#### Economics 103 – Statistics for Economists

Francis J. DiTraglia

University of Pennsylvania

Lecture 18

#### Two-sample Problem



Suppose  $X_1, \ldots, X_n \sim \text{iid } N(\mu_x, \sigma_x^2)$  independently of  $Y_1, \ldots, Y_m \sim \text{iid } N(\mu_y, \sigma_y^2)$ . What is  $E[\bar{X}_n - \bar{Y}_m]$ , the expectation of the sampling distribution of the difference of sample means?

- (a)  $\mu_X$
- (b)  $\mu_{x} \mu_{y}$
- (c)  $\mu_y$
- (d)  $\mu_x + \mu_y$
- (e) 0

$$E[\bar{X}_n - \bar{Y}_m] = E[\bar{X}_n] - E[\bar{Y}_m] = \mu_x - \mu_y$$

#### Two-sample Problem



Suppose  $X_1, \ldots, X_n \sim \text{iid } N(\mu_x, \sigma_x^2)$  independently of  $Y_1, \ldots, Y_m \sim \text{iid } N(\mu_y, \sigma_y^2)$ . What is  $Var[\bar{X}_n - \bar{Y}_m]$ , the variance of the sampling distribution of the difference of sample means?

- (a)  $\sigma_x^2 \sigma_y^2$
- (b)  $\sigma_x^2 + \sigma_y^2$
- (c)  $\sigma_x^2/n + \sigma_y^2/m$
- (d)  $\sigma_x^2/n \sigma_y^2/m$
- (e) 1

By independence:  $Var[\bar{X}_n - \bar{Y}_m] = Var[\bar{X}_n] + Var[\bar{Y}_m] = \frac{\sigma_x^2}{n} + \frac{\sigma_y^2}{m}$ 

#### Two-sample Problem



Suppose  $X_1, \ldots, X_n \sim \text{iid } N(\mu_x, \sigma_x^2)$  independently of  $Y_1, \ldots, Y_m \sim \text{iid } N(\mu_y, \sigma_y^2)$ . What is the sampling distribution of  $\bar{X}_n - \bar{Y}_m$ , the difference of sample means?

- (a)  $\chi^2$
- (b) t
- (c) F
- (d) Normal

Normal, by independence and linearity property of normal distributions.

## Sampling Distribution of $\bar{X}_n - \bar{Y}_m$

Suppose  $X_1, \ldots, X_n \sim \text{iid } N(\mu_x, \sigma_x^2)$  independently of  $Y_1, \ldots, Y_m \sim \text{iid } N(\mu_y, \sigma_y^2)$ . Then,

$$(\bar{X}_n - \bar{Y}_m) \sim N\left(\mu_x - \mu_y, \frac{\sigma_x^2}{n} + \frac{\sigma_y^2}{m}\right)$$

$$rac{\left(ar{X}_n - ar{Y}_m
ight) - \left(\mu_{\mathsf{x}} - \mu_{\mathsf{y}}
ight)}{\sqrt{rac{\sigma_{\mathsf{x}}^2}{n} + rac{\sigma_{\mathsf{y}}^2}{m}}} \sim \mathit{N}(0,1)$$

Shorthand: 
$$SE(\bar{X}_n - \bar{Y}_m) = \sqrt{\frac{\sigma_x^2}{n} + \frac{\sigma_y^2}{m}}$$

## CI for Difference of Population Means, $\sigma_x^2, \sigma_y^2$ Known

$$\frac{(\bar{X}_n - \bar{Y}_m) - (\mu_{\scriptscriptstyle X} - \mu_{\scriptscriptstyle Y})}{\mathit{SE}(\bar{X}_n - \bar{Y}_m)} \sim \mathit{N}(0, 1)$$

Thus, we construct a  $100 \times (1 - \alpha)\%$  CI for  $\mu_x - \mu_y$  as follows:

$$(ar{X}_{\it n} - ar{Y}_{\it m}) \pm \ {
m qnorm}(1 - lpha/2) \ {\it SE}(ar{X}_{\it n} - ar{Y}_{\it m})$$

Where 
$$SE(\bar{X}_n - \bar{Y}_m) = \sqrt{\frac{\sigma_x^2}{n} + \frac{\sigma_y^2}{m}}$$

#### Calculate the ME for the Difference of Means



I generated independent random samples of size 25 from two normal distributions in R. One had a population standard deviation of 4 and the other had a population standard deviation of 3. The sample means were approximately 4.2 and 3.1.

Calculate the ME for a 95% confidence interval for the difference of population means.

$$SE = \sqrt{\frac{3^2}{25} + \frac{4^2}{25}} = \frac{\sqrt{9+16}}{5} = 1$$

$$ME = qnorm(1 - 0.05/2) \times SE \approx 2 \times SE = 2$$

#### Calculate the LCL for the Difference of Means



I generated independent random samples of size 25 from two normal distributions in R. One had a population standard deviation of 4 and the other had a population standard deviation of 3. The sample means were approximately 4.2 and 3.1.

Calculate the LCL for a 95% confidence interval for the difference of population means.

$$LCL = (4.2 - 3.1) - ME = 1.1 - 2 = -0.9$$

#### Calculate the UCL for the Difference of Means



I generated independent random samples of size 25 from two normal distributions in R. One had a population standard deviation of 4 and the other had a population standard deviation of 3. The sample means were approximately 4.2 and 3.1.

Calculate the UCL for a 95% confidence interval for the difference of population means.

$$UCL = (4.2 - 3.1) + ME = 1.1 + 2 = 3.1$$

95% Confidence Interval: (-0.9, 3.1)

The actual population means were 4 and 3, respectively

## What if $\sigma_x^2, \sigma_y^2$ are Unknown?

Suppose  $X_1, \ldots, X_n \sim \text{iid } N(\mu_x, \sigma_x^2)$  independently of  $Y_1, \ldots, Y_m \sim \text{iid } N(\mu_y, \sigma_y^2)$ . Then,

$$rac{\left(ar{X}_n - ar{Y}_m
ight) - \left(\mu_{\mathsf{X}} - \mu_{\mathsf{y}}
ight)}{\sqrt{rac{S_{\mathsf{X}}^2}{n} + rac{S_{\mathsf{y}}^2}{m}}} \sim t(
u)$$

#### Formula for $\nu$ is Complicated and You Don't Need to Know it

#### Two possibilities:

- 1. Have R find the correct value of  $\nu$  for us
- 2. If m, n are large enough, approximately standard normal.

## Case of Equal, Unknown Variances

The book considers a case where  $\sigma_x^2 = \sigma_y^2 = \sigma^2$ , that is a common unknown variance. This is a very dangerous assumption. It is almost certainly false and can throw off our results in a serious way. You are not responsible for this case.

## Sampling Distributions Under Normality: One-sample

Suppose that  $X_1, \ldots, X_n \sim \text{iid } N(\mu, \sigma^2)$ . Then:

$$\left(\frac{n-1}{\sigma^2}\right)S^2 \sim \chi^2(n-1)$$

$$\frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \sim N(0,1)$$

$$\frac{\bar{X}_n - \mu}{S/\sqrt{n}} \sim t(n-1)$$

## Sampling Distributions Under Normality: Two-sample

Suppose  $X_1, \ldots, X_n \sim \text{iid } N(\mu_x, \sigma_x^2)$  independently of  $Y_1, \ldots, Y_m \sim \text{iid } N(\mu_y, \sigma_y^2)$ . Then:

$$\frac{(\bar{X}_n - \bar{Y}_n) - (\mu_x - \mu_y)}{\sqrt{\frac{\sigma_x^2}{n} + \frac{\sigma_y^2}{m}}} \sim N(0, 1)$$

$$\frac{\left(\bar{X}_{n}-\bar{Y}_{m}\right)-\left(\mu_{x}-\mu_{y}\right)}{\sqrt{\frac{S_{x}^{2}}{n}+\frac{S_{y}^{2}}{m}}} \sim t(\nu)$$

# But what if the population isn't Normal?

#### The Central Limit Theorem

Suppose that  $X_1, \ldots, X_n$  are a random sample from a population with unknown mean  $\mu$ . Then, provided that n is sufficiently large, the sampling distribution of  $\bar{X}_n$  is approximately  $N\left(\mu, \widehat{SE}(\bar{X}_n)^2\right)$ , even if the even if the underlying population is non-normal.

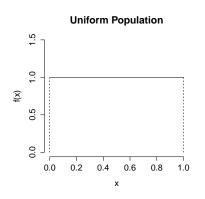
In Other Words...

$$rac{ar{X}_n - \mu}{\widehat{SE}(ar{X}_n)} pprox N(0,1)$$

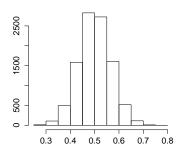
Use this to create approximate CIs for population mean!

You should be amazed by this.

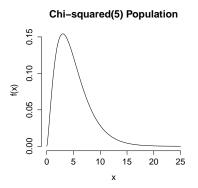
## Example: Uniform(0,1) Population, n = 20



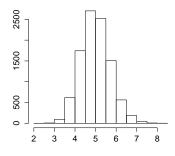
#### Sample Mean - Uniform Pop (n = 20)



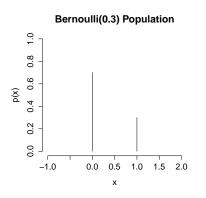
## Example: $\chi^2(5)$ Population, n=20



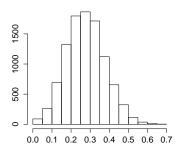
#### Sample Mean - Chisq(5) Pop (n=20)



## Example: Bernoulli(0.3) Population, n = 20



#### Sample Mean – Ber(0.3) Pop (n = 20)



## Who is the Chief Justice of the US Supreme Court?



- (a) Harry Reid
- (b) John Roberts
- (c) William Rehnquist
- (d) Stephen Breyer

## Are US Voters Really That Ignorant?

Pew: "What Voters Know About Campaign 2012"

#### The Data

Of 771 registered voters polled, only 39% correctly identified John Roberts as the current chief justice of the US Supreme Court.

#### Research Question

Is the majority of voters unaware that John Roberts is the current chief justice, or is this just sampling variation?

Assume Random Sampling...

## Confidence Interval for a Proportion

What is the appropriate probability model for the sample?

 $X_1, \ldots, X_n \sim \mathsf{iid} \; \mathsf{Bernoulli}(p), \; 1 = \mathsf{Know} \; \mathsf{Roberts} \; \mathsf{is} \; \mathsf{Chief} \; \mathsf{Justice}$ 

What is the parameter of interest?

p = Proportion of voters *in the population* who know Roberts is Chief Justice.

What is our estimator?

Sample Proportion:  $\widehat{p} = (\sum_{i=1}^{n} X_i)/n$ 

## Sample Proportion is the Sample Mean!

 $X_1, \ldots, X_n \sim \mathsf{iid} \; \mathsf{Bernoulli}(p)$ 

$$\widehat{p} = \frac{1}{n} \sum_{i=1}^{n} X_i = \bar{X}_n$$

$$E[\widehat{p}] = E\left(\frac{1}{n}\sum_{i=1}^{n}X_{i}\right) = \frac{1}{n}\sum_{i=1}^{n}E[X_{i}] = \frac{np}{n} = p$$

$$Var(\widehat{p}) = Var\left(\frac{1}{n}\sum_{i=1}^{n}X_{i}\right) = \frac{1}{n^{2}}\sum_{i=1}^{n}Var(X_{i}) = \frac{np(1-p)}{n^{2}} = \frac{p(1-p)}{n}$$

$$SE(\widehat{p}) = \sqrt{\frac{p(1-\widehat{p})}{n}}$$

$$\widehat{SE}(\widehat{p}) = \sqrt{\frac{\widehat{p}(1-\widehat{p})}{n}}$$

## Central Limit Theorem Applied to Sample Proportion

#### Central Limit Theorem: Intuition

Sample means are approximately normally distributed provided the sample size is large even if the population is non-normal.

#### CLT For Sample Mean

#### **CLT** for Sample Proportion

$$rac{ar{X}_n - \mu}{\widehat{SE}(ar{X}_n)} pprox N(0,1)$$

$$\frac{\widehat{p}-p}{\sqrt{\frac{\widehat{p}(1-\widehat{p})}{n}}}\approx N(0,1)$$

In this example, the population is Bernoulli(p) rather than normal.

The sample mean is  $\hat{p}$  and the population mean is p.

## Approximate 95% CI for Population Proportion

$$\frac{\widehat{p} - p}{\sqrt{\frac{\widehat{p}(1-\widehat{p})}{n}}} \approx N(0,1)$$

$$P\left(-2 \le \frac{\widehat{p} - p}{\sqrt{\frac{\widehat{p}(1-\widehat{p})}{n}}} \le 2\right) \approx 0.95$$

$$P\left(\widehat{p} - 2\sqrt{\frac{\widehat{p}(1-\widehat{p})}{n}} \le p \le \widehat{p} + 2\sqrt{\frac{\widehat{p}(1-\widehat{p})}{n}}\right) \approx 0.95$$

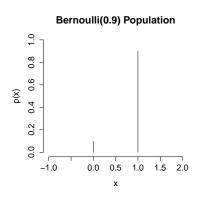
## $100 \times (1 - \alpha)$ CI for Population Proportion (p)

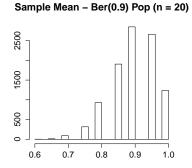
 $X_1, \ldots, X_n \sim \text{iid Bernoulli}(p)$ 

$$\widehat{p} \pm \operatorname{qnorm}(1 - \alpha/2) \sqrt{\frac{\widehat{p}(1 - \widehat{p})}{n}}$$

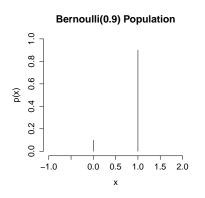
Approximation based on the CLT. Works well provided n is large and p isn't too close to zero or one.

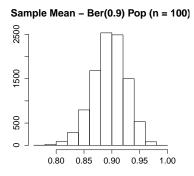
## Example: Bernoulli(0.9) Population, n = 20





## Example: Bernoulli(0.9) Population, n = 100





## Approximate 95% CI for Population Proportion



39% of 771 Voters Polled Correctly Identified Chief Justice Roberts

$$\widehat{SE}(\widehat{p}) = \sqrt{\frac{\widehat{p}(1-\widehat{p})}{n}} = \sqrt{\frac{(0.39)(0.61)}{771}}$$
  
 $\approx 0.018$ 

What is the ME for an approximate 95% confidence interval?

$$ME \approx 2 \times \widehat{SE}(\bar{X}_n) \approx 0.04$$

What can we conclude?

Approximate 95% CI: (0.35, 0.43)

#### Are Republicans Better Informed Than Democrats?

Pew: "What Voters Know About Campaign 2012"

Of the 239 Republicans surveyed, 47% correctly identified John Roberts as the current chief justice. Only 31% of the 238 Democrats surveyed correctly identified him. Is this difference meaningful or just sampling variation?

Again, assume random sampling.

## Confidence Interval for a Difference of Proportions

#### What is the appropriate probability model for the sample?

 $X_1, \ldots, X_n \sim \text{ iid Bernoulli}(p) \text{ independently of}$  $Y_1, \ldots, Y_m \sim \text{ iid Bernoulli}(q)$ 

#### What is the parameter of interest?

The difference of population proportions p-q

#### What is our estimator?

The difference of sample proportions:  $\hat{p} - \hat{q}$  where:

$$\widehat{p} = \frac{1}{n} \sum_{i=1}^{n} X_i \qquad \widehat{q} = \frac{1}{m} \sum_{i=1}^{m} Y_i$$

## Difference of Sample Proportions $\widehat{p} - \widehat{q}$ and the CLT

#### What We Have

Approx. sampling dist. for individual sample proportions from CLT:

$$\widehat{p} \approx N\left(p, \widehat{SE}(\widehat{p})^2\right), \quad \widehat{q} \approx N\left(q, \widehat{SE}(\widehat{q})^2\right)$$

What We Want

Sampling Distribution of the difference  $\widehat{p} - \widehat{q}$ 

Use Independence of the Two Samples

$$\widehat{p} - \widehat{q} \approx N \left( p - q, \widehat{SE}(\widehat{p})^2 + \widehat{SE}(\widehat{q})^2 \right)$$

$$\implies \widehat{SE}(\widehat{p} - \widehat{q}) = \sqrt{\widehat{SE}(\widehat{p})^2 + \widehat{SE}(\widehat{q})^2} = \sqrt{\frac{\widehat{p}(1 - \widehat{p})}{n} + \frac{\widehat{q}(1 - \widehat{q})}{m}}$$

## Approx. 95% CI for Difference of Population Proportions

$$\frac{(\widehat{p}-\widehat{q})-(p-q)}{\sqrt{\frac{\widehat{p}(1-\widehat{p})}{n}+\frac{\widehat{q}(1-\widehat{q})}{m}}}\approx N(0,1)$$

$$P\left(-2 \le \frac{(\widehat{p} - \widehat{q}) - (p - q)}{\sqrt{\frac{\widehat{p}(1 - \widehat{p})}{n} + \frac{\widehat{q}(1 - \widehat{q})}{m}}} \le 2\right) \approx 0.95$$

$$(\widehat{
ho}-\widehat{q}) \pm \mathtt{qnorm}(1-lpha/2) \sqrt{rac{\widehat{p}(1-\widehat{
ho})}{n} + rac{\widehat{q}(1-\widehat{q})}{m}}$$

$$100 \times (1-\alpha)$$
 CI for Diff. of Popn. Proportions  $(p-q)$ 

 $X_1, \ldots, X_n \sim \mathsf{iid} \; \mathsf{Bernoulli}(p) \; \mathsf{indep.} \; Y_1, \ldots, Y_n \sim \mathsf{iid} \; \mathsf{Bernoulli}(q)$ 

$$(\widehat{
ho}-\widehat{q}) \pm \mathtt{qnorm}(1-lpha/2) \sqrt{rac{\widehat{
ho}(1-\widehat{
ho})}{n} + rac{\widehat{q}(1-\widehat{q})}{m}}$$

Approximation based on the CLT. Works well provided n, m large and p, q aren't too close to zero or one.

## ME for approx. 95% for Difference of Proportions

47% of 239 Republicans vs. 31% of 238 Democrats identified Roberts

#### Republicans

## $\hat{p} = 0.47$ $\widehat{SE}(\widehat{p}) = \sqrt{\frac{\widehat{p}(1-\widehat{p})}{n}} \approx 0.032$ $\widehat{SE}(\widehat{q}) = \sqrt{\frac{\widehat{q}(1-\widehat{q})}{m}} \approx 0.030$

#### Democrats

$$\widehat{q} = 0.31$$
 $m = 238$ 
 $\widehat{SE}(\widehat{q}) = \sqrt{\frac{\widehat{q}(1-\widehat{q})}{m}} \approx 0.030$ 

#### Difference: (Republicans - Democrats)

$$\begin{split} \widehat{p} - \widehat{q} &= 0.47 - 0.31 = 0.16 \\ \widehat{SE}(\widehat{p} - \widehat{q}) &= \sqrt{\widehat{SE}(\widehat{p})^2 + \widehat{SE}(\widehat{q})^2} \approx 0.044 \implies ME \approx 0.09 \\ \hline \text{Approximate 95\% CI} & (0.07, 0.25) \end{split}$$
 What can we conclude?