

# Economics 103 – Statistics for Economists

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Lecture # 13

# Continuous RVs – Part III

## Last Time

- ▶ Expectation for Continuous RVs
- ▶ Normal Random Variable
- ▶ Linear Combination of Normal RV
- ▶ Areas under Normal pdfs

## Today

- ▶ Percentiles/Quantiles for Continuous RVs
- ▶ Linear Combination of *Several* Normal RVs
- ▶ Friends of Normal Distribution

## Recall: Def. of Cumulative Distribution Function (CDF)

$$F(x_0) \equiv P(X \leq x_0)$$

$$= \int_{-\infty}^{x_0} f(x) \, dx \text{ for Continuous RVs}$$

## Percentiles/Quantiles for Continuous RVs

Quantile Function  $Q(p)$  is the inverse of CDF  $F(x_0)$

Plug in a probability  $p$ , get out the value of  $x_0$  such that  $F(x_0) = p$

$$Q(p) = F^{-1}(p)$$

In other words:

$$Q(p) = \text{the value of } x_0 \text{ such that } \int_{-\infty}^{x_0} f(x) dx = p$$

Inverse exists as long as  $F(x_0)$  is *strictly increasing*.

## Example: Median

The median of a continuous random variable is  $Q(0.5)$ , i.e. the value of  $x_0$  such that

$$\int_{-\infty}^{x_0} f(x) dx = 1/2$$

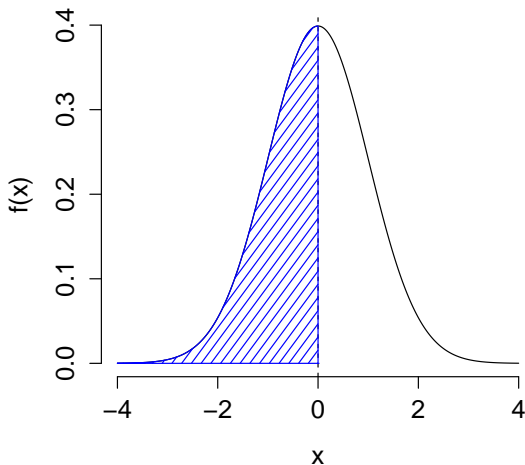
What is the median of a standard normal RV?



## What is the median of a standard normal RV?



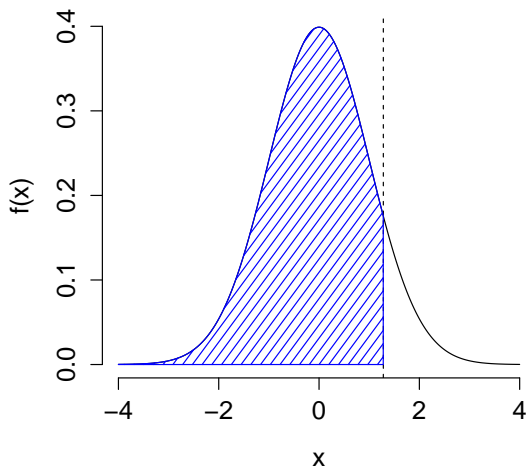
By symmetry,  $Q(0.5) = 0$ . R command: `qnorm()`





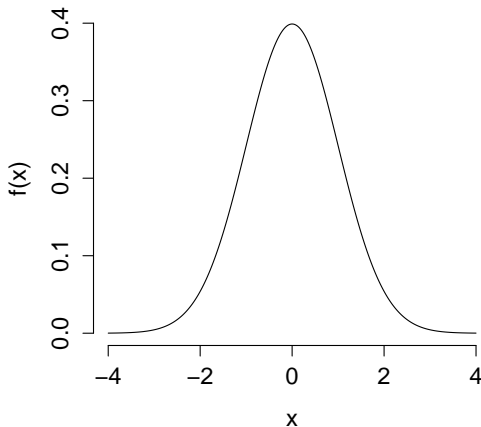
## 90th Percentile of a Standard Normal

$$\text{qnorm}(0.9) \approx 1.28$$



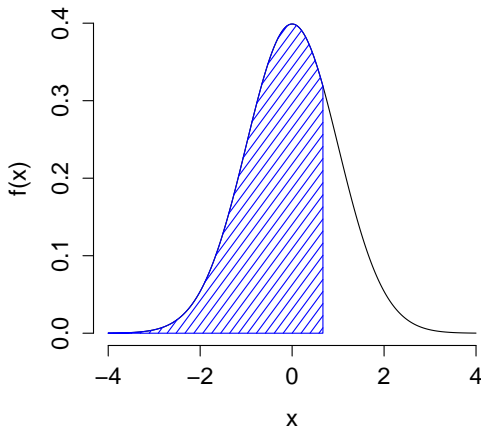
## Using Quantile Function to find Symmetric Intervals

Suppose  $X$  is a standard normal RV. What is the value of  $c$  such that  $P(-c \leq X \leq c) = 0.5$ ?



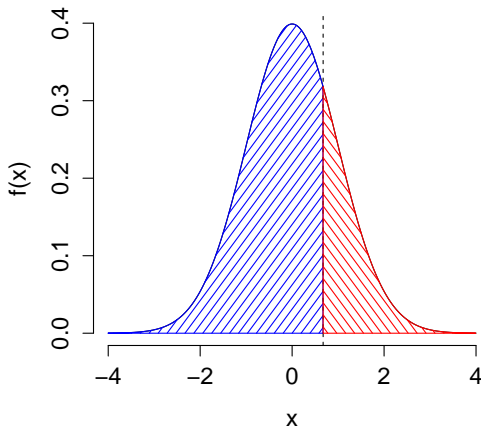
$$\text{qnorm}(0.75) \approx 0.67$$

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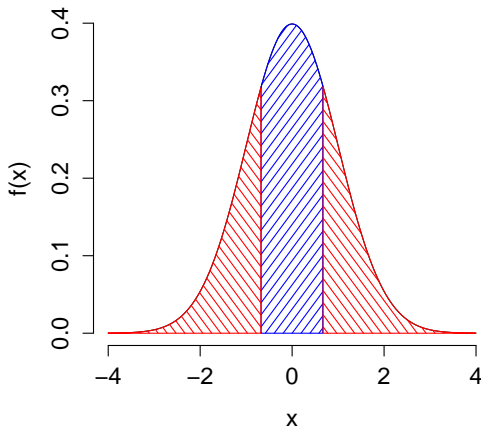
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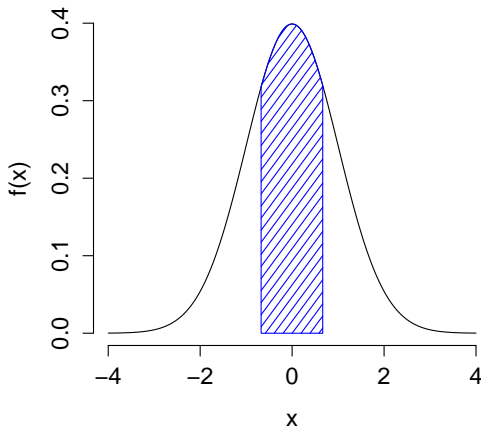
$$\text{pnorm}(0.67) - \text{pnorm}(-0.67) \approx ?$$

Suppose  $X$  is a standard normal RV. What is the value of  $c$  such that  $P(-c \leq X \leq c) = 0.5$ ?



$$\text{pnorm}(0.67) - \text{pnorm}(-0.67) \approx 0.5$$

Suppose  $X$  is a standard normal RV. What is the value of  $c$  such that  $P(-c \leq X \leq c) = 0.5$ ?



## 68% Central Interval for Standard Normal



Suppose  $X$  is a standard normal random variable. What value of  $c$  ensures that  $P(-c \leq X \leq c) \approx 0.68$ ?

## 95% Central Interval for Standard Normal



Suppose  $X$  is a standard normal random variable. What value of  $c$  ensures that  $P(-c \leq X \leq c) \approx 0.95$ ?



## R Commands for *Arbitrary* Normal Distributions

Let  $X \sim N(\mu, \sigma^2)$  . Then we can use R to evaluate the CDF and Quantile function of  $X$  as follows:

CDF $F(x)$	<code>pnorm(x, mean = <math>\mu</math>, sd = <math>\sigma</math>)</code>
Quantile Function $Q(p)$	<code>qnorm(p, mean = <math>\mu</math>, sd = <math>\sigma</math>)</code>

Notice that this means you don't have to transform  $X$  to a standard normal in order to find areas under its pdf using R.

## Example from Homework: $X \sim N(0, 16)$

One Way:

$$\begin{aligned} P(X \geq 10) &= 1 - P(X \leq 10) = 1 - P(X/4 \leq 10/4) \\ &= 1 - P(Z \leq 2.5) = 1 - \Phi(2.5) = 1 - \text{pnorm}(2.5) \\ &\approx 0.006 \end{aligned}$$

## Example from Homework: $X \sim N(0, 16)$

One Way:

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An Easier Way:

$$\begin{aligned}P(X \geq 10) &= 1 - P(X \leq 10) \\&= 1 - \text{pnorm}(10, \text{mean} = 0, \text{sd} = 4) \\&\approx 0.006\end{aligned}$$

Suppose  $X$  has mean  $\mu_x$  variance  $\sigma_x^2$  and is independent of  $Y$ , which has mean  $\mu_y$  variance  $\sigma_y^2$ . Let  $a, b$  be constants.

What is  $E[aX + bY]$ ?

Suppose  $X$  has mean  $\mu_x$  variance  $\sigma_x^2$  and is independent of  $Y$ , which has mean  $\mu_y$  variance  $\sigma_y^2$ . Let  $a, b$  be constants.

What is  $E[aX + bY]$ ?

$$E[aX + bY] = aE[X] + bE[Y] = a\mu_x + b\mu_y$$

What is  $Var(aX + bY)$ ?

Suppose  $X$  has mean  $\mu_x$  variance  $\sigma_x^2$  and is independent of  $Y$ , which has mean  $\mu_y$  variance  $\sigma_y^2$ . Let  $a, b$  be constants.

What is  $E[aX + bY]$ ?

$$E[aX + bY] = aE[X] + bE[Y] = a\mu_x + b\mu_y$$

What is  $Var(aX + bY)$ ?

$$Var(aX + bY) = a^2 Var(X) + b^2 Var(Y) = a^2 \sigma_x^2 + b^2 \sigma_y^2$$

By independence.

Now suppose  $X \sim N(\mu_x, \sigma_x^2)$  independent of  $Y \sim N(\mu_y, \sigma_y^2)$ . Let  $a, b$  be constants.

What is  $E[aX + bY]$ ?

Now suppose  $X \sim N(\mu_x, \sigma_x^2)$  independent of  $Y \sim N(\mu_y, \sigma_y^2)$ . Let  $a, b$  be constants.

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What is  $Var(aX + bY)$ ?

$$Var(aX + bY) = a^2 Var(X) + b^2 Var(Y) = a^2 \sigma_x^2 + b^2 \sigma_y^2$$

By independence.

## Here's the Surprising Thing:

If  $X$  and  $Y$  are independent Normal Random Variables and  $a, b$  are constants, then  $aX + bY$  is *also* a Normal Random Variable!

# Linear Combinations of Independent Normals

Let  $X \sim N(\mu_x, \sigma_x^2)$  independent of  $Y \sim N(\mu_y, \sigma_y^2)$ . Then if  $a, b, c$  are constants:

$$aX + bY + c \sim N(a\mu_x + b\mu_y + c, a^2\sigma_x^2 + b^2\sigma_y^2)$$

## Important

- ▶ Result assumes independence
- ▶ Particular to Normal Distribution
- ▶ Extends to more than two Normal RVs

Suppose  $X_1, X_2, \sim \text{iid } N(\mu, \sigma^2)$



Let  $\bar{X} = (X_1 + X_2)/2$ . What is the distribution of  $\bar{X}$ ?

- (a)  $N(\mu, \sigma^2/2)$
- (b)  $N(0, 1)$
- (c)  $N(\mu, \sigma^2)$
- (d)  $N(\mu, 2\sigma^2)$
- (e)  $N(2\mu, 2\sigma^2)$



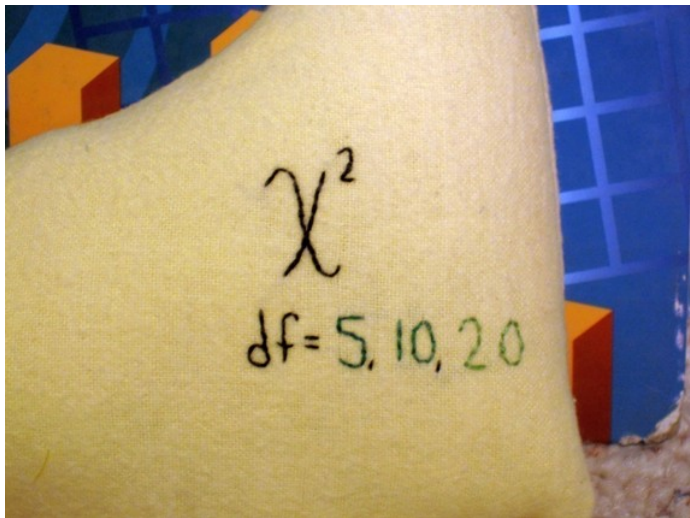
Figure : The Normal Distribution and Friends.

## Functions of Independent RVs are Independent

If  $X$  and  $Y$  are independent random variables and  $g$  and  $h$  are functions, then the random variables  $g(X)$  and  $h(Y)$  are also independent.



Figure : PDF for  $\chi^2$ -Distribution





## $\chi^2$ Random Variable

Let  $X_1, \dots, X_\nu \sim \text{iid } N(0, 1)$ . Then,

$$(X_1^2 + \dots + X_\nu^2) \sim \chi^2(\nu)$$

where the parameter  $\nu$  is the *degrees of freedom*

## $\chi^2$ Random Variable

Let  $X_1, \dots, X_\nu \sim \text{iid } N(0, 1)$ . Then,

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Support =  $(0, \infty)$

## $\chi^2$ PDFs

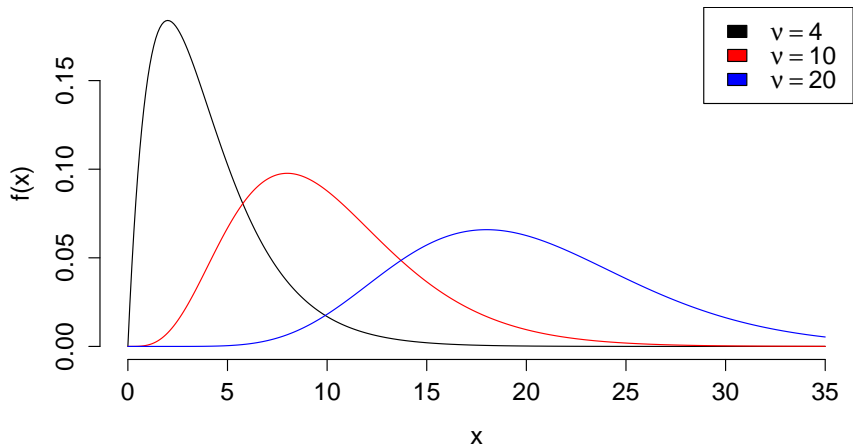
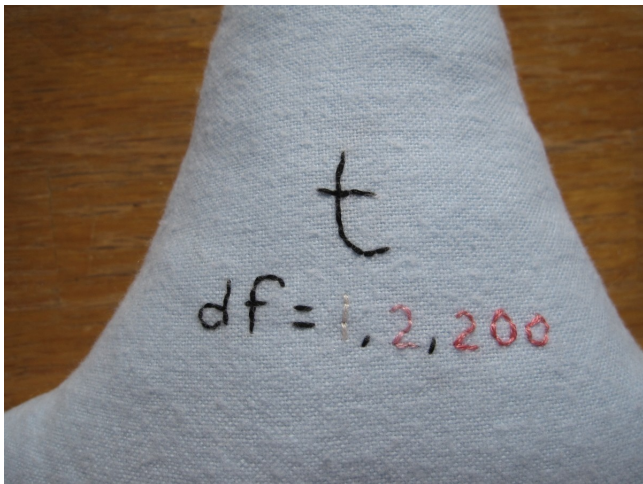




Figure : PDF for Student-t Distribution



## Student-t Random Variable

Let  $X \sim N(0, 1)$  independent of  $Y \sim \chi^2(\nu)$ . Then,

$$\frac{X}{\sqrt{Y/\nu}} \sim t(\nu)$$

where the parameter  $\nu$  is the degrees of freedom.

## Student-t Random Variable

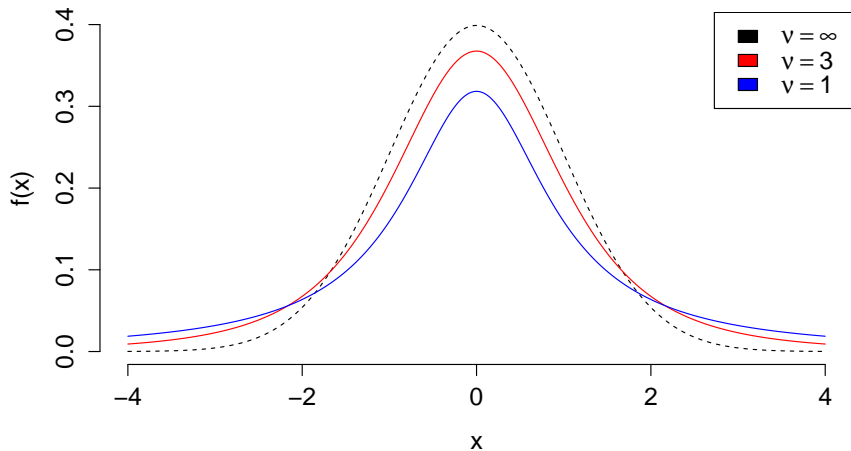
Let  $X \sim N(0, 1)$  independent of  $Y \sim \chi^2(\nu)$ . Then,

$$\frac{X}{\sqrt{Y/\nu}} \sim t(\nu)$$

where the parameter  $\nu$  is the degrees of freedom.

- ▶ Support =  $(-\infty, \infty)$
- ▶ As  $\nu \rightarrow \infty$ ,  $t \rightarrow$  Standard Normal.
- ▶ Symmetric around zero, but mean and variance may not exist!
- ▶ Degrees of freedom  $\nu$  control “thickness of tails”

# Student-t PDFs





## F Random Variable

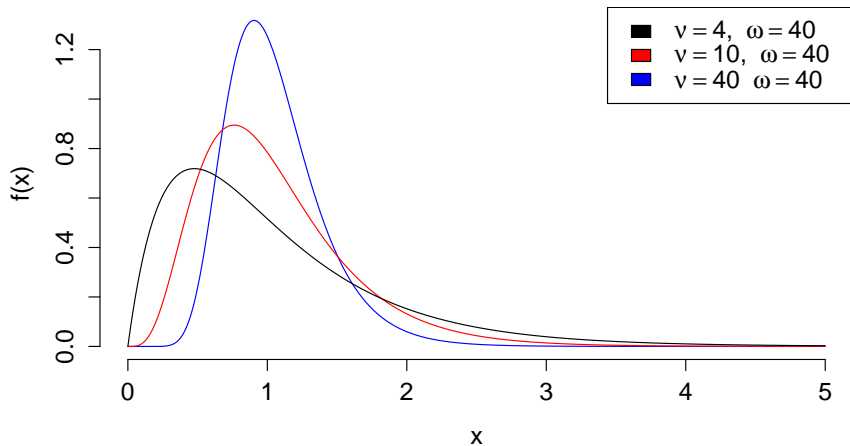
Suppose  $X \sim \chi^2(\nu)$  independent of  $Y \sim \chi^2(\omega)$ . Then,

$$\frac{X/\nu}{Y/\omega} \sim F(\nu, \omega)$$

where  $\nu$  is the numerator degrees of freedom and  $\omega$  is the denominator degrees of freedom.

Support =  $(0, \infty)$

## F PDFs



## R Commands – CDFs and Quantile Functions

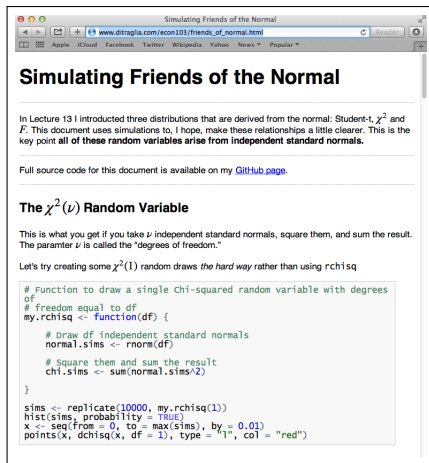
$F(x) = P(X \leq x)$  is the CDF,  $Q(p) = F^{-1}(p)$  the Quantile Function

	$F(x)$	$Q(p)$
$N(\mu, \sigma^2)$	<code>pnorm(x, mean = <math>\mu</math>, sd = <math>\sigma</math>)</code>	<code>qnorm(p, mean = <math>\mu</math>, sd = <math>\sigma</math>)</code>
$\chi^2(\nu)$	<code>pchisq(x, df = <math>\nu</math>)</code>	<code>qchisq(p, df = <math>\nu</math>)</code>
$t(\nu)$	<code>pt(x, df = <math>\nu</math>)</code>	<code>qt(p, df = <math>\nu</math>)</code>
$F(\nu, \omega)$	<code>pf(x, df1 = <math>\nu</math>, df2 = <math>\omega</math>)</code>	<code>qf(p, df1 = <math>\nu</math>, df2 = <math>\omega</math>)</code>

Mnemonic: “p” is for Probability, “q” is for Quantile.

[http://fditraglia.github.io/Econ103Public/Rtutorials/friends\\_of\\_normal.html](http://fditraglia.github.io/Econ103Public/Rtutorials/friends_of_normal.html)

Source Code on my [Github Page](#)



## Simulating Friends of the Normal

In Lecture 13 I introduced three distributions that are derived from the normal: Student-t,  $\chi^2$  and  $F$ . This document uses simulations to, I hope, make these relationships a little clearer. This is the key point **all of these random variables arise from independent standard normals**.

Full source code for this document is available on my [Github page](#).

### The $\chi^2(\nu)$ Random Variable

This is what you get if you take  $\nu$  independent standard normals, square them, and sum the result. The parameter  $\nu$  is called the "degrees of freedom."

Let's try creating some  $\chi^2(1)$  random draws *the hard way* rather than using `rchisq`

```
# Function to draw a single Chi-squared random variable with degrees
# of
# freedom equal to df
my.rchisq <- function(df) {
  # Draw df independent standard normals
  normal.sims <- rnorm(df)

  # Square them and sum the result
  chi.sims <- sum(normal.sims^2)
}

sims <- replicate(10000, my.rchisq(1))
hist(sims, probability = TRUE)
x <- seq(from = 0, to = max(sims), by = 0.01)
points(x, dchisq(x, df = 1), type = "l", col = "red")
```

## R Commands – PDFs and Random Draws

	$f(x)$	Make n iid Random Draws
$N(\mu, \sigma^2)$	<code>dnorm(x, mean = <math>\mu</math>, sd = <math>\sigma</math>)</code>	<code>rnorm(n, mean = <math>\mu</math>, sd = <math>\sigma</math>)</code>
$\chi^2(\nu)$	<code>dchisq(x, df = <math>\nu</math>)</code>	<code>rchisq(n, df = <math>\nu</math>)</code>
$t(\nu)$	<code>dt(x, df = <math>\nu</math>)</code>	<code>rt(n, df = <math>\nu</math>)</code>
$F(\nu, \omega)$	<code>df(x, df1 = <math>\nu</math>, df2 = <math>\omega</math>)</code>	<code>rf(n, df1 = <math>\nu</math>, df2 = <math>\omega</math>)</code>

Mnemonic: “d” is for Density, “r” is for Random.

Example:  $X_1, X_2, X_3 \sim \text{iid } N(0, 1)$

What is the distribution of  $Y_1 = X_1^2 + X_2^2$ ?

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Sum of squares of two indep. std. normals  $\Rightarrow Y_1 \sim \chi^2(2)$

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Sum of squares of two indep. std. normals  $\Rightarrow Y_1 \sim \chi^2(2)$

What is the distribution of  $Y_2 = (Y_1/2)/(X_3^2)$ ?



Example:  $X_1, X_2, X_3 \sim \text{iid } N(0, 1)$

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What is the distribution of  $Y_2 = (Y_1/2)/(X_3^2)$ ?

$Y_1 \sim \chi^2(2)$  and  $X_3^2 \sim \chi^2(1)$

Example:  $X_1, X_2, X_3 \sim \text{iid } N(0, 1)$

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$Y_1 \sim \chi^2(2)$  and  $X_3^2 \sim \chi^2(1)$

Hence  $Y_2 = \text{ratio of two indep. } \chi^2 \text{ RVs, each divided by its degrees of freedom} \Rightarrow Y_2 \sim F(2, 1)$

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What is the distribution of  $Z = X_3/\sqrt{Y_1/2}$ ?

Example:  $X_1, X_2, X_3 \sim \text{iid } N(0, 1)$

What is the distribution of  $Y_1 = X_1^2 + X_2^2$ ?

Sum of squares of two indep. std. normals  $\Rightarrow Y_1 \sim \chi^2(2)$

What is the distribution of  $Y_2 = (Y_1/2)/(X_3^2)$ ?

$Y_1 \sim \chi^2(2)$  and  $X_3^2 \sim \chi^2(1)$

Hence  $Y_2 =$  ratio of two indep.  $\chi^2$  RVs, each divided by its degrees of freedom  $\Rightarrow Y_2 \sim F(2, 1)$

What is the distribution of  $Z = X_3/\sqrt{Y_1/2}$ ?

Ratio of standard normal and square root of independent  $\chi^2$  RV divided by its degrees of freedom  $\Rightarrow Z \sim t(2)$

Suppose  $X_1, X_2, \sim \text{iid } N(\mu, \sigma^2)$



Let  $Y = (X_1 - \mu)^2 + (X_2 - \mu)^2$ . What is the distribution of  $Y/\sigma^2$ ?

- (a)  $F(2, 1)$
- (b)  $\chi^2(2)$
- (c)  $t(2)$
- (d)  $N(\mu, \sigma)$
- (e) None of the above

$$Y_1 \sim \chi^2(2), \quad Y_2 \sim F(2, 1), \quad Z \sim t(2)$$

What is the median of  $Y_1$ ?

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`qchisq(0.5, df = 2)`  $\approx 1.4$

$$Y_1 \sim \chi^2(2), \quad Y_2 \sim F(2, 1), \quad Z \sim t(2)$$

What is the median of  $Y_1$ ?

`qchisq(0.5, df = 2)`  $\approx 1.4$

What is  $P(Y_2 \leq 5)$ ?



$$Y_1 \sim \chi^2(2), \quad Y_2 \sim F(2, 1), \quad Z \sim t(2)$$

What is the median of  $Y_1$ ?

`qchisq(0.5, df = 2)`  $\approx 1.4$

What is  $P(Y_2 \leq 5)$ ?

`pf(5, df1 = 2, df2 = 1)`  $\approx 0.7$

$$Y_1 \sim \chi^2(2), \quad Y_2 \sim F(2, 1), \quad Z \sim t(2)$$

What is the median of  $Y_1$ ?

`qchisq(0.5, df = 2)`  $\approx 1.4$

What is  $P(Y_2 \leq 5)$ ?

`pf(5, df1 = 2, df2 = 1)`  $\approx 0.7$

What value of  $c$  gives  $P(-c \leq Z \leq c) = 0.5$ ?

Use Symmetry (like normal)

`c = qt(0.75, df = 2)`  $\approx 0.8$

$$Y_1 \sim \chi^2(2), \quad Y_2 \sim F(2, 1), \quad Z \sim t(2)$$

What is the median of  $Y_1$ ?

$$\text{qchisq}(0.5, \text{df} = 2) \approx 1.4$$

What is  $P(Y_2 \leq 5)$ ?

$$\text{pf}(5, \text{df1} = 2, \text{df2} = 1) \approx 0.7$$

What value of  $c$  gives  $P(-c \leq Z \leq c) = 0.5$ ?

Use Symmetry (like normal)

$$c = \text{qt}(0.75, \text{df} = 2) \approx 0.8$$

$$\text{or equivalently } -c = \text{qt}(0.25, \text{df} = 2) \approx -0.8$$