Regression & Analysis of Variance

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Overview

We will analyze data from 3 distinct datasets **Abalone**, **Cigs**, **BirthWeight**, we will make hypothesis and present observations & interpret results from findings from our analysis

Abalone:

Hypothesis: I believe that a simple linear regression model with normal error assumption is appropriate to describe the relationship between the height of abalones and their ages, and particularly, that a larger height is associated with an older age, we will use data from abalone.csv to test this hypotheis

```
# importing data
file1 <- "http://www.math.mcgill.ca/yyang/regression/data/abalone.csv"
abalone <- read.csv(file1, header = TRUE)

# functions for summary statistics

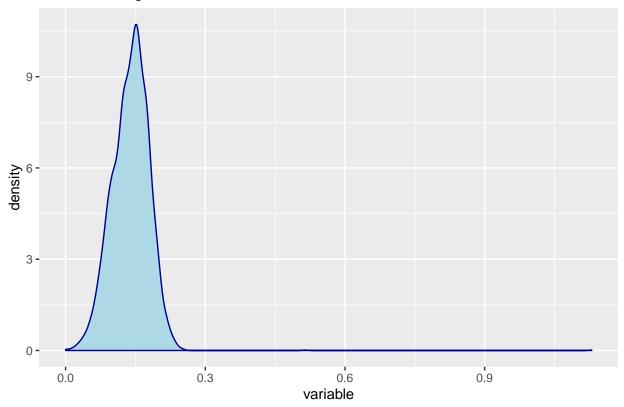
Summary_Table<-function(data,variable)
{
    data %>% summarise(Avg = mean(variable),
        Med = median(variable),
        Q25 = quantile(variable,0.25), Q75 = quantile(variable,0.75),
        StD = sd(variable), Var=var(variable), Min=min(variable),
        Max=max(variable))%>%kable()
}

Plot_Distribution<-function(data,variable,title="")
{
    ggplot(data, aes(x=variable))+geom_density(color="darkblue", fill="lightblue")+ggtitle(title)}
}</pre>
```

Univariate Analysis: Height

```
Plot_Distribution(abalone,abalone #Height, "Abalone Height Distribution")
```

Abalone Height Distribution



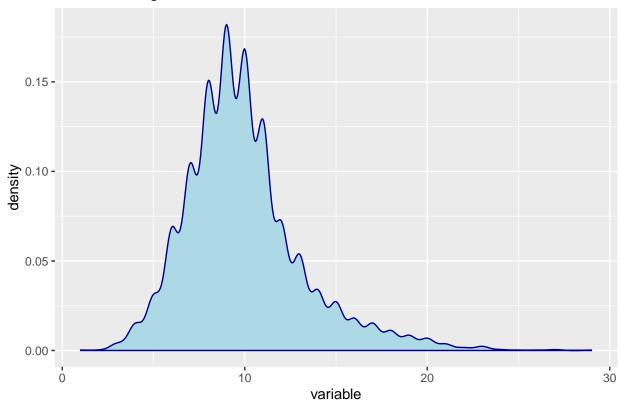
Summary_Table(abalone,abalone\$Height)

Avg	Med	Q25	Q75	StD	Var	Min	Max
0.1395164	0.14	0.115	0.165	0.0418271	0.0017495	0	1.13

Univariate Analysis: Ring

Plot_Distribution(abalone,abalone\$Rings,"Abalone Ring Distribution")

Abalone Ring Distribution

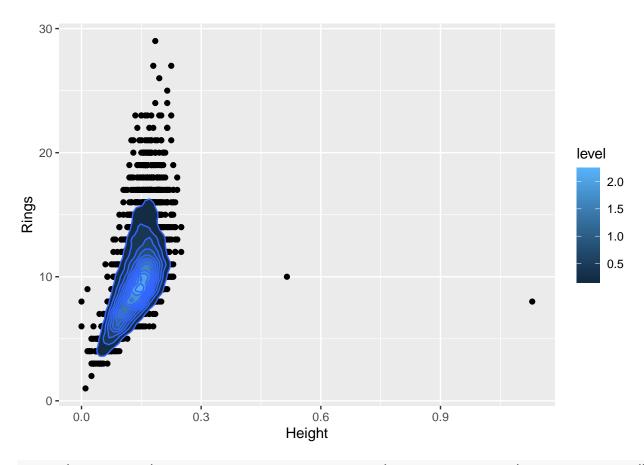


Summary_Table(abalone,abalone\$Rings)

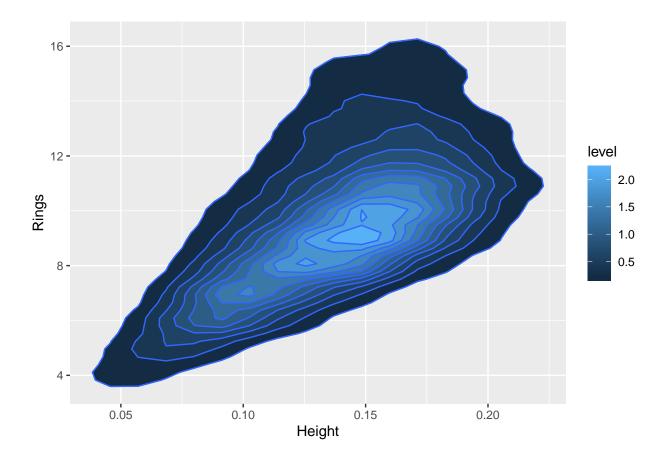
Avg	Med	Q25	Q75	StD	Var	Min	Max
9.933685	9	8	11	3.224169	10.39527	1	29

Bivariate Analysis: Height vs Rings

ggplot(abalone,aes(x=Height,y=Rings))+geom_point()+stat_density_2d(aes(fill = ..level..), geom = "polyg"



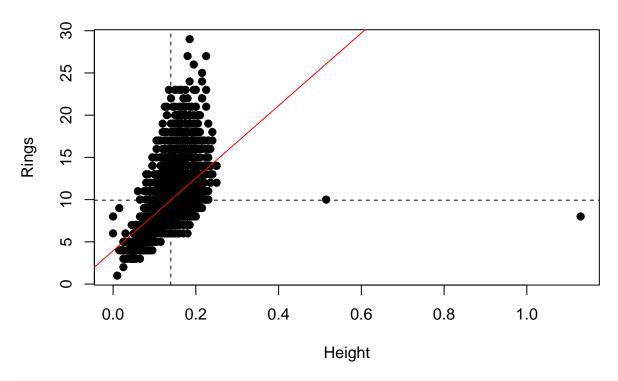
ggplot(abalone,aes(x=Height,y=Rings,fill = ..level..), geom = "polygon")+geom_density_2d()+stat_density



Fitting Linear Model

```
plot(abalone$Height,abalone$Rings,pch=19,xlab='Height',ylab='Rings')
abline(v=mean(abalone$Height),h=mean(abalone$Rings),lty=2)
fit.RP<-lm(abalone$Rings~abalone$Height)
title('Line of best fit for Abalone Data')
abline(coef(fit.RP),col='red')</pre>
```

Line of best fit for Abalone Data

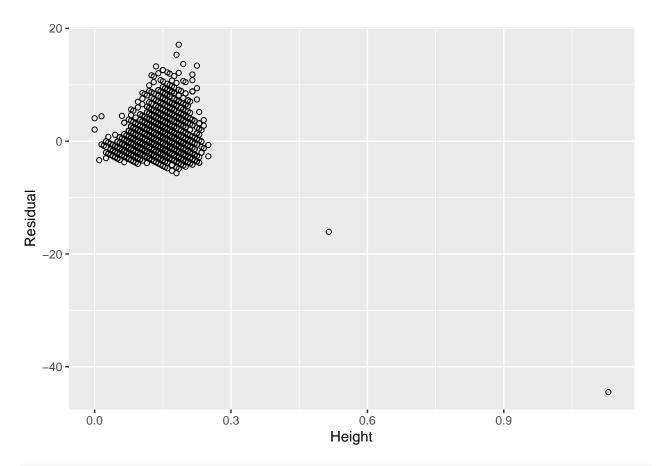


summary(fit.RP)

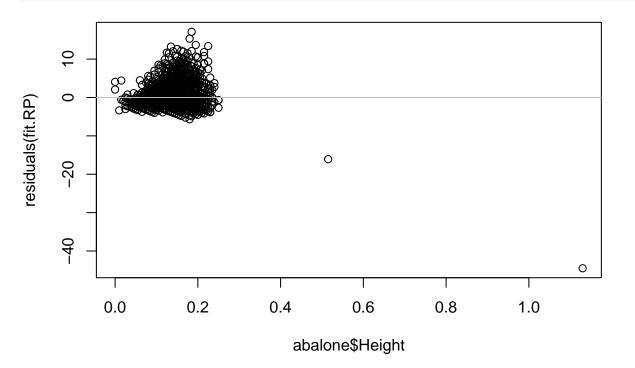
```
##
## lm(formula = abalone$Rings ~ abalone$Height)
##
## Residuals:
##
      Min
               1Q Median
                               3Q
                                      Max
## -44.496 -1.657 -0.607
                            0.839 17.112
##
## Coefficients:
                 Estimate Std. Error t value Pr(>|t|)
                                       27.30
                   3.9385
                              0.1443
                                               <2e-16 ***
## (Intercept)
## abalone$Height 42.9714
                              0.9904
                                       43.39
                                               <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 2.677 on 4175 degrees of freedom
## Multiple R-squared: 0.3108, Adjusted R-squared: 0.3106
## F-statistic: 1882 on 1 and 4175 DF, p-value: < 2.2e-16
```

Model Aquecuacy Checking (Residual Analysis) & Diagnostic

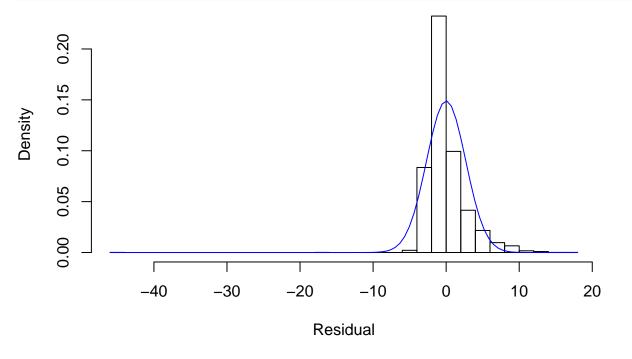
```
ggplot(data = data.frame(x = abalone$Height, y = residuals(fit.RP)), aes(x = x, y = y))+geom_point(shap
```



plot(abalone\$Height,residuals(fit.RP))
abline(h=0,col="gray")

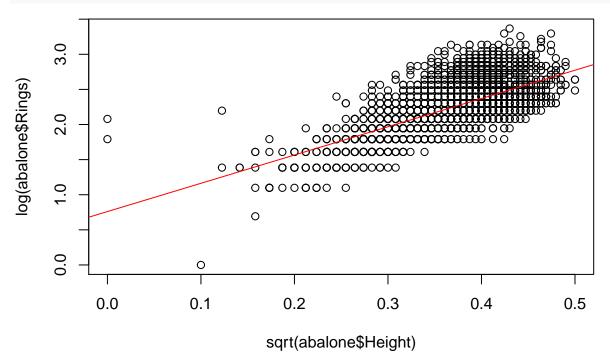


```
hist(residuals(fit.RP),breaks=40,freq=FALSE,xlab="Residual",main="")
curve(dnorm(x,mean=0,sd=sd(residuals(fit.RP))),add=TRUE,col="blue")
```



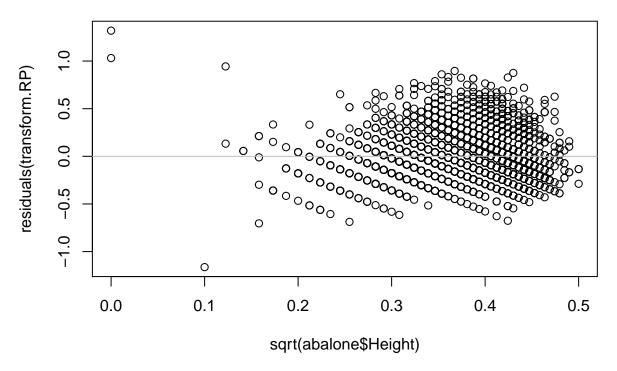
Removing Outliers & Data Transformation

```
abalone<-filter(abalone, Height<0.4)
plot(y=log(abalone$Rings), x=sqrt(abalone$Height))
transform.RP<-lm(log(abalone$Rings)~sqrt(abalone$Height))
abline(coef(transform.RP),col='red')
```

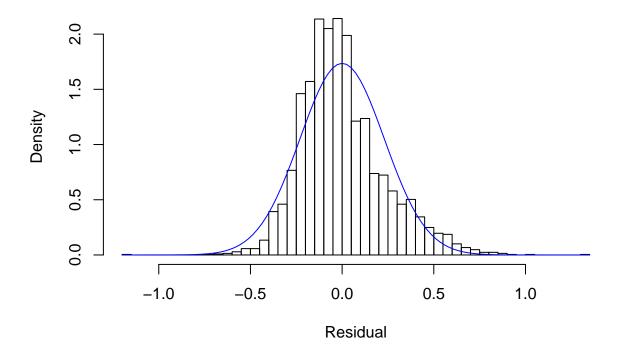


```
summary(transform.RP)
```

```
##
## Call:
## lm(formula = log(abalone$Rings) ~ sqrt(abalone$Height))
##
## Residuals:
##
        Min
                  1Q
                       Median
                                    3Q
                                            Max
   -1.16301 -0.14958 -0.03369
                              0.11898
##
## Coefficients:
##
                        Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                         0.76049
                                    0.02407
                                              31.59
                                                      <2e-16 ***
## sqrt(abalone$Height)
                         4.02516
                                    0.06452
                                              62.38
                                                      <2e-16 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 0.2299 on 4173 degrees of freedom
## Multiple R-squared: 0.4826, Adjusted R-squared: 0.4824
## F-statistic: 3892 on 1 and 4173 DF, p-value: < 2.2e-16
plot(sqrt(abalone$Height),residuals(transform.RP))
abline(h=0,col='gray')
```

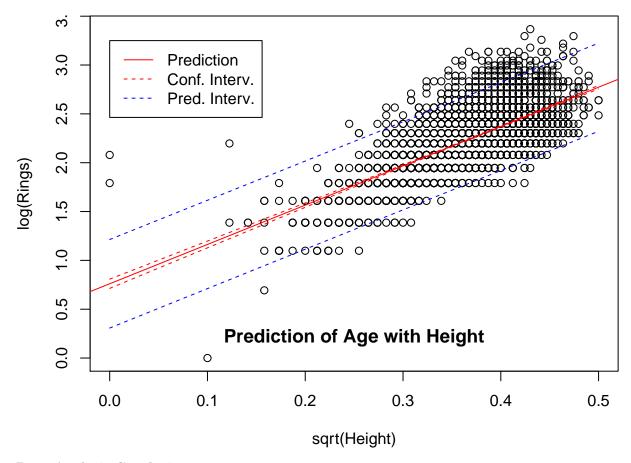


hist(residuals(transform.RP), breaks=40, freq=FALSE, xlab="Residual", main="")
curve(dnorm(x, mean=0, sd=sd(residuals(transform.RP))), add=TRUE, col="blue")



Building Confidence & Prediction Interval

```
par(mar=c(4,4,0,1))
x<-sqrt(abalone$Height)</pre>
y<-log(abalone$Rings)
fit.RP < -lm(y \sim x)
xnew < -seq(0,0.5,by=0.01)
\#Confidence\ interval
ynew.interval<-predict(fit.RP,newdata=data.frame(x=xnew),interval='confidence')</pre>
plot(x,y,xlab='sqrt(Height)',ylab='log(Rings)')
abline(coef(fit.RP),col='red')
lines(xnew,ynew.interval[,2],lty=2,col='red')
lines(xnew,ynew.interval[,3],lty=2,col='red')
#Prediction interval
yonew.interval<-predict(fit.RP,newdata=data.frame(x=xnew),interval='prediction')</pre>
lines(xnew,yonew.interval[,2],lty=2,col='blue')
lines(xnew,yonew.interval[,3],lty=2,col='blue')
legend(0,3.25,c('Prediction','Conf. Interv.','Pred. Interv.'),col=c('red','red','blue'),lty=c(1,2,2))
title('Prediction of Age with Height',line=-17)
```



Data Analysis Conclusions:

After fitting a standard linear model and analyzing residuals, it was clear that there was not a full linear relationship between the two variable as there was sign of non-constant variance which slightly violated linear model assumptions we had made about the residuals being mean 0 with constant variance our residuals tended to have a right skew. Once log transformation of the response variable was made as well as squared root transformation of the predictor, our model tended to perform a lot better

Cigs:

Hypothesis:

We will investigate and build a model for the relationship between response CO2 produced for cigarette and predictors of Tar, Weight & Nicotine. I believe a Multiple Linear Regression Model will help explain the relationship and will be an adequate enough model with significant relationships between Tar and Weight

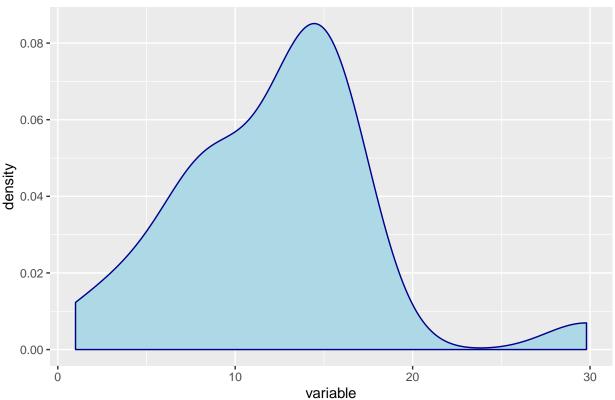
```
# importing data
file1 <- "http://www.math.mcgill.ca/yyang/regression/data/cigs.csv"
cigs <- read.csv(file1, header = TRUE)

CO<-cigs$CO
TAR<-cigs$TAR
NICOTINE<-cigs$NICOTINE
WEIGHT<-cigs$WEIGHT</pre>
```

Univariate Analysis

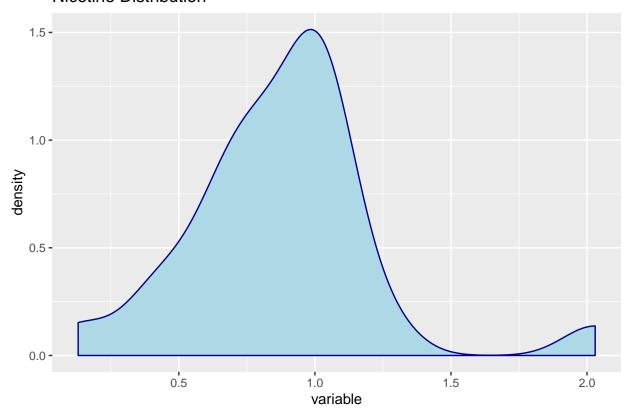
Plot_Distribution(cigs,TAR,"Tar Distribution")

Tar Distribution



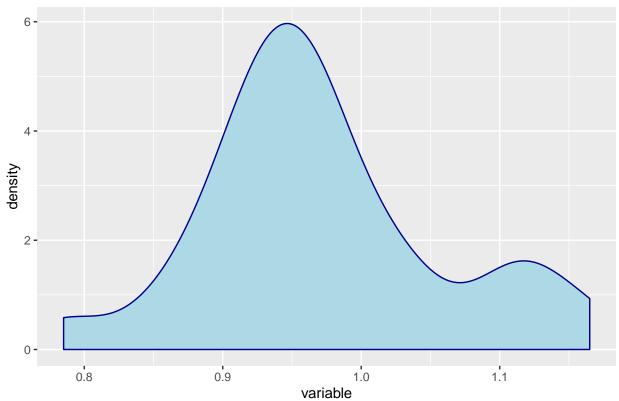
Plot_Distribution(cigs,NICOTINE,"Nicotine Distribution")

Nicotine Distribution



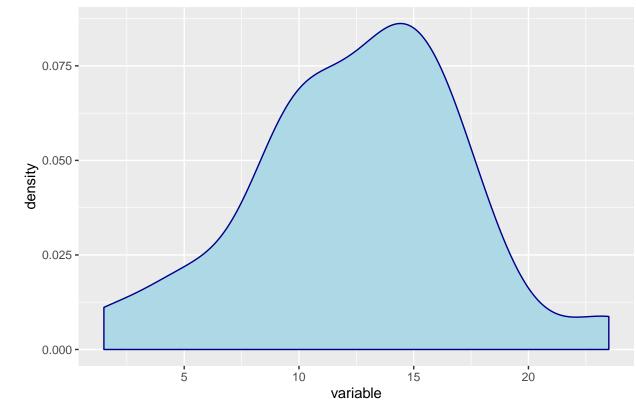
Plot_Distribution(cigs,WEIGHT,"Weight Distribution")

Weight Distribution



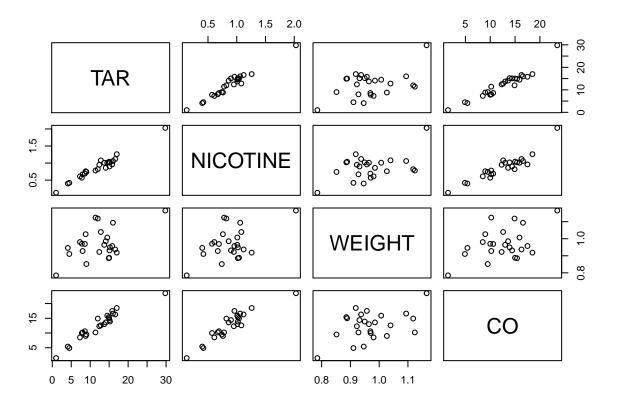
Plot_Distribution(cigs,CO,"CO2 Distribution")





Bivariate Analysis

plot(cigs)



Regression Analysis

```
summary(lm(CO~TAR+NICOTINE+WEIGHT))
```

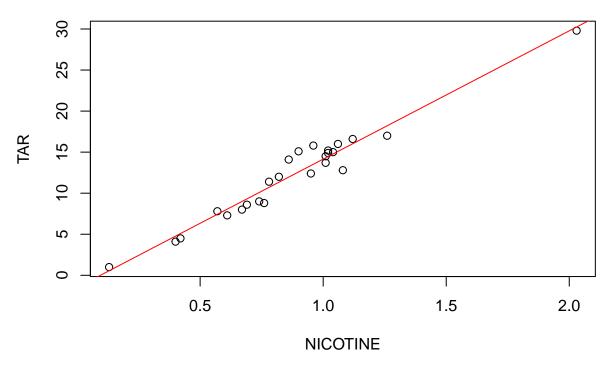
```
##
## Call:
## lm(formula = CO ~ TAR + NICOTINE + WEIGHT)
##
## Residuals:
       Min
                 1Q
                      Median
## -2.89261 -0.78269 0.00428 0.92891 2.45082
## Coefficients:
              Estimate Std. Error t value Pr(>|t|)
## (Intercept)
                3.2022
                           3.4618
                                    0.925 0.365464
                0.9626
                           0.2422
                                    3.974 0.000692 ***
## NICOTINE
                           3.9006 -0.675 0.507234
               -2.6317
## WEIGHT
               -0.1305
                           3.8853 -0.034 0.973527
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 1.446 on 21 degrees of freedom
## Multiple R-squared: 0.9186, Adjusted R-squared: 0.907
## F-statistic: 78.98 on 3 and 21 DF, p-value: 1.329e-11
```

anova(lm(CO~TAR+NICOTINE+WEIGHT))

```
## Analysis of Variance Table
## Response: CO
            Df Sum Sq Mean Sq F value
                                          Pr(>F)
             1 494.28 494.28 236.4843 6.651e-13 ***
## TAR
                                          0.5023
## NICOTINE
                 0.97
                         0.97
                                0.4661
## WEIGHT
             1
                 0.00
                         0.00
                                0.0011
                                          0.9735
## Residuals 21 43.89
                         2.09
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
```

Analyzing Muticollinearity

```
plot(y=TAR,x=NICOTINE)
abline(coef(lm(TAR~NICOTINE)),col='red')
```

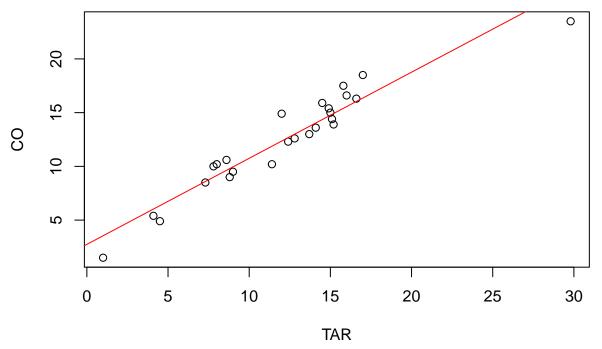


cor(TAR,NICOTINE)

[1] 0.9766076

Picking Model

```
anova(lm(CO~TAR+WEIGHT))
## Analysis of Variance Table
##
## Response: CO
##
                                            Pr(>F)
             Df Sum Sq Mean Sq F value
              1 494.28
                        494.28 242.4892 2.308e-13 ***
## TAR
## WEIGHT
              1
                  0.03
                           0.03
                                  0.0123
                                            0.9127
## Residuals 22
                 44.84
                           2.04
##
                     '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
sigma(lm(CO~TAR+WEIGHT))
## [1] 1.427713
sigma(lm(CO~TAR))
## [1] 1.396721
plot(y=CO,x=TAR)
abline(coef(lm(CO~TAR)),col='red')
```



Data Analysis Conclusions:

After analyzing predictors, conclusions were made about multicollinearity between Tar & Nicotine thus we opted to choose tar as one of the variable in our model, upon fitting a MLR model with tar, weight, nicotine it was clear that the only significant preditor was tar. Analysis on resdisual was done to confirm this finding

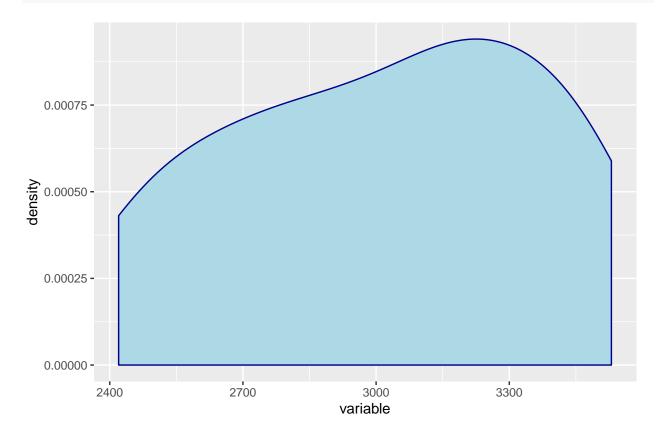
BirthWeight (Smokers vs Non-Smokers):

Hypothesis: I believe that there is a difference in Birth Weight in relation to Gestion time, in between Smokers & Non-Smokers

data.source<-"http://www.math.mcgill.ca/yyang/regression/data/birthsmokers.csv"
birthsmokers<-read.csv(file=data.source)</pre>

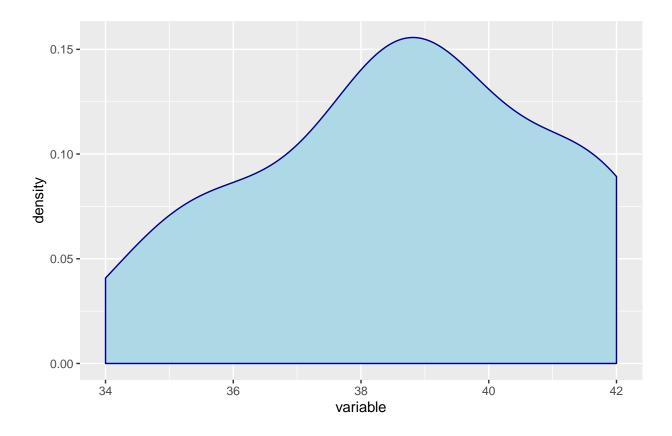
Univariate Analysis: Weight

Plot_Distribution(birthsmokers, birthsmokers\$Wgt)

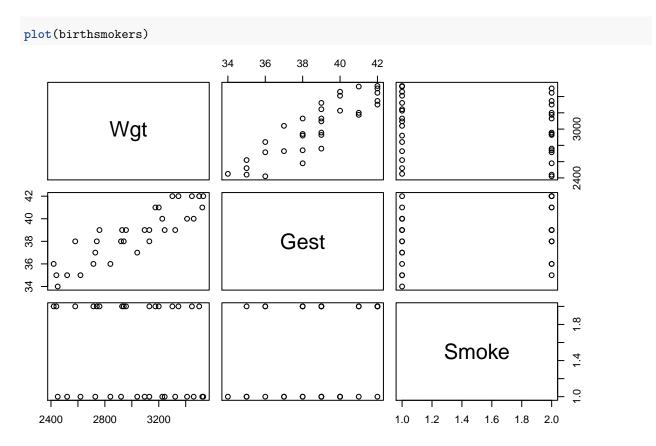


Univariate Analysis: Gestations

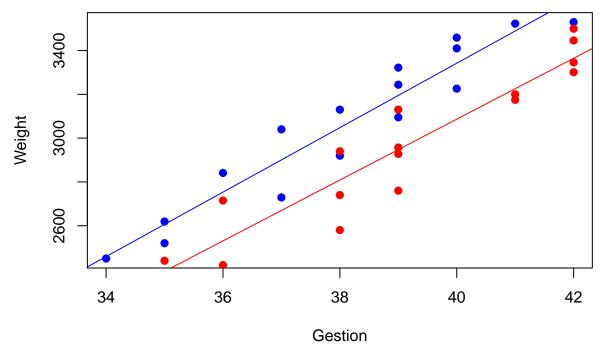
Plot_Distribution(birthsmokers,birthsmokers\$Gest)



Bivariate Analysis



```
plot(x=subset(birthsmokers,Smoke=="no")$Gest,y=subset(birthsmokers,Smoke=="no")$Wgt,col="blue",pch=19,x
points(x=subset(birthsmokers,Smoke=="yes")$Gest,y=subset(birthsmokers,Smoke=="yes")$Wgt,col="red",pch=1
abline(coef(lm(Wgt~Gest,data=subset(birthsmokers,Smoke=="no"))),col='blue')
abline(coef(lm(Wgt~Gest,data=subset(birthsmokers,Smoke=="yes"))),col='red')
```



Regression Analysis

```
summary(lm(Wgt~Gest,data=subset(birthsmokers,Smoke=="no")))
```

```
##
## Call:
## lm(formula = Wgt ~ Gest, data = subset(birthsmokers, Smoke ==
       "no"))
##
##
## Residuals:
      Min
               1Q Median
                               3Q
                                      Max
## -171.52 -101.59
                    23.28
                            83.63 139.48
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -2546.14
                           457.29 -5.568 6.93e-05 ***
                            11.97 12.294 6.85e-09 ***
## Gest
                 147.21
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 106.9 on 14 degrees of freedom
## Multiple R-squared: 0.9152, Adjusted R-squared: 0.9092
## F-statistic: 151.1 on 1 and 14 DF, p-value: 6.852e-09
```

```
summary(lm(Wgt~Gest,data=subset(birthsmokers,Smoke=="yes")))
##
## Call:
## lm(formula = Wgt ~ Gest, data = subset(birthsmokers, Smoke ==
       "yes"))
##
##
## Residuals:
##
      Min
                1Q Median
                                3Q
                                       Max
  -228.53 -64.86 -19.10
                                   184.53
##
                             93.89
##
## Coefficients:
##
              Estimate Std. Error t value Pr(>|t|)
## (Intercept) -2474.56
                           553.97 -4.467 0.000532 ***
## Gest
                 139.03
                             14.11
                                     9.851 1.12e-07 ***
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
## Residual standard error: 126.6 on 14 degrees of freedom
## Multiple R-squared: 0.8739, Adjusted R-squared: 0.8649
## F-statistic: 97.04 on 1 and 14 DF, p-value: 1.125e-07
confint(lm(Wgt~Gest,data=subset(birthsmokers,Smoke=="no")))
##
                    2.5 %
                              97.5 %
## (Intercept) -3526.9338 -1565.3420
## Gest
                 121.5251
                            172.8887
confint(lm(Wgt~Gest,data=subset(birthsmokers,Smoke=="yes")))
##
                    2.5 %
                              97.5 %
## (Intercept) -3662.7169 -1286.4114
## Gest
                 108.7586
                            169.2989
```

Data Analysis Conclusions:

There is strong evidence of difference in correlation between Non Smoker & Smokers as observed from the data. More specifically there is 8 units of difference between predictors & response for Smokers vs Non-Smokers