#### Economics 103 – Statistics for Economists

Francis J. DiTraglia

University of Pennsylvania

Lecture # 12

# Continuous Distributions – Part II

# Last Time: Continuous RVs, Probability As Area

#### Probability Density Function (pdf)

- $f(x) \ge 0$  for all x in the support
- $f(x) \neq P(X = x)$ , can be greater than one

#### Cumulative Distribution Function

- $F(x_0) \equiv P(X \le x_0) = \int_{-\infty}^{x_0} f(x) \ dx$
- ▶ First Fundamental Theorem of Calculus: f(x) = F'(x)

# Last Time: Uniform(0,1) RV

#### Intuition

Equally likely to take on any value on its support: [0,1]

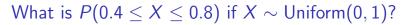
Probability Density Function

f(x) = 1 for  $x \in [0, 1]$ , zero otherwise

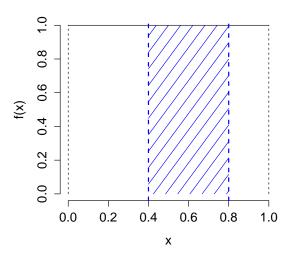
Cumulative Distribution Function

$$F(x_0) = \begin{cases} 0, x_0 < 0 \\ x_0, 0 \le x_0 \le 1 \\ 1, x_0 > 1 \end{cases}$$

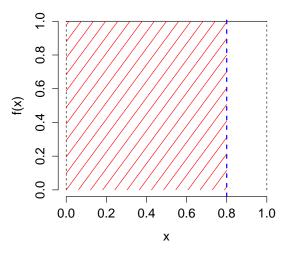
# Key Idea: Probability of Intervals



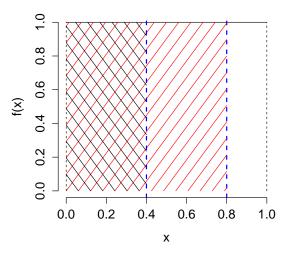




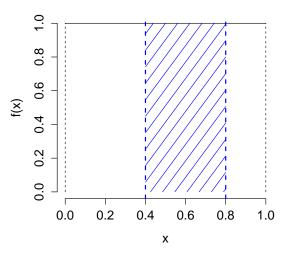
# $F(0.8) = P(X \le 0.8)$



# F(0.8) - F(0.4) = ?



# $F(0.8) - F(0.4) = P(0.4 \le X \le 0.8) = 0.4$



# Probability of Interval for Continuous RV

$$P(a \le X \le b) = \int_a^b f(x) \ dx = F(b) - F(a)$$

This is just the Second Fundamental Theorem of Calculus.

# Expected Value for Continuous RVs

$$\int_{-\infty}^{\infty} x f(x) \ dx$$

Remember: Integrals Replace Sums!

# Example: Uniform(0,1) Random Variable



$$E[X] = \int_{-\infty}^{\infty} xf(x) dx = \int_{0}^{1} x \cdot 1 dx$$
$$= \frac{x^{2}}{2} \Big|_{0}^{1} = 1/2 - 0 = 1/2$$

#### Expected Value of a Function of a Continuous RV

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)f(x) \ dx$$

# Example: Uniform(0,1) RV

$$E[X^{2}] = \int_{-\infty}^{\infty} x^{2} f(x) dx = \int_{0}^{1} x^{2} \cdot 1 dx$$
$$= \frac{x^{3}}{3} \Big|_{0}^{1} = 1/3$$

Once we have defined expected value for continuous RVs, we can use everything we know about variance, covariance, etc. from discrete RVs!

#### Variance of Continuous RV

$$Var(X) = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) \ dx$$

where

$$\mu = E[X] = \int_{-\infty}^{\infty} x f(x) \ dx$$

Shortcut formula still holds for continuous RVs!

$$Var(X) = E[X^2] - (E[X])^2$$

# Example: Uniform(0,1) Random Variable



$$Var(X) = E[(X - E[X])^2] = E[X^2] - (E[X])^2$$
  
= 1/3 - (1/2)^2  
= 1/12  
 $\approx 0.083$ 

# Much More Complicated Without the Shortcut Formula!

$$Var(X) = E\left[ (X - E[X])^2 \right] = \int_{-\infty}^{\infty} (x - \mu)^2 f(x) dx$$

$$= \int_0^1 (x - 1/2)^2 \cdot 1 dx = \int_0^1 (x^2 - x + 1/4) dx$$

$$= \left( \frac{x^3}{3} - \frac{x^2}{2} + \frac{x}{4} \right) \Big|_0^1 = 1/3 - 1/2 + 1/4$$

$$= 4/12 - 6/12 + 3/12 = 1/12$$

# We're Won't Say More About These, But Just So You're Aware of Them...

#### Joint Density

$$P(a \le X \le b \cap c \le Y \le d) = \int_{c}^{d} \int_{a}^{b} f(x, y) \, dxdy$$

#### Marginal Densities

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) \ dy, \qquad f_Y(y) = \int_{-\infty}^{\infty} f(x, y) \ dx$$

Independence in Terms of Joint and Marginal Densities

$$f_{XY}(x, y) = f_X(x)f_Y(y)$$

Conditional Density

$$f_{Y|X} = f_{XY}(x,y)/f_X(x)$$

# We've now covered everything on the

Random Variables Handout

# The Most Important RV of All

#### Normal Random Variable

Notation:  $X \sim N(\mu, \sigma^2)$ 

Parameters:  $\mu = E[X]$ ,  $\sigma^2 = Var(X)$ 

Support:  $(-\infty, +\infty)$ 

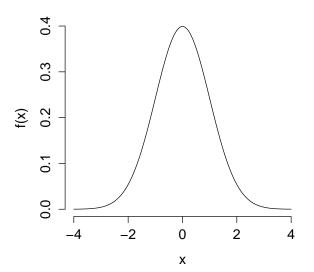
Probability Density Function

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{1}{2} \left(\frac{x-\mu}{\sigma}\right)^2\right\}$$

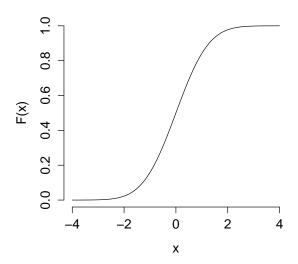
No Explicit Formula for CDF (use computer instead)

$$F(x_0) = \int_{-\infty}^{x_0} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left\{-\frac{1}{2} \left(\frac{x-\mu}{\sigma}\right)^2\right\} dx$$

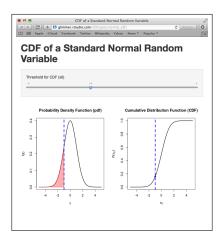
# Normal PDF Centered at the Mean (Here $\mu=0$ , $\sigma^2=1$ )



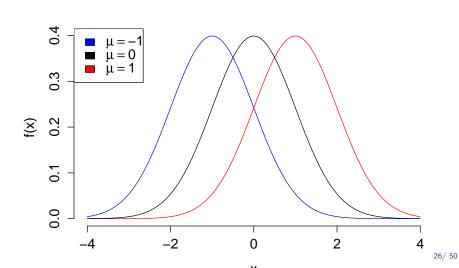
# Normal CDF ( $\mu = 0$ , $\sigma^2 = 1$ )



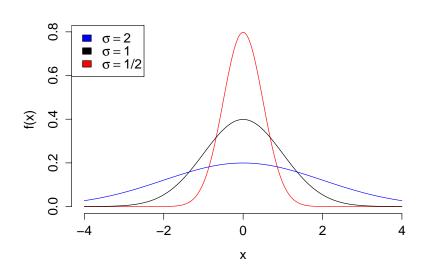
# http://glimmer.rstudio.com/fditraglia/normal\_cdf/



# Different Means, Same Variance



# Same Mean, Different Variances



#### Linear Function of Normal RV is a Normal RV

Suppose that  $X \sim N(\mu, \sigma^2)$ . Then if a and b constants,

$$a + bX \sim N(a + b\mu, b^2\sigma^2)$$

#### **Important**

- Using what know know about expectations of linear functions, no surprise what mean and variance are.
- Surprise is that the linear combination is normal
- Linear trans. does not preserve, e.g., Bernoulli or Binomial.

#### Example



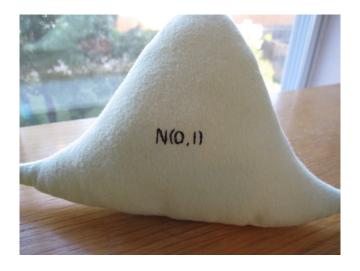
Suppose  $X \sim N(\mu, \sigma^2)$  and let  $Z = (X - \mu)/\sigma$ . What is the distribution of Z?

- (a)  $N(\mu, \sigma^2)$
- (b)  $N(\mu, \sigma)$
- (c)  $N(0, \sigma^2)$
- (d)  $N(0,\sigma)$
- (e) N(0,1)



Figure: Standard Normal Distribution (PDF)

# Standard Normal Distribution: N(0,1)



# Standard Normal Distribution: N(0,1)

Mean = 0, Variance = Standard Deviation <math>= 1

$$f(x) = \frac{1}{\sqrt{2\pi}}e^{-x^2/2}$$

Special symbol for Standard Normal CDF (no closed form):

$$\Phi(x_0) = \int_{-\infty}^{x_0} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \ dx$$

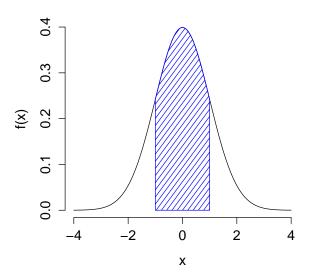
R Command:  $\Phi(x_0) = pnorm()$ 

# Where does the Empirical Rule come from?

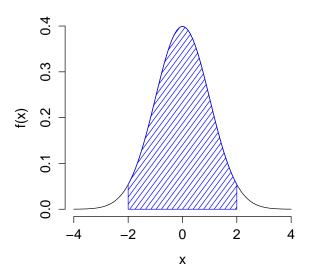
#### **Empirical Rule**

Approximately 68% of observations within  $\mu \pm \sigma$ Approximately 95% of observations within  $\mu \pm 2\sigma$ Nearly all observations within  $\mu \pm 3\sigma$ 

# Middle 68% of $N(0,1) \Rightarrow \text{approx.} (-1,1)$



# Middle 95% of $N(0,1) \Rightarrow \text{approx.} (-2,2)$



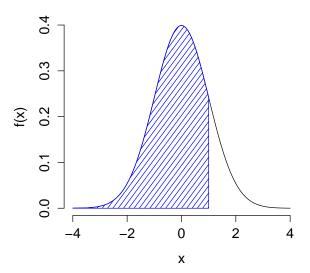
# More Formally...

$$\int_{-1}^{1} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \ dx \approx 0.68$$

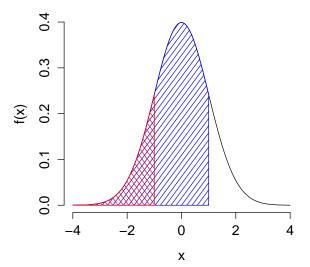
$$\int_{-2}^{2} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \ dx \approx 0.95$$

But how do we know this?

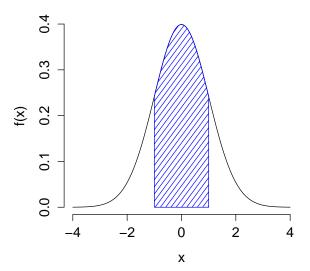
# $\Phi(1) = \texttt{pnorm(1)} \approx 0.84$



$$\Phi(1) - \Phi(-1) = pnorm(1) - pnorm(-1) \approx 0.84 - 0.16$$



$$\Phi(1) - \Phi(-1) = pnorm(1) - pnorm(-1) \approx 0.68$$



#### Suppose $X \sim N(0,1)$

$$P(-2 \le X \le 2) = \Phi(2) - \Phi(-2)$$

$$= pnorm(2) - pnorm(-2)$$

$$\approx 0.95$$

$$P(-3 \le X \le 3) = \Phi(3) - \Phi(-3)$$

$$= pnorm(3) - pnorm(-3)$$

$$\approx 1$$

## What if $X \sim N(\mu, \sigma^2)$ ?

$$P(X \le a) = P(X - \mu \le a - \mu)$$

$$= P\left(\frac{X - \mu}{\sigma} \le \frac{a - \mu}{\sigma}\right)$$

$$= P\left(Z \le \frac{a - \mu}{\sigma}\right)$$

Where Z is a standard normal random variable, i.e. N(0,1).



Which of these equals  $P(Z \le (a - \mu)/\sigma)$  if  $Z \sim N(0, 1)$ ?

- (a)  $\Phi(a)$
- (b)  $1 \Phi(a)$
- (c)  $\Phi(a)/\sigma \mu$
- (d)  $\Phi\left(\frac{\mathsf{a}-\mu}{\sigma}\right)$
- (e) None of the above.

# What if $X \sim N(\mu, \sigma^2)$ ?

$$P(X \le a) = P(X - \mu \le a - \mu)$$

$$= P\left(\frac{X - \mu}{\sigma} \le \frac{a - \mu}{\sigma}\right)$$

$$= P\left(Z \le \frac{a - \mu}{\sigma}\right)$$

$$= \Phi\left(\frac{a - \mu}{\sigma}\right)$$

$$= pnorm((a - \mu)/\sigma)$$

Where Z is a standard normal random variable, i.e. N(0,1).



Which of these is  $P(X \ge b)$ ?

- (a)  $\Phi(b)$
- (b)  $1 \Phi\left(\frac{b-\mu}{\sigma}\right)$
- (c)  $1 \Phi(b)$
- (d)  $1 (\Phi(b)/\sigma \mu)$

$$P(X \ge b) = 1 - P(X \le b) = 1 - P\left(\frac{X - \mu}{\sigma} \le \frac{b - \mu}{\sigma}\right)$$

$$= 1 - P\left(Z \le \frac{b - \mu}{\sigma}\right) = 1 - \Phi\left(\frac{b - \mu}{\sigma}\right)$$

$$= 1 - \operatorname{pnorm}((b - \mu)/\sigma)$$

Where Z is a standard normal random variable.

$$P(a \le X \le b) = P\left(\frac{a-\mu}{\sigma} \le \frac{X-\mu}{\sigma} \le \frac{b-\mu}{\sigma}\right)$$

$$= P\left(\frac{a-\mu}{\sigma} \le Z \le \frac{b-\mu}{\sigma}\right)$$

$$= \Phi\left(\frac{b-\mu}{\sigma}\right) - \Phi\left(\frac{a-\mu}{\sigma}\right)$$

$$= pnorm((b-\mu)/\sigma) - pnorm((a-\mu)/\sigma)$$

Where Z is a standard normal random variable.



What is 
$$P(\mu - \sigma \le X \le \mu + \sigma)$$
?

$$P(\mu - \sigma \le X \le \mu + \sigma) = P\left(-1 \le \frac{X - \mu}{\sigma} \le 1\right)$$

$$= P(-1 \le Z \le 1)$$

$$= \Phi(1) - \Phi(-1)$$

$$= pnorm(1) - pnorm(-1)$$

$$\approx 0.68$$



What is 
$$P(\mu - 2\sigma \le X \le \mu + 2\sigma)$$
?

$$P(\mu - 2\sigma \le X \le \mu + 2\sigma) = P\left(-2 \le \frac{X - \mu}{\sigma} \le 2\right)$$

$$= P\left(-2 \le Z \le 2\right)$$

$$= \Phi(2) - \Phi(-2)$$

$$= pnorm(2) - pnorm(-2)$$

$$\approx 0.95$$