FINAL EXAMINATION ECON 103, STATISTICS FOR ECONOMISTS

May 7th, 2014

You will have 120 minutes to complete this exam. Graphing calculators, notes, and textbooks are not permitted.

I pledge that, in taking and preparing for this exam, I have abided by the University of Pennsylvania's Code of Academic Integrity. I am aware that any violations of the code will result in a failing grade for this course.

Name:			
Student ID #: -			
Signature:			

Question:	1	2	3	4	5	6	Total
Points:	30	20	45	45	20	40	200
Score:							

Instructions: Answer all questions in the space provided, continuing on the back of the page if you run out of space. Show your work for full credit but be aware that writing down irrelevant information will not gain you points. Be sure to sign the academic integrity statement above and to write your name and student ID number on *each page* in the space provided. Make sure that you have all pages of the exam before starting.

Warning: If you continue writing after we call time, even if this is only to fill in your name, twenty-five points will be deducted from your final score. In addition, two points will be deducted for each page on which you do not write your name and student ID.

- 1. Let $X_1, X_2, X_3 \stackrel{iid}{\sim} \text{Bernoulli}(1/2)$ and define $Y = X_1 + X_2$ and $Z = X_2 \cdot X_3$.
 - (a) (5 points) What is the support of Y? What is its pmf? Is it a "named" random variable? If so, which one and what are its parameters? Explain briefly.

(b) (5 points) Calculate $E[(Y-1)^2]$.

(c) (5 points) What is the support of Z? What is its pmf? Is it a "named" random variable? If so, which one and what are its parameters? Explain briefly.

(d) (15 points) Express the joint pmf of Y and Z in tabular form. Place the realizations for Y in the rows of the table and the realizations for Z in the columns.

Name: _____

Student ID #: _____

- 2. For each statement, indicate whether it is TRUE or FALSE and briefly explain why.
 - (a) (4 points) For any two RVs X and Y, E[XY] = Cov(X, Y) E[X]E[Y].

(b) (4 points) If $Z \sim N(0,1)$ then P(Z=0) > P(Z=5).

(c) (4 points) If I run the command $x \leftarrow rnorm(1)$ at the R console followed by replicate(100, x) the result will be 100 iid N(0,1) random draws.

(d) (4 points) I rejected the null hypothesis H_0 : $\mu = 5$ against the two-sided alternative H_1 : $\mu \neq 5$ with $\alpha = 0.05$. Based on this result, I know that if I were to construct a 95% confidence interval for μ it would not contain zero.

(e) (4 points) The p-value for my test was 0.03. This means that if I had set $\alpha = 0.05$ I would have rejected the null hypothesis.

Name: _____

Student ID #: _____

3. This question is based on a dataframe called **profs** containing the salaries of a random sample of 397 professors at US universities in 2008. Here are the first few rows:

> head(profs)

	rank	sex	salary
1	Prof	Male	140
2	Prof	Male	173
3	AsstProf	Male	80
4	Prof	Male	115
5	Prof	Male	142
6	AssocProf	Male	97

Each row contains data for a single professor. The column rank is a categorical variable indicating a professor's rank: Prof stands for "Full Professor," while AssocProf and AsstProf stand for "Associate" and "Assistant" Professor, respectively. The column sex indicates a professor's sex, Male or Female, while salary gives her salary measured in thousands of US dollars. There are no missing observations.

(a) (10 points) Write R code to carry out the following operations *separately* for each sex: count the number of professors, calculate the sample mean of salary, and calculate the sample variance of salary.

(b) (10 points) The results of the preceding part are summarized in the following table

	Female	Male
Sample Size	39	358
Sample Mean	101	115
Sample Variance	676	926

Use this information to construct an approximate 95% confidence interval for the difference of population mean salaries for male and female professors in the US.

(c) (10 points) Suppose that, rather than constructing a confidence interval, we wanted to carry out a two-sided test of the null hypothesis that the population mean salary is equal across male and female professors. What is the value of your test statistic? What R command would you use to calculate the p-value? Approximately what value would you get if you ran this command?

Name: _____ Student ID #: ____

(d) (15 points) The following table *breaks down* the results of part (b) by rank. It also provides approximate 95% CIs for the difference of means (Male - Female):

	Sample	Size	Sample Mean		Sample Var.		Approx. 95% CI	
	Female	Male	Female	Male	Female	Male	(Male - Female)	
AsstProf	11	56	78	81	88	63	3 ± 6	
AssocProf	10	54	89	95	327	165	6 ± 12	
Prof	18	248	122	127	388	796	5 ± 10	

Explain these results. How to they compare to your confidence interval and p-value from parts (b) and (c)? If you find differences, can you propose an explanation? What overall conclusions, if any, can we draw from this dataset?

Name:	Student ID #:	

4. In the basement of the McNeil building I have a storeroom containing thousands of digital scales. I know from past experience that some scales from this manufacturer are defective: they give biased readings. All of the scales, both biased and unbiased, are imperfect in that they make small, independent, normally distributed errors with a standard deviation of two grams. This means that if you weigh the same object repeatedly, you will get a slightly different result each time and the average deviation from the mean will be about two grams. For unbiased scales, the errors have zero mean: if you weigh the same object a very large number of times on an unbiased scale, the average weight will equal the true weight. For biased scales, on the other hand, the errors have mean one gram: if you weigh the same object a very large number of times on a a biased scale, the average weight will be one gram higher than the true weight. For each scale in my storeroom, I carry out the following procedure. First I weigh the same 10 gram mass repeatedly and record the results. In total, I make 16 measurements. These constitute a random sample from the following population:

$$X_1, \dots, X_{16} \stackrel{iid}{\sim} N(\mu, \sigma^2 = 4)$$

If the scale I'm testing is unbiased, then $\mu = 10$, the true weight. If the scale is biased then $\mu = 11$, one gram higher than the true weight. I then carry out an exact test of the null hypothesis H_0 : $\mu = 10$ against the one-sided alternative H_1 : $\mu > 10$ with $\alpha = 0.025$. If I reject the null, I decide that the scale is defective and throw it away. Otherwise I keep it for use in the candy weighing experiment in Econ 103.

(a) (10 points) What formula should I use for my test statistic? Approximately what is my critical value? What is my decision rule?

(b) (5 points) Suppose I'm testing one of the scales and, unbeknownst to me, it happens to be one of the *unbiased* ones. What is the probability that I will reject the null? Explain your answer.

Vame:	Student ID #:

(c)	(15 points)	Suppose I'	m testing on	e of the scales	and, unbek	nownst	to me, it	hap-
	pens to be	one of the	$biased \ {\rm ones}.$	What is the	probability	that I	will reject	t the
	null? Expla	in vour ans	swer.					

(d) (5 points) If I wanted to increase the probability that I will successfully detect biased scales, how should I change my testing procedure? Explain.

Name: ______ Student ID #: _____

(e)	(5 points) From past experience, I know that roughly 10% of scales produced by
	this manufacturer are biased. If I test a large number of scales, approximately what
	percentage of them will I end up throwing away? Explain.

(f) (5 points) Continuing from the previous part, among the scales that I end up throwing away, approximately what percentage of them will *actually* be biased? Explain.

5. (20 points) Suppose that $X_1, \ldots, X_n \stackrel{iid}{\sim} \text{Bernoulli}(p)$, and that n is sufficiently large that the sampling distribution of \widehat{p} is approximately normal by the CLT. Write an R function called prop.test that returns the p-value for the refined test of H_0 : $p = p_0$ against H_1 : $p \neq p_0$. Your function should take two arguments. The first, \mathbf{x} , is the sample data: a vector of zeros and ones corresponding to the realizations x_1, \ldots, x_n of X_1, \ldots, X_n . The second is $\mathbf{p}.0$, a value between zero and one that indicates which null hypothesis the user wishes to test. For example if $\mathbf{p}.0 = 0.5$ then we are testing H_0 : p = 0.5. You may assume that there are no missing values in \mathbf{x} .

Name: Studen	t ID #:
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6. Garth is concerned that students registered for Monday recitation sections of Econ 103 may have an unfair advantage on quizzes because they have additional days to study compared to students in Friday sections. To decide whether he should adjust the course curve to compensate, he fits a number of linear regression models, the results of which appear in Table 1. Each regression is calculated using a dataframe called gradebook:

	quiz.avg	${\tt midterm.avg}$	Monday
1	71	85.8	1
2	80	74.6	1
82	2 58	66.3	0
83	3 73	96.1	1

(To preserve privacy the preceding values are fake but the regression results given below are based on the real course gradebook.) Each row of gradebook corresponds to a student in Econ 103. The column quiz.avg contains that student's overall quiz average in percentage points, while midterm.avg contains the average of her scores on the two midterms, also in percentage points. Finally, Monday is a dummy variable that takes on the value one for students in Monday recitation sections, zero otherwise.

(a) (5 points) Which set of regression results allows us to determine the sample mean of quiz.avg for Monday and Friday students students separately? What is the sample mean of each group? What is the difference of means?

(b) (5 points) Continuing from the previous part, construct a 95% confidence interval for the difference of means between Monday and Friday students. Is the difference statistically significant at the 5% level? Explain briefly.

Name:	Student ID #:

(c) (5 points) What is the sample correlation between quiz.avg and midterm.avg? Approximately how accurately does a simple linear regression model with the predictor midterm.avg alone predict quiz.avg?

(d) (10 points) Garth fears that confounders could be responsible for the difference in average quiz scores between Monday and Friday recitation sections. For example, if Differential Equations meets on Friday from 10-11:30am, then students taking this course will be forced to register for a Monday recitation section of Econ 103. Since taking difficult math courses is likely correlated with greater mathematical ability and hence grades in Econ 103, this could bias the comparison. Accordingly, Garth decides to control for differences in students' ability using midterm.avg, fitting a regression in which quiz.avg is linearly related to midterm.avg but the intercept is allowed to differ for Monday and Friday students. The slope is not allowed to vary across groups. Which of the four regression models given below contains the results? Explain the fitted relationships given by the model and briefly discuss the results in light of Garth's concerns.

Name: ______ Student ID #: _____

(e) (15 points) Before making a final decision about whether to adjust the recitation quiz scores, Garth decides to try one more thing. Rather than comparing Monday recitations to Friday recitations, he calculates the sample mean of quiz.avg and the associated standard error for *each* recitation section:

FridayA	${ t Monday A}$	${ t Monday B}$	FridayB	${\tt MondayC}$
60.25	66.33	67.73	68.50	71.17
(2.4)	(2.6)	(2.3)	(4.3)	(2.2)

The standard errors appear in parentheses. (The letter designations are arbitrary: for privacy reasons I don't want to reveal precisely which recitation section is which.) Based on this new information, the full set of regression results, and your answers to the preceding parts, do you think Garth should adjust quiz.avg upwards for students registered for Friday recitation sections? If so, by how much? Use what you've learned in Econ 103 to make your case.

Name:	Student ID #:	

Table 1: Regression Results

Regression 1:

```
lm(formula = quiz.avg ~ Monday + midterm.avg + Monday:midterm.avg,
   data = gradebook)
                  coef.est coef.se
(Intercept)
                  39.19
                           10.39
                           12.48
Monday
                   4.46
midterm.avg
                   0.36
                            0.15
Monday:midterm.avg -0.01
                            0.18
---
n = 83, k = 4
residual sd = 10.15, R-Squared = 0.23
Regression 2:
lm(formula = quiz.avg ~ Monday, data = gradebook)
           coef.est coef.se
(Intercept) 63.77
                     2.19
Monday
           4.86
                     2.65
___
n = 83, k = 2
```

Regression 3:

residual sd = 11.19, R-Squared = 0.04

Regression 4:

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