

# ECONOMETRIC MODELS II - CAT II

May 3, 2025

## QUESTION ONE

### Fixed Effects Model Setup

In general, the Fixed Effects (FE) model is given by;

$$Y_{it} = \alpha_i + \beta_1 X_{1,it} + \beta_2 X_{2,it} + \cdots + \beta_k X_{k,it} + U_{it} \quad (1)$$

for  $i = 1, \dots, n$  (individuals) and  $t = 1, \dots, T$  (time periods), which can be written as;

$$Y_{it} = \alpha_i + \mathbf{X}'_{it}\boldsymbol{\beta} + U_{it} \quad (2)$$

where;

- (a)  $Y_{it}$  is the dependent variable
- (b)  $\alpha_i$  is the individual-specific fixed effect
- (c)  $\mathbf{X}_{it} = (X_{1,it}, X_{2,it}, \dots, X_{k,it})'$  is a  $k \times 1$  vector of regressors
- (d)  $\boldsymbol{\beta}$  is a  $k \times 1$  vector of parameters
- (e)  $U_{it}$  is the idiosyncratic error.

### (i) Deriving the FE estimator, $\hat{\boldsymbol{\beta}}$ (4 Marks)

**Step 1: Eliminating  $\alpha_i$  using the "Within Transformation"**

1. Averaging Equation (2) over time for each individual  $i$ , we have;

$$\bar{Y}_i = \alpha_i + \bar{\mathbf{X}}'_i \boldsymbol{\beta} + \bar{U}_i \quad (3)$$

2. Subtracting the time average from Equation (3) [the original equation], we have;

$$Y_{it} - \bar{Y}_i = (\mathbf{X}_{it} - \bar{\mathbf{X}}_i)' \boldsymbol{\beta} + (U_{it} - \bar{U}_i) \quad (4)$$

3. This gives the transformed FE model defined as;

$$\ddot{Y}_{it} = \ddot{\mathbf{X}}'_{it} \boldsymbol{\beta} + \ddot{U}_{it} \quad (5)$$

where;

- (a)  $\ddot{Y}_{it} = Y_{it} - \bar{Y}_i$
- (b)  $\ddot{\mathbf{X}}_{it} = \mathbf{X}_{it} - \bar{\mathbf{X}}_i$
- (c)  $\ddot{U}_{it} = U_{it} - \bar{U}_i$ .

### Step 2: Estimation

Stacking for all individuals and time periods, the FE estimator is obtained as follows;

1. From Equation (5), we have the estimate of  $\hat{\beta}$  defined as;

$$\hat{\beta}_{\text{FE}} \ddot{\mathbf{X}}'_{it} = \ddot{Y}_{it} \quad (6)$$

2. Multiplying both sides of Equation (6) by  $\ddot{\mathbf{X}}_{it}$ , we have;

$$\hat{\beta}_{\text{FE}} \ddot{\mathbf{X}}_{it} \ddot{\mathbf{X}}'_{it} = \ddot{\mathbf{X}}_{it} \ddot{Y}_{it} \quad (7)$$

3. Summing both sides of Equation (7) over all  $i$  and  $t$ , we have;

$$\hat{\beta}_{\text{FE}} \left( \sum_{i=1}^n \sum_{t=1}^T \ddot{\mathbf{X}}_{it} \ddot{\mathbf{X}}'_{it} \right) = \left( \sum_{i=1}^n \sum_{t=1}^T \ddot{\mathbf{X}}_{it} \ddot{Y}_{it} \right) \quad (8)$$

4. And this yields the required estimator for  $\hat{\beta}$  as;

$$\hat{\beta}_{\text{FE}} = \left( \sum_{i=1}^n \sum_{t=1}^T \ddot{\mathbf{X}}_{it} \ddot{\mathbf{X}}'_{it} \right)^{-1} \left( \sum_{i=1}^n \sum_{t=1}^T \ddot{\mathbf{X}}_{it} \ddot{Y}_{it} \right) \quad (9)$$

### (ii) Model with Dummy Variables

(3 Marks)

Alternatively, dummy variables can be included for each individual  $i$  (excluding one or using one as a reference to avoid collinearity). The model (in Equation (5)) then becomes;

$$Y_{it} = \gamma_1 D_{i1} + \gamma_2 D_{i2} + \cdots + \gamma_{n-1} D_{i,n-1} + \mathbf{X}'_{it} \beta + U_{it} \quad (10)$$

where;

- (a)  $D_{ij} = 1$  if observation  $i$  corresponds to individual  $j$ , and 0 otherwise
- (b)  $\gamma_j$  estimates  $\alpha_j$  relative to the base individual (excluded or referenced).

In matrix form, we have;

$$\mathbf{Y} = \mathbf{D}\boldsymbol{\gamma} + \mathbf{X}\beta + \mathbf{U} \quad (11)$$

where;

- (b)  $\mathbf{D}$  is the  $nT \times (n-1)$  dummy matrix
- (b)  $\boldsymbol{\gamma}$  is the vector of fixed effects
- (c)  $\mathbf{X}$  is the  $nT \times k$  regressor matrix
- (d)  $\beta$  is the slope coefficient vector.

as required

## QUESTION TWO

### (i) Copy and Complete Table 1

(2 Marks)

Given that;

- (a) Model 1 has log-likelihood =  $-11680$
- (b) Model 2 has log-likelihood =  $-9218$
- (c) The p-value is  $< 2.2 \times 10^{-16}$

The table 1 is completed by computing the following;

(1) **Degrees of freedom (df):**

- (a) Model 1 having 2 predictors  $\rightarrow 2 \text{ df} + 1 \text{ intercept} = 3 \text{ parameters}$
- (b) Model 2 having 4 predictors  $\rightarrow 4 \text{ df} + 1 \text{ intercept} = 5 \text{ parameters}$
- (c) Therefore, the Df (difference) =  $5 - 3 = 2$

(2) **Likelihood Ratio Test Statistic (Chi-sq):**

$$\text{Chisq} = 2 \times (\text{LogLik}_{\text{Model 2}} - \text{LogLik}_{\text{Model 1}}) = 2 \times (-9218 - (-11680)) = 2 \times 2462 = 4924$$

Therefore, the completed table is given as follows;

Table 1: Likelihood Ratio Test					
Model	#Df	LogLik	Df	Chisq	Pr(>Chisq)
1	3	-11680			
2	5	-9218	2	4924	$< 2.2 \times 10^{-16}$

### (ii) Which Model Fits the Data Better? Explain

(2 Marks)

Model 2 fits the data better because the Log-likelihood is higher ( $-9218$  against  $-11680$ ).

Also, the Likelihood Ratio test yields a very large chi-square value (4924) with 2 df, implying a better fit.

Finally, since the p-value is  $< 2.2 \times 10^{-16}$ , is highly significant, we conclude that Model 1 is sufficiently, a smaller model. Therefore, Model 2 significantly improves the fit by including SLOPE and ET.

### (iii) R Program

```
# (iii-1) Read in the dataset
data <- read.csv("data.csv")
```

(1 Mark)

```
# (iii-2) Fit the two logistic regression models
model1 <- glm(STATUS ~ TASP + ELEVATION, data = data, family=binomial)
model2 <- glm(STATUS ~ TASP + SLOPE + ELEVATION + ET, data = data, family= binomial)
```

(4 Marks)

```
# (iii-3) Run the likelihood ratio test
anova(model1, model2, test="Chisq")
```

(2 Marks)

## QUESTION THREE

### (i) Model Fit Comparison

(2 Marks)

To assess model fit, we compare R-squared, SEs, F-statistics, as in Table 1;

Table 2: Model Fit Statistics		
Criterion	OLS Model	Fixed Effects Model
Residual Std. Error	0.2834	0.283
Multiple R-squared	0.0152	0.9986
Adjusted R-squared	0.0150	0.9986
F-statistic	82.22	$3.54 \times 10^{-5}$
P-value	$< 2.2 \times 10^{-16}$	$< 2.2 \times 10^{-16}$

From Table 2, the FE model fits the data much better because;

- (a)  $R^2$  increased from 0.015 to 0.9986. Thus, the model accounts for almost all the variability in  $\ln(\text{hr})$ .
- (b) There is much higher F-statistic for Fixed Effects.
- (c) There is very slight improvement in residual standard error.

Therefore, the Fixed Effects account for the unobserved year-specific effects, which are significant. Hence, it is a better fit to the data.

### (ii) Fixed Effects Model Equation

(3 Marks)

Given the following information;

- (a) Dependent variable:  $\ln(\text{hr})$
- (b) Independent variable:  $\ln(\text{wg})$
- (c) Fixed Effects: Dummy variables for year

The fitted Fixed Effects model is defined as;

$$\ln(\text{hr})_{it} = \beta_1 \cdot \ln(\text{wg})_{it} + \sum_{y=1979}^{1988} \alpha_y \cdot D_y \quad (12)$$

where;

- (a)  $D_y$  are year dummies (one omitted to avoid multicollinearity)
- (b)  $\beta_1 = 0.0825$  is the coefficient of  $\ln(\text{wg})$
- (c)  $\alpha_y$  are the fixed effects for each year

Alternatively, the fitted Fixed Effects model is defined as;

$$\begin{aligned} \ln(\text{hr})_{it} = & 0.082531 \cdot \ln(\text{wg})_{it} \\ & + 7.454946 \cdot \text{factor}(\text{year})1979 + 7.445318 \cdot \text{factor}(\text{year})1980 \\ & + 7.452165 \cdot \text{factor}(\text{year})1981 + 7.430298 \cdot \text{factor}(\text{year})1982 \\ & + 7.397614 \cdot \text{factor}(\text{year})1983 + 7.421926 \cdot \text{factor}(\text{year})1984 \\ & + 7.452550 \cdot \text{factor}(\text{year})1985 + 7.444488 \cdot \text{factor}(\text{year})1986 \\ & + 7.458266 \cdot \text{factor}(\text{year})1987 + 7.463137 \cdot \text{factor}(\text{year})1988 \end{aligned} \quad (13)$$

### (iii) R Scatterplot Output Sketch (2 Marks)

```
scatterplot(yhat ~ lnwg | LaborSupply$year, boxplots = FALSE,
            xlab = "Hourly wage", ylab = "yhat", smooth = FALSE)
```

#### Interpretation and Expectations

- (a) There should be multiple small scatter plots, specifically, one for each year, showing  $\ln(\text{wg})$  on the horizontal or x-axis against the fitted values  $\hat{y}$ , on the vertical or y-axis.
- (b) Each panel should show the linear relationship for that year, indicating pure scatter of predicted values.
- (c) Expectation: An upward linear trend is expected in each panel, reflecting the positive relationship between wage and hours worked [as indicated by the  $\ln(\text{wg})$  coefficient].

### (iv) R Program

#### (iv-1) Read Dataset (1 Mark)

```
# Load necessary packages
library(car)      # For scatterplot
library(plm)      # For fixed effects
library(dplyr)    # Optional for data manipulation

# Read the dataset
LaborSupply <- read.csv("LaborSupply.csv")
```

#### (iv-2) Fit Models (4 Marks)

```
# OLS Model
ols_model <- lm(lnhr ~ lnwg, data = LaborSupply)
summary(ols_model)

# Fixed Effects Model using plm
fe_model <- plm(lnhr ~ lnwg+factor(year), data=LaborSupply, index=c("id", "year"),
               model = "within")
summary(fe_model)
```