

Machine Learning

Classification: The Laplace Approximation

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Labs No new lab week 11

Lectures **Cancel** Ethical Dilemmas in Machine Learning

Lectures **Workgroups** will hopefully take place during spring instead

Lectures **Cancel** 10/12 \Rightarrow last lecture on the 9/12

$$\text{posterior} \propto \text{likelihood} \times \text{prior}$$

- If we pick the conjugate prior to the likelihood parameter then the posterior is in the same family as the prior
- This means that we do not have to compute the proportionality (evidence)
- We can just multiply likelihood and prior and identify terms

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- This means that we do not have to compute the proportionality (evidence)
- We can just multiply likelihood and prior and identify terms
- *what if conjugacy does not make sense?*

- Classification (Task)
- Logistic Regression (Model)
- Bayesian Logistic Regression (Model)
- Laplace Approximation (Inference)

Classification

- Data $\{\mathcal{D}, \mathcal{C}\}$
 - Variates: $\mathcal{D} = \{x_i\}_{i=1}^N$
 - Features: $x \in \mathbb{R}^D$
 - Labels: $\mathcal{C} \in \{\mathcal{C}_1, \dots, \mathcal{C}_k\}$
- **Task:** given a set of observations and their corresponding class can we associate the correct class to new observations?

Image This is an image Stella, she was a pug!

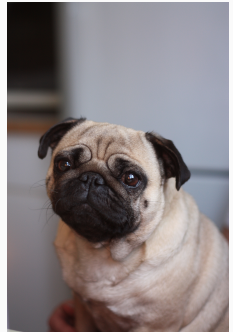


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Question does the appearance of the image make her a pug?

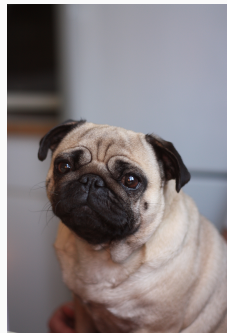


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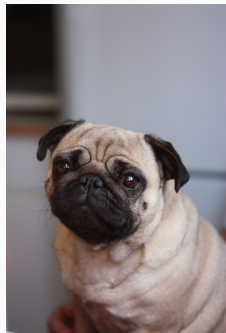
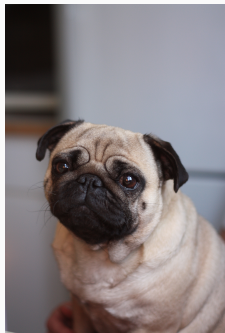


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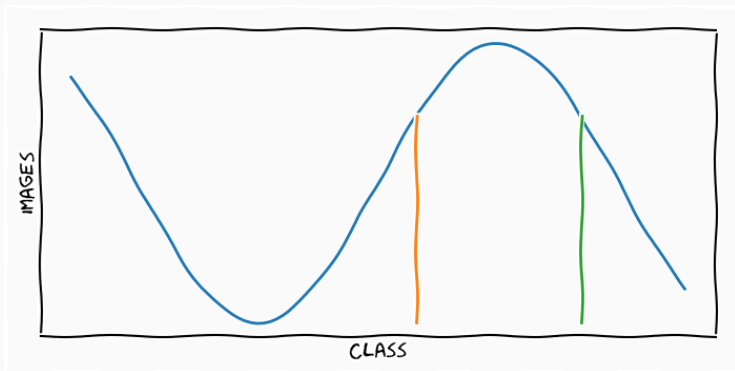
Question does the appearance of the image make her a pug?

Question does her being a pug make an image of her appear like this?

Question is it possible that this image is not of a pug?



Generative model



- *Machine Learning is an inverse problem*

$$\begin{aligned} p(\mathcal{D}, \mathcal{C}) &= p(\mathcal{D}|\mathcal{C})p(\mathcal{C}) \\ &\Rightarrow p(\mathcal{C}|\mathcal{D}) \end{aligned}$$

1. **Formulate** the likelihood and the prior
2. **Derive** the posterior
3. Get updated belief through posterior with new data

- Lets assume we want to update our belief of class \mathcal{C}_1 from the information in x

$$p(\mathcal{C}_1|x) = \frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x)}$$

Classification

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- What is the evidence in this case?

$$p(x) = \int p(x|\mathcal{C})p(\mathcal{C})d\mathcal{C} = \sum_{i=1}^k p(x|\mathcal{C}_i)p(\mathcal{C}_i)$$

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- Posterior

$$p(\mathcal{C}_1|x) = \frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{\sum_{i=1}^k p(x|\mathcal{C}_i)p(\mathcal{C}_i)}$$

- Lets assume we have only two classes i.e. $\mathcal{C} \in \{\mathcal{C}_1, \mathcal{C}_2\}$

$$p(x) = p(x|\mathcal{C}_1)p(\mathcal{C}_1) + p(x|\mathcal{C}_2)p(\mathcal{C}_2)$$

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- Posterior:

$$p(\mathcal{C}_1|x) = \frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x|\mathcal{C}_1)p(\mathcal{C}_1) + p(x|\mathcal{C}_2)p(\mathcal{C}_2)} =$$

Binary Classification

- Lets assume we have only two classes i.e. $\mathcal{C} \in \{\mathcal{C}_1, \mathcal{C}_2\}$

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- Posterior:

$$\begin{aligned} p(\mathcal{C}_1|x) &= \frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x|\mathcal{C}_1)p(\mathcal{C}_1) + p(x|\mathcal{C}_2)p(\mathcal{C}_2)} = \\ &= \frac{\left(\frac{1}{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}\right)}{\left(\frac{1}{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}\right)} \cdot \frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x|\mathcal{C}_1)p(\mathcal{C}_1) + p(x|\mathcal{C}_2)p(\mathcal{C}_2)} = \end{aligned}$$

Binary Classification

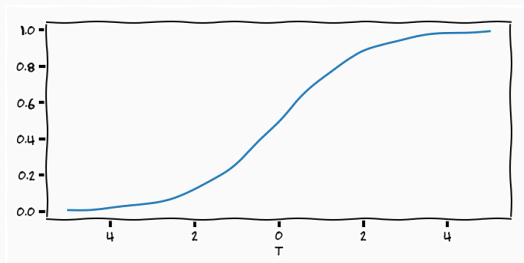
- Lets assume we have only two classes i.e. $\mathcal{C} \in \{\mathcal{C}_1, \mathcal{C}_2\}$

$$p(x) = p(x|\mathcal{C}_1)p(\mathcal{C}_1) + p(x|\mathcal{C}_2)p(\mathcal{C}_2)$$

- Posterior:

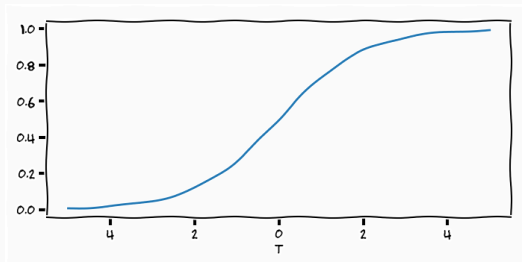
$$\begin{aligned} p(\mathcal{C}_1|x) &= \frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x|\mathcal{C}_1)p(\mathcal{C}_1) + p(x|\mathcal{C}_2)p(\mathcal{C}_2)} = \\ &= \frac{\left(\frac{1}{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}\right)}{\left(\frac{1}{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}\right)} \cdot \frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x|\mathcal{C}_1)p(\mathcal{C}_1) + p(x|\mathcal{C}_2)p(\mathcal{C}_2)} = \\ &= \frac{1}{1 + \frac{p(x|\mathcal{C}_2)p(\mathcal{C}_2)}{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}} \end{aligned}$$

Binary Classification



$$y = \frac{1}{1 + e^{-t}}$$

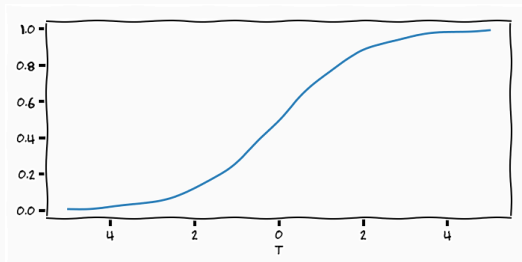
Binary Classification



$$y = \frac{1}{1 + e^{-t}}$$

$$p(C_1|x) = \frac{1}{1 + \frac{p(x|C_2)p(C_2)}{p(x|C_1)p(C_1)}}$$

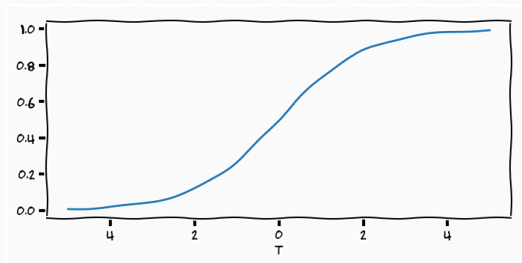
Binary Classification



$$y = \frac{1}{1 + e^{-t}}$$

$$p(C_1|x) = \frac{1}{1 + \frac{p(x|C_2)p(C_2)}{p(x|C_1)p(C_1)}} = \frac{1}{1 + \exp\left(\log\left(\frac{p(x|C_2)p(C_2)}{p(x|C_1)p(C_1)}\right)\right)}$$

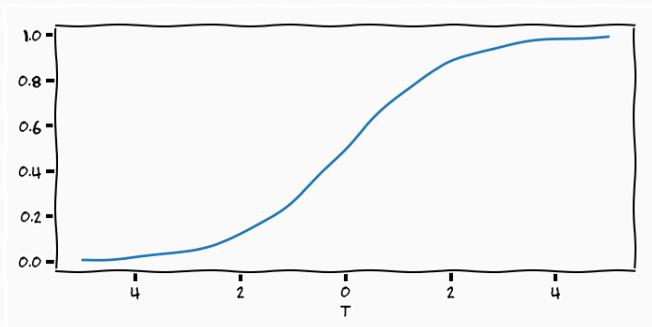
Binary Classification



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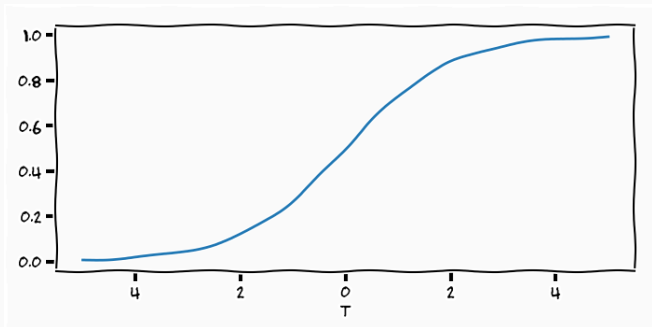
$$\begin{aligned} p(C_1|x) &= \frac{1}{1 + \frac{p(x|C_2)p(C_2)}{p(x|C_1)p(C_1)}} = \frac{1}{1 + \exp\left(\log\left(\frac{p(x|C_2)p(C_2)}{p(x|C_1)p(C_1)}\right)\right)} \\ &= \frac{1}{1 + \exp\left(-\log\left(\frac{p(x|C_1)p(C_1)}{p(x|C_2)p(C_2)}\right)\right)} \end{aligned}$$

Binary Classification



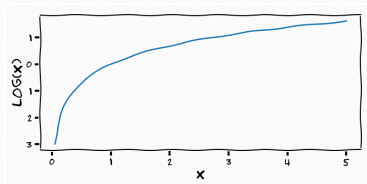
$$t = \log \left(\frac{p(x|C_1)p(C_1)}{p(x|C_2)p(C_2)} \right)$$

Binary Classification

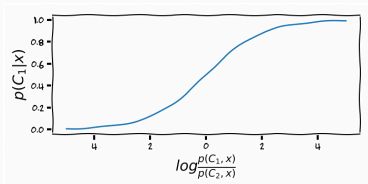


$$t = \log \left(\frac{p(x|C_1)p(C_1)}{p(x|C_2)p(C_2)} \right) = \log \left(\frac{p(x, C_1)}{p(x, C_2)} \right)$$

Binary Classification

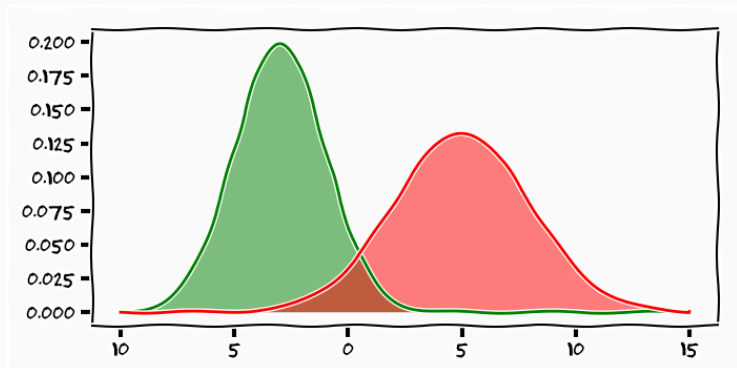


- $p(x, C_1) > p(x, C_2)$
 - $p(C_1|x) > 0.5$
- $p(x, C_1) < p(x, C_2)$
 - $p(C_1|x) < 0.5$
- $p(x, C_1) = p(x, C_2)$
 - $p(C_1|x) = 0.5$



- $p(x, C_1) = 0$
 - $p(C_1|x) = 0$
- $p(x, C_2) \rightarrow 0$
 - $p(C_1|x) \rightarrow 1$
- $p(x, C_1) = p(x, C_2) = 0$
 - Undefined

Likelihoods



- We haven't specified the model yet, let's make a Gaussian likelihood

$$p(x|C_i) = \mathcal{N}(x|\mu_i, \sigma_i) = \frac{1}{\sqrt{2\pi\sigma_i^2}} e^{-\frac{(x-\mu_i)^2}{2\sigma_i^2}}$$

$$t = \log \left(\frac{p(x|C_1)p(C_1)}{p(x|C_2)p(C_2)} \right)$$

$$t = \log \left(\frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x|\mathcal{C}_2)p(\mathcal{C}_2)} \right) = \log \left(\frac{\frac{1}{\sqrt{2\pi\sigma_1^2}} e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}} p(\mathcal{C}_1)}{\frac{1}{\sqrt{2\pi\sigma_2^2}} e^{-\frac{(x-\mu_2)^2}{2\sigma_2^2}} p(\mathcal{C}_2)} \right)$$

$$\begin{aligned} t &= \log \left(\frac{p(x|\mathcal{C}_1)p(\mathcal{C}_1)}{p(x|\mathcal{C}_2)p(\mathcal{C}_2)} \right) = \log \left(\frac{\frac{1}{\sqrt{2\pi\sigma_1^2}} e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}} p(\mathcal{C}_1)}{\frac{1}{\sqrt{2\pi\sigma_2^2}} e^{-\frac{(x-\mu_2)^2}{2\sigma_2^2}} p(\mathcal{C}_2)} \right) \\ &= \log \left(\frac{A}{B} \right) = \log(A) - \log(B) \end{aligned}$$

$$\log(A) = \log \left(\frac{1}{\sqrt{2\pi\sigma_1^2}} e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}} p(\mathcal{C}_1) \right)$$

$$\begin{aligned}\log(A) &= \log\left(\frac{1}{\sqrt{2\pi\sigma_1^2}} e^{-\frac{(x-\mu_1)^2}{2\sigma_1^2}} p(\mathcal{C}_1)\right) = \\ &= -\frac{1}{2}\log(2\pi\sigma_1^2) - \frac{(x-\mu_1)^2}{2\sigma_1^2} + \log(p(\mathcal{C}_1))\end{aligned}$$

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$$\log(A) - \log(B)$$

$$\log(A) - \log(B) = -\frac{1}{2}\log\left(\frac{2\pi\sigma_1^2}{2\pi\sigma_2^2}\right) + \log\left(\frac{p(\mathcal{C}_1)}{p(\mathcal{C}_2)}\right)$$

$$\begin{aligned}\log(A) - \log(B) = & -\frac{1}{2}\log\left(\frac{2\pi\sigma_1^2}{2\pi\sigma_2^2}\right) + \log\left(\frac{p(C_1)}{p(C_2)}\right) \\ & - \frac{x^2 - 2x\mu_1 + \mu_1^2}{2\sigma_1^2} + \frac{x^2 - 2x\mu_2 + \mu_2^2}{2\sigma_2^2}\end{aligned}$$

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$$p(C_1|x) = \frac{1}{1 + e^{-t}}$$

Posterior analysis

- The posterior over \mathcal{C}_1 (and its the same for \mathcal{C}_2) depends on x as

$$-x^2 \left(\frac{1}{2\sigma_1^2} - \frac{1}{2\sigma_2^2} \right) + x \left(\frac{\mu_1}{\sigma_1^2} - \frac{\mu_2}{\sigma_2^2} \right)$$

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- if $\sigma_1 = \sigma_2$

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 - the posterior is **linear** in x

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- if $\sigma_1 = \sigma_2$
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Posterior analysis

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- if $\sigma_1 = \sigma_2$
 - the posterior is **linear** in x
- if $\mu_2 = \mu_1$ and $\sigma_1 = \sigma_2$
 - the posterior does not depend on x

$$p(\mathcal{C}_1|x) = \frac{1}{1 + e^{-\log \frac{p(\mathcal{C}_1)}{p(\mathcal{C}_2)}}} = \frac{p(\mathcal{C}_1)}{p(\mathcal{C}_1) + p(\mathcal{C}_2)} = \frac{p(\mathcal{C}_1)}{p(\mathcal{C})} = p(\mathcal{C}_1)$$

Posterior analysis

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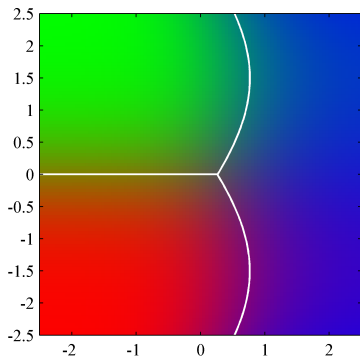
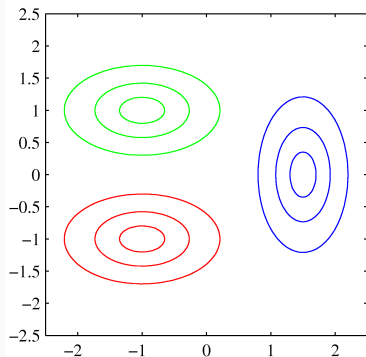
$$-x^2 \left(\frac{1}{2\sigma_1^2} - \frac{1}{2\sigma_2^2} \right) + x \left(\frac{\mu_1}{\sigma_1^2} - \frac{\mu_2}{\sigma_2^2} \right)$$

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$$p(\mathcal{C}_1|x) = \frac{1}{1 + e^{-\log \frac{p(\mathcal{C}_1)}{p(\mathcal{C}_2)}}} = \frac{p(\mathcal{C}_1)}{p(\mathcal{C}_1) + p(\mathcal{C}_2)} = \frac{p(\mathcal{C}_1)}{p(\mathcal{C})} = p(\mathcal{C}_1)$$

- *if the observations does not provide me with any information to update my belief my posterior belief is equal to my prior belief*

Posterior analysis¹



- Red & Green share the same covariance \Rightarrow linear separation
- Blue & (Red & Green) different covariance \Rightarrow curved separation

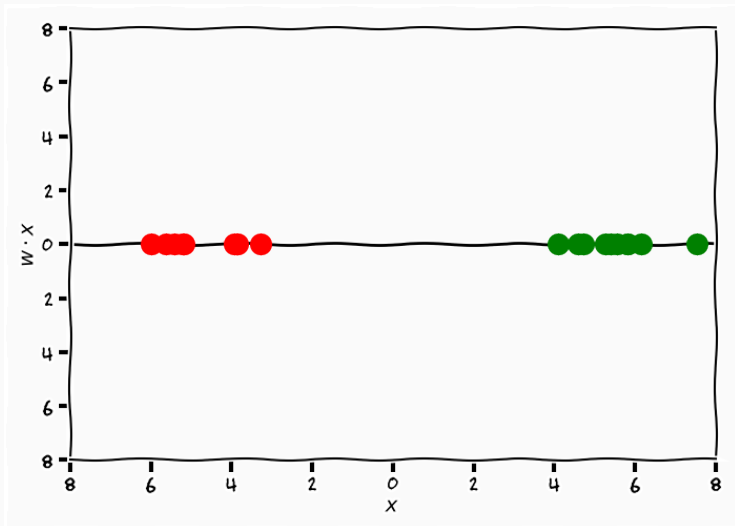
¹Bishop, C. M. (2006). Figure 4.11

Logistic Regression

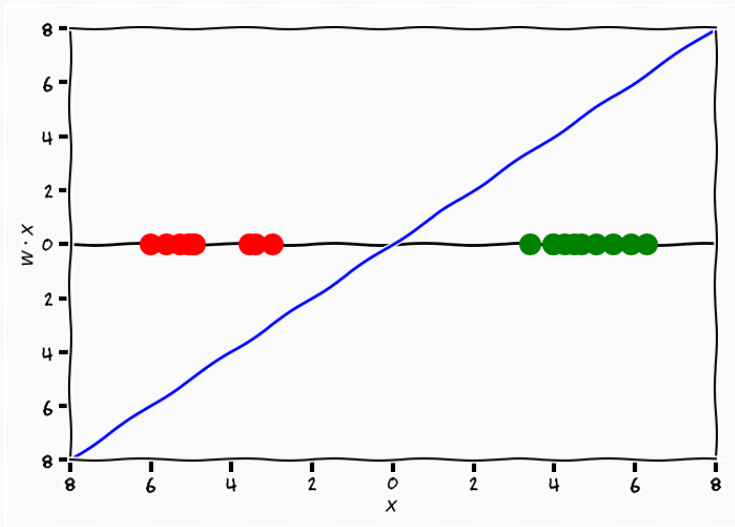
$$p(\mathcal{C}_1|x) = \frac{1}{1 + \exp(-f(x))}$$

- We can derive the posterior through principle by Bayes Rule
 - This forces us to make our assumptions clear
- We can directly parametrise the posterior as we know its form
 - this is called **logistic regression**

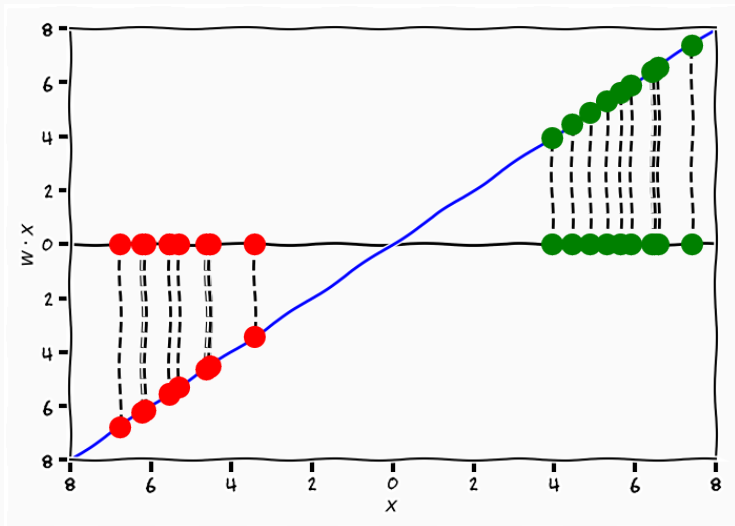
Logistic Regression



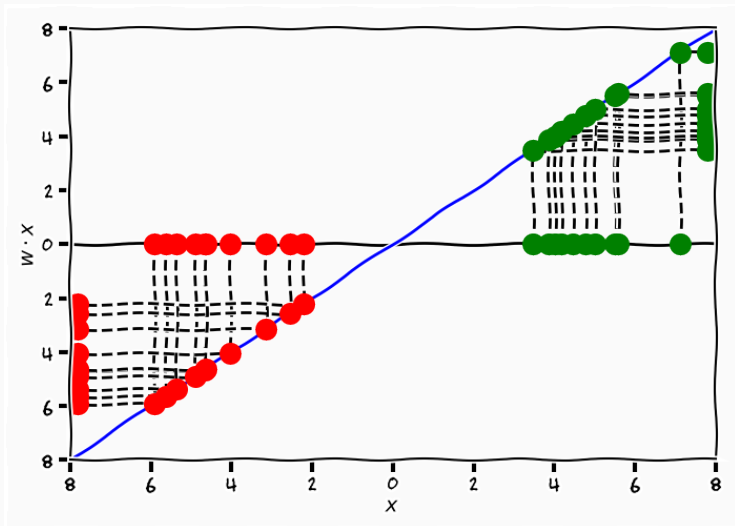
Logistic Regression



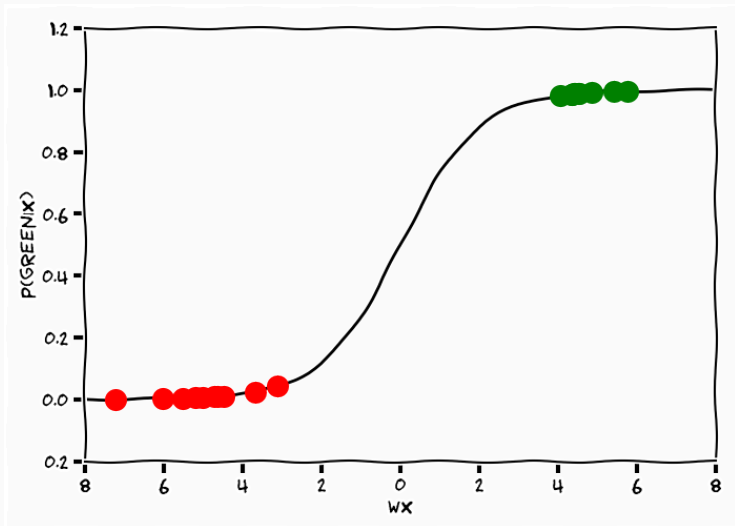
Logistic Regression



Logistic Regression



Logistic Regression



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 - Each Gaussian has

$$\mu_i \in \mathbb{R}^{100}$$

$$\Sigma_i \in \mathbb{R}^{100 \times 100}$$

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- Binary classification implies 20200 parameters

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 - Each Gaussian has

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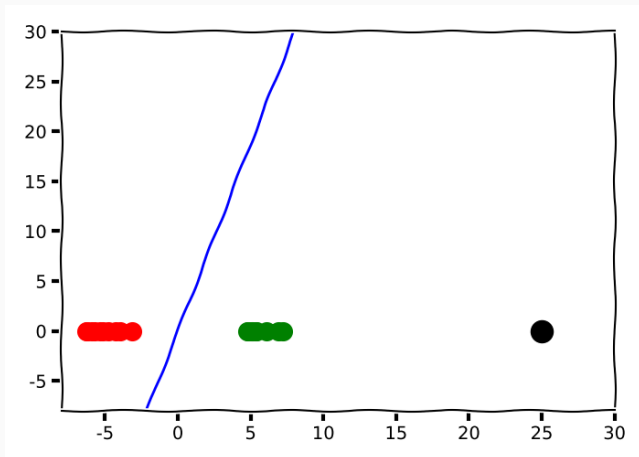
- Binary classification implies 20200 parameters
- Logistic regression implies 101

$$\mathbf{w} \in \mathbb{R}^{100}$$

$$p(C|\mathcal{D})$$

- Reaching this through principle makes it a posterior
- If we parametrise it directly we have to see it as a likelihood
 - we do not **model** \mathcal{D}
 - we do not quantify our uncertainty in \mathcal{D}
 - denominator in Bayes Rule $p(\mathcal{D})$

Why Not



$$t = \log \left(\frac{p(x|C_1)p(C_1)}{p(x|C_2)p(C_2)} \right)$$

$$p(C_1|x) = \frac{1}{1 + \exp(-f(x))} = \sigma(x)$$

- We seek a function that is positive for C_1 and negative for C_2
- Linear classifier

$$f(x) = w^T x$$

- Maximum Likelihood

$$\hat{w} = \operatorname{argmax}_w p(C|\mathcal{D}) = \operatorname{argmin}_w -\log(p(C|\mathcal{D}))$$

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- We have seen how to do linear regression in a principled way
 - specify prior over \mathbf{w}
 - derive posterior

Bayesian Logistic Regression

We want to use the same motivation as we did for normal regression

- Prior

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{m}_0, \mathbf{S}_0)$$

- Likelihood

$$p(c_i | \mathbf{w}, \mathbf{x}_i) = \sigma(\mathbf{w}^T \mathbf{x}_i)$$

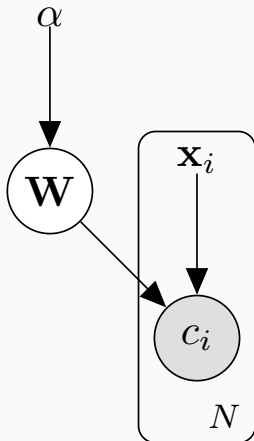
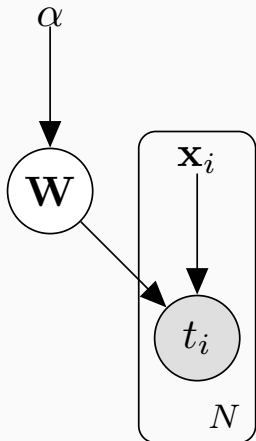
$$p(\mathbf{c} | \mathbf{w}, \mathbf{x}) = \prod_i^N \sigma(\mathbf{w}^T \mathbf{x}_i)^{c_i} \cdot (1 - \sigma(\mathbf{w}^T \mathbf{x}_i))^{1-c_i}$$

You can also do a feature mapping if you want and replace \mathbf{x} with $\Phi(\mathbf{x})$

$$\mathcal{N}(\boldsymbol{\mu}_N, \mathbf{S}_N) \propto \log \left(\prod_i^N \sigma(x_i)^{c_i} \cdot (1 - \sigma(x_i))^{1-c_i} \right) \\ - \frac{1}{2}(\mathbf{w} - \mathbf{m}_0)^T \mathbf{S}_0^{-1}(\mathbf{w} - \mathbf{m}_0) - \log(Z)$$

- Previously we could use conjugacy to reach analytical posterior
- In this case we have changed the likelihood and the Gaussian prior is non-conjugate
- This posterior is intractable

Bayesian Linear Regression



Laplace Approximation

Laplace Approximation

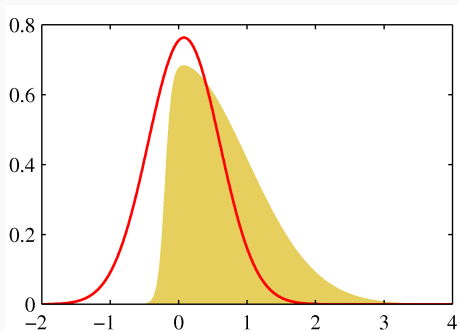
$$p(z|x) = \frac{1}{Z} f(z) = \frac{f(z)}{\int f(z) dz}$$

- Will use $p(z)$ to refer to the posterior
- $p(z)$ is unknown as we cannot compute Z
- $f(z)$ is possible to compute if we have likelihood and prior

$$f(z) = p(x|z)p(z)$$

$$\log p(z) = \log \left(\frac{1}{Z} f(z) \right) = \log(f(z)) + \text{const w.r.t. } z$$

- $p(z)$ and $f(z)$ will have the same modes
 - *is this always true?*
- **Idea**: we can approximate each mode with a distribution we can normalise



- Find the mode of the posterior
- Fit Gaussian to this mode

Taylor Expansion

$$f(x) = f(x_0) + \frac{\partial}{\partial x} f(x_0)(x - x_0) + \frac{1}{2} \frac{\partial^2}{\partial x^2} f(x_0)(x - x_0)^2 + \mathcal{O}((x - x_0)^3)$$

- A Taylor expansion is an approximation of a function around a specific value
- If we expand around a maxima x_0

$$\frac{\partial}{\partial x} f(x_0) = 0$$

- This leads to

$$f(x) = f(x_0) - \frac{1}{2} \left| \frac{\partial^2}{\partial x^2} f(x_0) \right| (x - x_0)^2 + \mathcal{O}((x - x_0)^3)$$

$$f(\mathbf{w}) = p(\mathbf{t}|\mathbf{w})p(\mathbf{w})$$

- we want to find the mode of this, i.e. the maxima

$$\hat{\mathbf{w}} = \operatorname{argmax}_{\mathbf{w}} p(\mathbf{t}|\mathbf{w})p(\mathbf{w})$$

- This we know as the Maximum-a-Posterior (MAP) estimate

Laplace Approximation

1. Find mode of $p(z)$

$$\frac{\partial}{\partial z} p(z_0) = \frac{\partial}{\partial z} f(z_0) = 0$$

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$$\log f(z) \approx \log f(z_0) - \frac{1}{2} \frac{\partial^2}{\partial^2} \log(f(z_0))(z - z_0)^2$$

Laplace Approximation

1. Find mode of $p(z)$

$$\frac{\partial}{\partial z} p(z_0) = \frac{\partial}{\partial z} f(z_0) = 0$$

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3. Take exponential to get function

$$f(z) \approx f(z_0) e^{\underbrace{-\frac{1}{2} \frac{\partial^2}{\partial^2} \log(f(z_0))(z-z_0)^2}_A} = f(z_0) e^{-\frac{1}{2} A (z-z_0)^2}$$

Laplace Approximation

$$f(z) \approx f(z_0)e^{-\frac{1}{2}A(z-z_0)^2}$$

- we want to find an approximation, to $p(z)$ so we need to normalise to a distribution

$$p(z) = \frac{1}{Z}f(z) \approx q(z)$$

Laplace Approximation

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- we want to find an approximation, to $p(z)$ so we need to normalise to a distribution

$$p(z) = \frac{1}{Z}f(z) \approx q(z)$$

- assume that $q(z)$ is Gaussian, i.e. $f(z_0) = p(\text{mean})$

$$q(z) = \left(\frac{A}{2\pi}\right)^{\frac{1}{2}} e^{-\frac{A}{2}(z-z_0)^2}$$

Laplace Approximation

- One dimensional

$$q(z) = \left(\frac{A}{2\pi} \right)^{\frac{1}{2}} e^{-\frac{A}{2}(z-z_0)^2}$$

- D dimensional

$$q(\mathbf{z}) = \frac{|\mathbf{A}|}{(2\pi)^{\frac{D}{2}}} e^{-\frac{1}{2}(\mathbf{z}-\mathbf{z}_0)^T \mathbf{A}(\mathbf{z}-\mathbf{z}_0)} = \mathcal{N}(\mathbf{z}|\mathbf{z}_0, \mathbf{A}^{-1})$$

$$\mathbf{A} = -\nabla \nabla \log f(\mathbf{z})|_{\mathbf{z}=\mathbf{z}_0}$$

- Where \mathbf{A} is the Hessian matrix

Bayesian Logistic Regression

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You can also do a feature mapping if you want and replace \mathbf{x} with $\Phi(\mathbf{x})$

$$q(\mathbf{w}|\mathbf{t}) = \mathcal{N}(\mathbf{m}_N, \mathbf{S}_N) \approx p(\mathbf{w}|\mathbf{t}) \propto p(\mathbf{t}|\mathbf{w})p(\mathbf{w}) = f(\mathbf{w})$$

- Compute $f(\mathbf{w})$

$$\begin{aligned} \log p(\mathbf{w}|\mathbf{t}) = & \log \left(\prod_i^N \sigma(\mathbf{w}^T \mathbf{x}_i)^{t_i} \cdot (1 - \sigma(\mathbf{w}^T \mathbf{x}_i))^{1-t_i} \right) \\ & - \frac{1}{2}(\mathbf{w} - \mathbf{m}_0)^T \mathbf{S}_0^{-1}(\mathbf{w} - \mathbf{m}_0) - \log(Z) + \text{const} \end{aligned}$$

- The stationary point is the MAP estimate

$$\mathbf{S}_N^{-1} = -\nabla \nabla \log p(\mathbf{w}|\mathbf{t})|_{\mathbf{w}=\mathbf{w}_{MAP}} = \mathbf{S}_0^{-1} + \sum_{n=1}^N \sigma(\mathbf{w}^T \mathbf{x}) (1 - \sigma(\mathbf{w}^T \mathbf{x})) \mathbf{x} \mathbf{x}^T$$

- we can compute the Hessian around \mathbf{w}_{MAP}
- this leads to the final approximation

$$q(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{w}_{MAP}, \mathbf{S}_N)$$

Summary

- Compute a mode of the posterior distribution, i.e MAP estimate

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- Compute a mode of the posterior distribution, i.e MAP estimate
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- Identify elements in expansion as parameters of a Gaussian
- Normalise to a distribution
- *You can do exactly the same thing with a GP*

$$\begin{aligned} p(\mathcal{C}_1|\mathbf{x}, \mathbf{t}) &= \int p(\mathcal{C}_1|\mathbf{x}, \mathbf{w})p(\mathbf{w}|\mathbf{t})d\mathbf{w} \\ &\approx \int p(\mathcal{C}_1|\mathbf{x}, \mathbf{w})q(\mathbf{w})d\mathbf{w} \end{aligned}$$

- To compute predictions we can use our new approximate posterior in place of the true posterior

Summary

- Classification often means non-conjugate prior
- Laplace Approximation
 - match modes with the true posterior
- As we often know the MAP estimate of different models we can often apply this method relatively easy
- Can be really bad if the posterior is far from Gaussian

eof

References



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