

Machine Learning

Classification: The Laplace Approximation

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Introduction

Coursework 2

- The second coursework will start next week.
- This week there are no lab sessions

Today

- Classification (Task)
- Logistic Regression (Model)
- Bayesian Logistic Regression (Model)
- Laplace Approximation (Inference)

Conjugacy

$\mathsf{posterior} \propto \mathsf{likelihood} \times \mathsf{prior}$

- If we pick the conjugate prior to the likelhood parameter then the posterior is in the same family as the prior
- This means that we do not have to compute the proportionality (evidence)
- We can just multiply likelihood and prior and identify terms

Conjugacy

$$\mathcal{N}(\mathbf{w}|\boldsymbol{\mu}_{\textit{N}}, \boldsymbol{\Sigma}_{\textit{N}}) \propto \mathcal{N}(\boldsymbol{\mu}(\mathbf{w}), \boldsymbol{\Sigma}_{1}) \mathcal{N}(\mathbf{w}|\boldsymbol{\mu}, \boldsymbol{\Sigma}_{2})$$

- Multiply right-hand side
- Identify the terms on the right-hand side

Conjugacy

$$\mathcal{N}(\mathbf{w}|\boldsymbol{\mu}_{N}, \boldsymbol{\Sigma}_{N}) \propto \mathcal{N}(\boldsymbol{\mu}(\mathbf{w}), \boldsymbol{\Sigma}_{1}) \mathcal{N}(\mathbf{w}|\boldsymbol{\mu}, \boldsymbol{\Sigma}_{2})$$

- Multiply right-hand side
- Identify the terms on the right-hand side
- what if conjugacy does not make sense?

- Data $\{\mathcal{D}, \mathcal{C}\}$
 - Variates: $\mathcal{D} = \{x_i\}_{i=1}^N$
 - Features: $x \in \mathbb{R}^D$
 - Labels: $C \in \{C_1, \dots, C_k\}$
- Task: given a set of observations and their corresponding class can we associate the correct class to new observations?

Image This is an image Stella, she was a pug!

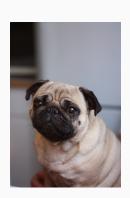


Image This is an image Stella, she was a pug!

Question does the appearance of the image make her a pug?

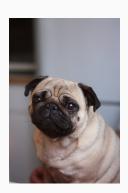


Image This is an image Stella, she was a pug!

Question does the appearance of the image make her a pug?

Question does her being a pug make an image of her appear like this?



 $\bullet\,$ The image appears this way because she was a pug

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 - Pug + Camera + lots of other stuff \rightarrow Image

$$f(x) = I$$

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- To handle all possible cases we need to think about how the data was created
 - Pug + Camera + lots of other stuff \rightarrow Image

$$f(x) = I$$

 When we formulate models we want to formulate the generation of the data

$$p(\mathcal{D}, \mathcal{C}) = p(\mathcal{D}|\mathcal{C})p(\mathcal{C})$$

- 1. Formulate the likelihood and the prior
- 2. Formulate the posterior
- 3. Get updated belief through posterior with new data

ullet Lets assume we want to update our belief of class \mathcal{C}_1 from the information in x

$$p(\mathcal{C}_1|x) = \frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x)}$$

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$$p(C_1|x) = \frac{p(x|C_1)p(C_1)}{p(x)}$$

What is the evidence in this case?

$$p(x) = \int p(x|\mathcal{C})p(\mathcal{C})d\mathcal{C} = \sum_{i=1}^{k} p(x|\mathcal{C}_i)p(\mathcal{C}_i)$$

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$$p(x) = \int p(x|\mathcal{C})p(\mathcal{C})d\mathcal{C} = \sum_{i=1}^{k} p(x|\mathcal{C}_i)p(\mathcal{C}_i)$$

Posterior

$$p(C_1|x) = \frac{p(x|C_1)p(C_1)}{\sum_{i=1}^k p(x|C_i)p(C_i)}$$

 \bullet Lets assume we have only two classes i.e. $\mathcal{C} \in \{\mathcal{C}_1, \mathcal{C}_2\}$

$$p(x) = p(x|\mathcal{C}_1)p(\mathcal{C}_1) + p(x|\mathcal{C}_2)p(\mathcal{C}_2)$$

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• Posterior:

$$p(\mathcal{C}_1|x) = \frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x|\mathcal{C}_1)p(\mathcal{C}_1) + p(x|\mathcal{C}_2)p(\mathcal{C}_2)} =$$

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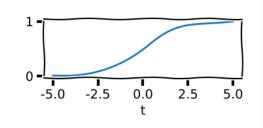
$$= \frac{\left(\frac{1}{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}\right)}{\left(\frac{1}{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}\right)} \cdot \frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x|\mathcal{C}_1)p(\mathcal{C}_1) + p(x|\mathcal{C}_2)p(\mathcal{C}_2)} =$$

ullet Lets assume we have only two classes i.e. $\mathcal{C} \in \{\mathcal{C}_1, \mathcal{C}_2\}$

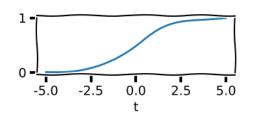
$$p(x) = p(x|\mathcal{C}_1)p(\mathcal{C}_1) + p(x|\mathcal{C}_2)p(\mathcal{C}_2)$$

• Posterior:

$$\begin{split} \rho(\mathcal{C}_1|x) &= \frac{\rho(x|\mathcal{C}_1)\rho(\mathcal{C}_1)}{\rho(x|\mathcal{C}_1)\rho(\mathcal{C}_1) + \rho(x|\mathcal{C}_2)\rho(\mathcal{C}_2)} = \\ &= \frac{\left(\frac{1}{\rho(x|\mathcal{C}_1)\rho(\mathcal{C}_1)}\right)}{\left(\frac{1}{\rho(x|\mathcal{C}_1)\rho(\mathcal{C}_1)}\right)} \cdot \frac{\rho(x|\mathcal{C}_1)\rho(\mathcal{C}_1)}{\rho(x|\mathcal{C}_1)\rho(\mathcal{C}_1) + \rho(x|\mathcal{C}_2)\rho(\mathcal{C}_2)} = \\ &= \frac{1}{1 + \frac{\rho(x|\mathcal{C}_2)\rho(\mathcal{C}_2)}{\rho(x|\mathcal{C}_1)\rho(\mathcal{C}_1)}} \end{split}$$

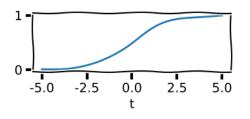


$$y = \frac{1}{1 + e^{-t}}$$



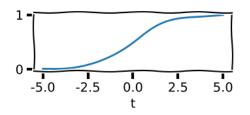
$$y = \frac{1}{1 + e^{-t}}$$

$$p(\mathcal{C}_1|x) = \frac{1}{1 + \frac{p(x|\mathcal{C}_2)p(\mathcal{C}_2)}{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}}$$



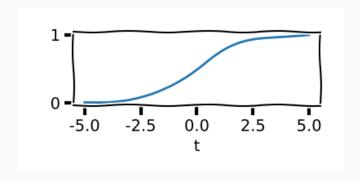
$$y = \frac{1}{1 + e^{-t}}$$

$$p(\mathcal{C}_1|x) = \frac{1}{1 + \frac{p(x|\mathcal{C}_2)p(\mathcal{C}_2)}{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}} = \frac{1}{1 + \exp\left(\log\left(\frac{p(x|\mathcal{C}_2)p(\mathcal{C}_2)}{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}\right)\right)}$$

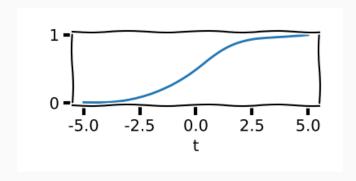


$$y = \frac{1}{1 + e^{-t}}$$

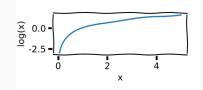
$$\begin{split} \rho(\mathcal{C}_1|x) &= \frac{1}{1 + \frac{\rho(x|\mathcal{C}_2)\rho(\mathcal{C}_2)}{\rho(x|\mathcal{C}_1)\rho(\mathcal{C}_1)}} = \frac{1}{1 + \exp\left(\log\left(\frac{\rho(x|\mathcal{C}_2)\rho(\mathcal{C}_2)}{\rho(x|\mathcal{C}_1)\rho(\mathcal{C}_1)}\right)\right)} \\ &= \frac{1}{1 + \exp\left(-\log\left(\frac{\rho(x|\mathcal{C}_1)\rho(\mathcal{C}_1)}{\rho(x|\mathcal{C}_2)\rho(\mathcal{C}_2)}\right)\right)} \end{split}$$

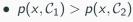


$$t = \log \left(\frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x|\mathcal{C}_2)p(\mathcal{C}_2)} \right)$$



$$t = \log \left(\frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x|\mathcal{C}_2)p(\mathcal{C}_2)} \right) = \log \left(\frac{p(x,\mathcal{C}_1)}{p(x,\mathcal{C}_2)} \right)$$





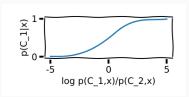
•
$$p(C_1|x) > 0.5$$

•
$$p(x, C_1) < p(x, C_2)$$

•
$$p(C_1|x) < 0,5$$

•
$$p(x, \mathcal{C}_1) = p(x, \mathcal{C}_2)$$

•
$$p(C_1|x) = 0.5$$



•
$$p(x, C_1) = 0$$

•
$$p(\mathcal{C}_1|x)=0$$

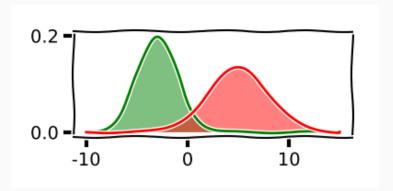
•
$$p(x, \mathcal{C}_2) \rightarrow 0$$

•
$$p(\mathcal{C}_1|x) \rightarrow 1$$

•
$$p(x, \mathcal{C}_1) = p(x, \mathcal{C}_2) = 0$$

Undefined

Likelihoods



 We haven't specified the model yet, lets make a Gaussian likelihood

$$p(x|C_i) = \mathcal{N}(x|\mu_i, \sigma_i) = \frac{1}{\sqrt{2\pi\sigma_i^2}} e^{-\frac{(x-\mu_i)^2}{2\sigma_i^2}}$$

Posterior

$$t = \log \left(\frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x|\mathcal{C}_2)p(\mathcal{C}_2)} \right)$$

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$$= \log\left(\frac{A}{B}\right) = \log(A) - \log(B)$$

$$\log(A) = \log\left(\frac{1}{\sqrt{2\pi\sigma_1^2}}e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}}p(\mathcal{C}_1)\right)$$

$$\log(A) = \log\left(\frac{1}{\sqrt{2\pi\sigma_1^2}}e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}}p(C_1)\right) =$$

$$= -\frac{1}{2}\log(2\pi\sigma_1^2) - \frac{(x-\mu_1)^2}{2\sigma_1^2} + \log(p(C_1))$$

$$\log(A) = \log\left(\frac{1}{\sqrt{2\pi\sigma_1^2}}e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}}p(C_1)\right) =$$

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$$= -\frac{1}{2}\log(2\pi\sigma_1^2) - \frac{x^2 - 2x\mu_1 + \mu_1^2}{2\sigma_1^2} + \log(p(C_1))$$

$$\log(A) - \log(B)$$

$$\log(A) - \log(B) = -\frac{1}{2}\log\left(\frac{2\pi\sigma_1^2}{2\pi\sigma_2^2}\right) + \log\left(\frac{p(C_1)}{p(C_2)}\right)$$

$$\begin{split} \log(A) - \log(B) &= -\frac{1}{2} \log \left(\frac{2\pi\sigma_1^2}{2\pi\sigma_2^2} \right) + \log \left(\frac{p(\mathcal{C}_1)}{p(\mathcal{C}_2)} \right) \\ &- \frac{x^2 - 2x\mu_1 + \mu_1^2}{2\sigma_1^2} + \frac{x^2 - 2x\mu_2 + \mu_2^2}{2\sigma_2^2} \end{split}$$

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$$-\frac{x^2 - 2x\mu_1 + \mu_1^2}{2\sigma_1^2} + \frac{x^2 - 2x\mu_2 + \mu_2^2}{2\sigma_2^2} =$$
$$= \log\left(\frac{\sigma_2}{\sigma_1}\right) + \log\left(\frac{p(\mathcal{C}_1)}{p(\mathcal{C}_2)}\right)$$

$$\log(A) - \log(B) = -\frac{1}{2} \log \left(\frac{2\pi\sigma_1^2}{2\pi\sigma_2^2} \right) + \log \left(\frac{p(C_1)}{p(C_2)} \right)$$

$$- \frac{x^2 - 2x\mu_1 + \mu_1^2}{2\sigma_1^2} + \frac{x^2 - 2x\mu_2 + \mu_2^2}{2\sigma_2^2} =$$

$$= \log \left(\frac{\sigma_2}{\sigma_1} \right) + \log \left(\frac{p(C_1)}{p(C_2)} \right)$$

$$- x^2 \left(\frac{1}{2\sigma_1^2} - \frac{1}{2\sigma_2^2} \right) + x \left(\frac{\mu_1}{\sigma_1^2} - \frac{\mu_2}{\sigma_2^2} \right) - \left(\frac{\mu_1^2}{2\sigma_1^2} - \frac{\mu_2^2}{\sigma_2^2} \right)$$

$$\log(A) - \log(B) = -\frac{1}{2} \log\left(\frac{2\pi\sigma_1^2}{2\pi\sigma_2^2}\right) + \log\left(\frac{p(C_1)}{p(C_2)}\right)$$

$$-\frac{x^2 - 2x\mu_1 + \mu_1^2}{2\sigma_1^2} + \frac{x^2 - 2x\mu_2 + \mu_2^2}{2\sigma_2^2} =$$

$$= \log\left(\frac{\sigma_2}{\sigma_1}\right) + \log\left(\frac{p(C_1)}{p(C_2)}\right)$$

$$-x^2\left(\frac{1}{2\sigma_1^2} - \frac{1}{2\sigma_2^2}\right) + x\left(\frac{\mu_1}{\sigma_1^2} - \frac{\mu_2}{\sigma_2^2}\right) - \left(\frac{\mu_1^2}{2\sigma_1^2} - \frac{\mu_2^2}{\sigma_2^2}\right) =$$

= t

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$$-\frac{x^2 - 2x\mu_1 + \mu_1^2}{2\sigma_1^2} + \frac{x^2 - 2x\mu_2 + \mu_2^2}{2\sigma_2^2} =$$

$$= \log\left(\frac{\sigma_2}{\sigma_1}\right) + \log\left(\frac{p(C_1)}{p(C_2)}\right)$$

$$-x^2\left(\frac{1}{2\sigma_1^2} - \frac{1}{2\sigma_2^2}\right) + x\left(\frac{\mu_1}{\sigma_1^2} - \frac{\mu_2}{\sigma_2^2}\right) - \left(\frac{\mu_1^2}{2\sigma_1^2} - \frac{\mu_2^2}{\sigma_2^2}\right) =$$

 $p(\mathcal{C}_1|x) = \frac{1}{1 + e^{-t}}$

• The posterior over \mathcal{C}_1 (and its the same for \mathcal{C}_2) depends on x as

$$-x^{2}\left(\frac{1}{2\sigma_{1}^{2}}-\frac{1}{2\sigma_{2}^{2}}\right)+x\left(\frac{\mu_{1}}{\sigma_{1}^{2}}-\frac{\mu_{2}}{\sigma_{2}^{2}}\right)$$

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- if $\sigma_1 = \sigma_2$
 - the posterior is linear in x

• The posterior over \mathcal{C}_1 (and its the same for \mathcal{C}_2) depends on x as

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- if $\sigma_1 = \sigma_2$
 - the posterior is linear in x
- if $\mu_2 = \mu_1$ and $\sigma_1 = \sigma_2$
 - the posterior does not depend on x

$$p(C_1|x) = \frac{1}{1 + e^{-\log\frac{p(C_1)}{p(C_2)}}} = \frac{p(C_1)}{p(C_1) + p(C_2)} = \frac{p(C_1)}{p(C)} = p(C_1)$$

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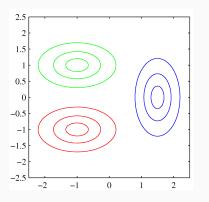
$$-x^2 \left(\frac{1}{2\sigma_1^2} - \frac{1}{2\sigma_2^2}\right) + x \left(\frac{\mu_1}{\sigma_1^2} - \frac{\mu_2}{\sigma_2^2}\right)$$

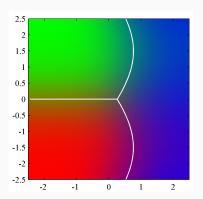
- if $\sigma_1 = \sigma_2$
 - the posterior is linear in x
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$$p(C_1|x) = \frac{1}{1 + e^{-\log\frac{p(C_1)}{p(C_2)}}} = \frac{p(C_1)}{p(C_1) + p(C_2)} = \frac{p(C_1)}{p(C)} = p(C_1)$$

 if the observations does not provide me with any information to update my belief my posterior belief is equal to my prior belief

Posterior analysis¹



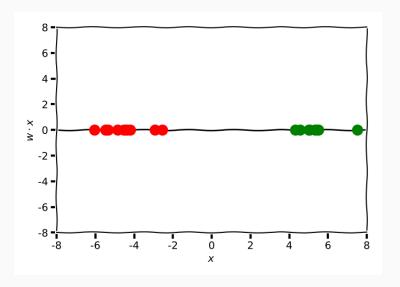


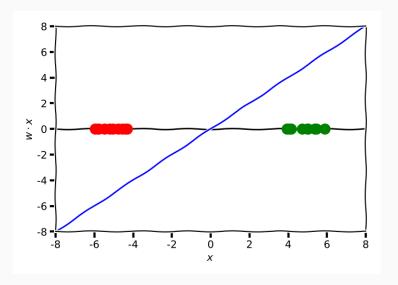
- ullet Red & Green share the same covariance \Rightarrow linear separation
- Blue & (Red & Green) different covariance \Rightarrow curved separation

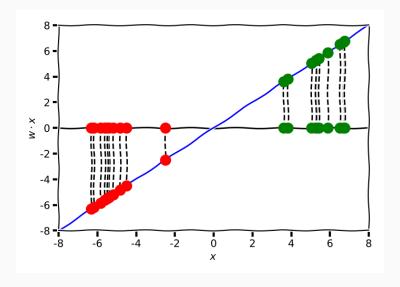
¹Bishop, C. M. (2006). Figure 4.11

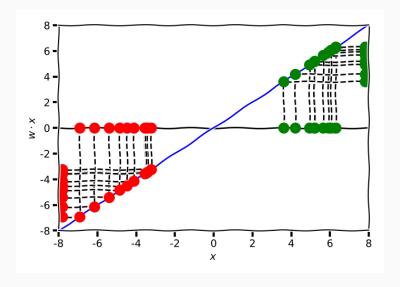
$$p(\mathcal{C}_1|x) = \frac{1}{1 + exp(-f(x))}$$

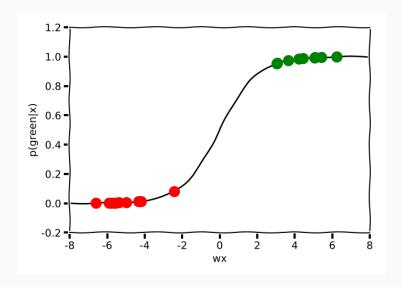
- We can derive the posterior through principle by Bayes Rule
 - This forces us to make our assumptions clear
- We can directly parametrise the posterior as we know its form
 - this is called logistic regression











 \bullet If we have $\textbf{x} \in \mathbb{R}^{100}$

- \bullet If we have $x \in \mathbb{R}^{100}$
 - Each Gaussian has

$$\mu_i \in \mathbb{R}^{100}$$

$$\Sigma_i \in \mathbb{R}^{100 \times 100}$$

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• Binary classification implies 20200 parameters

- If we have $\mathbf{x} \in \mathbb{R}^{100}$
 - Each Gaussian has

$$\mu_i \in \mathbb{R}^{100}$$

$$\Sigma_i \in \mathbb{R}^{100 \times 100}$$

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- However good our priors are this is likely to require a lot of data to learn

- If we have $\mathbf{x} \in \mathbb{R}^{100}$
 - Each Gaussian has

$$\mu_i \in \mathbb{R}^{100}$$

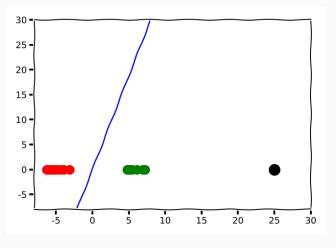
$$\Sigma_i \in \mathbb{R}^{100 \times 100}$$

- Binary classification implies 20200 parameters
- However good our priors are this is likely to require a lot of data to learn
- Logistic regression implies 101

$$\textbf{w} \in \mathbb{R}^{100}$$

- Reaching this through principle makes it a posterior
- If we parametrise it directly we have to see it as a likelihood
 - ullet we do not model ${\mathcal D}$
 - ullet we do not quantify our uncertainty in ${\mathcal D}$
 - denominator in Bayes Rule $p(\mathcal{D})$

Why Not



$$t = \log \left(\frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x|\mathcal{C}_2)p(\mathcal{C}_2)} \right)$$

$$p(\mathcal{C}_1|x) = \frac{1}{1 + exp(-f(x))} = \sigma(x)$$

- ullet We seek a function that is positive for \mathcal{C}_1 and negative for \mathcal{C}_2
- Linear classifier

$$f(x) = w^{\mathrm{T}}x$$

Maximum Likelihood

$$\hat{w} = \operatorname{argmax}_{w} p(\mathcal{C}|\mathcal{D}) = \operatorname{argmin}_{w} - \log(p(\mathcal{C}|\mathcal{D}))$$

$$E(w) = -\log p(\mathcal{C}|\mathcal{D}) = -\log \left(\prod_{i}^{N} \sigma(x_i)^{c_i} \cdot (1 - \sigma(x_i))^{1 - c_i} \right)$$

- we want to minimise the above
- take derivatives with respect to w

$$\frac{\delta E(w)}{\delta w} = \sum_{i}^{N} (\sigma(x_i) - c_i) x_i$$

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- $(\sigma(x_i) c_i)$ is the classification error
 - if 0 no gradient
 - if \neq 0 then the gradient goes in the direction of the data-point x
- Using the gradient we can now update w and find a solution

Bayesian Logistic Regression

• We have made no assumptions about the function

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- We have made no assumptions about the function
- We have seen how to do linear regression in a principled way
 - specify prior over w
 - derive posterior

We want to use the same motivation as we did for normal regression

Prior

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{m}_0, \mathbf{S}_0)$$

Likelihood

$$p(c_i|\mathbf{w},\mathbf{x}_i) = \sigma(\mathbf{w}^{\mathrm{T}}\mathbf{x}_i)$$

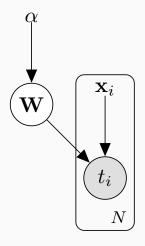
$$p(\mathbf{c}|\mathbf{w},\mathbf{x}) = \prod_{i}^{N} \sigma(\mathbf{w}^{\mathrm{T}}\mathbf{x}_i)^{c_i} \cdot (1 - \sigma(\mathbf{w}^{\mathrm{T}}\mathbf{x}_i))^{1-c_i}$$

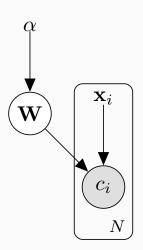
You can also do a feature mapping if you want and replace \mathbf{x} with $\Phi(\mathbf{x})$

$$\mathcal{N}(\boldsymbol{\mu}_N, \mathbf{S}_N) \neq \log \left(\prod_i^N \sigma(x_i)^{c_i} \cdot (1 - \sigma(x_i))^{1 - c_i} \right)$$
$$- \frac{1}{2} (\mathbf{w} - \mathbf{m}_0)^{\mathrm{T}} \mathbf{S}_0^{-1} (\mathbf{w} - \mathbf{m}_0) - \log(Z)$$

- Previously we could use conjugacy to reach analytical posterior
- In this case we have changed the likelihood and the Gaussian prior is non-conjugate
- This posterior is intractable

Bayesian Linear Regression





• want to reach posterior over weights

$$p(\mathbf{w}|\mathbf{t}) \propto p(\mathbf{t}|\mathbf{w})p(\mathbf{w})$$

we cannot use conjugacy anymore as likelihood is sigmoid

$$p(z|x) = \frac{1}{Z}f(z) = \frac{f(z)}{\int f(z)dz}$$

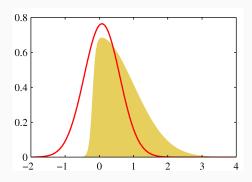
- Will use p(z) to refer to the posterior
- p(z) is unknown as we cannot compute Z
- f(z) is possible to compute if we have likelihood and prior

$$f(z) = p(x|z)p(z)$$

$$\log p(z) = \log \left(\frac{1}{Z}f(z)\right) = \log(f(z)) + \text{const w.r.t. } z$$

- p(z) and f(z) will have the same modes
 - is this always true?
- Idea: we can approximate each mode with a distribution we can normalise

Laplace Approximation Ch. 4.4 [1]



- Find the mode of the posterior
- Fit Gaussian to this mode

Taylor Expansion

$$f(x) = f(x_0) + \frac{\partial}{\partial x} f(x_0)(x - x_0) + \frac{1}{2} \frac{\partial^2}{\partial x^2} f(x_0)(x - x_0)^2 + \mathcal{O}((x - x_0)^3)$$

- A Taylor expansion is an approximation of a function around a specific value
- If we expand around a maxima x_0

$$\frac{\partial}{\partial x}f(x_0)=0$$

• This leads to

$$f(x) = f(x_0) - \frac{1}{2} \left| \frac{\partial^2}{\partial x^2} f(x_0) \right| (x - x_0)^2 + \mathcal{O}((x - x_0)^3)$$

$$f(\mathbf{w}) = p(\mathbf{t}|\mathbf{w})p(\mathbf{w})$$

• we want to find the mode of this, i.e. the maxima

$$\hat{\mathbf{w}} = \operatorname{argmax}_{\mathbf{w}} p(\mathbf{t}|\mathbf{w}) p(\mathbf{w})$$

This we know as the Maximum-a-Posterior (MAP) estimate

1. Find mode of p(z)

$$\frac{\partial}{\partial z}p(z_0)=\frac{\partial}{\partial z}f(z_0)=0$$

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2. Make Taylor Expansion around mode

$$\log f(z) \approx \log f(z_0) - \frac{1}{2} \frac{\partial^2}{\partial^2} \log(f(z_0))(z-z_0)^2$$

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3. Take exponential to get function

$$f(z) \approx f(z_0)e^{-\frac{1}{2}\underbrace{\frac{\partial^2}{\partial^2}\log(f(z_0))}_{A}(z-z_0)^2} = f(z_0)e^{-\frac{1}{2}A(z-z_0)^2}$$

$$f(z) \approx f(z_0)e^{-\frac{1}{2}A(z-z_0)^2}$$

• we want to find an approximation, to p(z) so we need to normalise to a distribution

$$p(z) = \frac{1}{Z}f(z) \approx q(z)$$

$$f(z) \approx f(z_0)e^{-\frac{1}{2}A(z-z_0)^2}$$

• we want to find an approximation, to p(z) so we need to normalise to a distribution

$$p(z) = \frac{1}{Z}f(z) \approx q(z)$$

• assume that q(z) is Gaussian, i.e. $f(z_0) = p(\text{mean})$

$$q(z) = \left(\frac{A}{2\pi}\right)^{\frac{1}{2}} e^{-\frac{A}{2}(z-z_0)^2}$$

One dimensional

$$q(z) = \left(\frac{A}{2\pi}\right)^{\frac{1}{2}} e^{-\frac{A}{2}(z-z_0)^2}$$

D dimensional

$$q(\mathbf{z}) = \frac{|\mathbf{A}|}{(2\pi)^{\frac{D}{2}}} e^{-\frac{1}{2}(\mathbf{z} - \mathbf{z}_0)^{\mathrm{T}} \mathbf{A} (\mathbf{z} - \mathbf{z}_0)} = \mathcal{N}(\mathbf{z} | \mathbf{z}_0, \mathbf{A}^{-1})$$
$$\mathbf{A} = -\nabla \nabla \log f(\mathbf{z})|_{\mathbf{z} = \mathbf{z}_0}$$

• Where A is the Hessian matrix

We want to use the same motivation as we did for normal regression

Prior

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Likelihood

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$$p(\mathbf{c}|\mathbf{w},\mathbf{x}) = \prod_{i}^{N} \sigma(\mathbf{w}^{\mathrm{T}}\mathbf{x}_i)^{c_i} \cdot (1 - \sigma(\mathbf{w}^{\mathrm{T}}\mathbf{x}_i))^{1-c_i}$$

You can also do a feature mapping if you want and replace x with $\Phi(x)$

$$q(\mathbf{w}|\mathbf{t}) = \mathcal{N}(\mathbf{m}_N, S_N) \approx p(\mathbf{w}|\mathbf{t}) \propto p(\mathbf{t}|\mathbf{w})p(\mathbf{w}) = f(\mathbf{w})$$

• Compute $f(\mathbf{w})$

$$\log p(\mathbf{w}|\mathbf{t}) = \log \left(\prod_{i}^{N} \sigma(\mathbf{w}^{\mathrm{T}} \mathbf{x}_{i})^{t_{i}} \cdot (1 - \sigma(\mathbf{w}^{\mathrm{T}} \mathbf{x}_{i}))^{1 - t_{i}} \right)$$
$$- \frac{1}{2} (\mathbf{w} - \mathbf{m}_{0})^{\mathrm{T}} \mathbf{S}_{0}^{-1} (\mathbf{w} - \mathbf{m}_{0}) - \log(Z)$$

The stationary point is the MAP estimate

$$\mathbf{S}_{N}^{-1} = -\nabla\nabla\log\rho(\mathbf{w}|\mathbf{t})|_{\mathbf{w} = \mathbf{w}_{MAP}} = \mathbf{S}_{0}^{-1} + \sum_{n=1}^{N}\sigma(\mathbf{w}^{\mathrm{T}}\mathbf{x})(1 - \sigma(\mathbf{w}^{\mathrm{T}}\mathbf{x}))\mathbf{x}\mathbf{x}^{\mathrm{T}}$$

- \bullet we can compute the Hessian around $\mathbf{w}_{\mathsf{MAP}}$
- this leads to the final approximation

$$q(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{w}_{MAP}, \mathbf{S}_N)$$

Summary

Compute a mode of the posterior distribution, i.e MAP estimate

- Compute a mode of the posterior distribution, i.e MAP estimate
- Perform Taylor expansion around mode

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- Identify elements in expansion as parameters of a Gaussian
- Normalise to a distribution
- You can do exactly the same thing with a GP

Prediction

$$egin{aligned}
ho(\mathcal{C}_1|\mathbf{x},\mathbf{t}) &= \int
ho(\mathcal{C}_1|\mathbf{x},\mathbf{w})
ho(\mathbf{w}|\mathbf{t}) \mathrm{d}\mathbf{w} \ &pprox \int
ho(\mathcal{C}_1|\mathbf{x},\mathbf{w}) q(\mathbf{w}) \mathrm{d}\mathbf{w} \end{aligned}$$

 To compute predictions we can use our new approximate posterior in place of the true posterior

- Classification often means non-conjugate prior
- Laplace Approximation
 - match modes with the true posterior
- As we often know the MAP estimate of different models we can often apply this method relatively easy
- Can be really bad if the posterior is far from Gaussian
- We can fit several modes and make a mixture

eof

References



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