

# Spatial Point Patterns

## Interaction

Josep Lluís Carrasco

Departament de Fonaments Clínics  
Universitat de Barcelona

# Outline

## 1 Distance Methods

## 2 Models for non-Poisson patterns

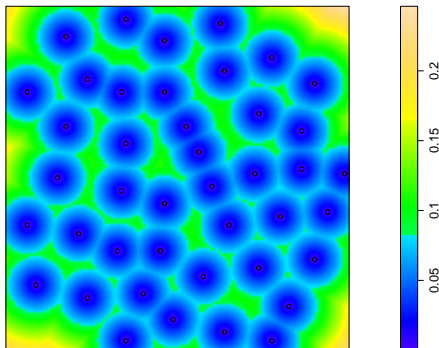
# Outline

- 1 Distance Methods
- 2 Models for non-Poisson patterns

# Empty space distances

- The distance from a fixed reference location  $u$  in the window to the nearest data point.

Empty space distances



# Empty space distances

$$d(u, \mathbf{x}) = \min \{ \|u - x_i\| : x_i \in \mathbf{x} \}$$

- Cumulative distribution function

$$F(r) = P[d(u, \mathbf{X}) \leq r]$$

- Empirical distribution function

$$F^*(r) = \frac{1}{m} \sum_j \mathbf{1}\{d(u_j, x) \leq r\}$$

- Empirical *edge corrected* distribution function

$$\hat{F}(r) = \sum_j e(u_j, r) \mathbf{1}\{d(u_j, x) \leq r\}$$

where  $e(u_j, r)$  is a weight to correct the edge effect.

# Edge correction

- Simplest approach: **border method**.
- Only use those points at a distance  $r$  to the area boundary.
- For each  $r$  the function is estimated as

$$\hat{F}^b(r) = \frac{\hat{F}(r) \cap A_{(-r)}}{|A_{(-r)}|}$$

- Other option: edge effect as a censoring issue (Baddeley and Gill, 1997)

# Empty space distances

- Homogeneous Poisson Point Process

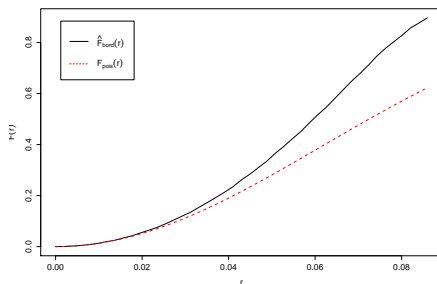
$$F_{pois}(r) = 1 - \exp(-\lambda\pi r^2)$$

- Regular pattern

$$\hat{F}(r) > F_{pois}(r)$$

- Clustered pattern

$$\hat{F}(r) < F_{pois}(r)$$



# Nearest neighbour distances

$$t_i = \min_{j \neq i} \|x_i - x_j\|$$

- Cumulative distribution function

$$G(r) = P[d(u, \mathbf{X} \setminus \{u\}) \leq r | u \in \mathbf{X}]$$

- Empirical distribution function

$$G^*(r) = \frac{1}{n} \sum_i \mathbf{1}\{t_i \leq r\}$$

- Empirical *edge corrected* distribution function

$$\hat{G}(r) = \sum_i e(x_i, r) \mathbf{1}\{t_i \leq r\}$$



# Nearest neighbour distances

- Homogeneous Poisson point process

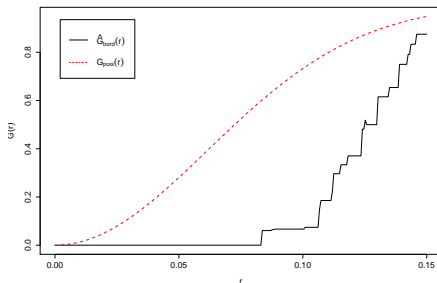
$$G_{Pois}(r) = 1 - \exp(-\lambda\pi r^2)$$

- Regular pattern

$$\hat{G}(r) < G_{Pois}(r)$$

- Clustered pattern

$$\hat{G}(r) > G_{Pois}(r)$$



# Pairwise distances

$$s_{ij} = \|x_i - x_j\|$$

- Expected number of other points of the process within a distance  $r$

$$E[n(\mathbf{X}, r)] = \lambda \cdot K(r)$$

where  $K(r)$  is the Ripley's K-function.

- Estimator of  $K(r)$

$$\hat{E}[n(\mathbf{X}, r)] = \frac{1}{N} \sum_i \sum_{j \neq i} \mathbf{1}\{\|x_i - x_j\| \leq r\}$$

$$\hat{\lambda} = \frac{N}{\text{area}(A)}$$

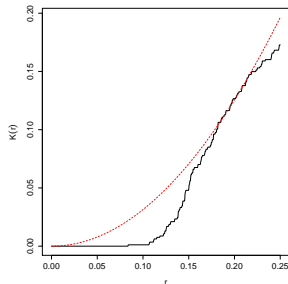
$$\hat{K}(r) = \frac{\text{area}(A)}{N^2} \sum_i \sum_{j \neq i} \mathbf{1}\{\|x_i - x_j\| \leq r\} e(x_i, x_j, r)$$

- Isotropic method.**  $e(x_i, x_j, r)$  is the proportion of a disc centered at point  $x_i$  and passing through point  $x_j$  that lies within the study area.

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- Clustered pattern

$$\hat{K}(r) > K_{Pois}(r)$$



# Additional related functions

- L-function

$$L(r) = \sqrt{\frac{K(r)}{\pi}} \quad L_{Pois}(r) = r$$

- *Pair correlation function*. Probability of observing a pair of points separated by a distance  $r$ , divided by the corresponding probability for a Poisson process.

$$g(r) = \frac{K'(r)}{2\pi r} \quad g_{Pois}(r) = 1$$

Clustering  $g(r) > 1$ ; Regularity  $g(r) < 1$

- J function

$$J(r) = \frac{1 - G(r)}{1 - F(r)} \quad J_{Pois} = 1$$

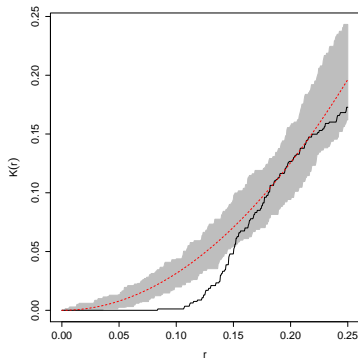
# Envelopes and Montecarlo tests

- 1 Suppose the reference curve is the theoretical K function for CSR.
- 2 Generate M independent simulations of CSR inside the study region A.
- 3 Compute the estimated K functions for each of these realizations.
- 4 Obtain the pointwise upper and lower envelopes of these simulated curves.

$$L(r) = \min_j \hat{K}^{(j)}(r)$$

$$U(r) = \max_j \hat{K}^{(j)}(r)$$

# Envelopes and Montecarlo tests



# Envelopes and Montecarlo tests

- For any fixed value of  $r$  the significance level is  $\alpha = 2 / (M + 1)$ .
- $r$  must be fixed in advance
- If the test was used as *Does  $K(r)$  lie outside for any  $r$ ?* the significance level would be higher.
- Simultaneous Monte Carlo test  $\rightarrow$  simultaneous critical bands with a global significance level  $1 / (M + 1)$ .
- ① For each estimate  $\hat{K}_i(r)$ ,  $i = 1, \dots, M$  compute its maximum deviation from Poisson  $K$  function

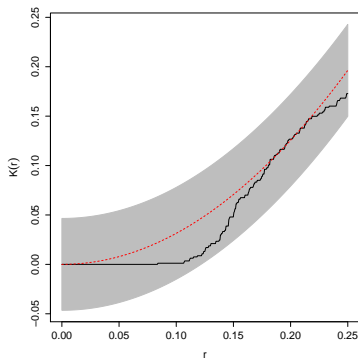
$$D_i = \max |\hat{K}_i(r) - K_{Pois}(r)|$$

②  $D_{max} = \max \{D_i\}$

$$L(r) = \pi r^2 - D_{max}$$

$$U(r) = \pi r^2 + D_{max}$$

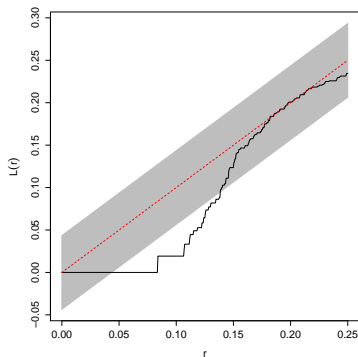
# Envelopes and Montecarlo tests





# Envelopes and Montecarlo tests

A more powerful test is obtained if the variance is stabilized using  $\sqrt{K(r)}$ .



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# Neyman-Scott processes

- ① Parent locations follow a Poisson process
- ② Numbers of children per parent are i.i.d..
- ③ Locations of children around their respective parent are i.i.d.
  - Matérn process.
    - Parent points come from a Poisson process with intensity  $\lambda(x)$ .
    - Each parent has a *Poisson* ( $\mu$ ) number of offspring.
    - Children has an independently and uniformly distributed displacement from its parent in a disc of radius  $r$  centred around the parent.
  - Thomas process.
    - Parent points come from a Poisson process with intensity  $\lambda(x)$ .
    - Each parent has a *Poisson* ( $\mu$ ) number of offspring.
    - Children has an independently and isotropic Gaussian  $N(0, \sigma^2 \mathbf{I})$  distributed displacement from its parent in a disc of radius  $r$  centred around the parent.

# Cox processes

- Poisson process with a random intensity function.
- Most used: log-Gaussian Cox Process.
- Intensities  $\lambda(\mathbf{u}) \sim N(\exp(\beta\mathbf{X}), C(r))$
- Covariance between locations takes the form:

$$C(r) = \sigma^2 f(r, \alpha)$$

- $\sigma^2$  and  $\alpha$  controls the scale and the strength of the autocorrelation.
- $f(r, \alpha)$  is a decay function. Two cases are:
  - Exponential.  $f(r, \alpha) = \exp(-r/\alpha)$
  - Gaussian  $f(r, \alpha) = \exp(-(r/\alpha)^2)$

# Model fitting: estimation method

- **Minimum contrast** (Diggle, 1983)

$$D(\theta) = \int_a^b |\hat{K}(r)^q - K_\theta(r)^q|^p dr$$

- $a$ ,  $b$  and  $q$  are used to control the sampling fluctuations in the estimates of  $K$ .
  - $a$ ,  $b$ : limits of the range of distances.
  - $q > 0$  is used to transform the  $K$  function to more powerful options.
  - If  $q = 0.5$  the contrast uses the  $L$  function.
  - In spatial processes commonly  $p = 2$ .
- It gives consistent estimates.
- It can be very computationally intensive.

Diggle, PJ (1983). Statistical Analysis of Spatial Point Patterns. Chapter 5. Academic Press, London 1983

# Model fitting: estimation method

- **Composite likelihood** (Guan, 2006).
- It maximises a composite likelihood based on knowledge of the second moment of intensity (covariance between locations).
- **Palm likelihood** (Tanaka et al, 2008).
- It assumes a Palm distribution for the point process.
- Asymptotic normality of estimates (if the likelihood is correct).

Guan, Y. (2006). Journal of the American Statistical Association 101, 1502-1512

Tanaka, U. and Ogata, Y. and Stoyan, D. (2008). Biometrical Journal 50, 43-57



# Gibbs point processes

- Spatial point process models that are constructed by writing down their probability densities.

$$f(\mathbf{X}) = \alpha \left[ \prod_{i=1}^n b(x_i) \right] \left[ \prod_{i < j} c(x_i, x_j) \right]$$

- $\alpha$  is a normalising constant.
- $b(x_i)$  are the *first order* terms.
- $c(x_i, x_j)$  are the *second order* terms.
- Intensity.  $\lambda(u) = b(u) [\prod_{i=1}^n c(u, x_i)]$
- Gibbs point processes can model regular patterns.
- Estimation approach: Maximum pseudolikelihood

# Gibbs point processes for regular patterns

## Hard core process

- Intensity.  $b(u) = \lambda$
- Pairwise interaction.

$$c(u, v) = \begin{cases} 1 & \text{if } \|u - v\| > r \\ 0 & \text{if } \|u - v\| \leq r \end{cases}$$

## Strauss process

- Pairwise interaction.

$$c(u, v) = \begin{cases} 1 & \text{if } \|u - v\| > r \\ \gamma & \text{if } \|u - v\| \leq r \end{cases}$$

$$0 \leq \gamma \leq 1$$

- Intensity.  $b(u) = \lambda \cdot \gamma^{t(u)}$

where  $t(u)$  is the number of points at a distance lower than  $r$ .

# Gibbs point processes for cluster patterns

## Geyer process

- Pairwise interaction.

$$c(u, v) = \begin{cases} 1 & \text{if } \|u - v\| > r \\ \gamma & \text{if } \|u - v\| \leq r \end{cases}$$

- Intensity.  $b(u) = \lambda \cdot \gamma^{\min(s, t(u))}$

where

- $t(u)$  is the number of points at a distance lower than  $r$ .
- $s$  is the saturation parameter.

# Gibbs point processes for cluster patterns

## Area Interaction process

- Two points interact if their discs of radius  $r$  overlaps.
- Pairwise interaction.

$$c(u, v) = \begin{cases} 1 & \text{if } \|u - v\| > 2r \\ \gamma & \text{if } \|u - v\| \leq 2r \end{cases}$$

- Intensity.  $b(u) = \lambda \cdot \gamma^{-A}$

where  $A$  is the area of the region formed by the union of discs of radius  $r$  centred at the observed points.