Economics 103 – Statistics for Economists

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Lecture 15

Sampling Distributions and Estimation – Part II

Unbiased means "Right on Average"

Bias of an Estimator

Let $\widehat{\theta}_n$ be a sample estimator of a population parameter θ_0 . The bias of $\widehat{\theta}_n$ is $E[\widehat{\theta}_n] - \theta_0$.

Unbiased Estimator

A sample estimator $\widehat{\theta}_n$ of a population parameter θ_0 is called unbiased if $E[\widehat{\theta}_n] = \theta_0$

$$\widehat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n \left(X_i - \bar{X} \right)^2$$

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Making a habit of playing the lottery. Expectation is negative.

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Procedure = Random Variable

Making a habit of playing the lottery. Expectation is negative.

Result of that Procedure = Constant

How much you win in a *particular* lottery. Could be greater than or less than cost of ticket in any *individual* instance.

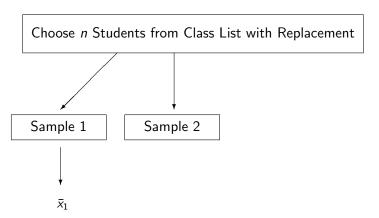
Choose n Students from Class List with Replacement

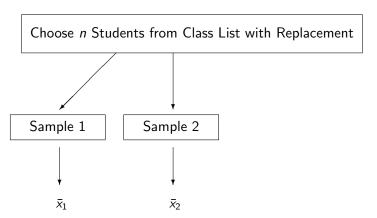
Choose *n* Students from Class List with Replacement

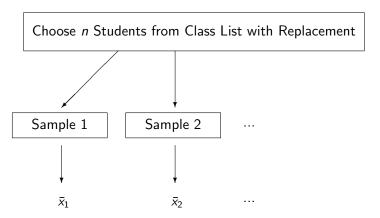
 \bar{x}_1

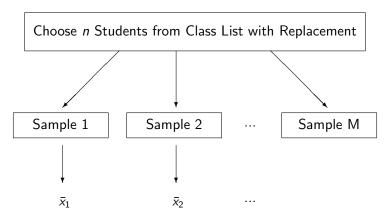
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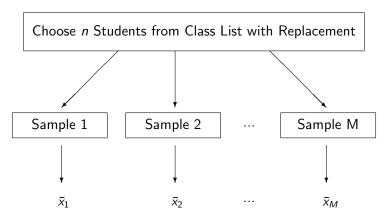
Sample 1

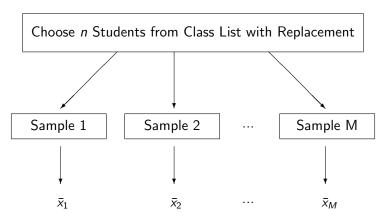




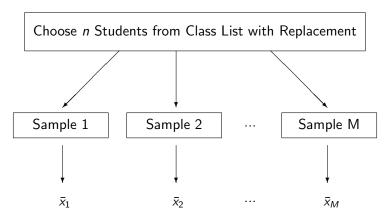








Repeat M times \rightarrow get M different sample means



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Sampling Dist: long run relative frequencies of the \bar{x}_i

Height of Econ 103 Students

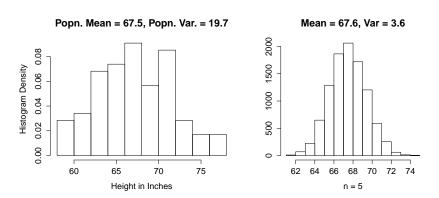


Figure : Left: Population, Right: Sampling distribution of X_5

How Large is the Average Family?



How many brothers and sisters are in your family, including yourself?

The average number of children per family was about 2.0 twenty years ago.

Biased Sample!

Biased Sample!

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Biased Sample!

- ► Zero children ⇒ didn't send any to college
- Sampling by children so large families oversampled

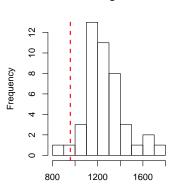
Candy Weighing: 44 Estimates, Each With n = 5

$\widehat{\theta} = 20 \times$	(X ₁	+		+	<i>χ</i> ₋ `	١
0 - 20 X	(1	Τ.	٠.	\top	^ 5	ı

Summary of Sampling Dist.		
Overestimates	42	
Exactly Correct	0	
Underestimates	2	
$E[\hat{\theta}]$	1269 grams	
$SD(\widehat{ heta})$	189 grams	

Actual Mass: $\theta_0 = 958$ grams

Histogram



Est. Weight of All Candies (grams)

100 Candies Total:

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20 Fun Size Snickers Bars (large)

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So What Happened?

Not a random sample! The Snickers bars were oversampled.

Could we have avoided this? How?



Let $X_1, X_2, \dots X_n \sim iid$ mean μ , variance σ^2 and define $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$. True or False:

 $ar{X}_n$ is an unbiased estimator of μ

- (a) True
- (b) False



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TRUE!



Let $X_1, X_2, \dots X_n \sim iid$ mean μ , variance σ^2 . True or False:

 X_1 is an unbiased estimator of μ

- (a) True
- (b) False



Let $X_1, X_2, ... X_n \sim iid$ mean μ , variance σ^2 . True or False:

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- (a) True
- (b) False

TRUE!

$$E[\bar{X}_n] = E\left[\frac{1}{n}\sum_{i=1}^n X_i\right] = \frac{1}{n}\sum_{i=1}^n E[X_i] = \mu$$

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$$Var(\bar{X}_n) = Var\left(\frac{1}{n}\sum_{i=1}^n X_i\right) = \frac{1}{n^2}\sum_{i=1}^n Var(X_i) = \frac{\sigma^2}{n}$$

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$$Var(X_1) = \sigma^2$$

Efficiency - Compare Unbiased Estimators by Variance

Let $\widehat{\theta}_1$ and $\widehat{\theta}_2$ be unbiased estimators of θ_0 . We say that $\widehat{\theta}_1$ is *more* efficient than $\widehat{\theta}_2$ if $Var(\widehat{\theta}_1) < Var(\widehat{\theta}_2)$.

Except in very simple situations, unbiased estimators are hard to come by. In fact, in many interesting applications there is a *tradeoff* between bias and variance:

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Mean Squared Error (MSE):

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Mean Squared Error (MSE): Squared Bias plus Variance

$$MSE(\widehat{\theta}) = Bias(\widehat{\theta})^2 + Var(\widehat{\theta})$$

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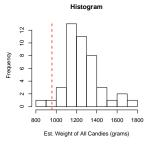
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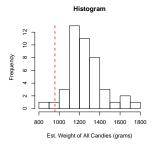
$$MSE(\widehat{\theta}) = Bias(\widehat{\theta})^2 + Var(\widehat{\theta})$$

= $(E[\widehat{\theta}] - \theta_0)^2 + Var(\widehat{\theta})$

$E[\hat{\theta}]$	1269 grams
$ heta_{0}$	958 grams
$SD(\widehat{ heta})$	189 grams

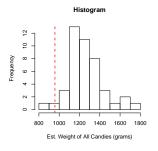


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$$\begin{array}{lll} \mathsf{Bias} & = & 1269 \; \mathsf{grams} - 958 \; \mathsf{grams} \\ & = & & & & & & & \\ \end{array}$$

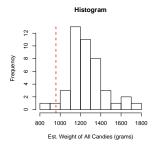
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Bias =
$$1269 \text{ grams} - 958 \text{ grams}$$

= 311 grams
MSE =

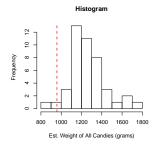
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Bias =
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= 311 grams
MSE = $Bias^2 + Variance$
=

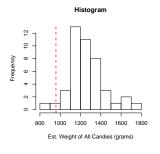
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MSE = $Bias^2 + Variance$
= $(311^2 + 189^2) \text{ grams}^2$
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MSE = $Bias^2 + Variance$
= $(311^2 + 189^2) \text{ grams}^2$
= $1.3244 \times 10^5 \text{ grams}^2$

	Histogram
Frequency	6 - 5 - 5 - 5 - 5 - 5 - 5 - 5 - 5 - 5 -
	800 1000 1200 1400 1600 1800
	Est. Weight of All Candies (grams)

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Bias = 1269 grams - 958 grams
= 311 grams
MSE = Bias² + Variance
=
$$(311^2 + 189^2)$$
 grams²
= 1.3244×10^5 grams²
RMSE = $\sqrt{\text{MSE}} = 364$ grams

Finite Sample versus Asymptotic Properties of Estimators

Finite Sample Properties

For *fixed sample size n* what are the properties of the sampling distribution of $\widehat{\theta}_n$? (E.g. bias and variance.)

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For *fixed sample size n* what are the properties of the sampling distribution of $\widehat{\theta}_n$? (E.g. bias and variance.)

Asymptotic Properties

What happens to the sampling distribution of $\widehat{\theta}_n$ as the sample size n gets larger and larger? (That is, $n \to \infty$).

Why Asymptotics?

Law of Large Numbers

Make precise what we mean by "bigger samples are better."

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Central Limit Theorem

As $n \to \infty$ nearly all sampling distributions behave behave like a normal random variable!

Consistency

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If an estimator $\widehat{\theta}_n$ (which is a RV) converges to θ_0 (a constant) as $n \to \infty$, we say that $\widehat{\theta}_n$ is consistent for θ_0 .

What does it mean for a RV to converge to a constant?

For this course we'll use *MSE Consistency*:

$$\lim_{n\to\infty}\mathsf{MSE}(\widehat{\theta}_n)=0$$

This makes sense since $MSE(\widehat{\theta}_n)$ is a *constant*, so this is just an ordinary limit from calculus.

Let $X_1, X_2, \dots X_n \sim iid$ mean μ , variance σ^2 . Then the sample mean

$$\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$$

is consistent for the population mean μ .

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$$= 0 + \sigma^2/n$$

$$\to 0$$

Hence \bar{X}_n is consistent for μ

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Suppose $X_1, X_2, \dots, X_n \sim \text{iid } N(\mu, \sigma^2)$. What is the sampling distribution of \bar{X}_n ?

- (a) $\chi^2(n)$
- (b) *t*(*n*)
- (c) F(n,n)
- (d) $N(\mu, \sigma^2/n)$
- (e) Not enough information to determine.

But still, how can something random converge to something constant?

Sampling Distribution of \bar{X}_n Collapses to μ

Look at an example where we can directly calculate not only the mean and variance of the sampling distribution of \bar{X}_n , but the sampling distribution itself:

$$X_1, X_2, \ldots, X_n \sim \text{iid } N(\mu, \sigma^2) \Rightarrow \bar{X}_n \sim N(\mu, \sigma^2/n)$$

Sampling Distribution of \bar{X}_n Collapses to μ

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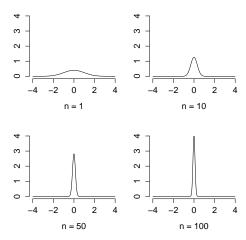


Figure : Sampling Distributions for \bar{X}_n where $X_i \sim \text{iid } N(0,1)$

Another Visualization: Keep Adding Observations

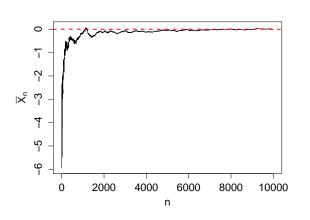


Figure : Running sample means: $X_i \sim \text{iid } N(0, 100)$

n	\bar{X}_n
1	-2.69
2	-3.18
3	-5.94
4	-4.27
5	-2.62
10	-2.89
20	-5.33
50	-2.94
100	-1.58
500	-0.45
1000	-0.13
5000	-0.05
10000	0.00
10000	0.00

Although I showed two examples involving normal RVs, the LLN holds IN GENERAL!