ENV 710: Lab 7

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```
library(ggplot2)
require(mosaic)
```

Problem 1

The null hypothesis is that the average weight of confiscated elephant tusks is the same across the years 1970, 1990, and 2010. While the alternative hypothesis is that the average weight of confiscated elephant tusks has decreased over time in the three year groups.

To test this hypothesis, we will use a one-way ANOVA design with the year as a factor because there is only one independent variable for this model. Besides, the number of the subjects in this situation is too small to conduct a random effect analysis. Thus, I construct the model as follows:

```
Tusk <- read.csv("TuskData.csv")
lm1 <- lm(Tusk.kg ~ factor(Year), data = Tusk)
summary(lm1)</pre>
```

```
##
## Call:
## lm(formula = Tusk.kg ~ factor(Year), data = Tusk)
##
## Residuals:
##
       Min
                1Q Median
                                3Q
                                       Max
  -3.6964 -0.9710 -0.0348 0.9109
                                   3.9427
## Coefficients:
                    Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                                 0.3438
                                          34.08
                                                  <2e-16 ***
                     11.7150
## factor(Year)1990
                     -8.7916
                                 0.4862
                                        -18.08
                                                  <2e-16 ***
## factor(Year)2010
                     -9.5400
                                 0.4862 -19.62
                                                  <2e-16 ***
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
## Residual standard error: 1.537 on 57 degrees of freedom
## Multiple R-squared: 0.8931, Adjusted R-squared: 0.8894
## F-statistic: 238.1 on 2 and 57 DF, p-value: < 2.2e-16
```

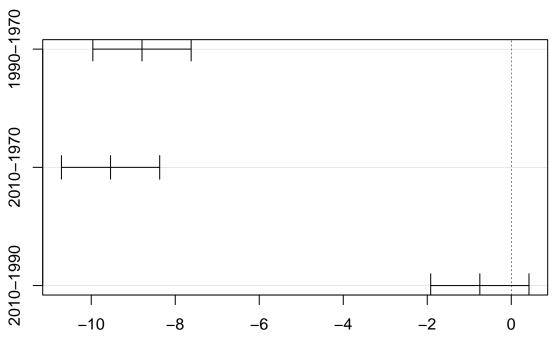
The ANOVA results show that there is a statistically significant difference in the mean weight of elephant tusks across the three years (F(2, 57) = 238.1, p < 2.2e-16). The mean weight of tusks appears to have decreased over time, with the lowest average weight in 2010.

To follow up on this result, I have performed a Tukey's HSD test and its illustration to determine which pairs of years have significantly different mean weights.

TukeyHSD(lm1)

```
##
     Tukey multiple comparisons of means
       95% family-wise confidence level
##
##
## Fit: aov(formula = x)
##
## $'factor(Year)'
##
                   diff
                                lwr
                                           upr
                                                   p adj
## 1990-1970 -8.7915546
                         -9.961482 -7.6216274 0.0000000
## 2010-1970 -9.5400032 -10.709930 -8.3700759 0.0000000
## 2010-1990 -0.7484485 -1.918376 0.4214787 0.2803896
plot(TukeyHSD(lm1))
```

95% family-wise confidence level



Differences in mean levels of factor(Year)

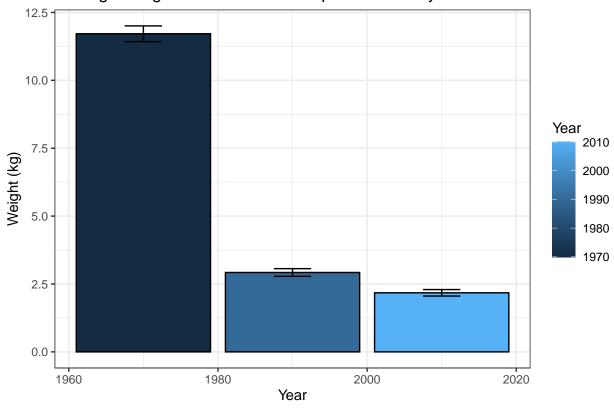
In this example, there is not a statistically significant difference between year 2010-1990 as illustrated by the fact that the CI overlaps 0 and the p-value is greater than 0.05. By contrast, there appears to be significant differences in biomass between year 2010-1970 and year 1990-1970.

Besides, I have also visualized the results via a barplot of the means of tusk weight and standard errors per year group with error bars.

```
# Calculate mean and standard error for each year
mean_data <- aggregate(Tusk$Tusk.kg, by=list(Year=Tusk$Year), FUN=mean)
names(mean_data)[2] <- "Mean"
se_data <- aggregate(Tusk$Tusk.kg, by=list(Year=Tusk$Year), FUN=sd)
se_data$se <- se_data$x / sqrt(length(Tusk$Tusk.kg))

# Create bar plot with error bars
ggplot(mean_data, aes(x=Year, y=Mean, fill=Year)) +
    geom_bar(stat="identity", position="dodge", color="black") +
    geom_errorbar(aes(ymin=Mean-se_data$se, ymax=Mean+se_data$se), width=5) +
    labs(title="Average Weight of Confiscated Elephant Tusks by Year", x="Year", y="Weight (kg)") +
    theme_bw()</pre>
```

Average Weight of Confiscated Elephant Tusks by Year

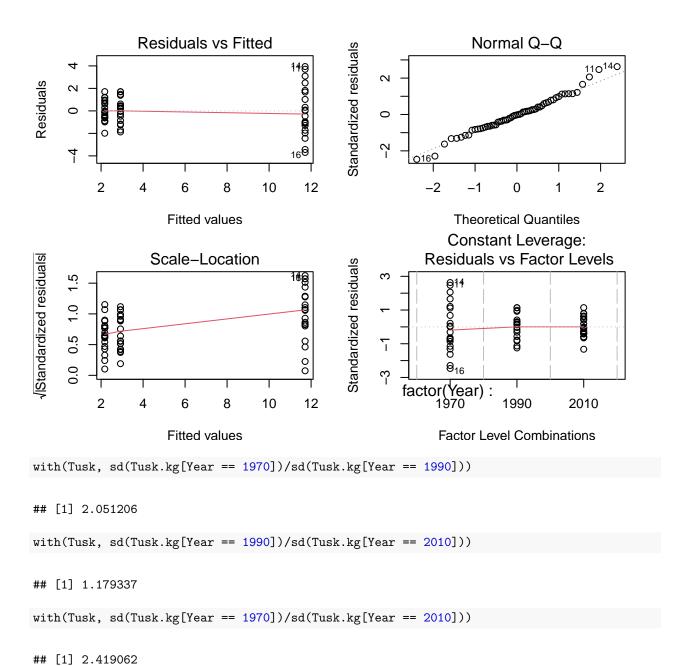


Then, I checked the assumptions of the statistical test using diagnosis plots. First, In the residual plot, the residuals randomly scattered around the centerline, indicating residuals roughly normally and approximately independently distributed with a mean of 0 and constant variance. Second, the qq plot shows a linear trend, indicating the normality of the dependent variables is good.

Third, however, the figure of standardized residual plot is worrying since the variance seems to increase with the treatment means, which might be evidence of heteroscedasticity. Thus, I double checked the ratios of the sample standard deviations and find some of them exceed the limit of 2. To remedy this, I transform the dependent variable into a log format and rerun the model. But the ratio also cannot meet the requirement. Thus, I decide to ignore this problem since other plots are good.

At last, this plot, the leverage plot does not reflect any outliers, although observations 14 and 16 have more leverage than other observations.

```
par(mfrow=c(2,2), mar = c(3.8, 4, 3, 2))
plot(lm1)
```



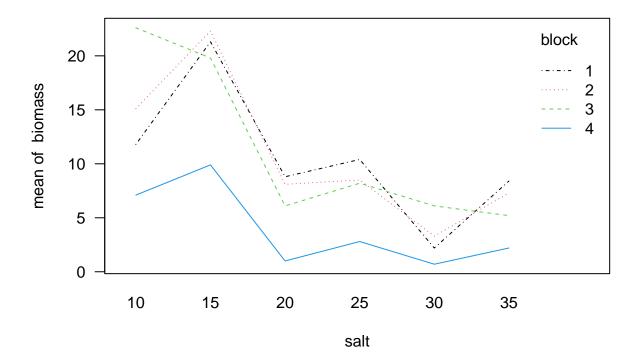
Problem 2

The null hypothesis is that the mean biomass growth is the same across all six levels of salt addition. And the alternative hypothesis is that the mean biomass growth is different across at least one pair of salt addition levels.

To test this hypothesis, I use a fixed complete block design with salt addition as the factor and block as the blocking variable. That is because the assignment of treatments is carried out separately within each block with every treatment included at least once in every block. Besides, the number of the subjects in this situation is 4, which is too small to conduct a random effect analysis.

Moreover, I have made interaction plots to determine whether an interaction term should be included in the model. The lines shown in the plot are nearly parallel to each other, indicating that the model of interaction.

```
# salt <- read.csv("./labs/lab7 linear models with nominal explanatory variables/salt.csv")
salt <- read.csv("salt.csv")
salt$inter <- interaction(salt$salt, salt$block)
with(salt, interaction.plot(salt, block, biomass, col = c(1,
2, 3, 4), las = 1, cex = 0.9))</pre>
```



Thus, I construct the model as follows:

```
lm2 <- lm(biomass ~ factor(salt) + factor(block), data = salt)
TukeyHSD(lm2)</pre>
```

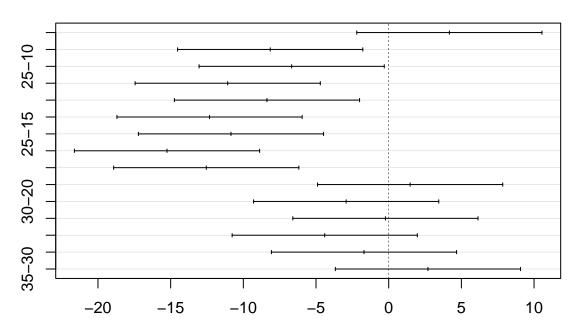
```
## Tukey multiple comparisons of means
## 95% family-wise confidence level
##
## Fit: aov(formula = x)
##
## $'factor(salt)'
## diff lwr upr p adj
## 15-10 4.175 -2.195427 10.5454267 0.3242073
```

```
## 20-10 -8.150 -14.520427 -1.7795733 0.0089392
## 25-10 -6.675 -13.045427 -0.3045733 0.0374481
## 30-10 -11.075 -17.445427 -4.7045733 0.0005436
## 35-10 -8.375 -14.745427 -2.0045733 0.0071740
## 20-15 -12.325 -18.695427 -5.9545733 0.0001755
## 25-15 -10.850 -17.220427 -4.4795733 0.0006697
## 30-15 -15.250 -21.620427 -8.8795733 0.0000151
## 35-15 -12.550 -18.920427 -6.1795733 0.0001439
## 25-20
         1.475 -4.895427 7.8454267 0.9715211
## 30-20 -2.925 -9.295427 3.4454267 0.6742374
## 35-20 -0.225 -6.595427 6.1454267 0.9999965
## 30-25 -4.400 -10.770427
                           1.9704267 0.2749136
## 35-25 -1.700 -8.070427 4.6704267 0.9487968
## 35-30
         2.700 -3.670427 9.0704267 0.7393906
##
## $'factor(block)'
##
            diff
                        lwr
                                  upr
                                          p adj
## 2-1 0.2833333 -4.330839
                            4.897506 0.9979405
## 3-1 0.8500000 -3.764172 5.464172 0.9501884
## 4-1 -6.5333333 -11.147506 -1.919161 0.0048516
## 3-2 0.5666667 -4.047506 5.180839 0.9841982
## 4-2 -6.8166667 -11.430839 -2.202494 0.0034273
## 4-3 -7.3833333 -11.997506 -2.769161 0.0017187
```

summary(1m2)

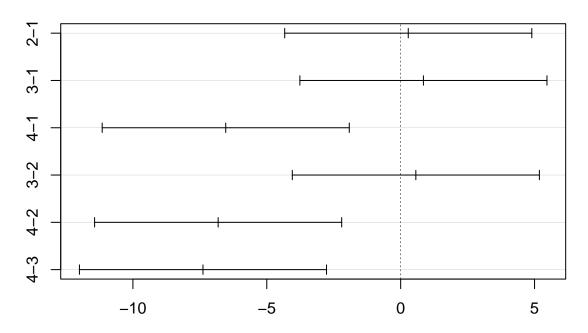
```
##
## Call:
## lm(formula = biomass ~ factor(salt) + factor(block), data = salt)
##
## Residuals:
##
      Min
               1Q Median
                               3Q
                                      Max
  -3.7000 -1.5729 0.0375 1.4812 6.2500
##
## Coefficients:
##
                 Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                  15.5000
                            1.6981
                                      9.128 1.64e-07 ***
## factor(salt)15
                  4.1750
                              1.9608
                                       2.129 0.050205 .
## factor(salt)20 -8.1500
                              1.9608 -4.157 0.000844 ***
## factor(salt)25 -6.6750
                              1.9608 -3.404 0.003923 **
## factor(salt)30 -11.0750
                              1.9608 -5.648 4.63e-05 ***
## factor(salt)35 -8.3750
                              1.9608 -4.271 0.000669 ***
## factor(block)2
                  0.2833
                            1.6009
                                       0.177 0.861893
## factor(block)3
                   0.8500
                             1.6009
                                       0.531 0.603237
                              1.6009 -4.081 0.000984 ***
## factor(block)4 -6.5333
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
## Residual standard error: 2.773 on 15 degrees of freedom
## Multiple R-squared: 0.8862, Adjusted R-squared: 0.8255
## F-statistic: 14.6 on 8 and 15 DF, p-value: 8.593e-06
```

95% family-wise confidence level



Differences in mean levels of factor(salt)

95% family-wise confidence level



Differences in mean levels of factor(block)

```
ad <- read.csv("Sales.csv", header = T)
ad <- with(ad, data.frame(sales, city = factor(city), campaign = factor(campaign),
time = factor(time)))

mod.sales <- lm(sales ~ campaign * time, data = ad)
summary(mod.sales)</pre>
```

```
##
## Call:
## lm(formula = sales ~ campaign * time, data = ad)
##
## Residuals:
              1Q Median
##
      Min
                            ЗQ
                   26.6 241.9
## -378.4 -199.0
                                317.8
##
## Coefficients:
                   Estimate Std. Error t value Pr(>|t|)
##
## (Intercept)
                      718.6
                                 123.8
                                        5.804 5.53e-06 ***
                                 175.1 -0.802
## campaign2
                     -140.4
                                                  0.430
## time2
                       85.6
                                 175.1
                                        0.489
                                                  0.629
## time3
                      -23.2
                                 175.1
                                        -0.133
                                                  0.896
                      -10.4
## campaign2:time2
                                                  0.967
                                 247.6 -0.042
## campaign2:time3
                      -17.6
                                 247.6 -0.071
                                                  0.944
## ---
## Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
```

```
##
## Residual standard error: 276.8 on 24 degrees of freedom
## Multiple R-squared: 0.1135, Adjusted R-squared: -0.07113
## F-statistic: 0.6148 on 5 and 24 DF, p-value: 0.6896
require(lme4)
## Loading required package: lme4
##
## Attaching package: 'lme4'
## The following object is masked from 'package:mosaic':
##
       factorize
require(lmerTest)
## Loading required package: lmerTest
## Warning: package 'lmerTest' was built under R version 4.2.3
## Attaching package: 'lmerTest'
## The following object is masked from 'package:lme4':
##
##
       lmer
## The following object is masked from 'package:mosaic':
##
##
       rand
## The following object is masked from 'package:stats':
##
##
       step
with(ad, tapply(sales, list(time), mean))
             2
       1
## 648.4 728.8 616.4
with(ad, tapply(sales, list(campaign), mean))
          1
## 739.4000 589.6667
```

```
mod.rep <- lmer(sales ~ campaign + time + (1 | city/campaign),</pre>
data = ad)
summary(mod.rep)
## Linear mixed model fit by REML. t-tests use Satterthwaite's method [
## lmerModLmerTest]
## Formula: sales ~ campaign + time + (1 | city/campaign)
##
     Data: ad
## REML criterion at convergence: 286.4
## Scaled residuals:
       Min
               1Q
                     Median
                                   30
                                           Max
## -1.72966 -0.41252 0.08216 0.58379 1.88255
##
## Random effects:
## Groups
                             Variance Std.Dev.
                 Name
## campaign:city (Intercept) 75045.0 273.94
                 (Intercept) 1245.7
                                       35.29
## Residual
                               339.9
                                       18.44
## Number of obs: 30, groups: campaign:city, 10; city, 10
##
## Fixed effects:
              Estimate Std. Error
                                        df t value Pr(>|t|)
## (Intercept) 723.267
                          123.707
                                     8.024 5.847 0.00038 ***
```

-149.733

80.400

-32.000

-0.033 0.000

(Intr) cmpgn2 time2

-0.033 0.000 0.500

Correlation of Fixed Effects:

campaign2
time2

campaign2 -0.707

time3

time2

time3

##

##

174.819

8.245

8.245

Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1

8.000 -0.857 0.41663

18.000 9.751 1.32e-08 ***

18.000 -3.881 0.00110 **