

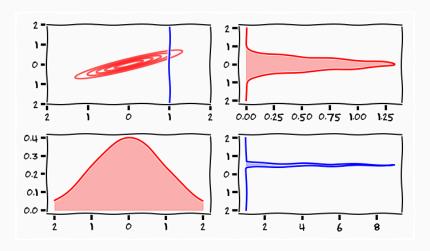
Distributions

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# Introduction

#### **Basic Probabilities**



## The Rules of Probability

#### Sum Rule

$$p(X) = \sum_{Y} p(X, Y)$$

**Product Rule** 

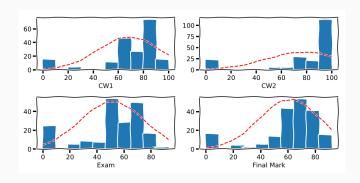
$$p(X, Y) = p(Y|X)p(X)$$

⇒ Bayes Rule

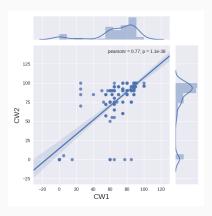
$$p(X|Y) = \frac{P(Y|X)p(X)}{p(Y)}$$

#### COMS30007 2018

# Marginals



# Marginal



$$p(\mathsf{CW1}, \mathsf{CW2}) = \sum_{x=1}^{100} p(\mathsf{CW1}, \mathsf{CW2}, \mathsf{Exam} = x) = \sum_{x=1}^{100} p(\mathsf{CW1}, \mathsf{CW2} | \mathsf{Exam} = x) p(\mathsf{Exam} = x)$$

#### Exam

$$p(Exam = 100 | CW1 = 20, CW2 = 30)$$

- What is the probability of me getting Exam=100 if CW1=20 and CW2=30
- As you will get a result on the exam the probability for all results sums to 1

$$\sum_{x=0}^{x=100} p(\mathsf{Exam} = x | \mathsf{CW1} = 20, \mathsf{CW2} = 30) = 1.0$$

#### Coursework

$$p(\text{Exam} = 70|\text{CW1} = 70) = \sum_{x=0}^{x=100} p(\text{Exam} = 70, \text{CW2} = x|\text{CW1} = 70)$$

 What is the probability that I will get Exam=70 if I got 70 on the coursework CW1

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- Remember that each conditional is a probability
- However rare it is that you get 100% on both courseworks the conditional probability over all possibe exam results will sum to one

$$\sum_{x=1}^{x=100} p(\mathsf{Exam} = x | \mathsf{CW1} = 100, \mathsf{CW2} = 100) = 1.0$$

- Remember that each conditional is a probability
- However rare it is that you get 100% on both courseworks the conditional probability over all possibe exam results will sum to one

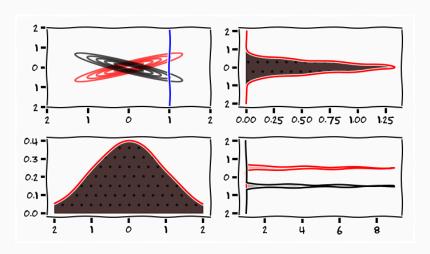
$$\sum_{x=1}^{x=100} p(\mathsf{Exam} = x | \mathsf{CW1} = 100, \mathsf{CW2} = 100) = 1.0$$

What shows that it is rare is that the probability for getting

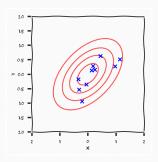
$$\sum_{x=1}^{x=100} p(\mathsf{Exam} = x, \mathsf{CW1} = 100, \mathsf{CW2} = 100)$$
$$= p(\mathsf{CW1} = 100, \mathsf{CW2} = 100) \le 1.0$$

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# Dangers of Marginals

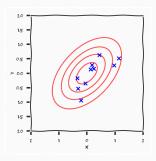


# **Learning with Distributions**



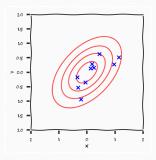
• Our goal is to understand realisations of a system

https://en.wikipedia.org/wiki/All\_models\_are\_wrong



- Our goal is to understand realisations of a system
- If we can, then we can "equate" our model with the system

https://en.wikipedia.org/wiki/All\_models\_are\_wrong



- Our goal is to understand realisations of a system
- If we can, then we can "equate" our model with the system
- Importantly not as truth, but as a useful hypothesis related to our assumptions<sup>1</sup>

https://en.wikipedia.org/wiki/All\_models\_are\_wrong

$$p(\theta|Y) = \frac{p(Y|\theta)p(\theta)}{p(Y)}$$

$$\underbrace{p(\theta|Y)}_{\text{posterior}} = \underbrace{P(Y|\theta)}_{\text{likelihood}} \cdot \underbrace{p(\theta)}_{\text{prior}} \cdot \underbrace{\frac{1}{p(Y)}}_{\text{evidence}}$$

**Likelihood** how likely is the data to come from the model specific model indexed by  $\theta$ 

$$\underline{p(\theta|Y)} = \underbrace{P(Y|\theta)}_{\text{likelihood}} \cdot \underbrace{p(\theta)}_{\text{prior}} \cdot \underbrace{\frac{1}{p(Y)}}_{\text{evidence}}$$

**Likelihood** how likely is the data to come from the model specific model indexed by  $\theta$ 

Prior what do I believe the specific model to be, i.e. how likely do I believe different  $\theta$  to be

$$\underbrace{p(\theta|Y)}_{\text{posterior}} = \underbrace{P(Y|\theta)}_{\text{likelihood}} \cdot \underbrace{p(\theta)}_{\text{prior}} \cdot \underbrace{\frac{1}{p(Y)}}_{\text{evidence}}$$

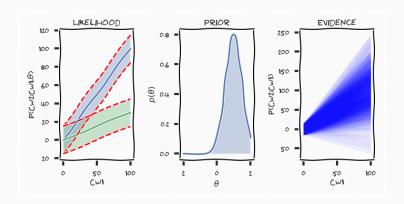
**Likelihood** how likely is the data to come from the model specific model indexed by  $\theta$ 

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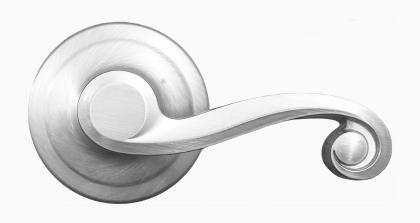
**Evidence** how likely do I think the data to be under all models weighted by how likely I think the specific models are

$$\underline{p(\theta|Y)} = \underbrace{P(Y|\theta)}_{\text{likelihood}} \cdot \underbrace{p(\theta)}_{\text{prior}} \cdot \underbrace{\frac{1}{p(Y)}}_{\text{evidence}}$$

- **Likelihood** how likely is the data to come from the model specific model indexed by  $\theta$ 
  - Prior what do I believe the specific model to be, i.e. how likely do I believe different  $\theta$  to be
  - **Evidence** how likely do I think the data to be under all models weighted by how likely I think the specific models are
  - **Posterior** which distributions of models do I believe have generated this data



$$\mathsf{CW2} = heta \cdot \mathsf{CW1} \pm 15\%$$
  $heta \sim \mathcal{N} oxed{(1.0, 0.5)}$ 



# **Discrete Distributions**

#### Bernoulli Distribution

 $\bullet$  Distribution over binary random variable  $x \in \{0,1\}$ 

$$p(x=1|\mu)=\mu$$

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$$p(x=1|\mu)=\mu$$

• Due to binary outcome

$$p(x=0|\mu)=1-\mu$$

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$$p(x=1|\mu)=\mu$$

Due to binary outcome

$$p(x=0|\mu)=1-\mu$$

Distribution

$$Bern(x|\mu) = \mu^x (1-\mu)^{1-x}$$



- ullet We want to figure out what  $\mu$  is for a specific coin
- Toss the coin N times,  $\mathcal{D} = \{x_1, x_2, \dots, x_N\}$

$$p(\mathcal{D}|\mu) = \prod_{n=1}^{N} p(x_n|\mu) = \prod_{n=1}^{N} \mu^{x_n} (1-\mu)^{1-x_n}$$

• What happens if we blindly trust this one experiment?

#### Maximum Likelihood

$$\mu_{ML} = \operatorname{argmax}_{\mu} p(\mathcal{D}|\mu) = \frac{1}{N} \sum_{n=1}^{N} x_n$$

- if we get 3 heads in a row, we believe it will always be heads
- ullet we need to include an assumption as a prior over  $\mu$

$$p(\mu|\mathcal{D}) = \frac{p(\mathcal{D}|\mu)p(\mu)}{p(\mathcal{D})}$$

Also gives us an uncertainty related to our knowledge

#### **Posterior**

$$p(\mu|\mathcal{D}) = \frac{p(\mathcal{D}|\mu)p(\mu)}{p(\mathcal{D})}.$$

ullet if we can specify a prior  $p(\mu)$  we can reach the posterior belief

#### **Posterior**

$$p(\mu|\mathcal{D}) = \frac{p(\mathcal{D}|\mu)p(\mu)}{p(\mathcal{D})}.$$

- ullet if we can specify a prior  $p(\mu)$  we can reach the posterior belief
- what do we know about coins?

#### **Posterior**

$$p(\mu|\mathcal{D}) = \frac{p(\mathcal{D}|\mu)p(\mu)}{p(\mathcal{D})}.$$

- if we can specify a prior  $p(\mu)$  we can reach the posterior belief
- what do we know about coins?
- how do I make that knowledge mathematicall explicit?

 $\bullet$  If we have a prior belief  $\mu$  we want the posterior belief to have the same functional form

 $posterior \propto likelihood \times prior$ 

ullet If we have a prior belief  $\mu$  we want the posterior belief to have the same functional form

posterior 
$$\propto$$
 likelihood  $\times$  prior

Likelihood

$$p(\mathcal{D}|\mu) = \prod_{n=1}^{N} p(x_n|\mu) = \prod_{n=1}^{N} \mu^{x_n} (1-\mu)^{1-x_n}$$

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Likelihood

$$p(\mathcal{D}|\mu) = \prod_{n=1}^{N} p(x_n|\mu) = \prod_{n=1}^{N} \mu^{x_n} (1-\mu)^{1-x_n}$$

Conjugate prior

$$p(\mu|\theta) = f_1(\theta)\mu^{f_2(\theta)}(1-\mu)^{f_3(\theta)}$$
$$\int_0^1 p(\mu|\theta)d\mu = 1$$

ullet If we have a prior belief  $\mu$  we want the posterior belief to have the same functional form

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Likelihood

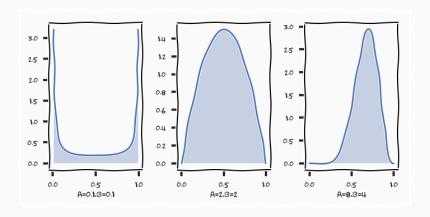
$$p(\mathcal{D}|\mu) = \prod_{n=1}^{N} p(x_n|\mu) = \prod_{n=1}^{N} \mu^{x_n} (1-\mu)^{1-x_n}$$

Conjugate prior

$$p(\mu|\theta) = f_1(\theta)\mu^{f_2(\theta)}(1-\mu)^{f_3(\theta)}$$
$$\int_0^1 p(\mu|\theta)d\mu = 1$$

Does this make philosophical sense?

### **Beta Distribution**



$$\mathsf{Beta}(\mu|a,b) = \frac{\Gamma(a+b)}{\Gamma(a) + \Gamma(b)} \mu^{a-1} (1-\mu)^{b-1}$$

$$p(\mu|\mathcal{D}) \propto p(\mathcal{D}|\mu)p(\mu)$$

$$p(\mu|\mathcal{D}) \propto p(\mathcal{D}|\mu)p(\mu)$$

$$= \prod_{i=1}^{N} \mathsf{Bern}(x|\mu)\mathsf{Beta}(\mu|a,b)$$

$$\begin{split} p(\mu|\mathcal{D}) &\propto p(\mathcal{D}|\mu) p(\mu) \\ &= \prod_{i=1}^{N} \mathsf{Bern}(x|\mu) \mathsf{Beta}(\mu|a,b) \\ &= \prod_{i=1}^{N} \mu^{x} (1-\mu)^{1-x} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \mu^{a-1} (1-\mu)^{b-1} \end{split}$$

$$\begin{split} p(\mu|\mathcal{D}) &\propto p(\mathcal{D}|\mu) p(\mu) \\ &= \prod_{i=1}^{N} \mathsf{Bern}(x|\mu) \mathsf{Beta}(\mu|a,b) \\ &= \prod_{i=1}^{N} \mu^{\mathsf{X}} (1-\mu)^{1-\mathsf{X}} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \mu^{a-1} (1-\mu)^{b-1} \\ &= \mu^{\sum_{i} \mathsf{X}_{i}} (1-\mu)^{\sum_{i} (1-\mathsf{X}_{i})} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \mu^{a-1} (1-\mu)^{b-1} \end{split}$$

$$\begin{split} \rho(\mu|\mathcal{D}) &\propto \rho(\mathcal{D}|\mu) \rho(\mu) \\ &= \prod_{i=1}^{N} \mathrm{Bern}(x|\mu) \mathrm{Beta}(\mu|a,b) \\ &= \prod_{i=1}^{N} \mu^{x} (1-\mu)^{1-x} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \mu^{a-1} (1-\mu)^{b-1} \\ &= \mu^{\sum_{i} x_{i}} (1-\mu)^{\sum_{i} (1-x_{i})} \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \mu^{a-1} (1-\mu)^{b-1} \\ &= \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \mu^{\sum_{i} x_{i} + a} (1-\mu)^{\sum_{i} (1-x_{i}) + b - 1}. \end{split}$$

#### Churn the handle

$$p(\mu|\mathcal{D}) = \frac{p(\mathcal{D}|\mu)p(\mu)}{p(\mathcal{D})} = \underbrace{\frac{p(\mathcal{D}|\mu)p(\mu)}{\int p(\mathcal{D}|\mu)p(\mu)d\mu}}_{\text{This is hard}}$$

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### Conjugacy

• We know the functional form of the posterior

#### Churn the handle

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### Conjugacy

- We know the functional form of the posterior
- We know that the posterior is proportional to the likelihood times the prior

#### Churn the handle

$$p(\mu|\mathcal{D}) = \frac{p(\mathcal{D}|\mu)p(\mu)}{p(\mathcal{D})} = \underbrace{\frac{p(\mathcal{D}|\mu)p(\mu)}{\int p(\mathcal{D}|\mu)p(\mu)d\mu}}_{\text{This is hard}}$$

### Conjugacy

- We know the functional form of the posterior
- We know that the posterior is proportional to the likelihood times the prior
- Use these facts to avoid the integral

 Because we know the form of the posterior, we can identify its parameters

$$\mathsf{Beta}(\mu|a_n,b_n) \propto \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \mu^{\sum_{i} x_i + a} \underbrace{\sum_{i} (1-x_i) + b_{-1}}_{b_n}$$

$$\mathsf{Beta}(\mu|a,b) = \frac{\Gamma(a+b)}{\Gamma(a) + \Gamma(b)} \mu^{a-1} (1-\mu)^{b-1}$$

 Because we know the form of the posterior, we can identify its parameters

$$\mathsf{Beta}(\mu|a_n,b_n) \propto \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \mu^{\underbrace{\sum_i x_i + a}_{a_n}} (1-\mu)^{\underbrace{\sum_i (1-x_i) + b}_{b_n} - 1}$$

$$\mathsf{Beta}(\mu|a,b) = \frac{\Gamma(a+b)}{\Gamma(a) + \Gamma(b)} \mu^{a-1} (1-\mu)^{b-1}$$

This leads to the following posterior

$$\mathsf{Beta}(\mu|a_n,b_n) = \frac{\Gamma(\sum_i x_i + a + \sum_i (1 - x_i) + b)}{\Gamma(\sum_i x_i + a) \Gamma(\sum_i (1 - x_i) + b)} \mu^{\sum_i x_i + a} (1 - \mu)^{\sum_i (1 - x_i) + b - 1}$$

## Lab 2

bin/bin/computerlab.jpg

### Multinomial

• If we have a variable that can take K different states

$$\boldsymbol{x} = [0, 0, 1, 0, 0, 0]^{\mathrm{T}}$$

#### Multinomial

• If we have a variable that can take K different states

$$\bm{x} = [0, 0, 1, 0, 0, 0]^{\mathrm{T}}$$

Multinomial

$$p(\mathbf{x}|\boldsymbol{\mu}) = \prod_{k=1}^K \mu_k^{\mathbf{x}_k}$$
$$\boldsymbol{\mu} = [\mu_1, \dots, \mu_k]^{\mathrm{T}}, \sum_k \mu_k = 1$$

#### Multinomial

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$$\mathbf{x} = [0, 0, 1, 0, 0, 0]^{\mathrm{T}}$$

Multinomial

$$p(\mathbf{x}|\boldsymbol{\mu}) = \prod_{k=1}^K \mu_k^{\mathbf{x}_k}$$
  $\boldsymbol{\mu} = [\mu_1, \dots, \mu_k]^{\mathrm{T}}, \sum_k \mu_k = 1$ 

Likelihood

$$p(\mathsf{D}|\boldsymbol{\mu}) = \prod_{n=1}^{N} \prod_{k=1}^{K} \mu_k^{\mathsf{x}_{nk}}$$

Multinomial

$$p(\mathbf{x}|\boldsymbol{\mu}) = \prod_{k=1}^K \mu_k^{\mathbf{x}_k}$$

Multinomial

$$p(\mathbf{x}|\boldsymbol{\mu}) = \prod_{k=1}^K \mu_k^{x_k}$$

• Conjugate prior

$$p(\mu|lpha) \propto \prod_{k=1}^K \mu_k^{lpha_k-1}$$

Multinomial

$$p(\mathsf{x}|oldsymbol{\mu}) = \prod_{k=1}^K \mu_k^{\mathsf{x}_k}$$

Conjugate prior

$$p(\mu|lpha) \propto \prod_{k=1}^K \mu_k^{lpha_k-1}$$

• Dirichlet Distribution

$$Dir(\boldsymbol{\mu}|\boldsymbol{\alpha}) = \frac{\Gamma(\alpha_0)}{\Gamma(\alpha_1) \cdot \ldots \cdot \Gamma(\alpha_K)} \prod_{k=1}^K \mu_k^{\alpha_k - 1}$$

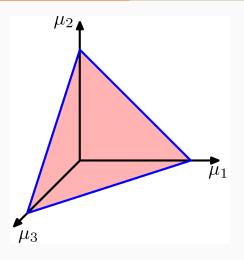
Posterior

$$p(\mu|\mathcal{D}, \alpha) \propto p(\mathcal{D}|\mu)p(\mu|\alpha) \propto \prod_{k=1}^{K} \mu_k^{\alpha_k + m_k + 1}$$
 $m_k = \sum_n x_{nk}$ 

Normalised Form

$$p(|\mathcal{D},\alpha) = \frac{\Gamma(\alpha_0 + N)}{\Gamma(\alpha_1 + m_1) \cdot \ldots \cdot \Gamma(\alpha_K + m_K)} \prod_{k=1}^K \mu_k^{\alpha_k + m_k + 1}$$

## **Dirichlet Prior**



Spans the plane  $\mu_1 + \mu_2 + \mu_3 = 1$ 

$$p(\mu|\mathcal{D}, lpha) = rac{p(\mathcal{D}|oldsymbol{\mu})p(oldsymbol{\mu}|oldsymbol{lpha})}{p(\mathcal{D}|oldsymbol{lpha})}$$

• all these priors have parameters, where do they come from?

$$p(\mu|\mathcal{D}, lpha) = rac{p(\mathcal{D}|oldsymbol{\mu})p(oldsymbol{\mu}|oldsymbol{lpha})}{p(\mathcal{D}|oldsymbol{lpha})}$$

- all these priors have parameters, where do they come from?
- either we know them

$$p(\mu|\mathcal{D}, lpha) = rac{p(\mathcal{D}|oldsymbol{\mu})p(oldsymbol{\mu}|oldsymbol{lpha})}{p(\mathcal{D}|oldsymbol{lpha})}$$

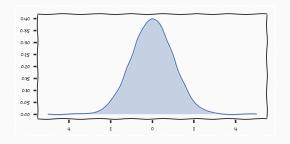
- all these priors have parameters, where do they come from?
- either we know them
- if we don't then place a prior over the priors parameters and go again

$$p(\mu|\mathcal{D}, lpha) = rac{p(\mathcal{D}|\mu)p(\mu|lpha)}{p(\mathcal{D}|lpha)}$$

- all these priors have parameters, where do they come from?
- either we know them
- if we don't then place a prior over the priors parameters and go again
- the idea is to build up a hierarchy until you can input your knowledge/assumptions

# **Continous Distributions**

#### **Gaussian Distribution**

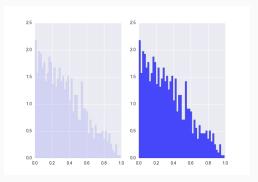


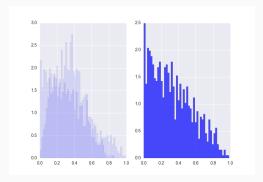
$$p(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{\frac{D}{2}} |\boldsymbol{\Sigma}|^{\frac{1}{2}}} e^{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^{\mathrm{T}} \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})}$$

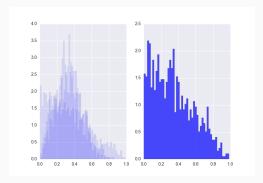
### Central Limit Theorem<sup>2</sup>

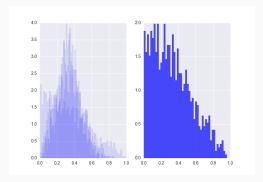
The central limit theorem states that the distribution of the sum (or average) of a large number of independent, identically distributed variables will be approximately normal, regardless of the underlying distribution.

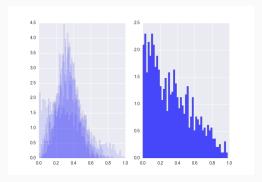
<sup>&</sup>lt;sup>2</sup>https://www.youtube.com/watch?v=wadzsURQFT4

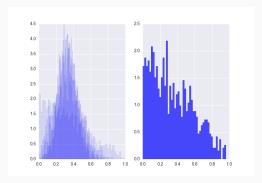


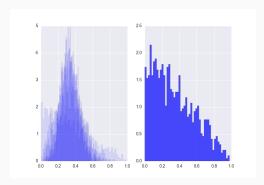


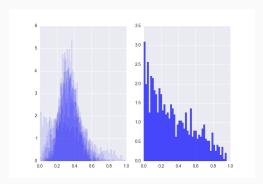


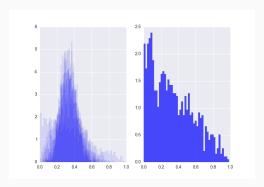


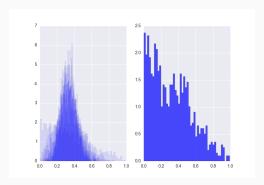


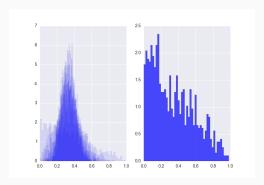


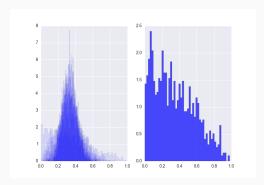


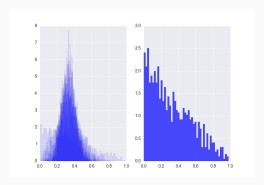


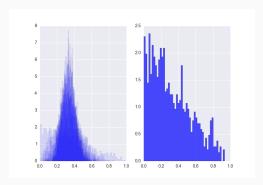


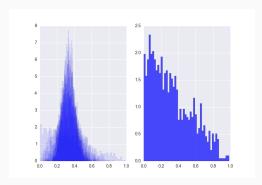


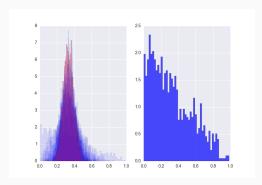






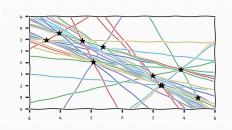






#### Cerces





The search for Cerces

Gauss made the assumption that Piazzi's measurment errors where independent draws from a unknown distribution that was fixed. This we often know as i.i.d Independent and Identically Distributed

## Conjugate Prior<sup>3</sup>

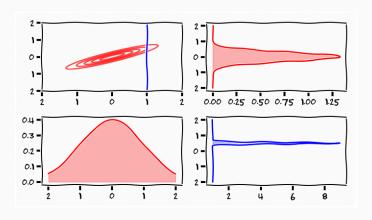
- Gaussians are self-conjugate
  - Gaussian likelihood + Gaussian Prior → Gaussian Posterior
- Gaussian distribution
  - ullet Conjugate prior for  $\mu$  is Gaussian
  - Conjugate prior for Σ is Inverse-Wishard

<sup>3</sup>https://en.wikipedia.org/wiki/Conjugate\_prior

#### Gaussian Distribution

$$p(x_1,x_2) = \mathcal{N}\left(\left[\begin{array}{c} \mu_1 \\ \mu_2 \end{array}\right], \left[\begin{array}{cc} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{array}\right]\right)$$
 Posterior  $p(x_1|x_2) \propto p(x_2|x_1)p(x_1)$  Marginal  $p(x_1) = \int p(x_1,x_2)\mathrm{d}x_2$  Conditional  $p(x_1|x_2) = \frac{p(x_1,x_2)}{p(x_2)}$ 

#### **Gaussian Identities**



Tuesday Lecture

## Exponential Family [1] Ch. 2.4

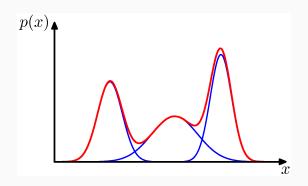
- Most distributions are parametetrised using exponentials
- Exponential family natural parametrisation

$$p(\mathbf{x}|\boldsymbol{\eta}) = h(\mathbf{x})g(\boldsymbol{\eta})e^{\boldsymbol{\eta}^{\mathrm{T}}\mathbf{u}(\mathbf{x})}$$

• Conjugate prior

$$p(\boldsymbol{\eta}|\boldsymbol{\chi},\nu) = f(\boldsymbol{\chi},\nu)g(\boldsymbol{\chi})^{\nu}e^{\nu\boldsymbol{\eta}^{\mathrm{T}}\boldsymbol{\chi}}$$

## Mixtures of Gaussians [1] Ch 2.3.9

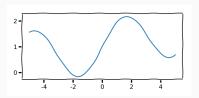


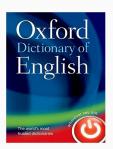
$$p(\mathsf{x}) = \sum_{k=1}^K p(k) \underbrace{p(\mathsf{x}|k)}_{\mathcal{N}(\mu_k, \Sigma_k)}$$

### **Stochastic Process**



#### Stochastic Process





#### Kologrovs Existence Theorem

Defines what a distribution needs to forfill in order for a process to exist. Each finite instantiaton of the process is this distribution.

# **Example**

### **Decisions**



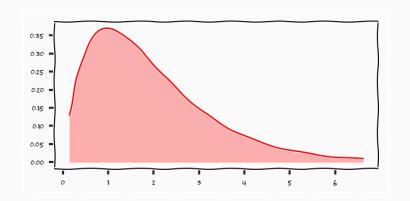
4

<sup>&</sup>lt;sup>4</sup>Reservoir Dogs Tipping Scene YouTube

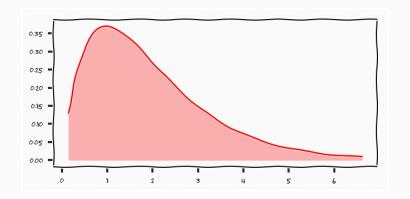
## Tip.

p(y)

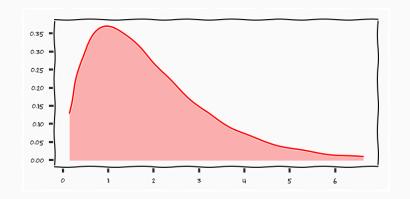
- what do I believe about tip before I see data?
- what is a sensible tip?



- ullet I believe that  $1\pounds$  is a sensible tip
- You cannot tip negative
- There is potentially an upper bound



- I believe that 1£ is a sensible tip
- You cannot tip negative
- There is potentially an upper bound
- This is not a model, its just a belief in a variable



- I believe that 1£ is a sensible tip
- You cannot tip negative
- There is potentially an upper bound
- This is not a model, its just a belief in a variable
- a model relates new phenomenon to knowledge



- it is quite hard to say something about tip without any other knowledge
- Assumption the value of tip is related to the quality of the food

#### Likelihoods

 how likely do I think the observed data y is to come from this specific x.

Tipping if I know the quality of the food what do I believe the tip should be

## What is the tip that I should expect to get?

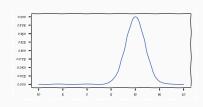
$$\mathbb{E}_{p(x)}[p(y|x)] = \int p(y|x)p(x)dx = p(y)$$

- What should I expect to get in tip
- I have an idea of the general distribution of quality of food
- Understanding is when we can relate knowledge to new phenomenon

#### Hierarchical distribution

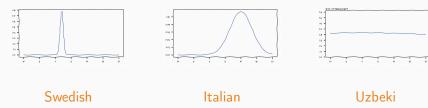
- Its quite hard to think of a prior over quality of food
- Can we parametrise the quality?

$$p(x|c) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu_c)(x-\mu_c)}{2\sigma^2}}$$



#### Hierarchical distribution

- Its quite hard to think of a prior over quality of food
- Can we parametrise the quality?
- Lets assume that if we know the cusine we have an idea



#### Hierarchical distribution

- Its quite hard to think of a prior over quality of food
- Can we parametrise the quality?
- Lets assume that if we know the cusine we have an idea
- Relating to knowledge!

## Tipping model

$$p(y, x, c) = p(y|x)p(x|c)p(c)$$

- Graphical Model shows dependency structure
- Shows "minimal" factorisation of joint distribution (model)



### Decisions

### Knowing the tip

- Which cusine did they eat if?
  - p(c|y)
- What was the quality of the food?
  - p(x|y)



## **Summary**

#### Summary

- Distributions allows us to make our assumptions explicit
- Conjugacy implies that the posterior and the prior is in the same family
- Now we have the tools that allows us to do Machine Learning

eof

# **Appendix**

#### **Posterior**

$$p(x_1|x_2) \propto p(x_2|x_1)p(x_1)$$

- 1. Multiply right-hand side
- 2. Look at the exponents
- 3. Find the three terms, constant, mixed and quadratic
- 4. Complete the square to find the parameters

$$p(x_1) = \int p(x_1, x_2) dx_2 = \mathcal{N}(\mu_1, \Sigma_{11})$$

- 1. Write out the exponent of the joint distribution
- 2. Complete Square and collect terms with  $x_1 \mu_1$  (as we know the result)
- 3. Compute integral by knowing that densities always integrates to one

#### Conditional

$$p(x_1|x_2) = \frac{p(x_1,x_2)}{p(x_2)}$$

- 1. Factorise the problem as  $p(x_1, x_2) = p(x_1|x_2)p(x_2)$
- 2. We know the marginal and the joint
- Use Schur complement to re-write the covariance matrix on block form

## References



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