

Taylor Series

$\frac{1}{x} = \sum_{n=0}^{\infty} x^n$
 $e^x = \sum_{n=0}^{\infty} \frac{x^n}{n!}$

$\cos x = \sum_{n=0}^{\infty} (-1)^n \frac{x^{2n}}{(2n)!}$

$\sin x = \sum_{n=0}^{\infty} (-1)^n \frac{x^{2n+1}}{(2n+1)!}$

$\ln(1+x) = \sum_{n=1}^{\infty} (-1)^{n+1} \frac{x^n}{n}$

Permutations and Combinations

$P(n, k) = \frac{n!}{(n-k)!}$

$C(n, k) = \binom{n}{k} = \frac{n!}{(n-k)!k!}$

Laplace Transforms

$F^*(s) = \int_0^{\infty} f(t)e^{-st}dt$

$f(t) = \int_0^{\infty} F^*(s)e^{st}ds$

Convolution Property

$f(t) * g(t) = \int_0^t f(t-x)g(x)dx \leftrightarrow F^*(s)G^*(s)$

Z-Transform

Mapping of discrete function f_n into complex fuction with variable z .
 $F(z) = \sum_{n=0}^{\infty} f_n z^n$

Probability and Conditional

$P(A \cup B) = P(A) + P(B) - P(A \cap B)$

$P(A|B) = \frac{P(A \cap B)}{P(B)}$

A, B are independent if $P(A, B) = P(A)P(B)$

Total Probability

$P(B) = \sum_i P(A_i)P(B|A_i)$

Bayes’ Rule

$P(A_i|B) = \frac{P(A_i \cap B)}{P(B)} = \frac{P(A_i)P(B|A_i)}{\sum_j P(A_j)P(B|A_j)}$

PMF (Probability Mass Function)

$p_X(x) = p(\{s \in \Omega \text{ s.t. } X(s) = x\})$

$\sum_x p_X(x) = 1$

Bernoulli Random Variable

$X = 1$ on success, $X = 0$ on failure.

$p(X = x) = p$, if $x = 1$

$p(X = x) = 1 - p$, if $x = 0$

Geometric Random Variable

Counts #trials until first success.

$p_X(x) = (1 - p)^{x-1}p$, $x = 1, 2, \dots$

$p(X \geq s + 1|X \geq t) = p(X \geq s)$

Binomial Random Variable

Counts #success in n identical independent experiments.

$p_X(x) = \binom{n}{x}p^x(1 - p)^{n-x}$, when $0 \leq x \leq n$

$p_X(x) = 0$, otherwise

Poisson Random Variable

Model occurrence of event over time interval assuming event happens at rate λ

$p_X(x) = e^{-\lambda} \frac{\lambda^x}{x!}$, when $x = 0, 1, \dots$

PDF (Probability Density Function)

$\int_{-\infty}^{\infty} f_X(x)dx = 1$

CDF (Cumulative Distribution Function)

$F_X(x) = P(X \leq x)$

$\lim_{x \rightarrow -\infty} F_X(x) = 0$

$\lim_{x \rightarrow \infty} F_X(x) = 1$

$P(a < X \leq b) = F_X(b) - F_X(a)$

Uniform Distribution

$f_X(x) = \frac{1}{b-a}$, when $a \leq x \leq b$

$f_X(x) = 0$, otherwise

Exponential Distribution

Memoryless continuous distribution.

$f_X(x) = \lambda e^{-\lambda x}$, when $x \geq 0$

$F_X(x) = 1 - e^{-\lambda x}$, when $x \geq 0$

$F_X(x) = 0$, otherwise

$P(X > x) = e^{-\lambda x}$

Expectation

$E[X] = \sum_x xp(x)$

$E[X] = \int_{-\infty}^{\infty} xf_X(x)dx$

If $Y = g(X)$, $E[Y] = \sum_x g(x)p(x)$,

$E[Y] = \int_{-\infty}^{\infty} g(x)f_X(x)dx$

$E[X + Y] = E[X] + E[Y]$

$E[aX] = aE[X]$

$E[XY] = E[X]E[Y]$, if X, Y are independent.

For X, Y with joint PMF $p(x, y)$ or PDF

$f_{X,Y}(x, y)$, $E[XY] = \sum_{(x,y)} xy p(x, y)$,

$E[XY] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f_{X,Y}(x, y) dx dy$

Conditional Expectation

X, Y are random variables,

$E[Y|X] = \sum_y y P(Y = y|X = x) = \sum_y y p_{Y|X}(y|x)$,

$E[Y|X] = \int_{-\infty}^{\infty} y f_{Y|X}(y|x) dy$

Unconditional Expectation

$E[Y] = \sum_x E[Y|X]p_X(x)$

$E[Y] = \int_{-\infty}^{\infty} E[Y|X]f_X(x)dx$

Variance

$Var[X] = E[(X - E[X])^2] = \sum_x (x - E[X])^2 p(x) = \int_{-\infty}^{\infty} (x - E[X])^2 f_X(x) dx$

$Var[X] = E[X^2] - E[X]^2$

$Var[X + Y] = Var[X] + Var[Y]$, if X, Y are independent.

Expectations and Variances

Binomial: $np, np(1 - p)$

Geometric: $\frac{1}{p}, \frac{1-p}{p^2}$

Uniform: $\frac{a+b}{2}, \frac{(b-a)^2}{12}$

Exponential: $\frac{1}{\lambda}, \frac{1}{\lambda^2}$

Poisson: λ, λ

Covariance: measure of joint probability

$Cov(X, Y) = E[(X - E[X])(Y - E[Y])]$

$Cov(X, Y) = E[XY] - E[X]E[Y]$

If X, Y are independent, $Cov(X, Y) = 0$

Correlation: scaled version of covariance

$\rho(X, Y) = \frac{Cov(X, Y)}{\sqrt{Var(X)Var(Y)}}$, range $[-1, 1]$

Q1: A given program has an execution time that is uniformly distributed between 10 and 20 seconds.

The number of interrupts that occur during execution is a Poisson random variable with parameter t where t is the program execution time. The probability distribution of the number of interrupts is therefore $P(N = k) = (\lambda t)ke^{\lambda t}$.

(a) What is $E[N|T = t]$, where N is the number of interrupts the program experiences, and T is the running time of the program.

$E[N|T = t] = \lambda t$, since for fixed running time, the number of interrupts is a Poisson random variable with mean λt .

(b) Find the expected number of interrupts the program experiences during a randomly selected execution.

$E[N] = \int_{10}^{20} E[N|T = t]f_T(t)dt = \int_{10}^{20} \frac{\lambda t}{10} dt = 15\lambda$

Q2: Suppose that you made a webpage and you are collecting the statistics from the visitors. There are m types of visitors. Each visit is equally likely to be any of the m types. Find the expected number of visitors needed in order to have at least one of each type. Hint: Let X denote the number of visitors needed. It is useful to represent X by $X = \sum_{i=1}^m X_i$ where each X_i is a geometric random variable.

Suppose the current visitor pool contains i different types. Let X_i denote the number of additional visitors needed until it contains $i + 1$ types. The X_i is are independent geometric random variables with parameter $(m - i)/m$, $i = 0, 1, \dots, m - 1$.
 $E[X] = E[\sum_{i=1}^m X_i] = \sum_{i=1}^m E[X_i] = \sum_{i=1}^m \frac{m}{m-i}$

Q3: A Markov chain $\{X_n, n \geq 0\}$ with states 0, 1, 2, has the transition probability matrix

$\begin{bmatrix} \frac{1}{2} & \frac{1}{3} & \frac{1}{6} \\ 0 & \frac{1}{3} & \frac{2}{3} \\ \frac{1}{2} & 0 & \frac{1}{2} \end{bmatrix}$ If $P[X_0 = 0] = P[X_0 = 1] = 1$, find

the state probability vector $P[X_3 = 2]$.

Cubing the transition probability matrix, we obtain

$P^3 = \begin{bmatrix} \frac{13}{36} & \frac{11}{54} & \frac{47}{108} \\ \frac{9}{12} & \frac{27}{9} & \frac{13}{36} \end{bmatrix}$

$P[X_3 = 2] = \frac{1}{4} \cdot \frac{47}{108} + \frac{1}{4} \cdot \frac{11}{27} + \frac{1}{2} \cdot \frac{13}{36}$

Q4: A workstation tries to transmit frames through Ethernet. Suppose that whether or not collision occurs in the current transmission depends on the result of the last two trans- missions the workstation had. That is, suppose that if collisions have occurred in both of the past two transmissions, then with probability 0.7 a collision will occur in the current transmission; if a collision occurs in last transmission but not the transmission before the last one, then a collision will occur in the current transmission with probability 0.5; if a collision occurred in the transmission before the last one but not the last one, then one will occur in the current transmission with probability 0.4; if there have been no collision in the past two transmissions, then a collision will occur in the current transmission with probability 0.2. (Hint: Note that the state description needs to include status of last two transmissions).

(b) Find the transition probability matrix.

$P = \begin{bmatrix} 0.7 & 0 & 0.3 & 0 \\ 0.5 & 0 & 0.5 & 0 \\ 0 & 0.4 & 0 & 0.6 \\ 0 & 0.2 & 0 & 0.8 \end{bmatrix}$

(c) What fraction of frames suffer a collision?

Solve $\pi = \pi P$ to obtain the stationary state probabilities. Then the fraction of frames suffering a collision is $\pi_0 + \pi_2$ (or $\pi_0 + \pi_1$).