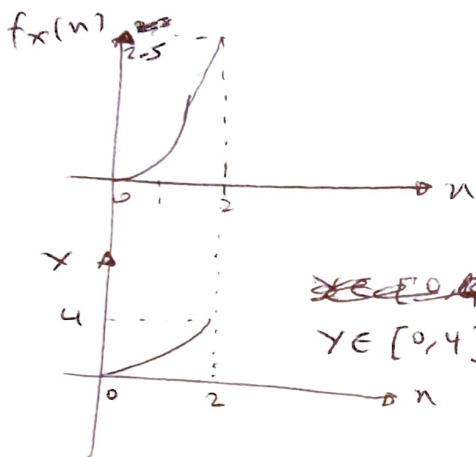


1.

$$f_X(n) = \begin{cases} \frac{5}{32} n^4 & n \in (0, 2] \\ 0 & \text{else.} \end{cases}$$

$$Y = X^2$$



(i) Using CDF formula

$$F_X(n=\alpha) = \begin{cases} 0 & \alpha \leq 0 \\ \frac{1}{32} \alpha^5 & \alpha \in (0, 2] \\ 1 & \alpha > 2 \end{cases}$$

~~Example~~ $P(Y \leq \alpha)$

$$F_Y(y=\alpha) = P(Y \leq \alpha)$$

or $F_Y(y=\alpha) = P(X^2 \leq \alpha) \Rightarrow$

or $F_Y(y=\alpha) = P(X \leq \sqrt{\alpha})$

or $F_Y(y=\alpha) = F_X(\sqrt{\alpha}) = \begin{cases} 0 & \sqrt{\alpha} \leq 0 \\ \frac{1}{32} (\sqrt{\alpha})^5 & \sqrt{\alpha} \in (0, 2] \\ 1 & \sqrt{\alpha} > 2 \end{cases}$

1 a.

or $F_Y(y=\alpha) = \begin{cases} 0 & \alpha \leq 0 \\ \frac{1}{32} \alpha^{\frac{5}{2}} & \alpha \in (0, 4] \\ 1 & \alpha > 4 \end{cases}$

Ans

$$f_X(y) = \left. \frac{d}{dy} (F_X(y=x)) \right|_{y=x} =$$

$$= \begin{cases} 0 & y < 0 \\ \frac{5}{64} y^{\frac{3}{2}} & y \in [0, 4] \\ 0 & y > 4 \end{cases}$$

(ii) Using Transformation ('cute') formula:

Ans

$$f_X(y) = \frac{d}{dy} F_X(y) = \frac{d}{dy} F_X(g^{-1}(y))$$

$$= \frac{d}{dn} (F_X(g^{-1}(y))) \cdot \left| \frac{dn}{dy} \right|$$

Here since ~~$g(x) = y$~~ $y = g(n) = n^2$,

$$g^{-1}(y) = \sqrt{y}.$$

$$\frac{dy}{dn} = \frac{d}{dn} (n^2) = 2n$$

$$\frac{dn}{dy} = \frac{d}{dy} (\sqrt{y}) = \frac{1}{2\sqrt{y}}$$

$$f_X(y) = \left. f_X(g^{-1}(y)) \cdot \left| \frac{1}{2\sqrt{y}} \right| \right|_{y=x}$$

$$= \begin{cases} 0 & \sqrt{y} \leq 0 \\ \frac{5}{32} (\sqrt{y})^4 \cdot \left| \frac{1}{2\sqrt{y}} \right| & \sqrt{y} \in (0, 2] \\ 0 & \sqrt{y} > 2 \end{cases}$$

1b (again!)

1-3

$$f_Y(y) = \begin{cases} 0 & y \leq 0 \\ \frac{5}{64} y^{\frac{3}{2}} & y \in (0, 4] \\ 0 & y > 4 \end{cases}$$

Ans (Again).

(i) Using $f_Y(y)$

$$E[Y] = \int_{y=0}^{y=4} y f_Y(y) dy$$

$$E[Y] = \int_{y=0}^4 y \times \frac{5}{64} y^{\frac{3}{2}} dy$$

$$E[Y] = \left. \frac{5}{64} \cdot y^{\frac{7}{2}} \cdot \frac{2}{\frac{7}{2}} \right|_{y=0}^{y=4}$$

$$E[Y] = \frac{10}{64 \times 7} \cdot 4^{\frac{7}{2}}$$

(1c)

$$E[Y] = \frac{20}{7} \quad \underline{\underline{\text{Ans}}}$$

(ii) using Law of The Unconscious Statistician (LOTUS):

$$E[Y] = E[X^2] = \int_{n=0}^{n=2} x^2 \cdot f_X(n) dn$$

$$E[Y] = \int_{n=0}^{n=2} n^2 \cdot \frac{5}{32} n^4 dn$$

$$E[Y] = \left. \frac{5}{32} \cdot \frac{n^7}{7} \right|_{n=0}^{n=2}$$

$$E[Y] = \frac{20}{7} \quad \underline{\underline{\text{Ans (again!)}}}$$

2.

Roll #1

$$X = \{1, 2, 3, 4, 5, 6\}$$

Roll again?

Roll #2

$$Y = \{1, 2, 3, 4, 5, 6\}$$

Let A be the event that you cash in on the first die throw.

$$E(W) = E(X) + E(Y)$$

$$W = AX + \bar{A}Y$$

then

$$W = AX + \bar{A}Y \quad (1)$$

~~Cashing in on the first throw.~~

In order to maximize W , we need to optimize the event (algorithm) A .

Intuitively, we would want to throw again if we get 'lower' values of X and hold back from throwing again if we get a high value of Y .

For quantifying A , we can utilize a variable g (for good throw) which stands for the number of allowed 'high-valued' throws allowed.

TABLE 1: $E(X|g)$ vs g .

g	$X g$	$E(X g)$
1	{6}	6
2	{6, 5}	5.5
3	{6, 5, 4}	5
4	{6, 5, 4, 3}	4.5
5	{6, 5, 4, 3, 2}	4
6	{6, 5, 4, 3, 2, 1}	3.5

Here, $g=2$,

i.e.

hold back upon getting

the highest 2

values of X : 5 or 6.

In such a case,

$$E(X|g=2) = \frac{5+6}{2} = 5.5$$

From the Table 1, we can derive $E(X|g)$ as a simple linear function of g :

$$E(X|g) = 6.5 - \frac{g}{2} \quad (2)$$

Taking expectation on both sides of (1):

$$E(W) = E(AX) + E(\bar{A}X)$$

and using g to quantify event A :

$$E(W) = E(X|g) \cdot P(g) + E(Y|\bar{g}) \cdot P(\bar{g}) \quad (3)$$

where $X|g$ = value of X given that only the highest g values of X are allowed for holding back from throwing again.

$P(g)$ = Probability of obtaining any of the highest g values on the first die throw.

$X|\bar{g}$ = gives that the ^{none of the} highest g values were landed on the first throw, the value landed on the second throw..

It may be noted that ~~that~~ ~~that~~ ~~that~~ $E(Y)$ is not dependent on the value of g or \bar{g} . So $E(Y|\bar{g}) = E(Y)$.

So $E(W) = \left(6.5 - \frac{1}{2}g\right)\left(\frac{g}{6}\right) + \left(\frac{1+2+3+4+5+6}{6}\right)\left(1 - \frac{g}{6}\right)$

$E(W) \equiv -\frac{g^2}{12} + \frac{g}{2} + 3.5$ (4)

arg (Max $E(W)$) = arg $\left(\frac{d}{dg} E(W) = 0\right)$

$g_{Max} = \arg\left(-\frac{g}{6} + \frac{1}{2} = 0\right)$

$g_{Max} = 3$ (5)

~~which is~~

Putting g_{Max} into (4):

Max. $E(W) = -\frac{3^2}{12} + \frac{3}{2} + 3.5$

Max. $E(W) = 4.25$ Ans.

An expected \$4.25 can be obtained by
thereby again ^{only} the first three lands 1, 2 or 3.



3.

outcomes	X_i	$P(X_i)$
T	1	$(1/2)$
HT	2	$(1/2)^2$
HHT	3	$(1/2)^3$
\vdots	\vdots	\vdots
$\underbrace{H \dots H T}_{(n-1)}$	n	$(1/2)^n$

$$E(X) = \sum_{i=1}^{\infty} P(X_i) \cdot X_i \cdot P(X_i)$$

$$\Rightarrow E(X) = \sum_{i=1}^{\infty} 2^{i-1} \cdot \left(\frac{1}{2}\right)^i$$

$$\Rightarrow E(X) = \sum_{i=1}^{\infty} 2^{-1}$$

3(a)

$$\Rightarrow E(X) \rightarrow \infty$$

~~Ex~~ ~~$E(X)$~~

$$P(X > 65) = P(K > \log_2(65) + 1)$$

$$\Leftrightarrow X = 2^{K-2}$$

$$\Rightarrow P(X > 65) = P(K > 7)$$

$$\Rightarrow P(X > 65) = 1 - P(K \leq 7)$$

$$\Rightarrow P(X > 65) = 1 - \sum_{k=2}^{K=7} \left(\frac{1}{2}\right)^k$$

$$\Rightarrow P(X > 65) = 1 - \frac{\left(\frac{1}{2}\right)^1 \cdot \left(1 - \left(\frac{1}{2}\right)^7\right)}{1 - \frac{1}{2}}$$

$$\Rightarrow P(X > 65) = 1 - \left(1 - \left(\frac{1}{2}\right)^7\right)$$

3(b)

$$\Rightarrow P(X > 65) = \left(\frac{1}{2}\right)^7 \quad \underline{\text{Ans}}$$

~~E(X)~~ =

If $X_{\text{max}} = 2^{30}$, then the player should not bet after
 $k = 31$ trials.

$$E(X) = \sum_{k=1}^{30} 2^{k-1} \cdot \left(\frac{1}{2}\right)^k + 2^{30} \cdot \left(\frac{1}{2}\right)^{30} \cdot 1$$

$$E(X) = \left(\frac{1}{2}\right) \times 30 + 1 = 16$$

↓
 We win
 irrespective of the
 outcome on the 31st
 trial.

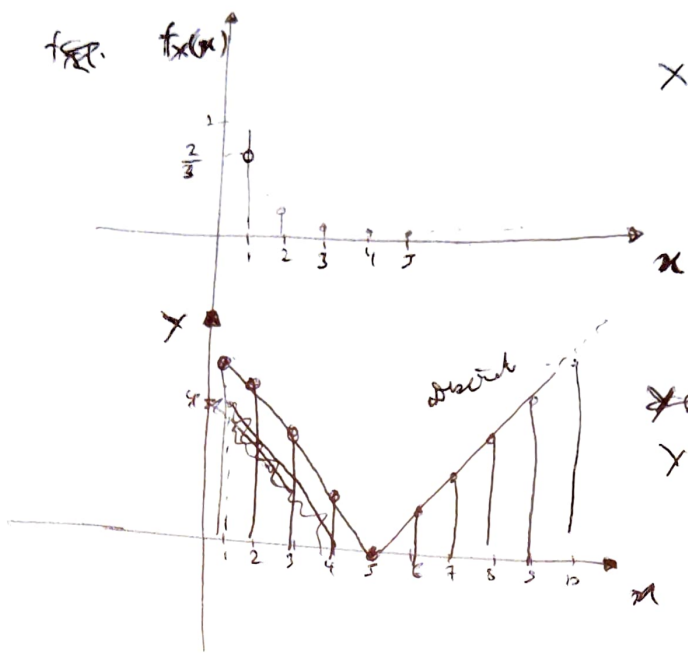
If it's a $\frac{1}{2}$, then
 the game ends and I take
 home 2^{30} .

If not, I would
 still have won the
 2^{30} and forfeit
 the game after
 the 31st trial.

————— X ————— X ————— X —————

4. $X \sim \text{geometric}(\frac{1}{3}) \Rightarrow f_X(n) = (\frac{2}{3})(\frac{1}{3})^{n-1} \quad n \geq 1, 2, 3, \dots$

$Y = |X-5|$



$X \in \{1, 2, 3, 4, \dots, \infty\}$

$X \in [0, \infty)$

$Y \in \{0, 1, 2, \dots, \infty\}$

$F_X(n) = 0 \quad n \leq 0$

$\frac{2}{3} \quad n=1$

$\frac{2}{3} + \frac{2}{3} \cdot \frac{1}{3} \quad n=2$

$\sum_{i=1}^n (\frac{2}{3})(\frac{1}{3})^{i-1} \quad n=n$

$\frac{2}{3} \cdot (1 - (\frac{1}{3})^{n+1})$

$1 - \frac{1}{3}$

X	X_i	$F_X(n)$
0	5	$(\frac{2}{3})(\frac{1}{3})^4$
1	4,6	$(\frac{2}{3})(\frac{1}{3})^3$
2	3,7	$(\frac{2}{3})(\frac{1}{3})^2$
3	2,8	$(\frac{2}{3})(\frac{1}{3})^1$
4	1,9	$(\frac{2}{3})(\frac{1}{3})^0$
5	0	$1 - (\frac{1}{3})^5$

4.1

Y_i	X_i	$\text{pmf}(Y_i)$
0	5	$(\frac{2}{3})(\frac{1}{3})^4$
1	4,6	$(\frac{2}{3})(\frac{1}{3})^3 + (\frac{2}{3})(\frac{1}{3})^5$
2	3,7	$(\frac{2}{3})(\frac{1}{3})^2 + (\frac{2}{3})(\frac{1}{3})^6$
3	2,8	$(\frac{2}{3})(\frac{1}{3})^1 + (\frac{2}{3})(\frac{1}{3})^7$
4	1,9	$(\frac{2}{3}) + (\frac{2}{3})(\frac{1}{3})^8$
5	0	$(\frac{2}{3})(\frac{1}{3})^9$
...
N	N+5	$(\frac{2}{3})(\frac{1}{3})^{N+4}$

$F_Y(y) = P(|X-5| \leq y)$

$F_Y(y) = P(n \leq \alpha+5) + P(5-n \leq \alpha)$

$F_Y(y) = P(n \leq \alpha+5) + P(n > 5-\alpha)$

$F_Y(y) = P_X(\alpha+5) + (1 - F_X(5-\alpha))$

$$F_X(4) =$$

$$n = 1, 2, 3, 4$$

5.

5.)

$$P_X(k) = \begin{cases} 0.5 & k=1 \\ 0.3 & k=2 \\ 0.2 & k=3 \\ 0 & \text{otherwise} \end{cases}$$

5(a)

$$E(X) = 0.5 + 0.3 \times 2 + 0.2 \times 3 = 1.7$$

A₁

$$E(X^2) = 0.5 + 0.3 \times 4 + 0.2 \times 9 = 3.5$$

$$V(X) = E(X^2) - (E(X))^2 = 3.5 - 1.7^2 =$$

5(b)

$$V(X) = 0.61$$

A₂

5(b)

$$\sigma(X) = \sqrt{0.61}$$

A₂

$$Y = \frac{2}{X}$$

$$E(Y) = 0.5 \times \frac{2}{1} + 0.3 \times \frac{2}{2} + 0.2 \times \frac{2}{3}$$

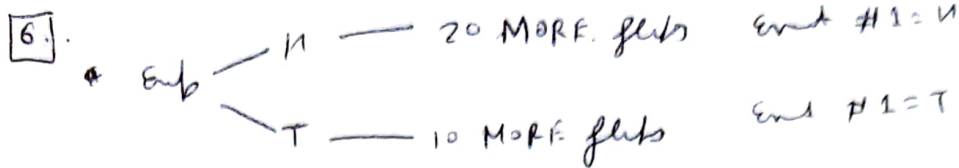
$$E(Y) = 1 + 0.3 + \frac{0.4}{3}$$

5(c)

$$E(Y) =$$

$$1.4333$$

A₁



$X = \# \text{ heads in ALL flips (including first one)}$.

$$a) P(X=5) = P(X=5 | \#1=H) \cdot P(\#1=H) + P(X=5 | \#1=T) \cdot P(\#1=T)$$

6(a)

$$P(X=5) = {}^{20}C_4 \cdot \left(\frac{1}{2}\right)^{20} \cdot \left(\frac{1}{2}\right) + {}^{10}C_5 \cdot \left(\frac{1}{2}\right)^{10} \cdot \left(\frac{1}{2}\right)$$

$$= 0.1254 \quad \underline{Ans}$$

$$b) P(\#1=H | X=5) = \frac{P(X=5 | \#1=H) \cdot P(\#1=H)}{P(X=5)}$$

$$P(\#1=H | X=5) = \frac{{}^{20}C_4 \cdot \left(\frac{1}{2}\right)^{20} \cdot \left(\frac{1}{2}\right)}{0.1254}$$

6(b)

$$P(\#1=H | X=5) = 0.0184 \quad \underline{Ans}$$

$$c) P(\#Z_{end}=H | X=5) = \frac{P(X=5 | \#Z_{end}=H) \cdot P(\#Z_{end}=H)}{P(X=5)}$$

$$P(\#Z_{end}=H | X=5) = \frac{P(X=5 | \#Z_{end}=H, \#1=H) \cdot P(\#1=H) + P(X=5 | \#Z_{end}=H, \#1=T) \cdot P(\#1=T)}{0.1254} \cdot \left(\frac{1}{2}\right)$$

$$P(\#Z_{end}=H | X=5) = \left[{}^{19}C_3 \cdot \left(\frac{1}{2}\right)^{19} \cdot \left(\frac{1}{2}\right) + {}^9C_4 \cdot \left(\frac{1}{2}\right)^9 \cdot \left(\frac{1}{2}\right) \right] \left(\frac{1}{2}\right)$$

$$P(\#Z_{end}=H | X=5) = 0.4945 \quad \underline{Ans}$$

1/7.

$$\text{eg } P(AB) = P(A) \cdot P(B) \quad (1)$$

$$\text{check if } P(AB|C) = P(A|C) \cdot P(B|C) \quad (?)$$

2/7

FALSE: Ans

Counter example: (found on stack overflow).

exp: Flip 2 coins. ~~noted~~ ~~as a repeated~~

Event A: Coin 1 is H

Event B: Coin 2 is H

Event C: Both coins ^{show} the same face.

$$P(AB) = \frac{1}{4} \quad P(A) \cdot P(B) = \frac{1}{4} \quad \text{i.e. A and B are independent events}$$

$\Omega_A, \Omega_B:$

TT	TH
HT	HH

$$P(A|C) = \frac{1}{2} \quad P(B|C) = \frac{1}{2}$$

 $\Omega_C:$

TT	HH
----	----

$$\text{But } P(AB|C) = \frac{1}{2} \neq P(A|C) \cdot P(B|C)$$

_____ X _____ X _____

1/8

given

$$P(A+B) = P(A) + P(B)$$

(1)

 \Rightarrow

$$P(AB) = 0$$

$$A \cap B = \emptyset$$

(1a)

check if

$$P(A+B|C) = P(A|C) + P(B|C)$$

(2)

~~C is a subset of~~

$$C \subseteq \Omega$$

i) $A \cap B$ in Ω ~~exists~~ does not existthen $A \cap B$ does not exist in any subset of Ω , such as C , too.

So $P(AB|C) = 0$ too, like $P(AB) = 0$ ~~and~~ ^{can} ~~is already~~ ^{from (1a)} ~~seen~~ ^{to}

$$So \quad P(A+B|C) = P(A|C) + P(B|C) - P(AB|C)$$

But $P(AB|C) = 0$

1/8

So

$$P(A+B|C) = P(A|C) + P(B|C)$$

Ans

Hence proved.

The statement is TRUE!

X ————— X ————— X

2/102

2.1

given A is ALMOST INDEPENDENT of B if

$$P(A|B) \in [P(A) - 0.01, P(A) + 0.01]$$

(a) if given A is 'AI' of B , or $P(A|B) \in [P(A) - 0.01, P(A) + 0.01]$ (2)
check if \bar{A} is also 'AI' of B .

we know that $P(A|B) + P(\bar{A}|B) = 1$

$$\Rightarrow P(\bar{A}|B) = 1 - P(A|B)$$

$$\Rightarrow P(\bar{A}|B) = 1 - [P(A) - 0.01, P(A) + 0.01]$$

$$\Rightarrow P(\bar{A}|B) = [1 - P(A) - 0.01, 1 - P(A) + 0.01]$$

$$\Rightarrow P(\bar{A}|B) = [P(\bar{A}) - 0.01, P(\bar{A}) + 0.01]$$

2/2 (a) P & A

\Rightarrow

\bar{A} is 'AI' of B

if A is 'AI' of B

Ans.

(b) given A is 'AI' of B i.e. $P(A|B) \in [P(A) - 0.01, P(A) + 0.01]$

i.e. $P(A|B) \in [P(A) - 0.01, P(A) + 0.01]$ (2) (1)

check if B is also 'AI' of A

i.e. $P(B|A) \in [P(B) - 0.01, P(B) + 0.01]$ (2) (?)

$$P(B|A) = \frac{P(A|B) \cdot P(B)}{P(A)}$$

$$\Rightarrow P(B|A) = \frac{[P(A) - 0.01, P(A) + 0.01] \cdot P(B)}{P(A)}$$

$$P(B|A) = \left[P(B) - \underbrace{0.01 \frac{P(B)}{P(A)}}_{\downarrow}, P(B) + 0.01 \frac{P(B)}{P(A)} \right]$$

For B to be 'AI' of A too,

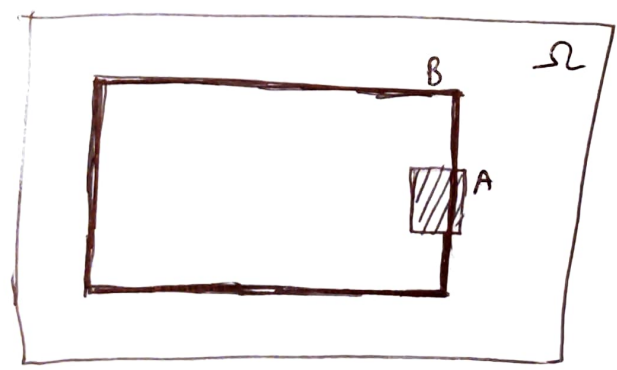
we require that $\frac{P(B)}{P(A)} \leq 1$

which is NOT a general & ~~correct~~ correct ~~assumption~~ inequality in general.

(b) \therefore The statement is false;
 1. A given A is 'AI' of B
 ~~B is 'AI' of A~~

Ans

Counterexample



Here $B \square$
 $A \blacksquare$

$P(B) = 0.75$

Let's take $P(A) = 0.01$
 ~~$P(A) = 0.01$~~
 Say $P(A) = 0.01$

Assume that 1% of A is outside B .

Then that is $P(A)$

By construction / hypothesis $\left\{ \begin{array}{l} P(B) = 0.75 \\ P(A) = \text{some smaller value} \\ P(A \cap B) = 0.01 P(A) \end{array} \right.$

$P(A|B) =$

$P(A|B) = \frac{P(A \cap B)}{P(B)}$

$\therefore P(A|B) = \frac{0.01 P(A)}{0.75}$

$\therefore P(A|B) = \frac{0.01}{0.75} P(A)$

$P(B|A) = \frac{P(A \cap B)}{P(A)} = \frac{0.01 P(A)}{P(A)} = 0.01$

$\therefore P(B|A) = 0.01$

$\neq \left(P(B) - 0.01, P(B) + 0.01 \right)$

————— X ————— X —————