

# Lecture 15. Bayesian classification

COMP90051 Statistical Machine Learning

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# This lecture

- Bayesian ideas in discrete settings
  - \* Beta-Binomial conjugacy
- Bayesian classification
  - \* non-conjugacy necessitates approximation

# How to apply Bayesian view to discrete data?

- First off consider models which *generate* the input
  - \* cf. *discriminative* models, which *condition* on the input
  - \* I.e.,  $p(y \mid \mathbf{x})$  vs  $p(\mathbf{x}, y)$ , Logistic Regression vs Naïve Bayes
- For simplicity, start with most basic setting
  - \*  $n$  coin tosses, of which  $k$  were heads
  - \* only have  $\mathbf{x}$  (sequence of outcomes), but no ‘classes’  $\mathbf{y}$
- Methods apply to **generative models** over discrete data
  - \* e.g., topic models, generative classifiers (Naïve Bayes, mixture of multinomials)

# Discrete Conjugate prior: Beta-Binomial

- Conjugate priors also exist for discrete spaces
- Consider  $n$  coin tosses, of which  $k$  were heads
  - \* let  $p(\text{head}) = q$  from a single toss (*Bernoulli dist*)
  - \* Inference question is the coin biased, i.e., is  $q \approx 0.5$

- Several draws, use

*Binomial dist*

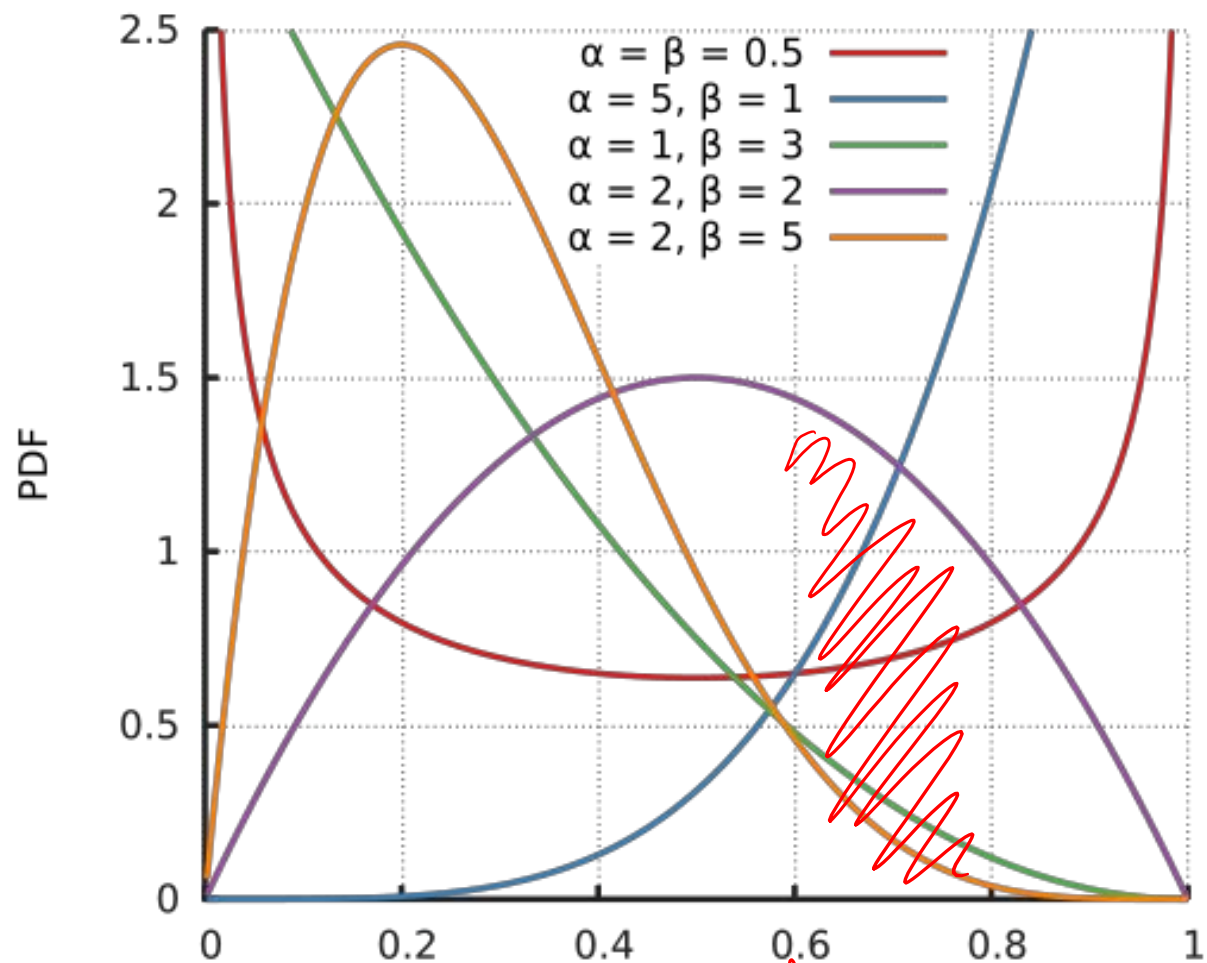
$$p(k|n, q) = \binom{n}{k} q^k (1 - q)^{n-k}$$

- \* and its conjugate prior, *Beta dist*

$$p(q) = \text{Beta}(q; \alpha, \beta)$$

$$= \frac{\gamma(\alpha + \beta)}{\gamma(\alpha)\gamma(\beta)} q^{\alpha-1} (1 - q)^{\beta-1}$$

# Beta distribution



Sourced from [https://en.wikipedia.org/wiki/Beta\\_distribution](https://en.wikipedia.org/wiki/Beta_distribution)

# Beta-Binomial conjugacy

$$p(k|n, q) = \binom{n}{k} q^k (1 - q)^{n-k}$$

$$p(q) = \text{Beta}(q; \alpha, \beta)$$

$$= \frac{\gamma(\alpha + \beta)}{\gamma(\alpha)\gamma(\beta)} q^{\alpha-1} (1 - q)^{\beta-1}$$

Sweet! We know the normaliser for Beta

Bayesian posterior

$$\begin{aligned} p(q|k, n) &\propto p(k|n, q)p(q) \\ &\propto q^k (1 - q)^{n-k} q^{\alpha-1} (1 - q)^{\beta-1} \\ &= q^{k+\alpha-1} (1 - q)^{n-k+\beta-1} \\ &\propto \text{Beta}(q; k + \alpha, n - k + \beta) \end{aligned}$$

trick: ignore  
constant factors  
(normaliser)

# Uniqueness up to normalisation

- A trick we've used many times:

*When an unnormalized distribution is proportional to a recognised distribution, we say it must be that distribution*

- If  $f(\theta) \propto g(\theta)$  for  $g$  a distribution,  $\frac{f(\theta)}{\int_{\Theta} f(\theta) d\theta} = g(\theta)$ .

- Proof:  $f(\theta) \propto g(\theta)$  means that  $\exists C$   

$$f(\theta) = C \cdot g(\theta)$$

$$\int_{\Theta} f(\theta) d\theta = C \int_{\Theta} g(\theta) d\theta = C$$

and the result follows from LHS1/LHS2 = RHS1/RHS2

# Laplace's Sunrise Problem

*Every morning you observe the sun rising. Based solely on this fact, what's the probability that the sun will rise tomorrow?*

- Use Beta-Binomial, where  $q$  is the  $\Pr(\text{sun rises in morning})$ 
  - \* posterior  $p(q|k, n) = \text{Beta}(q; k + \alpha, n - k + \beta)$
  - \*  $n = k$  = observer's age in days
  - \* let  $\alpha = \beta = 1$  (*uniform prior*)
- Under these assumptions



$$p(q|k) = \text{Beta}(q; k + 1, 1)$$
$$E_{p(q|k)} [q] = \frac{k + 1}{k + 2}$$

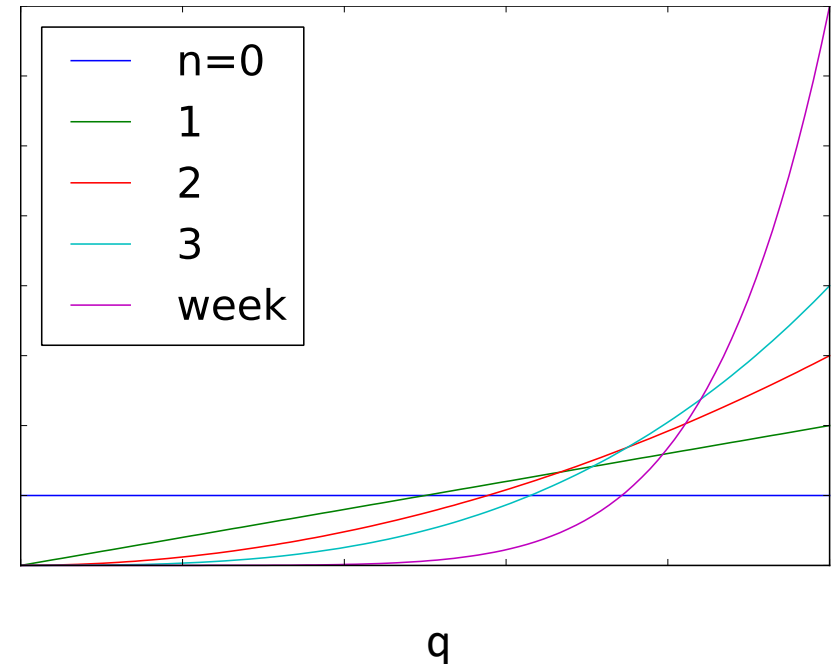
'smoothed' count of days  
where sun rose / did not



# Sunrise Problem (cont.)

Consider a human life-span

| Day (n, k)                     | $k+\alpha$ | $n-k+\beta$ | $E[q]$  |
|--------------------------------|------------|-------------|---------|
| 0                              | 1          | 1           | 0.5     |
| 1                              | 2          | 1           | 0.667   |
| 2                              | 3          | 1           | 0.75    |
| ...                            |            |             |         |
| 365                            | 366        | 1           | 0.997   |
| 2920<br>( <del>80</del> years) | 2921       | 1           | 0.99997 |



Effect of prior diminishing with data, *but never disappears completely.*

# Suite of useful conjugate priors

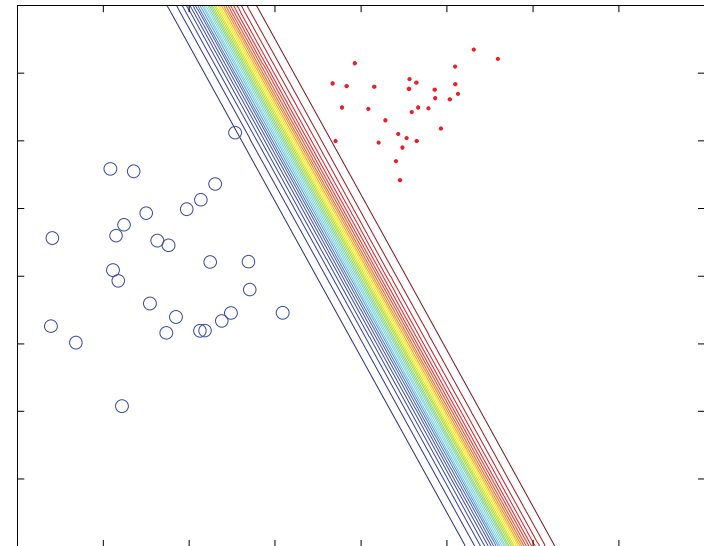
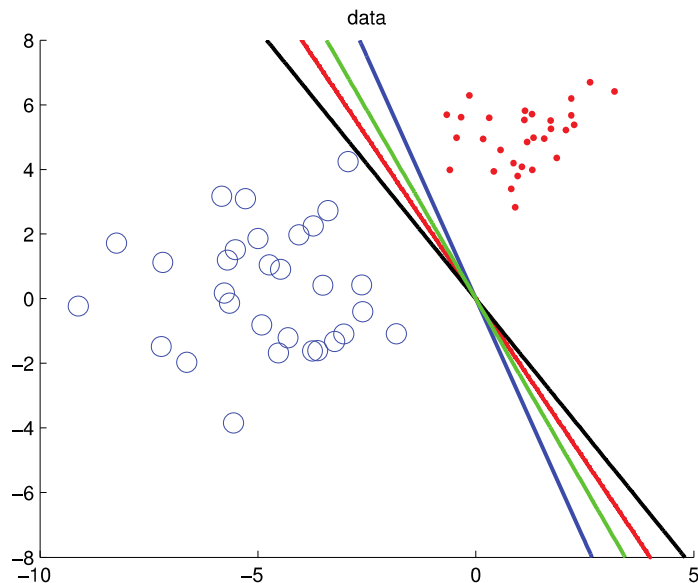
|                | likelihood  | conjugate prior   |
|----------------|-------------|---|
| regression     | Normal      | Normal (for mean)   |
|                | Normal      | Inverse Gamma (for variance)<br>or Inverse Wishart (covariance) |
| classification | Binomial    | Beta  |
|                | Multinomial | Dirichlet   |
| counts         | Poisson     | Gamma   |

# Bayesian Logistic Regression

*Discriminative classifier, which conditions on inputs. How can we do Bayesian inference in this setting?*

# Now for Logistic Regression...

- Similar problems with parameter uncertainty compared to regression
  - \* although predictive uncertainty in-built to model outputs



# No conjugacy

- Can we use conjugate prior? E.g.,
  - \* Beta-Binomial for *generative* binary models
  - \* Dirichlet-Multinomial for multiclass (similar formulation)
- Model is *discriminative*, with parameters defined using logistic sigmoid\*

$$p(y|q, \mathbf{x}) = q^y (1 - q)^{1-y}$$

$$q = \sigma(\mathbf{x}'\mathbf{w})$$

- \* need prior over  $\mathbf{w}$ , not  $q$
- \* **no known conjugate prior** (!), thus use a Gaussian prior

\* Or softmax for multiclass; same problems arise and similar solution

# Approximation

- No known solution for the normalising constant

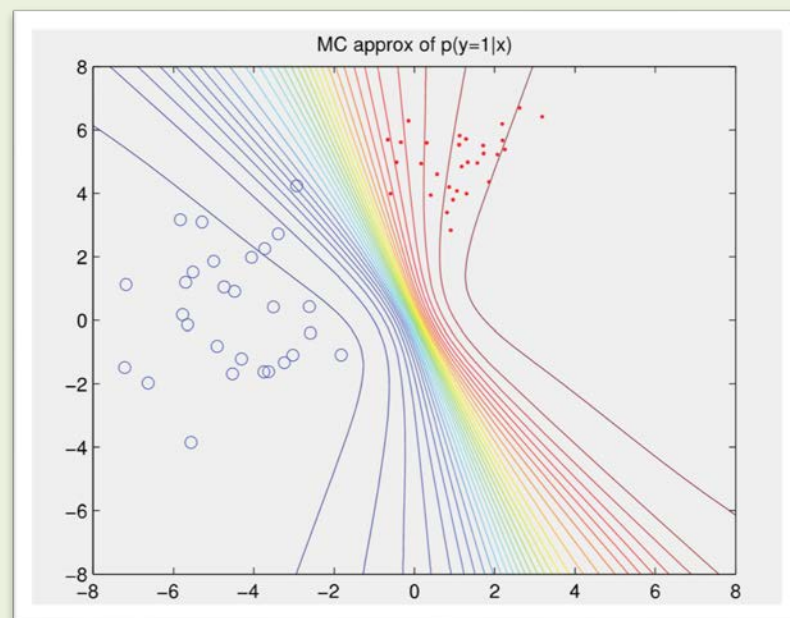
$$p(\mathbf{w}|\mathbf{X}, \mathbf{y}) \propto p(\mathbf{w})p(\mathbf{y}|\mathbf{X}, \mathbf{w})$$

$$= \text{Normal}(\mathbf{0}, \sigma^2 \mathbf{I}) \prod_{i=1}^n \sigma(\mathbf{x}'_i \mathbf{w})^{y_i} (1 - \sigma(\mathbf{x}'_i \mathbf{w}))^{1-y_i}$$

- Resolve by *approximation*

## Laplace approx.:

- assume posterior  $\simeq$  Normal about mode
- can compute normalisation constant, draw samples etc.



Murphy Fig 8.6 p258

# Summary

- Bayesian ideas in discrete settings
  - \* Beta-Binomial conjugacy
- Bayesian classification
  - \* non-conjugacy necessitates approximation
- Next time: probabilistic graphical models