

ENV 710

binomial logistic regression
or log-linear models



roadmap

- download: `pheno.csv`
- load packages:
COUNT
DHARMa
boot
vcdExtra
ResourceSelection
ggplot2
pscl

where we are

multivariate linear models



interactions

centering/scaling explanatory
variables

random effects and mixed models



generalized linear models

logistic regression

- binary response variable
 - success or failure (1 or 0)
 - model the probability of success (π)

$$\pi = \beta_0 + \beta_1 X_1$$

- model log odds

$$\text{logit}(\pi) = \log[\pi/(1 - \pi)]$$

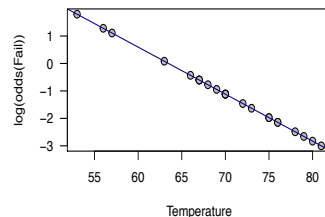
$$\text{logit}(\pi) = \beta_0 + \beta_1 X_k + \dots + \beta_k X_k$$

- calculate probabilities

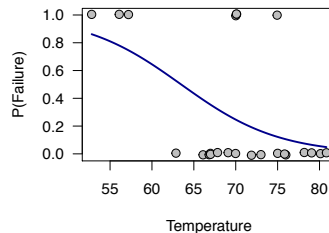
$$\pi = \frac{\exp(\beta_0 + \beta_1 X_k + \dots + \beta_k X_k)}{1 + \exp(\beta_0 + \beta_1 X_k + \dots + \beta_k X_k)}$$

goal is to model the probability of 'success'

transform response to be linearly related to the IV's through the logit transform



back transform to probability using the inverse logit



logistic regression for binomial counts

- **binomial count** = sum of independent binary responses
- $Y \sim \text{binomial}(n, \pi)$, with population proportion, π , and n trials
- observed proportion of 1's is Y/n – binomial proportion
- continuous proportions of amounts cannot be modeled
 - e.g., the proportion of fat that is saturated fat
 - numerator and denominator are not integers, and no n is involved
 - when response variable is a continuous proportion, use ordinary regression methods

- model how population proportion depends on the explanatory variables through a nonlinear link function

$$\text{logit}(\pi) = \beta_0 + \beta_1 x_1 + \beta_2 x_2$$

- interpretation of model is same as for binary logistic regression, which is a special case of the binomial model in which all n_i 's are 1

e.g., Titanic

	survive	cases	age	sex	class
1	1	1	0	0	1
2	13	13	0	0	2
3	14	31	0	0	3
4	5	5	0	1	1
5	11	11	0	1	2
6	13	48	0	1	3
7	140	144	1	0	1
8	80	93	1	0	2
9	76	165	1	0	3
10	57	175	1	1	1
11	14	168	1	1	2
12	75	462	1	1	3

binomial regression

$Y = x$ out of n trials

	pclass	survived	sex	age	surv
1	1st	survived	female	29.00	1
2	1st	survived	male	0.91	1
3	1st	died	female	2.00	0
4	1st	died	male	30.00	0
5	1st	died	female	25.00	0
6	1st	survived	male	48.00	1

logistic regression

$Y = 1$ or 0



I – Phenology

Do minimum nighttime temperature and height of trees influence proportion of trees that flower?

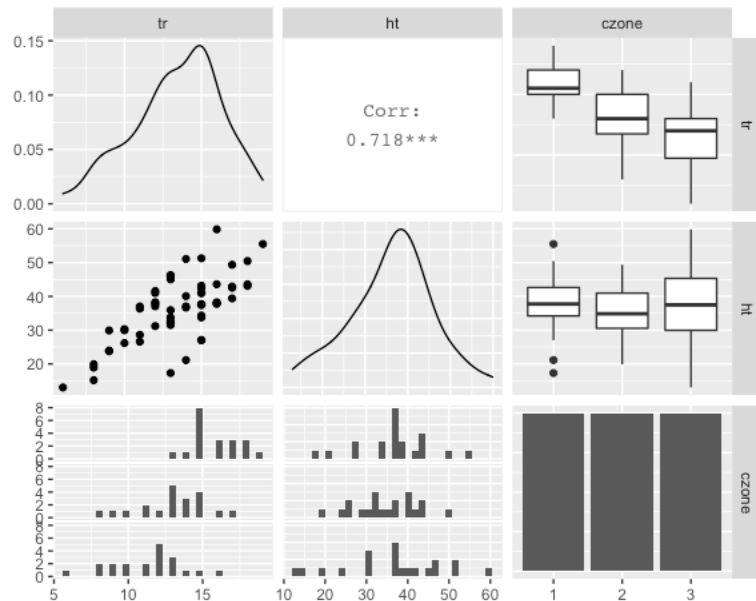
- 20 samples of 20 trees surveyed in each zone
- `czone`: 3 climate zones (hi, med, low nighttime temps)
- `height`: average tree height of each sample of trees
- `n`: number of trees sampled



I – Phenology

Do minimum nighttime temperature and height of trees influence proportion of trees that flower?

	tr	ht	czone	n
1	13	18.09	1	20
2	13	25.31	1	20
3	15	26.39	1	20
4	15	29.98	1	20
5	15	30.61	1	20
6	15	33.27	1	20
...



```
glm(cbind(tr, n - tr) ~ factor(czone) * ht,  
     family = binomial, data = trees)
```

```
b0 <- glm(cbind(tr, n-tr)~factor(czone)*ht, family=binomial,
          data=trees)
```

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)	
(Intercept)	-0.326077	0.515127	-0.633	0.52673	
factor(czone)2	-1.403413	0.733661	-1.913	0.05576	.
factor(czone)3	-0.937770	0.616617	-1.521	0.12830	
ht	0.047032	0.014231	3.305	0.00095	***
factor(czone)2:ht	0.021752	0.020764	1.048	0.29482	
factor(czone)3:ht	-0.004873	0.016824	-0.290	0.77208	

Null deviance: 114.2732 on 59 degrees of freedom
 Residual deviance: 5.8137 on 54 degrees of freedom
 AIC: 211.19

```
b1<-update(b0, .~.-factor(czone):ht)
```

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)	
(Intercept)	-0.38963	0.27057	-1.440	0.15	
factor(czone)2	-0.66316	0.16544	-4.008	6.11e-05	***
factor(czone)3	-1.11444	0.16438	-6.779	1.21e-11	***
ht	0.04884	0.00683	7.152	8.58e-13	***

Null deviance: 114.2732 on 59 degrees of freedom
 Residual deviance: 8.1554 on 56 degrees of freedom
 AIC: 209.53

I – Phenology

Compare models

```
lrtest(b0, b1)
```

Likelihood ratio test

Model 1: cbind(tr, n - tr) ~ factor(czone) * ht

Model 2: cbind(tr, n - tr) ~ factor(czone) + ht

	#Df	LogLik	Df	Chisq	Pr(>Chisq)
1	6	-99.595			
2	4	-100.766	-2	2.3417	0.3101

Check model fit

```
pchisq(b1$deviance, b1$df.residual, lower.tail = F)
1
```

Check for overdispersion

```
d2 = sum(residuals(b1,"pearson")^2)
disp = d2/df.residual(b1)
phi = sqrt(disp)
```

```
[1] 0.3801822
```

```
DHARMA::testDispersion(b1)
```

DHARMA nonparametric dispersion test via sd of residuals fitted vs. simulated

```
data: simulationOutput
dispersion = 0.67906, p-value = 0.024
alternative hypothesis: two.sided
```

I – Phenology

Coefficients as log-odds

```
coef(b1)
```

Intercept)	factor(czone)2	factor(czone)3	ht
-0.38962645	-0.66315645	-1.11443652	0.04884507

Coefficients as odds

```
exp(coef(b1))
```

(Intercept)	factor(czone)2	factor(czone)3	ht
0.6773098	0.5152225	0.3281001	1.0500577

Coefficients as probability

```
inv.logit(coef(b1))
```

(Intercept)	factor(czone)2	factor(czone)3	ht
0.4038072	0.3400309	0.2470447	0.5122088

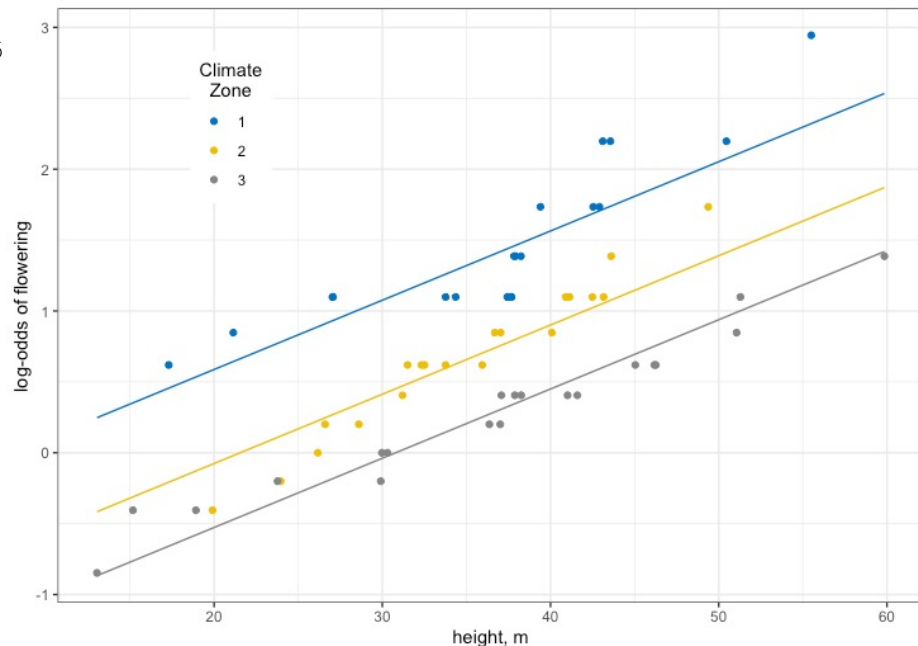
I – Phenology

Coefficients as log-odds

`coef(b1)`

Intercept)	factor(czone) 2	factor(czone) 3	ht
-0.38962645	-0.66315645	-1.11443652	0.048845

- Intercept is log-odds of flowering at czone 1 when tree height is 0
- czone 2 is the difference in the log-odds of flowering for czone 2 compared to czone 1: log-odds of flowering decreases by -0.66 from czone 1 to 2.
- log-odds of flowering in czone 2 is $-0.389 + 0.663 = -1.052$.
- ht is the change in log-odds of flowering for every additional meter of tree height.



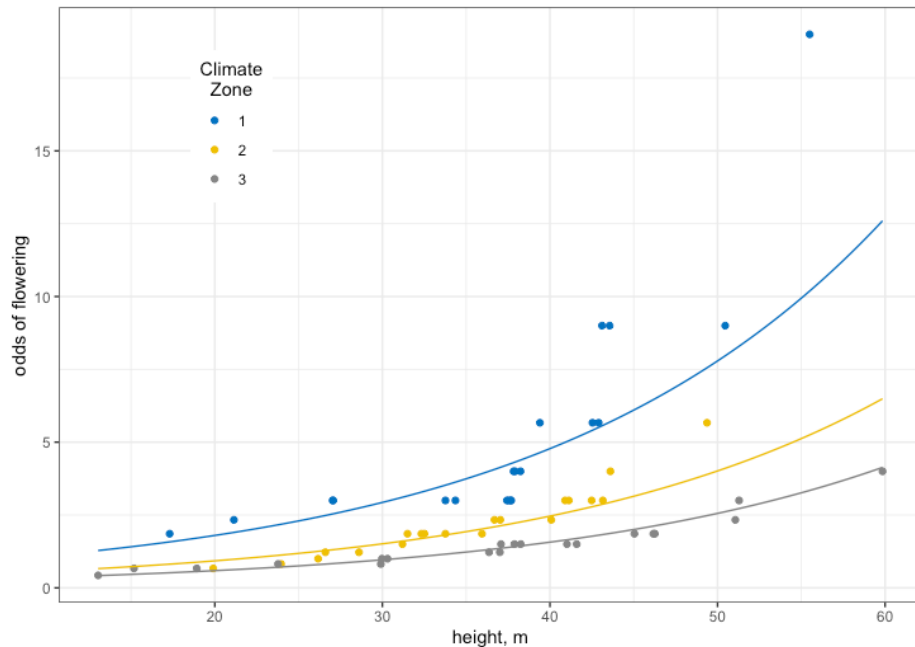
I – Phenology

Coefficients as odds

```
exp(coef(b1))
```

(Intercept)	factor(czone)2	factor(czone)3	ht
0.6773098	0.5152225	0.3281001	1.0500577

- Intercept is the odds of flowering at czone 1 when tree height is 0
- czone 2 is the odds ratio of flowering between czone 2 and czone 1
- mean odds of flowering in czone 2 is 0.349
[$\exp(-0.38962645 + -0.66315645)$]
- ht, the odds of flowering increase by a factor of 1.05 for every meter, or approximately 5% increase in odds of flowering for each additional unit of tree height



I – Phenology

Coefficients as probability

```
inv.logit(coef(b1))
```

(Intercept)	factor(czone)2	factor(czone)3	ht
0.4038072	0.3400309	0.2470447	0.5122088

- log-odds coefficients can be converted into probabilities with the `inv.logit()`, but can't be directly interpreted because the relationship is not linear
- best to evaluate probability at specific values of predictors

$$p = \frac{1}{(1 + 1/e^\beta)}$$

Probability of tree flowering in climate zone 2 at 15, 35, and 50 m of height

```
cfs <- coef(b1)
inv.logit(cfs[1] + cfs[2] + cfs[3]*0 +
          cfs[4]*c(15, 35, 50))
```

```
[1] 0.4206497 0.6585400 0.8005077
```

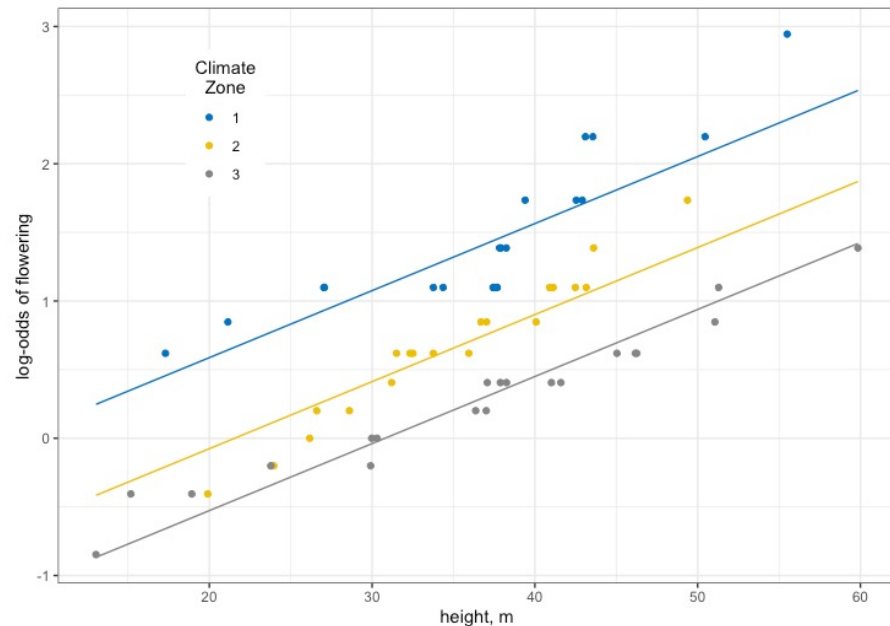
I – Phenology

*“I’ll do algebra, I’ll do trigonometry, I’ll even do statistics...
But graphing is where I draw the line.”*

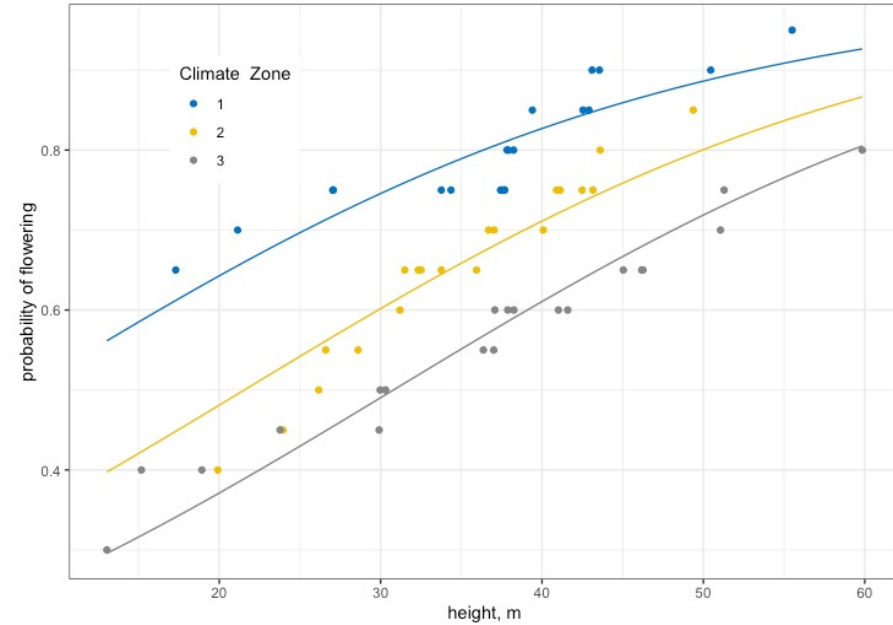
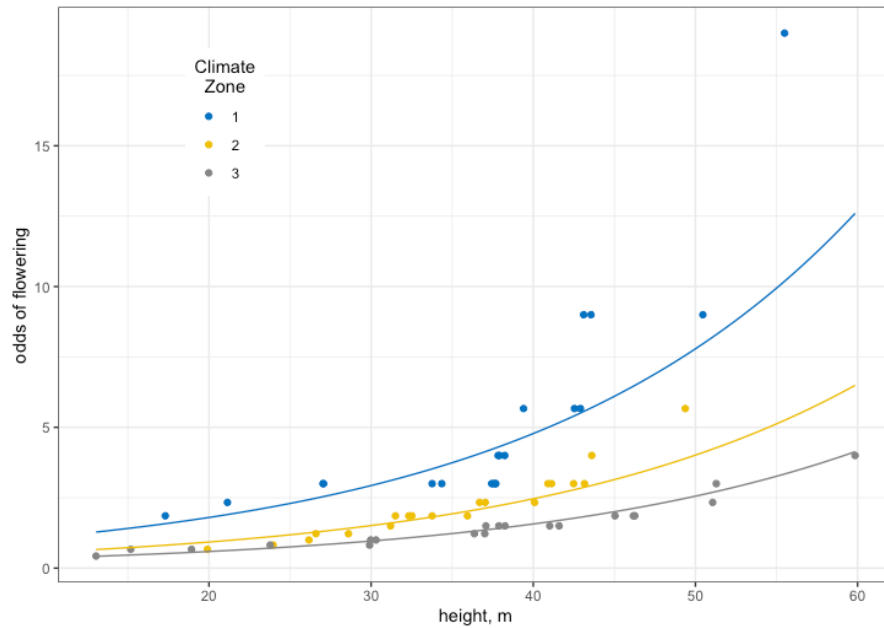
```
jcoPalette <- c("#0073C2FF", "#EFC000FF", "#868686FF", "#CD534CFF",  
"#7AA6DCFF", "#003C67FF", "#8F7700FF", "#3B3B3BFF", "#A73030FF",  
"#4A6990FF")
```

```
b <- coef(b1)
```

```
ggplot(data = trees, aes(x = ht, y = log(tr/(n-tr)))) +  
  geom_point(aes(color = czone)) +  
  stat_function(fun = function(x){b[1]+b[4]*x}, colour =  
jcoPalette[1]) +  
  stat_function(fun = function(x){b[1]+b[2]+b[4]*x}, colour =  
jcoPalette[2]) +  
  stat_function(fun = function(x){b[1]+b[3]+b[4]*x}, colour =  
jcoPalette[3]) +  
  theme_bw() + scale_colour_manual("Climate\n Zone",  
values=jcoPalette) +  
  theme(legend.position = c(0.2, 0.8)) +  
  labs(x = "height, m", y = "log-odds of flowering")
```



I – Phenology



I – Phenology

```
## plot as odds with ggplot
ggplot(data = trees, aes(x = ht, y = tr/(n-tr))) +
  geom_point(aes(color = czone)) +
    stat_function(fun = function(x){exp(b[1]+b[4]*x)},
      colour = jcoPalette[1]) +
    stat_function(fun = function(x){exp(b[1]+b[2]+b[4]*x)},
      colour = jcoPalette[2]) +
    stat_function(fun = function(x){exp(b[1]+b[3]+b[4]*x)},
      colour = jcoPalette[3]) +
  theme_bw() + scale_colour_manual("Climate\n Zone", values=jcoPalette) +
  theme(legend.position = c(0.2, 0.8)) +
  labs(x = "height, m", y = "odds of flowering")

## plot as probabilities with ggplot
ggplot(data = trees, aes(x = ht, y = tr/n)) +
  geom_point(aes(color = czone)) +
    stat_function(fun = function(x){inv.logit(b[1]+b[4]*x)},
      colour = jcoPalette[1]) +
    stat_function(fun = function(x){inv.logit(b[1]+b[2]+b[4]*x)},
      colour = jcoPalette[2]) +
    stat_function(fun = function(x){inv.logit(b[1]+b[3]+b[4]*x)},
      colour = jcoPalette[3]) +
  theme_bw() + scale_colour_manual("Climate\n Zone", values=jcoPalette) +
  theme(legend.position = c(0.2, 0.8)) +
  labs(x = "height, m", y = "probability of flowering")
```

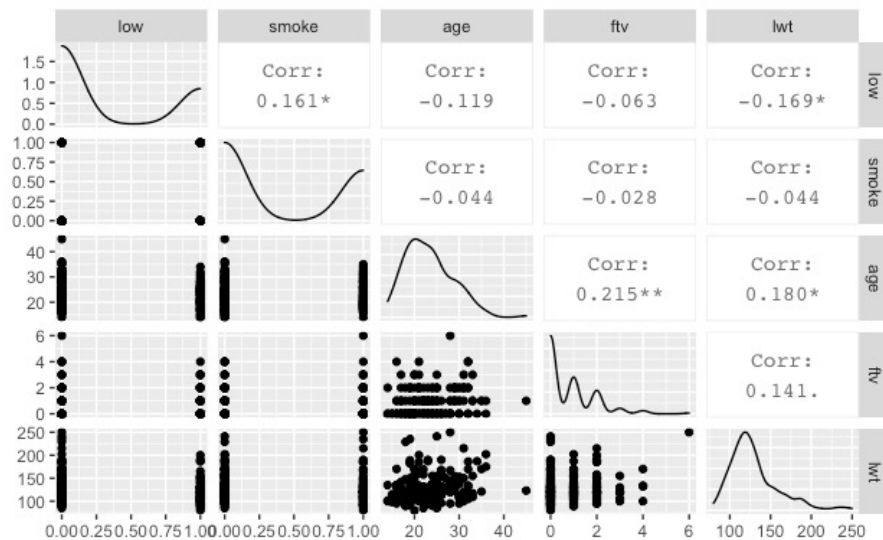

2 – Low birth weight

What variables result in low birth weight?

- low: 1=low birthweight baby; 0=normal weight
- smoke: 1=history of mother smoking; 0=mother nonsmoker
- age: age of mother: 14-45
- lwt: weight (lbs) at last menstrual period: 80-250 lbs
- ftv: number of physician visits in 1st trimester: 0-6

Build model, reduce it, check its goodness of fit, and interpret the results

```
require(COUNT)
data(lbw)
lbw$lwt <- as.numeric(lbw$lwt)
ggpairs(lbw, columns = c("low", "smoke", "age",
                        "ftv", "lwt"))
```



2 – Low birth weight

```
lr1 <- glm(low ~ factor(smoke) + age + ftv + lwt,  
           family = "binomial", data = lbw)  
lr2 <- update(lr1, .~-ftv)  
lr3 <- update(lr2, .~-age)  
summary(lr3)
```

```
Call:  
glm(formula = low ~ factor(smoke) + lwt,  
     family = "binomial", data = lbw)
```

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	0.619201	0.795870	0.778	0.4366
factor(smoke)1	0.676579	0.324685	2.084	0.0372 *
lwt	-0.013301	0.006088	-2.185	0.0289 *

```
Null deviance: 234.67 on 188 degrees of freedom  
Residual deviance: 224.36 on 186 degrees of freedom  
AIC: 230.36
```

```
lrtest(lr1, lr2, lr3)  
Likelihood ratio test
```

```
Model 1: low ~ factor(smoke) + age + ftv + lwt  
Model 2: low ~ factor(smoke) + age + lwt  
Model 3: low ~ factor(smoke) + lwt  
#Df  LogLik Df  Chisq Pr(>Chisq)  
1    5 -111.42  
2    4 -111.45 -1  0.0645      0.7995  
3    3 -112.18 -1  1.4634      0.2264
```

```
pchisq(lr3$deviance, lr3$df.residual,  
lower.tail = F)  
[1] 0.02868237
```

```
> pR2(lgr4)  
fitting null model for pseudo-r2  
      llh      llhNull      G2  
-112.17946472 -117.33599810  10.31306675  
      McFadden      r2ML      r2CU  
0.04394673      0.05310445      0.07468003
```

2 – Low birth weight

```
lr1 <- glm(low ~ factor(smoke) + age + ftv + lwt,  
           family = "binomial", data = lbw)  
lr2 <- update(lr1, .~-ftv)  
lr3 <- update(lr2, .~-age)  
summary(lr3)
```

Call:

```
glm(formula = low ~ factor(smoke) + lwt,  
     family = "binomial", data = lbw)
```

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	0.619201	0.795870	0.778	0.4366
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```
Null deviance: 234.67 on 188 degrees of freedom  
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```

How does smoking affect the odds of having a low birthweight baby?

What is the effect of low mother weight on the odds of having a low birthweight baby?

What is the probability of having a low birthweight baby for a smoker weighing in the lowest 25% of women?

2 – Low birth weight

```
lr1 <- glm(low ~ factor(smoke) + age + ftv + lwt,  
           family = "binomial", data = lbw)  
lr2 <- update(lr1, .~-ftv)  
lr3 <- update(lr2, .~-age)  
summary(lr3)
```

```
Call:  
glm(formula = low ~ factor(smoke) + lwt,  
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```

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
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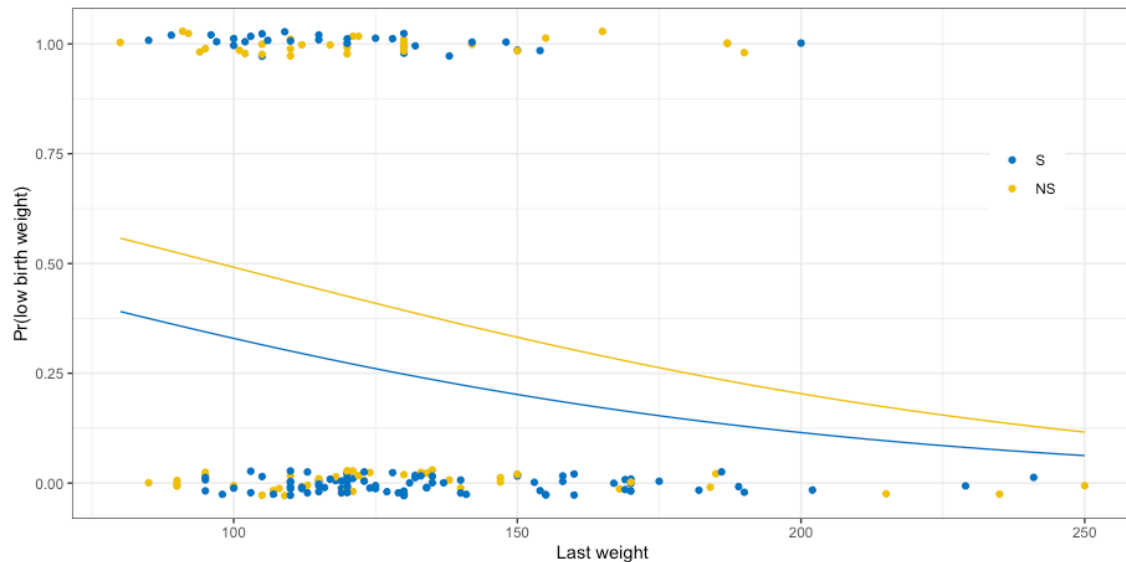
```
Null deviance: 234.67 on 188 degrees of freedom  
Residual deviance: 224.36 on 186 degrees of freedom  
AIC: 230.36
```

Smoking increases the odds of having a low birthweight baby by 1.97 times or 97% ($z = 2.08$, $p = 0.037$).

Every additional pound of weight decreases the odds of a woman having a low birthweight child by 1.3% ($z = -2.19$, $p = 0.029$).

The probability of having a low birthweight child for a smoker in the 25th percentile of weights is 45.8%.

2 – Low birth weight



```
ggplot(lbw, aes(x = as.numeric(lwt), y = as.numeric(low),  
  color=as.factor(smoke))) +  
  stat_function(fun = function(x){inv.logit(lr3c[1]+lr3c[3]*x)},  
    colour = jcoPalette[1]) +  
  stat_function(fun = function(x){inv.logit(lr3c[1]+lr3c[2] + lr3c[3]*x)},  
    colour = jcoPalette[2]) +  
  geom_point(position=position_jitter(height=0.03, width=0)) +  
  xlab("Last weight") + ylab("Pr(low birth weight)") + theme_bw() +  
  theme(legend.position=c(0.9, 0.7)) +  
  scale_colour_manual("", values=jcoPalette, labels = c("S", "NS"))
```



Questions?