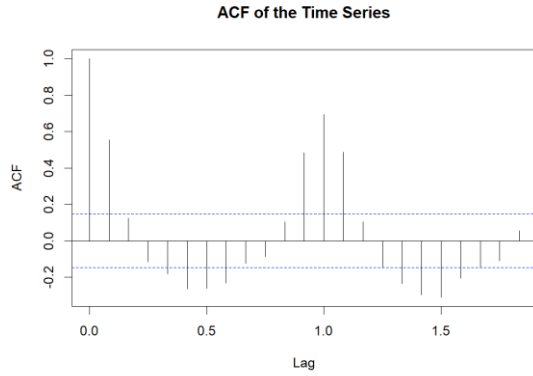


Figure 3.1. The ACF Plot of the DTS.

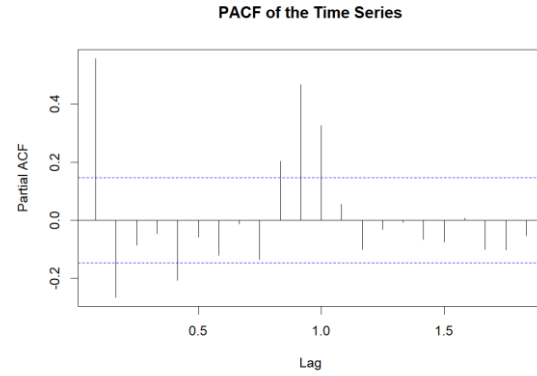


Figure 3.2. The PACF Plot of the DTS.

The ACF is defined by the function at lag  $k$ :

$$\rho_k = \frac{E[(x_t - \mu)(x_{t-k} - \mu)]}{E[x_t - \mu]^2}$$

Where  $E$  – the expected value,  $x_t$  – DTS, and  $\mu$  – the mean of the time series. (Adhikari and Agrawal, 2013).

While PACF is defined as follows:

$$\phi_k = \text{Corr}(x_t, x_{t+k} | x_{t+1}, x_{t+2}, \dots, x_{t+k-1})$$

Where  $\phi_k$  is the PACF,  $x_t$  – DTS,  $k$  – the lag order; and Corr is the correlation. (Frain, 1992).

**Task 4.** Check for autocorrelation. Compare the results with the task 3.

After performing the Ljung-Box test with the code below, where the null hypothesis ( $H_0$ ) suggests that the residuals are distributed independently, while the alternative ( $H_1$ ) is there is significant autocorrelation:

```
Box.test(detrended_series, type = 'Ljung-Box')
```

The resulted p-value =  $7.772e-14 < 0.05$  provides enough evidence to reject the null hypothesis, and concluding that there is significant autocorrelation on at least one lag. This is consistent with both ACF and PACF plots which have noticeable spikes at lag 1.0 suggesting that there is significant correlation at that lag.

**Task 5.** Check for Stationarity and compare the results with the task 3.

To accurately assess the non-/stationarity of the DTS, two tests were performed.

1. Augmented Dickey-Fuller Test, where  $H_0$  suggests that the series has a unit root, and therefore, is not stationary, while the alternative is the stationarity.  
After running the test:  
`adf.test(detrended_series)`  
The resulting p-value of  $0.01 < 0.05$  provides enough evidence to reject  $H_0$ . Thus, the DTS is stationary.

2. To ensure the accuracy of such statement, the second test, Kwiatkowski-Phillips-Schmidt-Shin (KPSS), was performed:  
`kpss.test(detrended_series)`  
 The  $H_0$  is stationary DTS, while  $H_1$  is non-stationarity. The p-value = 0.1 does not provide enough evidence to reject  $H_0$ , implying the stationarity of the DTS.

Both results are consistent with the observations from ACF and PACF in task 3. Therefore, there is significant evidence that the series are stationary.

#### Task 6. Normality test.

To ensure the accuracy of the assessment, three tests for normality were performed. In all three tests,  $H_0$  suggests a normally distributed time series, while  $H_1$  states that the DTS is not normally distributed.

1. Lilliefors (Kolmogorov-Smirnov):  
`lillie.test(coredata(detrended_series))`      # p-value = 2.975e-05
2. Anderson-Darling:  
`ad.test(as.numeric(coredata(detrended_series)))`      # p-value = 4.002e-10
3. Shapiro-Francia:  
`sf.test(detrended_series)`      # p-value = 1.032e-05

Small p-values in all the three tests provide enough evidence to reject  $H_0$ . Thus, the series is not normally distributed.

#### Task 7. Fit an ARMA model & determine the best lag.

As noted in task 2, the DTS has a clear seasonality. To fit the data into the suitable model, the seasonality must be accounted for. Thus, Seasonal AutoRegressive Integrated Moving Average (SARIMA) model would be the best candidate. The model is defined in R by the formula:

$$model = ARMA(p, d, q)[P, D, Q](s)$$

Where,  $p$  – is the number of AR terms,  $d$  – differencing,  $q$  – the number of MA terms; while  $P$  – is the number of seasonal AR terms,  $D$  – seasonal differencing, and  $Q$  – number of seasonal MA terms with specified period  $s$  (Fraim, 1992). After implementing the model in R:

```
sarima_model = arima(detrended_series, order = c(1, 0, 1), seasonal = list(order = c(0, 0, 1), period = 12))
```

where  $p=1$ ,  $d=0$ ,  $q=1$ ,  $P=D=0$ ,  $Q=1$ ,  $s=12$ . The parameters were selected from ACF ( $q$ ,  $Q$ , because lag 1 is a significant spike, outside the blue line) and PACF ( $p$ ,  $P$ , because lag 1 is a significant spike, outside the blue line).  $D$  and  $d$  were set to 0 as the time series was detrended in task 2.

After experimenting with the parameters above, the final model was constructed:

```
sarima_model = arima(detrended_series, order = c(1, 0, 2), seasonal = list(order = c(0, 1, 2), period = 12))
```

The initial model's Akaike Information Criterion (AIC) was 283.31, while the final model's AIC = 204.92, the lower value suggesting that the final model fits the series better.

The AIC of the final model is even slightly better than the auto.arma suggested model which produced AIC = 205.02

**Task 8.** The coefficients, the equation for the model. Make 1-step ahead prediction.

The final model's coefficients were:

```
ar1  ma1  ma2  sma1  sma2
0.9383 -0.6924 -0.1264 -0.7106 -0.1064
```

This implies that the final model's mathematical equation is:

$$(1 - 0.9383B)x_t = (1 - 0.6924B - 0.1264B^2)(1 - 0.7106B^{12} - 0.1064B^{24})\epsilon_t$$

Where B is the backshift operator,  $x_t$  – the DTS, and  $\epsilon_t$  – is the white noise. (Adhikari and Agrawal, 2013).

To forecast with the model 1 day ahead, the following code was used:

```
forecast(sarima_model, h=1)
```

And produce the following result:

```
Point Forecast  Lo 80  Hi 80  Lo 95  Hi 95
Nov 2024      0.3328849 -0.2026685 0.8684383 -0.4861734 1.151943
```

Indicating that the November's inflation rate would be around 99.5138266% (slight decrease), or 101.151943% (slight increase) compared to previous month (i.e., October) with 95% confidence, and the most likely value being 100.3328849%.

**Task 9.** Check for residual mean, finite & constant variance, zero autocovariance.

To access whether the fitted model is a good fit for the data, the residuals should behave like white noise as outline in the task specifications. Therefore, three tests were performed to examine these criteria:

1. **Mean:** `mean(residuals(sarima_model))` # -0.006575348, which is close enough to 0.
2. Studentized Breusch-Pagan test for **constant variance:** `bptest(residuals(sarima_model) ~ fitted(sarima_model))` # p-value = 0.3843, i.e., not enough evidence to reject  $H_0$ , thus suggesting that the variance is constant.  
 $H_0$  – Homoscedasticity (variance is constant)  
 $H_1$  – heteroscedasticity (variance is not constant)
3. Ljung-Box for **autocorrelation:** `Box.test(residuals(sarima_model), lag = 12, type = "Ljung-Box")` # see Task 4 for the hypotheses; p-value = 0.7837 => not enough evidence to reject  $H_0$ , thus, suggesting no significant autocorrelation.

To conclude, the tests above showcase that the residuals behave as white noise implying that the constructed SARIMA model is the good fit for the data.

## References:

Adhikari, R. & Agrawal, R.K. (2013), "Deep learning for time series modelling". Available at: <https://arxiv.org/pdf/1302.6613> (Accessed: 1 December 2024).

Frain, J. (1992), *Lecture Notes on Univariate Time Series Analysis and Box Jenkins Forecasting*. Available at: [https://www.tcd.ie/Economics/staff/frainj/main/MSc%20Material/Session\\_notes/UNIVAR4.pdf](https://www.tcd.ie/Economics/staff/frainj/main/MSc%20Material/Session_notes/UNIVAR4.pdf) (Accessed: 5 December 2024).

OpenAI (2023) *ChatGPT<sup>1</sup> response to user queries*, 27 November - 8 December. Available at: <https://chat.openai.com/> (Accessed: 27 November 2024).

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<sup>1</sup> The code was partially written using ChatGPT.