statistics

- we've already done some stats using numpy:
 - o get 1000 *continuous* (float) values evenly distributed over the interval [0, 1), this is called a continuous uniform random distribution:

```
import numpy as np
c = np.random.random(1000)
```

• get 1000 *discrete* (integer) values evenly distributed over the interval [0, 10), this is called a discrete uniform random distribution:

```
d = np.random.randint(0, 10, 1000)
```

• we know how to check how these values are distributed, by visualizing them:

```
import matplotlib.pyplot as plt
def cf():
    plt.close('all')
f, ax = plt.subplots()
ax.hist(c, bins=30)
```

- if nbins is too low, you can't capture enough of the variability of your data in the plotted distribution
- if nbins is too high, you capture too much of the variability, get a very noisy distribution
- guess and test, but a decent rule of thumb for continuous distributions is nbins = np.sqrt(n)
- can also use bins='auto' to automatically calculate reasonable bin size, but sometimes seems
 to use too few bins, other possibilities include e.g. bins='sqrt'
- for discrete distributions, you should probably use as many bins as you have possible values:

```
ds.min() # check that we got what we asked for
ds.max()
f, ax = plt.subplots()
edges=np.arange(0, 10+1) # +1 for the right bin edge
ax.hist(ds, bins=edges)
```

• otherwise, you'll end up with artificial gaps between discrete values:

```
f, ax = plt.subplots()
edges2 = np.arange(0, 10+0.5, 0.5) # 0 to 10 inclusive, 0.5 steps
ax.hist(ds, bins=edges2) # notice the artificial gaps
```

- plotting the distribution of your data is important
 - can reveal outliers, and maybe sources of error in the data collection
 - many stats tests make assumptions about how your data are distributed, and if your data don't satisfy those assumptions, you should use a different stats test
 - good to get into the habit of plotting distribs
- in addition to uniform distrib, the other very common continuous distribution is the normal (Gaussian) distrib

```
mu, sigma = 0, 1
s = np.random.normal(loc=mu, scale=sigma, size=1000) # s for "sample"
f, ax = plt.subplots()
ax.hist(s, bins=30)
```

o generate bimodally distributed (having 2 peaks) data by combining two normal distributions

```
s1 = np.random.normal(loc=0, scale=1, size=1000)
s2 = np.random.normal(loc=5, scale=0.5, size=1000)
# confirm we got approximately what we asked for:
s1.mean() # approx 0
s1.std() # approx 1
s2.mean() # approx 5
s2.std() # approx 0.5
bimodal = np.concatenate([s1, s2]) # combine both into a single array
f, ax = plt.subplots()
ax.hist(bimodal, bins=30)
```

- are bimodal.mean() and bimodal.std() meaningful in this case? no, they're poor descriptors
 of this bimodal distribution, best way to tell is to plot and inspect the distribution
- matplotlib hist vs numpy hist:
 - to plot histograms, we've been using ax.hist() or plt.hist() from matplotlib
 - sometimes you might want to calculate a histogram without plotting it
 - np.histogram() returns the count in each bin, and the bin edges
 - n, edges = np.histogram(bimodal, bins=30)
 - then you can programatically do things like find what the peak value is, and where it is:
 - n.max(), n.argmax()
- scipy.stats
 - numpy can generate random samples from different kinds of distributions, but scipy.stats has a lot more stats functionality
 - o import scipy.stats as stats
 - stats? shows a big list of all the stats related objects and functions in scipy.stats
 - instead of just asking for a random sample of numbers from a particular kind of distribution, scipy.stats provides "random variables" as objects, which you can then not only sample, but also call their methods:

```
rv = stats.norm() # create a continuous normal random variable object
rv.mean() # returns exactly 0.0
rv.std() # returns exactly 1.0
rv = stats.norm(loc=5, scale=0.5)
rv.mean() # returns exactly 5
rv.std() # returns exactly 0.5
s = rv.rvs(1000) # sample 1000 random values from rv
f, ax = plt.subplots()
ax.hist(s, bins=30) # similar to what we got before from np.random.normal()
```

note that each time you sample a random value, you get different values out:

```
ax.hist(rv.rvs(1000), bins=30) # each call adds a new sampling to the plot
ax.hist(rv.rvs(1000), bins=30)
ax.hist(rv.rvs(1000), bins=30)
```

- the benefit of using a random variable object is that it provides an exact representation of a particular type of distribution
- o to access it analytically as a function of x, call the .pdf() method
 - rv.pdf(x) PDF = probability density function, or more typically, just "distribution"
 - probability always has to sum to 1, so area under the curve == 1
 - let's plot the exact representation of the normal distribution over top of the normalized histogram of our 1000 sampled values from that distribution:

```
f, ax = plt.subplots()
ax.hist(s, bins=30, normed=True) # plot a normalized distrib, area == 1
x = np.arange(3, 7, 0.01) # evenly spaced x values from 3 to 7
y = rv.pdf(x) # exact distribution
ax.plot(x, y)
ax.set_xlabel('x')
ax.set_ylabel('probability')
ax.set_title('mu=5, sigma=0.5, n=1000')
f.canvas.set_window_title('sampled and exact distributions')
```

· stats tests:

- o you've collected a bunch of data, presumably sampled from some natural process
- o you plot the distribution of your data, and see that it's roughly normally distributed
- how can you check if the mean of your data is significantly different from, say, 0?
 - do a stats test, which gives you p-value (probability) of null hypothesis
 - if p-value < some threshold (at least 0.05), null hypothesis is false, mean of your data is significantly different from 0
 - in this case, use a "1-sample t-test", stats.ttest_1samp(a, popmean) where a is the sample of observations and popmean is the population mean you want to test it against.

```
rv = stats.norm(loc=2, scale=10) # mean is 2, std is 10
s = rv.rvs(50) # acquire small amount of data
f, ax = plt.subplots()
ax.hist(s, bins='auto') # does it look normal? barely
t, p = stats.ttest_lsamp(s, 0) # p > 0.05, cannot reject null hypothesis
```

having higher n, i.e. more data, gives you more statistical power, i.e. better able to detect a weak effect:

```
s = rv.rvs(500) # acquire more data from same source
f, ax = plt.subplots()
ax.hist(s, bins='auto') # does it look normal? yes
t, p = stats.ttest_1samp(s, 0) # p < 0.05, can reject null hypothesis</pre>
```

or, having a stronger effect allows you to get away with less data:

```
rv = stats.norm(loc=4, scale=5) # 2x the mean, 1/2 the std
s = rv.rvs(50) # acquire small amount of data
f, ax = plt.subplots()
ax.hist(s, bins='auto') # does it look normal? barely
t, p = stats.ttest_1samp(s, 0) # p < 0.05, can reject null hypothesis</pre>
```

- if you have two samples of data, e.g. control vs. treatment, are they significantly different?
- do a 2-sample t-test stats.ttest_ind(), safest is called "Welch's", which doesn't assume the two samples have equal variance (standard deviation squared)

```
s1 = stats.norm.rvs(loc=0, scale=1, size=1000) # control
s2 = stats.norm.rvs(loc=0.5, scale=1, size=500) # treatment
```

```
f, ax = plt.subplots()
ax.hist(s1, bins='auto')
ax.hist(s2, bins='auto')
t, tp = stats.ttest_ind(s1, s2, equal_var=False) # Welch's
# tp << 0.05, reject null hypothesis, samples are significantly different</pre>
```

- t-test is a "parametric" test, assumes data come from some distribution that can be described by some set of parameters, in this case mean and std of normal distrib
- o there are also "non-parametric" tests, which assume nothing about the underlying distributions
- o this makes them safe in their assumptions, but gives them less statistical power
- one common non-parametric test is the Kolmogorov-Smirnov (2-sample) test, which only assumes that your two data sets are drawn from a continuous distribution:

```
d, kp = stats.ks_2samp(s1, s2)
# kp << 0.05, but higher than tp</pre>
```

 another common non-parametric test is Mann-Whitney U test, which doesn't even assume continuous distributions (allows for discrete distributions as well), but you pay for that with less statistical power:

```
u, up = stats.mannwhitneyu(s1, s2)
# up << 0.05, but higher than tp</pre>
```

- nice comparison of KS vs. MW: https://www.quora.com/What-are-the-differences-between-the-Kolmogorov-Smirnov-test-and-the-Mann-Whitney-U-test
- visually checking distributions for normality is important, but you can also quantitatively test for normality with tests, e.g. Kolmogorov-Smirnov 1-sample test, stats.kstest()
- null hypothesis says that the sample comes from the specified theoretical distribution:

```
mu, sigma = bimodal.mean(), bimodal.std() # blindly assume it's normal
d, p = stats.kstest(bimodal, 'norm', args=(mu, sigma))
# p = 0.0, reject null, not normal
s = stats.norm.rvs(loc=-2, scale=2, size=200)
d, p = stats.kstest(s, 'norm', args=(-2, 2))
# p > 0.05, can't reject null, likely normal
```

for more advanced R-like statistical modelling, see the statsmodels package (http://statsmodels.org)

stats exercises

- 1. Load in some example data from stats.csv. What's the easiest way to load such data? If you forget how, the internet is your friend.
- 2. The data should have two columns: control and treatment. If they aren't already 1D arrays, convert them to 1D arrays and assign them those names. Check the number of valid entries in each array. Remove any NaN values. The internet is still your friend.
- 3. Plot the distributions of both the control and treatment in the same figure. Give them labels. Do they both look normal? Do they look significantly different? Are their means and standard deviations different? Save the figure to a stats.png file.
- 4. Choose an appropriate stats test to see if the two distributions are significantly different. What happens if you choose an inappropriate test?
- 5. Use the KS test separately on each of the two distributions to check for normality.