

Economics 103 – Statistics for Economists

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Lecture # 12

Continuous RVs – Part III

Last Time

- ▶ Expectation for Continuous RVs
- ▶ Normal Random Variable
- ▶ Linear Combination of Normal RV
- ▶ Areas under Normal pdfs

Today

- ▶ Percentiles/Quantiles for Continuous RVs
- ▶ Linear Combination of *Several* Normal RVs
- ▶ Friends of Normal Distribution

Recall: Def. of Cumulative Distribution Function (CDF)

$$F(x_0) \equiv P(X \leq x_0)$$

$$= \int_{-\infty}^{x_0} f(x) dx \text{ for Continuous RVs}$$

Percentiles/Quantiles for Continuous RVs

Quantile Function $Q(p)$ is the inverse of CDF $F(x_0)$

Plug in a probability p , get out the value of x_0 such that $F(x_0) = p$

$$Q(p) = F^{-1}(p)$$

In other words:

$$Q(p) = \text{the value of } x_0 \text{ such that } \int_{-\infty}^{x_0} f(x) dx = p$$

Inverse exists as long as $F(x_0)$ is *strictly increasing*.

Example: Median

The median of a continuous random variable is $Q(0.5)$, i.e. the value of x_0 such that

$$\int_{-\infty}^{x_0} f(x) dx = 1/2$$

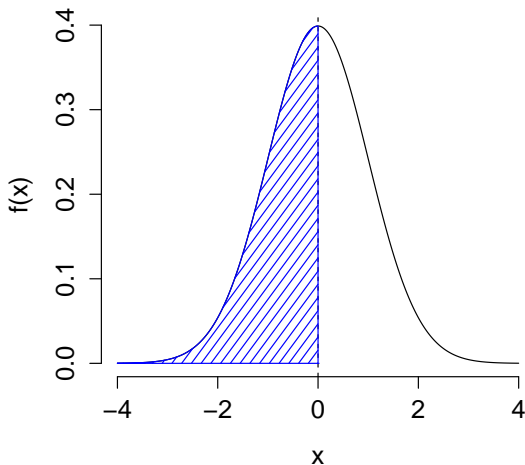
What is the median of a standard normal RV?



What is the median of a standard normal RV?

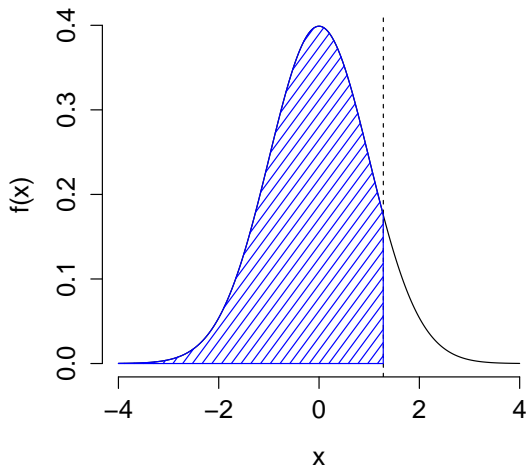


By symmetry, $Q(0.5) = 0$. R command: `qnorm()`



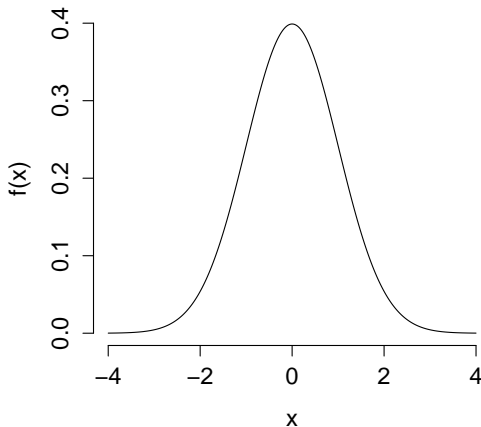
90th Percentile of a Standard Normal

$$\text{qnorm}(0.9) \approx 1.28$$



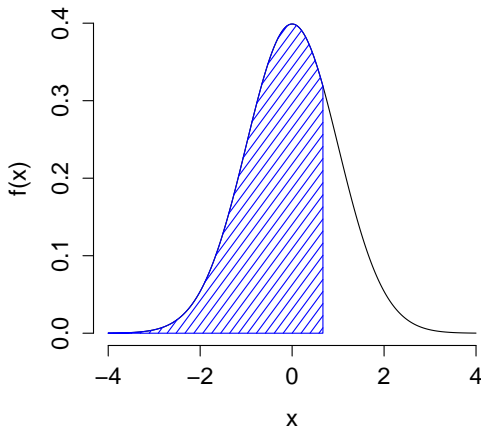
Using Quantile Function to find Symmetric Intervals

Suppose X is a standard normal RV. What is the value of c such that $P(-c \leq X \leq c) = 0.5$?



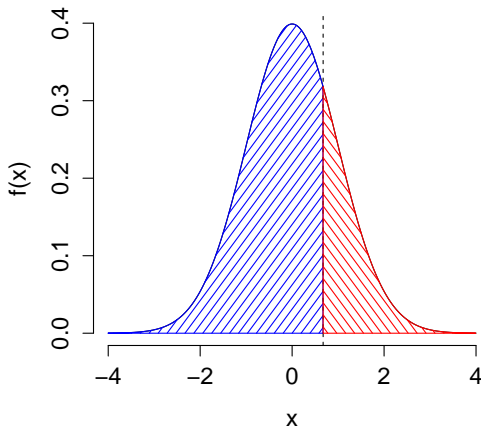
$$\text{qnorm}(0.75) \approx 0.67$$

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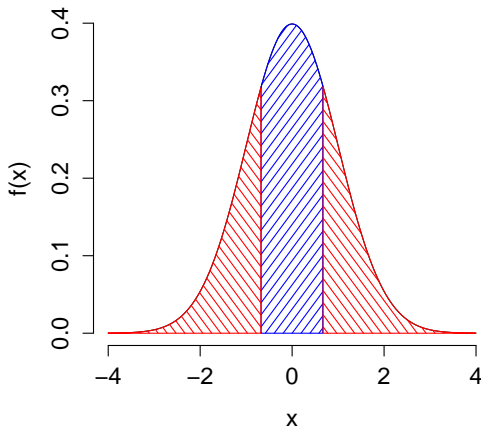
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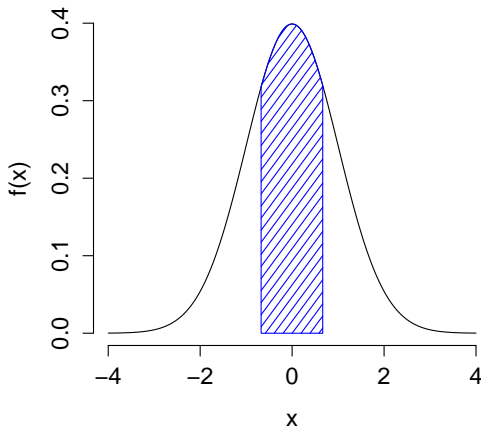
$$\text{pnorm}(0.67) - \text{pnorm}(-0.67) \approx ?$$

Suppose X is a standard normal RV. What is the value of c such that $P(-c \leq X \leq c) = 0.5$?



$$\text{pnorm}(0.67) - \text{pnorm}(-0.67) \approx 0.5$$

Suppose X is a standard normal RV. What is the value of c such that $P(-c \leq X \leq c) = 0.5$?



68% Central Interval for Standard Normal



Suppose X is a standard normal random variable. What value of c ensures that $P(-c \leq X \leq c) \approx 0.68$?

95% Central Interval for Standard Normal



Suppose X is a standard normal random variable. What value of c ensures that $P(-c \leq X \leq c) \approx 0.95$?

R Commands for *Arbitrary* Normal Distributions

Let $X \sim N(\mu, \sigma^2)$. Then we can use R to evaluate the CDF and Quantile function of X as follows:

CDF $F(x)$	<code>pnorm(x, mean = μ, sd = σ)</code>
Quantile Function $Q(p)$	<code>qnorm(p, mean = μ, sd = σ)</code>

Notice that this means you don't have to transform X to a standard normal in order to find areas under its pdf using R.

Example from Homework: $X \sim N(0, 16)$

One Way:

$$\begin{aligned} P(X \geq 10) &= 1 - P(X \leq 10) = 1 - P(X/4 \leq 10/4) \\ &= 1 - P(Z \leq 2.5) = 1 - \Phi(2.5) = 1 - \text{pnorm}(2.5) \\ &\approx 0.006 \end{aligned}$$

Example from Homework: $X \sim N(0, 16)$

One Way:

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An Easier Way:

$$\begin{aligned}P(X \geq 10) &= 1 - P(X \leq 10) \\&= 1 - \text{pnorm}(10, \text{mean} = 0, \text{sd} = 4) \\&\approx 0.006\end{aligned}$$

Suppose X has mean μ_x variance σ_x^2 and is independent of Y , which has mean μ_y variance σ_y^2 . Let a, b be constants.

What is $E[aX + bY]$?

Suppose X has mean μ_x variance σ_x^2 and is independent of Y , which has mean μ_y variance σ_y^2 . Let a, b be constants.

What is $E[aX + bY]$?

$$E[aX + bY] = aE[X] + bE[Y] = a\mu_x + b\mu_y$$

What is $Var(aX + bY)$?

Suppose X has mean μ_x variance σ_x^2 and is independent of Y , which has mean μ_y variance σ_y^2 . Let a, b be constants.

What is $E[aX + bY]$?

$$E[aX + bY] = aE[X] + bE[Y] = a\mu_x + b\mu_y$$

What is $Var(aX + bY)$?

$$Var(aX + bY) = a^2 Var(X) + b^2 Var(Y) = a^2 \sigma_x^2 + b^2 \sigma_y^2$$

By independence.

Now suppose $X \sim N(\mu_x, \sigma_x^2)$ independent of $Y \sim N(\mu_y, \sigma_y^2)$. Let a, b be constants.

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What is $E[aX + bY]$?

$$E[aX + bY] = aE[X] + bE[Y] = a\mu_x + b\mu_y$$

What is $Var(aX + bY)$?

$$Var(aX + bY) = a^2 Var(X) + b^2 Var(Y) = a^2 \sigma_x^2 + b^2 \sigma_y^2$$

By independence.

Here's the Surprising Thing:

If X and Y are independent Normal Random Variables and a, b are constants, then $aX + bY$ is *also* a Normal Random Variable!

Linear Combinations of Independent Normals

Let $X \sim N(\mu_x, \sigma_x^2)$ independent of $Y \sim N(\mu_y, \sigma_y^2)$. Then if a, b, c are constants:

$$aX + bY + c \sim N(a\mu_x + b\mu_y + c, a^2\sigma_x^2 + b^2\sigma_y^2)$$

Important

- ▶ Result assumes independence
- ▶ Particular to Normal Distribution
- ▶ Extends to more than two Normal RVs

Suppose $X_1, X_2, \sim \text{iid } N(\mu, \sigma^2)$



Let $\bar{X} = (X_1 + X_2)/2$. What is the distribution of \bar{X} ?

- (a) $N(\mu, \sigma^2/2)$
- (b) $N(0, 1)$
- (c) $N(\mu, \sigma^2)$
- (d) $N(\mu, 2\sigma^2)$
- (e) $N(2\mu, 2\sigma^2)$



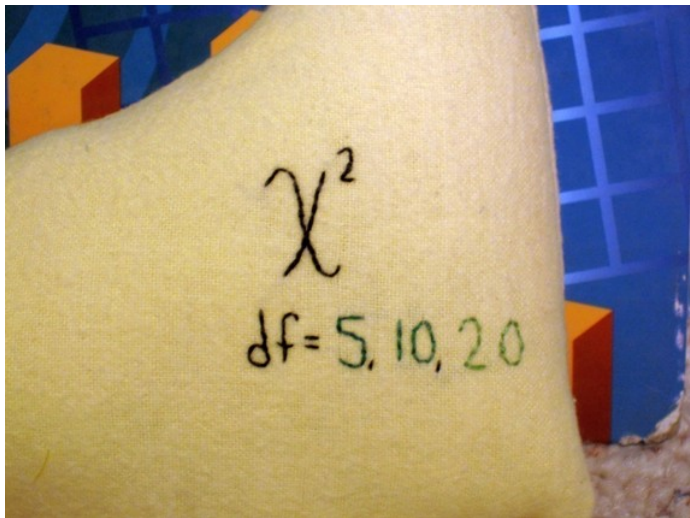
Figure : The Normal Distribution and Friends.

Functions of Independent RVs are Independent

If X and Y are independent random variables and g and h are functions, then the random variables $g(X)$ and $h(Y)$ are also independent.



Figure : PDF for χ^2 -Distribution



χ^2 Random Variable

Let $X_1, \dots, X_\nu \sim \text{iid } N(0, 1)$. Then,

$$(X_1^2 + \dots + X_\nu^2) \sim \chi^2(\nu)$$

where the parameter ν is the *degrees of freedom*

χ^2 Random Variable

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Support = $(0, \infty)$

χ^2 PDFs

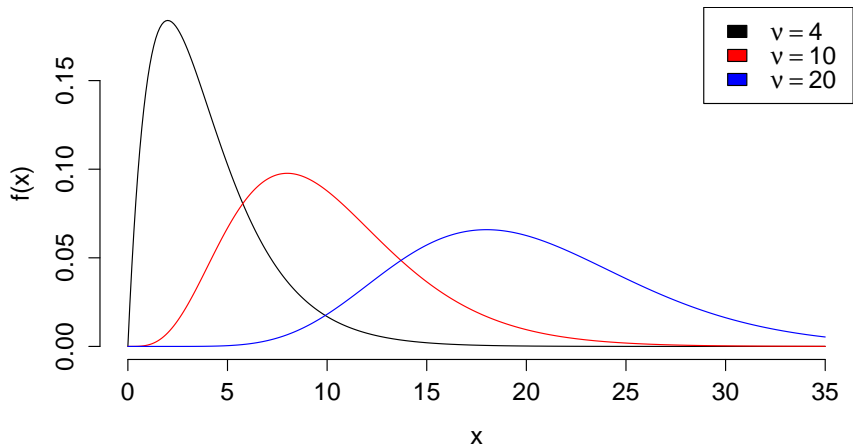
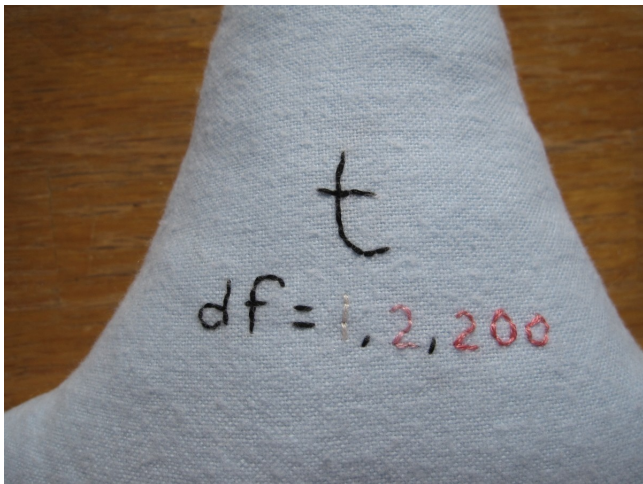




Figure : PDF for Student-t Distribution



Student-t Random Variable

Let $X \sim N(0, 1)$ independent of $Y \sim \chi^2(\nu)$. Then,

$$\frac{X}{\sqrt{Y/\nu}} \sim t(\nu)$$

where the parameter ν is the degrees of freedom.

Student-t Random Variable

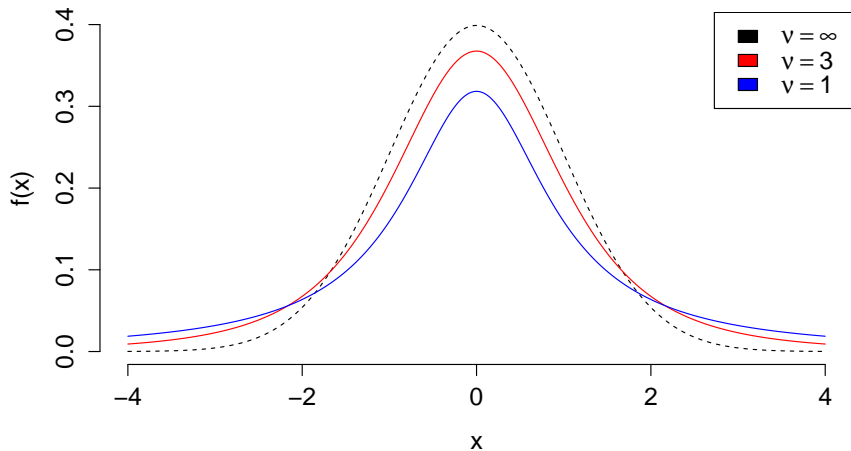
Let $X \sim N(0, 1)$ independent of $Y \sim \chi^2(\nu)$. Then,

$$\frac{X}{\sqrt{Y/\nu}} \sim t(\nu)$$

where the parameter ν is the degrees of freedom.

- ▶ Support = $(-\infty, \infty)$
- ▶ As $\nu \rightarrow \infty$, $t \rightarrow$ Standard Normal.
- ▶ Symmetric around zero, but mean and variance may not exist!
- ▶ Degrees of freedom ν control “thickness of tails”

Student-t PDFs



F Random Variable

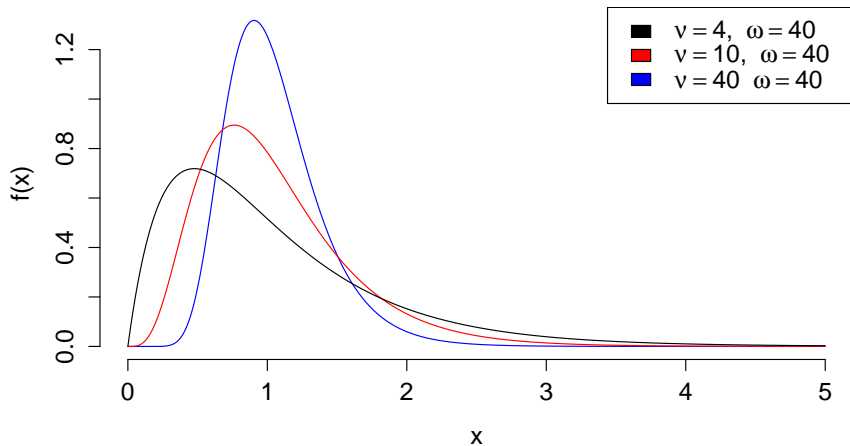
Suppose $X \sim \chi^2(\nu)$ independent of $Y \sim \chi^2(\omega)$. Then,

$$\frac{X/\nu}{Y/\omega} \sim F(\nu, \omega)$$

where ν is the numerator degrees of freedom and ω is the denominator degrees of freedom.

Support = $(0, \infty)$

F PDFs



R Commands – CDFs and Quantile Functions

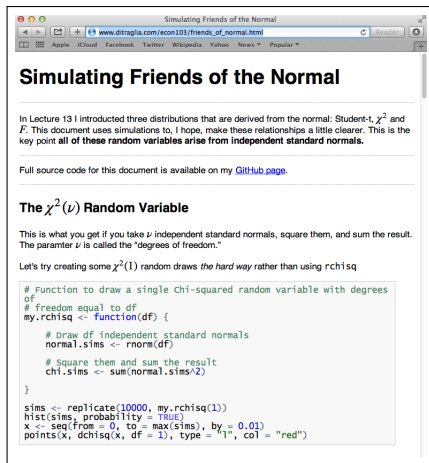
$F(x) = P(X \leq x)$ is the CDF, $Q(p) = F^{-1}(p)$ the Quantile Function

	$F(x)$	$Q(p)$
$N(\mu, \sigma^2)$	<code>pnorm(x, mean = μ, sd = σ)</code>	<code>qnorm(p, mean = μ, sd = σ)</code>
$\chi^2(\nu)$	<code>pchisq(x, df = ν)</code>	<code>qchisq(p, df = ν)</code>
$t(\nu)$	<code>pt(x, df = ν)</code>	<code>qt(p, df = ν)</code>
$F(\nu, \omega)$	<code>pf(x, df1 = ν, df2 = ω)</code>	<code>qf(p, df1 = ν, df2 = ω)</code>

Mnemonic: “p” is for Probability, “q” is for Quantile.

http://fditraglia.github.io/Econ103Public/Rtutorials/friends_of_normal.html

Source Code on my [Github Page](#)



Simulating Friends of the Normal

In Lecture 13 I introduced three distributions that are derived from the normal: Student-t, χ^2 and F . This document uses simulations to, I hope, make these relationships a little clearer. This is the key point **all of these random variables arise from independent standard normals**.

Full source code for this document is available on my [Github page](#).

The $\chi^2(\nu)$ Random Variable

This is what you get if you take ν independent standard normals, square them, and sum the result. The parameter ν is called the "degrees of freedom."

Let's try creating some $\chi^2(1)$ random draws *the hard way* rather than using `rchisq`

```
# Function to draw a single Chi-squared random variable with degrees
# of
# freedom equal to df
my.rchisq <- function(df) {
  # Draw df independent standard normals
  normal.sims <- rnorm(df)

  # Square them and sum the result
  chi.sims <- sum(normal.sims^2)
}

sims <- replicate(10000, my.rchisq(1))
hist(sims, probability = TRUE)
x <- seq(from = 0, to = max(sims), by = 0.01)
points(x, dchisq(x, df = 1), type = "l", col = "red")
```

R Commands – PDFs and Random Draws

	$f(x)$	Make n iid Random Draws
$N(\mu, \sigma^2)$	<code>dnorm(x, mean = μ, sd = σ)</code>	<code>rnorm(n, mean = μ, sd = σ)</code>
$\chi^2(\nu)$	<code>dchisq(x, df = ν)</code>	<code>rchisq(n, df = ν)</code>
$t(\nu)$	<code>dt(x, df = ν)</code>	<code>rt(n, df = ν)</code>
$F(\nu, \omega)$	<code>df(x, df1 = ν, df2 = ω)</code>	<code>rf(n, df1 = ν, df2 = ω)</code>

Mnemonic: “d” is for Density, “r” is for Random.

Example: $X_1, X_2, X_3 \sim \text{iid } N(0, 1)$

What is the distribution of $Y_1 = X_1^2 + X_2^2$?

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Sum of squares of two indep. std. normals $\Rightarrow Y_1 \sim \chi^2(2)$

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Sum of squares of two indep. std. normals $\Rightarrow Y_1 \sim \chi^2(2)$

What is the distribution of $Y_2 = (Y_1/2)/(X_3^2)$?

Example: $X_1, X_2, X_3 \sim \text{iid } N(0, 1)$

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Sum of squares of two indep. std. normals $\Rightarrow Y_1 \sim \chi^2(2)$

What is the distribution of $Y_2 = (Y_1/2)/(X_3^2)$?

$Y_1 \sim \chi^2(2)$ and $X_3^2 \sim \chi^2(1)$

Example: $X_1, X_2, X_3 \sim \text{iid } N(0, 1)$

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$Y_1 \sim \chi^2(2)$ and $X_3^2 \sim \chi^2(1)$

Hence $Y_2 = \text{ratio of two indep. } \chi^2 \text{ RVs, each divided by its degrees of freedom} \Rightarrow Y_2 \sim F(2, 1)$

Example: $X_1, X_2, X_3 \sim \text{iid } N(0, 1)$

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What is the distribution of $Z = X_3/\sqrt{Y_1/2}$?

Example: $X_1, X_2, X_3 \sim \text{iid } N(0, 1)$

What is the distribution of $Y_1 = X_1^2 + X_2^2$?

Sum of squares of two indep. std. normals $\Rightarrow Y_1 \sim \chi^2(2)$

What is the distribution of $Y_2 = (Y_1/2)/(X_3^2)$?

$Y_1 \sim \chi^2(2)$ and $X_3^2 \sim \chi^2(1)$

Hence $Y_2 =$ ratio of two indep. χ^2 RVs, each divided by its degrees of freedom $\Rightarrow Y_2 \sim F(2, 1)$

What is the distribution of $Z = X_3/\sqrt{Y_1/2}$?

Ratio of standard normal and square root of independent χ^2 RV divided by its degrees of freedom $\Rightarrow Z \sim t(2)$

Suppose $X_1, X_2, \sim \text{iid } N(\mu, \sigma^2)$



Let $Y = (X_1 - \mu)^2 + (X_2 - \mu)^2$. What is the distribution of Y/σ^2 ?

- (a) $F(2, 1)$
- (b) $\chi^2(2)$
- (c) $t(2)$
- (d) $N(\mu, \sigma)$
- (e) None of the above

$$Y_1 \sim \chi^2(2), \quad Y_2 \sim F(2, 1), \quad Z \sim t(2)$$

What is the median of Y_1 ?

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`qchisq(0.5, df = 2)` ≈ 1.4

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What is the median of Y_1 ?

`qchisq(0.5, df = 2)` ≈ 1.4

What is $P(Y_2 \leq 5)$?

$$Y_1 \sim \chi^2(2), \quad Y_2 \sim F(2, 1), \quad Z \sim t(2)$$

What is the median of Y_1 ?

`qchisq(0.5, df = 2)` ≈ 1.4

What is $P(Y_2 \leq 5)$?

`pf(5, df1 = 2, df2 = 1)` ≈ 0.7

$$Y_1 \sim \chi^2(2), \quad Y_2 \sim F(2, 1), \quad Z \sim t(2)$$

What is the median of Y_1 ?

`qchisq(0.5, df = 2)` ≈ 1.4

What is $P(Y_2 \leq 5)$?

`pf(5, df1 = 2, df2 = 1)` ≈ 0.7

What value of c gives $P(-c \leq Z \leq c) = 0.5$?

Use Symmetry (like normal)

`c = qt(0.75, df = 2)` ≈ 0.8

$$Y_1 \sim \chi^2(2), \quad Y_2 \sim F(2, 1), \quad Z \sim t(2)$$

What is the median of Y_1 ?

$$\text{qchisq}(0.5, \text{df} = 2) \approx 1.4$$

What is $P(Y_2 \leq 5)$?

$$\text{pf}(5, \text{df1} = 2, \text{df2} = 1) \approx 0.7$$

What value of c gives $P(-c \leq Z \leq c) = 0.5$?

Use Symmetry (like normal)

$$c = \text{qt}(0.75, \text{df} = 2) \approx 0.8$$

$$\text{or equivalently } -c = \text{qt}(0.25, \text{df} = 2) \approx -0.8$$