

Taylor Series

1/x = sum\_{n=0}^inf x^n  
e^x = sum\_{n=0}^inf x^n/n!

cos x = sum\_{n=0}^inf (-1)^n x^{2n}/(2n)!

sin x = sum\_{n=0}^inf (-1)^n x^{2n+1}/(2n+1)!

ln(1 + x) = sum\_{n=1}^inf (-1)^{n+1} x^n/n

Permutations and Combinations

P(n, k) = n!/(n-k)!

C(n, k) = (n choose k) = n!/(n-k)!k!

Laplace Transforms

F\*(s) = integral\_0^inf f(t)e^{-st}dt

f(t) = integral\_0^inf F\*(s)e^{st}ds

Convolution Property

f(t) \* g(t) = integral\_0^t f(t-x)g(x)dx <-> F\*(s)G\*(s)

Z-Transform

Mapping of discrete function f\_n into complex fuction with variable z.

F(z) = sum\_{n=0}^inf f\_n z^n

Probability and Conditional

P(A union B) = P(A) + P(B) - P(A intersection B)

P(A|B) = P(A intersection B)/P(B)

A, B are independent if P(A, B) = P(A)P(B)

Total Probability

P(B) = sum\_i P(A\_i)P(B|A\_i)

Bayes' Rule

P(A\_i|B) = P(A\_i intersection B)/P(B) = P(A\_i)P(B|A\_i)/sum\_j P(A\_j)P(B|A\_j)

PMF (Probability Mass Function)

p\_X(x) = p({s in Omega s.t. X(s) = x})

sum\_x p\_X(x) = 1

Bernoulli Random Variable

X = 1 on success, X = 0 on failure.

p(X = x) = p, if x = 1

p(X = x) = 1 - p, if x = 0

Geometric Random Variable

Counts #trials until first success.

p\_X(x) = (1 - p)^{x-1} p, x = 1, 2, ...

p(X >= s + 1 | X >= t) = p(X >= s)

Binomial Random Variable

Counts #success in n identical independent experiments.

p\_X(x) = (n choose x) p^x (1 - p)^{n-x}, when 0 <= x <= n

p\_X(x) = 0, otherwise

Poisson Random Variable

Model occurrence of event over time interval assuming event happens at rate lambda

p\_X(x) = e^{-lambda} lambda^x / x!, when x = 0, 1, ...

PDF (Probability Density Function)

integral\_{-inf}^inf f\_X(x)dx = 1

CDF (Cumulative Distribution Function)

F\_X(x) = P(X <= x)

lim\_{x -> -inf} F\_X(x) = 0

lim\_{x -> inf} F\_X(x) = 1

P(a < X <= b) = F\_X(b) - F\_X(a)

Uniform Distribution

f\_X(x) = 1/(b-a), when a <= x <= b

f\_X(x) = 0, otherwise

Exponential Distribution

Memoryless continuous distribution.

f\_X(x) = lambda e^{-lambda x}, when x >= 0

F\_X(x) = 1 - e^{-lambda x}, when x >= 0

F\_X(x) = 0, otherwise

P(X > x) = e^{-lambda x}

Expectation

E[X] = sum\_x x p(x)

E[X] = integral\_{-inf}^inf x f\_X(x) dx

If Y = g(X), E[Y] = sum\_x g(x) p(x),

E[Y] = integral\_{-inf}^inf g(x) f\_X(x) dx

E[X + Y] = E[X] + E[Y]

E[aX] = a E[X]

E[XY] = E[X]E[Y], if X, Y are independent.

For X, Y with joint PMF p(x, y) or PDF

f\_{X,Y}(x, y), E[XY] = sum\_{(x,y)} xy p(x, y),

E[XY] = integral\_{-inf}^inf integral\_{-inf}^inf xy f\_{X,Y}(x, y) dx dy

Conditional Expectation

X, Y are random variables,

E[Y|X] = sum\_y y P(Y = y|X = x) = sum\_y y p\_{Y|X}(y|x),

E[Y|X] = integral\_{-inf}^inf y f\_{Y|X}(y|x) dy

Unconditional Expectation

E[Y] = sum\_x E[Y|X] p\_X(x)

E[Y] = integral\_{-inf}^inf E[Y|X] f\_X(x) dx

Variance

Var[X] = E[(X - E[X])^2] = sum\_x (x - E[X])^2 p(x) =

integral\_{-inf}^inf (x - E[X])^2 f\_X(x) dx

Var[X] = E[X^2] - E[X]^2

Var[X + Y] = Var[X] + Var[Y] + 2Cov(X, Y)

Var[XY] = E[X^2 Y^2] - E^2[XY] = E[X^2]E[Y^2] +

Cov(X^2, Y^2) - (E[X]E[Y] + Cov(X, Y))^2

Expectations and Variances

Binomial: np, np(1 - p)

Geometric: 1/p, 1/p^2

Uniform: (a+b)/2, ((b-a)^2)/12

Exponential: 1/lambda, 1/lambda^2

Poisson: lambda, lambda

Covariance: measure of joint probability

Cov(X, Y) = E[(X - E[X])(Y - E[Y])]

Cov(X, Y) = E[XY] - E[X]E[Y]

If X, Y are independent, Cov(X, Y) = 0

Correlation: scaled version of covariance

rho(X, Y) = Cov(X, Y) / sqrt(Var(X)Var(Y)), range [-1, 1]

Stationary Process: F\_X(x; t) = F\_X(x; t + tau)

Independent Processes:

F\_X(x; t) = F\_{X\_1}(x\_1, t\_1) F\_{X\_2}(x\_2, t\_2) ... F\_{X\_n}(x\_n, t\_n)

f\_X(x, t) = product\_{i=1}^n f\_{X\_i}(x\_i, t\_i) (Continuous State)

p\_X(x, t) = product\_{i=1}^n p\_{X\_i}(x\_i, t\_i) (Discrete State)

Markovian Property: P[X(t\_{n+1}) <=

x\_{n+1} | X(t\_n) = x\_n, X(t\_{n+2}) = x\_{n+2}, ... , X(t\_0) =

x\_0] = P[X(t\_{n+1}) <= x\_{n+1} | X(t\_n) = x\_n]

Discrete Time Markov Chains (DTMC):

p\_{ij} = P[X\_n = j | X\_{n-1} = i] (Homogenous)

P = (p\_{ij}) = [ p00 p01 ... ; p10 p11 ... ; : : : ]

sum\_j p\_{ij} = 1 for each row.

Initial State Probabilities:

pi^{(0)} = (pi\_0^{(0)}, pi\_1^{(0)}, ...), where pi\_j^{(0)} = P[X\_0 = j]

n-Step Transition Probabilities:

p\_{ij}^{(n)} = P[X\_n = j | X\_0 = i] = P[X\_{n+k} = j | X\_k = i]

Chapman-Kolmogorov: p\_{ij}^{(n)} = sum\_k p\_{ik}^{(s)} p\_{kj}^{(t)}

Limiting Distribution: pi = lim\_{n -> inf} pi^{(0)} P^n

pi = pi P and sum\_j pi\_j = 1

Continuous Time Markov Chains (CTMC):

State transitions permitted at arbitrary time instances. Time spent in a state is exponentially distributed.

State Transition Probability:

p\_{ij}(t) = p(X(tau + t) = j | X(tau) = i)

Chapman-Kolmogorov Equation:

p\_{ij}(s + t) = sum\_k p\_{ik}(s) p\_{kj}(t)

Transition Probability: H(t) = {p\_{ij}(t)}

H(s + t) = H(s)H(t)

H(t + Delta t) = H(t)H(Delta t)

H(t + Delta t) = H(t)[H(Delta t) - I]

dH(t)/dt = H(t) lim\_{Delta t -> 0} [H(Delta t) - I / Delta t]

Q = lim\_{Delta t -> 0} [H(Delta t) - I / Delta t]

dH(t)/dt = H(t)Q

Transition Rate Matrix: Q, infinitesimal

generator of H(t). H(t) = e^{Qt}.

Diagonal Elements <= 0.

q\_{ii} = lim\_{Delta t -> 0} [p\_{ii}(Delta t) - 1] / Delta t.

Off-Diagonal Elements >= 0.

q\_{ij} = lim\_{Delta t -> 0} [p\_{ij}(Delta t) - 0] / Delta t, for i != j.

In each row, sum of off-diagonal = magnitude of diagonal: q\_{ii} = -sum\_{i != j} q\_{ij}

State Probabilities:

pi(t) = pi(0)H(t), pi(t) = pi(0)e^{Qt}

Stationary Distribution: pi Q = 0, sum\_j pi\_j = 1

Birth-Death Process: At state k, lambda\_k, mu\_k are birth and death rates. Transition Matrix: Q =

[ -lambda\_0 lambda\_0 0 0 ... ; mu\_1 -(lambda\_1 + mu\_1) lambda\_1 0 ... ; 0 mu\_2 -(lambda\_2 + mu\_2) lambda\_2 0 ... ; ... ... ... ... ... ]

Equilibrium Solution: pi Q = 0 and sum\_j pi\_j = 1

Differential Difference Equations: d pi\_k(t) / dt =

lambda\_{k-1} pi\_{k-1}(t) + mu\_{k+1} pi\_{k+1}(t) - (lambda\_k + mu\_k) pi\_k(t)

d pi\_0(t) / dt = mu\_1 pi\_1(t) - lambda\_0 pi\_0(t)

d pi\_k(t) / dt = flow in - flow out

flow in = lambda\_{k-1} pi\_{k-1} + mu\_{k+1} pi\_{k+1}

flow out = (lambda\_k + mu\_k) pi\_k

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Q1: Program has execution time uniformly distributed between 10 & 20 seconds. Number of interrupts during execution is Poisson random variable with parameter lambda t where t is program execution time. Probability distribution of the number of interrupts is P(N = k) = (lambda t)^k e^{-lambda t}.

(a) What is E[N | T = t], where N is the number of interrupts, and T is running time.

E[N | T = t] = lambda t, since for fixed running time, the number of interrupts is a Poisson random variable with mean lambda t.

(b) Expected number interrupts during randomly selected execution.

E[N] = integral\_{10}^{20} E[N | T = t] f\_T(t) dt = integral\_{10}^{20} lambda t f\_T(t) dt = 15 lambda

Q2: We have m types of visitors with equal probability. Find expected visitors to have one of each type. Hint: Let X denote number of visitors needed. Represent X by X = sum\_{i=1}^m X\_i where each X\_i is a geometric random variable.

Suppose the current visitor pool contains i different types. Let X\_i denote the number of additional visitors needed until it contains i + 1 types. The X\_i is are independent geometric random variables with parameter (m - i) / m, i = 0, 1, ... , m - 1.

E[X] = E[sum\_{i=1}^m X\_i] = sum\_{i=1}^m E[X\_i] = sum\_{i=1}^m m / (m - i)

Q3: A Markov chain {X\_n, n >= 0} with states 0, 1, 2 has transition probability matrix

[ 1/2 1/3 1/6 ; 0 1/3 2/3 ; 1/2 0 1/2 ] If P[X\_0 = 0] = P[X\_0 = 1] = 1/4,

find state probability vector P[X\_3 = 2].

Cubing transition probability matrix

P^{(3)} = [ 13/36 11/54 47/108 ; 4/9 4/27 11/27 ; 5/12 2/9 13/36 ]

P[X\_3 = 2] = 1/4 \* 47/108 + 1/4 \* 11/27 + 1/2 \* 13/36

Q4: Whether or not collision occurs depends on result of last two transmissions If collisions occurred in both of the past two, collision will occur with probability 0.7; collision occurs in last but not the transmission before the last one, then a collision will occur with probability 0.5; collision occurred in transmission before the last one but not the last one, collision with probability 0.4; no collision in the past two transmissions, collision with probability 0.2.

(b) Find the transition probability matrix.

P = [ 0.7 0 0.3 0 ; 0.5 0 0.5 0 ; 0 0.4 0 0.6 ; 0 0.2 0 0.8 ]

(c) What fraction of frames suffer a collision?

Solve  $\pi = \pi P$  to obtain the stationary state probabilities. Then the fraction of frames suffering a collision is  $\pi_0 + \pi_2$  (or  $\pi_0 + \pi_1$ ).

**Q5:** A shop has room for 2 customers. Customers arrive at Poisson rate 3/hour and service times are exponential random variables with mean 0.25 hours.

(a) Average number of customers in shop?

We have birth-death process with  $\lambda = 3$  and  $\mu = 4$ .

$$E[k] = \sum_{k=0}^2 k\pi_k. \text{ Solve for } \pi_k, \pi_1 = \frac{\lambda}{\mu}\pi_0,$$

$$\pi_2 = \frac{\lambda}{\mu}\pi_1, \sum_{j=0}^2 \pi_j = 1, \text{ so } \pi_0 = 16/37,$$

$$E[k] = 30/37.$$

(b) What is proportion of customers who get

served?  $\pi_0 + \pi_1 = 28/37$ .

**Q6:** Packets arrive at router Poisson rate 3/ms and

time to forward exponential with mean 0.2 ms.

Fraction of time buffer empty?

State is num packets in router,  $\lambda = 3, \mu = 5$ . Solve

$$p_1 = \frac{\lambda}{\mu}p_0, p_2 = \frac{\lambda}{\mu}p_1, \dots, p_{k+1} = \frac{\lambda}{\mu}p_k,$$

$$\sum_{k=0}^{\infty} p_k = 1, \text{ thus } p_0 = 2/5.$$

**Q7:** 2 machines produce products at  $n$

products/hour. Lifetime of machine follows

exponential with mean  $1/x$  hours, time to fix

machine follows exponential with mean  $1/y$  hours.

Expected long term producing rate?

State is num machines working.  $\lambda_0 = \lambda_1 = y,$

$\mu_1 = x, \mu_2 = 2x$  bc if either machine fails, num

working machines reduce from 2 to 1. Solve

$$p_1 = \frac{y}{x}p_0, p_2 = \frac{y}{2x}p_1, p_0 + p_1 + p_2 = 1, \text{ thus}$$

$$p_0 = \frac{2x^2}{2x^2+2xy+y^2}. \text{ Expected num products}$$

produced per hour:  $p_0 \times 0 + p_1 \times n + p_2 \times 2n$